

### Section 1: Hospital Pricing and Selection on Observables

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# Background on Hospital Pricing

Defining characteristic of hospital services: it's complicated!

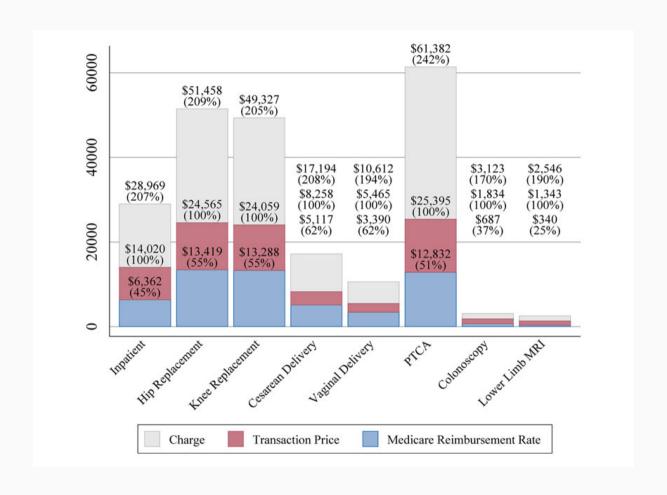
	-			
1/04/11	1	041000	LEARLINK DUO-VENT	17.00
1/04/11	1	0406462	TUBE COMNECTING STERIL 6FT	27.00
1/05/11	Ť	0406462	TUBE CONNECTING STERIL OFT	27.00
1/05/11	1	3005741	ACCU-CHEK CCRV	18.00
1/05/11	1	3019692	SURGICEL 2X14 STRIP EACH	451.00
1/05/11	-	3005741	ACCU-CHEK CCRV	
1/05/11	-	3005741	ACCU-CHEK CCRV	18.00
1/05/11	1-	3019692	SURGICEL 2X14 STRIP EACH	18.00_
	1	3005741	ACCU-CHEK CCRV	451.00-
1/05/11	1	3005741	ACCU-CREK CCRV	18.00
1/05/11	10	2900025	OKYGEN HOURLY	18.00
1/05/11	1	0402230	LEUKINS TUEE SPECIM TRAP	560.00
1/05/11	1	0416826	SET EXTENSION 1-VALVE	77.00
1/05/11	1	0406793	SUCTION YANKAUER	12.00
1/05/11	_ 1	0416018	SECOND DEEL SET LUER LOCK	44.00
		042-64	DUE LOCK	5.00

Brill, Steven. 2013. "Bitter Pill: Why Medical Bills are Killing Us." \*Time Magazine\*.

Lots of different payers paying lots of different prices:

- Medicare fee-for-service prices
- Medicaid payments
- Private insurance negotiations (including Medicare Advantage)
- But what about the price to patients?

Price  $\neq$  charge  $\neq$  cost  $\neq$  patient out-of-pocket spending



Source: Health Care Pricing Project

Not clear what exactly is negotiated...

#### Fee-for-service

- price per procedure
- percentage of charges
- markup over Medicare rates

#### Capitation

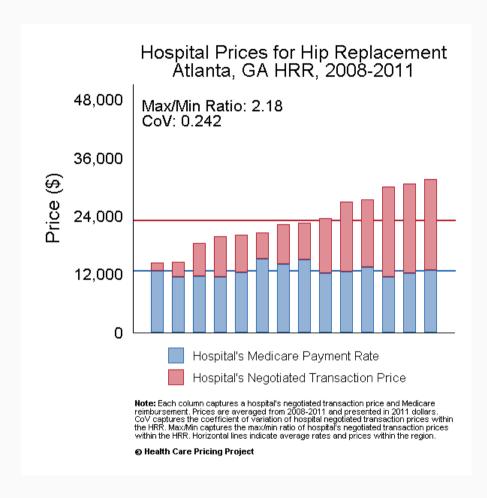
- payment per patient
- pay-for-performance
- shared savings

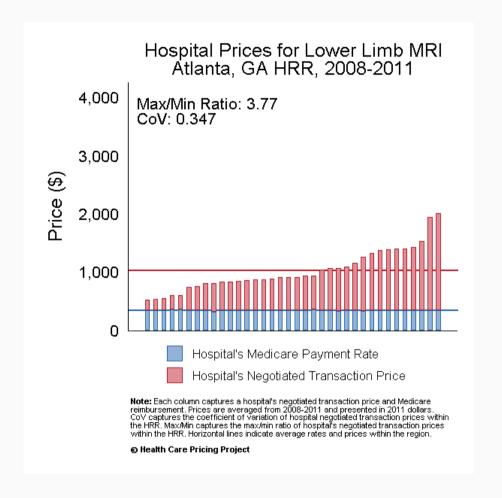
# Hospital prices in real life

#### A few empirical facts:

- 1. Hospital services are expensive
- 2. Prices vary dramatically across different areas
- 3. Lack of competition is a major reason for high prices

# Hospital prices in real life





Source: Health Care Pricing Project

# **Understanding HCRIS Data**

#### What is HCRIS?

Healthcare Cost Report Information System ('cost reports')

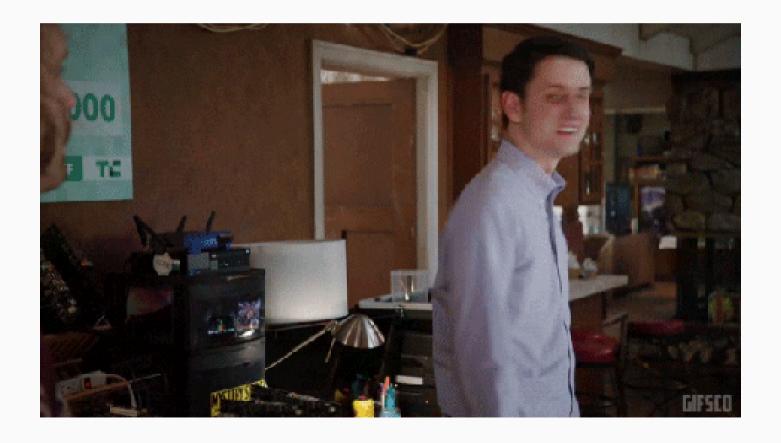
- Nursing Homes (SNFs)
- Hospice
- Home Health Agencies
- Hospitals

# **Hospital Cost Reports**

TAT	EMENT OF PATIENT REVENUES	PROVIDER CCN:	PERIOD:	WORKSHEET G-2,	
AND OPERATING EXPENSES		TROVIDER CCIV.	FROM	PARTS I & II	
			TO	. I AKIBI KII	
			10		
PART	I - PATIENT REVENUES				
	T THILLY REVERGES				
		INPATIENT OUTPATIENT		TOTAL	
	REVENUE CENTER	1	2	3	1
	GENERAL INPATIENT ROUTINE CARE SERVICES				
1	Hospital				1
2	Subprovider IPF				2
3	Subprovider IRF				3
4	Subprovider (Other)				4
5	Swing bed - SNF				5
6	Swing bed - NF				6
7	Skilled nursing facility				7
8	Nursing facility				8
9	Other long term care				9
10	Total general inpatient care services (sum of lines 1-9)				10
	INTENSIVE CARE TYPE INPATIENT HOSPITAL SERVICES				
11	Intensive care unit				11
12	Coronary care unit				12
13	Burn intensive care unit				13
14	Surgical intensive care unit				14
15	Other special care (specify)				15
16	Total intensive care type inpatient hospital services (sum of				16
	of lines 11-15)				
17	Total inpatient routine care services (sum of lines 10 and 16)				17
18	Ancillary services				18
19	Outpatient services				19
20	Rural Health Clinic (RHC)				20
21	Federally Qualified Health Center (FQHC)				21
22	Home health agency				22

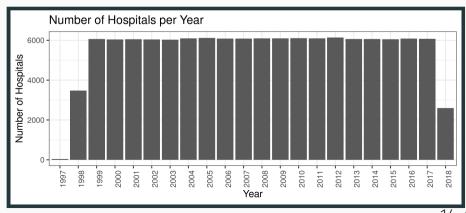
#### The Data

Let's work with the HCRIS GitHub repository. But forming the dataset is up to you this time.

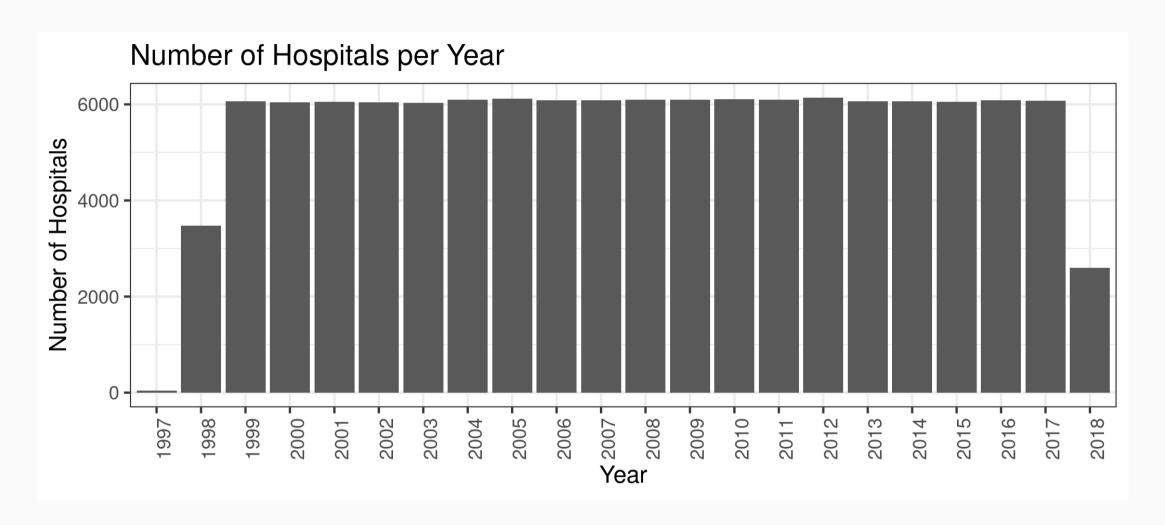


#### The Data

```
hcris.data %>%
  ggplot(aes(x=as.factor(year))) +
  geom_bar() +
  labs(
    x="Year",
    y="Number of Hospitals",
    title="Number of Hospitals per Year"
  ) + theme_bw() +
  theme(axis.text.x = element_text(angle = 90, hjust=1))
```

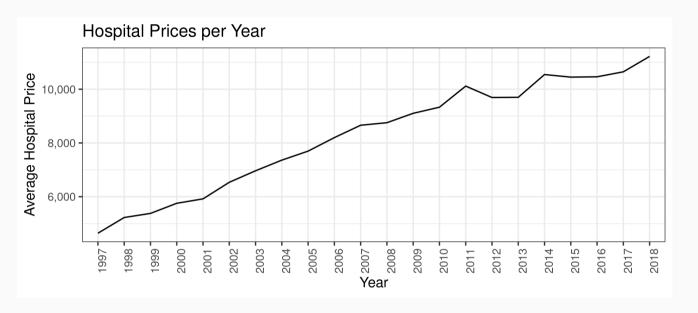


# Number of hospitals

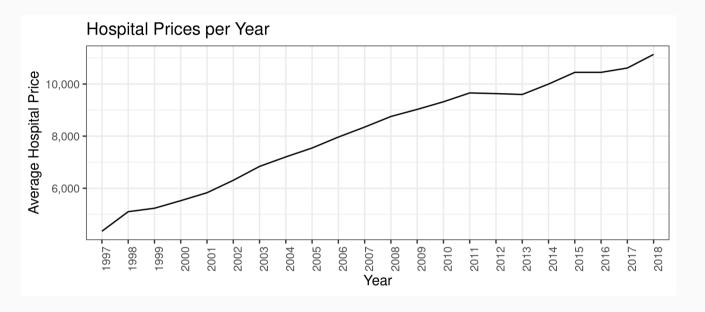


# Estimating hospital prices

# Estimating hospital prices



# Estimating hospital prices



### Potential Outcomes Framework

### What is a "Potential Outcome"



Treatment effect =Y(1)-Y(0)= \$15,000

#### What is a "Potential Outcome"



Treatment effect =Y(1)-Y(0)=?



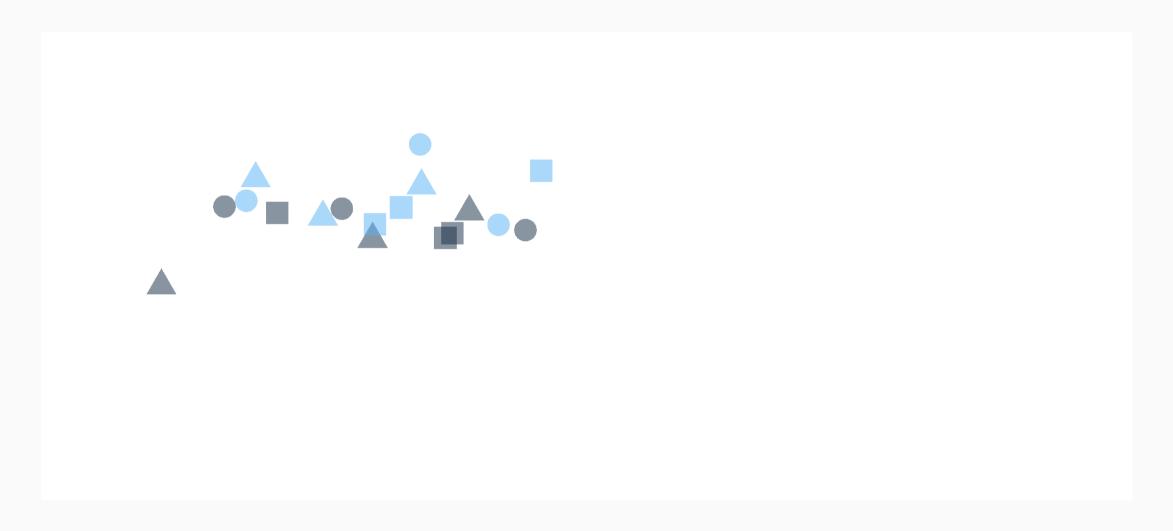
$$Y(0) = ?$$

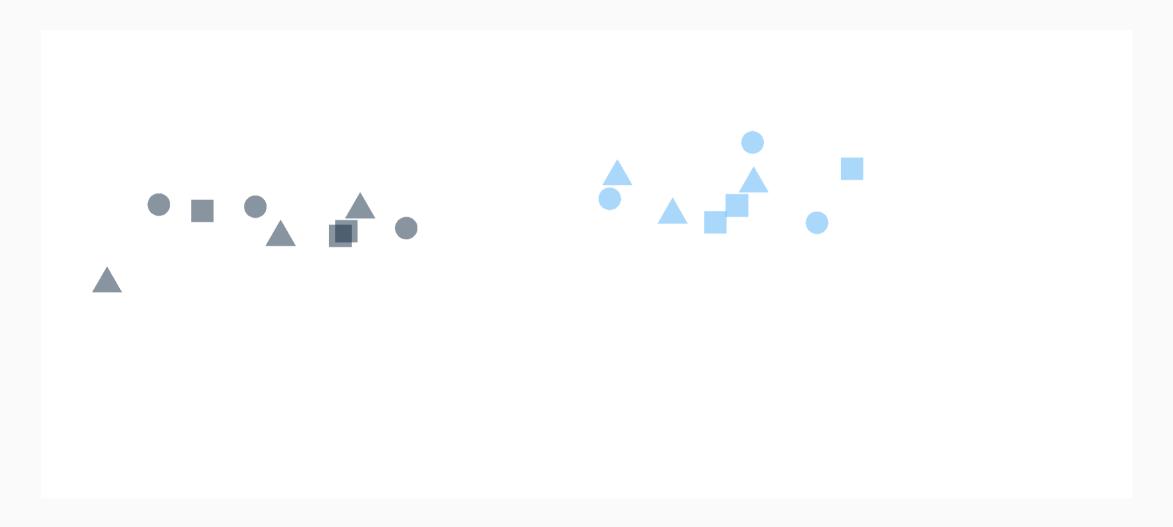
# Do we ever observe the potential outcomes?

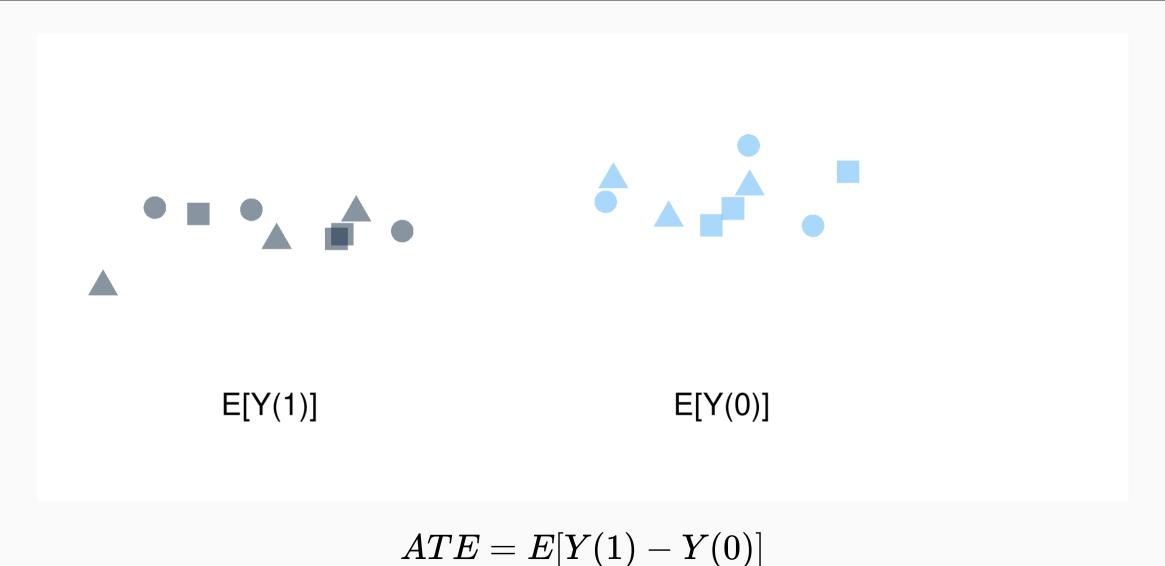


Without a time machine...hard to get individual treatment effects.

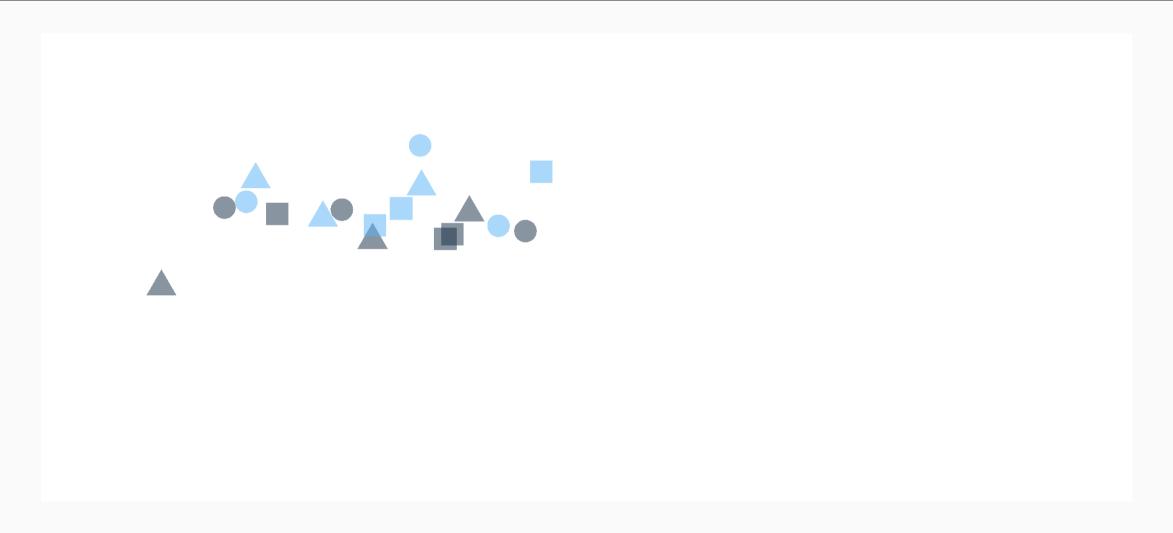


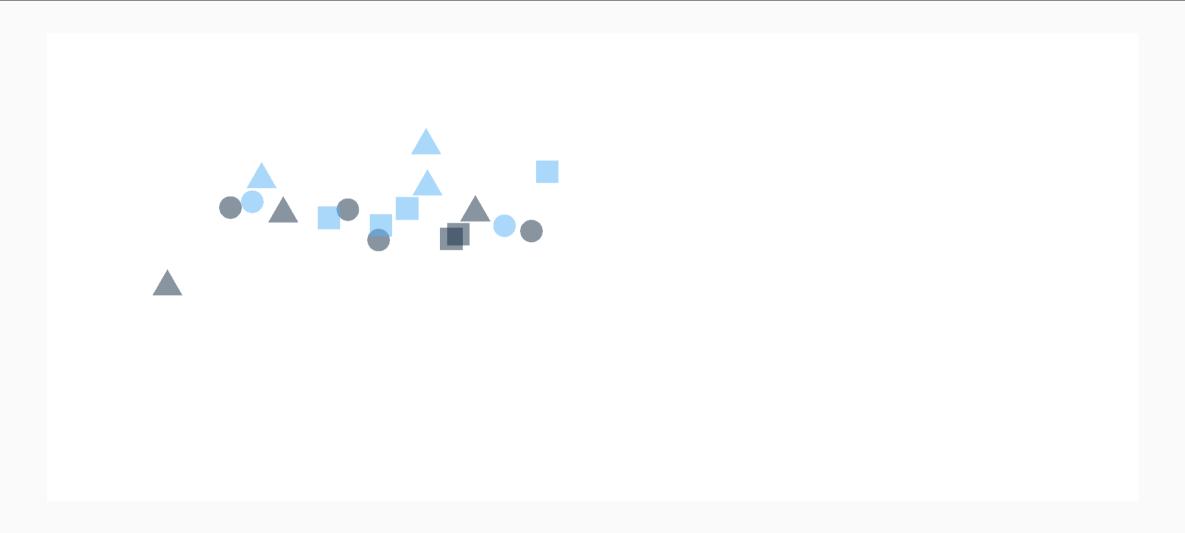


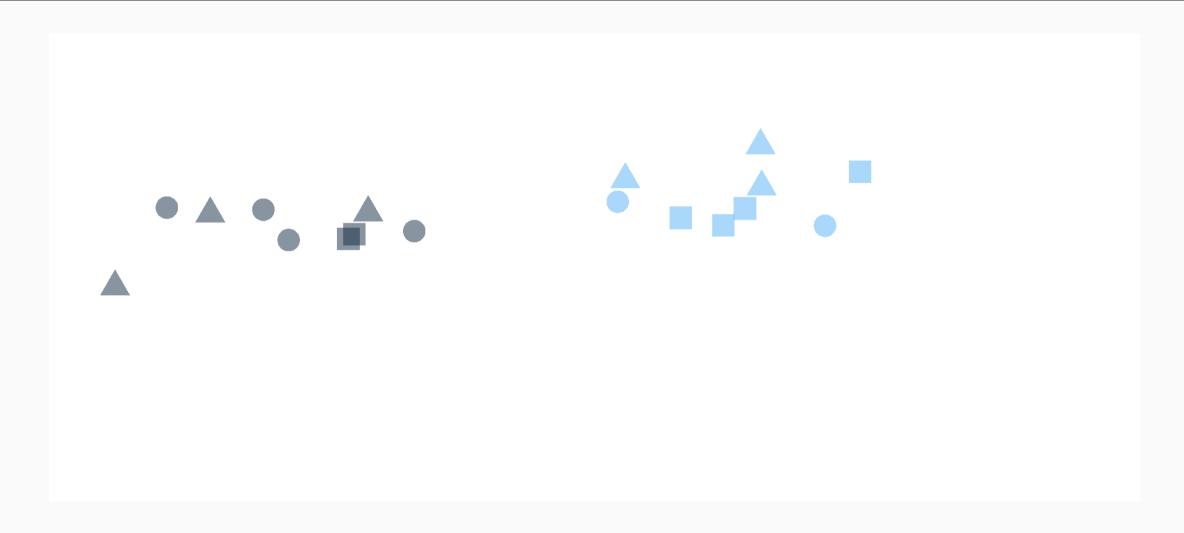


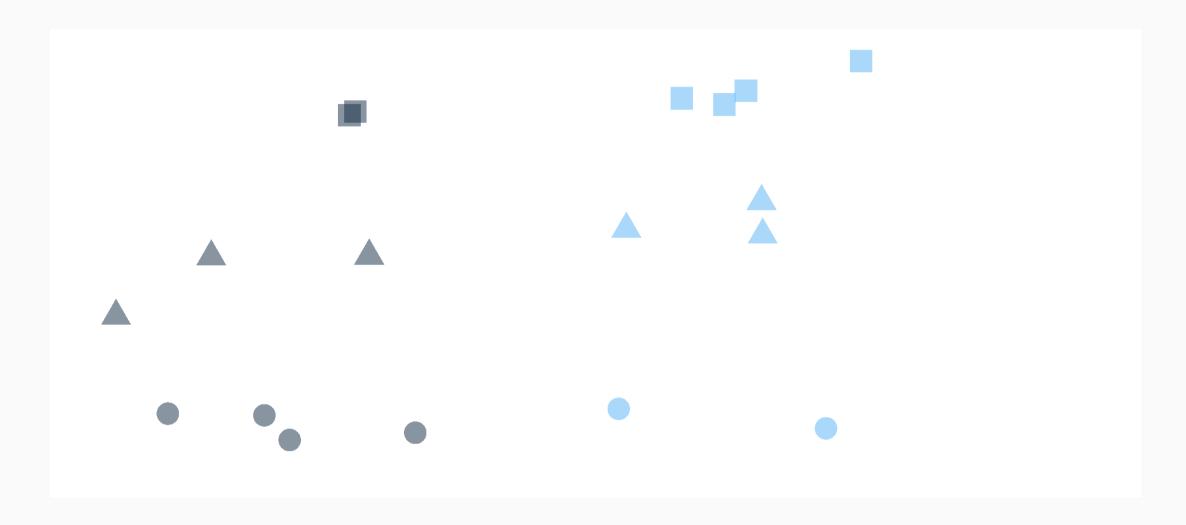


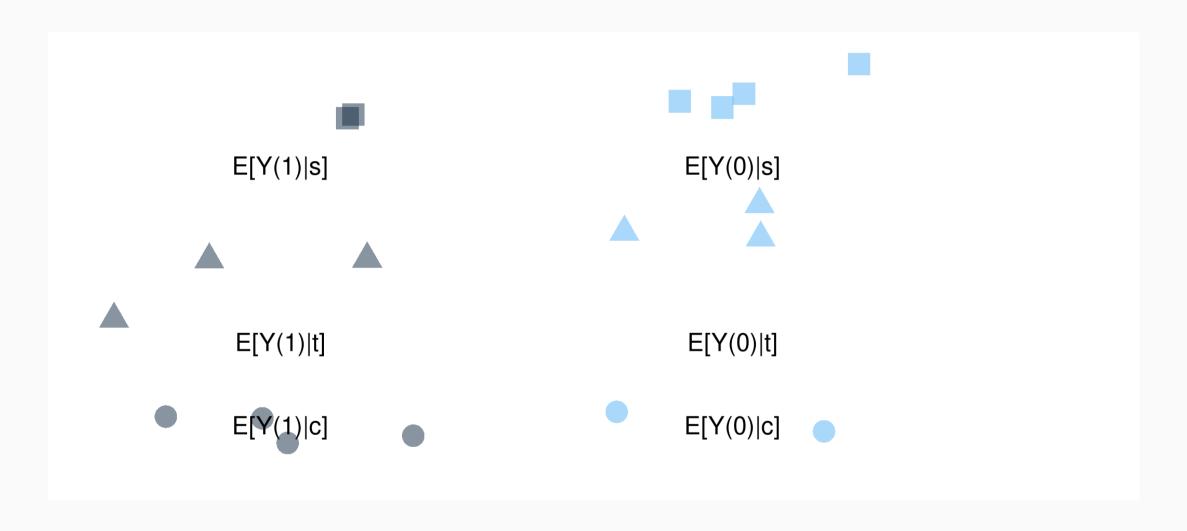
### Selection on Observables



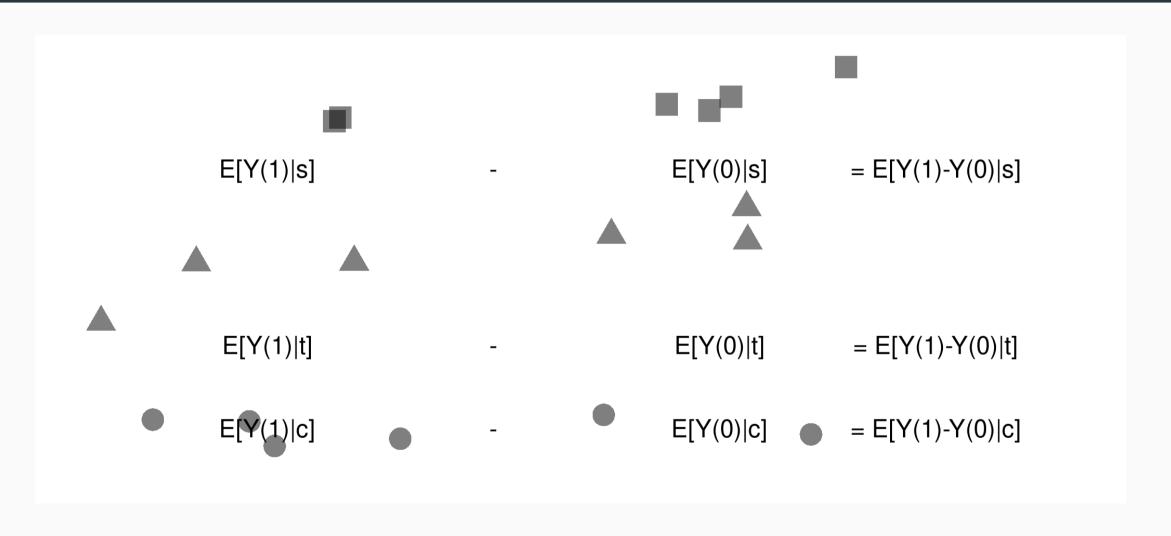




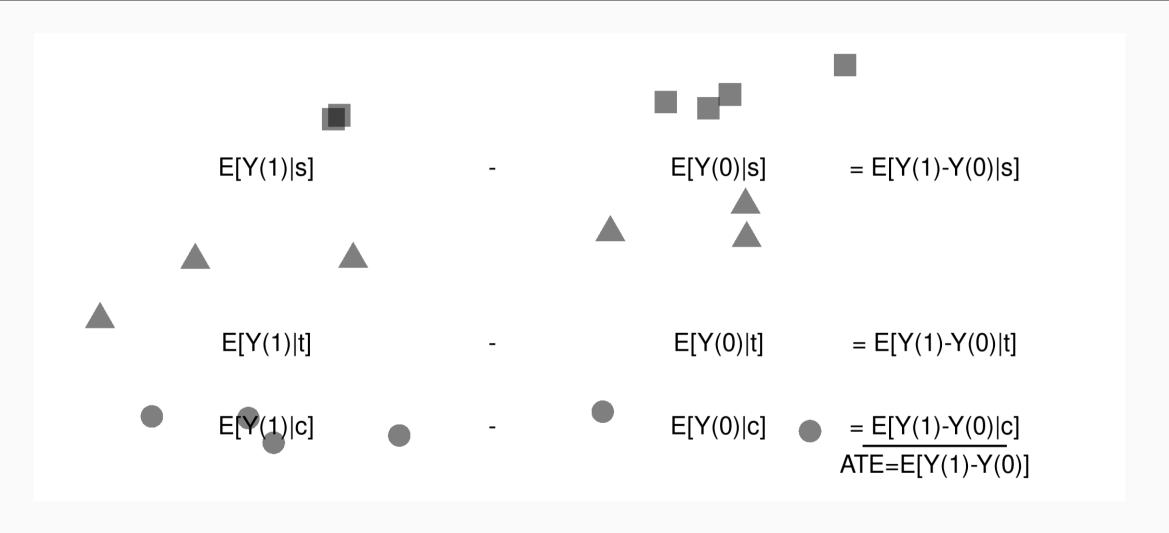




# Assumption 1: Selection on Observables



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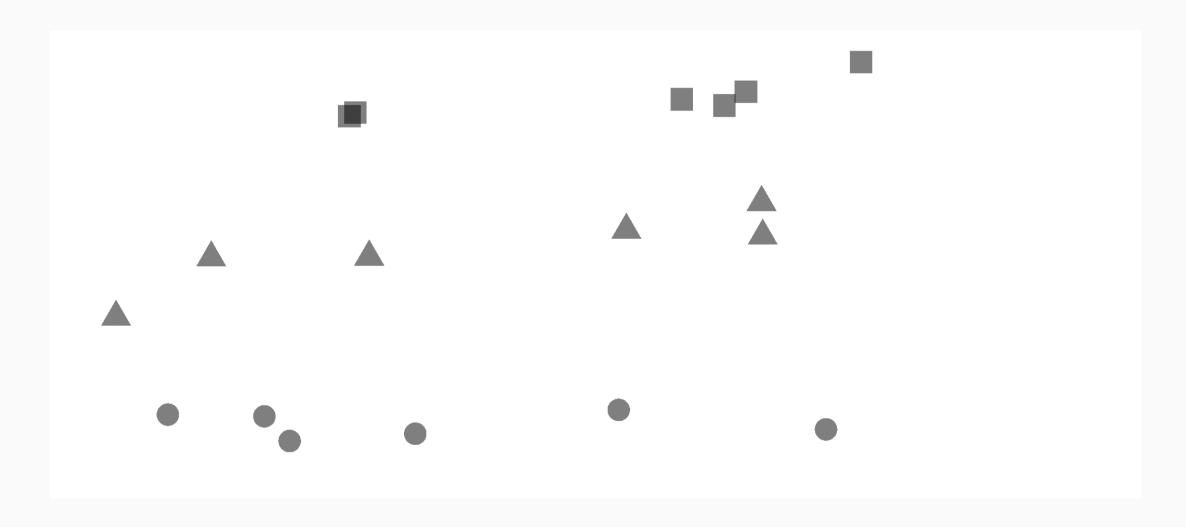
# Assumption 1: Selection on Observables

More formally:

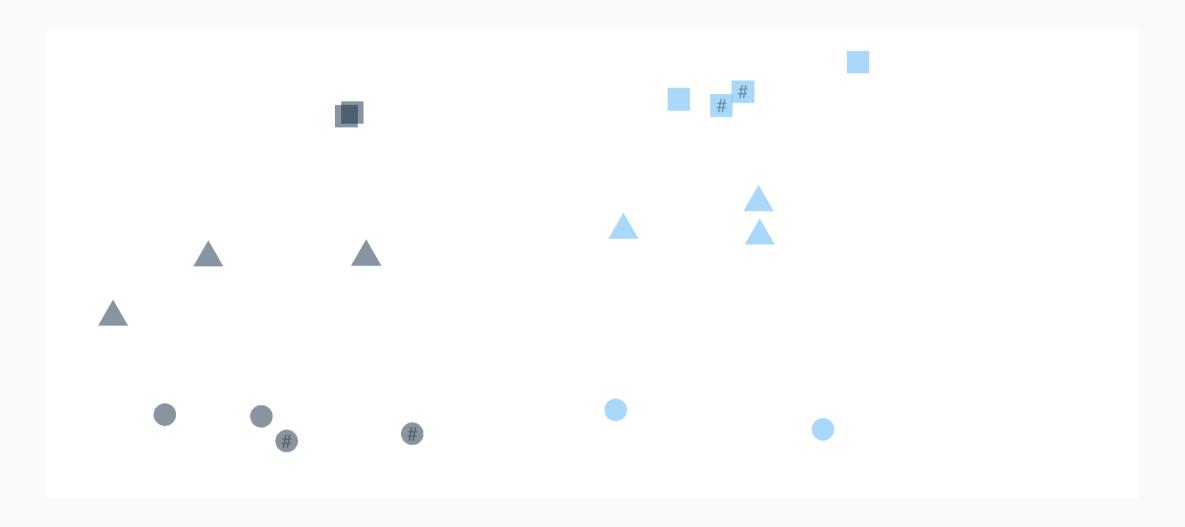
- E[Y(1)|W, shape] = E[Y(1)|shape]
- $Y(1), Y(0) \perp \!\!\! \perp W | shape$

In words...nothing unobserved that determines treatment selection and affects your outcome of interest.

### Violation of Selection on Observables



#### Violation of Selection on Observables

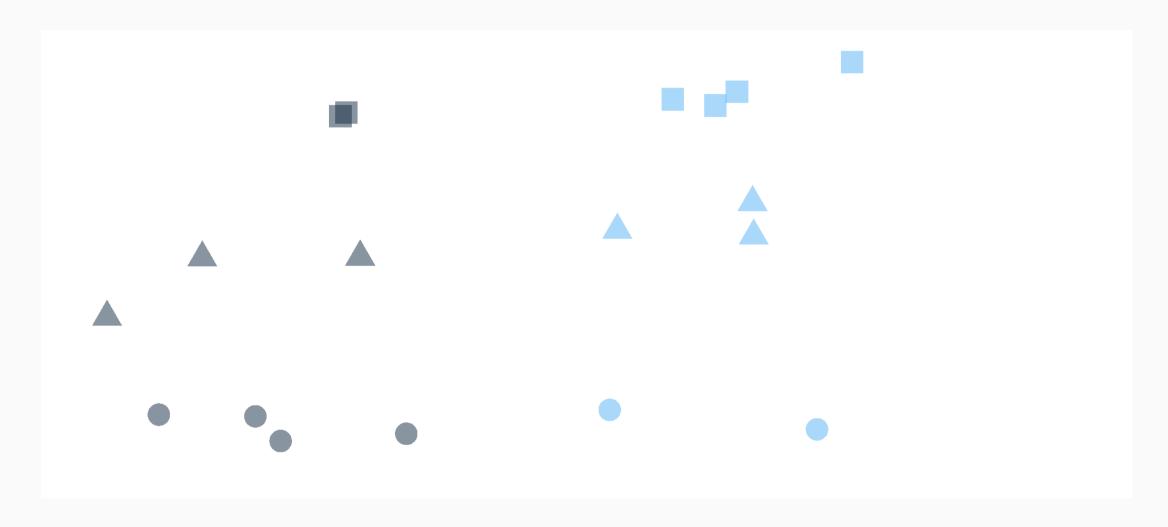


### **Assumption 2: Common Support**

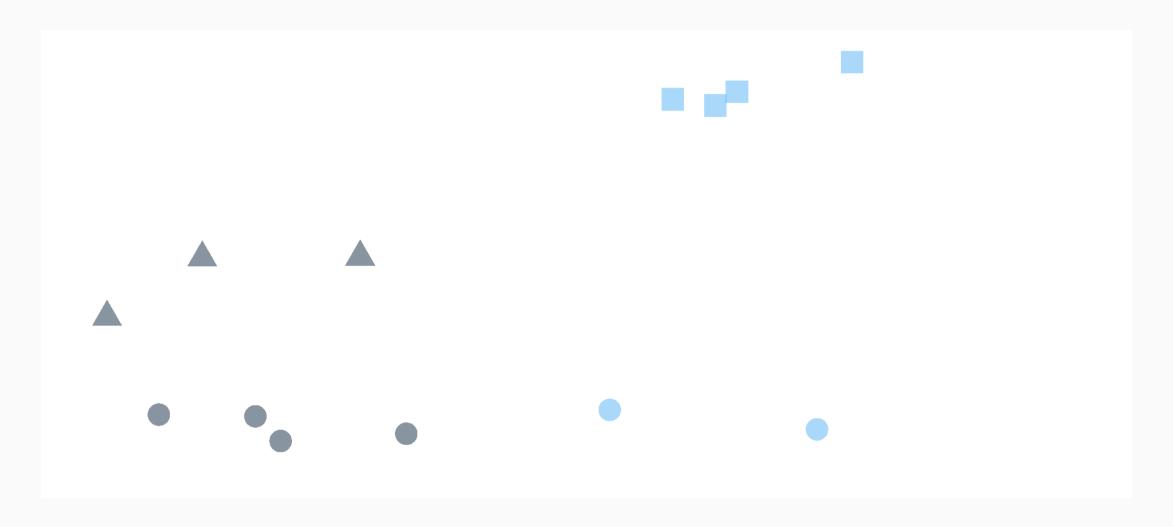
Someone of each type must be in both the treated and untreated groups

$$0<\Pr(W=1|X)<1$$

# Assumption 2: Common Support



# Violation of Common Support



#### Causal Inference with Observational Data

#### Fundamental Problem of Causal Inference

Potential outcomes:

- ullet  $Y_i(1)$  for  $W_i=1$
- $Y_i(0)$  for  $W_i=0$

Average treatment effect is  $E[Y_i(1)-Y_i(0)]$ , but:

- $E[Y_i|W_i=1] 
  eq E[Y_i(1)]$
- $E[Y_i|W_i=0] 
  eq E[Y_i(0)]$

#### Fundamental Problem of Causal Inference

- We don't observe the counterfactual outcome...what would have happened if a treated unit was actually untreated.
- ALL attempts at causal inference represent some attempt at estimating the counterfactual outcome. We need an estimate for  $E[Y_i(0)]$  among those that were treated, and vice versa for  $E[Y_i(1)]$ .

#### Causal inference with observational data

Solution for now: find covariates  $X_i$  such that the following assumptions are plausible:

1. Selection on observables:

$$Y_i(1), Y_i(0) \perp \!\!\! \perp W_i | X_i$$

2. Common support:

$$0<\Pr(W_i=1|X_i)<1$$

#### Causal inference with observational data

With selection on observables and common support:

1. Matching estimators:

$$E[Y_i(1) - Y_i(0)] = E[E[Y_i|W_i = 1, X_i] - E[Y_i|W_i = 0, X_i]]$$

2. Regression estimators:

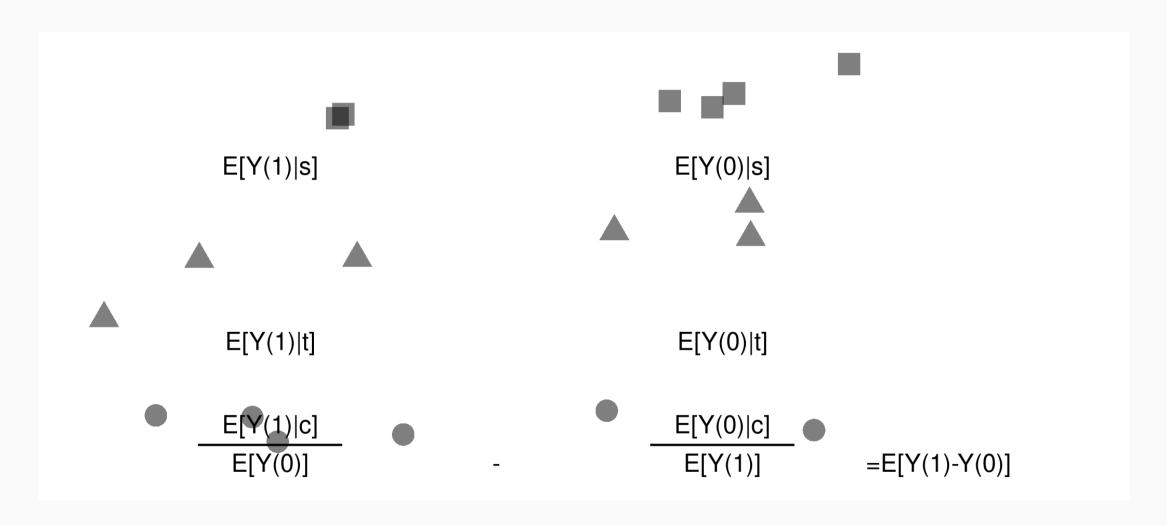
$$E[Y_i(1) - Y_i(0)] = E[E[Y_i|W_i = 1, X_i]] - E[E[Y_i|W_i = 0, X_i]]$$

What's the difference?

# Matching



# Regression



# **Estimation options**

- Matching
- Weighting
- Regression
- Doubly-robust weighting + regression (won't cover)

## Matching: The process

- 1. For each observation i, find the m "nearest" neighbors,  $J_m(i)$ .
- 2. Impute  $\hat{Y}_i(0)$  and  $\hat{Y}_i(1)$  for each observation:

$$\hat{Y_i}(0) = \left\{ egin{array}{ll} Y_i & ext{if} & W_i = 0 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 1 \end{array} 
ight.$$

$$\hat{Y}_i(1) = \left\{ egin{array}{ll} Y_i & ext{if} & W_i = 1 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 0 \end{array} 
ight.$$

3. Form "matched" ATE:

$$\hat{\delta}^{ ext{match}} = rac{1}{N} \sum_{i=1}^{N} \left( \hat{Y}_i(1) - \hat{Y}_i(0) 
ight)$$

# Matching: Defining "nearest"

1. Euclidean distance:

$$\sum_{k=1}^K (X_{ik}-X_{jk})^2$$

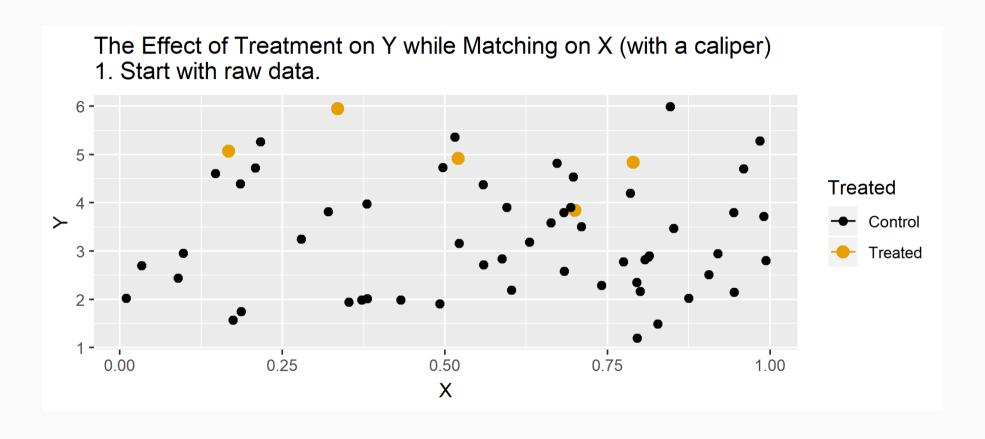
2. Scaled Euclidean distance:

$$\sum_{k=1}^K rac{1}{\sigma_{X_k}^2} (X_{ik} - X_{jk})^2$$

3. Mahalanobis distance:

$$(X_i-X_j)'\Sigma_X^{-1}(X_i-X_j)$$

# Animation for matching



# Weighting

- 1. Estimate propensity score ps  $\leftarrow$  glm(W~X, family=binomial, data), denoted  $\hat{\pi}(X_i)$
- 2. Weight by inverse of propensity score

$$\hat{\mu}_1 = rac{\sum_{i=1}^N rac{Y_i W_i}{\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{W_i}{\hat{\pi}(X_i)}}$$
 and  $\hat{\mu}_0 = rac{\sum_{i=1}^N rac{Y_i (1-W_i)}{1-\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{1-W_i}{1-\hat{\pi}(X_i)}}$ 

3. Form "inverse-propensity weighted" ATE:

$$\hat{\delta}^{IPW} = \hat{\mu}_1 - \hat{\mu}_0$$

## Regression

- 1. Regress  $Y_i$  on  $X_i$  among  $W_i=1$  to form  $\hat{\mu}_1(X_i)$
- 2. Regress  $Y_i$  on  $X_i$  among  $W_i=0$  to form  $\hat{\mu}_0(X_i)$
- 3. Form difference in predictions:

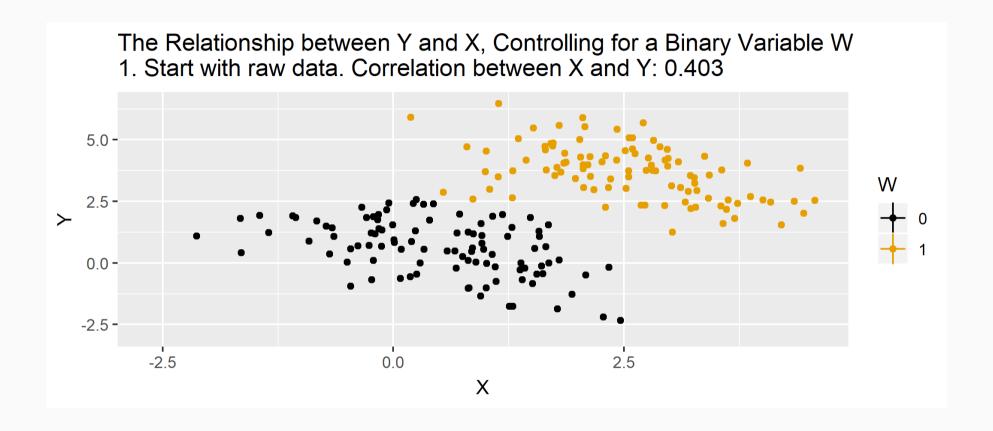
$$\hat{\delta}^{reg} = rac{1}{N} \sum_{i=1}^N \left(\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)
ight)$$

### Regression

Or estimate in one step,

$$Y_i = \delta W_i + eta X_i + W_i imes \left(X_i - ar{X}
ight) \gamma + arepsilon_i$$

# Animation for regression



#### Simulated data

Now let's do some matching, re-weighting, and regression with simulated data:

```
n \leftarrow 5000

select.dat \leftarrow tibble(

x = runif(n, 0, 1),

z = rnorm(n, 0, 1),

w = (x>0.65),

y = -2.5 + 4*w + 1.5*x + rnorm(n,0,1),

w_alt = (x + z > 0.35),

y_alt = -2.5 + 4*w_alt + 1.5*x + 2.25*z + rnorm(n,0,1)
```

## Simulation: nearest neighbor matching

## Matched number of observations.....

## Matched number of observations (unweighted). 5016

```
nn.est1 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=1,
                           estimand="ATE")
summary(nn.est1)
## Estimate ... 4.0175
## AI SE..... 0.52954
## T-stat..... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                              5000
## Original number of treated obs.....
                                              1732
```

5000

## Simulation: nearest neighbor matching

```
nn.est2 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=2,
                          estimand="ATE")
summary(nn.est2)
## Estimate ... 4.0175
## AI SE.... 0.52954
## T-stat.... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                              5000
## Original number of treated obs.....
                                              1732
## Matched number of observations.....
                                              5000
## Matched number of observations (unweighted). 5016
```

## Simulation: regression

```
reg1.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=1)}
reg1 \( \times \text{ lm(y \( \times \text{ x, data=reg1.dat)}} \)

reg0.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=0)}
reg0 \( \times \text{ lm(y \( \times \text{ x, data=reg0.dat)}} \)
pred1 \( \times \text{ predict(reg1,new=select.dat)} \)
pred0 \( \times \text{ predict(reg0,new=select.dat)} \)
mean(pred1-pred0)
```

## [1] 4.076999

#### Violation of selection on observables

#### NN Matching

```
##
## Estimate ... 7.6642
## AI SE.... 0.052903
## T-stat.... 144.87
## p.val.... < 2.22e-16
##
## Original number of observations .... 5000
## Original number of observations .... 2748
## Matched number of observations (unweighted). 23014</pre>
```

#### <u>Regression</u>

```
reg1.dat \leftarrow select.dat %>% filter(w_alt=1)
reg1 \leftarrow lm(y_alt \simeq x, data=reg1.dat)

reg0.dat \leftarrow select.dat %>% filter(w_alt=0)
reg0 \leftarrow lm(y_alt \simeq x, data=reg0.dat)
pred1_alt \leftarrow predict(reg1,new=select.dat)
pred0_alt \leftarrow predict(reg0,new=select.dat)
mean(pred1_alt-pred0_alt)
```

**##** [1] 7.646532

# Pricing and Hospital Profit Status

#### Penalized hospitals

```
final.hcris ← hcris.data %>% ungroup() %>%
  filter(price_denom>100, !is.na(price_denom),
        price_num>0, !is.na(price_num),
        price<100000,
        beds>30, year=2012) %>%
mutate( hvbp_payment = ifelse(is.na(hvbp_payment),0,hvbp_payment),
        hrrp_payment = ifelse(is.na(hrrp_payment),0,abs(hrrp_payment)),
        penalty = (hvbp_payment-hrrp_payment<0))</pre>
```

#### Summary stats

Always important to look at your data before doing any formal analysis. Ask yourself a few questions:

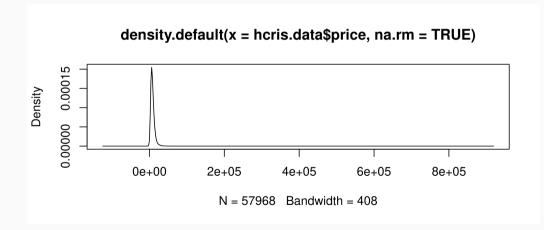
- 1. Are the magnitudes reasonable?
- 2. Are there lots of missing values?
- 3. Are there clear examples of misreporting?

#### **Summary stats**

```
summary(hcris.data$price)

### Min. 1st Qu. Median Mean 3rd Qu. Max. NA's
## -123697 4783 7113 Inf 10230 Inf 63662

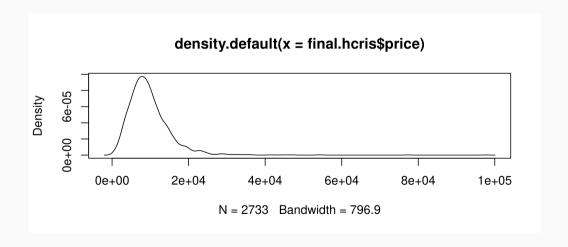
plot(density(hcris.data$price, na.rm=TRUE))
```



```
summary(final.hcris$price)

## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 340.8 6129.9 8705.4 9646.9 11905.4 97688.8

plot(density(final.hcris$price))
```



# Dealing with problems

We've adopted a very brute force way to deal with outlier prices. Other approaches include:

- 1. Investigate very closely the hospitals with extreme values
- 2. Winsorize at certain thresholds (replace extreme values with pre-determined thresholds)
- 3. Impute prices for extreme hospitals

# Differences among penalized hospitals

- Mean price among penalized hospitals: 9,896.31
- Mean price among non-penalized hospitals: 9,560.41
- Mean difference: 335.9

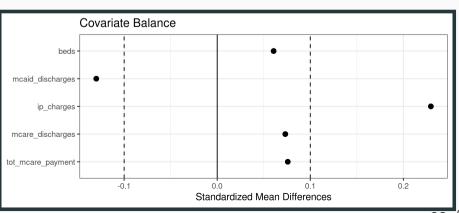
## Comparison of hospitals

Are penalized hospitals sufficiently similar to non-penalized hospitals?

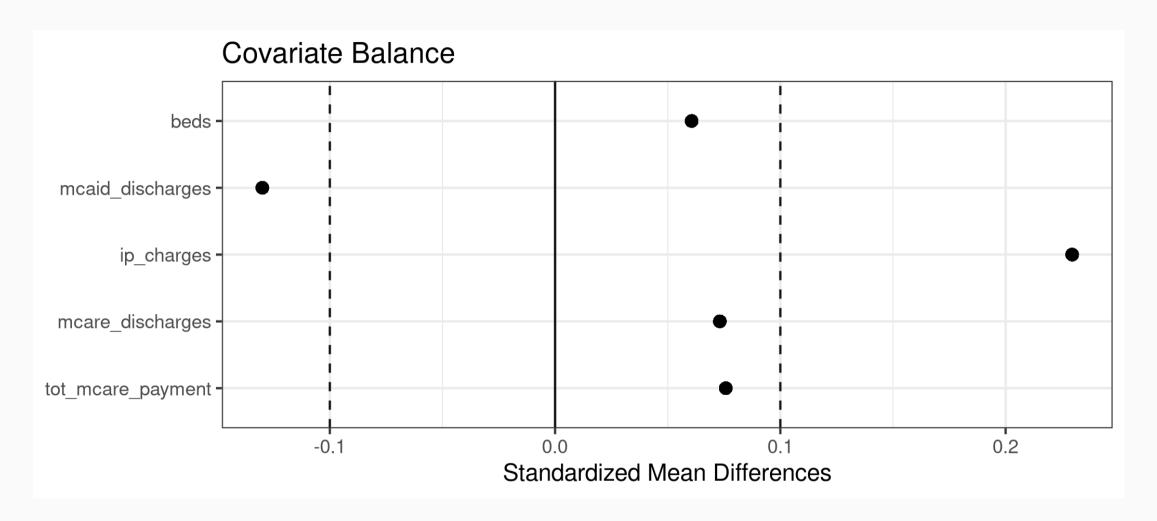
Let's look at covariate balance using a love plot, part of the library(cobalt) package.

## Love plots without adjustment

```
love.plot(bal.tab(lp.covs,treat=lp.vars$penalty), colors="black", shapes="circle", threshold=0.1) +
theme bw() + theme(legend.position="none")
```



# Love plots without adjustment



### Using matching to improve balance

Some things to think about:

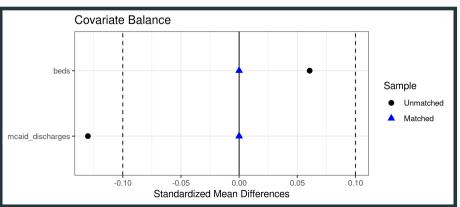
- exact versus nearest neighbor
- with or without ties (and how to break ties)
- measure of distance

## 1. Exact Matching

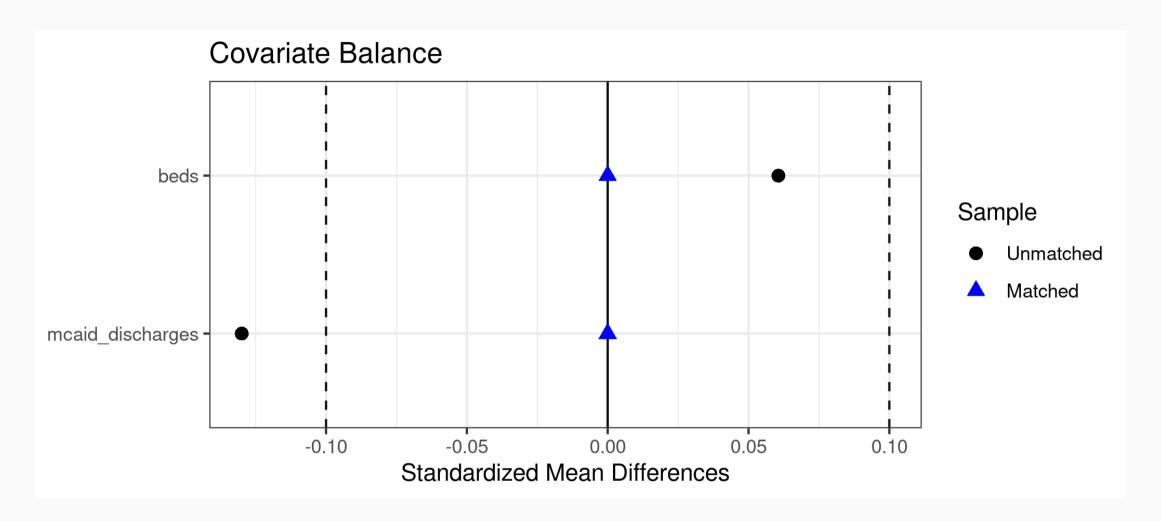
## attr(,"class")
## [1] "Match"

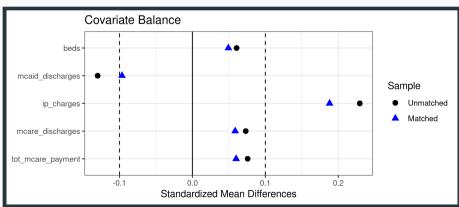
# 1. Exact Matching (on a subset)

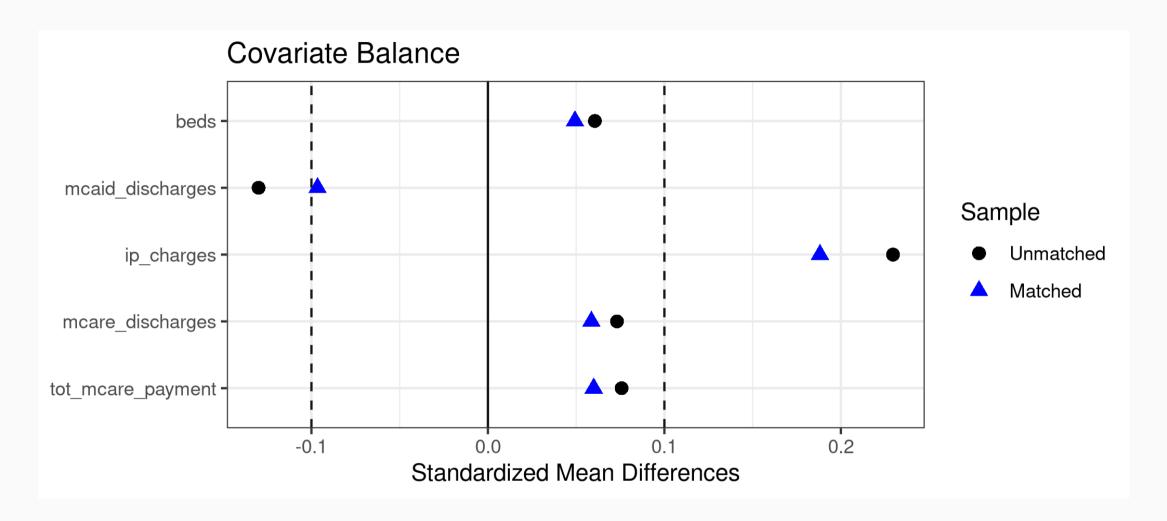
## 1. Exact Matching (on a subset)

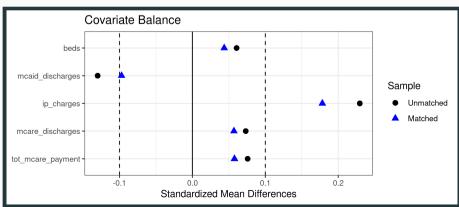


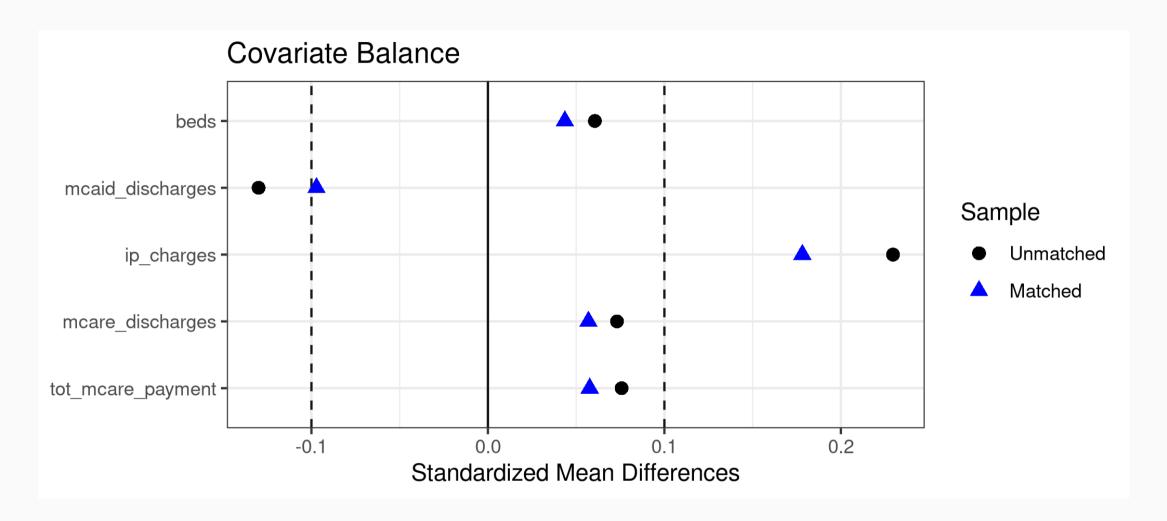
## 1. Exact Matching (on a subset)





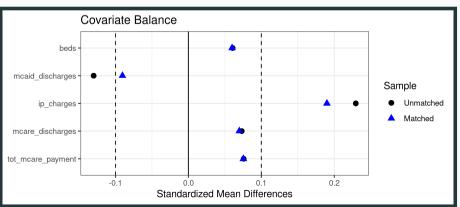




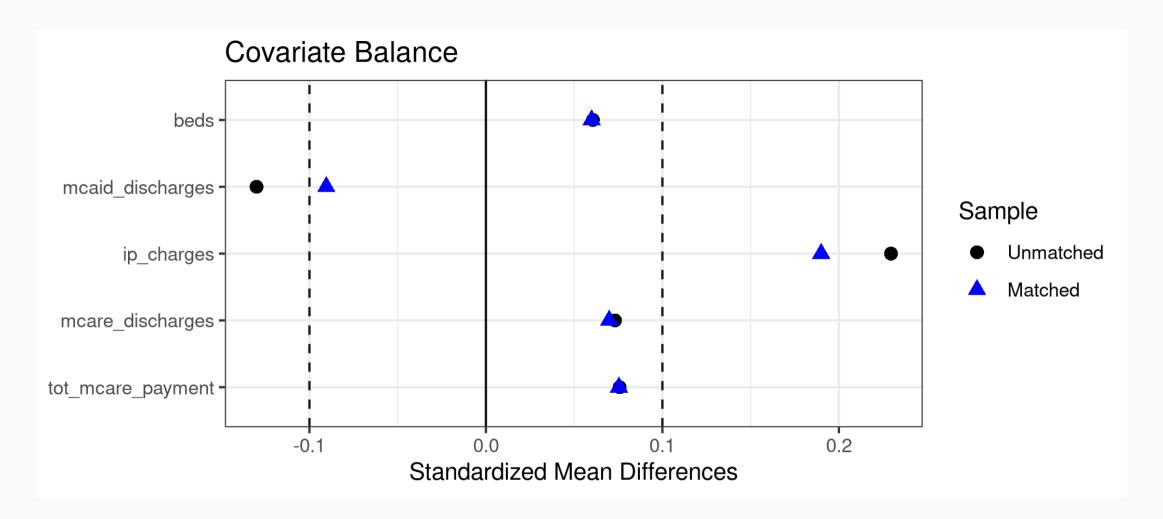


## 2. Nearest neighbor matching (Mahalanobis)

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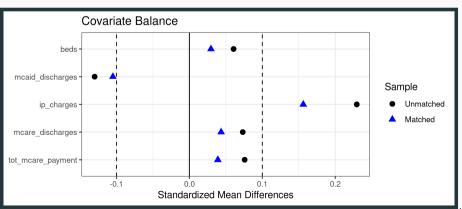


## 2. Nearest neighbor matching (Mahalanobis)

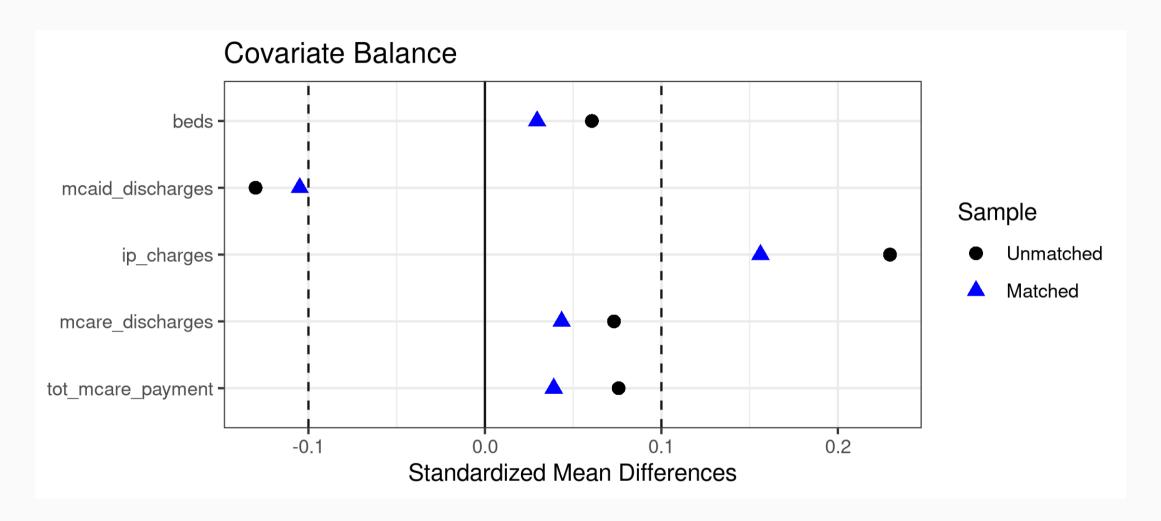


### 2. Nearest neighbor matching (propensity score)

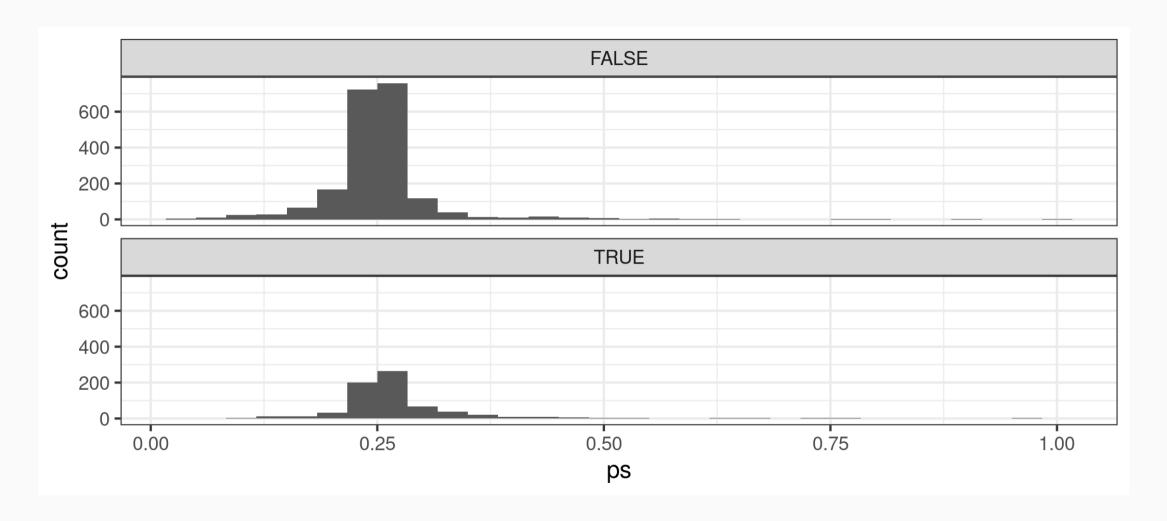
## 2. Nearest neighbor matching (propensity score)



## 2. Nearest neighbor matching (propensity score)



# 3. Weighting



### Results: Exact matching

```
##
## Estimate... 1777.6
## AI SE.... 34.725
## T-stat.... 51.191
## p.val.... < 2.22e-16
##
## Original number of observations.... 2707
## Original number of treated obs.... 698
## Matched number of observations (unweighted). 12
## Matched number of observations (unweighted). 12
##
## Number of obs dropped by 'exact' or 'caliper' 2695</pre>
```

### Results: Nearest neighbor

#### • Inverse variance

```
##
## Estimate... -526.95
## AI SE..... 223.06
## T-stat.... -2.3623
## p.val.... 0.01816
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2711
```

### Results: Nearest neighbor

#### Mahalanobis

```
##
## Estimate... -492.82
## AI SE..... 223.55
## T-stat.... -2.2046
## p.val.... 0.027485
##
## Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 2708
```

### Results: Nearest neighbor

#### Propensity score

```
##
## Estimate... -201.03
## AI SE..... 275.76
## T-stat.... -0.72898
## p.val..... 0.46601
##
##
Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 14795
```

### Results: IPW weighting

```
lp.vars \leftarrow lp.vars %>%
mutate(ipw = case_when(
    penalty=1 ~ 1/ps,
    penalty=0 ~ 1/(1-ps),
    TRUE ~ NA_real_
))
mean.t1 \leftarrow lp.vars %>% filter(penalty=1) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t0 \leftarrow lp.vars %>% filter(penalty=0) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t1$mean_p - mean.t0$mean_p
```

```
## [1] -196.8922
```

### Results: IPW weighting with regression

```
ipw.reg ← lm(price ~ penalty, data=lp.vars, weights=ipw)
summarv(ipw.reg)
##
## Call:
### lm(formula = price ~ penalty, data = lp.vars, weights = ipw)
##
## Weighted Residuals:
     Min
          1Q Median
                       3Q
                               Max
## -18691 -4802 -1422 2651 94137
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 9876.4 147.8 66.808 <2e-16 ***
## penaltyTRUE -196.9 211.2 -0.932 0.351
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7829 on 2705 degrees of freedom
## Multiple R-squared: 0.0003211, Adjusted R-squared: -4.85e-05
## F-statistic: 0.8688 on 1 and 2705 DF, p-value: 0.3514
```

### Results: Regression

```
## [1] -5.845761
```

### Results: Regression in one step

### Results: Regression in one step

```
###
## Call:
## lm(formula = price ~ penalty + beds + mcaid discharges + ip charges +
      mcare discharges + tot mcare payment + beds diff + mcaid diff +
###
      ip diff + mcare diff + mpay diff, data = reg.dat)
###
##
## Residuals:
     Min
             10 Median
                          3Q
                               Max
## -38175 -2900
                 -597
                        2105 67409
##
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept) 8.466e+03 1.711e+02 49.482 < 2e-16 ***
## penaltyTRUE
              -5.846e+00 2.124e+02 -0.028 0.97804
## beds
                  1.107e+00 1.421e+00 0.779 0.43618
## mcaid discharges -4.714e-01 7.296e-02 -6.462 1.23e-10 ***
## ip charges
                    6.426e-06 1.285e-06 5.002 6.04e-07 ***
## mcare discharges -8.122e-01 9.257e-02 -8.774 < 2e-16 ***
                                        13.857 < 2e-16 ***
## tot_mcare_payment 9.502e-05
                              6.858e-06
## beds diff
                    2.517e+00 2.986e+00
                                        0.843 0.39931
## mcaid diff
             1.058e-01 1.570e-01
                                        0.674 0.50050
## ip_diff
                   -4.534e-06 2.027e-06 -2.237 0.02539 *
                                        2.657 0.00793 **
## mcare diff
             4.806e-01 1.809e-01
## mpay diff
                   -5.452e-05 1.321e-05 -4.128 3.78e-05 ***
## ---
```

### **Summary of ATEs**

- 1. Exact matching: 1777.63
- 2. NN matching, inverse variance: -526.95
- 3. NN matching, mahalanobis: -492.82
- 4. NN matching, pscore: -201.03
- 5. Inverse pscore weighting: -196.89
- 6. IPW regression: -196.89
- 7. Regression: -5.85
- 8. Regression 1-step: -5.85

### So what have we learned?

### Key assumptions for causal inference

- 1. Selection on observables
- 2. Common support

These become more nuanced but the intuition is the same in almost all questions of causal inference.

### Causal effect assuming selection on observables

If we assume selection on observables holds, then we only need to condition on the relevant covariates to identify a causal effect. But we still need to ensure common support...

- 1. Matching
- 2. Reweighting
- 3. Regression