

Biostatistics

Applications in Medicine

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Syllabus

1. General review

- a. What is Biostatistics?
- b. Population/Sample/Sample size
- c. Type of Data – quantitative and qualitative variables
- d. Common probability distributions
- e. Work example – Malaria in Tanzania

2. Applications in Medicine

- a. Construction and analysis of diagnostic tools – Binomial distribution, sensitivity, specificity, ROC curve, Rogal-Gladen estimator
- b. Estimation of treatment effects - generalized linear models
- c. Survival analysis - Weibull regression, Kaplan-Meier curve, log-rank test, Cox's proportional hazards model

3. Applications in Genetics, Genomics, and other 'omics data

- a. Genetic association studies – Hardy-Weinberg test, homozygosity, minor allele frequencies, additive model, multiple testing correction
- b. Methylation association studies – M versus beta values, estimation of biological age
- c. Gene expression studies based on RNA-seq experiments – Tests based on Poisson and Negative-Binomial

4. Other Topics

- a. Estimation of Species diversity – Diversity indexes, Poisson mixture models
- b. Serological analysis – Gaussian (skew-normal) mixture models
- c. Advanced sample size and power calculations

Exercise: data about recovery from a SARS-CoV-2 infection

16 patients from a Beijing hospital between

January 28 and February 9, 2020

time to end of symptoms

time to negative PCR test
(Homework)

Package survival

Fit a Weibull regression model with time to end of symptoms as the outcome and age and gender as the covariate

Assess the validity of the model by testing a Gumbel distribution in the residuals

Parametric analysis

versus

Non-parametric analysis

Parametric analysis



Non-parametric analysis



Non-parametric methods

**Comparison of different
survival curves**

Log-rank test
Peto-Peto test

Kolmogorov-Smirnov test

**Semi-parametric
regression**

Cox's proportional hazard
model

Comparison of different survival curves

Two treatments under comparison

Time to clinical response

$$H_0 : S_1(t) = S_2(t) \text{ versus } H_0 : S_1(t) \neq S_2$$

Log-rank test as a Mantel-Haenszel test for categorical data

Do you know other tests

Mantel-Haenszel test

Analysis of the association in $K \times 2 \times 2$ contingency tables (an extension of Fisher's exact test to K tables 2×2).

Stratum	Treatment	Responded	Not Responded
1	A		
	B		
2	A		
	B		
3	A		
	B		

In stratum i

$$\Delta_i = \frac{\pi_{1i}(1 - \pi_{2i})}{(1 - \pi_{1i})\pi_{2i}}$$

π_{1i} = prob. of response to treatment 1

π_{2i} = prob. of response to treatment 2

$$H_0 : \Delta_1 = \dots = \Delta_K = 1(t) \text{ versus } H_1 : \exists_{i,j} \Delta_i \neq \Delta_j = 1$$

under the assumption of $\Delta_1 = \dots = \Delta_K = \Delta$

Log-rank test

Adaptation of the classical Mantel-Haenszel test for $k \times 2 \times 2$ contingency tables where k is the number of different timepoints in which it was observed the event of interest



Basic idea

There are k 2×2 tables like this one

Group	Number of “deaths” at $t_{(i)}$	Number of “survivors” beyond $t_{(i)}$	Total
1	d_{1i}	$n_{1i} - d_{1i}$	n_{1i}
2	d_{2i}	$n_{2i} - d_{2i}$	n_{2i}
Total	d_i	$n_i - d_i$	n_i

Conditional probability (see Fisher's exact test)

$H_0 : S_1(t) = S_2(t)$ versus $H_1 : S_1(t) \neq S_2(t)$

$H_0 : \pi_{1i} = \pi_{2i} = \pi$ versus $H_1 : \pi_{1i} \neq \pi_{2i}$

π_{1i} = probability of "death" at time $t_{(i)}$ in group 1

π_{2i} = probability of "death" at time $t_{(i)}$ in group 2

$d_{li} | \pi_{li}, n_{li} \rightsquigarrow \text{Binomial}(n = n_{li}, \pi = \pi_{li}), l = 1, 2$

$d_i | \pi_{li}, n_{li}, H_0 \rightsquigarrow \text{Binomial}(n = n_i, \pi = \pi_i)$

Basic idea

Calculate the distribution of d_{1i} conditional to the total marginals

Group	Number of “deaths” at $t_{(i)}$	Number of “survivors” beyond $t_{(i)}$	Total
1	d_{1i}	$n_{1i} - d_{1i}$	n_{1i}
2	d_{2i}	$n_{2i} - d_{2i}$	n_{2i}
Total	d_i	$n_i - d_i$	n_i

Conditional probability (see Fisher's exact test)

$d_{1i} | d_i, n_i, n_{1i}, H_0 \rightsquigarrow \text{Hypergeometric}(N = n_i, M = d_i, n = n_{1i})$

$$P[d_{1i} = d | d_i, n_i, n_{1i}, H_0] = \frac{\binom{d_i}{d} \binom{n_i - d_i}{n_{1i} - d}}{\binom{n_i}{n_{1i}}}$$

$$E[d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i}$$

$$\text{Var}[d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i} \left(1 - \frac{d_i}{n_i}\right) \frac{n_i - n_{1i}}{n_i - 1}$$

Test statistic

Incorporating information from k 2×2 contingency tables

$$U = \sum_{i=1}^k (d_{1i} - e_{1i})$$

$$e_{1i} = E [d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i}$$

$$E [U | H_0] = 0$$

$$\nu_{1i} = Var [d_{1i} | d_i, n_i, n_{1i}, H_0]$$

$$Var [U | H_0] = \sum_{i=1}^k \nu_{1i}$$

$$= n_{1i} \frac{d_i}{n_i} \left(1 - \frac{d_i}{n_i} \right) \frac{n_i - n_{1i}}{n_i - 1}$$

Log-rank test

For large samples

$$Q = \frac{U - \overline{E(U)}}{\sqrt{var(U)}} \mid H_0 \rightsquigarrow \text{Normal}(\mu = 0, \sigma = 1)$$

$$Q^* = \frac{U^2}{var(U)} \mid H_0 \rightsquigarrow \chi_{(1)}^2$$

Decision rule

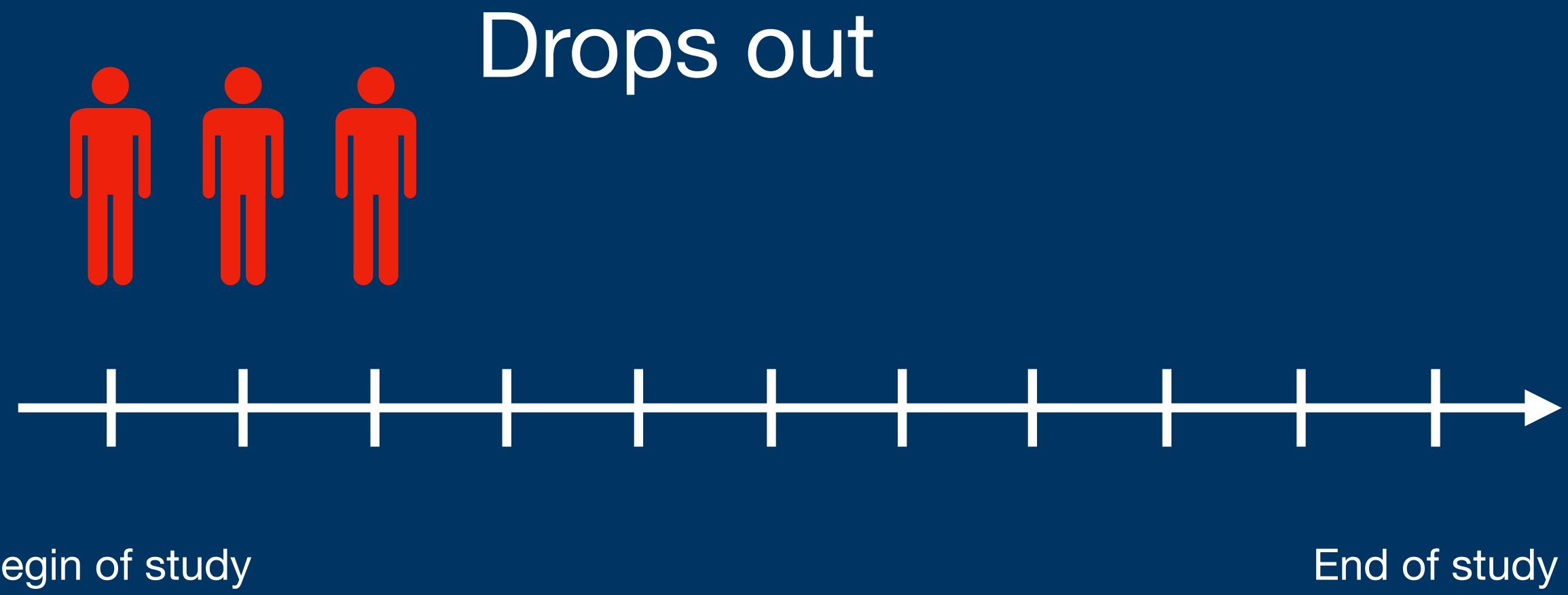
$$p = P [Q^* > q_{obs} \mid H_0] \quad \left\{ \begin{array}{ll} \text{do not reject } H_0, & \text{if } p > \alpha \\ \text{reject } H_0, & \text{otherwise} \end{array} \right.$$

Exercise 1: rituximab clinical trial data

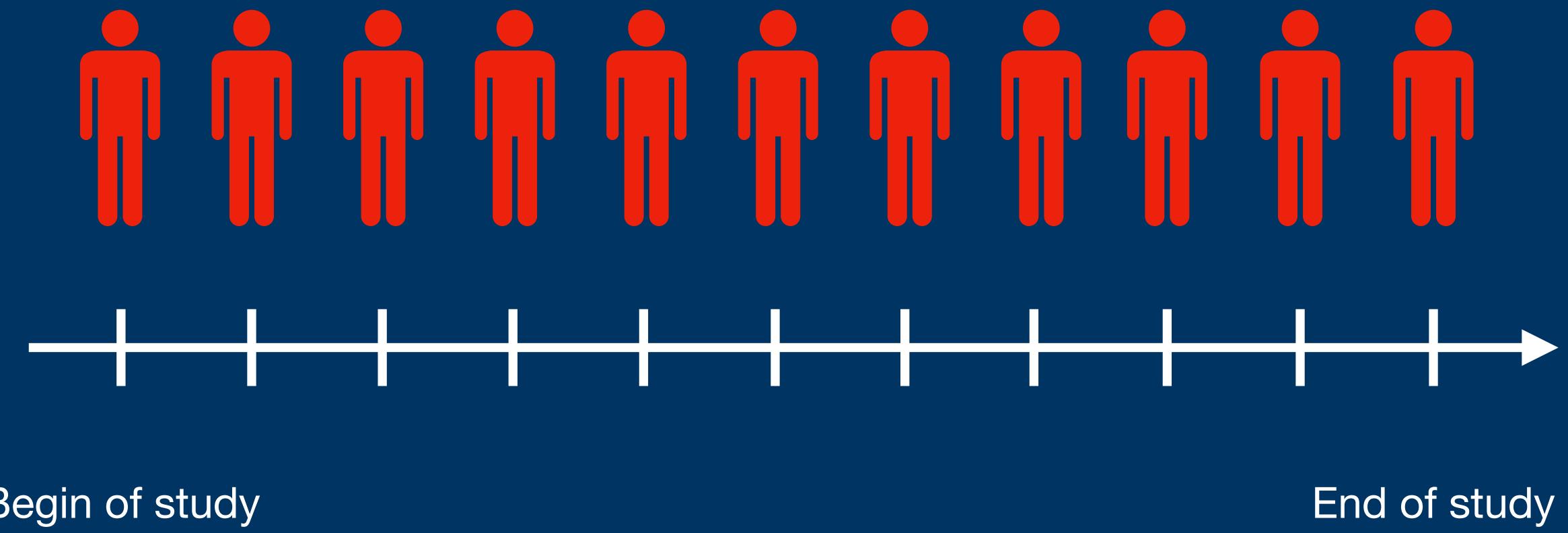
Compare the survival curves for time to first response for men versus women using log-rank (survdiff function from survival package)

Draw your conclusions.

What type of incomplete data?



What type of incomplete data?



Exercise 2: rituximab clinical trial data

Identify and quantify censored data concerning time to first treatment response

Should you consider interval censoring in this case?

Basic mathematical formulation of the problem

Right/left censored data

$$\{t_i, d_i\}, i = 1, \dots, n$$

$$t_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is right censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$t_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is left censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$d_i = \begin{cases} 0, & \text{if } t_i \text{ is censored} \\ 1, & \text{if } t_i \text{ is completely observed} \end{cases}$$

Interval censored data

$$\{a_i, b_i, d_i\}, i = 1, \dots, n$$

$$a_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$b_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

In practice

Package survival

Survival time

$$\text{time}_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is right or interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \\ t_i^+, & \text{if } t_i \text{ is left censored} \end{cases}$$

$$t_i \in (t_i^*, t_i^+)$$

$$\text{time2}_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is interval censored} \\ 0, & \text{otherwise} \end{cases}$$

Event indicator

$$d_i = \begin{cases} 0, & \text{if } t_i \text{ is right censored} \\ 1, & \text{if } t_i \text{ is completely observed} \\ 2, & \text{if } t_i \text{ is left censored} \\ 3, & \text{if } t_i \text{ is interval censored} \end{cases}$$

Likelihood function of a parametric model under different censoring mechanisms

$$T_i | \theta \rightsquigarrow F(\theta)$$

Weibull, Gamma, Lognormal, Log-logistic, etc

Right censored data

$$L\left(\theta | \{t_i, d_i\}\right) \equiv \prod_{i=1}^n f_\theta(t_i)^{d_i} S_\theta(t_i)^{1-d_i}$$

Left censored data

$$L\left(\theta | \{t_i, d_i\}\right) \equiv \prod_{i=1}^n f_\theta(t_i)^{d_i} F_\theta(t_i)^{1-d_i}$$

Interval censored data

$$L\left(\theta | \{a_i, b_i, d_i\}\right) \equiv \prod_{i=1}^n f_\theta(a_i)^{d_i} (F_\theta(b_i) - F_\theta(a_i))^{1-d_i}$$

Parametric estimation

$$\hat{\theta} = \operatorname{argmax}_{\theta} L(\theta | \{t_i, d_i\})$$

No closed-form expressions

Numerical solutions for the maximum likelihood equations

Exercise: rituximab clinical trial data

Estimate lognormal, weibull and log-logistic model to data on time to treatment response using the “survreg” function of package survival.

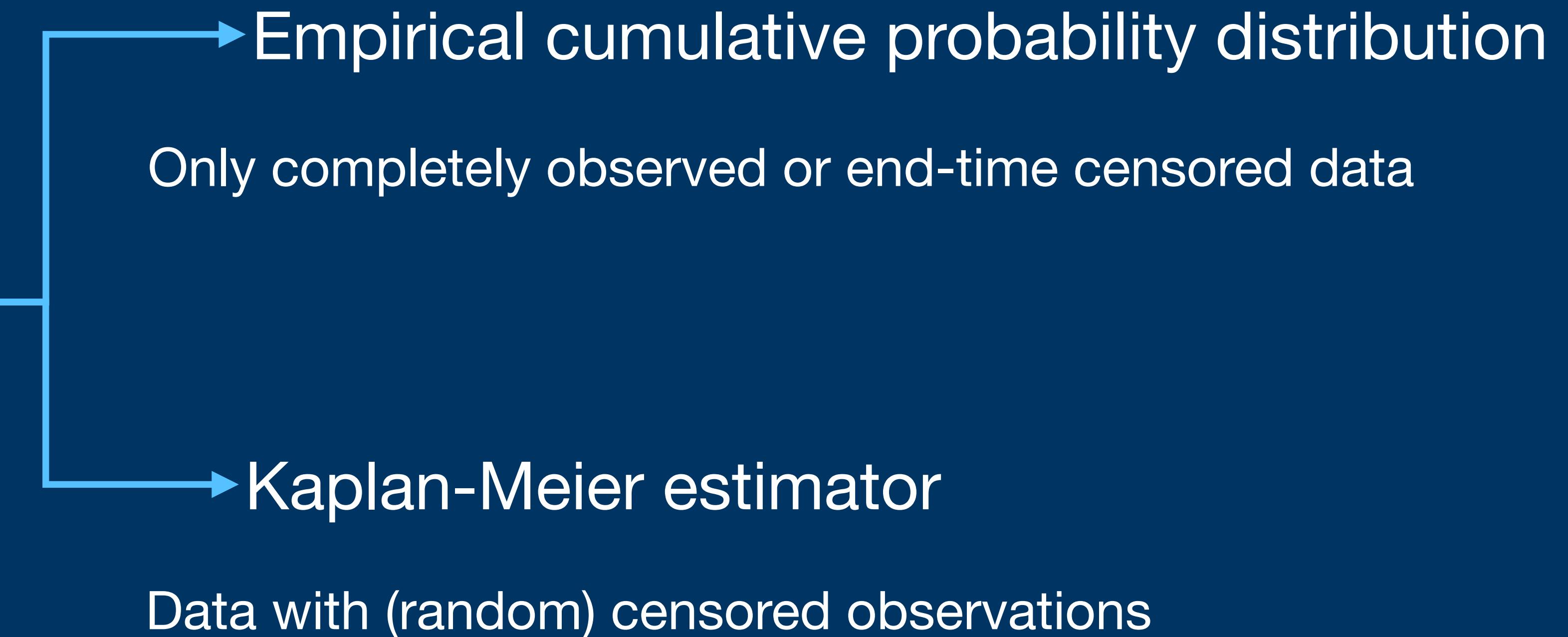
What is the best model for the data?

Checking the adequacy of a parametric model

Can we use the Kolmogorov-Smirnov test directly to data?

Kaplan-Meier estimator for the survival function

Non-parametric
estimation of the
cumulative probability
distribution



Kaplan-Meier estimator for the survival function

$$\hat{S}(t) = \prod_{i:t_{(i)} \leq t} \left(1 - \frac{d_i}{n_i}\right) \quad t \in (0, t_{\max})$$

d_i = number of individuals in which the event was observed at $t_{(i)}$

n_i = number of individuals without the event of interest at $t_{(i-1)}$

$\{t_{(i)}, i = 1, \dots, r\}$ = unique times when the event of interest was observed

Kaplan-Meier estimator for the survival function

$$\hat{S}(t_{(1)}) = 1 - \frac{d_1}{n_1}$$

n_1 = number of individuals without the event of interest at time 0 = n

$$\hat{S}(t_{(i)}) = \hat{S}(t_{(i-1)}) \left(1 - \frac{d_i}{n_i} \right)$$

Exercise 3: rituximab clinical trial data

Estimate the survival curve of time to treatment response using the Kaplan-Meier estimator

Compare the Kaplan-Meier estimated survival curve to the survival curve predicted by the best parametric model from Exercise 2.