

Part 4: Hospital Prices and Penalties

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Penalized hospitals

```
hcris.data ← read rds(here("data/HCRIS Data.rds"))
hcris.data ← hcris.data %>%
 mutate( discount factor = 1-tot discounts/tot charges,
          price num = (ip charges + icu charges + ancillary charges)*discount factor - tot mcare payment,
          price denom = tot discharges - mcare discharges,
          price = price num/price denom)
final.hcris ← hcris.data %>% ungroup() %>%
  filter(price denom>100, !is.na(price denom),
         price_num>0, !is.na(price num).
         price<100000,
         beds>30, year=2012) %>%
  mutate( hvbp_payment = ifelse(is.na(hvbp_payment), 0, hvbp_payment),
          hrrp payment = ifelse(is.na(hrrp payment), 0, abs(hrrp payment)),
    penalty = (hvbp payment-hrrp payment<0))</pre>
```

Summary stats

Always important to look at your data before doing any formal analysis. Ask yourself a few questions:

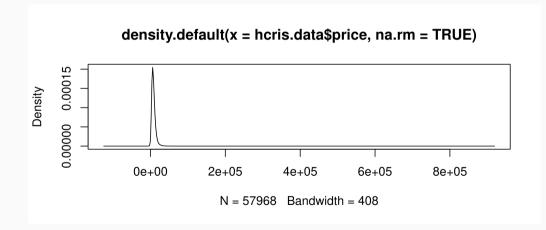
- 1. Are the magnitudes reasonable?
- 2. Are there lots of missing values?
- 3. Are there clear examples of misreporting?

Summary stats

```
summary(hcris.data$price)

### Min. 1st Qu. Median Mean 3rd Qu. Max. NA's
## -123697 4783 7113 Inf 10230 Inf 63662

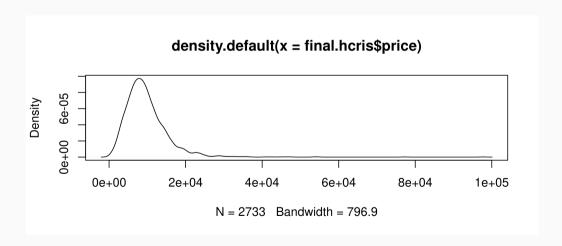
plot(density(hcris.data$price, na.rm=TRUE))
```



```
summary(final.hcris$price)

## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 340.8 6129.9 8705.4 9646.9 11905.4 97688.8

plot(density(final.hcris$price))
```



Dealing with problems

We've adopted a very brute force way to deal with outlier prices. Other approaches include:

- 1. Investigate very closely the hospitals with extreme values
- 2. Winsorize at certain thresholds (replace extreme values with pre-determined thresholds)
- 3. Impute prices for extreme hospitals

Differences among penalized hospitals

- Mean price among penalized hospitals: 9,896.31
- Mean price among non-penalized hospitals: 9,560.41
- Mean difference: 335.9

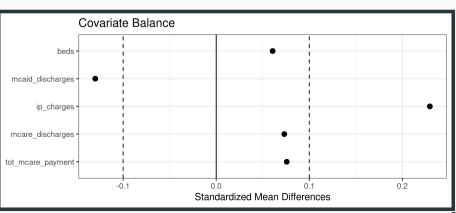
Comparison of hospitals

Are penalized hospitals sufficiently similar to non-penalized hospitals?

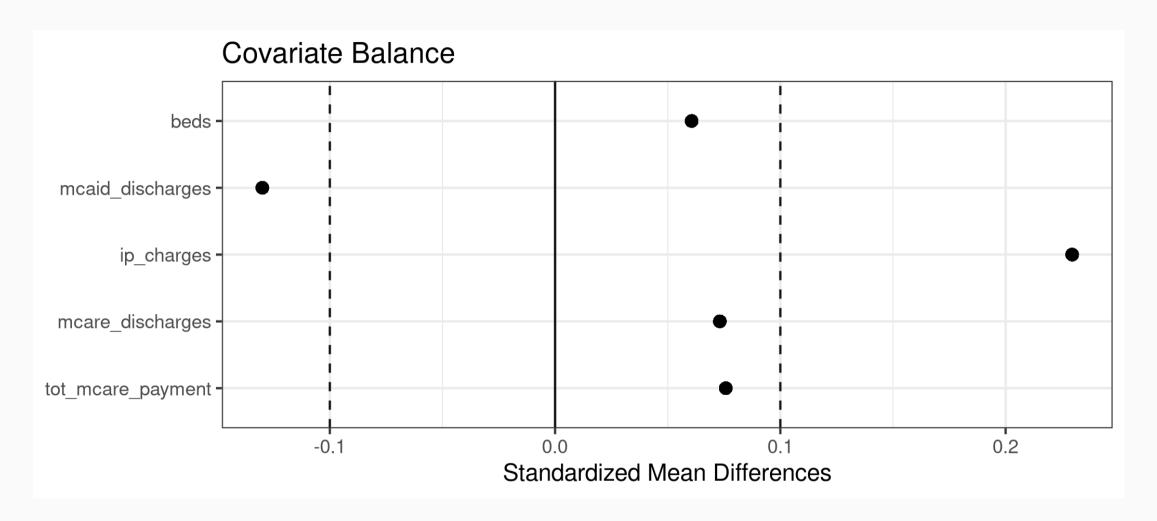
Let's look at covariate balance using a love plot, part of the library(cobalt) package.

Love plots without adjustment

```
love.plot(bal.tab(lp.covs,treat=lp.vars$penalty), colors="black", shapes="circle", threshold=0.1) +
    theme bw() + theme(legend.position="none")
```



Love plots without adjustment



Using matching to improve balance

Some things to think about:

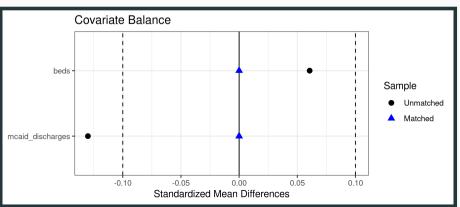
- exact versus nearest neighbor
- with or without ties (and how to break ties)
- measure of distance

1. Exact Matching

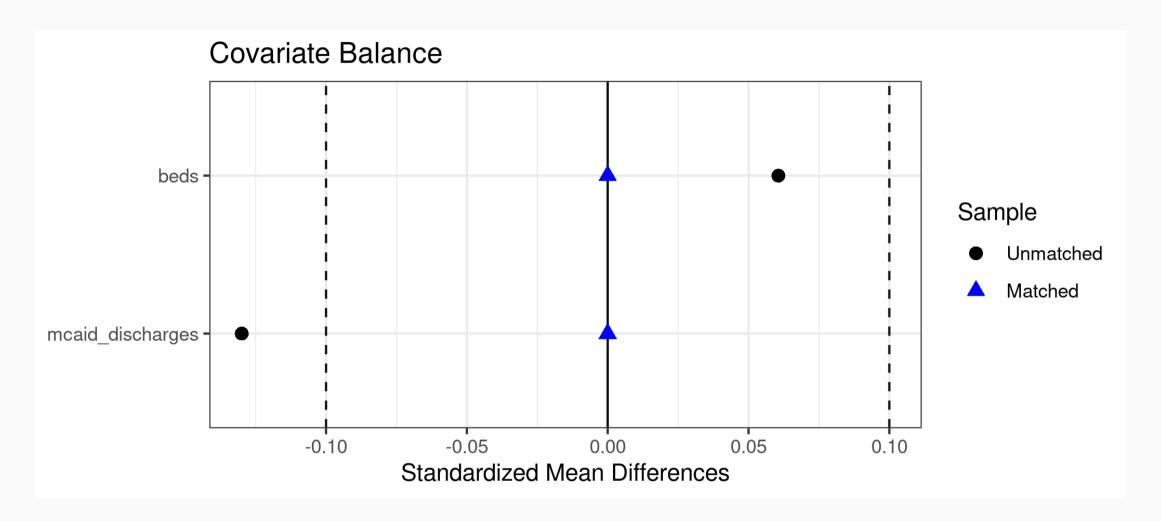
attr(,"class")
[1] "Match"

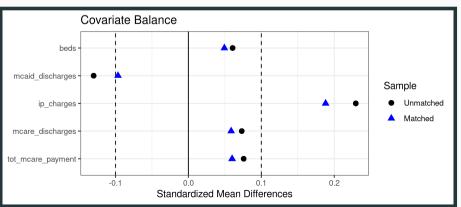
1. Exact Matching (on a subset)

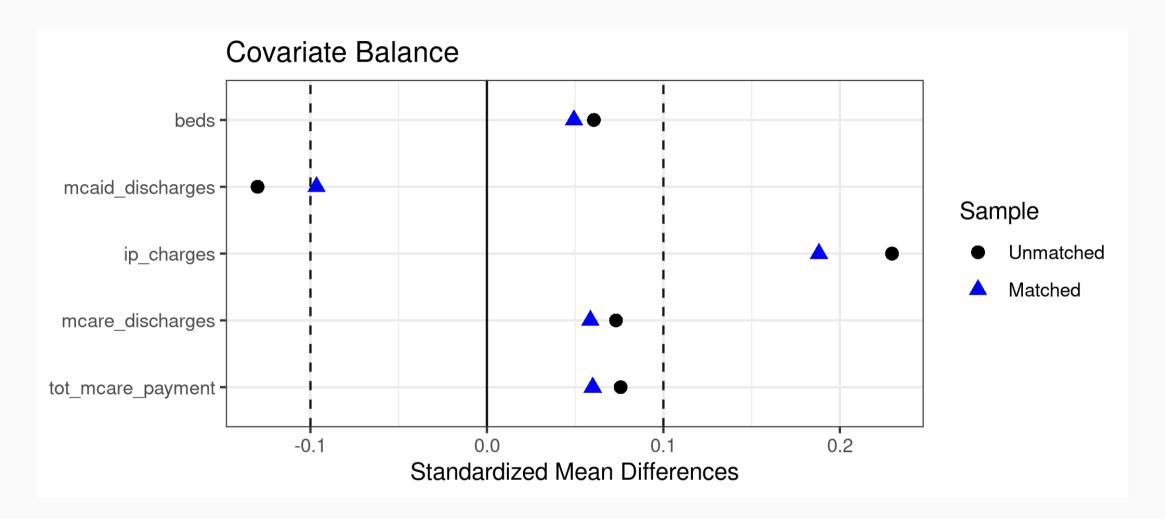
1. Exact Matching (on a subset)

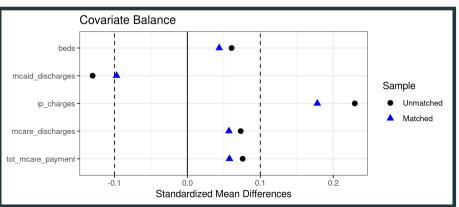


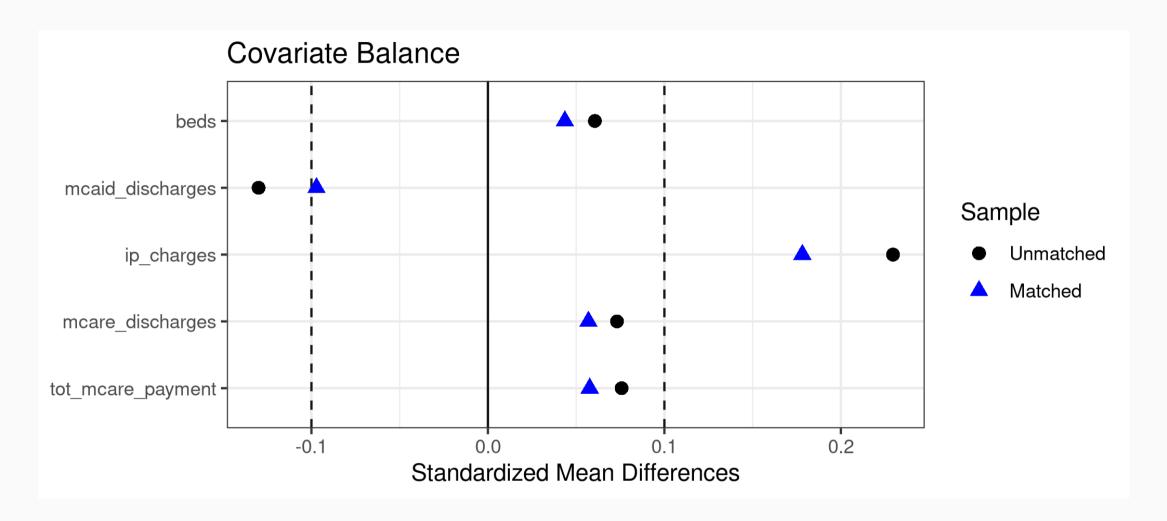
1. Exact Matching (on a subset)





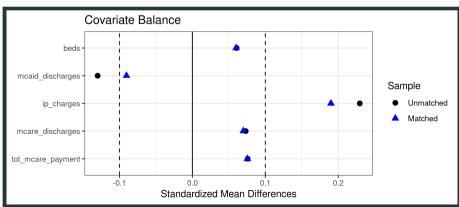




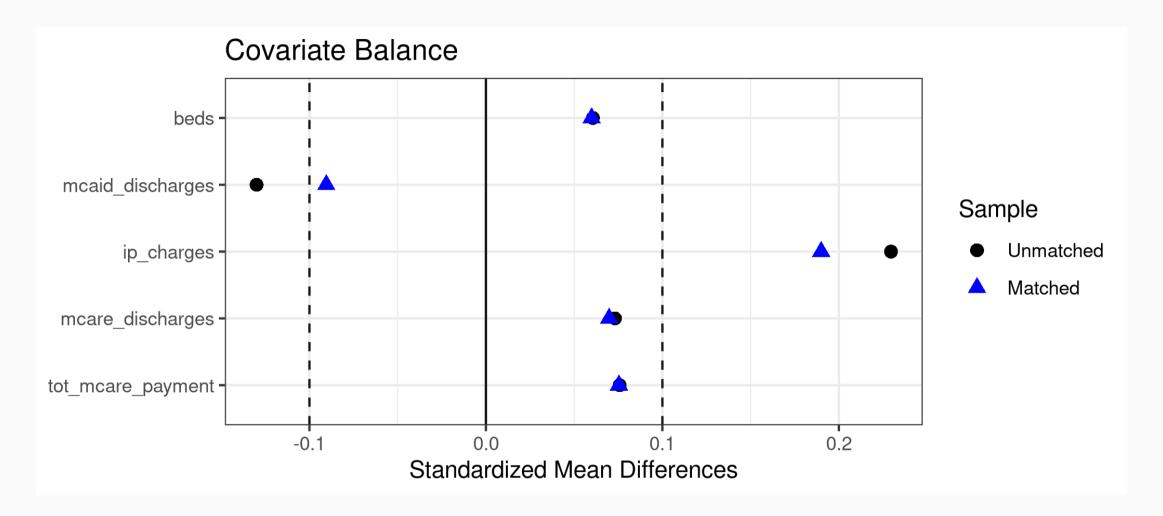


2. Nearest neighbor matching (Mahalanobis)

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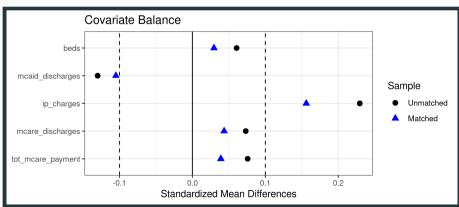


2. Nearest neighbor matching (Mahalanobis)

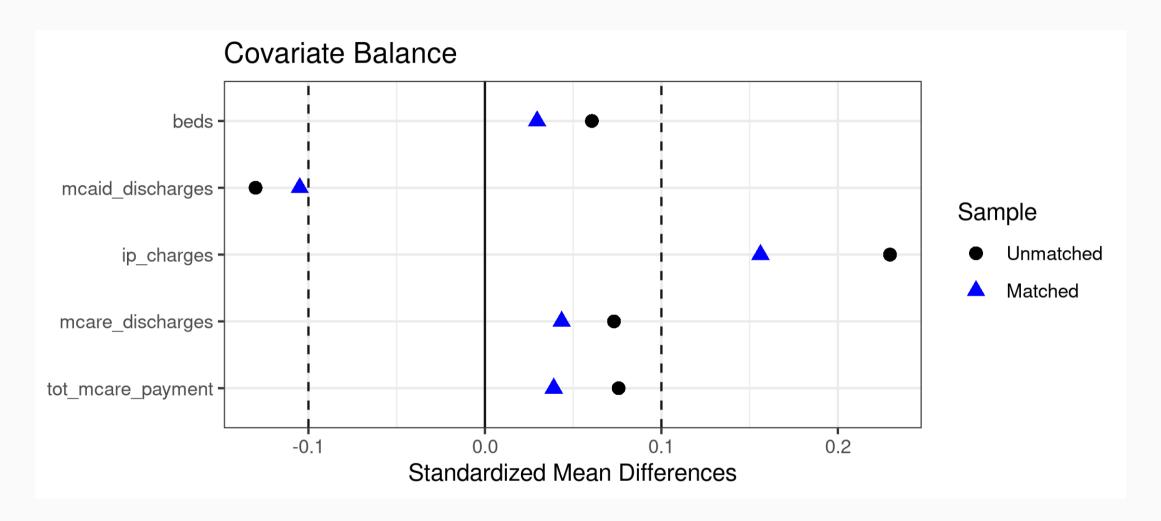


2. Nearest neighbor matching (propensity score)

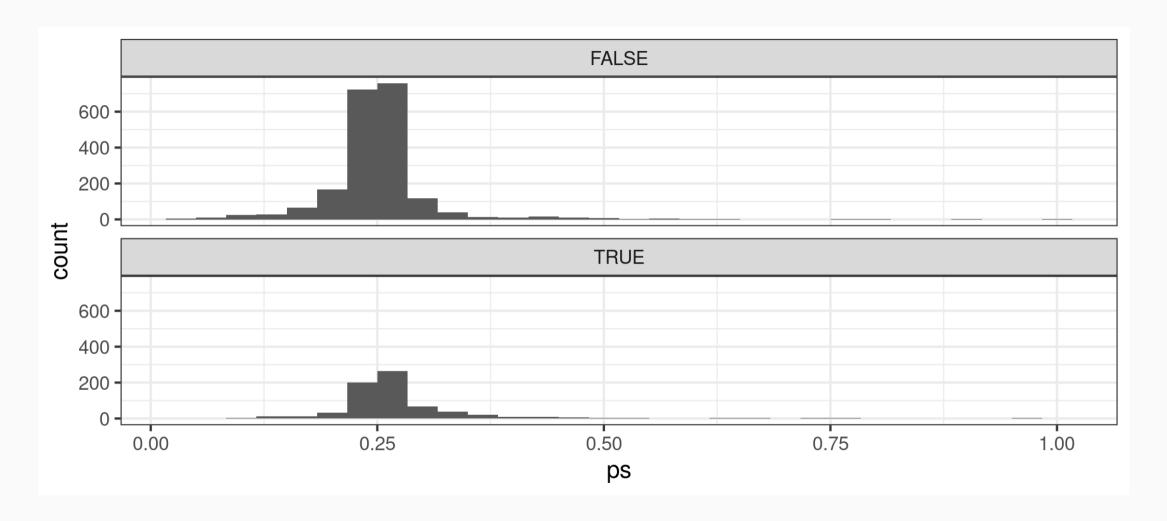
2. Nearest neighbor matching (propensity score)



2. Nearest neighbor matching (propensity score)



3. Weighting



Results: Exact matching

```
##
## Estimate... 1777.6
## AI SE..... 34.725
## T-stat.... 51.191
## p.val..... < 2.22e-16
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 12
## Matched number of observations (unweighted). 12
##
## Number of obs dropped by 'exact' or 'caliper' 2695</pre>
```

Results: Nearest neighbor

• Inverse variance

```
##
## Estimate... -526.95
## AI SE..... 223.06
## T-stat.... -2.3623
## p.val.... 0.01816
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2711
```

Results: Nearest neighbor

Mahalanobis

```
##
## Estimate... -492.82
## AI SE..... 223.55
## T-stat.... -2.2046
## p.val.... 0.027485
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2708
```

Results: Nearest neighbor

Propensity score

```
##
## Estimate... -201.03
## AI SE..... 275.76
## T-stat.... -0.72898
## p.val..... 0.46601
##
##
Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 14795
```

Results: IPW weighting

```
lp.vars \leftarrow lp.vars %>%
mutate(ipw = case_when(
    penalty=1 ~ 1/ps,
    penalty=0 ~ 1/(1-ps),
    TRUE ~ NA_real_
))
mean.t1 \leftarrow lp.vars %>% filter(penalty=1) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t0 \leftarrow lp.vars %>% filter(penalty=0) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t1$mean_p - mean.t0$mean_p
```

```
## [1] -196.8922
```

Results: IPW weighting with regression

```
ipw.reg ← lm(price ~ penalty, data=lp.vars, weights=ipw)
summarv(ipw.reg)
##
## Call:
### lm(formula = price ~ penalty, data = lp.vars, weights = ipw)
##
## Weighted Residuals:
     Min
          1Q Median
                       3Q
                               Max
## -18691 -4802 -1422 2651 94137
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 9876.4 147.8 66.808 <2e-16 ***
## penaltyTRUE -196.9 211.2 -0.932 0.351
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7829 on 2705 degrees of freedom
## Multiple R-squared: 0.0003211, Adjusted R-squared: -4.85e-05
## F-statistic: 0.8688 on 1 and 2705 DF, p-value: 0.3514
```

Results: Regression

```
## [1] -5.845761
```

Results: Regression in one step

Results: Regression in one step

```
###
## Call:
## lm(formula = price ~ penalty + beds + mcaid discharges + ip charges +
      mcare discharges + tot mcare payment + beds diff + mcaid diff +
###
      ip diff + mcare diff + mpay diff, data = reg.dat)
###
##
## Residuals:
     Min
             10 Median
                          3Q
                                Max
## -38175 -2900
                 -597
                        2105 67409
##
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept) 8.466e+03 1.711e+02 49.482 < 2e-16 ***
## penaltyTRUE
                   -5.846e+00 2.124e+02 -0.028 0.97804
## beds
                  1.107e+00 1.421e+00 0.779 0.43618
## mcaid discharges -4.714e-01 7.296e-02 -6.462 1.23e-10 ***
## ip charges
                    6.426e-06 1.285e-06 5.002 6.04e-07 ***
## mcare discharges -8.122e-01 9.257e-02 -8.774 < 2e-16 ***
                                        13.857 < 2e-16 ***
## tot_mcare_payment 9.502e-05
                              6.858e-06
## beds diff
                    2.517e+00 2.986e+00
                                         0.843 0.39931
## mcaid diff
             1.058e-01 1.570e-01
                                        0.674 0.50050
## ip_diff
                   -4.534e-06 2.027e-06 -2.237 0.02539 *
                                        2.657 0.00793 **
## mcare diff
             4.806e-01 1.809e-01
## mpay diff
                   -5.452e-05 1.321e-05 -4.128 3.78e-05 ***
## ---
```

Summary of ATEs

- 1. Exact matching: 1777.63
- 2. NN matching, inverse variance: -526.95
- 3. NN matching, mahalanobis: -492.82
- 4. NN matching, pscore: -201.03
- 5. Inverse pscore weighting: -196.89
- 6. IPW regression: -196.89
- 7. Regression: -5.85
- 8. Regression 1-step: -5.85

Summary of ATEs

Why such large differences between linear (unweighted) regression and other approaches?

Problem is due to common support. Without weighting, the treated group looks very different than the control group, and standard OLS (without weights) doesn't do anything to account for this.

So what have we learned?

Key assumptions for causal inference

- 1. Selection on observables
- 2. Common support

These become more nuanced but the intuition is the same in almost all questions of causal inference.

Causal effect assuming selection on observables

If we assume selection on observables holds, then we only need to condition on the relevant covariates to identify a causal effect. But we still need to ensure common support...

- 1. Matching
- 2. Reweighting
- 3. Regression