

# **Biostatistics**

## **Applications in Medicine**

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# Syllabus

## 1. General review

- a. What is Biostatistics?
- b. Population/Sample/Sample size
- c. Type of Data – quantitative and qualitative variables
- d. Common probability distributions
- e. Work example – Malaria in Tanzania

## 2. Applications in Medicine

- a. Construction and analysis of diagnostic tools – Binomial distribution, sensitivity, specificity, ROC curve, Rogal-Gladen estimator
- b. Estimation of treatment effects - generalized linear models
- c. Survival analysis - Weibull regression, Kaplan-Meier curve, log-rank test, Cox's proportional hazards model

## 3. Applications in Genetics, Genomics, and other 'omics data

- a. Genetic association studies – Hardy-Weinberg test, homozygosity, minor allele frequencies, additive model, multiple testing correction
- b. Methylation association studies – M versus beta values, estimation of biological age
- c. Gene expression studies based on RNA-seq experiments – Tests based on Poisson and Negative-Binomial

## 4. Other Topics

- a. Estimation of Species diversity – Diversity indexes, Poisson mixture models
- b. Serological analysis – Gaussian (skew-normal) mixture models
- c. Advanced sample size and power calculations

# Exercise:

## data about recovery from a SARS-CoV-2 infection

16 patients from a Beijing hospital between  
January 28 and February 9, 2020



time to end of symptoms

time to negative PCR test  
(Homework)

Package survival

Fit a Weibull regression model with time to end of symptoms as the outcome and age  
and gender as the covariate

Assess the validity of the model by testing a Gumbel distribution in the residuals

**Parametric analysis**

**versus**

**Non-parametric analysis**

# Parametric analysis

**Complete  
data**



**Incomplete  
data**

# Non-parametric analysis



# Non-parametric methods

## Comparison of different survival curves

Log-rank test  
Peto-Peto test

Kolmogorov-Smirnov test

## Semi-parametric regression

Cox's proportional hazard model

# Comparison of different survival curves

Two treatments under comparison

Time to clinical response

$$H_0 : S_1(t) = S_2(t) \text{ versus } H_0 : S_1(t) \neq S_2$$

Log-rank test as a Mantel-Haenszel test for categorical data

Do you know other tests



# Mantel-Haenszel test

Analysis of the association in  $K \times 2 \times 2$  contingency tables (an extension of Fisher's exact test to  $K$  tables  $2 \times 2$ ).

Stratum	Treatment	Responded	Not Responded
1	A		
	B		
2	A		
	B		
3	A		
	B		

In stratum  $i$

$$\Delta_i = \frac{\pi_{1i}(1 - \pi_{2i})}{(1 - \pi_{1i})\pi_{2i}}$$

$\pi_{1i}$  = prob. of response to treatment 1

$\pi_{2i}$  = prob. of response to treatment 2

$$H_0 : \Delta_1 = \dots = \Delta_K = 1 \text{ (t) versus } H_1 : \exists_{i,j} \Delta_i \neq \Delta_j = 1$$

under the assumption of  $\Delta_1 = \dots = \Delta_K = \Delta$

# Log-rank test

Adaptation of the classical Mantel-Haenszel test for  $k \times 2 \times 2$  contingency tables where  $k$  is the number of different timepoints in which it was observed the event of interest



Basic idea

There are k 2 x 2 tables like this one

Group	Number of “deaths” at $t_{(i)}$	Number of “survivors” beyond $t_{(i)}$	Total
1	$d_{1i}$	$n_{1i} - d_{1i}$	$n_{1i}$
2	$d_{2i}$	$n_{2i} - d_{2i}$	$n_{2i}$
Total	$d_i$	$n_i - d_i$	$n_i$

## Conditional probability (see Fisher's exact test)

$$H_0 : S_1(t) = S_2(t) \text{ versus } H_1 : S_1(t) \neq S_2(t)$$

$$H_0 : \pi_{1i} = \pi_{2i} = \pi \text{ versus } H_1 : \pi_{1i} \neq \pi_{2i}$$

$\pi_{1i}$  = probability of "death" at time  $t_{(i)}$  in group 1

$\pi_{2i}$  = probability of "death" at time  $t_{(i)}$  in group 2

$$d_{li} \mid \pi_{li}, n_{li} \rightsquigarrow \text{Binomial}(n = n_{li}, \pi = \pi_{li}), l = 1, 2$$

$$d_i \mid \pi_{li}, n_{li}, H_0 \rightsquigarrow \text{Binomial}(n = n_i, \pi = \pi_i)$$

## Basic idea

Calculate the distribution of  $d_{1i}$  conditional to the total marginals

Group	Number of “deaths” at $t_{(i)}$	Number of “survivors” beyond $t_{(i)}$	Total
1	$d_{1i}$	$n_{1i} - d_{1i}$	$n_{1i}$
2	$d_{2i}$	$n_{2i} - d_{2i}$	$n_{2i}$
Total	$d_i$	$n_i - d_i$	$n_i$

## Conditional probability (see Fisher's exact test)

$$d_{1i} | d_i, n_i, n_{1i}, H_0 \rightsquigarrow \text{Hypergeometric}(N = n_i, M = d_i, n = n_{1i})$$

$$P [d_{1i} = d | d_i, n_i, n_{1i}, H_0] = \frac{\binom{d_i}{d} \binom{n_i - d_i}{n_{1i} - d}}{\binom{n_i}{n_{1i}}}$$

$$E [d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i} \qquad \text{Var} [d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i} \left(1 - \frac{d_i}{n_i}\right) \frac{n_i - n_{1i}}{n_i - 1}$$

## Test statistic

Incorporating information from k 2 x 2 contingency tables

$$U = \sum_{i=1}^k (d_{1i} - e_{1i})$$

$$e_{1i} = E [d_{1i} | d_i, n_i, n_{1i}, H_0] = n_{1i} \frac{d_i}{n_i}$$

$$E [U | H_0] = 0$$

$$v_{1i} = Var [d_{1i} | d_i, n_i, n_{1i}, H_0]$$

$$Var [U | H_0] = \sum_{i=1}^k v_{1i}$$

$$= n_{1i} \frac{d_i}{n_i} \left( 1 - \frac{d_i}{n_i} \right) \frac{n_i - n_{1i}}{n_i - 1}$$

# Log-rank test

For large samples

$$Q = \frac{U - \overbrace{E(U)}^{=0}}{\sqrt{\text{var}(U)}} \mid H_0 \rightsquigarrow \text{Normal}(\mu = 0, \sigma = 1)$$

$$Q^* = \frac{U^2}{\text{var}(U)} \mid H_0 \rightsquigarrow \chi^2_{(1)}$$

Decision rule

$$p = P [Q^* > q_{obs} \mid H_0]$$

$$\begin{cases} \text{do not reject } H_0, & \text{if } p > \alpha \\ \text{reject } H_0, & \text{otherwise} \end{cases}$$

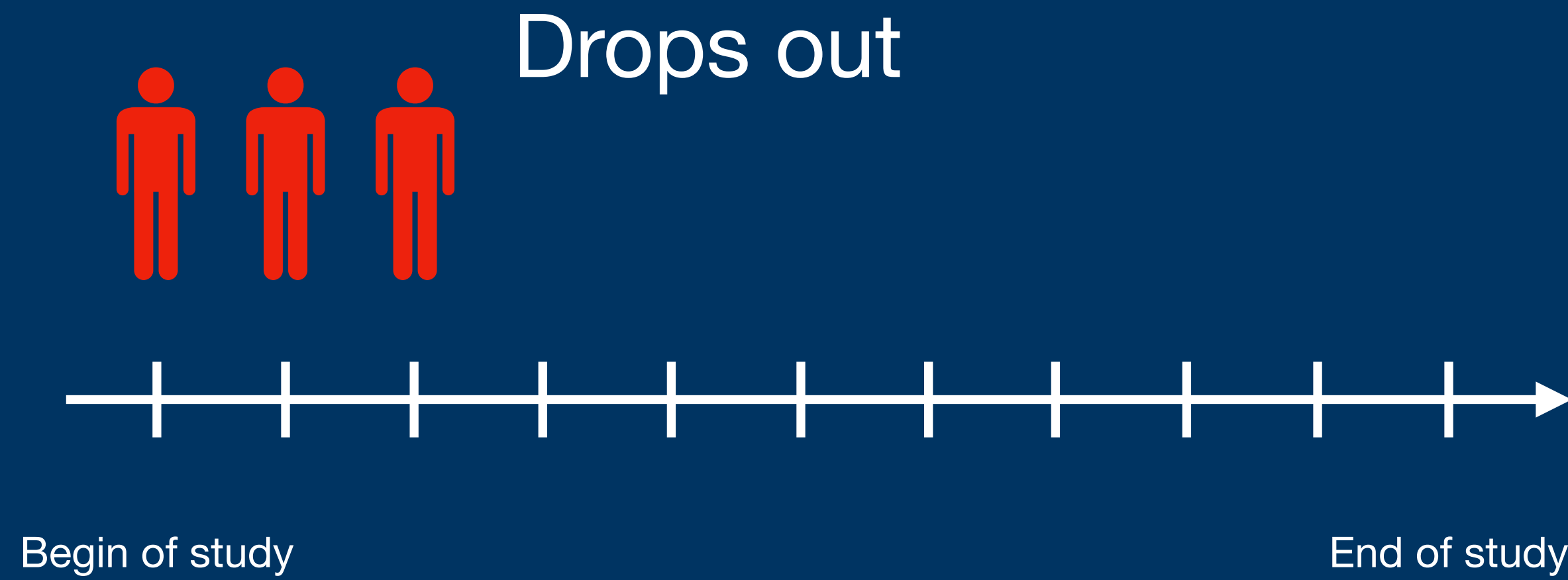


## Exercise 1: rituximab clinical trial data

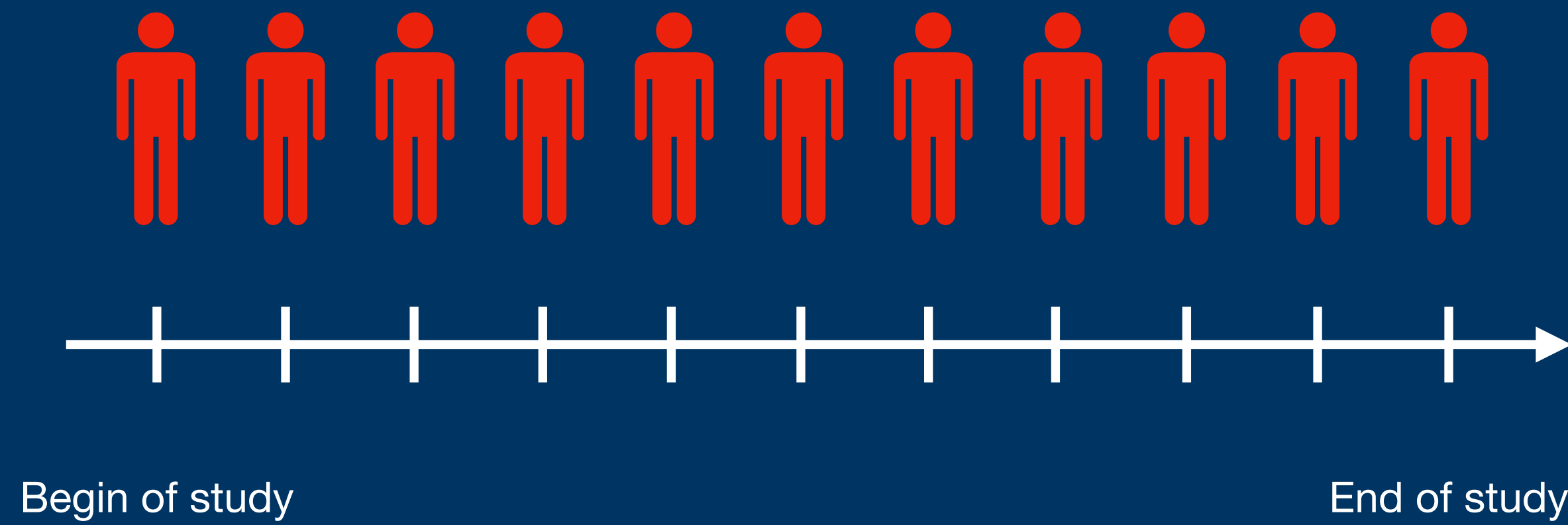
Compare the survival curves for time to first response for men versus women using log-rank (survdif function from survival package)

Draw your conclusions.

# What type of incomplete data?



# What type of incomplete data?



## Exercise 2: rituximab clinical trial data

Identify and quantify censored data concerning time to first treatment response

Should you consider interval censoring in this case?

# Basic mathematical formulation of the problem

## Right/left censored data

$$\{t_i, d_i\}, i = 1, \dots, n$$

$$t_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is right censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$t_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is left censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

## Interval censored data

$$\{a_i, b_i, d_i\}, i = 1, \dots, n$$

$$a_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$b_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \end{cases}$$

$$d_i = \begin{cases} 0, & \text{if } t_i \text{ is censored} \\ 1, & \text{if } t_i \text{ is completely observed} \end{cases}$$

## In practice

### Package survival

#### Survival time

$$\text{time}_i = \begin{cases} t_i^*, & \text{if } t_i \text{ is right or interval censored} \\ t_i, & \text{if } t_i \text{ is completely observed} \\ t_i^+, & \text{if } t_i \text{ is left censored} \end{cases}$$

$$t_i \in (t_i^*, t_i^+)$$

$$\text{time2}_i = \begin{cases} t_i^+, & \text{if } t_i \text{ is interval censored} \\ 0, & \text{otherwise} \end{cases}$$

#### Event indicator

$$d_i = \begin{cases} 0, & \text{if } t_i \text{ is right censored} \\ 1, & \text{if } t_i \text{ is completely observed} \\ 2, & \text{if } t_i \text{ is left censored} \\ 3, & \text{if } t_i \text{ is interval censored} \end{cases}$$

# Likelihood function of a parametric model under different censoring mechanisms

$$T_i | \theta \rightsquigarrow F(\theta)$$

Weibull, Gamma, Lognormal, Log-logistic, etc

Right censored data

$$L\left(\theta | \{t_i, d_i\}\right) \equiv \prod_{i=1}^n f_{\theta}(t_i)^{d_i} S_{\theta}(t_i)^{1-d_i}$$

Left censored data

$$L\left(\theta | \{t_i, d_i\}\right) \equiv \prod_{i=1}^n f_{\theta}(t_i)^{d_i} F_{\theta}(t_i)^{1-d_i}$$

Interval censored data

$$L\left(\theta | \{a_i, b_i, d_i\}\right) \equiv \prod_{i=1}^n f_{\theta}(a_i)^{d_i} \left(F_{\theta}(b_i) - F_{\theta}(a_i)\right)^{1-d_i}$$

## Parametric estimation

$$\hat{\theta} = \operatorname{argmax}_{\theta} L\left(\theta \mid \{t_i, d_i\}\right)$$

No closed-form expressions

Numerical solutions for the maximum likelihood equations



## Exercise: rituximab clinical trial data

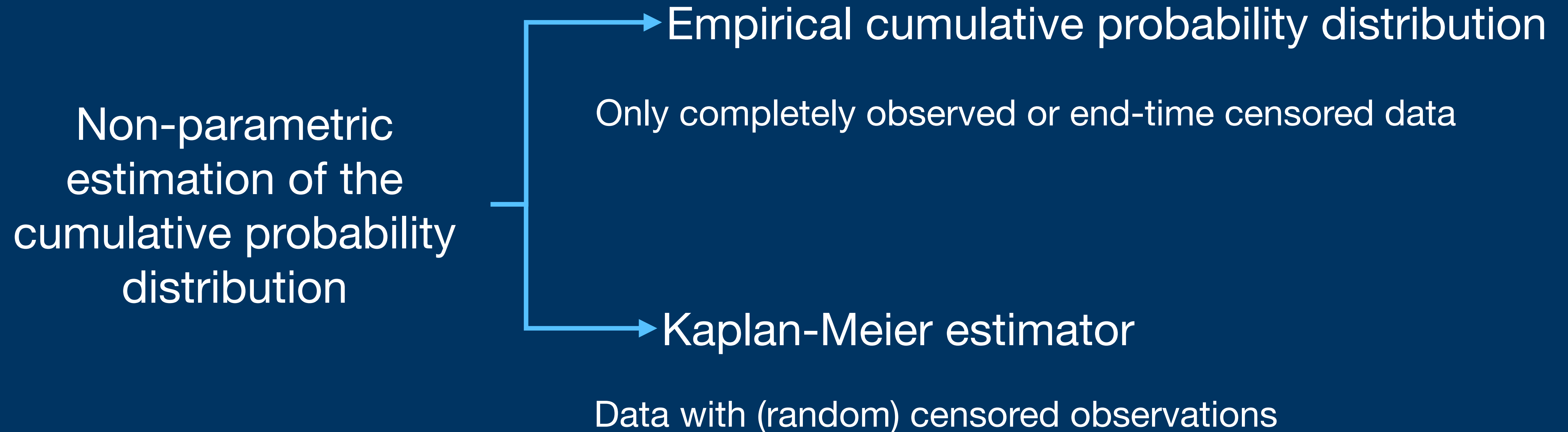
Estimate lognormal, weibull and log-logistic model to data on time to treatment response using the “survreg” function of package survival.

What is the best model for the data?

# Checking the adequacy of a parametric model

Can we use the Kolmogorov-Smirnov test directly to data?

# Kaplan-Meier estimator for the survival function



## Kaplan-Meier estimator for the survival function

$$\hat{S}(t) = \prod_{i: t_{(i)} \leq t} \left( 1 - \frac{d_i}{n_i} \right) \quad t \in (0, t_{\max})$$

$d_i$  = number of individuals in which the event was observed at  $t_{(i)}$

$n_i$  = number of individuals without the event of interest at  $t_{(i-1)}$

$\{t_{(i)}, i = 1, \dots, r\}$  = unique times when the event of interest was observed

## Kaplan-Meier estimator for the survival function

$$\hat{S}\left(t_{(1)}\right) = 1 - \frac{d_1}{n_1}$$

$n_1$  = number of individuals without the event of interest at time 0 =  $n$

$$\hat{S}\left(t_{(i)}\right) = \hat{S}\left(t_{(i-1)}\right) \left(1 - \frac{d_i}{n_i}\right)$$

## Exercise 3: rituximab clinical trial data

Estimate the survival curve of time to treatment response using the Kaplan-Meier estimator

Compare the Kaplan-Meier estimated survival curve to the survival curve predicted by the best parametric model from Exercise 2.