

ConCR-TMLE R Paper

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Abstract

An abstract of less than 150 words.

Introduction

Data Structure

Consider a survival analysis on an interval $[0, t_{max}]$ with competing risks. Let T_j^a denote counterfactual time-to-event variables for event j and intervention a , for competing events $j \in \mathcal{J} = \{1, 2, \dots, J\}$ and an intervention $a \in \mathcal{A}$. Our counterfactual data structure can then be denoted by

$$(T_j^a, \mathbf{X} : a \in \mathcal{A}, j \in \mathcal{J})$$

where $\mathbf{X} \in \mathbb{R}^d$ is a d -dimensional vector of baseline covariates. For a single time-point binary intervention, as in many randomized control trials, $\mathcal{A} = \{0, 1\}$ and the corresponding counterfactual data is

$$(T_j^1, T_j^0, \mathbf{X} : j \in \mathcal{J})$$

```
## sample data.table for counterfactual data
```

Let O denote the corresponding coarsened observed data where $O \sim P_0$. The observed data would include the time-to-censoring C , and observed intervention A . The time to first event (censoring or otherwise) we denote as $\tilde{T} = \min_j (C, T_j : j \in \mathcal{J})$ with $\Delta = (\arg\min_j T_j) \times \mathbf{1}(\min_j T_j \leq C)$ marking which outcome is observed ($\Delta = 0$ being that censoring occurred). The observable right-censored survival data with competing events can then be represented as

$$O = (\tilde{T}, \Delta, A, \mathbf{X})$$

```
## sample data.table of the observed data correlating to the above
## counterfactual data example
```

This observed data also allows the “long-format” formulation, where with single time-point intervention variable A and baseline covariate vector \mathbf{X} ,

$$O = (N_j(t), N_c(t), A, \mathbf{X} : j \in \mathcal{J}, t \leq \tilde{T})$$

Here $N_j(t) = \mathbf{1}(\tilde{T} \leq t, \Delta = j)$ and $N_c(t) = \mathbf{1}(\tilde{T} \leq t, \Delta = 0)$ denote counting processes for event j and censoring respectively.

Under coarsening at random (CAR), the observed data likelihood can be factorized as

$$p(O) = p(\mathbf{X}) \pi(A | \mathbf{X}) \lambda_c(\tilde{T} | A, \mathbf{X})^{\mathbf{1}(\Delta=0)} S_c(\tilde{T}^- | A, \mathbf{X}) \prod_{j=1}^J S(\tilde{T}^- | A, \mathbf{X}) \lambda_j(\tilde{T} | A, \mathbf{X})^{\mathbf{1}(\Delta=j)}$$

where $\lambda_c(t | A, \mathbf{X})$ is the hazard of the censoring process and $\lambda_j(t | A, \mathbf{X})$ is the hazard of the j^{th} event process. Additionally

$$S_c(t | a, \mathbf{x}) = \exp \left(- \int_0^t \lambda_c(s | a, \mathbf{x}) ds \right)$$

while in a pure competing risks setting

$$S(t | a, \mathbf{x}) = \exp \left(- \int_0^t \sum_{j=1}^J \lambda_j(s | a, \mathbf{x}) ds \right)$$

and

$$\begin{aligned} F_j(t | a, \mathbf{x}) &= \int_0^t S(s- | a, \mathbf{x}) \lambda_j(s | a, \mathbf{x}) ds \\ &= \int_0^t \exp \left(- \int_0^s \sum_{j=1}^J \lambda_j(u | a, \mathbf{x}) du \right) \lambda_j(s | a, \mathbf{x}) ds. \end{aligned}$$

Target Parameter

For a target parameter of the treatment regime a^* , cause $k \in J$ cumulative risk at time τ

$$\Psi_{a^*, k, \tau}(P_0) = \mathbb{E}[F_k(\tau | A = a^*, \mathbf{X})]$$

the corresponding efficient influence function $D_{a^*, k, \tau}^*(P)(O)$ is

$$\begin{aligned} &\sum_{j=1}^J \int_0^\tau \frac{\mathbf{1}(A = a^*) \mathbf{1}(s \leq \tau)}{\pi(A | \mathbf{X}) S_c(s- | A, \mathbf{X})} \left(\mathbf{1}(\delta = k) - \frac{F_k(\tau | A, \mathbf{X}) - F_k(s | A, \mathbf{X})}{S(s | A, \mathbf{X})} \right) \\ &\quad \left(N_j(ds) - \mathbf{1}(\tilde{T} \geq s) \lambda_j(s | A, \mathbf{X}) \right) ds \\ &\quad + F_k(t | A = a^*, \mathbf{X}) - \Psi_{a^*, k, \tau}(P_0) \end{aligned}$$

with a clever covariate $h_{a^*, k, j, \tau, s}$

$$h_{a^*, k, j, \tau, s} = \frac{\mathbf{1}(A = a^*) \mathbf{1}(s \leq \tau)}{\pi(A | \mathbf{X}) S_c(s- | A, \mathbf{X})} \left(\mathbf{1}(\delta = k) - \frac{F_k(\tau | A, \mathbf{X}) - F_k(s | A, \mathbf{X})}{S(s | A, \mathbf{X})} \right)$$

The components of the data distribution that must be estimated are $g(A | \mathbf{X})$ and $S_c(t | A, \mathbf{X})$, $\lambda_j(t | A, \mathbf{X})$, $F_j(t | A, \mathbf{X})$, and $S(t | A, \mathbf{X})$

Estimation

Cross-Validation Specification

Let $D_n = \{O_i\}_{i=1}^n$ be an observed sample of n i.i.d observations of $O \sim P_0$. For V -fold cross validation, let $B_n = \{1, \dots, V\}^n$ be a random vector that assigns the n observations into V validation folds. For each $v \in \{1, \dots, V\}$ we then define training set $D_v^T = \{O_i : B_n(i) = v\}$ with the corresponding validation set $D_v^V = \{O_i : B_n(i) \neq v\}$.

Stratified Cross-Validation

```
# StrataIDs <- as.numeric(factor(paste0(Data[["Trt"]], ":", Data[["Event"]]])))
# CVFolds <- origami::make_folds(n = Data, fold_fun = origami::folds_vfold, strata_ids = StrataIDs)
```

Propensity Score Estimation

For the conditional distribution of A given \mathbf{X} , $\pi(\cdot | \mathbf{X})$, and $\hat{\pi} : D_n \rightarrow \hat{\pi}(D_n)$, let L_π be a loss function such that the risk $\mathbb{E}_0[L_\pi(\hat{\pi}, O)]$ is minimized when $\hat{\pi} = \pi_0$. For instance, with a binary A , we may specify the negative log loss $L_\pi(\hat{\pi}, O) = -\log(\hat{\pi}(1 | \mathbf{X})^A \hat{\pi}(0 | \mathbf{X})^{1-A})$. Let \mathcal{M}_π be the set of candidate propensity score models. The discrete superlearner selector chooses the candidate propensity score model with the minimal cross validated risk

$$\hat{\pi}^{SL} = \operatorname{argmin}_{\hat{\pi} \in \mathcal{M}_\pi} \sum_{v=1}^V P_{D_v^\mathcal{V}} L_\pi(\hat{\pi}(D_v^\mathcal{T}), D_v^\mathcal{V})$$

```
getInitialEstimates <- function(...) {
  getPropScores <- function(...) {
    getSl3PropScores <- function {
      TrtTask <- sl3::make_sl3_Task(
        data = Data[, -c("Time", "Event", "ID")],
        covariates = colnames(CovDataTable),
        outcome = "Trt")

      TrtSL <- sl3::Lrnsl$new(learners = Models[["A"]], folds =
        CVFolds)
      TrtFit <- TrtSL$train(TrtTask)
    }
  }
}
```