

A Farewell to Army Segregation: The Effects of Racial Integration During the Korean War

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Abstract

The racial integration of the U.S. Army during the Korean War (1950-1953) is one of the largest and swiftest desegregation episodes in American history. Integration began in an effort to reinforce badly depleted all-white units, and went on to become Army-wide policy for reasons of military efficiency. The first part of this paper evaluates whether the Army achieved its goal of improving efficiency as measured by the survival rates of wounded soldiers. Using casualty data, I develop a novel wartime integration measure to quantify exogenous changes in racial integration over time and across regiments. Based on a modified difference-in-differences strategy, I find that a one standard deviation increase in regimental integration improved overall casualty survival rates by 3%. The second part of the paper explores the effects of wartime racial integration on the prejudicial attitudes of veterans after the war. To do so, I link individual soldiers to post-war social security and cemetery data using an unsupervised learning algorithm. With these linked samples, I show that a one standard deviation increase in wartime racial integration caused white veterans to live in more racially diverse neighborhoods and marry spouses with less distinctively white names. These results provide suggestive evidence that large-scale interracial contact reduces prejudice on a long-term basis.

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1 Introduction

The racial integration of the Eighth U.S. Army during the Korean War (1950-1953) is one of the largest and swiftest desegregation episodes in American history (Mershon and Schlossman, 1998). At the outset of the war, blacks and whites were rigidly segregated both by unit and by occupation: all-white units were largely responsible for combat, while blacks were assigned primarily to non-combat support roles.¹ Following the disastrous campaigns in the early years of the Korean War, many all-white combat units became severely depleted, and in a bid to bolster manpower, some of these depleted units began to accept black reinforcements (Blair, 1989). In this way, integration in Korea began informally after what was intended to be a temporary emergency measure, and not by decree.

The Army feared that racial strife would hinder combat performance in these newly integrated units, so as a precaution it sent teams of social scientists to foxholes in 1951—a time when integrated and segregated military units were still operating in parallel—to interview white officers commanding both racially mixed and segregated units (Hausrath, 1954). Code-named “Project Clear,” the primary concern of this study was to determine what effect, if any, racial integration had on combat performance. Project Clear offered some reassurance in this regard, as white officers perceived little difference in the performance of racially-mixed rifle squads² under their command. Given the circumstances, however, Project Clear had several drawbacks: black soldiers were never interviewed, interviews only took place once, and the project relied on self-reported measures of unit performance.

This raises the question what would more objective measures reveal about the performance of newly integrated units, or the treatment of blacks within those units? The first part of this paper addresses these questions by taking a novel quantitative approach. The approach requires a measure of integration that varies across units over the course of the war. I am able to construct such a measure thanks to a remarkably rich casualty dataset covering

¹ A notable exception is the all-black 24th Regiment, which was the last remaining of the four all-black “Buffalo Soldier” regiments from the American Civil War.

² Table 1 outlines the organizational structure of the U.S. military.

all 103,000 individual Korean War Army veterans who were killed or wounded during the war. Such casualty data make it possible to quantify the extent of integration: intuitively, if black and white soldiers were physically placed in the same foxholes, then their injuries should correlate highly. I find that between the first and last year of the war, casualty correlations rose from 2% to 91%. Using regimental war diaries, I present a case study that credibly shows integration also varied across units within a given time period.

Considering that the immediate goal of integration was to reinforce depleted all-white combat teams, I evaluate the effect of integration on team performance by modeling the survival rates of wounded soldiers. To do so, I develop a modified difference-in-differences strategy where the identifying assumption is that any departure from parallel trends in casualty survival can be attributed to continuous changes in integration levels. Using this strategy, I find that a one standard deviation (SD) increase in integration caused a 3% increase in the survival rates of wounded white soldiers relative to those who remained segregated. By contrast, I find that a one SD increase in integration caused a 2% *decrease* in the survival rates of wounded black soldiers. Given that blacks were initially confined to non-combat support roles, these results are best viewed as a convergence in hazardous combat assignments. What is less obvious is how racial integration affected overall survival rates. Based on a pooled sample of blacks and whites, I find that a one SD increase in integration caused a 3% increase in the *overall* survival rates of wounded soldiers. Taking the view that black and white soldiers are close substitutes, this result is consistent with the hypothesis that segregation impeded the efficient allocation of scarce labor.³

This positive effect of integration on overall survival rates is perhaps surprising given the vast number of empirical studies documenting the adverse effects of group heterogeneity on various outcomes, such as team performance or the provision of public goods. Alesina et al. (1999) show that ethnically fractionalized counties in the U.S. spend a smaller share

³ One limitation of the casualty data is that they do not include the overall size of regiments at a given point in the war. Some regimental war diaries contain data on regiment size, which I am in the process of converting into a machine-readable format.

on productive goods such as education and infrastructure. In a sample of U.S. firms, Adams and Ferreira (2009) find that directors in gender-diverse corporate boards have better attendance records but that these firms are more vulnerable to costly outside takeovers. Similar efficiency losses have also been tied to group fractionalization in military contexts. During the American Civil War (1861-1865), Costa and Kahn (2003) find that socially fractionalized infantry companies endured more arrests and desertions.⁴ A major difference between the Korean War and these other settings, however, is that in addition to composition, racial integration changed the size of regiments. In a depleted all-white combat unit, the marginal productivity of an additional black was relatively high compared to the over-strength support units that blacks were drawn from. Thus, the experiment I consider is what happens to regiment (firm) productivity, as measured by aggregate survival rates, when restrictions on where soldiers (labour) can be allocated are removed.

From a longer-term perspective, the natural restrictions that war places on selective migration and sorting present a rare opportunity to investigate the effects of interracial contact on prejudice.⁵ Without such restrictions, highly prejudiced individuals would be able to self-select out of contact, and standard econometric analysis would exaggerate the social benefit of intergroup contact. The Army’s unique ability to compel its soldiers is in stark contrast to civilian settings, where large-scale desegregation in turn propelled white flight. For example, Reber (2005) finds that actively busing black students into previously all-white schools *accelerated* the exodus of white families moving into the suburbs; similarly, during the Second Great Migration (1940-1970), Boustan (2007) estimates that 2.5 white urban residents departed for every black arrival.

Although enlistment contracts required soldiers to remain in the Korean Peninsula, one possibility is that integration led to a kind of preemptive white flight: would-be volunteers

⁴ By contrast, reduced enlistment terms during the Korean War kept desertion rates below 2%, compared to 10% in the Civil War.

⁵ The datasets I use in this paper are not rich enough to distinguish taste-based discrimination from statistical discrimination. For ease of exposition, I use the term “prejudice” as a stand-in for either form of discrimination.

with a distaste for integration were dissuaded from joining the Army as news of racial integration reached the home front. However, I find no evidence that this took place. Between the first and second halves of the war, the fraction of wounded white volunteers who came from the Deep South—historically the most prejudiced region in the country—*increased* by 0.5 percentage points.⁶ Moreover, once deployed, the vast majority of soldiers held low ranks and therefore had no say as to which unit they served in. In my sample, 80% of soldiers held the rank of corporal or lower (see Table 2 for a hierarchy of ranks), and results from balance tests indicate that white Southerners were just as likely to be integrated as white Northerners.

One of the fascinating aspects of racial integration during the Korean War is that we have contemporaneous evidence about the effect of integration on prejudice. Researchers from Project Clear interviewed white officers about their attitudes towards blacks and found that—compared to their segregated peers—officers commanding integrated platoons reported *less* anti-black prejudice (Bogart, 1992). Given that integration reduced prejudice contemporaneously, did it also produce any lasting effects?

To address this issue, the second part of the paper studies the long-run effects of wartime integration by examining the post-war behavior of veterans toward people outside their race group. The two behavioral measures I consider are (i) where veterans chose to live and (ii) whom they married. To study veterans’ residential choice, I link veterans in the casualty file to social security data using the expectation maximization (EM) algorithm advocated by Abramitzky et al. (2019). These linked data enable me to observe the veteran’s last zip code of residence, for which I in turn construct a residential similarity index using block-level Census data. Based on a sample of veterans born after 1930, I find that a one SD increase in wartime integration increased residential integration by 0.1 SD—roughly the difference in integration between present-day Central Harlem and the Upper East Side. To further mitigate sample-selectivity concerns, I repeat the analysis on a sample restricted to draftees.

⁶ The definition of the Deep South used here are the states of Georgia, Alabama, South Carolina, Mississippi, and Louisiana.

An arguably higher bar for prejudice reduction is exogamy—marrying outside one’s own race or ethnic group. To obtain data on wives, I again use the EM algorithm but this time to link soldiers to national cemetery data. Since at this stage I do not have direct data on the wife’s race, I construct a race index based on her given name, year of birth and state of death. In terms of intermarriage, I find that a SD increase in wartime integration led to a 2 percentage point (or 0.08 SD) increase in the wife’s non-white name index. For both residential choice and intermarriage, I replicate this analysis for veterans born before 1930 and find no effect.

The circumstances under which integration in Korea came about conform to the conditions of the so-called “contact hypothesis” (Allport, 1954). In particular, Allport (1954) hypothesizes that intergroup contact will reduce prejudice if two groups have (i) equal status (e.g. age/rank); (ii) common goals (e.g. a common enemy); (iii) intergroup cooperation (e.g. combat teams); and (iv) support of authorities (e.g. officers). A consistent finding from social psychology is that this type of intergroup contact does indeed reduce prejudice, but these studies are typically conducted in laboratory settings and rely on non-behavioral measures of prejudice (Pettigrew and Tropp, 2006). In the past 15 years, economists have generalized the contact hypothesis beyond the laboratory across a variety of revealed-preference measures (e.g. classrooms: Rao, 2019; dorm rooms: Boisjoly et al., 2006, Camargo et al., 2010, Burns et al., 2015, and Air Force Academy squadrons: Carrell et al., 2016).⁷

One drawback of real-world applications is that they are often contaminated by network effects (Camargo et al., 2010). If, say, a white individual were assigned a black roommate and reports having formed more black friendships afterward, the researcher cannot distinguish from this between a change in prejudicial attitudes or simply an overall increased exposure to blacks. Network effects could in theory be eliminated if black and white strangers were integrated and then isolated again. To such a scenario, the Korean War provides a suitable approximation: whites and blacks from different counties and states were integrated in the

⁷In a closely related study, Schindler and Westcott (2015) find that U.K. counties that experienced a one SD increase in exposure to black G.I.s during WWII had 0.03 SD fewer members of far-right nationalist parties in 2010. This paper differs from Schindler and Westcott (2015) in that I consider a treatment group who had direct contact with blacks.

Korean Peninsula only to be separated again upon returning to the United States. Thus, in terms of where veterans ultimately lived or whom they married, network effects are less pervasive compared to other physical settings, and any change in behavior toward the out-group plausibly reflects a genuine change in prejudice—be it in attitudes or beliefs. Overall, these results relating to veterans’ post-war behavior provide suggestive evidence that large-scale interracial contact reduces prejudice on a long-term basis.

The remainder of the paper is organized as follows. Section 2 describes the origins of racial segregation in the U.S. military and why, during the Korean War, this system was ultimately abandoned. Section 3 presents the casualty data and the six other datasets that veterans are matched to. Section 4 introduces the casualty similarity index as a measure of wartime integration and outlines the empirical strategy. Section 5 investigates outcomes during the war—namely, the effects of integration on casualty survival. Section 6 shifts the focus to the post-war period by investigating where integrated veterans ultimately chose to reside and whom they married. Section 7 concludes.

2 Historical Background

2.1 The origins of racial segregation in the military

Blacks served in major U.S. conflicts, beginning in the colonial militias and continuing through the American Revolution (Mershon and Schlossman, 1998). Following the founding of the United States, the history of blacks in the military closely mirrors that of American civil society. In 1783, federal and state legislators passed laws to restrict military service to whites only, though these prohibitions were ignored to meet manpower needs in times of extreme crisis. During the Civil War, 186,000 blacks served in segregated all-black units (Bowers et al., 1997). In the years following the Civil War, public interest in advancing civil rights for both free blacks and newly liberated slaves ran high, and the position of blacks in the military improved considerably.

Congress passed measures in 1866 which guaranteed blacks the right to serve, and in 1869, which required the Army to maintain four all-black regiments on a permanent basis. A federal statute in 1867 made blacks eligible for militia duty, and many states—all of which had previously excluded non-whites except during emergencies—thereafter admitted blacks to serve in all-black militias. These Reconstruction Era (1863-1877) laws reflected the new paradigm apparent in other public institutions including schools, transportation and hospitals: blacks had to accept a *de facto* “separate but equal” doctrine in order to gain access to public benefits. This doctrine was codified into law by the U.S. Supreme Court in the famous 1896 case *Plessy v. Ferguson*, which established the legal precedent for racial policy in both the military and civilian society for the next seventy years.

2.2 Blacks in the military during WWI and WWII

The late-nineteenth century saw a resurgence in racial hostility, and the white public’s enthusiasm for enforcing the rights of blacks waned. From 1890 to 1910, several states once again excluded blacks from militia service. But the looming American involvement in World War I in April 1917 rekindled the national debate concerning the role of blacks within the military. By this time, blacks made up 10% of the national population, and it was clear that the manpower required to support the war effort in Europe would need to include blacks. Yet Army officials were vehemently opposed to placing blacks in combat roles. Those blacks who did join the American Expeditionary Forces did so in segregated units, most of which were confined to non-combat support roles. Even the four traditional all-black regiments were outright excluded from deployment in Europe (Dalessandro et al., 2009).

2.3 Black activism and Executive Order 9981

Between 1916 to 1940, 1.6 million blacks relocated from the rural South to urban centers in the North and hundreds of thousands more to industrial centers elsewhere in the South. These massive demographic shifts dramatically altered the American political landscape.

Although race relations remained an afterthought for most whites, blacks began to grow as a political force, as those who left the South were now free from poll taxes, literacy requirements, and violence. Blacks increasingly demanded a more prominent role within the military (see Figure 1), which many viewed as a frontier in the battle for civil rights. In response to mounting political pressure, the military removed barriers to black participation, with the Navy even going so far as to integrate its training programs. However, the Army only expanded black opportunities within the circumscribed limits of racial segregation,⁸ and the Marine Corps continued to exclude blacks outright.

The growing activism and political influence of the black electorate began to turn the tide of racial policies within the military. In July 1948, President Harry Truman issued Executive Order 9981, which stipulated equal treatment and opportunity regardless of race or ethnicity across all branches of the military. Truman made it clear that this decree ultimately meant blacks and whites were to be integrated. Like subsequent integration directives in the context of public schools, however, Executive Order 9981 was too vague to have an immediate impact. Although the Air Force did begin integrate, the Army and the Marine Corps remained largely unchanged.

2.4 The Korean War (1950-1953)

Despite the fact that the legal groundwork for integration had been laid by the outset of the Korean War in June 1950, blacks and whites nevertheless remained rigidly segregated at the battalion level down (see Table 1). Traditional Army structure is formed in threes: three battalions to a regiment, three companies to a battalion, and so on. This meant that a regiment could contain, say, two all-white battalions and one all-black battalion, but smaller units were either all-white or all-black.⁹

In addition to being segregated across units, blacks and whites were also segregated by

⁸ A brief exception occurred following the Battle of the Bulge (December 1944-January 1945), when 2,500 blacks were reassigned to depleted all-white infantry companies, though these units remained segregated at the platoon level.

⁹ All-black units could have either black or white officers.

occupation. Most of the combat roles were assigned to all-white units, while black units were largely confined to non-combat support roles. This occupational segregation stemmed in part from official dissatisfaction with the performance of the all-black 24th Regiment early in the war (Bowers et al., 1997). These all-black units were typically over-strength. One of the few advances in racial policy the Army had made during the intra-war years was to remove race quotas that limited black manpower; yet because Truman's executive order prohibited the formation of new all-black units, the Army simply over-staffed those already in existence.

The early years of the Korean War were one of the most deadly campaigns in American military history, and many of the all-white combat units became severely depleted. This can be seen in Figure 2, which plots the monthly casualties throughout the war, and shows that the majority of both black and white casualties occurred early on.

In a bid to bolster manpower, some commanders of severely depleted all-white units reached informal agreements with personnel officers to accept black soldiers as replacements. One of those units was the 9th Regiment. The Lieutenant Colonel in command of the 9th Regiment's 1st battalion describes his decision to accept black reinforcements:

I was very, very low on men—less than half strength—and raised hell to get more troops. The division [personnel officer] called and, knowing that I had previously commanded a battalion of black troops [in the Twenty-fifth Infantry], said he had almost two hundred who would transfer to the infantry if they could serve with me. I agreed. In fact, I was proud to have them. [The Second Division commander, Major General] Keiser asked me if I realized what a can of worms I was opening up, to which I said, “So what? They are good fighting me. I need men.”

One feature evident from this passage is that the initial decision to integrate was made by high-ranking officers. This meant that whichever units accepted blacks had more to do with the idiosyncrasies of a few commanders than the average characteristics of the thousands of

enlisted men in those units.

Another turning point in the road to integration was the appointment of General Matthew Ridgway as the commander of the Eighth Army in January 1951. In Ridgway’s view, segregation was “both un-American and un-Christian for free citizens to be taught to downgrade themselves in this way, as if they were unfit to associate with their fellows or to accept leadership themselves.” In May 1951, Ridgway sent a telegraph to the Pentagon outlining his plans for Army-wide integration. It was to begin with infantry units, followed by non-infantry combat units, and finally non-combat support units.

In October 1951, the all-black 24th Regiment was disbanded and redistributed across the Army. One of the novel features about troop replacement during the Korean War is that for the first time, reinforcements were introduced individually instead of being rotated in as entire units. Integration proceeded in both directions, with blacks being gradually introduced into previously all-white units, and whites being gradually introduced into previously all-black units.

3 Data

The analysis considers three main outcomes of interest—namely, casualty survival, residential sorting and intermarriage—of which the latter two require merging multiple datasets. In this section, I describe these datasets and explain how they are merged; for a summary, Figure 3 shows which datasets are linked to which.

3.1 Data on integration and survival: Korean War casualty file

The Office of the Adjutant General’s Korean War (TAGOKOR) casualty file contains individual records on all 103,000 U.S. Army officers and enlisted men who were recorded as casualties (i.e. killed or wounded) during the Korean War. This casualty information was recorded in theater using IBM punch cards (Figure 4) by teams of specialized units

known as Mobile Machine Records Units (MRUs). These MRUs were dispatched to collect logistical information on individual units—including casualties—which ensured casualties were recorded within days of their occurrence.

TAGOKOR is unique among public-use American casualty files in that it is the only one that reports individual-level information on killed *and* living soldiers. This information on Korean War survivors is what makes it possible to study the behavior of a subset of soldiers after the war. Broadly speaking, the data in TAGOKOR fall into two categories: military and civilian. The military data details twenty different categories of casualty, including those killed, wounded, hospitalized, missing in action, and captured. Moreover, it records each soldier’s rank, military occupational specialty (e.g. “cook, rifleman”), service number, as well as their race, regiment and exact date of casualty. These last three pieces of information are crucial because they enable me to construct an integration measure for each regiment at different points in the war. In terms of civilian data, TAGOKOR details each soldier’s first and last name, middle initial and county of residence at the time of the war. Year of birth is only recorded for soldiers who were killed in action (KIA), but this information turns out to still be of use for matching surviving soldiers to their post-war outcomes. Table 3 contains an example of a typical observation contained in the TAGOKOR casualty file.

The summary statistics reported in Table 4 highlight the circumstances under which racial integration took place. Panel A summarizes each variable in the period before formal integration was announced on on October 1, 1951. Comparing columns (1) and (2), we see that whites were 3.1 percentage points less likely to survive their injuries than blacks, reflecting the initial occupational segregation between the two groups, and the need to reinforce depleted all-white battalions early in the war. In Panel B, we see this same casualty survival gap decrease to 0.5 percentage points, as blacks gradually joined whites in more hazardous combat roles. That the gap remains positive is evidence that black lives were not gratuitously put at risk once they were integrated.

Table 4 also speaks to the possible importance of white flight in the Korean War context.

Although soldiers already stationed in the Korean peninsula had little control as to whom they served with, it is possible that potential volunteers with a distaste for integration were dissuaded from joining the Army as news of racial integration reached the home front. Indeed, as the war progressed, we see by comparing across Panels A and B that the number of whites soldiers originating from the South—historically the most prejudiced region in country—declined by 6.6 percentage points. However, the bulk of this change stems from the introduction of the draft. Of the 5 million American soldiers who served during the Korean War, 1.5 million were draftees. These draftees were overrepresented in the Army relative to other branches of the armed forces because volunteers usually chose to enter the Navy or the Air Force. A unique feature of TAGOKOR is that one can indirectly observe the soldier’s draft status. TAGOKOR reports each soldier’s service number, and those with the prefix “US” were draftees. In columns (3) and (4), I restrict the sample to soldiers who volunteered for the Army (i.e. prefix not equal to “US”). Now we see that the fraction of volunteers who came from the South only declined by 2.7 percentage points. To further mitigate sample selectivity concerns, I repeat my analysis of residential sorting below with a sample restricted to draftees.

3.2 Data on last zip code of residence: Social Security Death Index

The Social Security Death Index (SSDI) contains records for over 90 million Americans who died between 1940 to 2010. This file was created from the Social Security Administration’s (SSA) Death Master File extract. It contains records for deceased individuals who had been assigned social security numbers (SSN) and whose deaths were reported to the SSA. From 1973 onward, the SSDI covers between 93% to 96% of deaths of individuals aged 65 or older (Hill and Rosenwaike, 2001). This high coverage rate for older cohorts implies that the vast majority of Korean War veterans appear in these records.

Table 5 presents a typical observation contained in the SSDI. The SSDI does not contain

any military information that could be used to uniquely identify Korean War veterans. However, the SSDI does contain enough personally identifiable information for it to be linked to veterans in the casualty file. In particular, for each deceased individual, the SSDI reports first and last names, middle initial (if death occurred after 2000), the state where SSN was issued, and dates of birth and death.

The SSDI also reports the last zip code of residence prior to death, which is the geographical unit at which I model a veteran's residential choice. For this model, the outcome of interest is the extent of racial integration in the veteran's last zip code. For each zip code, I construct a residential similarity index using block-level census data. More formally, let B_{bt} (W_{bt}) be the number of blacks (whites) in Census block b in year t , and let D_z be the set of blocks that belong to zip code z . The total number of blacks and whites in zip code z in year t is given by

$$B_{zt}^{Total} \equiv \sum_{b \in D_z} B_{bt}$$

$$W_{zt}^{Total} \equiv \sum_{b \in D_z} W_{bt}.$$

The residential similarity index for zip code z in year t is defined as

$$ZipSim_{zt}^{BW} \equiv 1 - \frac{1}{2} \sum_{b \in D_z} \left| \frac{B_{bt}}{B_{zt}^{Total}} - \frac{W_{bt}}{W_{zt}^{Total}} \right|. \quad (1)$$

By construction, $ZipSim_{zt}^{BW}$ can take on values between 0 and 1. The term to the right of the minus sign in Equation (1) is the familiar dissimilarity Index first proposed by Duncan and Duncan (1955). Hence $ZipSim_{zt}^{BW}$ is the complement the dissimilarity index. An $ZipSim_{zt}^{BW}$ value of 0 indicates complete segregation (i.e. none of the census blocks contain both blacks and whites), and a value of 1 indicates complete integration (i.e. the distribution of blacks and whites across census blocks is identical). Intuitively, $ZipSim_{zt}^{BW}$ is the proportion of blacks (or whites) who would have had to have be relocated to other

Census blocks in order to completely segregate the two groups.

To simplify the analysis, I assume that a veteran chooses which housing market to reside in (e.g. New York-Newark-Jersey City versus Santa Fe, New Mexico) for reasons unrelated to his racial preferences—be it work opportunities, family ties, and so forth. Here I define a housing market as a core-based statistical area (CBSA) so that the veteran’s choice set is limited to all zip codes within a given CBSA. I account for the veteran’s choice set by normalizing $ZipSim_{zt}^{BW}$ within a CBSA. In the case of, say, the New York housing market, the within-CBSA normalized zip code similarity index y_{zt} is given by

$$y_{zt} = \frac{ZipSim_{zt}^{BW} - E_{NY}(ZipSim_{zt}^{BW})}{\sigma_{NY}(ZipSim_{zt}^{BW})}. \quad (2)$$

The distribution of y_{zt} is shown in Figure 5. That y_{zt} is roughly symmetrical indicates that, within a housing market, the similarity index is distributed symmetrically. I also use the block-level data to compute and normalize the black share of each zip code as well. To account for neighborhood quality, I matched zip codes to annual house price data published on Zillow.com.¹⁰

In summary, I link casualty data to social security data, which in turn I link to Census and house price data. The linked dataset contains 16,081 individual Korean War veterans with data on individual characteristics (e.g. age, race), wartime characteristics (e.g. measure of integration, rank, occupational specialty, county of residence), the proportion of blacks and whites in the last zip code of residence, and the extent of integration between these two groups.

3.3 Data on wives: national cemeteries

The National Cemetery Administration maintains and updates data on all military veterans and their dependents who are buried in one of the 138 national cemeteries across the

¹⁰ House price data is available at <https://www.zillow.com/research/data/>.

United States. As of November 2018, these data contain records on 250,000 unique veteran/wife pairs for a sub-sample of Korean War Army veterans. Each record details both the veteran’s and his wife’s first and last name, middle initial, date of birth, date of death and the veteran’s military rank. Table 6 details a typical observation contained in a national cemetery record. To assign an integration measure to these veterans in national cemeteries, I link them to individuals soldiers in the casualty file using an unsupervised learning method. A limitation of this dataset is that it does not contain information on the race of the wife, and matching directly to administrative data that contains race such as SS-5 application forms decimates the sample size. To avoid this problem, I construct a “blackness/non-whiteness” name index for each wife. Mathematically, this name index is the probability that the wife is black/non-white given her first name, year of birth and state of burial. Plots of these data are shown in Figure 6.

4 Methodology

4.1 Variation in integration across time and regiments

The rich casualty data contained in TAGOKOR make it possible to quantify the extent of wartime racial integration; intuitively, if black and white soldiers were physically placed in the same foxholes, then their injuries should correlate. This phenomenon is evident in Figure 7, which plots daily black and white casualties in the early and late stages of the war. During the first three months of combat—at a time when blacks and whites were rigidly segregated in units of battalion-size (about 800 soldiers) or smaller—there is no discernible relationship between the timing of black and white casualties ($\rho = 0.02$, see Figure 7a). But these casualty patterns change sharply in the last year of the war. By this time, integration was virtually complete, and the relationship between black and white daily casualties is practically linear ($\rho = 0.91$, see Figure 7b).

This stark contrast is not driven by some unobserved trends, as integration also varied

considerably across regiments within a given time period. To illustrate this point, the story of the 9th Regiment presents a useful case study. The 9th Regiment was composed of three battalions: the all-white 1st and 2nd Battalions and the all-black 3rd Battalion. After the two all-white battalions sustained heavy losses in the first three months of the war, these units began to accept black reinforcements from the all-black 3rd Battalion. The photograph in Figure 8 shows a newly-integrated fire team from the 9th Regiment in a foxhole in November 1950. By this time, the other thirty regiments had not begun to integrate. Figure 9a plots daily black and white casualties between October 1950 to December 1950 for all regiments *except the 9th*. As expected, the casualty pattern is nearly identical to that of the previous three months (Figure 7a). For comparison, Figure 9a only plots daily black and white casualties between October 1950 to December 1950 for the newly-integrated 9th Regiment. The casualty pattern now bears a striking resemblance to that of the last year of the war, by which time racial integration was a *fait accompli*.

4.2 Measuring integration

Throughout this paper, I quantify these plausibly exogenous changes in wartime integration using a regimental¹¹ similarity index. Let B_{rd} and W_{rd} be the number black and white casualties in regiment r on day d of the war. I divide the war into six equal periods denoted by $t \in \{1, \dots, 6\}$ —i.e. six periods of 185 days, or 37 months divided by 6. Let D_t denote the set of days in period t (e.g. $D_1 = \{1, \dots, 185\}$). The total number of blacks/whites casualties in regiment r during period t of the war is given by

$$B_{rt}^{Total} \equiv \sum_{d \in D_t} B_{rd}, \quad (3)$$

$$W_{rt}^{Total} \equiv \sum_{d \in D_t} W_{rd}, \quad (4)$$

¹¹ I use the term “wartime integration” or “regimental similarity” to distinguish this measure from the *residential* similarity of zip codes. The former is an explanatory variable, while the latter is an outcome.

for $t \in \{1, \dots, 6\}$. We can then define similarity index for period t as

$$s_{rt}^{BW} \equiv 1 - \frac{1}{2} \sum_{d \in D_t} \left| \frac{B_{rd}}{B_{rt}^{Total}} - \frac{W_{rd}}{W_{rt}^{Total}} \right|. \quad (5)$$

By construction, s_{rt}^{BW} can take on values between 0 and 1; its distribution is plotted in Figure 10. An s_{rt}^{BW} value of 0 indicates complete segregation (i.e. blacks and whites are never injured on the same day), and a value of 1 indicates complete integration (i.e. the distribution of casualties across days is identical for blacks and whites). Intuitively, s_{rt}^{BW} is the proportion of blacks (or whites) who would have had to have been injured on different days in order to completely segregate the two groups.

5 Results: Wartime Outcomes

5.1 Integration and casualty survival rates

During the Korean War, a wounded soldier would first be attended to at a battalion aid station. Each aid station typically contained one doctor and a mix of medics and litter bearers (i.e. soldiers who carried stretchers). The wounded would then be transported behind the front line to a larger collection station. Once the wounded had been stabilized, they would then be transported to a Mobile Army Surgical Hospital (MASH) unit or a division clearing station, depending on the type of wound, after which they were transported to an evacuation hospital. In the most serious cases, wounded soldiers would be airlifted to Japan for additional care.

The battalion aid station—the first line of care—was located only a few hundred yards from the front line. This proximity to the front line meant that the success of any evacuation attempt depended critically on battle conditions as medical evacuation vehicles routinely came under enemy fire. As one Army medic put it: “We were always in danger of being attacked by the enemy, overrun by the enemy, being shelled by artillery, shelled by mortar,

and grenades thrown into the station.” Thus, a successful evacuation depended critically on the ability of nearby combat troops to provide “suppressive fire.”

Early in the Korean War, many all-white battalions were severely depleted, which would have impaired evacuation efforts. Those all-white units that were willing to accept black troops and thus be returned to “safer” sizes would have been better able to evacuate wounded troops to safety. By contrast, for blacks, integration meant being pulled away from support roles and thrust into more dangerous combat roles. Overall, then, integration should have increased the probability that whites survived their injuries relative to segregated whites, as measured by the “white casualty survival rate”. Likewise, integrated blacks would have been less likely to survive relative to blacks who remained in support roles.

I test this relationship with a linear model:

$$y_{irt} = \beta_0 + \beta_1 s_{rt}^{BW} + \mathbf{X}_{irt} \boldsymbol{\beta}_2 + \lambda_t + \gamma_r + \varepsilon_{irt}, \quad (6)$$

where y_{irt} is a dummy for whether soldier i in regiment r in period t survives their injury, s_{rt}^{BW} is the normalized casualty similarity index, \mathbf{X}_{irt} is a vector of controls, and λ_t and γ_r are time and regiment fixed effects, respectively.

When these fixed effects are included, Equation (6) is a difference-in-differences model in spirit, only that any departure in parallel trends is a result of integration. An implication of this model is that during periods when integration levels were steady, two regiments should have parallel survival rates. It is infeasible to pinpoint periods where integration levels remain steady because the integration measure I construct moves too infrequently (i.e. every six months) to demonstrate parallel trends convincingly. However, information on the exact date that particular units were integrated is available for a handful of units. For example, Figure 11 shows a screen shot of the 9th Infantry Regiment’s war diary for September 1950. The text indicates that the all-black 3rd battalion was reunited with the badly depleted 1st and 2nd battalions of the 9th Infantry on September 16, 1950, after having spent the first eight

weeks of the war away from combat on reserve duty. Figure 12 plots the time series of the survival rates of the 9th and 23rd Infantry regiments which—as members of the 2nd Infantry Division—cooperated closely in combat throughout the war. The first dashed vertical line denotes September 16, 1950 (i.e. day 83 of the war), the day that the three battalions of 9th Regiment were united; by contrast, the 23rd Regiment did not have any all-black units. The second vertical dashed line denotes October 1, 1951 (i.e. day 463 of the war), the date that formal integration began. Two features are apparent from this figure. First, after formal integration began, the two regiments exhibit parallel trends. Second, there is some evidence that survival rates improved for whites in the 9th Regiment immediately after being joined by the all-black 3rd battalion.

The results of estimating Equation (6) are reported in Table 7. In Panel A, column (1), we see that a one standard deviation increase in the integration measure increased white survival probability by 2.2 percentage points. To account for the fact that both integration and survival rates increased as the war progressed, column (2) includes period fixed effects. Columns (3) then adds regiment fixed effects. The preferred specification is column (4), which adds individual-level information on rank and military occupational specialty. Holding a rank of private meant that a wounded soldier was 1.5 percentage points less likely to survive his injury, which reflects the fact that privates were put in more hazardous positions. Here we see that a standard deviation change in integration increased casualty survival by 2.2 percentage points. Given that mean survival rate for whites was 75%, this represents an increase of 2.9%. Panel B repeats this exercise for blacks. We see in column (4) that, as predicted, integration is associated with a *decrease* in black survival rates by 1.8 percentage points, or 2.3%.

That the direction of the signs for blacks and whites differs is to be expected given that integration implied a convergence in hazardous combat assignments. What is less obvious is the effect of integration on *overall* survival rates. I test for this directly by estimating Equation (6) using a pooled sample of blacks and whites. Table 8 reports results. In each

specification, integration has a positive effect on overall survival rates ranging from 1.9-2.4 percentage points, or 2.4-3.2%. In column (1), the coefficient on black is strongly positive, which is in part due to race being correlated with unobserved combat exposure. Column (2) adds regiment fixed effects, which attenuates the coefficient on black slightly, though it remains significantly different from zero. Recall that segregation occurred at units of battalion size or smaller, which means controlling for regiment does not completely account for differences in combat exposure by race. Column (3) adds period fixed effects, which further attenuates the black coefficient because over the course of the war, the share of blacks increased as the intensity of combat waned. Column (4) controls for private status and having a support military occupational specialty, and again the black coefficient remains positive though marginally significant.

Columns (5)-(6) explore the relationship between integration and the safety of blacks more directly. Column (5) is restricted to a sample of regiment-period pairs whose integration measure belong to the bottom quartile. In these relatively segregated regiments, being black was highly predictive of survival. By contrast, column (6) is restricted to a sample of regiment-period pairs whose integration measures belong to the top quartile. Here we see that blacks were no more or less likely to survive their wounds than whites, and the null hypothesis of no discrimination cannot be rejected. Across all specifications, holding a rank of private is strongly predictive of fatality, which is expected given that these soldiers were exposed to more dangerous roles and were the least experienced.

The last two results in Table 8 are apparent in Figure 13c, which shows the casualty survival rates of blacks and whites over time. The black casualty survival rate begins elevated and exceeds that of whites throughout most of the war. In the absence of discrimination, we should not expect to see a strong relationship between integration and survival because toward the end of the war, *all* units were integrated as a matter of policy, and not to reinforce vulnerable units.

6 Results: Post-War Outcomes

The integration that soldiers experienced during the Korean War satisfied the four conditions for the contact hypothesis outlined by Allport (1954). In particular, black and white soldiers had (i) equal status (conditional on rank) (ii) common goals (iii) a cooperative setting (iv) support from authorities. Recall that Allport was interested in preference formation from a sociological perspective. Given the four Allportian conditions were satisfied, this raises the question to what extent, if any, did integration reduce the prejudicial attitudes of veterans?

6.1 Integration and residential choice

6.1.1 A linear model for residential choice

In this subsection, I estimate the effect of wartime integration on the residential similarity of each soldier’s last zip code of residence. To simplify the analysis, I assume that a soldier chooses which housing market to reside in (e.g. New York City versus Santa Fe, New Mexico) for reasons other than his racial preferences—be it work opportunities, family, and so forth; I assess the plausibility of this simplifying assumption later in this section. But when choosing which neighborhood (i.e. “zip code”) to reside in *within* a given housing market, the veteran’s racial preferences enter into his utility function. Thus, in this subsection I model neighborhood choice conditional on having already chosen a housing market.

To capture this within-housing market neighborhood choice, I estimate a linear model:

$$y_{irstz}^{BW} = \beta_0 + \beta_1 s_{rt}^{BW} + \mathbf{X}_i \beta_2 + \mathbf{Z}_{iz} \beta_3 + \mathbf{C}_c \beta_4 + \gamma_r + \lambda_t + \zeta_s + \varepsilon_{irstz}, \quad (7)$$

where y_{irstz} is the residential similarity index (within-CBSA z score) for soldier i who served in regiment r , resided in state s during the war, was wounded in period t , originated from county c , and last resided in zip code z , s_{rt}^{BW} is the normalized integration measure, \mathbf{X}_i is a

vector of individual-level controls, \mathbf{Z}_{iz} is a vector of last zip-code level characteristics, and γ_r , ζ_s and λ_t are regiment, state-of-origin, and period-of-casualty fixed effects, respectively.

For ease of exposition, I will refer to the similarity index of the last zip code of residence as “residential similarity,” and the similarity index of wartime integration as the “treatment.”¹² Table 9 reports estimation results for equation (7) for a sample restricted to white veterans. In column (1) we see that a one SD increase in treatment leads to a 0.06 SD increase in residential similarity. Column (2) additionally controls for the share of casualties in a regiment that were black, which approximates the share of blacks that served in that regiment. Interestingly, the share of blacks has no explanatory power, whereas the treatment effect barely changes. Simply put, increasing the share of blacks in a regiment does nothing to influence residential sorting patterns if veterans remain segregated within that regiment. Column (3) adds additional information about the last zip code of residence: average house prices and the share of blacks—both normalized within-CBSA—as well as their respective quadratic terms. Clearly, house prices are endogenous since they correlate with unobserved neighborhood quality. The positive coefficient on house prices indicates that, conditional on racial composition, high-quality neighborhoods are more integrated. The treatment effect is attenuated to 0.04 but remains statistically significant. Column (4) adds controls for county of origin, which indicate that veterans from poorer, blacker and more rural counties are less integrated after the war. Crucially, adding these controls does not alter the estimated treatment effect, which is further evidence that racial integration during the Korean War was quasi-random. Intuitively, this regression implies that if one were to take two white soldiers who are close in age, come from similar counties within the same state, and served during the same period of the war, then the white veteran who was more integrated during the war is more willing to live in integrated neighborhoods after the war.

One potential threat to this empirical exercise is that veterans were integrated on the basis

¹² Strictly speaking, the wartime integration measure applies to the regiment in which soldiers served, not individual soldiers *per se*. As such, OLS estimates of Equation (7) more closely resemble intention-to-treat (ITT) effects. Typically, ITT effects represent lower bounds on the treatment effects they approximate in terms of magnitude.

of some unobserved factor that is also correlated with residential similarity. For example, perhaps soldiers from counties that were more integrated to begin with were prioritized for integration during the war. I explore this possibility in column (5), which is restricted to a sub-sample of veterans who last resided in a different state than the one they were living in at the time of the war. For these veterans, the treatment effect is somewhat larger (0.042) and remains statistically significant, whereas the county of origin controls lose their explanatory power.

Table 10 shows the estimation results for Equation (7) for black veterans. None of the treatment effects are statistically significant. That the sample sizes are about 10% that of whites means these regressions are under-powered. Moreover, the true effect size for blacks may be smaller than that of whites because blacks faced imposing barriers to their residential choices. In the years following the Korean War, commercial banks practiced so called “redlining” to systematically deny mortgages to black applicants. Likewise, zoning laws such as restricting the supply of rental properties disproportionately excluded potential low-income black residents. Although the treatment effects are statistically insignificant, it is worth noting that, as Table 9 showed for whites, the treatments effects here are stable across columns (1)-(2) and columns (3)-(4), which indicates that blacks were similarly assigned to integration on a quasi-random basis.

6.1.2 Residential choice and sample selectivity

A key challenge in studying the effects of interracial contact on prejudicial attitudes is selection bias. Even if interracial contact has a negligible effect on prejudice, but highly prejudiced individuals self-select out of contact, then standard statistical analysis will yield upwardly-biased coefficient estimates. Such sample selectivity is pervasive in many large-scale desegregation episodes. As discussed earlier, a common response to both the desegregation of U.S. public schools and the migration of millions of blacks to predominantly white regions of the U.S. was so-called “white flight”—that is, white out-migration in response to

black in-migration.

Several features of the Korean War mitigate these concerns. First, soldiers in Korea had little control as to which units they served in, particularly low-ranking combat soldiers. In my sample, 80% of soldiers held the two lowest ranks, Private or Corporal. It is nevertheless possible that integration in Korea led to a form of preemptive white flight: as word about integration reached the general public, white civilians who might otherwise have enlisted in the Army instead refused to serve alongside blacks. Two historical facts mitigate these concerns: first, the Army attempted to conceal the fact of integration. As Mershon and Schlossman (1998) put it, “The Army and the Truman administration, fearful of possible hostility to this momentous and rapid institutional change, went to great lengths to minimize publicity and maximize secrecy during the early 1950s.”

Second, as the war progressed, the Army came to increasingly rely on the draft to meet manpower requirements. This increased reliance on the draft is evident in my casualty sample. Between the first to last year of the war, the share of draftees in my casualty file climbed from 6% to 40%. If preemptive white flight is present in my sample, then the treatment effects for volunteers should be smaller than those of draftees.

I test this possibility directly by splitting the sample between volunteers and draftees and re-estimate Equation (7). Table 11 reports results for white veterans. Since paternity and marital deferments were in use, I restrict my sample to men who were younger than 25 years of age at the time of service. Columns (1)-(3) report results for white volunteers, and columns (4)-(6) report results for white draftees. Two things are notable: first, contrary to the white flight hypothesis, the estimated treatment effect for white draftees exceeds those of volunteers. Although the treatment effects for draftees are not statistically different from zero, a separate pooled regression indicates that the treatment effects do not differ significantly by draft status. Second, the effects sizes appear to be larger for this younger sample—ranging from 0.11-0.20 than the full sample in Table 9—which ranges from 0.04-0.06. These larger point estimates for young veterans point to heterogeneity in treatment

effects. These heterogeneous effects could stem from a number of sources. One possibility is that older veterans were already homeowners and as such less geographically mobile. Another possibility is that younger veterans had less social status, which contact theory predicts would make them more responsive to intergroup contact. With the data currently available to me, I am unable to distinguish between these distinct mechanisms.

Table 12 reports analogous estimation results for black veterans. Comparing columns (1)-(3) to columns (4)-(6), we again see no evidence that volunteers have larger treatment effects than draftees, although here again the much smaller sample size reduces statistical power.

6.2 Integration and intermarriage

6.2.1 Linear and fractional response models for choosing a spouse

A high bar for prejudice reduction is exogamy—an individual’s willingness to marry outside of their own race or ethnicity. I test the relationship between integration and exogamy with two models. The first is a linear model, which I estimate by OLS:

$$y_{irst} = \beta_0 + \beta_1 s_{rt}^{BW} + \mathbf{X}_i \beta_2 + \gamma_r + \lambda_t + \zeta_s + \varepsilon_{irst} \quad (8)$$

where y_{irst} is the wife’s non-white (or black) name index for soldier i who served in regiment r , resided in state s during the war, and was wounded in period t , s_{rt}^{BW} is the normalized integration measure, \mathbf{X}_i is a vector of individual-level controls, and γ_r , ζ_s and λ_t are regiment, state or origin, and period of casualty fixed effects, respectively.

By construction, the wife’s name index y_{irst} is bounded between zero and one (recall Figures 6a and 6b). In order to account for this characteristic of the left-hand-side variable,

I specify the following functional form for the expectation of the name index y_{irst} :

$$\begin{aligned} E(y_{irst}|s_{rt}^{BW}, \mathbf{X}_i, \gamma_r, \lambda_t, \zeta_s) &= \Phi(\beta_0 + \beta_1 s_{rt}^{BW} + \mathbf{X}_i \boldsymbol{\beta}_2 + \gamma_r + \lambda_t + \zeta_s), \\ &\equiv \Phi(\mathbf{W}_{irst} \boldsymbol{\Theta} + \gamma_r + \lambda_t + \zeta_s), \end{aligned} \quad (9)$$

where Φ is the standard normal cumulative distribution function, \mathbf{W}_{irst} denotes the row vector $[1 \ s_{rt}^{BW} \ \mathbf{X}_i]$, and $\boldsymbol{\Theta}^\top$ denotes the parameter vector $[\beta_0 \ \beta_1 \ \boldsymbol{\beta}_2^\top]$. I estimate the parameters in Equation (9) using quasi-maximum likelihood, where the likelihood for an observation is specified as the Bernoulli likelihood:

$$L_i(\boldsymbol{\Theta}, \gamma_r, \lambda_t, \zeta_s) = [\Phi(\mathbf{W}_{irst} \boldsymbol{\Theta} + \gamma_r + \lambda_t + \zeta_s)]^{y_{irst}} [1 - \Phi(\mathbf{W}_{irst} \boldsymbol{\Theta} + \gamma_r + \lambda_t + \zeta_s)]^{1-y_{irst}}. \quad (10)$$

The quasi-maximum likelihood estimates (QMLE) of $\boldsymbol{\Theta}, \gamma_r, \lambda_t, \zeta_s$ are consistent as long as the conditional expectation in Equation (9) is correctly specified even if the Bernoulli specification in Equation (10) is incorrect.¹³ The asymptotic variance-covariance matrix of the QMLE estimator only requires first-moment assumptions.

Estimation results for both the linear and fractional response model for white veterans are reported in Table 13. The sample is restricted to men who were 22 years of age or younger during the war. This age cut off is based on the median age of first marriage in 1950, which for men was 22.8 years.¹⁴ In Panel A, the dependent variable is the wife's non-white name index. In column (1), we see that a one standard deviation change in integration corresponds to a 1.8 percentage point increase in the name index, or 0.8 SD. This effect size exceeds that from a one standard deviation change in the black share of the county of origin (0.9 of a percentage point, or 0.04 SD). The age of the soldier does not have a statistically significant effect, which is to be expected because the sample is restricted by age. Column (2) adds rank fixed effects, which barely change the estimated effect sizes. Column (3) includes

¹³ See Papke and Wooldridge (1996) for more details regarding this approach

¹⁴ See <https://www.thespruce.com/estimated-median-age-marriage-2303878>

state-of-residence fixed effects. The effect size of integration again barely changes, although the effect of the black share of their county of origin becomes negative and statistically significant. Column (4) estimates the fraction response model in Equation 9. The marginal effect of 3.1 percentage points is comparable to the OLS results. Panel B repeats this analysis using the wife’s black name index as the dependent variable; results are broadly the same.

6.2.2 Placebo test results

If racial integration indeed made veterans more exogamous, then one should not observe any effect on veterans who were already married. Since I do not observe marital status before the war, I use age as a proxy for marital status. Table 14 reports estimation results for Equation (9) for a sample of men who were at least 23 years of age during the Korean War. Now we see that across specifications, the treatment effect is no longer statistically different from zero, even though the sample size is 36% larger. Notably, age at casualty now has a strong negative effect on the wife’s name index. As with residential sorting, the fact that the treatment effect is statistically significant for the younger sample but not the older sample—and that age is strongly negative within the older sample—has at least two mutually-compatible explanations: (i) interracial contact is more effective at shaping the tastes and/or beliefs of younger individuals, and (ii) older veterans were more likely to already be married. Again, the data currently available to me is not sufficiently rich to distinguish between these mechanisms.¹⁵

7 Conclusion

The racial integration of the U.S. Army during the Korean War was one of the largest and swiftest desegregation episodes in American history. Integration began informally, as a

¹⁵ The full count 1950 Census is scheduled to be released in April 2022, which will contain valuable spouse data. Future work could restrict the sample to soldier who were not married as late as 1950 and identify treatment effects off of men who served early in the war.

bid to reinforce badly depleted all-white units, and by the end of the war had become Army-wide policy. Unlike the desegregation of U.S. public schools—which was initiated expressly for the benefit of blacks—the Army integrated for reasons of military efficiency.

The first part of my paper examines whether the Army managed to improve efficiency as measured by the survival rates of wounded soldiers. To this end, I developed a novel wartime integration measure to quantify exogenous changes in racial integration across time and regiments. Using a modified difference-in-differences strategy, I found based on a pooled sample of blacks and whites that a standard deviation increase in integration caused a 3% increase in the *overall* survival rates of wounded soldiers. On the whole, these estimates indicate that the Army managed to improve efficiency once it abandoned racial segregation and could freely allocate scarce labor (soldiers). These survival rates are admittedly only a partial measure of overall military efficiency, as integration may have improved combat efficiency at the expense of some unobserved tasks. Therefore, in future work I will consider broader measures of efficiency such as casualty rates, unit citations, other-than-honorable (OTH) discharges, and awarded medals.

Historically, large-scale desegregation episodes have often had unintended consequences, with a common response to both school desegregation and black in-migration being so-called “white flight.” The second part of this paper is therefore devoted to understanding the long-run effects of wartime integration by observing the behavior of Korean War veterans in their post-war civilian lives. In particular, by matching veterans to social security and cemetery data, I am able to observe where veterans lived and whom they married. Here I found that a standard deviation increase in wartime integration caused young veterans to live in neighborhoods that were 0.1 standard deviations more integrated—roughly the difference in integration between present-day Central Harlem and the Upper East Side. Similarly, a standard deviation increase in wartime integration caused veterans to marry spouses with names that were less distinctly white. In both cases, I observe substantial effect heterogeneity by age, though it remains unclear whether these patterns reflect age-specific changes in

prejudice or the fact that older veterans were more likely to have already made residential and marital decisions prior to the war. Overall, these results on veterans' post-war behavior support the view that large-scale interracial contact served to reduce prejudice on a long-term basis.

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8 Tables

Table 1: Glossary of Army units

| Units | Strength | Typical commander |
|--------------------------|-----------------|------------------------------|
| region/theater | 4+ army groups | six-star rank |
| army group front | 2+ field armies | five-star general |
| field army | 100,000-300,000 | general |
| corps | 30,000-50,000 | lieutenant general |
| division | 10,000-25,000 | major general |
| †regiment/brigade/legion | 1,000-5,500 | colonel/brigadier general |
| *battalion/cohort | 300-800 | lieutenant colonel |
| *company | 80-150 | captain/major |
| *platoon | 15-45 | second lieutenant/lieutenant |
| *squad/section | 8-14 | sergeant |
| *fireteam | 2-4 | lance corporal/corporal |

Note: *Formally segregated before Oct 1951. †I observe which regiment individual soldiers served in.

Table 2: Glossary of Army ranks

| Pay Grade | Rank | Abbreviation | Classification | 2017 Pay Range |
|-----------|----------------------------|--------------|------------------|-----------------------|
| E-1 | Private | PVT | Enlisted Soldier | \$1,600 (monthly) |
| E-2 | Private Second Class | PV2 | Enlisted Soldier | \$1,793 (monthly) |
| E-3 | Private First Class | PFC | Enlisted Soldier | \$22,630 - \$25,510 |
| E-4 | Specialist | SPC | Enlisted Soldier | \$25,067 - \$30,427 |
| E-4 | Corporal | CPL | NCO | \$25,067 - \$30,427 |
| E-5 | Sergeant | SGT | NCO | \$27,338 - \$38,794 |
| E-6 | Staff Sergeant | SSG | NCO | \$29,840 - \$46,220 |
| E-7 | Sergeant First Class | SFC | NCO | \$34,502 - \$62,010 |
| E-8 | Master Sergeant | MSG | NCO | \$49,633 - \$70,787 |
| E-8 | First Sergeant | 1SG | NCO | \$49,633 - \$70,787 |
| E-9 | Sergeant Major | SGM | NCO | \$60,631 - \$94,136 |
| E-9 | Command Sergeant Major | CSM | NCO | \$60,631 - \$94,136 |
| E-9 | Sergeant Major of the Army | SMA | NCO | \$60,631 - \$94,136 |
| W-1 | Warrant Officer 1 | WO1 | Warrant Officer | \$35,597 - \$61,510 |
| W-2 | Chief Warrant Officer 2 | CW2 | Warrant Officer | \$40,554 - \$67,687 |
| W-3 | Chief Warrant Officer 3 | CW3 | Warrant Officer | \$45,828 - \$80,392 |
| W-4 | Chief Warrant Officer 4 | CW4 | Warrant Officer | \$50,184 - \$93,478 |
| W-5 | Chief Warrant Officer 5 | CW5 | Warrant Officer | \$89,233 - \$116,770 |
| O-1 | Second Lieutenant | 2LT | CO | \$36,418 - \$45,824 |
| O-2 | First Lieutenant | 1LT | CO | \$41,958 - \$58,068 |
| O-3 | Captain | CPT | CO | \$48,560 - \$79,002 |
| O-4 | Major | MAJ | Field Officer | \$55,231 - \$92,218 |
| O-5 | Lieutenant Colonel | LTC | Field Officer | \$64,012 - \$108,752 |
| O-6 | Colonel | COL | Field Officer | \$76,784 - \$135,936 |
| O-7 | Brigadier General | BG | General Officer | \$101,257 - \$151,283 |
| O-8 | Major General | MG | General Officer | \$121,860 - \$175,676 |
| O-9 | Lieutenant General | LTG | General Officer | \$172,224 - \$186,998 |
| O-10 | General | GEN | General Officer | \$15,583 (monthly) |
| O-10 | General of the Army | GA | General Officer | \$15,583 (monthly) |

Note: Pay ranges are annual unless otherwise specified. NCO and CO denote “Noncommissioned Officers” and “Commissioned Officers” respectively. Source: <https://www.federalpay.org/military/army/ranks>.

Table 3: Typical observation in TAGOKOR casualty file

| Field title | Value | Meaning |
|----------------------------|----------------|---------------------------|
| NAME OF CASUALTY | OWENS CALVIN R | OWENS CALVIN R |
| SERVICE PREFIX AND NUMBER | RA17280815 | RA17280815 |
| GRADE | PFC | Private First Class |
| GRADE CODE | 6 | Private First Class |
| BRANCH | IN | Infantry |
| PLACE OF CASUALTY | L5 | South Korea |
| DAY OF CASUALTY | 31 | 31 |
| MONTH OF CASUALTY | 3 | MARCH |
| YEAR OF CASUALTY | A | 1951 |
| STATE OF RESIDENCE | 73 | Kansas |
| COUNTY OF RESIDENCE | 209 | Wyandotte |
| TYPE OF CASUALTY | RTD | Returned to Duty (FECOM) |
| CASUALTY CODE | 4 | Wounded by missile |
| CASUALTY GROUP CODE | F | Returned to Duty (FECOM) |
| PLACE OF DISPOSITION | | |
| DAY OF DISPOSITION | 25 | 25 |
| BLANK | | |
| MONTH OF DISPOSITION | 5 | May |
| YEAR OF DISPOSITION | A | 1951 |
| YEAR OF BIRTH | | |
| MOS CODE | 4812 | Heavy Weapons Infantryman |
| TROOP SEQUENCE NUMBER | 7003 | IN DIV - 3RD |
| ELEMENT SEQUENCE | 60 | IN DIV INF REGT |
| UNIT | 7 | 7 Infantry |
| RACIAL GROUP CODE | 1 | White |
| COMPONENT | 1 | USA - RA (Reg Army) |
| LINE OF DUTY | | |
| DISPOSITION OF EVACUATIONS | | |

Note: MOS denotes “Military Occupational Specialty”. Searchable records are available at <https://aad.archives.gov/aad/series-description.jsp?s=531>.

Table 4: Summary statistics for TAGOKOR

| | Full Sample | | Volunteers | |
|-----------------------|-----------------------|----------------------|-----------------------|----------------------|
| | White (1) | Black (2) | White (3) | Black (4) |
| A. Before 1-Oct-1951 | | | | |
| Casualty survival (%) | 69.0 | 72.1 | 68.2 | 71.1 |
| North (%) | 51.7 | 35.8 | 51.1 | 37.1 |
| South (%) | 34.2 | 57.6 | 35.5 | 55.9 |
| Private (%) | 57.3 | 64.1 | 54.0 | 60.0 |
| Officer (%) | 5.6 | 1.7 | 6.2 | 1.9 |
| Draftee (%) | 10.6 | 12.9 | — | — |
| Year of birth† | 1927.5 | 1927.3 | 1927.5 | 1927.2 |
| | (<i>n</i> = 56, 223) | (<i>n</i> = 7, 227) | (<i>n</i> = 50, 244) | (<i>n</i> = 6, 295) |
| B. After 1-Oct-1951 | | | | |
| Casualty survival (%) | 79.2 | 79.7 | 79.3 | 80.1 |
| North (%) | 55.2 | 29.6 | 52.1 | 34.0 |
| South (%) | 27.6 | 62.6 | 32.8 | 40.6 |
| Private (%) | 63.2 | 79.0 | 48.9 | 71.2 |
| Officer (%) | 5.7 | 2.1 | 13.6 | 6.0 |
| Draftee (%) | 57.9 | 64.7 | — | — |
| Year of birth† | 1929.5 | 1929.8 | 1929.5 | 1930.5 |
| | (<i>n</i> = 29, 494) | (<i>n</i> = 4, 463) | (<i>n</i> = 12, 430) | (<i>n</i> = 1, 572) |

Note: Formal integration began on 1-Oct-1951. †Year of birth is reported for fatalities only.

Table 5: Typical observation in Social Security Death Index

| Field title | Value |
|--------------|------------|
| ssn | XXXXXXXXXX |
| lastname | OWENS |
| suffix | |
| firstname | CALVIN |
| middlename | R |
| yeardeath | 2005 |
| yearbirth | 1933 |
| stateofdeath | Colorado |
| zipcode | 80222 |
| ssnarea | 512 |
| daydeathmiss | 0 |
| daybirthmiss | 0 |
| dod | 2005-11-07 |
| dob | 1933-01-08 |
| ageatdeath | 72.8795 |
| fipsofdeath | Colorado |
| ssnfips | Kansas |

Note: This table depicts one of over 90 million individual records contained in the Social Security Death Index (SSDI). Although the SSDI reports social security numbers, here I have redacted it for privacy reasons. The last zip code of residence is linked to demographic data to study residential choice. The remaining fields are used to link the SSDI to the casualty file.

Table 6: Typical observation in national cemetery record

| Field title | Veteran | Wife |
|--------------|-----------------------------|-----------------------------|
| d_first_name | Calvin | Virginia |
| d_mid_name | R | C |
| d_last_name | Owens | Owens |
| d_suffix | | |
| d_birth_date | 1/8/1933 | 2/20/1933 |
| d_death_date | 11/7/2005 | 2/10/2015 |
| section_id | 16 | 16 |
| row_num | | |
| site_num | 193 | 193 |
| cem_name | FT. LOGAN NATIONAL CEMETERY | FT. LOGAN NATIONAL CEMETERY |
| cem_addr_one | 4400 WEST KENYON AVENUE | 4400 WEST KENYON AVENUE |
| cem_addr_two | | |
| city | DENVER | DENVER |
| state | CO | CO |
| zip | 80236 | 80236 |
| cem_phone | 303-761-0117 | 303-761-0117 |
| relationship | Veteran (Self) | Wife |
| v_first_name | Calvin | Calvin |
| v_mid_name | R | R |
| v_last_name | Owens | Owens |
| v_suffix | | |
| branch | US ARMY | US ARMY |
| rank | CPL | CPL |
| war | KOREA | KOREA |

Note: Searchable records are available at <https://catalog.data.gov/dataset?publisher=Department+of+Veterans+Affairs&tags=cemeteries>.

Table 7: The effects of integration on survival by race

| Dependent variable: casualty survival dummy | | | | |
|---|---------------------------------|---------------------------------|---------------------------------|----------------------------------|
| | (1) | (2) | (3) | (4) |
| A. Whites | | | | |
| Integration (s_{BW}) | 0.022 ^{***} (0.006) | 0.017 ^{**} (0.008) | 0.022 ^{***} (0.008) | 0.022 ^{***} (0.009) |
| Private | | | | -0.015 ^{***} (0.006) |
| Support role | | | | 0.007 (0.008) |
| Intercept | 0.745 | 0.679 | 0.653 | 0.655 |
| N | 71,184 | 71,184 | 71,184 | 71,184 |
| R^2 | 0.003 | 0.016 | 0.024 | 0.024 |
| B. Blacks | | | | |
| Integration (s_{BW}) | 0.006 (0.019) | -0.029 ^{**} (0.013) | -0.019 [*] (0.011) | -0.018 [*] (0.010) |
| Private | | | | -0.028 ^{**} (0.012) |
| Support role | | | | 0.009 (0.035) |
| Intercept | 0.771 | 0.701 | 0.598 | 0.609 |
| N | 9,702 | 9,702 | 9,702 | 9,702 |
| R^2 | 0.000 | 0.013 | 0.017 | 0.017 |
| 6-month FE | No | Yes | Yes | Yes |
| Regiment FE | No | No | Yes | Yes |
| #Regiments | 24 | 24 | 24 | 24 |
| #Periods | 6 | 6 | 6 | 6 |

Note: Standard errors calculated using two-way clustering by regiment and six-month period of casualty. Period 1 and 2nd Regiment are the omitted groups in the applicable columns. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 8: The effects of integration on survival (pooled sample)

| Dependent variable: casualty survival dummy | | | | | | |
|---|---------------------------------|--------------------------------|---------------------------------|----------------------------------|----------------------------------|----------------------------------|
| | | | | | Segregated | Integrated |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Integration (s_{BW}) | 0.020 ^{***} (0.006) | 0.024 ^{**} (0.010) | 0.019 ^{***} (0.007) | 0.019 ^{***} (0.007) | | |
| Black | 0.023 ^{***} (0.007) | 0.019 ^{**} (0.008) | 0.008 (0.006) | 0.010 [*] (0.006) | 0.055 ^{**} (0.023) | 0.008 (0.012) |
| Private | | | | -0.017 ^{***} (0.005) | -0.020 ^{***} (0.005) | -0.015 ^{***} (0.006) |
| Support role | | | | 0.009 (0.007) | -0.032 (0.023) | 0.045 ^{**} (0.023) |
| N | 84,745 | 84,745 | 84,745 | 84,745 | 23,403 | 21,986 |
| R^2_{Adj} | 0.003 | 0.014 | 0.022 | 0.022 | 0.031 | 0.014 |
| Regiment FE | | X | X | X | X | X |
| 6-month FE | | | X | X | X | X |
| #Regiments | 24 | 24 | 24 | 24 | 14 | 14 |
| #Periods | 6 | 6 | 6 | 6 | 5 | 6 |

Note: Standard errors calculated using two-way clustering by regiment and six-month period of casualty. Columns (1)-(4) uses the the full sample of regiments for which the integration measure could be constructed. In columns (5)-(6), “Segregated” and “Integrated” refer to regiments that in any given period were in the bottom and top integration quartiles, respectively. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 9: Effect of integration on residential sorting (whites)

| Dependent variable: similarity index of last zip code of residence | | | | | |
|--|-------------------------------|-------------------------------|---------------------------------|----------------------------------|---------------------------------|
| | Stayers and movers | | | | Movers |
| | (1) | (2) | (3) | (4) | (5) |
| Integration (s_{BW}) | 0.057 [*] (0.024) | 0.060 [*] (0.030) | 0.036 ^{**} (0.013) | 0.036 ^{**} (0.014) | 0.042 ^{**} (0.015) |
| Age at casualty | -0.001 (0.001) | -0.001 (0.002) | -0.002 [*] (0.002) | -0.002 (0.001) | -0.002 (0.002) |
| Black share by regiment/period | | -0.011 (0.035) | -0.031 (0.025) | -0.030 (0.027) | -0.085 (0.060) |
| Last zip mean house price | | | 0.213 ^{***} (0.013) | 0.211 ^{***} (0.013) | 0.176 ^{***} (0.028) |
| Last zip black share | | | 0.444 ^{***} (0.011) | 0.441 ^{***} (0.010) | 0.401 ^{***} (0.016) |
| Origin county mean wage | | | | -0.049 ^{***} (0.007) | -0.009 (0.012) |
| Origin county high school share | | | | 0.011 (0.152) | -0.219 (0.217) |
| Origin county black share | | | | -0.317 ^{**} (0.110) | 0.013 (0.183) |
| Origin county rural share | | | | -0.261 ^{***} (0.053) | -0.000 (0.073) |
| R^2_{Adj} | 0.036 | 0.036 | 0.303 | 0.305 | 0.227 |
| N | 14,545 | 14,545 | 11,063 | 11,063 | 4,124 |

Note: Panel A. shows OLS results for white veterans. Panel B. show OLS results for black veterans. All regressors except “Age at casualty” are normalized. All regression include period of casualty, regiment and state-of-origin fixed effects. Standard errors two-way clustered by regiment and period of casualty. The dependent variable is a zip code similarity index is constructed using block-level demographic data from the U.S. Census Bureau. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 10: Effect of integration on residential sorting (blacks)

| Dependent variable: similarity index of last zip code of residence | | | | | |
|--|--------------------|-------------------|---------------------|---------------------|-------------------|
| | Stayers and movers | | | | Movers |
| | (1) | (2) | (3) | (4) | (5) |
| Integration (s_{BW}) | 0.028 (0.042) | 0.027 (0.040) | 0.067 (0.063) | 0.055 (0.065) | 0.084 (0.119) |
| Age at casualty | 0.000 (0.004) | -0.000 (0.005) | 0.001 (0.010) | 0.001 (0.012) | -0.009 (0.012) |
| Black share by regiment/period | | 0.008 (0.071) | -0.048 (0.093) | -0.052 (0.101) | 0.060 (0.264) |
| Last zip mean house price | | | 0.104*** (0.030) | 0.102*** (0.026) | -0.062 (0.069) |
| Last zip black share | | | 0.080*** (0.012) | 0.080*** (0.016) | 0.026 (0.031) |
| Origin county mean wage | | | | -0.046 (0.041) | 0.065 (0.043) |
| Origin county high school share | | | | 0.739 (0.413) | 0.576 (1.105) |
| Origin county black share | | | | 0.040 (0.365) | 0.002 (0.333) |
| Origin county rural share | | | | 0.069 (0.282) | 0.624 (0.728) |
| R^2_{Adj} | 0.074 | 0.074 | 0.093 | 0.092 | 0.031 |
| N | 1,536 | 1,536 | 932 | 932 | 416 |

Note: Panel A. shows OLS results for white veterans. Panel B. show OLS results for black veterans. All regressors except “Age at casualty” are normalized. All regression include period of casualty, regiment and state-of-origin fixed effects. Standard errors two-way clustered by regiment and period of casualty. The dependent variable is a zip code similarity index is constructed using block-level demographic data from the U.S. Census Bureau. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 11: Effect of integration on residential sorting (whites)

Dependent variable: similarity index of last zip code of residence

| | Volunteers | | | Draftees | | |
|---------------------------------|--------------------|---------------------|---------------------|-------------------|---------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Integration (s_{BW}) | 0.107** (0.037) | 0.085** (0.026) | 0.082** (0.023) | 0.196 (0.117) | 0.157 (0.115) | 0.158 (0.118) |
| Black share by regiment/period | 0.057 (0.077) | 0.024 (0.041) | 0.025 (0.042) | -0.290 (0.241) | -0.184 (0.229) | -0.190 (0.228) |
| Last zip mean house price | | 0.244*** (0.009) | 0.244*** (0.012) | | 0.193*** (0.038) | 0.196*** (0.040) |
| Last zip black share | | 0.454*** (0.024) | 0.453*** (0.026) | | 0.478*** (0.031) | 0.478 (0.041) |
| Origin county mean wage | | | -0.026 (0.017) | | | 0.009 (0.018) |
| Origin county high school share | | | -0.297** (0.110) | | | -0.224 (0.157) |
| Origin county black share | | | -0.257* (0.102) | | | -0.213 (0.674) |
| Origin county rural share | | | -0.240* (0.094) | | | 0.114 (0.235) |
| R^2_{Adj} | 0.032 | 0.305 | 0.306 | 0.032 | 0.319 | 0.317 |
| N | 5,019 | 3,821 | 3,821 | 1,026 | 787 | 787 |

Note: Panel A. shows OLS results for white veterans. Panel B. show OLS results for black veterans. All regressors except “Age at casualty” are normalized. All regression include period of casualty, regiment and state-of-origin fixed effects. Standard errors two-way clustered by regiment and period of casualty. The dependent variable is a zip code similarity index is constructed using block-level demographic data from the U.S. Census Bureau. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 12: Effect of integration on residential sorting (blacks)

| Dependent variable: similarity index of last zip code of residence | | | | | | |
|--|-------------------|---------------------|---------------------|-------------------|-------------------|-------------------|
| | Volunteers | | | Draftees | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Integration (s_{BW}) | -0.060 (0.125) | 0.127 (0.120) | 0.110 (0.132) | -0.058 (0.732) | 0.249 (0.301) | 0.221 (0.258) |
| Black share by regiment/period | 0.058 (0.113) | -0.088 (0.113) | -0.099 (0.103) | 0.551 (0.234) | 0.628* (0.266) | 0.568 (0.374) |
| Last zip mean house price | | 0.264*** (0.082) | 0.263*** (0.078) | | 0.062 (0.112) | 0.077 (0.159) |
| Last zip black share | | 0.146*** (0.036) | 0.141*** (0.035) | | 0.048 (0.010) | 0.038 (0.128) |
| Origin county mean wage | | | -0.050 (0.069) | | | -0.151 (0.017) |
| Origin county high school share | | | 0.643 (0.669) | | | -2.583 (2.616) |
| Origin county black share | | | 0.265 (0.574) | | | -0.297 (1.919) |
| Origin county rural share | | | 0.112 (0.552) | | | -0.881 (1.448) |
| R^2_{Adj} | 0.080 | 0.129 | 0.116 | 0.163 | 0.079 | 0.039 |
| N | 482 | 280 | 280 | 134 | 77 | 77 |

Note: Panel A. shows OLS results for white veterans. Panel B. show OLS results for black veterans. All regressors except “Age at casualty” are normalized. All regression include period of casualty, regiment and state-of-origin fixed effects. Standard errors two-way clustered by regiment and period of casualty. The dependent variable is a zip code similarity index is constructed using block-level demographic data from the U.S. Census Bureau. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 13: Effect of integration on intermarriage

Dependent variable: wife's non-white/black name index

| | OLS | | | Fractional |
|--------------------------------|--------------------------------|--------------------------------|----------------------------------|---|
| | (1) | (2) | (3) | (4) |
| A. Wife's non-white name index | | | | |
| Integration (s_{BW}) | 0.018 ^{**} (0.005) | 0.020 ^{**} (0.007) | 0.018 [*] (0.008) | 0.335 ^{***} (0.128) [0.031] [*] |
| Black share origin county | 0.009 (0.006) | 0.011 (0.006) | -0.011 ^{***} (0.003) | -0.045 (0.039) |
| Age at casualty | -0.001 (0.012) | -0.001 (0.012) | -0.001 (0.012) | -0.009 (0.022) |
| R^2_{Adj} | 0.003 | 0.000 | 0.006 | 0.048 |
| B. Wife's black name index | | | | |
| Integration (s_{BW}) | 0.022 [*] (0.009) | 0.024 [*] (0.011) | 0.024 ^{**} (0.008) | 0.510 ^{***} (0.169) [0.038] ^{***} |
| Black share origin county | 0.006 (0.008) | 0.007 (0.009) | -0.015 (0.009) | -0.061 (0.046) |
| Age at casualty | 0.005 (0.008) | 0.004 (0.009) | 0.003 (0.008) | 0.004 (0.027) |
| R^2_{Adj} | 0.001 | -0.003 | 0.001 | 0.065 |
| Period FE | X | X | X | X |
| Rank FE | | X | X | X |
| State of res FE | | | X | X |
| N | 664 | 664 | 664 | 664 |

Note: Panel A. regresses the wife's non-white name index on the similarity measure. Panel B. regresses the wife's black name index on the similarity measure. All regressors except "Age at casualty" are normalized. Standard errors two-way clustered by state of origin and period of casualty. Columns (1)-(3) reports OLS estimates. Column (4) estimates the parameters in the fractional response model using pseudo-MLE. The marginal effect is shown in square brackets. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 14: Effect of integration on intermarriage

Dependent variable: wife's non-white/black name index

| | OLS | | | Fractional |
|--------------------------------|----------------------|----------------------|----------------------|------------------------------|
| | (1) | (2) | (3) | (4) |
| A. Wife's non-white name index | | | | |
| Integration (s_{BW}) | 0.002 (0.006) | 0.002 (0.007) | 0.007 (0.008) | -0.058 (0.110) [0.008] |
| Black share origin county | 0.003 (0.005) | 0.003 (0.007) | -0.007 (0.013) | -0.025 (0.040) |
| Age at casualty | -0.007*** (0.002) | -0.007*** (0.002) | -0.007*** (0.002) | -0.029*** (0.007) |
| R^2 | 0.011 | 0.015 | 0.026 | 0.039 |
| B. Wife's black name index | | | | |
| Integration (s_{BW}) | 0.005 (0.004) | 0.004 (0.007) | 0.008 (0.007) | -0.012 (0.131) [0.007] |
| Black share origin county | -0.002 (0.004) | -0.003 (0.005) | -0.019 (0.011) | -0.095** (0.046) |
| Age at casualty | -0.005** (0.002) | -0.005** (0.002) | -0.005** (0.002) | -0.030*** (0.009) |
| R^2 | 0.011 | 0.012 | 0.026 | 0.062 |
| Period FE | X | X | X | X |
| Rank FE | | X | X | X |
| State of res FE | | | X | X |
| N | 901 | 901 | 901 | 901 |

Note: Panel A. regresses the wife's non-white name index on the similarity measure. Panel B. regresses the wife's black name index on the similarity measure. All regressors except "Age at casualty" are normalized. Standard errors two-way clustered by state of origin and period of casualty. Columns (1)-(3) reports OLS estimates and the adjusted R^2 . Column (4) estimates the parameters in the fractional response model using pseudo-MLE and reports the pseudo- R^2 . The marginal effect is shown in square brackets. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

Table 15: Balance test for Similarity Index

| Dependent variable: Similarity Index | | | | |
|--------------------------------------|--------------|--------------------|--------------------|--------------------|
| | q_1 (1) | $q_2 - q_1$ (2) | $q_3 - q_1$ (3) | $q_4 - q_1$ (4) |
| A. First half of war | | | | |
| From north | 0.4559 | 0.0139 (0.009) | 0.0264 (0.008) | 0.0414 (0.009) |
| From south | 0.3105 | 0.0794 (0.008) | 0.0152 (0.008) | -0.0061 (0.008) |
| Age at casualty | 23.281 | -0.274 (0.14) | -0.869 (0.14) | -0.858 (0.15) |
| N | 25,978 | 25,978 | 25,978 | 25,978 |
| B. Second half of war | | | | |
| From north | 0.5313 | -0.0049 (0.011) | -0.0260 (0.011) | -0.0073 (0.012) |
| From south | 0.3011 | 0.0034 (0.010) | 0.0363 (0.010) | 0.0055 (0.011) |
| Age at casualty | 22.457 | 0.080 (0.14) | -0.072 (0.13) | -0.257 (0.14) |
| N | 15,058 | 15,058 | 15,058 | 15,058 |

Note: Panel A. shows the balance test for the first half of the war (25-Jun-1950 to 1951-12-22). Panel B. show the balance test for the second half of the war (1951-12-23 to 27-Jul-1953). Column (1) shows the average of the regiments whose similarity fall within the first quartile (i.e. least integrated). Columns (2)-(4) show the difference in mean of the top three quartiles relative to the bottom quartile. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

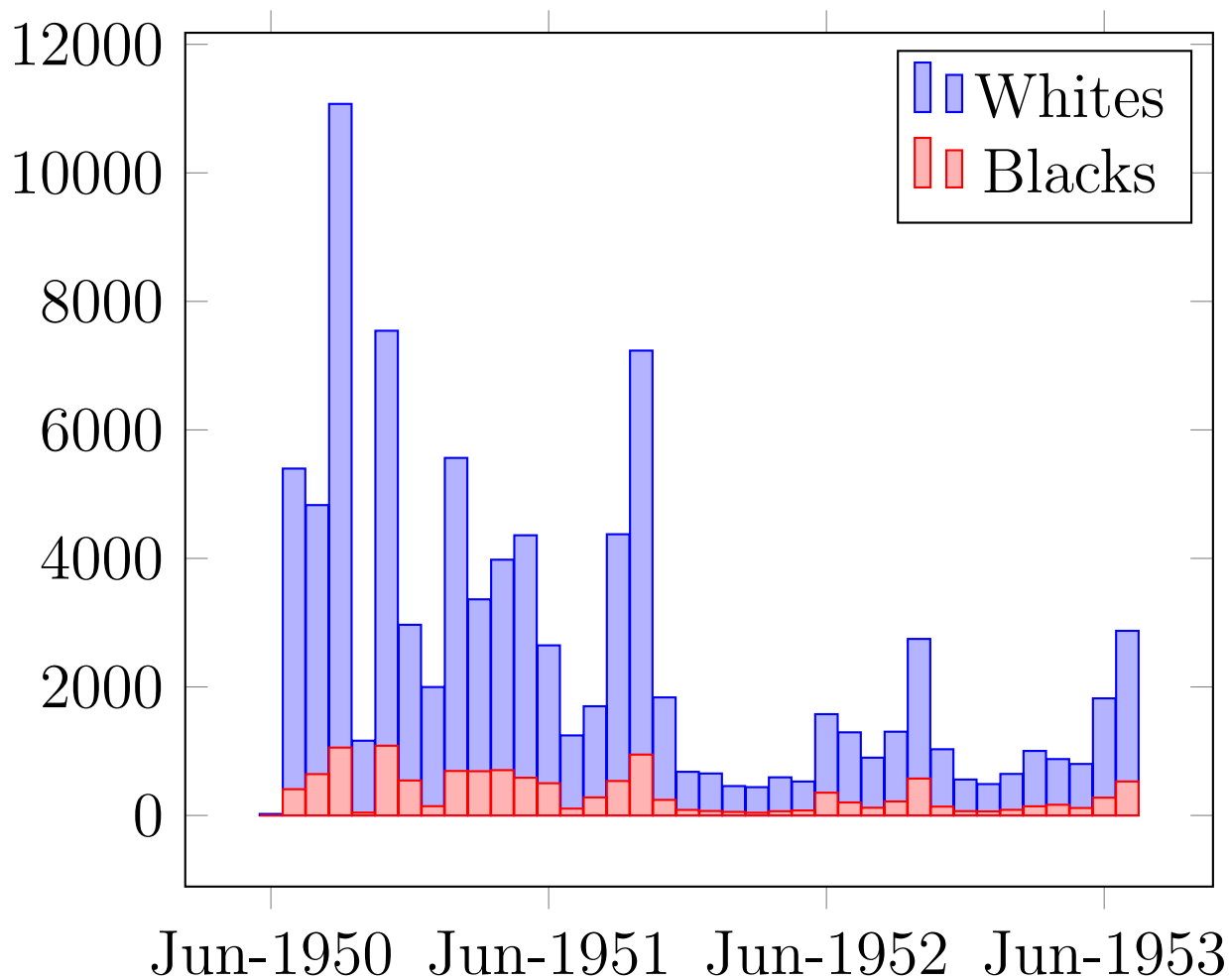
9 Figures

Figure 1: Racial discrimination and black protest



Note: Black custodians protesting discrimination by defense contractors. The “double V” symbol—one “V” formed by the men, the second by the brooms—became the rallying cry for civil rights during the Second World War. The first “V” signifies the struggle for victory over the Axis powers, while the second “V” signifies victory of discrimination at home. RG-208-NP-1KK-1. Office of War Information, National Archives.

Figure 2: Raw number of casualties in each month of the war



Note: Text (forthcoming)

Figure 3: Summary of merged datasets

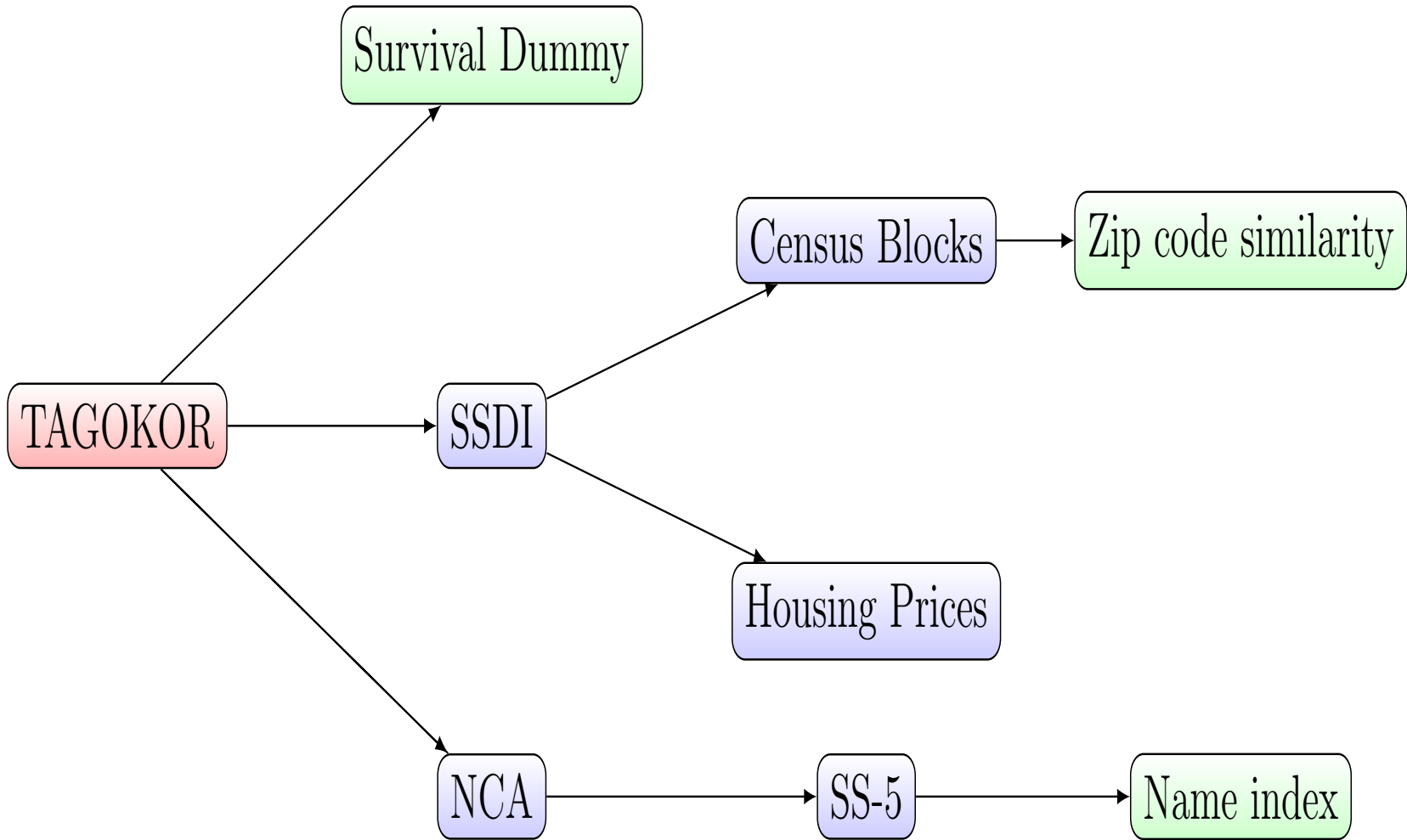
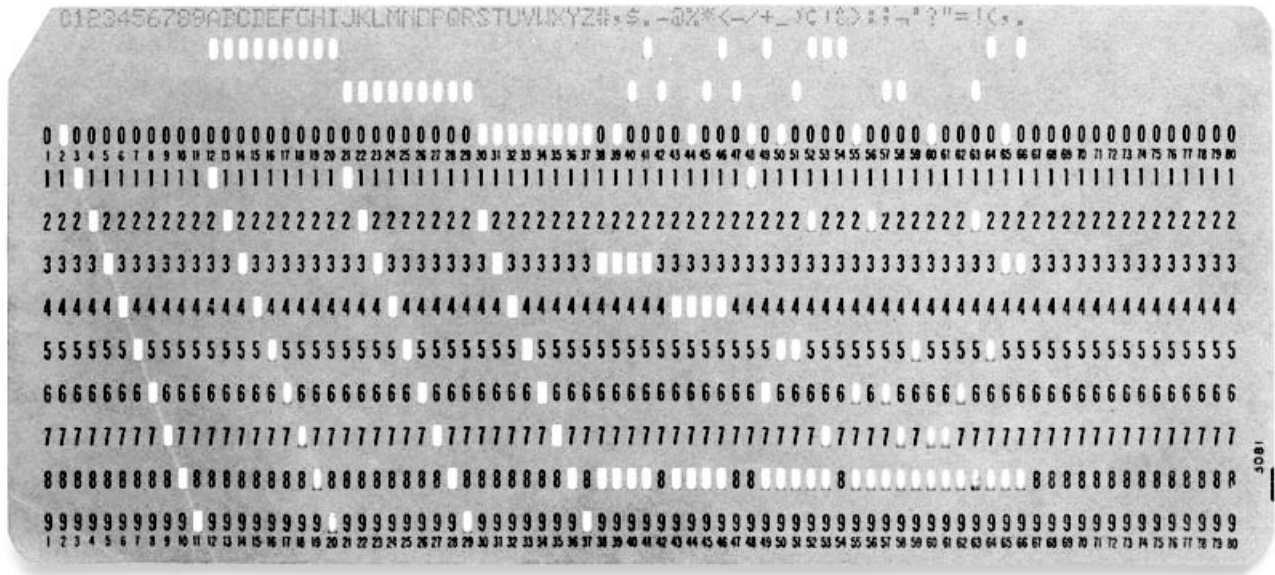


Figure 4: IBM casualty punch card



Note: IBM punch card used to document casualties during the war.

Figure 5: Similarity index for all U.S. zip codes (2010, standardized)

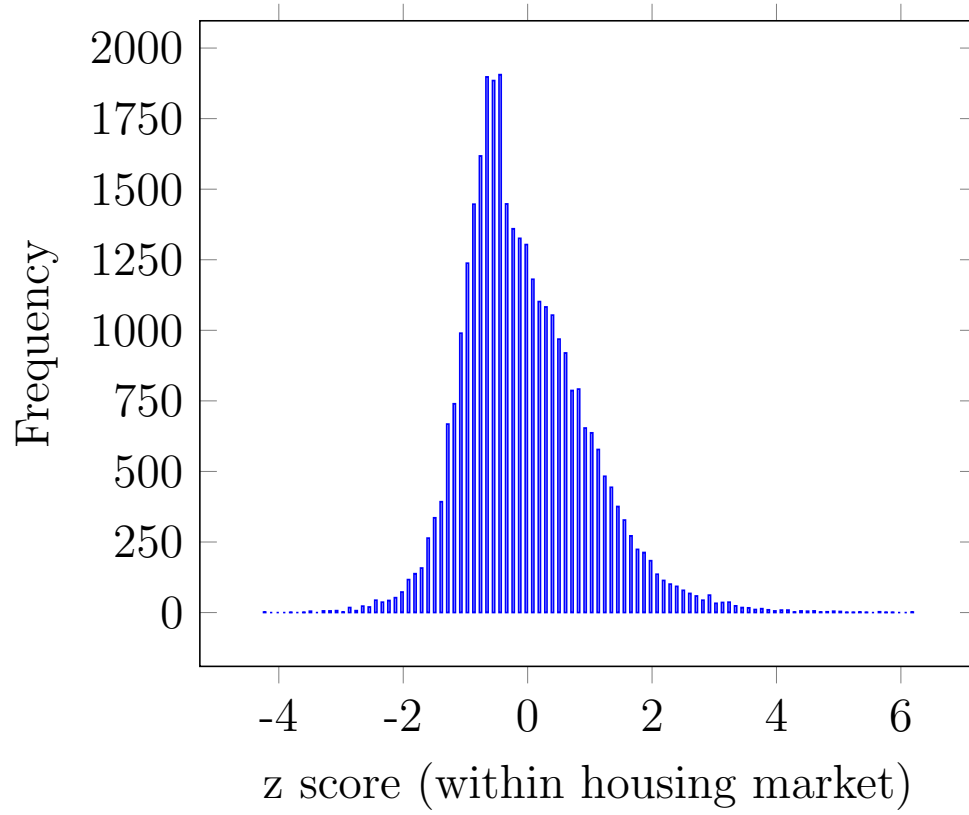
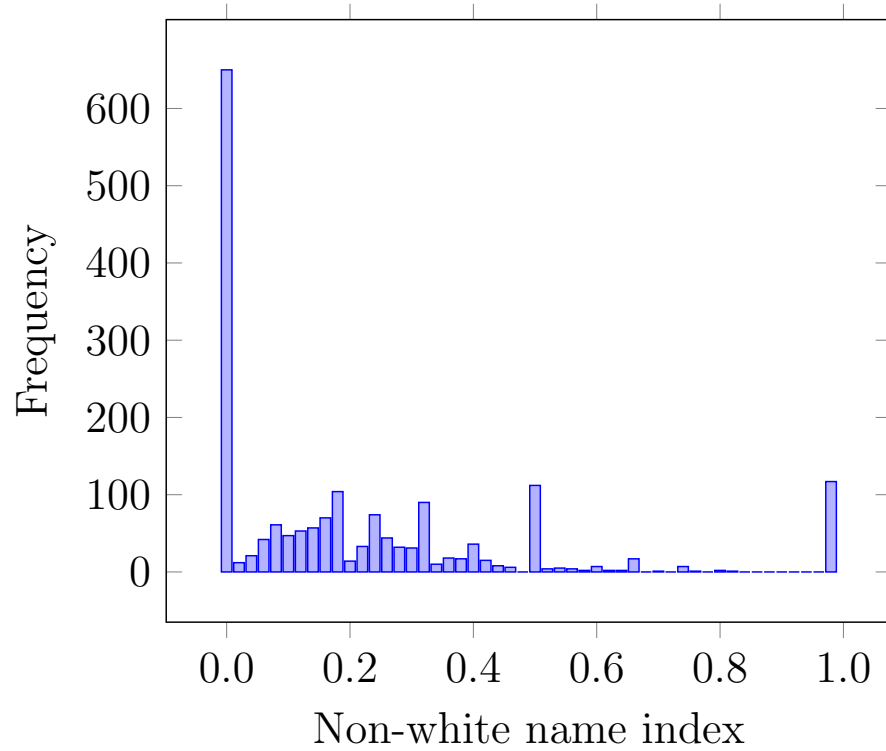
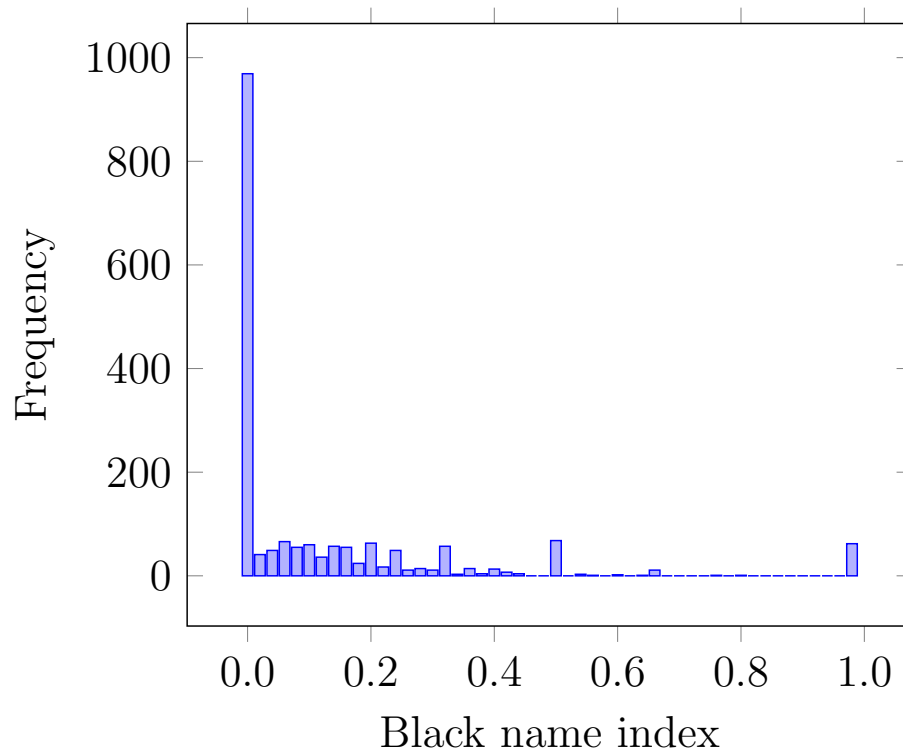


Figure 6: Racial content of wife's name



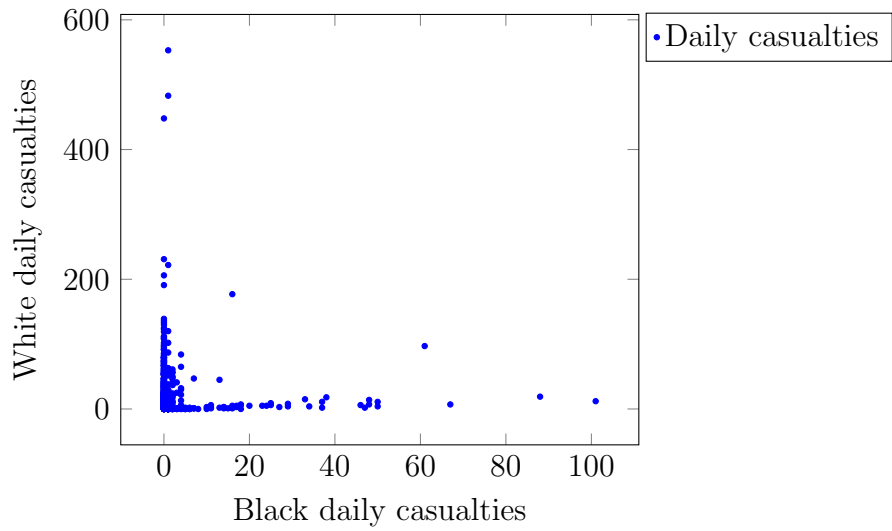
(a) Wife's black name index



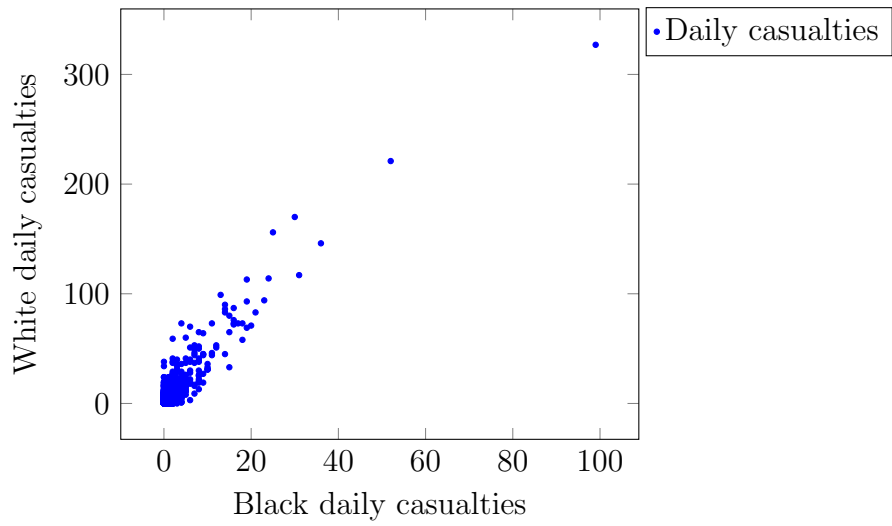
(b) Wife's non-white name index

Note: Name indexes for wives of white veterans who were 22 or younger during Korean War

Figure 7: Black white casualty correlation early and late in the Korean War



(a) First three months of war



(b) Last year of war

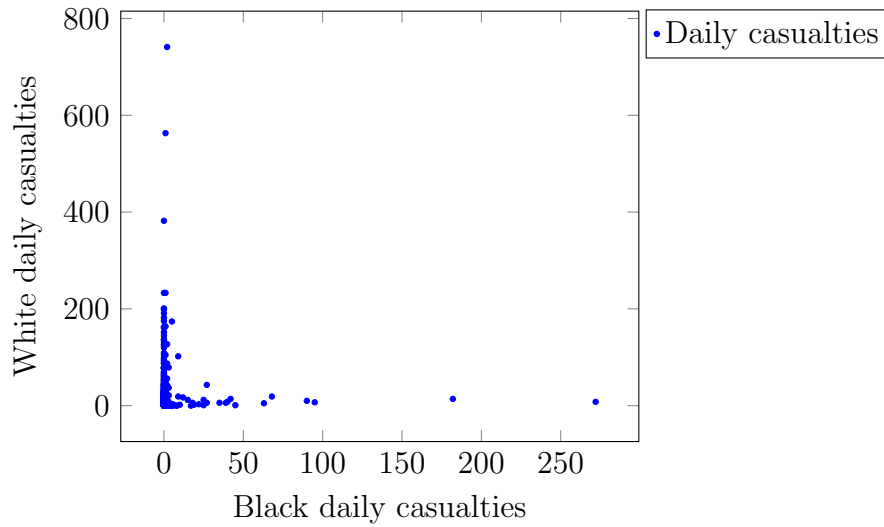
Note: Each dot represents a day in the war.

Figure 8: Example of early integration: 9th Infantry Regiment (November 1950)

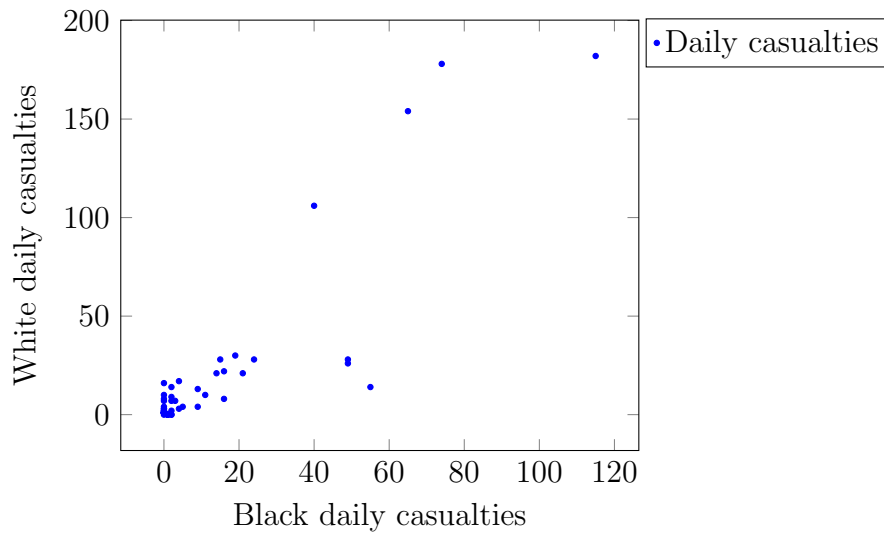


Note: Early in the Korean War, the all-white 1st and 2nd battalions of the 9th Infantry Regiment, like many other all-white combat units, became severely depleted. In desperation, the commanders of the 9th Infantry willingly integrated black replacement troops. This photo, taken sometime in November 1950, shows a recently integrated fire team in a foxhole.

Figure 9: Black white casualty correlation across regiments in the same time period



(a) Months 4-6 excluding 9th Regiment



(b) Months 4-6, 9th Regiment only

Note: Each dot represents a day in the war.

Figure 10: Histogram for regimental similarity index

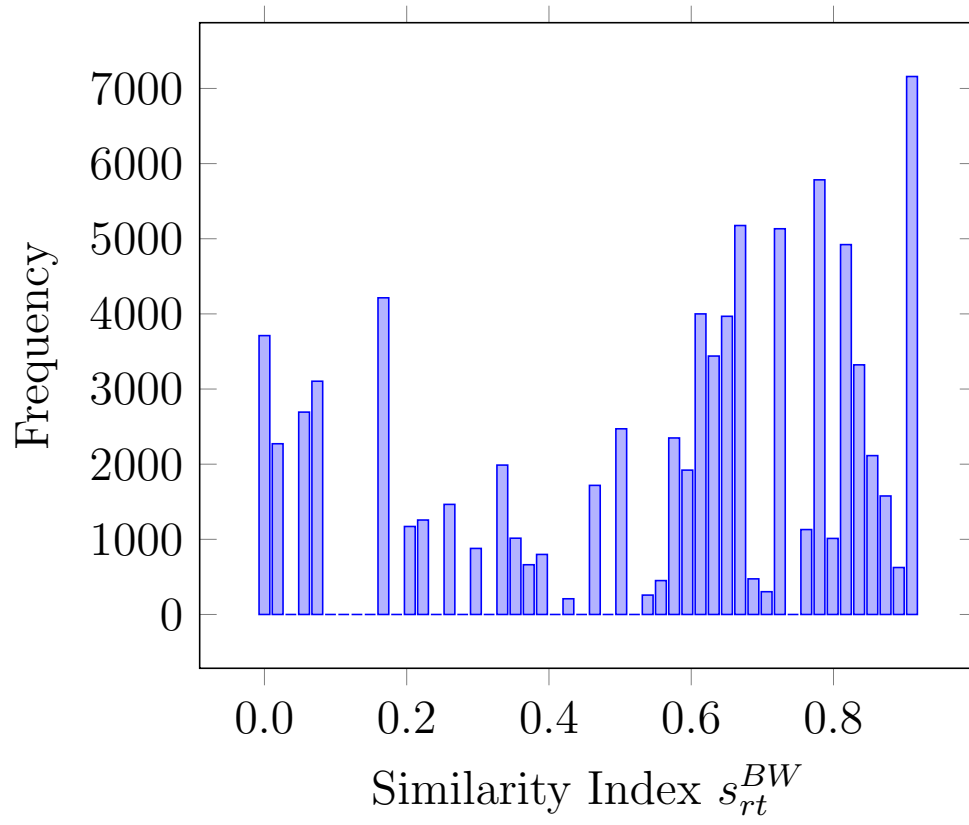
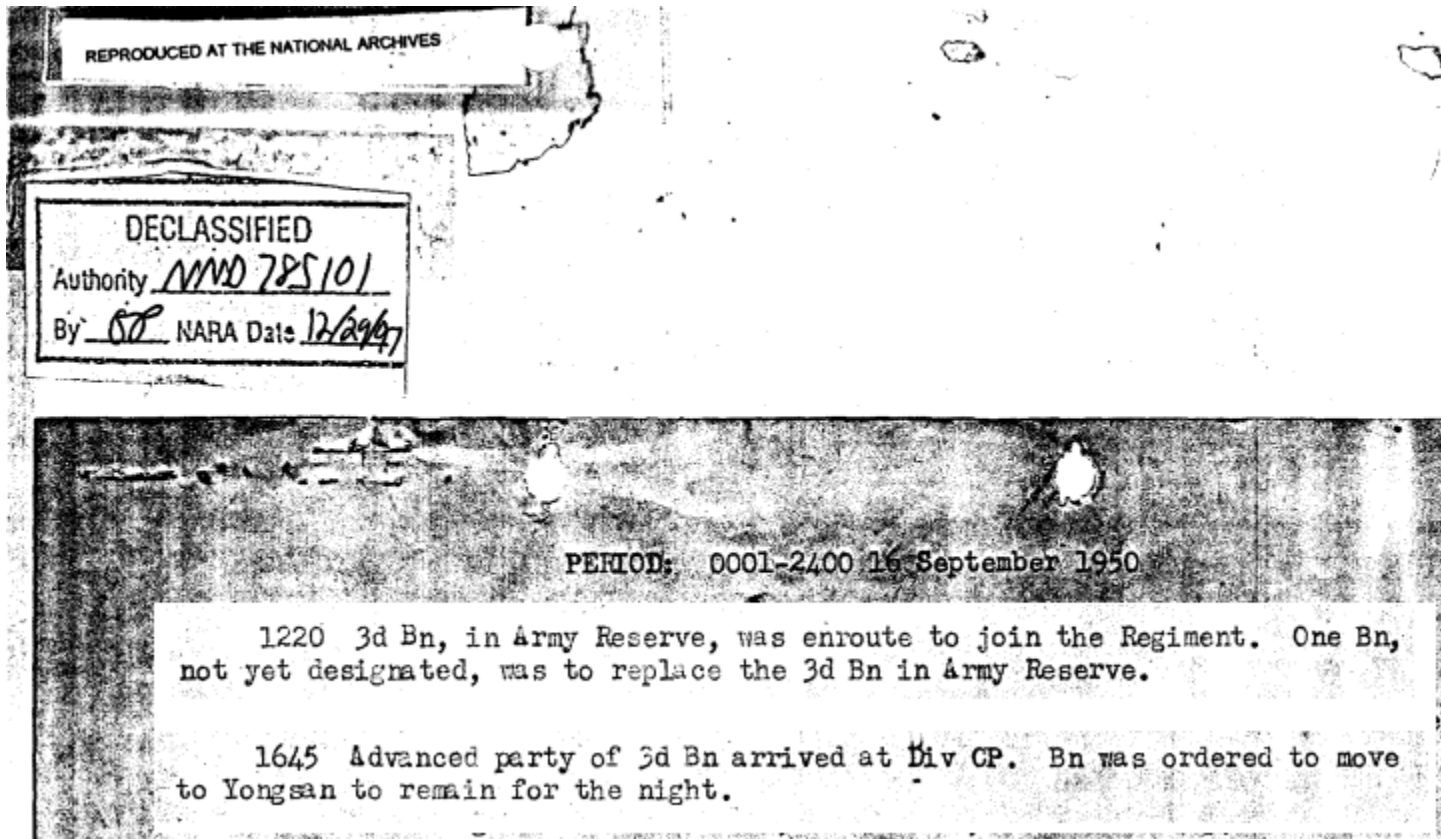
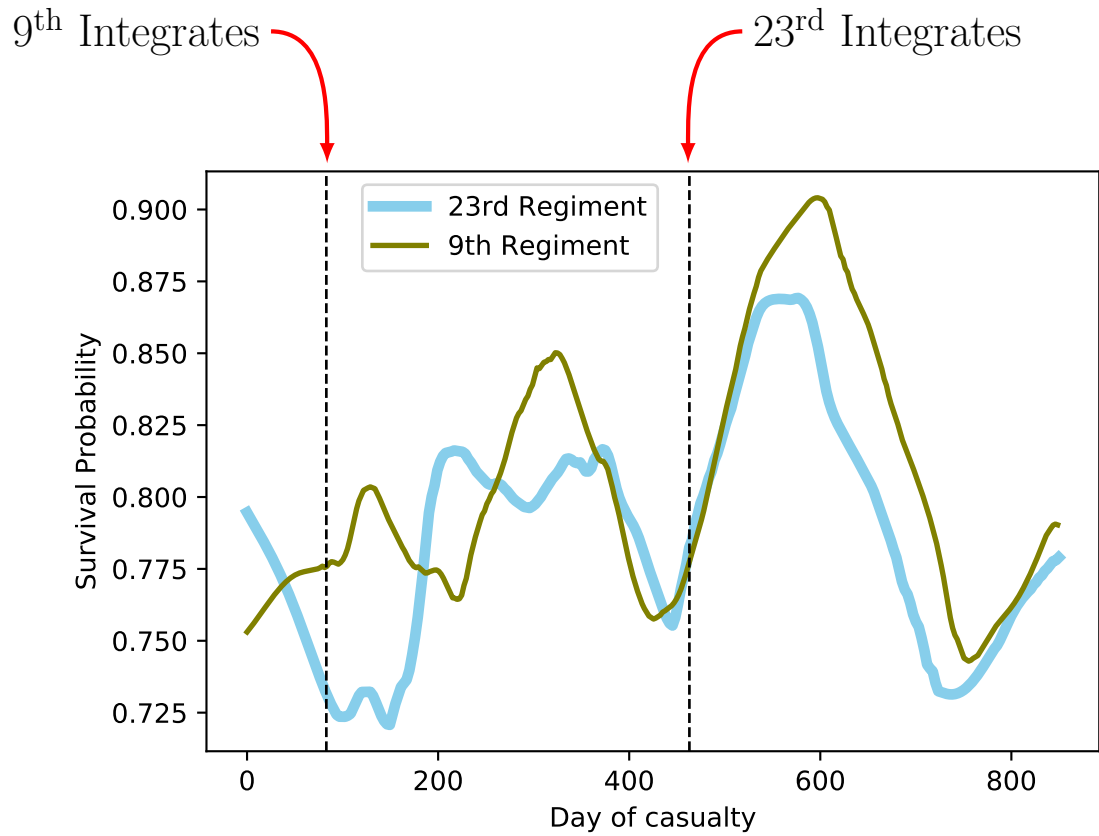


Figure 11: 9th Infantry Regimental War Diary (Sep 1950)



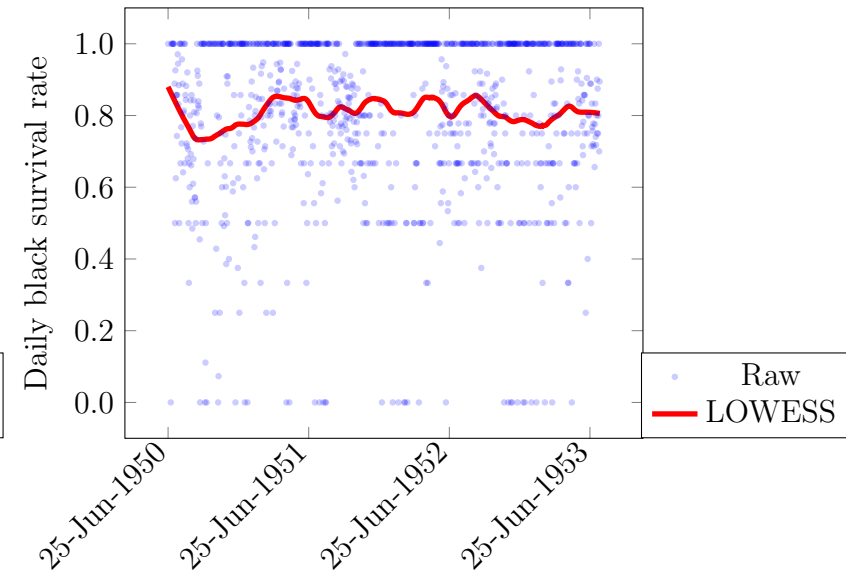
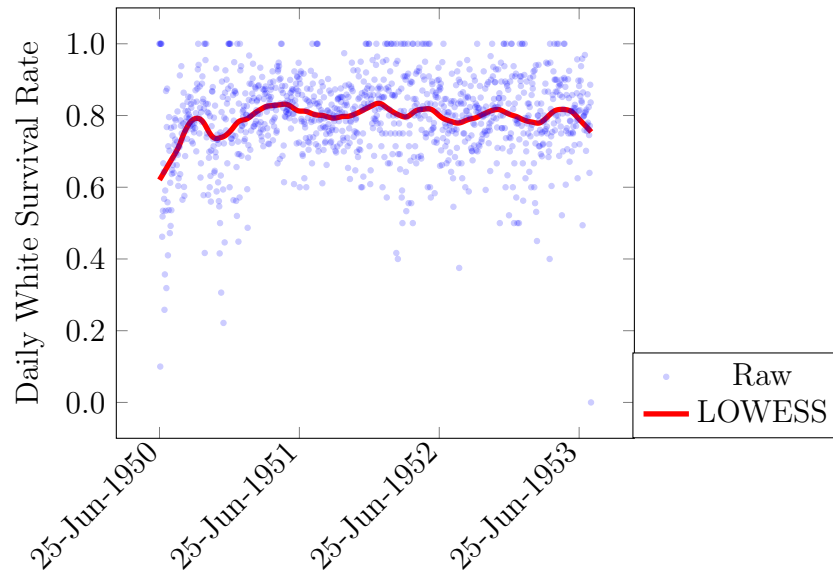
Note: Regimental war diary for the 9th Regiment

Figure 12: Casualty survival rates for 9th and 23rd regiments



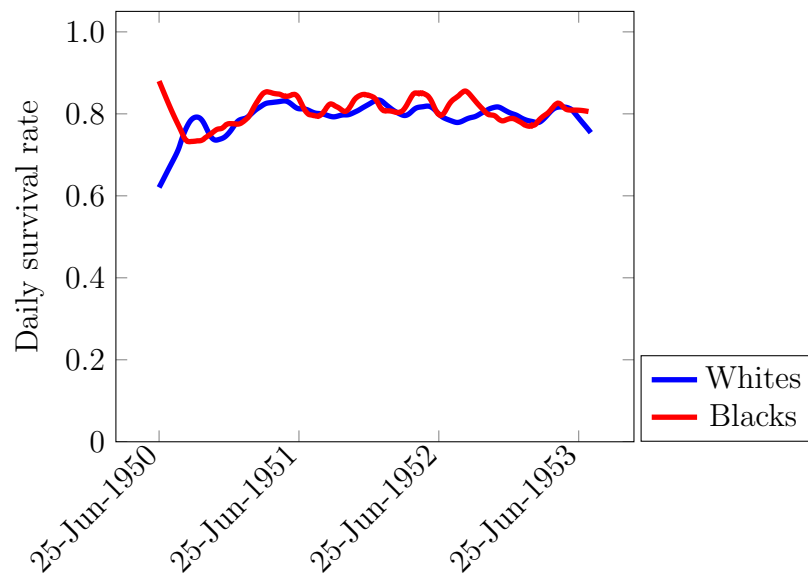
Note: The plot depicts the respective casualty survival rates of the 9th and 23rd Regiments, both of which belonged to the 2nd Infantry Division. The survival rates were estimated from daily casualty data using LOWESS. The first vertical line denotes 16-Sep-1950, and the second vertical line denotes 1-Oct-1951.

Figure 13: Daily survival rates through out the war by race



(a) Note: Each dot represents the average survival rate of wounded white soldiers on a given day in the war.

(b) Note: Each dot represents the average survival rate of wounded black soldiers on a given day in the war.



(c) Note: Figures 13a and 13b combined.