## 1. Moment Generating Function

$$M_X(t) = E(e^{tx})$$

continuous:

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} P_x(x) dx$$

discrete:

$$M_X(t) = \sum e^{tx} P_x(x)$$

$$\frac{dM_x(t)}{dt} = E(\frac{de^{tx}}{dt}) = E(xe^{tx}) \Rightarrow M_x^{(n)}(t) = E(x^n e^{tx})$$

## Properties:

(a) 
$$M_x(t) = E(\sum_{n=0}^{\infty} \frac{x^n t^n}{n!})$$
 according to Taylor Series  $M_x'(t) = E(0 + x + x^2 t + \frac{x^3 t^2}{2!} + \dots) = E(\sum_{n=0}^{\infty} \frac{x^{n+1} t^n}{n!})$   $\Rightarrow M_x^{(k)}(t) = E(\sum_{n=0}^{\infty} \frac{x^{n+k} t^n}{n!})$ 

(b) 
$$M_x^{(n)}(0) = E(X^n)$$
  
 $\Rightarrow VAR[x] = E[x^2] - E^2[x] = M_x''(0) - [M_x'(0)]^2$ 

- (c)  $M_x(t) = M_v(t) \Rightarrow XY$  has same distribution
- (d)  $M_r(0) = 1$
- (e)  $M_{ax+b}(t) = M_x(at)e^{bt}$  Proof:  $M_{ax+b}(t) = E(e^{t(ax+b)}) = E(e^{axt}e^{bt}) = E(e^{axt})e^{bt} = M_x(at)e^{bt}$
- (f)  $M_{X+Y}(bt) = M_X(t)M_Y(t)$  (X, Y independent)

#### 2. Gama Function

$$\Gamma(n) = \int_0^\infty u^{n-1} e^{-u} du = (n-1)!$$

Proof:  
1. 
$$\int_0^\infty e^{-x} dx = -e^{-x} \Big|_0^\infty = 1$$

$$2. \ \Gamma(n) = -\int_0^\infty x^{n-1} de^{-x}$$

2. 
$$\Gamma(n) = -\int_0^\infty x^{n-1} de^{-x}$$
  

$$= -(x^{n-1}e^{-x}|_0^\infty - \int_0^\infty e^{-x}(n-1)x^{n-1} dx)$$

$$= 0 + \int_0^\infty e^{-x} (n-1) x^{n-2} dx$$

$$= (n-1) \int_0^\infty e^{-x} x^{n-2} dx$$
  
=  $(n-1)\Gamma(n-1)$ 

$$= (n-1)\tilde{\Gamma}(n-1)$$

Use Induction to prove the formula

#### 3. Basic probability property

## (a) Basic Properties

1. 
$$P(E) = \frac{n(E)}{n(S)}, \in [0, 1]$$
  
2.  $P(\phi) = 0$ 

$$P(\phi) = 0$$

3. 
$$A \subseteq B, P(A) \leq P(B)$$

4. 
$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

## (b) Conditional Probability and Bayles Theorem

1. 
$$P(A|B) = \frac{P(A \cup B)}{P(B)}$$

1. 
$$P(A|B) = \frac{P(A \cup B)}{P(B)}$$
  
2.  $P(A|B) = \frac{P(B|A)P(A)}{P(B)} = \frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|A')P(A')}$ 

# (c) PDF and CDF

PDF: probability distribution function  $\Rightarrow P_X(x)$ CDF: cumulative distribution function  $\Rightarrow F_X(x) =$  $\phi_X(x) = N_X(x) = P(X < x)$ 

(d) Expectation

1. 
$$E(X) = \sum x P_X(x)$$
 or  $\int_{-\infty}^{\infty} x P_X(x)$ 

2. E(c) = c

3. E(aX) = aE(X)

4. E(X + Y) = E(X) + E(Y)

5.  $E(X) = M'_X(0)$ 

(e) Variance and Standard Deviation

1. 
$$VAR(X) = E(X^2) - E^2(X) = \sum [(x - E(X))]P_X(x) = E[(x - \mu)^2] = \sigma^2$$

2. VAR(c) = 0

3.  $VAR(aX) = a^2 VAR(x)$ 

4. 
$$VAR(X \pm Y) = VAR(X) + VAR(Y) \pm 2COV(x,y)$$
,  $COV(x,y) = E(XY) - E(X)E(Y)$ 

(f) z-score

 $Z = \frac{x-\mu}{\sigma}$ , measures the distance of x from expected value in standard units.

#### 4. Common Discrete Distributions

## (a) Binomial Distribution

DEF: n time Bernoulli trials combined. probability of success and fail is (p, 1-p). Probability of success remains the same through the trails. X is the r.v of success times.

Note as  $X \sim B(n, p)$ 

$$P_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$M_X(t) = (1-p+pe^t)^n$$

$$E(X) = np, \ VAR(X) = np(1-p)$$

# (b) Hyper Geometric Distribution

DEF: A sample of size n taken from a finite population of size N. The population has a subgroup of size  $r \ge n$  that is of interest. x is the number of members of the subgroup taken.

Note as  $X \sim H(N, n, r)$ 

$$P_X(x) = \frac{\binom{r}{x}\binom{N-r}{n-x}}{\binom{N}{n}}$$
 
$$E(X) = n\frac{r}{N}, \ VAR(X) = n\frac{r}{N}(1-\frac{r}{N})(\frac{N-n}{N-1})$$

When  $x \to \infty$ ,  $H(N,n,r) \to B(n,\frac{r}{N})$ . H samples without replacement while B samples with replacement.

## (c) Poisson Process

DEF: model the number of random occurance of some phenomenon in specific unit of space or time.

$$P_X(x) = \frac{e^{-\lambda} \lambda^x}{x!}$$

 $\lambda$ : average arrival in given time or space

$$E(X) = \lambda$$
,  $VAR(X) = \lambda$ 

Poisson simulates Binomial when n is large. usually when  $\lambda = np < 10$ , we use poisson as an approximation of Binomial.

## (d) Geometric Distribution

DEF: number of trials to get the first success in a sequence of Bernoulli trials where p is the success probability.

i. X is the r.v of number of total trials (x includes the first success)

$$P_X(x) = (1-p)^{x-1}p, x = 1, 2, 3...$$
  
 $E(X) = 1/p, VAR(X) = \frac{1-p}{p^2}$ 

ii. X is the r.v of number of failed trials (x excludes the first success)

$$P_X(x) = (1-p)^x p, x = 0, 1, 2, \dots$$

$$M_X(t) = \frac{p}{1 - e^t (1-p)}$$

$$E(X) = \frac{1-p}{p}, \ VAR(X) = \frac{1-p}{p^2}$$

(e) Negative Binomial Distribution DEF: X is the r.v of number of trials need to observe the  $r^{th}$  success in a sequence of Bernoulli trails where p is the success probability.

$$P_X(x) = {x-1 \choose r-1} p^r (1-p)^{x-r}, x = r, r+1, r+2...$$

$$M_X(t) = (\frac{p}{1-e^t (1-p)})^r$$

$$E(X) = \frac{r}{p}, VAR(X) = \frac{r(1-p)}{p^2}$$

Alternatively, X is the r.v of failures before the  $r^{th}$ success:

$$P_X(x) = {x+r-1 \choose r-1} p^r (1-p)^x, x = 0, 1, 2 \dots$$

$$M_X(t) = (\frac{1-p}{1-pe^t})^r$$

$$E(X) = \frac{r(1-p)}{p}, VAR(X) = \frac{r(1-p)}{p^2}$$

5. Chebychev's Theorem

$$P(\mu - k\sigma \le x \le \mu + k\sigma) \ge 1 - \frac{1}{k^2}$$

- 6. Continuous Random Variable
  - (a) Basic Properties

1.  $pdf : f_X(x) \ge 0$ 

2.  $cdf: F_X(x) \leq 0$ 3.  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ 

4.  $f_X(x) = \frac{d}{dx} F_X(x)$ 

(b) Mean, Medium and Variance mean:  $\mu = E(x) = \int_{-\infty}^{\infty} x f_X(x) dx$ medium m: solve function  $\int_{-\infty}^{m} f_X(x) dx = \frac{1}{2}$  $VAR(X) = E(X^{2}) - E^{2}(X) = \int_{-\infty}^{\infty} [x - E(X)] f_{X}(x) dx$ 

7. Common Continuous Distributions

(a) Uniform(rectangle) Distribution

$$f_X(x) = \frac{1}{b-a}, a \le x \le b$$

$$F_X(x) = \frac{x-a}{b-a}, a \le x \le b$$

$$E(X) = \frac{a+b}{2}, VAR(X) = \frac{(b-a)^2}{12}$$

(b) Exponential Distribution

$$f_X(x) = \frac{1}{\theta} e^{-\frac{x}{\theta}}, x \ge 0$$

$$F_X(x) = 1 - e^{-\frac{x}{\theta}}, x \ge 0$$

$$M_X(t) = \frac{1}{1 - \theta t}$$

$$E(X) = \theta, VAR(X) = \theta^2$$

Note as  $X \sim exp(\theta)$ 

(c) Gamma Distribution

$$f_X(x) = \frac{1}{\Gamma(n)\theta^n} x^{n-1} e^{-\frac{x}{\theta}}, x \ge 0$$

$$M_X(t) = \frac{1}{(1 - \theta t)^n}$$

$$E(X) = n\theta, VAR(X) = n\theta^2$$

Note as  $X \sim \Gamma(n, \theta)$ 

Exponential dist. is a special case of Gamma dist where n = 1. Gamma Distribution can be viewed as a sum of Exponential dists.

$$X \sim \Gamma(n, \theta) \Leftrightarrow X = \sum_{i=1}^{n} X_i, X_i \sim exp(\theta)$$

(d) Normal Distribution Standard normal distribution:  $X \sim N(0,1), \mu = 0, \sigma^2 = 1$ 

Normal distribution:

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, x \in (-\infty, \infty)$$

$$M_X(t) = e^{\mu x + \frac{1}{2}\sigma^2 t^2}$$

$$E(x) = \mu, VAR(X) = \sigma^2$$

8. Central Limit Theorem

Let  $\{X_1, X_2, ..., X_n\}$  be the independent random variable all of which have the same distribution and mean  $\mu$  and standard deviation  $\sigma$ . If n is large  $n \ge 30$ , then

$$S = X_1 + X_2 + \dots + X_n$$

will be approximately normal with mean  $n\mu$  and variance  $n\sigma^2$ .  $(S \sim N(n\mu, n\sigma^2))$ 

OR

$$S' = \frac{X_1 + X_2 + \dots + X_n}{n}$$

will be approximately normal with mean  $\mu$ , variance  $\frac{\sigma^2}{n}$ . $(S \sim N(\mu, \frac{\sigma^2}{n}))$ 

9. Finding CDF for Y = g(X)

1. g(x) is strictly increasing on the sample space for X Let h(y) be the inverse function of g(x). h(x) is also strictly increasing.

$$F_Y(y) = P(Y \le y)$$

$$= P(Y \le y)$$

$$= P(g(X) \le y)$$

$$= P[X \le h(y)]$$

$$= F_X(h(y))$$

Find the density function by differentiating

$$f_Y(y) = \frac{d}{dy} F_X(h(y))$$

$$= \frac{d}{dx} F_X(h(y)) \frac{dy}{dx}$$

$$= \frac{d}{dx} F_X(h(y)) h'(y)$$

2. g(x) is strictly decreasing on the sample space for X Let h(y) be the inverse function of g(x). h(x) is also strictly decreasing.

$$F_Y(y) = P(Y \le y)$$

$$= P(Y \le y)$$

$$= P(g(X) \le y)$$

$$= P[X \ge h(y)]$$

$$= 1 - F_X(h(y))$$

Find the density function by differentiating

$$f_Y(y) = \frac{d}{dy} [1 - F_X(h(y))]$$
$$= -\frac{d}{dx} F_X(h(y)) \frac{dy}{dx}$$
$$= -\frac{d}{dx} F_X(h(y)) h'(y)$$

IN ALL:

Density function for Y = g(X)

Let g(x) be strictly decreasing or increasing on the domain consisting of the sample space. Then:

$$f_Y(y) = \frac{d}{dx} F_X(h(y)) \cdot |h'(y)|$$