

### Section 1: Hospital Pricing and Selection on Observables

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# Background on Hospital Pricing

Defining characteristic of hospital services: it's complicated!

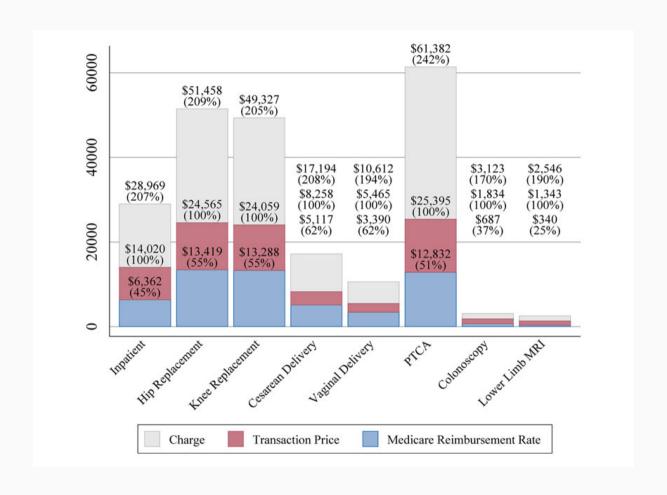
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Brill, Steven. 2013. "Bitter Pill: Why Medical Bills are Killing Us." \*Time Magazine\*.

Lots of different payers paying lots of different prices:

- Medicare fee-for-service prices
- Medicaid payments
- Private insurance negotiations (including Medicare Advantage)
- But what about the price to patients?

Price  $\neq$  charge  $\neq$  cost  $\neq$  patient out-of-pocket spending



Source: Health Care Pricing Project

Not clear what exactly is negotiated...

#### Fee-for-service

- price per procedure
- percentage of charges
- markup over Medicare rates

#### Capitation

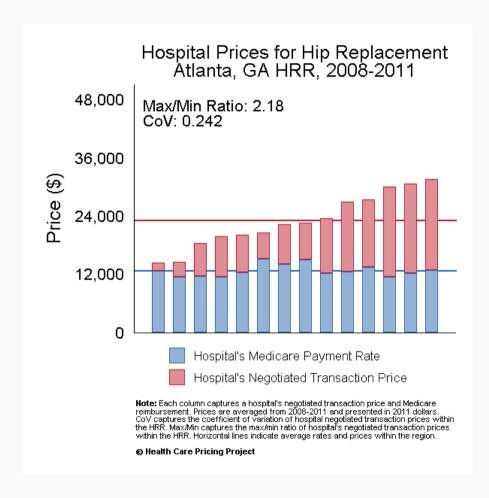
- payment per patient
- pay-for-performance
- shared savings

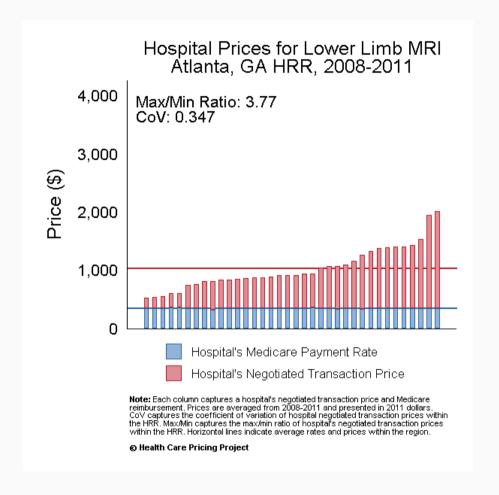
# Hospital prices in real life

#### A few empirical facts:

- 1. Hospital services are expensive
- 2. Prices vary dramatically across different areas
- 3. Lack of competition is a major reason for high prices

## Hospital prices in real life





Source: Health Care Pricing Project

# **Understanding HCRIS Data**

### What is HCRIS?

Healthcare Cost Report Information System ('cost reports')

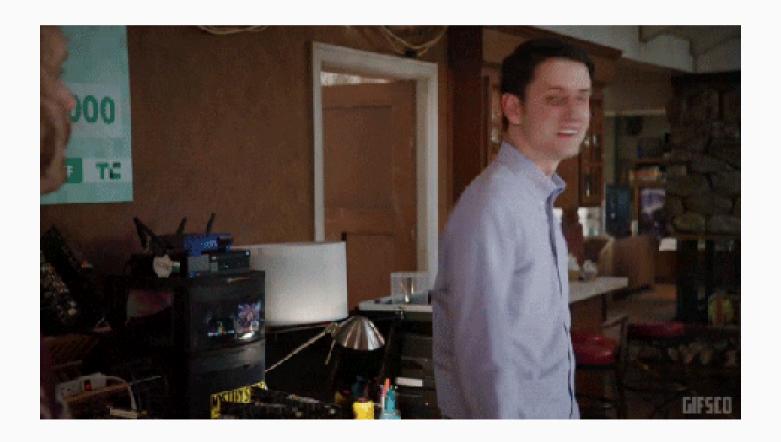
- Nursing Homes (SNFs)
- Hospice
- Home Health Agencies
- Hospitals

# **Hospital Cost Reports**

ART I -  GE 1 Ho 2 Su 3 Su 4 Su	PATIENT REVENUES  - PATIENT REVENUES  - PATIENT REVENUES  REVENUE CENTER  ENERAL INPATIENT ROUTINE CARE SERVICES  Jospital  ubprovider IPF  ubprovider IRF  ubprovider (Other)	PROVIDER CCN:  INPATIENT  1	PERIOD: FROM TO OUTPATIENT 2	WORKSHEET G-2, PARTS I & II  TOTAL 3	
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5 Sv					4
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6 Sv	wing bed - NF				6
7 Sk	killed nursing facility				7
8 N	Tursing facility				8
9 Ot	Other long term care				9
10 To	otal general inpatient care services (sum of lines 1-9)				10
IN	TENSIVE CARE TYPE INPATIENT HOSPITAL SERVICES				
11 In	ntensive care unit				11
12 Co	oronary care unit				12
13 Bt	Burn intensive care unit				13
14 Su	urgical intensive care unit				14
	Other special care (specify)				15
16 To	otal intensive care type inpatient hospital services (sum of				16
of	f lines 11-15)				
17 To	otal inpatient routine care services (sum of lines 10 and 16)				17
18 Aı	ancillary services				18
19 O	Outpatient services				19
20 Rt	tural Health Clinic (RHC)				20
21 Fe	ederally Qualified Health Center (FQHC)				21

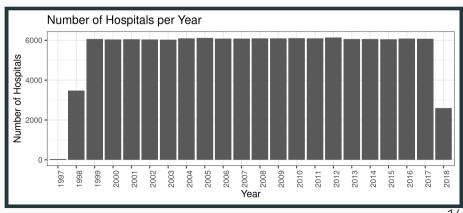
### The Data

Let's work with the HCRIS GitHub repository. But forming the dataset is up to you this time.

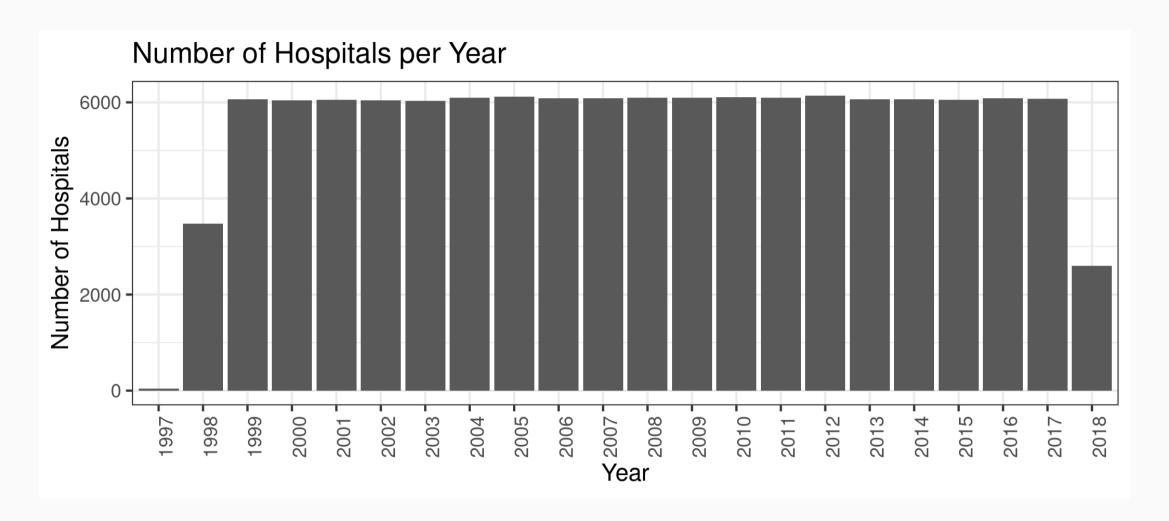


### The Data

```
hcris.data %>%
  ggplot(aes(x=as.factor(year))) +
  geom_bar() +
  labs(
    x="Year",
    y="Number of Hospitals",
    title="Number of Hospitals per Year"
  ) + theme_bw() +
  theme(axis.text.x = element_text(angle = 90, hjust=1))
```

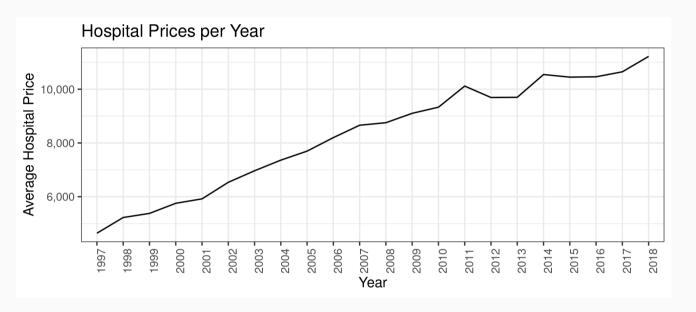


## Number of hospitals

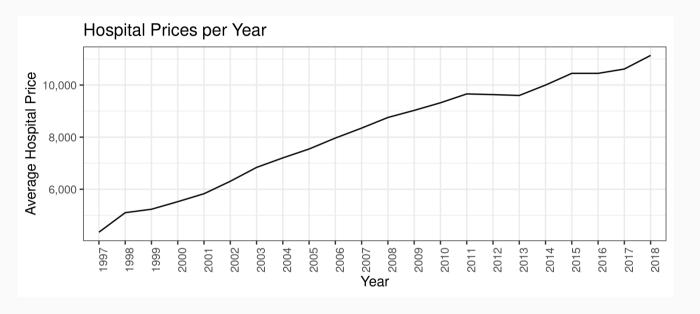


## Estimating hospital prices

## Estimating hospital prices



## Estimating hospital prices



### Causal Inference and Potential Outcomes

## Why causal inference?

Another example: What price should we charge for a night in a hotel?

#### **Machine Learning**

- Focuses on prediction
- High prices are strongly correlated with higher sales
- Increase prices to attract more people?

#### **Causal Inference**

- Focuses on counterfactuals
- What would sales look like if prices were higher?

### Goal of Causal Inference

- Goal: Estimate effect of some policy or program
- Key building block for causal inference is the idea of **potential outcomes**

### Some notation

#### **Observed outcome**

$$Y_i = Y_{1i} imes D_i + Y_{0i} imes (1-D_i)$$

or

$$Y_i = \left\{ egin{aligned} Y_{1i} ext{ if } D_i = 1 \ Y_{0i} ext{ if } D_i = 0 \end{aligned} 
ight.$$

Assumes **SUTVA** (stable unit treatment value assumption)...no interference across units

### Example of "Potential Outcomes"





$$Y_0$$
= \$60,000

Earnings due to Emory =  $Y_1 - Y_0$  = \$15,000

### Example of "Potential Outcomes"



Earnings due to Emory =  $Y_1 - Y_0$  = ?



$$Y_0$$
= ?

## Do we ever observe the potential outcomes?



Without a time machine...not possible to get individual effects.

### Fundamental Problem of Causal Inference

- We don't observe the counterfactual outcome...what would have happened if a treated unit was actually untreated.
- ALL attempts at causal inference represent some attempt at estimating the counterfactual outcome. We need an estimate for  $Y_0$  among those that were treated, and vice versa for  $Y_1$ .

# **Average Treatment Effects**

### Different treatment effects

Tend to focus on **averages**<sup>1</sup>:

• ATE: 
$$\delta_{ATE} = E[Y_1 - Y_0]$$

$$ullet$$
 ATT:  $\delta_{ATT}=E[Y_1-Y_0|D=1]$ 

• ATU: 
$$\delta_{ATU}=E[Y_1-Y_0|D=0]$$

<sup>&</sup>lt;sup>1</sup> or similar measures such as medians or quantiles

### **Average Treatment Effects**

#### • Estimand:

$$\delta_{ATE} = E[Y_1 - Y_0] = E[Y|D=1] - E[Y|D=0]$$

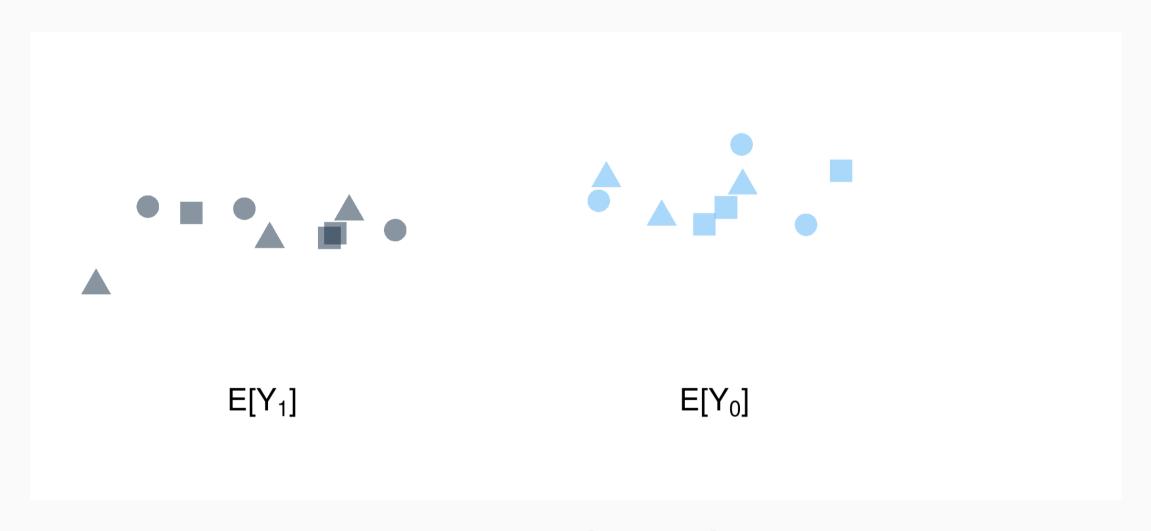
• Estimate:

$$\hat{\delta}_{ATE} = rac{1}{N_1} \sum_{D_i=1} Y_i - rac{1}{N_0} \sum_{D_i=0} Y_i,$$

where  $N_1$  is number of treated and  $N_0$  is number untreated (control)

• With random assignment and equal groups, inference/hypothesis testing with standard two-sample t-test

# ATE in pictures



$$ATE = E[Y_1 - Y_0]$$

# Selection Bias

### Selection bias

- The field of causal inference is all about different strategies to remove selection bias
- The first strategy (really, assumption) in this class: **selection on observables** or **conditional indpendence**

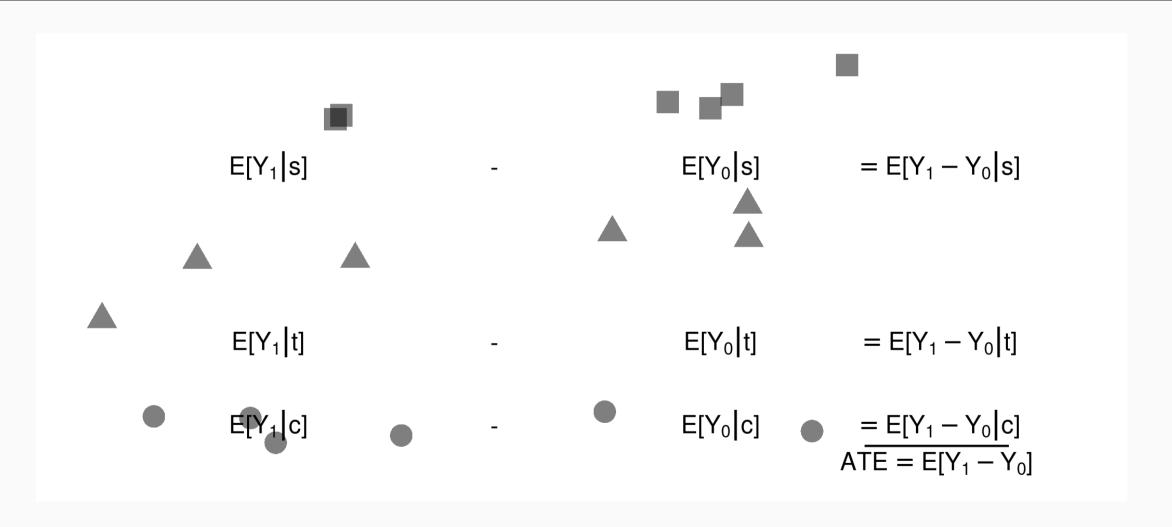
## Conditional independence

#### Some assumptions:

- $Y_{0i} = \alpha + \eta_i$
- $Y_{1i}-Y_{0i}=\delta$
- ullet There is some set of "controls",  $x_i$ , such that  $\eta_i=eta x_i+u_i$  and  $E[u_i|x_i]=0$  (conditional independence assumption, or CIA)

$$egin{aligned} Y_i &= Y_{1i} imes D_i + Y_{0i} imes (1 - D_i) \ &= \delta D_i + Y_{0i} D_i + Y_{0i} - Y_{0i} D_i \ &= \delta D_i + lpha + \eta_i \ &= \delta D_i + lpha + eta x_i + u_i \end{aligned}$$

# Selection in pictures



# Identifying causal effect

Find covariates  $X_i$  such that the following assumptions are plausible:

1. Selection on observables:

$$Y_{0i}, Y_{1i} \perp \!\!\! \perp D_i | X_i$$

2. Common support:

$$0<\Pr(D_i=1|X_i)<1$$

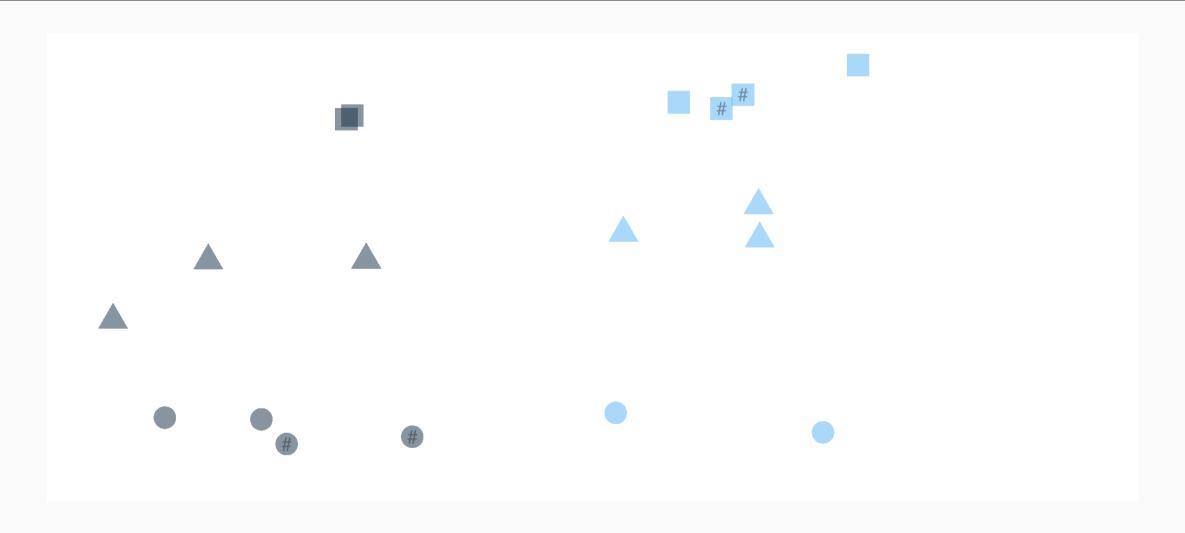
Then we can use  $X_i$  to group observations and use expectations for control as the predicted counterfactuals among treated, and vice versa.

## Assumption 1: Selection on Observables

$$E[Y_1|D,X] = E[Y_1|X]$$

In words...nothing unobserved that determines treatment selection and affects your outcome of interest.

#### Violation of Selection on Observables

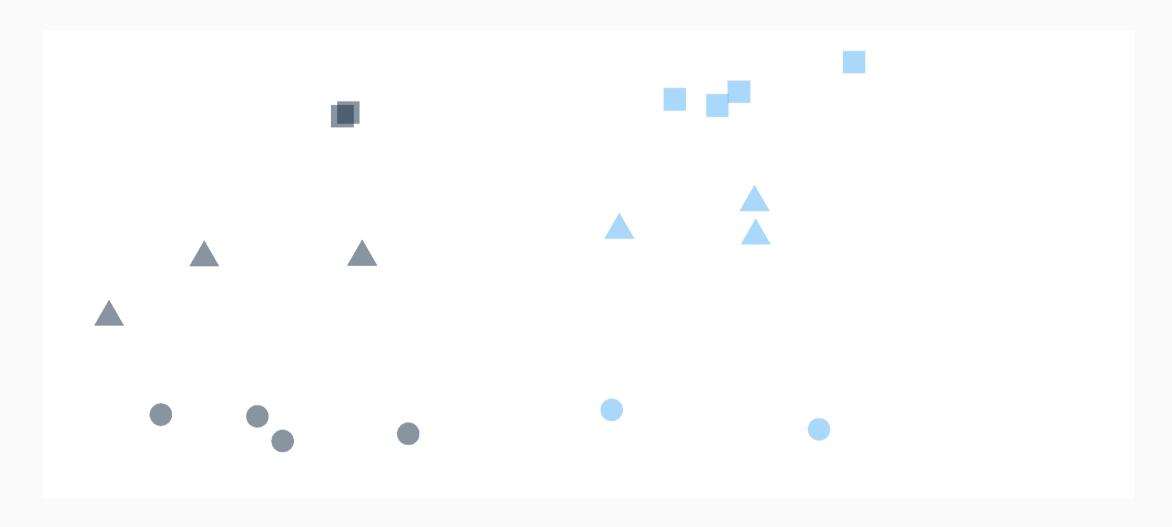


#### **Assumption 2: Common Support**

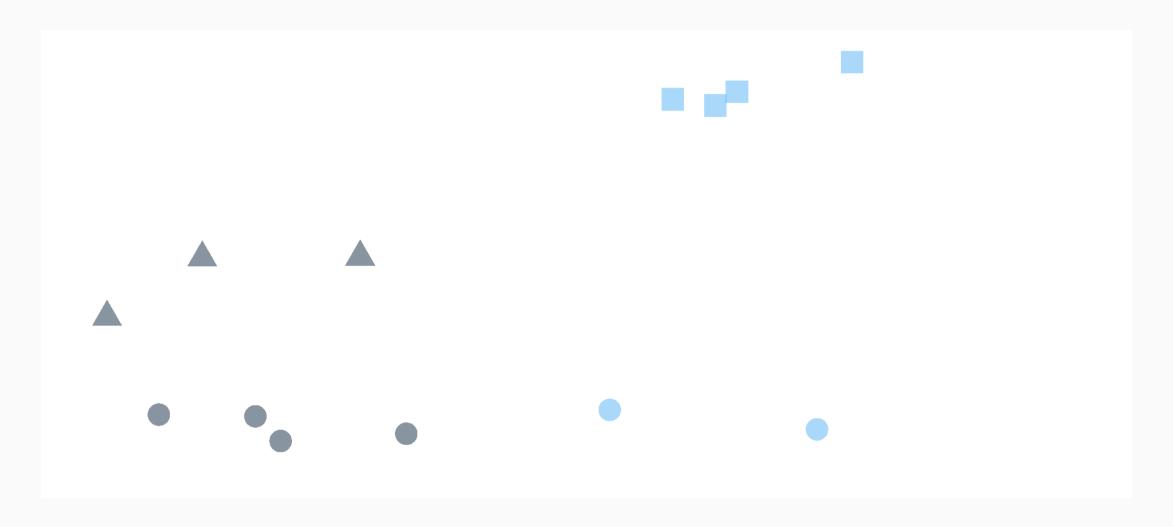
Someone of each type must be in both the treated and untreated groups

$$0<\Pr(D=1|X)<1$$

# Assumption 2: Common Support



# Violation of Common Support



## Role of selection in practice

• Look at Table 2.2 of Mastering Metrics

# Matchmaking

#### Causal inference with observational data

With selection on observables and common support:

1. Matching estimators:

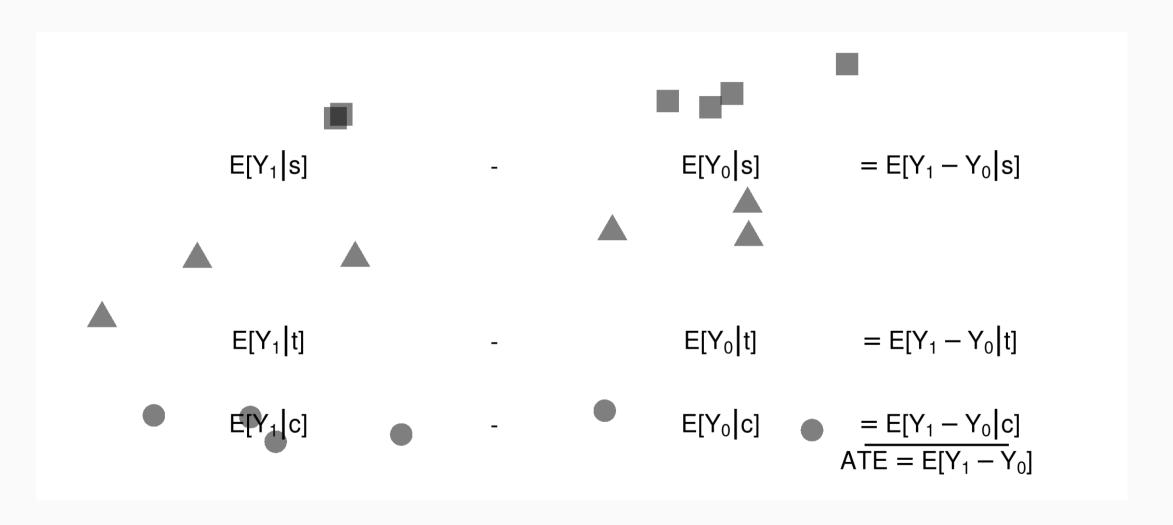
$$E[Y_i(1)-Y_i(0)]=E[E[Y_i|W_i=1,X_i]-E[Y_i|W_i=0,X_i]]$$

2. Regression estimators:

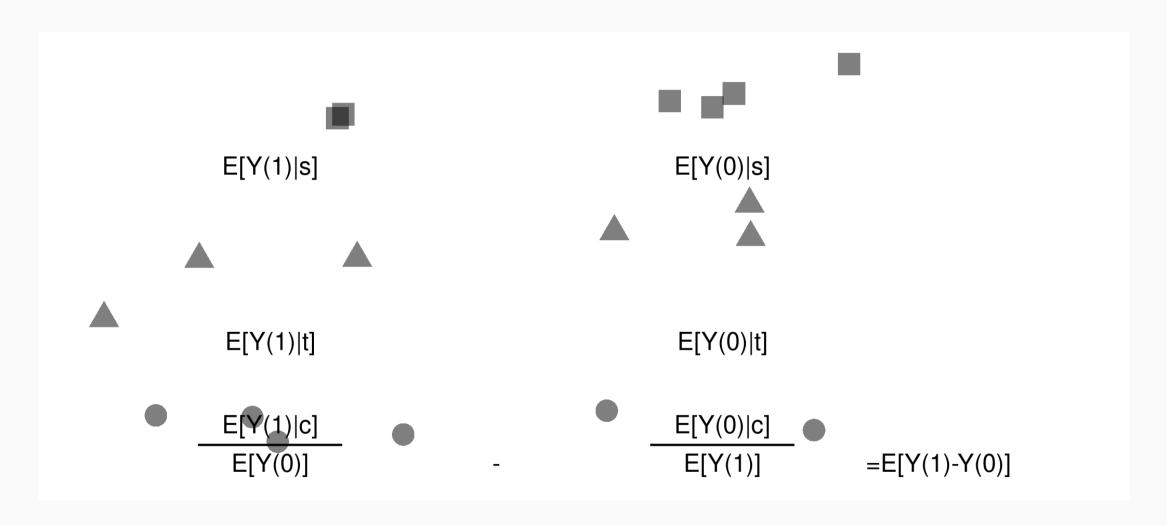
$$E[Y_i(1) - Y_i(0)] = E[E[Y_i|W_i = 1, X_i]] - E[E[Y_i|W_i = 0, X_i]]$$

What's the difference?

# Matching



# Regression



# **Estimation options**

- Matching
- Weighting
- Regression
- Doubly-robust weighting + regression (won't cover)

### Matching: The process

- 1. For each observation i, find the m "nearest" neighbors,  $J_m(i)$ .
- 2. Impute  $\hat{Y}_i(0)$  and  $\hat{Y}_i(1)$  for each observation:

$$\hat{Y}_i(0) = \left\{egin{array}{ll} Y_i & ext{if} & W_i = 0 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 1 \end{array}
ight.$$

$$\hat{Y}_i(1) = \left\{egin{array}{ll} Y_i & ext{if} & W_i = 1 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 0 \end{array}
ight.$$

3. Form "matched" ATE:

$$\hat{\delta}^{ ext{match}} = rac{1}{N} \sum_{i=1}^{N} \left( \hat{Y}_i(1) - \hat{Y}_i(0) 
ight)$$

# Matching: Defining "nearest"

1. Euclidean distance:

$$\sum_{k=1}^K (X_{ik}-X_{jk})^2$$

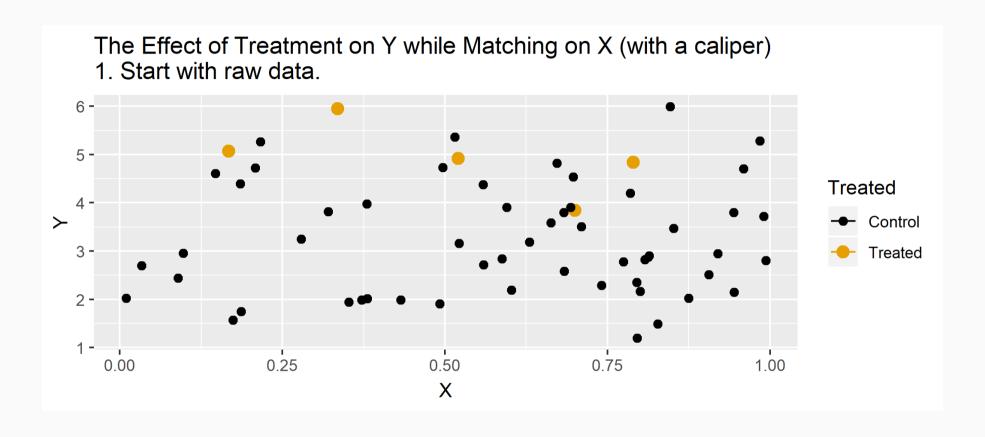
2. Scaled Euclidean distance:

$$\sum_{k=1}^K rac{1}{\sigma_{X_k}^2} (X_{ik} - X_{jk})^2$$

3. Mahalanobis distance:

$$(X_i-X_j)'\Sigma_X^{-1}(X_i-X_j)$$

# Animation for matching



## Weighting

- 1. Estimate propensity score ps  $\leftarrow$  glm(W~X, family=binomial, data), denoted  $\hat{\pi}(X_i)$
- 2. Weight by inverse of propensity score

$$\hat{\mu}_1 = rac{\sum_{i=1}^N rac{Y_i W_i}{\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{W_i}{\hat{\pi}(X_i)}}$$
 and  $\hat{\mu}_0 = rac{\sum_{i=1}^N rac{Y_i (1-W_i)}{1-\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{1-W_i}{1-\hat{\pi}(X_i)}}$ 

3. Form "inverse-propensity weighted" ATE:

$$\hat{\delta}^{IPW} = \hat{\mu}_1 - \hat{\mu}_0$$

#### Regression

- 1. Regress  $Y_i$  on  $X_i$  among  $W_i=1$  to form  $\hat{\mu}_1(X_i)$
- 2. Regress  $Y_i$  on  $X_i$  among  $W_i=0$  to form  $\hat{\mu}_0(X_i)$
- 3. Form difference in predictions:

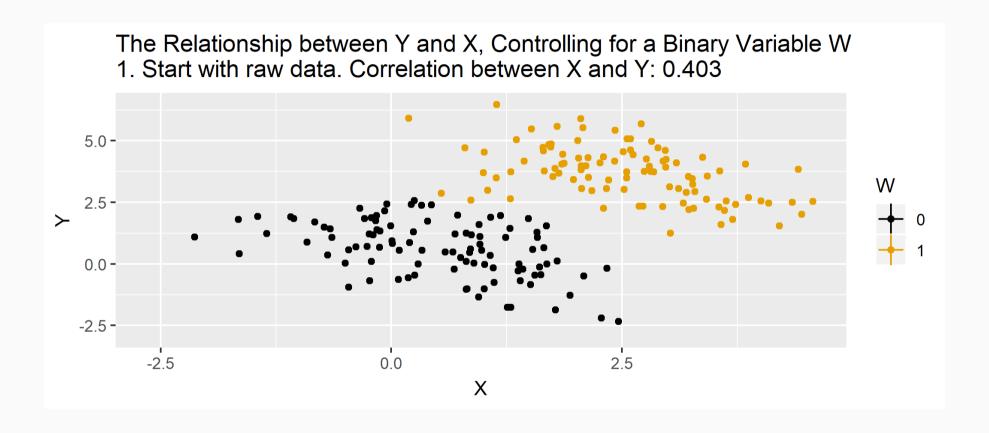
$$\hat{\delta}^{reg} = rac{1}{N} \sum_{i=1}^N \left(\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)
ight)$$

#### Regression

Or estimate in one step,

$$Y_i = \delta W_i + eta X_i + W_i imes \left(X_i - ar{X}
ight) \gamma + arepsilon_i$$

# Animation for regression



#### Simulated data

Now let's do some matching, re-weighting, and regression with simulated data:

```
n \leftarrow 5000

select.dat \leftarrow tibble(

x = runif(n, 0, 1),

z = rnorm(n, 0, 1),

w = (x>0.65),

y = -2.5 + 4*w + 1.5*x + rnorm(n,0,1),

w_alt = (x + z > 0.35),

y_alt = -2.5 + 4*w_alt + 1.5*x + 2.25*z + rnorm(n,0,1)
```

### Simulation: nearest neighbor matching

## Matched number of observations.....

## Matched number of observations (unweighted). 5016

```
nn.est1 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=1,
                          estimand="ATE")
summary(nn.est1)
## Estimate ... 4.0175
## AI SE..... 0.52954
## T-stat..... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                              5000
## Original number of treated obs.....
                                              1732
```

5000

### Simulation: nearest neighbor matching

## Matched number of observations (unweighted). 5016

```
nn.est2 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=2,
                          estimand="ATE")
summary(nn.est2)
## Estimate ... 4.0175
## AI SE..... 0.52954
## T-stat..... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                              5000
## Original number of treated obs.....
                                             1732
## Matched number of observations.....
                                             5000
```

### Simulation: regression

```
reg1.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=1)}
reg1 \( \times \text{ lm(y \( \times \text{ x, data=reg1.dat)}} \)

reg0.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=0)}
reg0 \( \times \text{ lm(y \( \times \text{ x, data=reg0.dat)}} \)
pred1 \( \times \text{ predict(reg1,new=select.dat)} \)
pred0 \( \times \text{ predict(reg0,new=select.dat)} \)
mean(pred1-pred0)
```

## [1] 4.076999

#### Violation of selection on observables

#### NN Matching

```
##
## Estimate... 7.6642
## AI SE.... 0.052903
## T-stat.... 144.87
## p.val.... < 2.22e-16
##
## Original number of observations.... 5000
## Original number of observations.... 2748
## Matched number of observations (unweighted). 23014</pre>
```

#### Regression

```
reg1.dat \leftarrow select.dat %>% filter(w_alt=1)
reg1 \leftarrow lm(y_alt \simeq x, data=reg1.dat)

reg0.dat \leftarrow select.dat %>% filter(w_alt=0)
reg0 \leftarrow lm(y_alt \simeq x, data=reg0.dat)
pred1_alt \leftarrow predict(reg1,new=select.dat)
pred0_alt \leftarrow predict(reg0,new=select.dat)
mean(pred1_alt-pred0_alt)
```

**##** [1] 7.646532

# Pricing and Hospital Profit Status

#### Penalized hospitals

```
final.hcris ← hcris.data %>% ungroup() %>%
  filter(price_denom>100, !is.na(price_denom),
        price_num>0, !is.na(price_num),
        price<100000,
        beds>30, year=2012) %>%
mutate( hvbp_payment = ifelse(is.na(hvbp_payment),0,hvbp_payment),
        hrrp_payment = ifelse(is.na(hrrp_payment),0,abs(hrrp_payment)),
        penalty = (hvbp_payment-hrrp_payment<0))</pre>
```

#### Summary stats

Always important to look at your data before doing any formal analysis. Ask yourself a few questions:

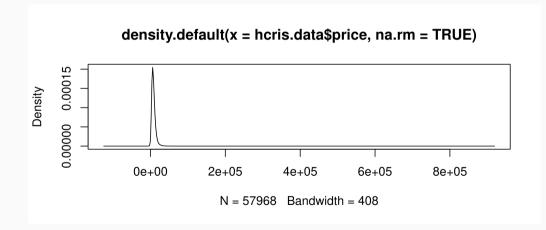
- 1. Are the magnitudes reasonable?
- 2. Are there lots of missing values?
- 3. Are there clear examples of misreporting?

#### **Summary stats**

```
summary(hcris.data$price)

### Min. 1st Qu. Median Mean 3rd Qu. Max. NA's
## -123697 4783 7113 Inf 10230 Inf 63662

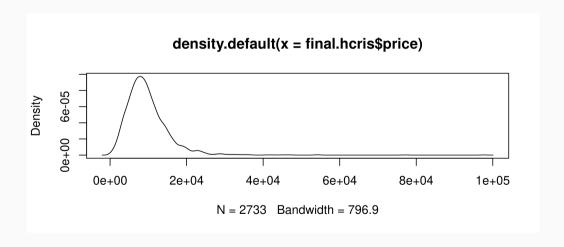
plot(density(hcris.data$price, na.rm=TRUE))
```



```
summary(final.hcris$price)

## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 340.8 6129.9 8705.4 9646.9 11905.4 97688.8

plot(density(final.hcris$price))
```



## Dealing with problems

We've adopted a very brute force way to deal with outlier prices. Other approaches include:

- 1. Investigate very closely the hospitals with extreme values
- 2. Winsorize at certain thresholds (replace extreme values with pre-determined thresholds)
- 3. Impute prices for extreme hospitals

# Differences among penalized hospitals

- Mean price among penalized hospitals: 9,896.31
- Mean price among non-penalized hospitals: 9,560.41
- Mean difference: 335.9

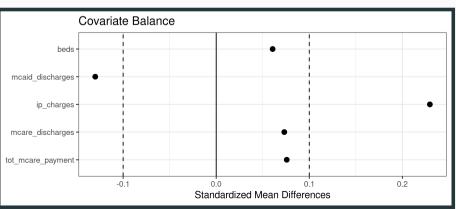
### Comparison of hospitals

Are penalized hospitals sufficiently similar to non-penalized hospitals?

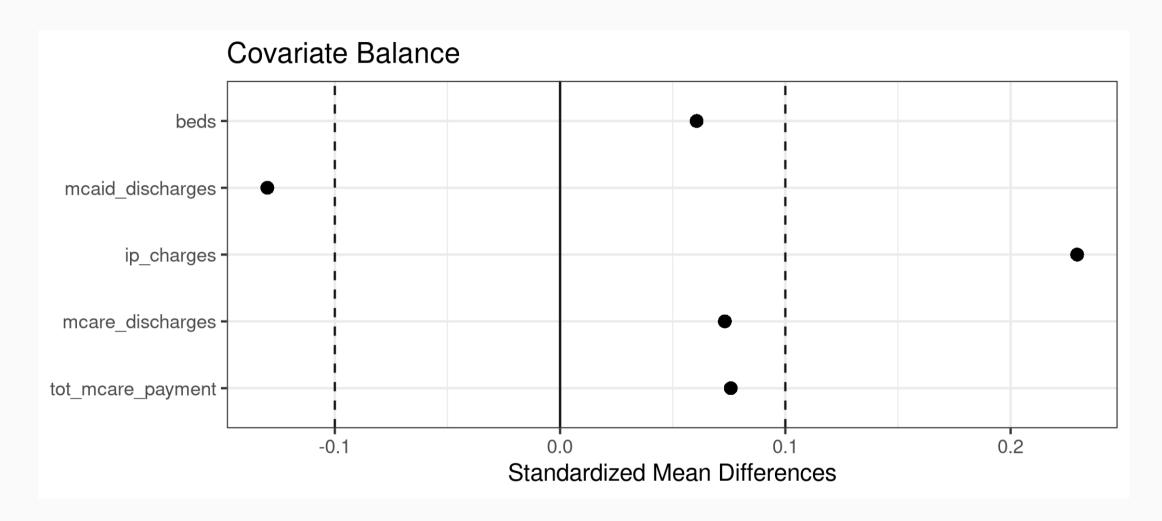
Let's look at covariate balance using a love plot, part of the library(cobalt) package.

### Love plots without adjustment

```
love.plot(bal.tab(lp.covs,treat=lp.vars$penalty), colors="black", shapes="circle", threshold=0.1) +
    theme bw() + theme(legend.position="none")
```



# Love plots without adjustment



#### Using matching to improve balance

Some things to think about:

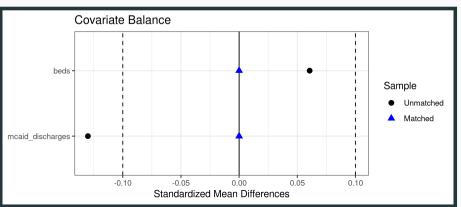
- exact versus nearest neighbor
- with or without ties (and how to break ties)
- measure of distance

### 1. Exact Matching

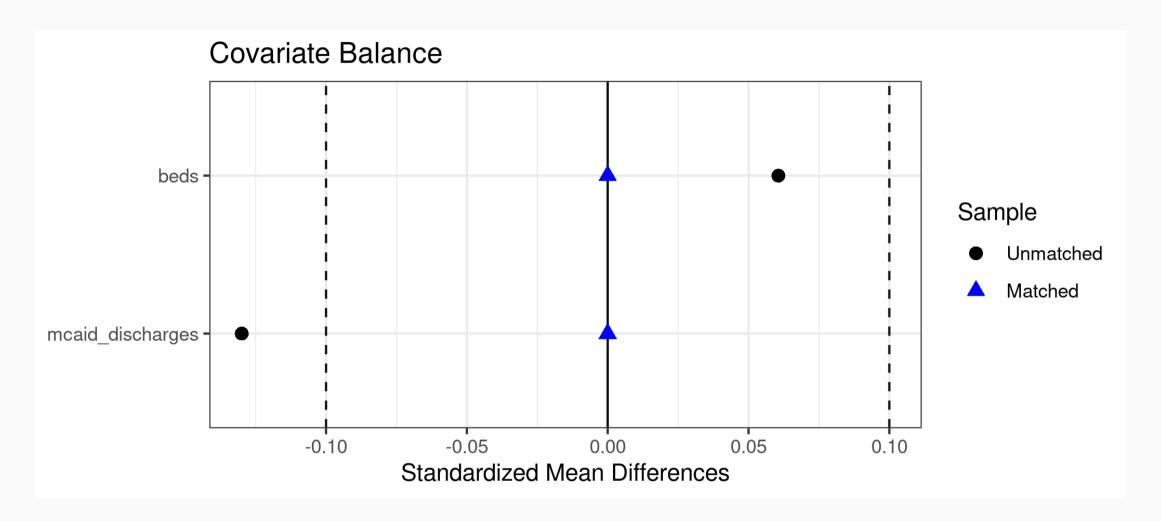
## [1] "Match"

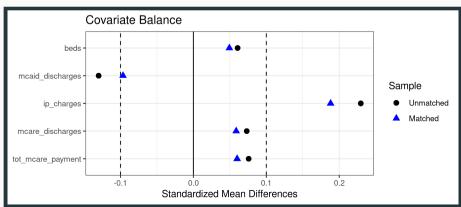
# 1. Exact Matching (on a subset)

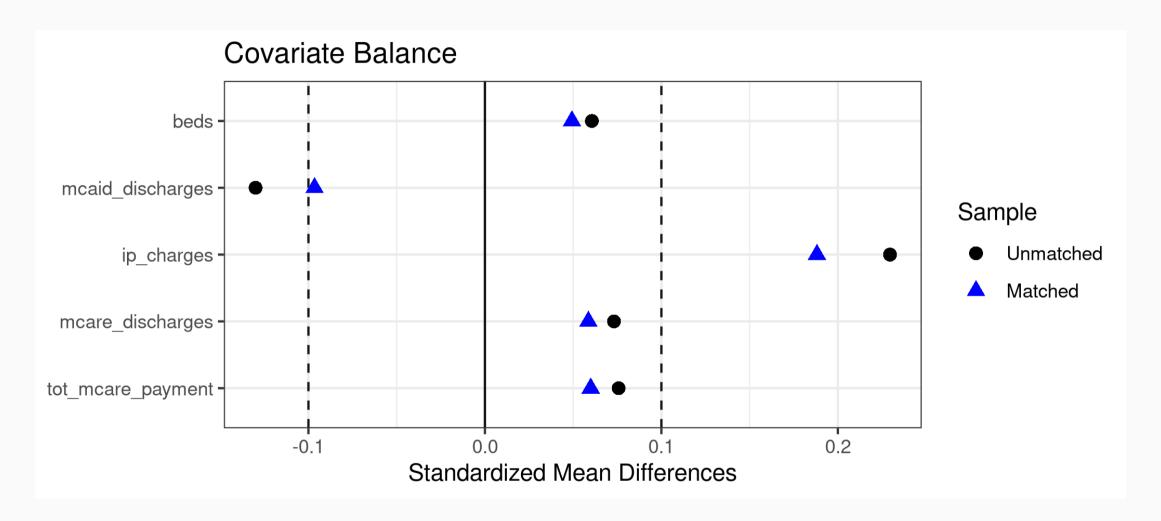
# 1. Exact Matching (on a subset)

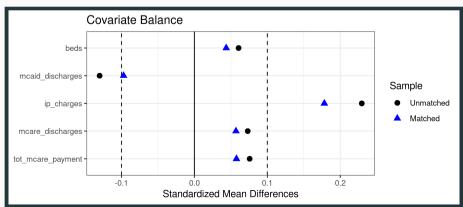


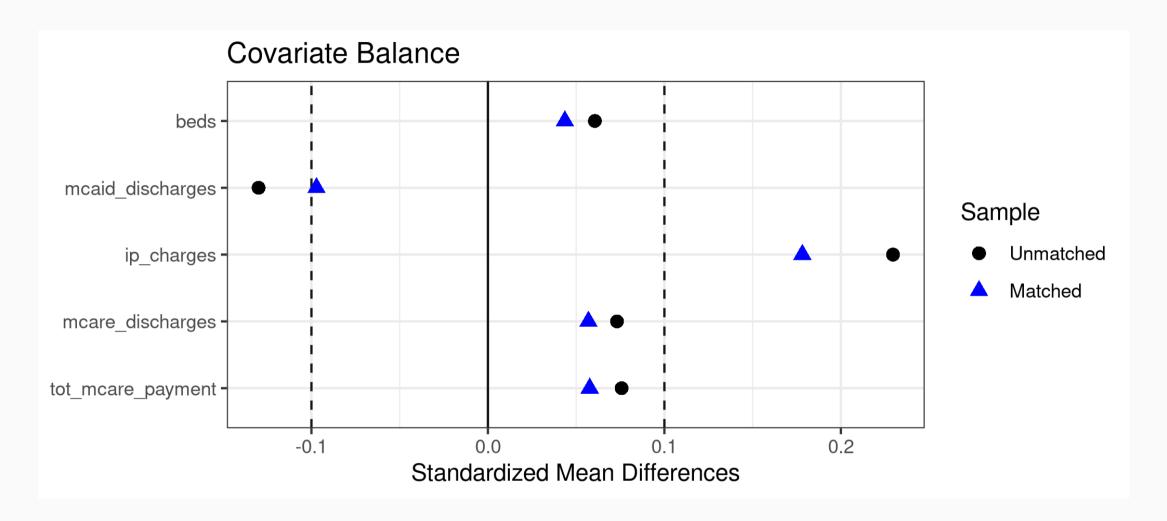
# 1. Exact Matching (on a subset)





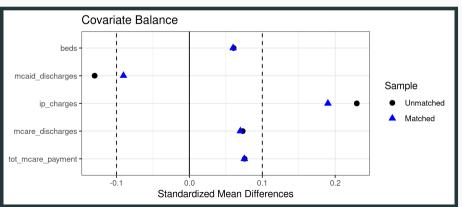




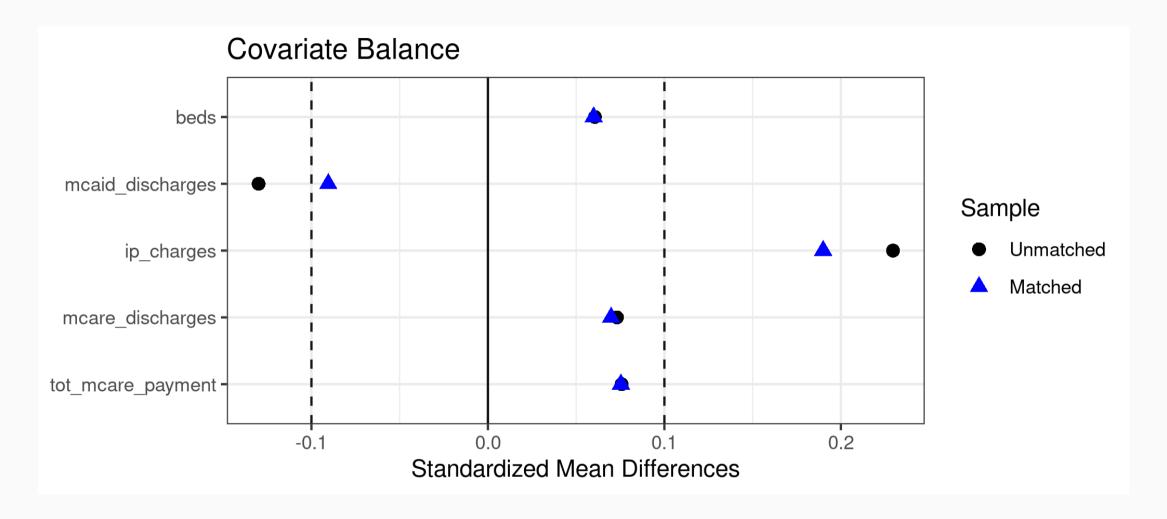


# 2. Nearest neighbor matching (Mahalanobis)

# 2. Nearest neighbor matching (Mahalanobis)

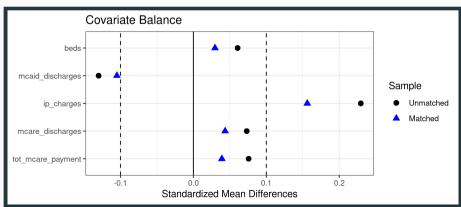


# 2. Nearest neighbor matching (Mahalanobis)

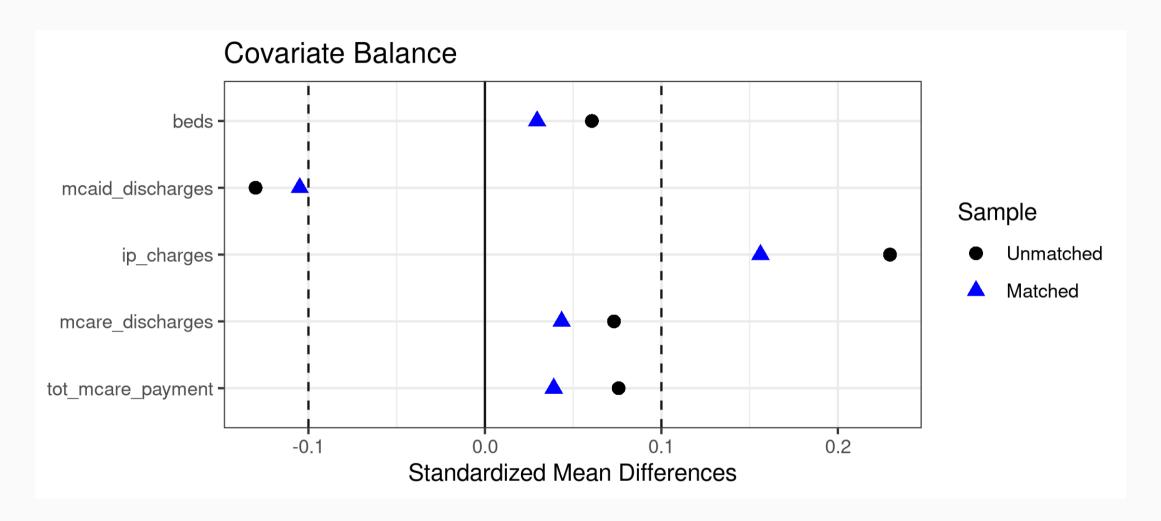


### 2. Nearest neighbor matching (propensity score)

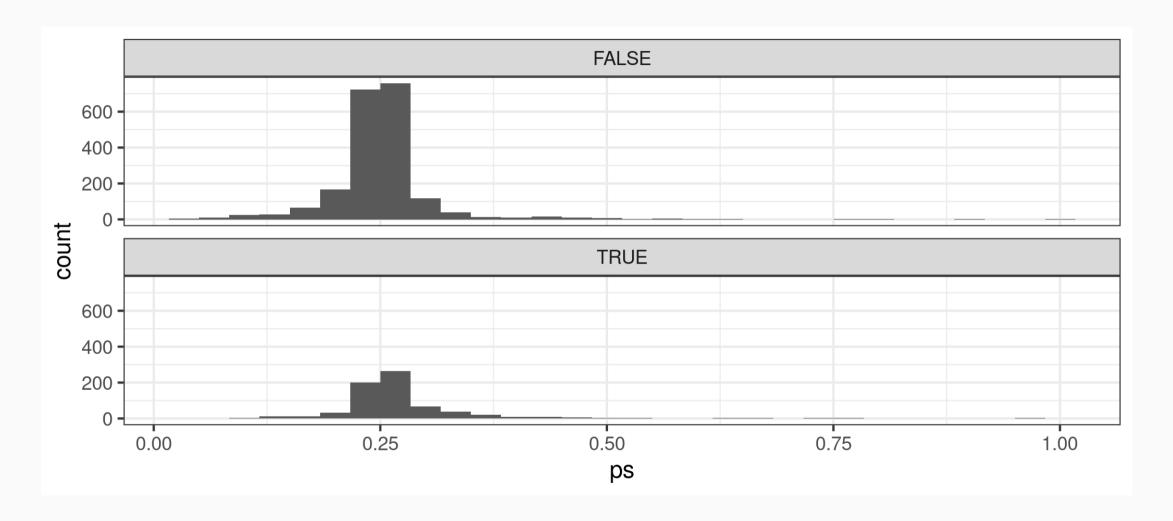
# 2. Nearest neighbor matching (propensity score)



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# 3. Weighting



### Results: Exact matching

```
##
## Estimate... 1777.6
## AI SE..... 34.725
## T-stat.... 51.191
## p.val..... < 2.22e-16
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 12
## Matched number of observations (unweighted). 12
##
## Number of obs dropped by 'exact' or 'caliper' 2695</pre>
```

### Results: Nearest neighbor

#### • Inverse variance

```
##
## Estimate... -526.95
## AI SE..... 223.06
## T-stat.... -2.3623
## p.val..... 0.01816
##
## Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 2711
```

### Results: Nearest neighbor

#### Mahalanobis

```
##
## Estimate... -492.82
## AI SE..... 223.55
## T-stat.... -2.2046
## p.val.... 0.027485
##
## Original number of observations...... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2708
```

### Results: Nearest neighbor

#### Propensity score

```
##
## Estimate... -201.03
## AI SE..... 275.76
## T-stat.... -0.72898
## p.val..... 0.46601
##
## Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 14795
```

### Results: IPW weighting

```
lp.vars \leftarrow lp.vars %>%
mutate(ipw = case_when(
    penalty=1 ~ 1/ps,
    penalty=0 ~ 1/(1-ps),
    TRUE ~ NA_real_
))
mean.t1 \leftarrow lp.vars %>% filter(penalty=1) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t0 \leftarrow lp.vars %>% filter(penalty=0) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t1$mean_p - mean.t0$mean_p
```

```
## [1] -196.8922
```

# Results: IPW weighting with regression

```
ipw.reg ← lm(price ~ penalty, data=lp.vars, weights=ipw)
summarv(ipw.reg)
##
## Call:
### lm(formula = price ~ penalty, data = lp.vars, weights = ipw)
##
## Weighted Residuals:
     Min
          1Q Median
                       3Q
                               Max
## -18691 -4802 -1422 2651 94137
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 9876.4 147.8 66.808 <2e-16 ***
## penaltyTRUE -196.9 211.2 -0.932 0.351
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7829 on 2705 degrees of freedom
## Multiple R-squared: 0.0003211, Adjusted R-squared: -4.85e-05
## F-statistic: 0.8688 on 1 and 2705 DF, p-value: 0.3514
```

### Results: Regression

```
## [1] -5.845761
```

### Results: Regression in one step

### Results: Regression in one step

```
###
## Call:
## lm(formula = price ~ penalty + beds + mcaid discharges + ip charges +
      mcare discharges + tot mcare payment + beds diff + mcaid diff +
###
      ip diff + mcare diff + mpay diff, data = reg.dat)
###
##
## Residuals:
     Min
             10 Median
                          3Q
                               Max
## -38175 -2900
                 -597
                        2105 67409
##
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept) 8.466e+03 1.711e+02 49.482 < 2e-16 ***
## penaltyTRUE
              -5.846e+00 2.124e+02 -0.028 0.97804
## beds
                  1.107e+00 1.421e+00 0.779 0.43618
## mcaid discharges -4.714e-01 7.296e-02 -6.462 1.23e-10 ***
## ip charges
                    6.426e-06 1.285e-06 5.002 6.04e-07 ***
## mcare discharges -8.122e-01 9.257e-02 -8.774 < 2e-16 ***
                                        13.857 < 2e-16 ***
## tot_mcare_payment 9.502e-05
                              6.858e-06
## beds diff
                    2.517e+00 2.986e+00
                                         0.843 0.39931
## mcaid diff
             1.058e-01 1.570e-01
                                        0.674 0.50050
## ip_diff
                   -4.534e-06 2.027e-06 -2.237 0.02539 *
                                        2.657 0.00793 **
## mcare diff
             4.806e-01 1.809e-01
## mpay diff
                   -5.452e-05 1.321e-05 -4.128 3.78e-05 ***
## ---
```

### Summary of ATEs

- 1. Exact matching: 1777.63
- 2. NN matching, inverse variance: -526.95
- 3. NN matching, mahalanobis: -492.82
- 4. NN matching, pscore: -201.03
- 5. Inverse pscore weighting: -196.89
- 6. IPW regression: -196.89
- 7. Regression: -5.85
- 8. Regression 1-step: -5.85

### So what have we learned?

### Key assumptions for causal inference

- 1. Selection on observables
- 2. Common support

These become more nuanced but the intuition is the same in almost all questions of causal inference.

### Causal effect assuming selection on observables

If we assume selection on observables holds, then we only need to condition on the relevant covariates to identify a causal effect. But we still need to ensure common support...

- 1. Matching
- 2. Reweighting
- 3. Regression