

Section 1: Hospital Pricing and Selection on Observables

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Background on Hospital Pricing

Defining characteristic of hospital services: it's complicated!

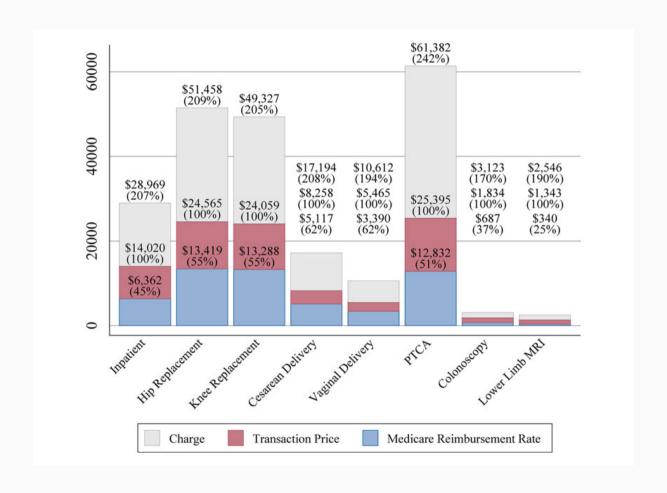
	-			
1/04/11	1	041000	LEARLINK DUO-VENT	17.00
1/04/11	1	0406462	TUBE COMNECTING STERIL 6FT	27.00
1/05/11	Ť	0406462	TUBE CONNECTING STERIL OFT	27.00
1/05/11	1	3005741	ACCU-CHEK CCRV	18.00
1/05/11	1	3019692	SURGICEL 2X14 STRIP EACH	451.00
1/05/11	-	3005741	ACCU-CHEK CCRV	
1/05/11	-	3005741	ACCU-CHEK CCRV	18.00
1/05/11	1-	3019692	SURGICEL 2X14 STRIP EACH	18.00_
	1	3005741	ACCU-CHEK CCRV	451.00-
1/05/11	1	3005741	ACCU-CREK CCRV	18.00
1/05/11	10	2900025	OKYGEN HOURLY	18.00
1/05/11	1	0402230	LEUKINS TUEE SPECIM TRAP	560.00
1/05/11	1	0416826	SET EXTENSION 1-VALVE	77.00
1/05/11	1	0406793	SUCTION YANKAUER	12.00
1/05/11	_ 1	0416018	SECOND DEEL SET LUER LOCK	44.00
		042-64	DUCK LOCK	5.00

Brill, Steven. 2013. "Bitter Pill: Why Medical Bills are Killing Us." *Time Magazine*.

Lots of different payers paying lots of different prices:

- Medicare fee-for-service prices
- Medicaid payments
- Private insurance negotiations (including Medicare Advantage)
- But what about the price to patients?

Price \neq charge \neq cost \neq patient out-of-pocket spending



Source: Health Care Pricing Project

Not clear what exactly is negotiated...

Fee-for-service

- price per procedure
- percentage of charges
- markup over Medicare rates

Capitation

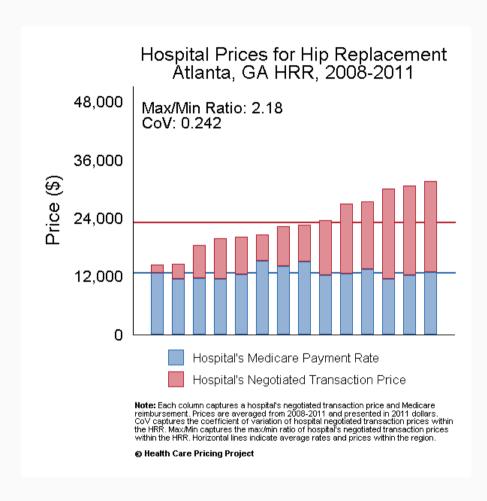
- payment per patient
- pay-for-performance
- shared savings

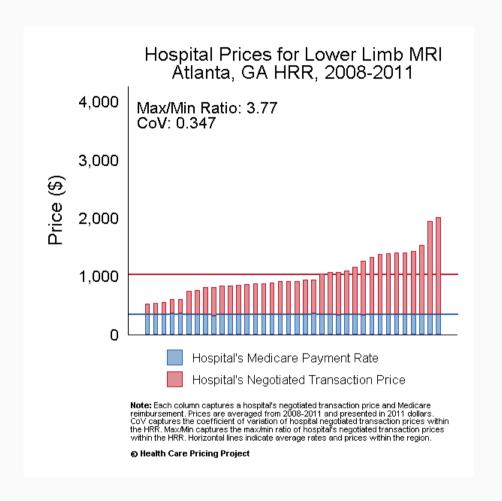
Hospital prices in real life

A few empirical facts:

- 1. Hospital services are expensive
- 2. Prices vary dramatically across different areas
- 3. Lack of competition is a major reason for high prices

Hospital prices in real life





Source: Health Care Pricing Project

Understanding HCRIS Data

What is HCRIS?

Healthcare Cost Report Information System ('cost reports')

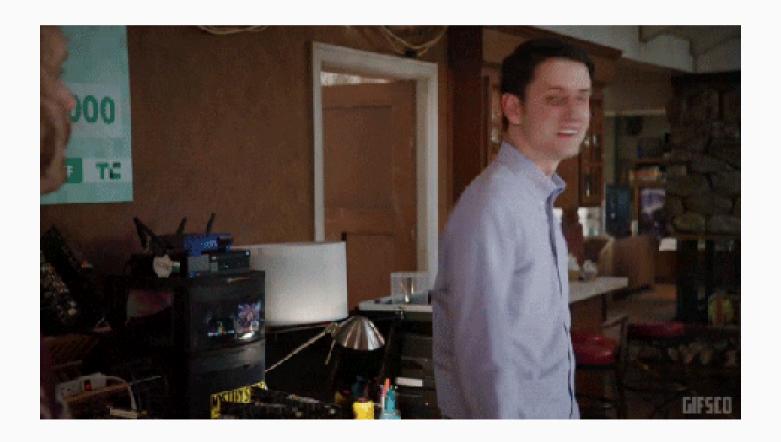
- Nursing Homes (SNFs)
- Hospice
- Home Health Agencies
- Hospitals

Hospital Cost Reports

10-1	2 FORM	FORM CMS-2552-10		4090 (Cont.) WORKSHEET G-2, PARTS I & II	
STATEMENT OF PATIENT REVENUES AND OPERATING EXPENSES		PROVIDER CCN:	PERIOD: FROM TO _		
PART	I - PATIENT REVENUES				
	REVENUE CENTER	INPATIENT 1	OUTPATIENT 2	TOTAL 3	
	GENERAL INPATIENT ROUTINE CARE SERVICES				
1	Hospital				1
2	Subprovider IPF				2
3	Subprovider IRF				3
4	Subprovider (Other)				4
5	Swing bed - SNF				5
6	Swing bed - NF				6
7	Skilled nursing facility				7
8	Nursing facility				8
9	Other long term care				9
10	Total general inpatient care services (sum of lines 1-9)				10
	INTENSIVE CARE TYPE INPATIENT HOSPITAL SERVICES				
11	Intensive care unit				11
12	Coronary care unit				12
13	Burn intensive care unit			1	13
14	Surgical intensive care unit				14
15					15
16	Total intensive care type inpatient hospital services (sum of				16
	of lines 11-15)				
17	Total inpatient routine care services (sum of lines 10 and 16)				17
18					18
19	Outpatient services				19
20	Rural Health Clinic (RHC)			2	20
21	Federally Qualified Health Center (FQHC)			2	21
22	Home health agency				22

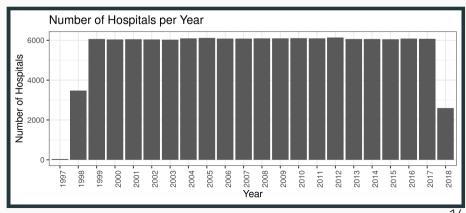
The Data

Let's work with the HCRIS GitHub repository. But forming the dataset is up to you this time.

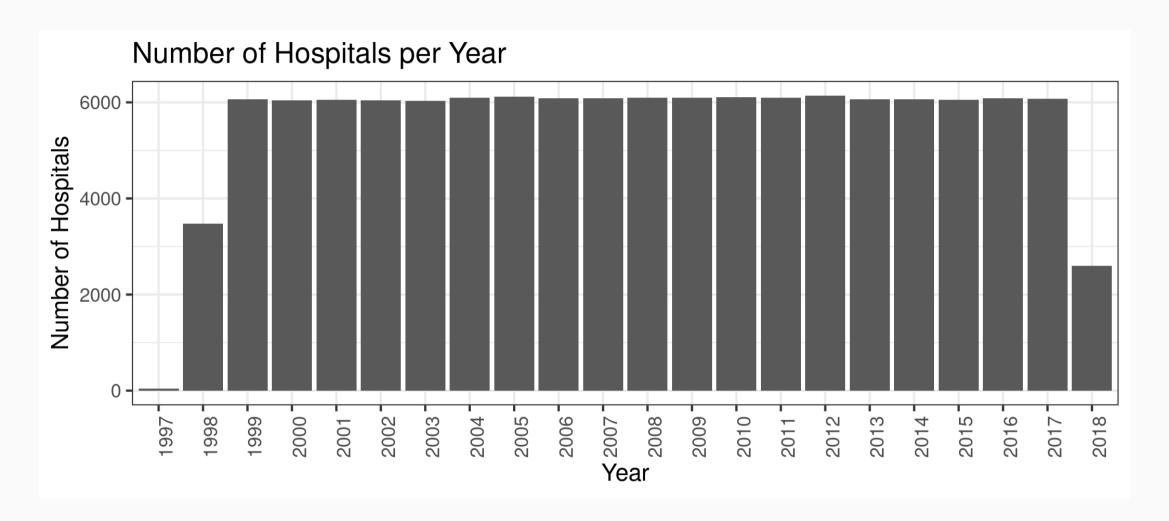


The Data

```
hcris.data %>%
  ggplot(aes(x=as.factor(year))) +
  geom_bar() +
  labs(
    x="Year",
    y="Number of Hospitals",
    title="Number of Hospitals per Year"
  ) + theme_bw() +
  theme(axis.text.x = element_text(angle = 90, hjust=1))
```

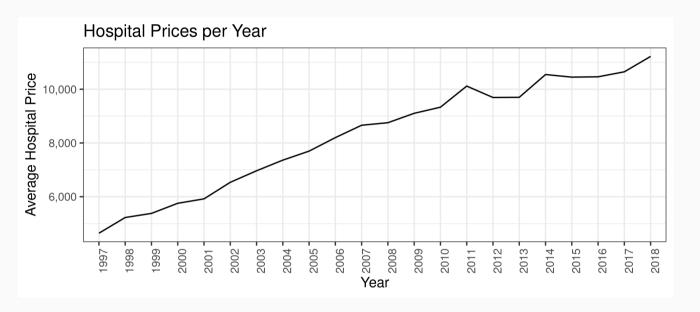


Number of hospitals

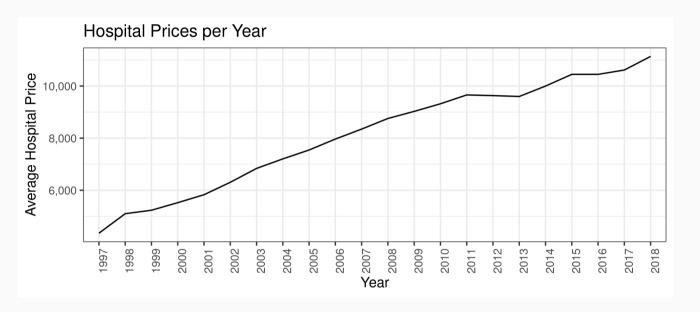


Estimating hospital prices

Estimating hospital prices



Estimating hospital prices



Causal Inference and Potential Outcomes

Why causal inference?

Another example: What price should we charge for a night in a hotel?

Machine Learning

- Focuses on prediction
- High prices are strongly correlated with higher sales
- Increase prices to attract more people?

Causal Inference

- Focuses on counterfactuals
- What would sales look like if prices were higher?

Goal of Causal Inference

- Goal: Estimate effect of some policy or program
- Key building block for causal inference is the idea of **potential outcomes**

Some notation

Observed outcome

$$Y_i = Y_{1i} imes D_i + Y_{0i} imes (1-D_i)$$

or

$$Y_i = \left\{ egin{aligned} Y_{1i} ext{ if } D_i = 1 \ Y_{0i} ext{ if } D_i = 0 \end{aligned}
ight.$$

Assumes **SUTVA** (stable unit treatment value assumption)...no interference across units

Example of "Potential Outcomes"





$$Y_0$$
= \$60,000

Earnings due to Emory = $Y_1 - Y_0$ = \$15,000

Example of "Potential Outcomes"



Earnings due to Emory = $Y_1 - Y_0$ = ?



$$Y_0$$
= ?

Do we ever observe the potential outcomes?



Without a time machine...not possible to get individual effects.

Fundamental Problem of Causal Inference

- We don't observe the counterfactual outcome...what would have happened if a treated unit was actually untreated.
- ALL attempts at causal inference represent some attempt at estimating the counterfactual outcome. We need an estimate for Y_0 among those that were treated, and vice versa for Y_1 .

Average Treatment Effects

Different treatment effects

Tend to focus on **averages**¹:

• ATE:
$$\delta_{ATE} = E[Y_1 - Y_0]$$

$$ullet$$
 ATT: $\delta_{ATT}=E[Y_1-Y_0|D=1]$

• ATU:
$$\delta_{ATU}=E[Y_1-Y_0|D=0]$$

¹ or similar measures such as medians or quantiles

Average Treatment Effects

• Estimand:

$$\delta_{ATE} = E[Y_1 - Y_0] = E[Y|D=1] - E[Y|D=0]$$

• Estimate:

$$\hat{\delta}_{ATE} = rac{1}{N_1} \sum_{D_i=1} Y_i - rac{1}{N_0} \sum_{D_i=0} Y_i,$$

where N_1 is number of treated and N_0 is number untreated (control)

 With random assignment and equal groups, inference/hypothesis testing with standard two-sample t-test

Selection Bias

Selection bias

• Without random assignment, there's a high probability that

$$E[Y_{0i}|D_i=1]
eq E[Y_{0i}|D_i=0]$$

• i.e., outcomes without treatment are different for the treated group

Omitted variables bias

- In a regression setting, selection bias is the same problem as omitted variables bias (OVB)
- Quick review: Goal of OLS is to find \hat{eta} to "best fit" the linear equation $y_i=lpha+x_ieta+\epsilon_i$

Regression review

$$egin{aligned} \min_{eta} \sum_{i=1}^{N} \left(y_{i} - lpha - x_{i}eta
ight)^{2} &= \min_{eta} \sum_{i=1}^{N} \left(y_{i} - (ar{y} - ar{x}eta) - x_{i}eta
ight)^{2} \ 0 &= \sum_{i=1}^{N} \left(y_{i} - ar{y} - (x_{i} - ar{x})\hat{eta}
ight) (x_{i} - ar{x}) \ 0 &= \sum_{i=1}^{N} (y_{i} - ar{y})(x_{i} - ar{x}) - \hat{eta} \sum_{i=1}^{N} (x_{i} - ar{x})^{2} \ \hat{eta} &= \frac{\sum_{i=1}^{N} (y_{i} - ar{y})(x_{i} - ar{x})}{\sum_{i=1}^{N} (x_{i} - ar{x})^{2}} = \frac{Cov(y, x)}{Var(x)} \end{aligned}$$

Omitted variables bias

$$egin{aligned} \hat{eta} &= rac{Cov(Y_i, s_i)}{Var(s_i)} \ &= rac{Cov(lpha + eta s_i + \gamma A_i + \epsilon_i, s_i)}{Var(s_i)} \ &= rac{eta Cov(s_i, s_i) + \gamma Cov(A_i, s_i) + Cov(\epsilon_i, s_i)}{Var(s_i)} \ &= eta rac{Var(s_i)}{Var(s_i)} + \gamma rac{Cov(A_i, s_i)}{Var(s_i)} + 0 \ &= eta + \gamma imes heta_{as} \end{aligned}$$

Removing selection bias without RCT

- The field of causal inference is all about different strategies to remove selection bias
- The first strategy (really, assumption) in this class: **selection on observables** or **conditional indpendence**

Intuition

Some assumptions:

- $Y_{0i} = \alpha + \eta_i$
- $Y_{1i}-Y_{0i}=\delta$
- ullet There is some set of "controls", x_i , such that $\eta_i=eta x_i+u_i$ and $E[u_i|x_i]=0$ (conditional independence assumption, or CIA)

$$egin{aligned} Y_i &= Y_{1i} imes D_i + Y_{0i} imes (1 - D_i) \ &= \delta D_i + Y_{0i} D_i + Y_{0i} - Y_{0i} D_i \ &= \delta D_i + lpha + \eta_i \ &= \delta D_i + lpha + eta x_i + u_i \end{aligned}$$

ATEs vs regression coefficients

What does this mean?

- ullet OLS puts more weight on observations with treatment D_i "unexplained" by X_i
- "Reverse" weighting such that the proportion of treated units are used to weight the ATU while the proportion of untreated units enter the weights of the ATT
- This is an average effect, but probably not the average we want

Matching and Weighting

Goal

Find covariates X_i such that the following assumptions are plausible:

1. Selection on observables:

$$Y_{0i}, Y_{1i} \perp \!\!\!\perp D_i | X_i$$

2. Common support:

$$0<\Pr(D_i=1|X_i)<1$$

Then we can use X_i to group observations and use expectations for control as the predicted counterfactuals among treated, and vice versa.

Assumption 1: Selection on Observables

$$E[Y_1|D,X] = E[Y_1|X]$$

In words...nothing unobserved that determines treatment selection and affects your outcome of interest.

Assumption 2: Common Support

Someone of each type must be in both the treated and untreated groups

$$0<\Pr(D=1|X)<1$$

Causal inference with observational data

With selection on observables and common support:

- 1. Matching estimators
- 2. Regression estimators What's the difference?

Estimation options

- Matching
- Weighting
- Regression
- Doubly-robust weighting + regression (won't cover)

Matching: The process

- 1. For each observation i, find the m "nearest" neighbors, $J_m(i)$.
- 2. Impute $\hat{Y}_i(0)$ and $\hat{Y}_i(1)$ for each observation:

$$\hat{Y_i}(0) = \left\{egin{array}{ll} Y_i & ext{if} & W_i = 0 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 1 \end{array}
ight.$$

$$\hat{Y}_i(1) = \left\{ egin{array}{ll} Y_i & ext{if} & W_i = 1 \ rac{1}{m} \sum_{j \in J_m(i)} Y_j & ext{if} & W_i = 0 \end{array}
ight.$$

3. Form "matched" ATE:

$$\hat{\delta}^{ ext{match}} = rac{1}{N} \sum_{i=1}^{N} \left(\hat{Y}_i(1) - \hat{Y}_i(0)
ight)$$

Matching: Defining "nearest"

1. Euclidean distance:

$$\sum_{k=1}^K (X_{ik}-X_{jk})^2$$

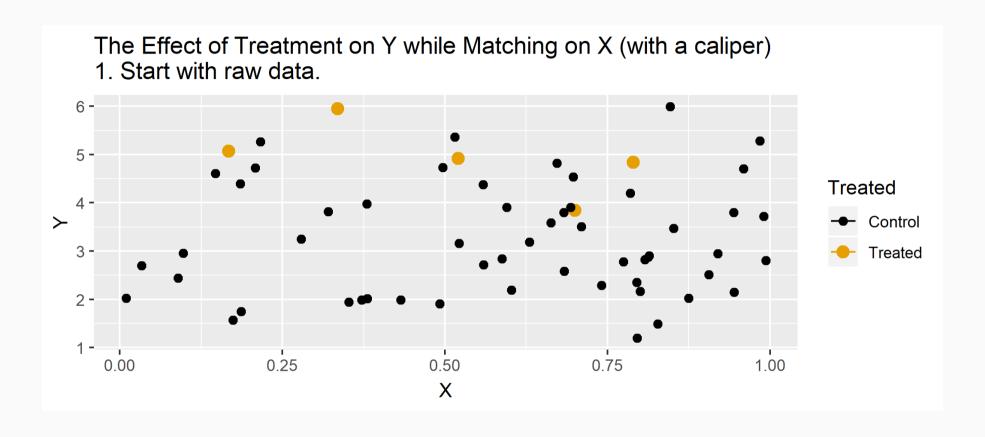
2. Scaled Euclidean distance:

$$\sum_{k=1}^K rac{1}{\sigma_{X_k}^2} (X_{ik} - X_{jk})^2$$

3. Mahalanobis distance:

$$(X_i-X_j)'\Sigma_X^{-1}(X_i-X_j)$$

Animation for matching



Weighting

- 1. Estimate propensity score ps \leftarrow glm(W~X, family=binomial, data), denoted $\hat{\pi}(X_i)$
- 2. Weight by inverse of propensity score

$$\hat{\mu}_1 = rac{\sum_{i=1}^N rac{Y_i W_i}{\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{W_i}{\hat{\pi}(X_i)}}$$
 and $\hat{\mu}_0 = rac{\sum_{i=1}^N rac{Y_i (1-W_i)}{1-\hat{\pi}(X_i)}}{\sum_{i=1}^N rac{1-W_i}{1-\hat{\pi}(X_i)}}$

3. Form "inverse-propensity weighted" ATE:

$$\hat{\delta}^{IPW} = \hat{\mu}_1 - \hat{\mu}_0$$

Regression

- 1. Regress Y_i on X_i among $W_i=1$ to form $\hat{\mu}_1(X_i)$
- 2. Regress Y_i on X_i among $W_i=0$ to form $\hat{\mu}_0(X_i)$
- 3. Form difference in predictions:

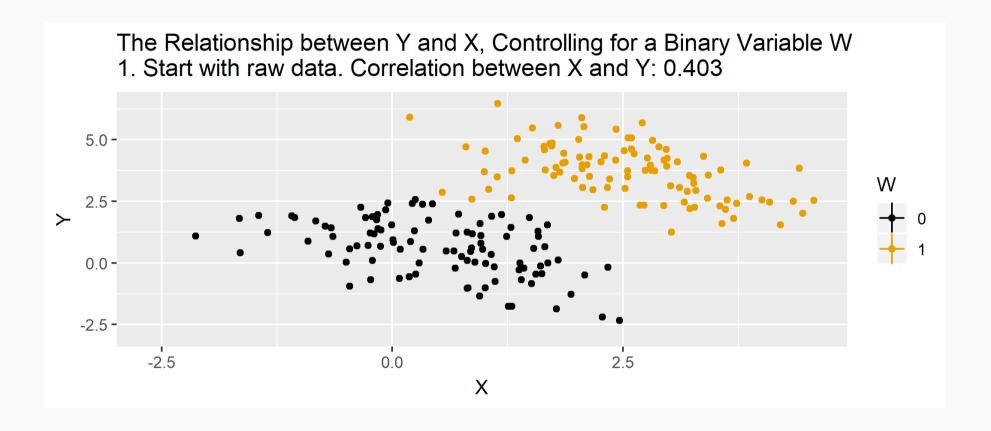
$$\hat{\delta}^{reg} = rac{1}{N} \sum_{i=1}^N \left(\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)
ight)$$

Regression

Or estimate in one step,

$$Y_i = \delta W_i + eta X_i + W_i imes \left(X_i - ar{X}
ight) \gamma + arepsilon_i$$

Animation for regression



Simulated data

Now let's do some matching, re-weighting, and regression with simulated data:

```
n \leftarrow 5000

select.dat \leftarrow tibble(

x = runif(n, 0, 1),

z = rnorm(n, 0, 1),

w = (x>0.65),

y = -2.5 + 4*w + 1.5*x + rnorm(n,0,1),

w_alt = (x + z > 0.35),

y_alt = -2.5 + 4*w_alt + 1.5*x + 2.25*z + rnorm(n,0,1)
```

Simulation: nearest neighbor matching

Matched number of observations.....

Matched number of observations (unweighted). 5016

```
nn.est1 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=1,
                          estimand="ATE")
summary(nn.est1)
## Estimate ... 4.0175
## AI SE.... 0.52954
## T-stat..... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                              5000
## Original number of treated obs.....
                                              1732
```

5000

Simulation: nearest neighbor matching

Matched number of observations (unweighted). 5016

```
nn.est2 ← Matching::Match(Y=select.dat$y,
                          Tr=select.dat$w.
                          X=select.dat$x,
                          M=1,
                          Weight=2,
                          estimand="ATE")
summary(nn.est2)
## Estimate ... 4.0175
## AI SE.... 0.52954
## T-stat..... 7.5869
## p.val..... 3.2863e-14
##
## Original number of observations.....
                                             5000
## Original number of treated obs.....
                                             1732
## Matched number of observations.....
                                             5000
```

Simulation: regression

```
reg1.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=1)}
reg1 \( \times \text{ lm(y \( \times \text{ x, data=reg1.dat)}} \)

reg0.dat \( \times \text{ select.dat } \%>\% \text{ filter(w=0)}
reg0 \( \times \text{ lm(y \( \times \text{ x, data=reg0.dat)}} \)
pred1 \( \times \text{ predict(reg1,new=select.dat)} \)
pred0 \( \times \text{ predict(reg0,new=select.dat)} \)
mean(pred1-pred0)
```

[1] 4.076999

Violation of selection on observables

NN Matching

```
##
## Estimate... 7.6642
## AI SE..... 0.052903
## T-stat.... 144.87
## p.val..... < 2.22e-16
##
## Original number of observations..... 5000
## Original number of observations..... 2748
## Matched number of observations (unweighted). 23014</pre>
```

Regression

```
reg1.dat \( \sepsilon \) select.dat \( \% \> \% \) filter(w_alt=1)
reg1 \( \subseteq \ln(y_alt \simpsilon x, \) data=reg1.dat)

reg0.dat \( \sepsilon \) select.dat \( \% \> \% \) filter(w_alt=0)
reg0 \( \subseteq \ln(y_alt \simpsilon x, \) data=reg0.dat)
pred1_alt \( \subseteq \) predict(reg1, new=select.dat)
pred0_alt \( \subseteq \) predict(reg0, new=select.dat)
mean(pred1_alt-pred0_alt)
```

[1] 7.646532

What covariates to use?

• There are such things as "bad controls"

Pricing and Hospital Profit Status

Penalized hospitals

Summary stats

Always important to look at your data before doing any formal analysis. Ask yourself a few questions:

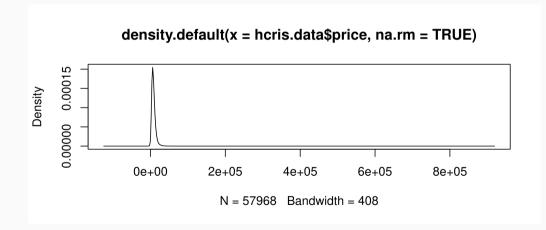
- 1. Are the magnitudes reasonable?
- 2. Are there lots of missing values?
- 3. Are there clear examples of misreporting?

Summary stats

```
summary(hcris.data$price)

## Min. 1st Qu. Median Mean 3rd Qu. Max. NA's
## -123697 4783 7113 Inf 10230 Inf 63662

plot(density(hcris.data$price, na.rm=TRUE))
```

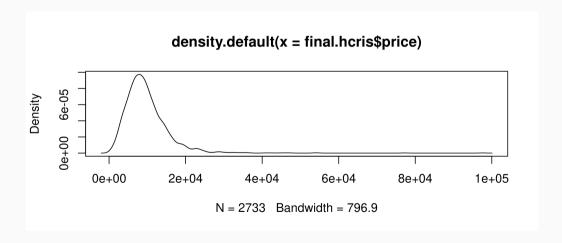


```
summary(final.hcris$price)

## Min. 1st Qu. Median Mean 3rd Qu. Max.

## 340.8 6129.9 8705.4 9646.9 11905.4 97688.8

plot(density(final.hcris$price))
```



Dealing with problems

We've adopted a very brute force way to deal with outlier prices. Other approaches include:

- 1. Investigate very closely the hospitals with extreme values
- 2. Winsorize at certain thresholds (replace extreme values with pre-determined thresholds)
- 3. Impute prices for extreme hospitals

Differences among penalized hospitals

- Mean price among penalized hospitals: 9,896.31
- Mean price among non-penalized hospitals: 9,560.41
- Mean difference: 335.9

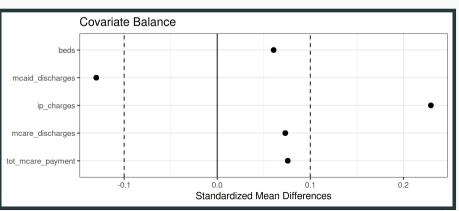
Comparison of hospitals

Are penalized hospitals sufficiently similar to non-penalized hospitals?

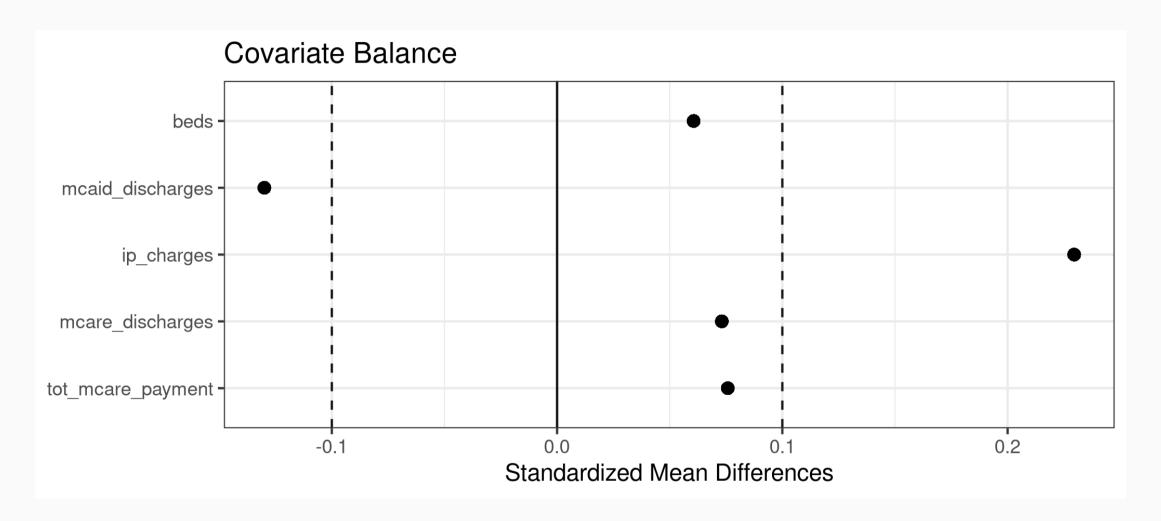
Let's look at covariate balance using a love plot, part of the library(cobalt) package.

Love plots without adjustment

```
love.plot(bal.tab(lp.covs,treat=lp.vars$penalty), colors="black", shapes="circle", threshold=0.1) +
theme bw() + theme(legend.position="none")
```



Love plots without adjustment



Using matching to improve balance

Some things to think about:

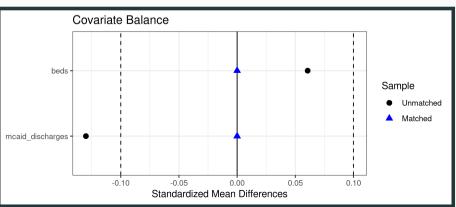
- exact versus nearest neighbor
- with or without ties (and how to break ties)
- measure of distance

1. Exact Matching

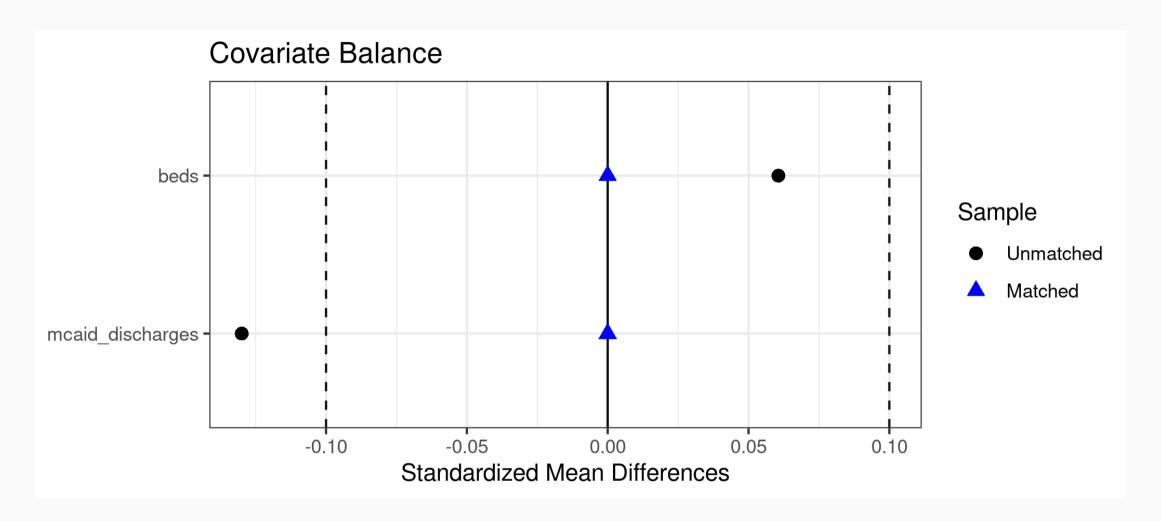
[1] "Match"

1. Exact Matching (on a subset)

1. Exact Matching (on a subset)

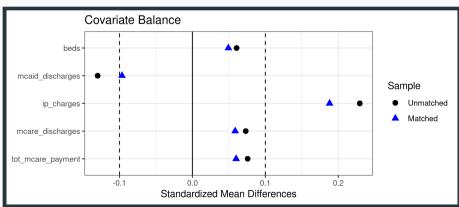


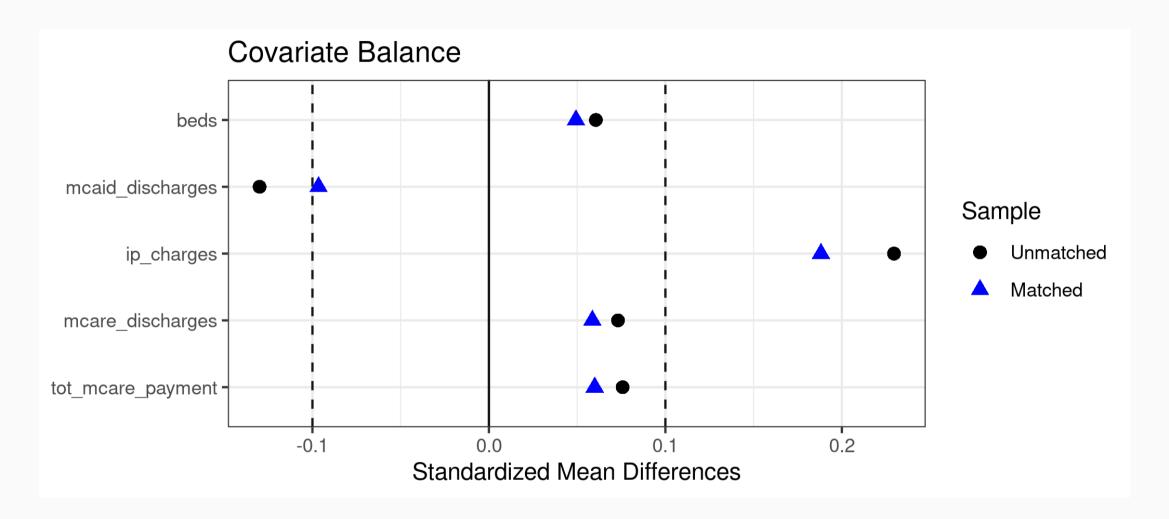
1. Exact Matching (on a subset)

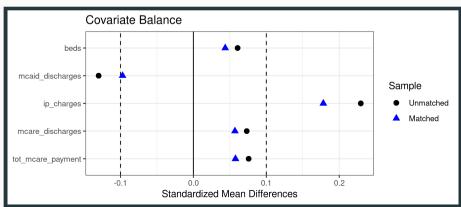


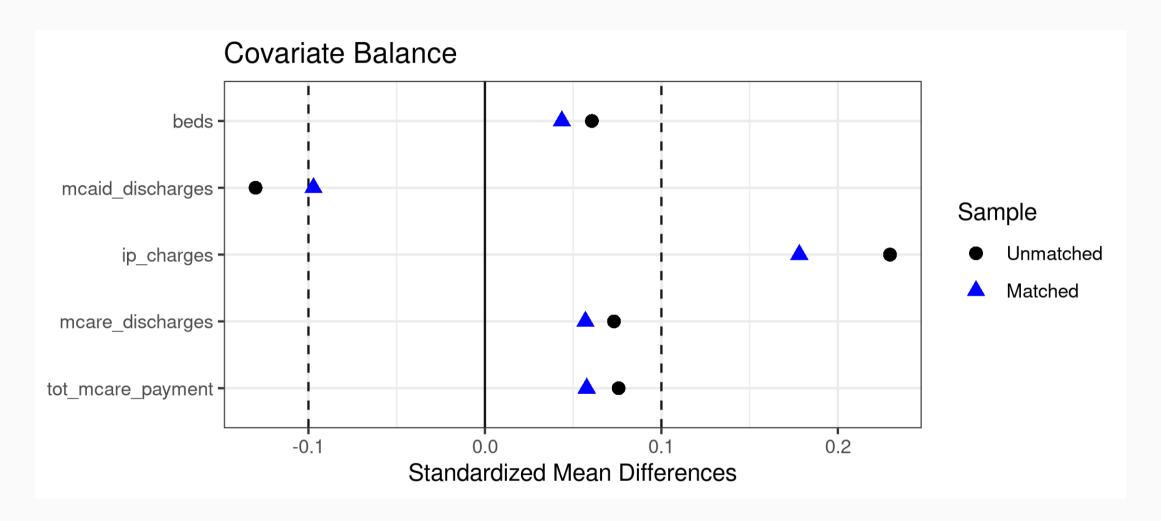
2. Nearest neighbor matching (inverse variance)

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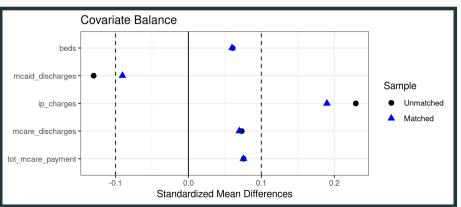




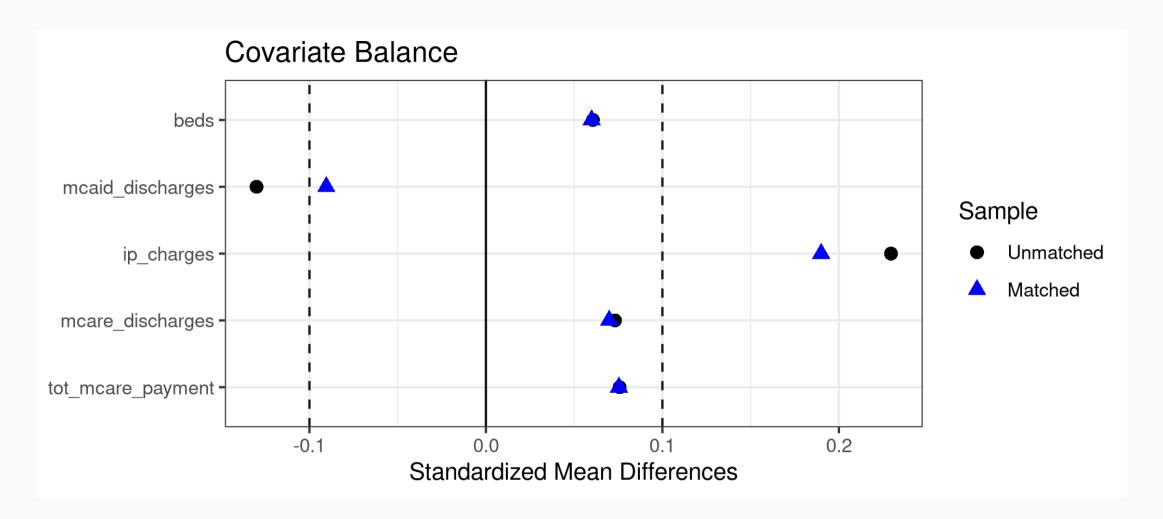


2. Nearest neighbor matching (Mahalanobis)

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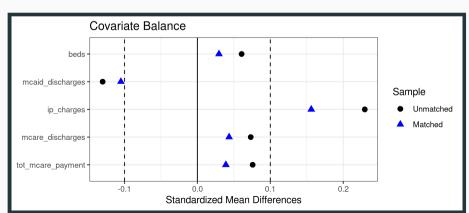


2. Nearest neighbor matching (Mahalanobis)

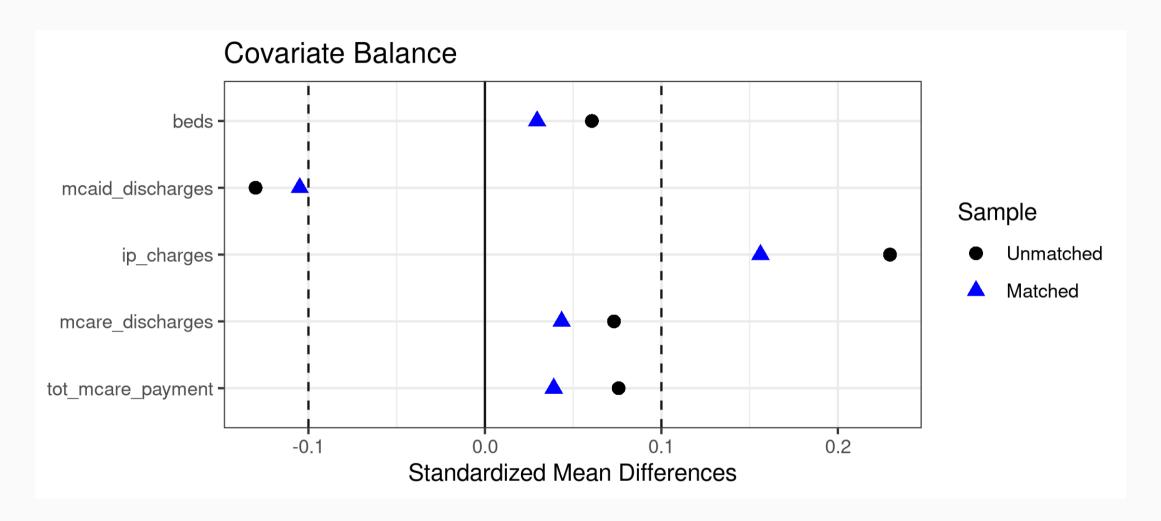


2. Nearest neighbor matching (propensity score)

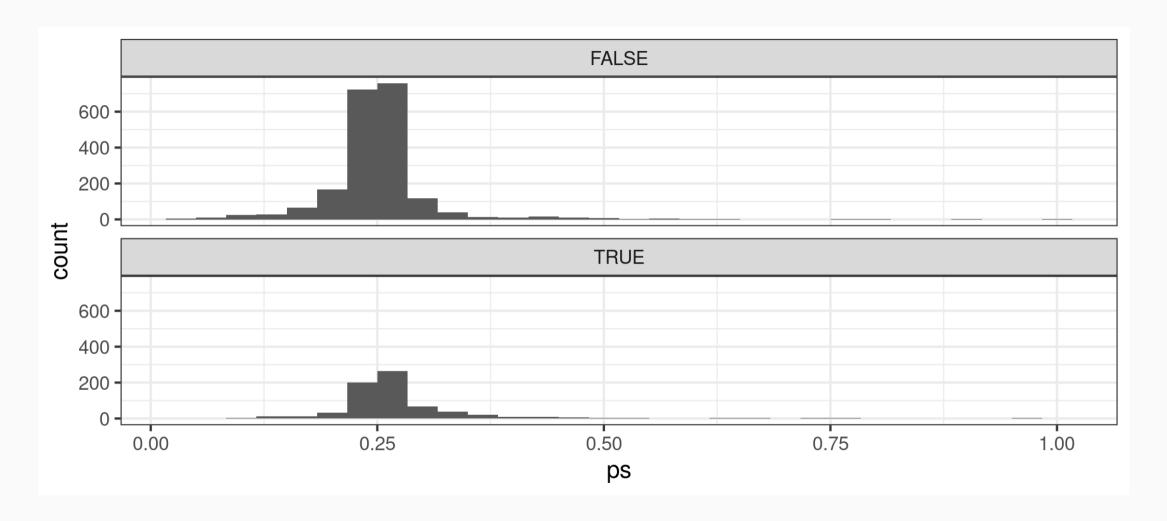
2. Nearest neighbor matching (propensity score)



2. Nearest neighbor matching (propensity score)



3. Weighting



Results: Exact matching

```
##
## Estimate... 1777.6
## AI SE..... 34.725
## T-stat.... 51.191
## p.val..... < 2.22e-16
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 12
## Matched number of observations (unweighted). 12
##
## Number of obs dropped by 'exact' or 'caliper' 2695</pre>
```

Results: Nearest neighbor

• Inverse variance

```
##
## Estimate... -526.95
## AI SE..... 223.06
## T-stat.... -2.3623
## p.val.... 0.01816
##
## Original number of observations..... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2711
```

Results: Nearest neighbor

Mahalanobis

```
##
## Estimate... -492.82
## AI SE..... 223.55
## T-stat.... -2.2046
## p.val.... 0.027485
##
## Original number of observations...... 2707
## Original number of treated obs..... 698
## Matched number of observations (unweighted). 2708
```

Results: Nearest neighbor

Propensity score

```
##
## Estimate... -201.03
## AI SE..... 275.76
## T-stat.... -0.72898
## p.val..... 0.46601
##
## Original number of observations...... 2707
## Original number of treated obs...... 698
## Matched number of observations (unweighted). 14795
```

Results: IPW weighting

```
lp.vars \leftarrow lp.vars %>%
mutate(ipw = case_when(
    penalty=1 ~ 1/ps,
    penalty=0 ~ 1/(1-ps),
    TRUE ~ NA_real_
))
mean.t1 \leftarrow lp.vars %>% filter(penalty=1) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t0 \leftarrow lp.vars %>% filter(penalty=0) %>%
    select(price, ipw) %>% summarize(mean_p=weighted.mean(price,w=ipw))
mean.t1$mean_p - mean.t0$mean_p
```

```
## [1] -196.8922
```

Results: IPW weighting with regression

```
ipw.reg ← lm(price ~ penalty, data=lp.vars, weights=ipw)
summarv(ipw.reg)
##
## Call:
### lm(formula = price ~ penalty, data = lp.vars, weights = ipw)
##
## Weighted Residuals:
     Min
          1Q Median
                       3Q
                               Max
## -18691 -4802 -1422 2651 94137
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 9876.4 147.8 66.808 <2e-16 ***
## penaltyTRUE -196.9 211.2 -0.932 0.351
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7829 on 2705 degrees of freedom
## Multiple R-squared: 0.0003211, Adjusted R-squared: -4.85e-05
## F-statistic: 0.8688 on 1 and 2705 DF, p-value: 0.3514
```

Results: Regression

```
## [1] -5.845761
```

Results: Regression in one step

Results: Regression in one step

```
###
## Call:
## lm(formula = price ~ penalty + beds + mcaid discharges + ip charges +
      mcare discharges + tot mcare payment + beds diff + mcaid diff +
###
      ip diff + mcare diff + mpay diff, data = reg.dat)
###
##
## Residuals:
     Min
             10 Median
                          3Q
                               Max
## -38175 -2900
                 -597
                        2105 67409
##
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept) 8.466e+03 1.711e+02 49.482 < 2e-16 ***
## penaltyTRUE
              -5.846e+00 2.124e+02 -0.028 0.97804
## beds
                  1.107e+00 1.421e+00 0.779 0.43618
## mcaid discharges -4.714e-01 7.296e-02 -6.462 1.23e-10 ***
## ip charges
                    6.426e-06 1.285e-06 5.002 6.04e-07 ***
## mcare discharges -8.122e-01 9.257e-02 -8.774 < 2e-16 ***
                                        13.857 < 2e-16 ***
## tot_mcare_payment 9.502e-05
                              6.858e-06
## beds diff
                    2.517e+00 2.986e+00
                                        0.843 0.39931
## mcaid diff
             1.058e-01 1.570e-01
                                        0.674 0.50050
## ip_diff
                   -4.534e-06 2.027e-06 -2.237 0.02539 *
                                        2.657 0.00793 **
## mcare diff
             4.806e-01 1.809e-01
## mpay diff
                   -5.452e-05 1.321e-05 -4.128 3.78e-05 ***
## ---
```

Summary of ATEs

- 1. Exact matching: 1777.63
- 2. NN matching, inverse variance: -526.95
- 3. NN matching, mahalanobis: -492.82
- 4. NN matching, pscore: -201.03
- 5. Inverse pscore weighting: -196.89
- 6. IPW regression: -196.89
- 7. Regression: -5.85
- 8. Regression 1-step: -5.85

So what have we learned?

Key assumptions for causal inference

- 1. Selection on observables
- 2. Common support

These become more nuanced but the intuition is the same in almost all questions of causal inference.

Causal effect assuming selection on observables

If we assume selection on observables holds, then we only need to condition on the relevant covariates to identify a causal effect. But we still need to ensure common support...

- 1. Matching
- 2. Reweighting
- 3. Regression