${\bf Advanced~Techniques~in~Regression~and~Unsupervised~Learning~for~Hepatitis} \\ {\bf Mortality~Analysis}$

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1 Introduction

Hepatitis is a viral liver infection that can lead to severe complications such as cirrhosis, liver cancer and even death. Understanding the factors influencing its progression is essential for improving patient outcomes, as early diagnosis and appropriate treatment are crucial to preventing patient mortality. For example, El-Serag (2012) highlighted that chronic hepatitis C can lead to liver cancer (a condition associated with a very high mortality rate), with prognosis depending on disease stage and liver condition [6]. Lok and McMahon (2009) emphasized that age, gender and cirrhosis presence significantly affect hepatitis B progression [7]. Early diagnosis and treatment, as noted by Choi et al. (2018), are essential in preventing complications such as cirrhosis and liver cancer, particularly in patients with coexisting conditions like HIV [8].

The goal of this study is to prevent death from hepatitis, as many patients remain undiagnosed or receive insufficient care. It also aims to assist healthcare professionals in making more informed decisions regarding treatment and surgical options. By analyzing detailed patient data, patterns can be identified that enable more accurate disease predictions, enhancing diagnostics and providing better guidance for medical decision-making, as discussed by Borgia et al. (2017) [9].

The outcome variable in this study is the mortality status of a patient. The input variables consist of various medical parameters and physical symptoms related to liver health. As a result, its primary objective remains the prediction of patient mortality due to liver complications within a high-dimensional feature space, where the interactions between all variables are considered. In order to achieve this task, we will perform different logistic regression techniques and Bayesian model averaging approach to try to both predict and infer a model. To better understand the most relevant factors affecting hepatitis, we performed principal component analysis to capture and summarize the most significant features. Additionally, cluster analysis was conducted to group individuals with similar clinical characteristics and symptoms based on previous obtained results.

2 Related works

The prediction of hepatitis progression has been studied using traditional statistical methods and machine learning techniques. Early studies, like McHutchison et al. (1998), used logistic regression [1], while later research, such as Nguyen (2020) and Majzoobi (2022), applied Support Vector Machines and Random Forests, respectively, to improve prediction accuracy [14], [15]. These hybrid approaches enhanced prediction accuracy but require large datasets and significant computational resources.

In addition, recent advances in deep learning, as seen in the work of Pubmed (2021), offer promising potential for automating the analysis of medical data [?]. However, despite their advantages, these methods have not yet been fully integrated into clinical practice, and manual intervention remains common. Combining traditional methods with machine learning techniques presents a balanced approach, offering both interpretability and improved accuracy.

3 Dataset

The database used for this project, named *hepatitis*, was sourced from *Kaggle*. It includes a sample of 150 individuals, with 20 distinct variables, all related to patients who are currently suffering from or have previously suffered from hepatitis. The recorded data encompasses various details such as the patient's age, gender, symptoms, treatment status and other relevant medical information. The collected variables can be categorized into different groups:

- Patient Characteristics: age, sex
- Symptoms: fatigue, malaise, anorexia
- Liver Complications: liver_big, liver_firm
- Treatment: steroid, antivirals
- Clinical Indicators:
 - Indicators of Portal Hypertension: spleen, spiders, varices, ascites
 - Severity Indicators: bilirubin, alkphosphate, sgot, albumin, protime, histology
- Response Variable: class

3.1 Features

This data will be analyzed to assess its impact on individuals with hepatitis. The collected information is summarized in the following table:

Variable Name	Description	Type	Categories
class	Whether the patient is deceased or alive	Binary	Deceased, Survived
age	Age of the patient	Discrete	-
sex	Gender of the patient	Binary	Male, Female
steroid	Whether the patient is taking corticosteroids	Binary	Yes, No
antivirals	Whether the patient is taking antivirals	Binary	Yes, No
fatigue	Whether the patient has generalized fatigue	Binary	Yes, No
malaise	Whether the patient has general discomfort	Binary	Yes, No
anorexia	Whether the patient has loss of appetite	Binary	Yes, No
liver_big	Whether the patient has liver enlargement (hepatomegaly)	Binary	Yes, No
liver_firm	Whether the patient has liver hardness (cirrhosis)	Binary	Yes, No
spleen	Whether the patient has spleen enlargement (splenomegaly)	Binary	Yes, No
spiders	Whether the patient has capillary dilation (telangiectasia)	Binary	Yes, No
ascites	Whether the patient has free fluid in the abdomen	Binary	Yes, No
varices	Whether the patient has varices in the abdomen	Binary	Yes, No
bilirubin	Bilirubin level	Continuous	-
alkphosphate	Alkaline phosphatase level	Continuous	-
sgot	Hepatic enzyme GOT level	Continuous	-
albumin	Albumin level	Continuous	-
protime	Blood coagulation time	Continuous	-
histology	Whether liver damage is observed at the cellular level	Binary	Yes, No

Table 1: Description of variables in the hepatitis dataset

3.2 Handling Missing Values

Once we have understood our database, we perform a Descriptive Statistics analysis. Based on the results, we have observed a significantly high number of missing values (NA). Specifically, the missing values represent 5.33% of the entire dataset. The variables obtained through blood analyses, in particular, show a notably high percentage of missing values. The case of the variables protime, alkphosphate and albumin are especially significant, with 42%, 18.6% and 10.6% of its data missing respectively. Given these observations, it is crucial to proceed with caution when applying various statistical techniques, as missing values could introduce complications. To address this, we opted to impute them utilizing the mice() function in R. Specifically, for binary variables, we employed logistic regression (logreg), while for continuous variables, we used predictive mean matching (pmm). On the other hand, no outliers were observed in the collected data, and thus, we assume that the data is accurate. For a more detailed understanding of the clinical variables, additional information about hepatitis can be found in Appendix A.

4 Regression

We aim to explore various logistic regression techniques to evaluate both predictive performance and model selection. In his paper, Dr. John McHutchison [1] performed logistic regression analysis to assess the probability of treatment response based on multiple variables, as well as Kaplan-Meier, survival analysis to evaluate the time to events such as cure or complications. Similarly, in our case, we will initially apply different regression techniques to identify the best predictive models. Following this, for a more in-depth analysis, we will employ additional criteria, including Bayesian Model Averaging, to infer the most significant covariates that influence our binary outcome.

4.1 LASSO logistic regression

We begin by considering our 19 covariates, including all possible two-way interactions. This results in a total of 190 parameters, while the dataset contains only 155 patients. Since the number of parameters exceeds the number of observations, the maximum likelihood estimators for the model parameters are likely to exhibit high variance. To address this issue, we apply LASSO regression, which helps reduce the number of covariates by imposing a penalty that encourages sparsity in the model. Let Y be the response vector, X the design matrix and β the vector of coefficients. Then its mathematical formulation is as follows:

$$\hat{\beta} = \arg\min_{\beta} \left(\frac{1}{2n} \|Y - X\beta\|^2 + \lambda \sum_{j=1}^{p} |\beta_j| \right).$$

After implementing LASSO, the number of non-zero parameters was substantially reduced. The optimal regularization parameter ($\lambda = 0.02$) was determined through cross-validation. Furthermore, the model achieved an AIC value of 32.62.

Taking into account these estimates, we proceed to perform Adaptive LASSO. Adaptive LASSO is a regularization technique that enhances variable selection by assigning weights (w_j) to each coefficient. Unlike the standard LASSO, Adaptive LASSO allows different penalties for each coefficient. Its mathematical formulation is as follows:

$$\hat{\beta}^{\text{adaptive}} = \arg\min_{\beta} \left(\frac{1}{2n} \|Y - X\beta\|^2 + \lambda \sum_{j=1}^p w_j |\beta_j| \right).$$

The regularization parameter, $\lambda = 53.062.490$, was selected through cross-validation. For the weights (w_j) , the initial LASSO estimates were used, assigning a minimal value to coefficients that were originally null to impose an infinite penalty. This approach ensures that the number of non-zero parameters is, at most, the same as in the LASSO model. The resulting AIC value is 85.5.

After fitting our models, we evaluate their performance by analyzing the ROC curves, which compare the accuracy of the LASSO and Adaptive LASSO regression models. These curves illustrate the trade-off between sensitivity, the proportion of actual positives correctly identified, and specificity, the proportion of actual negatives accurately classified. Together, sensitivity and specificity provide a balanced assessment of a model's ability to detect true positives while minimizing false positives. The area under the curve (AUC) serves as a key measure of discriminative ability. A higher AUC indicates better overall performance, with an AUC of 1 representing a perfect model and an AUC of 0.5 indicating no discriminative power.

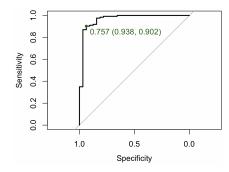


Figure 1: ROC curve for LASSO

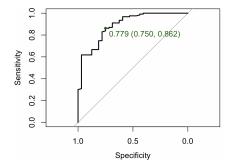


Figure 2: ROC curve for Adaptive LASSO

The LASSO model demonstrates high predictive accuracy, as evidenced by its sensitivity and specificity values. The AUC value of approximately 0.965 indicates near-perfect performance. This is further supported by the model's classification results, with 111 true positives and only 12 false negatives, alongside 30 true negatives and just 2 false positives. These metrics reflect strong performance in correctly predicting both positive and negative instances.

In contrast, the Adaptive LASSO model exhibits lower sensitivity and specificity, resulting in a weaker overall performance. The AUC value of 0.88 underscores this decline compared to the LASSO model. The adaptive model records more misclassifications, with 8 false positives and 17 false negatives, which contribute to its diminished predictive accuracy. Despite this, the Adaptive LASSO still achieves relatively strong performance, albeit not on par with the LASSO model.

	Predicted 0	Predicted 1		Predicted 0	Predicted 1
Fitted 0	30	2	Fitted 0	24	8
Fitted 1	12	111	Fitted 1	17	106

Figure 3: Confusion matrix for LASSO

Figure 4: Confusion matrix for Adaptive LASSO

4.2 Bayesian Model Averaging

In previous approach we performed the selection of our model through cross-validation as the aim was to obtain good predictions of the binary output of the database. Now, we will try to perform model inference to understand which covariates have more effect on the output using both BIC/EBIC criterion and Bayessian Model Averaging (BMA). The latter method will be used to obtain marginal posterior inclusion probabilities and understand which variables are indispensable for the model.

We begin by introducing the BIC and EBIC criteria for model selection. Both of these criteria are consistent for model selection and, unlike the cross-validation approach used previously, are particularly useful when the goal is to understand or infer a model. The BIC and EBIC aim to minimize the following expressions:

$$BIC: -2\log p(y\mid\theta) + \log(n)\|\beta\|_0$$

$$EBIC: -2\log p(y\mid\theta) + \log(n)\|\beta\|_0 + 2\log\left(\frac{d}{\|\beta\|_0}\right)$$

Both criterion have a penalty through the L_0 norm but take into consideration the dimensionality in the penalty. Initially we consider our 19 covariates and fit the model both via BIC criterion and EBIC criterion using the mombf package, BestBIC and BestEBIC functions. The selected significant individual covariates are the following for each criterion:

BIC: protime, albumin, spiders, bilirubin EBIC: protime, ascites, bilirubin

After this first approach, BMA is considered. Bayesian model averaging generates posterior inclusion probabilities for each covariate (and interaction if its the case) and also can generate point estimates. First, to consider a brief interpretation of the model, the individual 19 covariates are selected and using *mombf* package *model_selection* function posterior inclusion probabilities for each variable are computed and coefficient values for the highest probability model are analyzed. This results are shown in Table 4 at Apendix B.

In a second approach contemplating more complex models, we recover first order interactions approach used for predictive analysis, adding up to 19 principal covariates and 171 pairwise interactions. An important thing to consider is hierarchy between interactions: if we include an interaction in the model, both of its principal covariates should be included too. To perform this method we use BAS package for BMA models which allows forcing hierarchy constraints when computing probabilities for all possible models. In order to perform Bayesian model averaging and selection, first we set prior probabilities. For the model space, we define a Beta-Binomial prior with parameters a=1 and b=1, that is a uniform prior over model space. This prior assigns equal prior probabilities to all subsets of predictors. For the coefficients a CCH Mixture prior has been used, a mixture of Zellner's g-priors that introduces flexibility in specifying the prior beliefs about the regression coefficients.

The inclusion probabilities obtained via BMA with hierarchical interactions can be found in Table 5 at Appendix B.

4.2.1 Interpretation

We can interpret the significant variables obtained by each method. While information criterion only selected few variables, BMA provided us inclusion probabilities for every variable and in a more complex model even each interaction. Blood coagulation time named *protime* seems one of the most important variables as it has been selected in all 4 methods proposed. *Bilirubin* and *Albumin* levels also have been selected as an important variable in 4 and 3 methods respectively. An interesting result arouses from the *alk_phosphate* covariate that despite being neglected in 3 previous methods, has high inclusion posterior probabilities due to its interaction with other covariates, the most important being with binnary covariate *spiders*.

Now, we will interpret the following model obtained using BMA approach of the original 19 covariates in terms of odds ratios. The logistic regression model estimates the probability of survival, and the odds ratios (OR) represent the effect of each predictor variable on the odds of survival. The variables with highest marginal probabilities present the following coefficients:

Variable	Coefficient	Odds Ratio (OR)
Intercept	-4.614	$e^{-4.614} \approx 0.009$
protime	0.036	$e^{0.036} \approx 1.036$
albumin	1.240	$e^{1.240} \approx 3.45$
bilirubin	-0.564	$e^{-0.564} \approx 0.569$
sex=male	-0.0012	$e^{-0.0012} \approx 0.988$
spiders=True	-0.654	$e^{-0.654} \approx 0.5199$
anorexia=True	0.699	$e^{0.699} \approx 2.01$
steroid=True	0.454	$e^{0.454} \approx 1.57$

Table 2: Coefficient and Odds Ratio for each variable

An odds ratio greater than 1 indicates a positive relationship, while an odds ratio less than 1 indicates a negative relationship. Odds ratios for each predictor variable are summarized in Table 2.

After analyzing the previous table, the intercept has an odds ratio of $e^{-4.614} \approx 0.009$, meaning that when all predictor variables are set to their reference category, the odds of survival are very low, around 1%. Regarding the sex of the patient, being male slightly reduces the odds of survival by about 2% compared to females. Additionally, using steroids increases the odds of survival by a factor of 1.57 compared to not using steroids. This is very logical since steroids are useful medical treatment for liver diseases.

The presence of anorexia significantly increases the odds of survival, with an odds ratio of $e^{0.699} \approx 2.01$, meaning that having anorexia increases the odds of survival by a factor of 3.08 compared to not having anorexia. This result may seem counterintuitive, as loss of appetite is generally associated with a higher mortality rate. However, anorexia might indicate underlying health issues, which could prompt a patient to seek medical attention sooner. Early diagnosis and intervention could therefore improve the chances of survival. In this context, anorexia is more of an indicator suggesting liver problems, rather than a direct factor causing death. It may not directly determine survival, but it could be a useful signal for reducing the probability of mortality when addressed promptly.

On the other hand, spider-like lesions decrease the odds of survival. The odds ratio for spiders is $e^{-0.654} \approx 0.5199$, meaning that having spider-like lesions reduces the odds of survival to about 49% of the odds for those without spider-like lesions.

Biochemical indicators of severity can assist medical professionals in interpreting the results. The coefficient for bilirubin has an odds ratio of $e^{-0.564}\approx 0.569$, which means that for each unit increase in bilirubin, the odds of survival decrease to approximately 53.1% of the previous odds. In contrast, albumin has a positive effect on survival, with an odds ratio of $e^{1.240}\approx 3.45$. This indicates that for each unit increase in albumin, the odds of survival increase by a factor of 3.45. Lastly, prothrombin time also has a small positive effect on survival, with an odds ratio of $e^{0.036}\approx 1.036$, suggesting that for each unit increase in prothrombin time, the odds of survival increase by about 3.6%.

5 Unsupervised learning

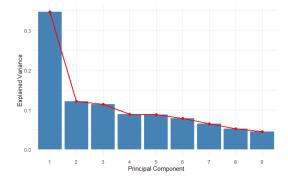
Our objective in this section is to cluster individuals based on similar clinical indicators to assess the relevance of these parameters in determining the mortality status of liver diseases. Given the frequent interconnections among medical variables, we will begin by reducing the number of parameters to pinpoint the most significant ones within each component. Previous studies have applied similar clustering techniques to identify key clinical parameters related to liver disease outcomes. For instance, [10] utilized clustering methods to identify prognostic factors for liver cirrhosis, while [11] explored the clustering of clinical features for predicting liver cancer survival. By reducing dimensionality and focusing on the most significant clinical variables, we aim to improve the interpretability and predictive accuracy of our model.

5.1 Principal Component Analysis

The continuous variables in the dataset reflect various indicators obtained from blood tests. Additionally, some binary variables describe physical symptoms related to liver diseases and hypertension. Our objective is to group the most highly correlated variables under the same principal component, such that each component will correspond to multiple clinical parameters that are interrelated. After calculating the covariance matrix, we find that the most correlated variables are spleen_palpable, spiders, ascites, varices, bilirubin, alk_phosphate, sgot, albumin and protime.

Dimensionality reduction methods include PCA and Probabilistic PCA. A Shapiro-Wilk test revealed that the variables do not follow a normal distribution with a significance level of $\alpha=0.05$, making Probabilistic PCA unsuitable due to its Gaussian assumption. In our case, traditional PCA is more adequate for analyzing the data. Mathematically, PCA relies on the eigenvalue decomposition of the covariance matrix to identify the directions (eigenvectors) that capture the most variability in the data. The data is then projected onto a lower-dimensional subspace defined by the principal components with the largest eigenvalues, preserving as much information as possible.

Before performing Principal Component Analysis, we examined the eigenvalues of the obtained correlation matrix. In order to determine the number of eigenvalues to be selected for our analysis, we use the elbow method. If we choose the number of eigenvalues based on the eigenvalue ratio, only one eigenvalue is selected, which will definitely lead to a difficult interpretation. Taking into account that one of the main objectives of PCA is to reduce dimensionality while keeping the maximum possible variance of data, we select the number of components based on the amount of variance explained by each component.



	Eigenvalue	Variance proportion
PC1	3.13	0.347
PC2	1.1	0.121
PC3	1.026	0.113
PC4	0.79	0.088
PC5	0.78	0.087
PC6	0.71	0.078
PC7	0.58	0.065
PC8	0.47	0.052
PC9	0.4	0.044

Figure 5: Explained variance by each principal component

Figure 6: Eigenvalues and Explained Variance for each Principal Component

After analyzing both previous figure and table, we reach the conclusion that the choice of 3 eigenvalues is the perfect decision leading to three different principal components. With three principal components, we manage to explain more than 57% of the total variance. If we add more eigenvalues the increase in the amount of variance explained does not increase considerably, leading to less than 9% per each principal component.

Once the number of components has been determined, we applied the Principal Component Analysis method to the initial data matrix. The following figure summarizes the results and allows the interpretation of each principal component based on the loadings:

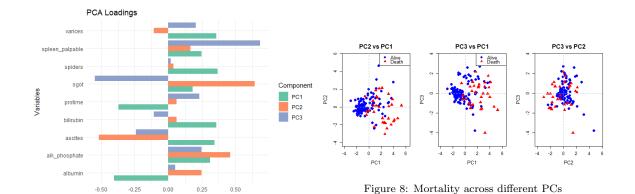


Figure 7: PCA loadings

Interpretation

- PC1: General Liver Dysfunction. PC1 reflects a general pattern of liver dysfunction, with strong correlations to symptoms like *spiders*, *ascites*, *varices*, *bilirubin*, and *alkaline phosphatase*, which are clinical indicators of severe liver disease. Negative correlations with *albumin* and *protime* suggest that worsening liver function leads to more severe symptoms. Higher PC1 values indicate more advanced liver dysfunction.
- PC2: Liver Enzyme Activity. PC2 is primarily associated with liver enzyme activity, showing strong correlations with sgot and alkaline phosphatase, which are markers of liver injury. Negative correlations with ascites suggest that PC2 is more focused on biochemical markers than physical symptoms. Weak correlations with albumin and protime further support this idea.
- PC3: Physical Symptoms of Liver Disease. PC3 is influenced mainly by spleen palpable, indicating it represents physical symptoms like splenomegaly, common in advanced liver disease. Negative correlations with sgot suggest lower enzyme activity in later stages of liver damage. Ascites and varices also contribute to PC3, but less strongly than spleen palpable. This component captures physical symptoms related to severe liver dysfunction.

Finally, we want to study the relationship of principal components with the mortality status of each patient, as it was our target variable in logistic regression. As mentioned before, higher values in PC1 reflects more severe liver disease. In the first plot of figure 4, deceased patients (red triangles) tend to cluster on the right side of the PC1 axis, where values are positive. This suggests a strong association between advanced liver dysfunction and mortality. Along the PC2 axis, which represents liver enzyme activity, there is no clear separation between alive and deceased patients. This indicates that PC2 contributes less to differentiating mortality compared to PC1.

Similar to the first plot, the second plot shows that deceased patients cluster on the right side of the PC1 axis, reinforcing the idea that advanced liver dysfunction is strongly associated with mortality. On the PC3 axis, positive values reflect more severe physical symptoms, such as splenomegaly (spleen palpable). While deceased patients show slightly higher PC3 values, the trend is not as strong as with PC1. Therefore, PC1 remains the most significant differentiator of mortality in this comparison.

The last plot examines the relationship between *liver enzyme activity* (PC2) and *physical symptoms* (PC3). In this case, there is no clear separation between alive and deceased patients, as deceased patients are scattered throughout the plot. Although PC3 contributes slightly to mortality patterns, PC2 and PC3 alone are less effective in distinguishing mortality compared to PC1.

5.2 Cluster Analysis

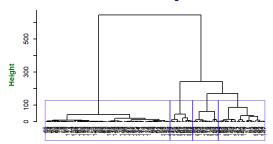
Based on previous results, we found that PCA was a promising method for dimensionality reduction and proceed with cluster analysis, through implementation of both K-means and agglomerative hierarchical methods. As previously mentioned, for the clustering, the values that each individual takes on those three principal components will be considered in order to group individuals that share similar symptoms and clinical indicators.

5.2.1 Hierarchical agglomerative clustering

In this case, the Ward method will be applied to the data to form clusters. The Ward method takes all individuals and calculates the distances between them using squared Euclidean distance. Once this is done, the two individuals with the smallest distance are grouped together. Then, the distances between the remaining n-1 elements (where n-2 of them are individual data points and the other element is a cluster of 2 individuals). This process continues until all individuals are grouped into a single cluster.

Since the applied method is hierarchical, a dendrogram can be constructed. The dendrogram obtained in this case is as follows in Figure 9. Once this is done, it is necessary to determine the number of clusters to work with. Via inspection, 4 clusters are considered, marked in blue in Figure 9. Then the selected clusters are analyzed in Figure 10 and are represented in 3D plot of Figure 15 found at Apendix C.

Hierarchical Clustering: 4 Clusters



Cluster	PC1	PC2	PC3	Size
C1	-1.25	-0.13	-0.11	88
C2	1.49	0.51	1.34	33
C3	0.93	1.23	-1.2	18
C4	2.74	-1.75	-0.82	16

Figure 9: Dendrogram of the hierarchical clustering.

Figure 10: Ward's Method and PCA

In this table, the average values of each factor for the individuals in each cluster are shown, that is, the centroids, as well as the number of individuals in each cluster. Looking at the results, we can say that the value of PC1 in cluster 4 is significant. The centroids of each cluster were analyzed based on the principal components as follows:

- Cluster 1: This group shows low liver dysfunction (PC1 = -1.25), low liver enzyme activity (PC2 = -0.13), and few physical symptoms (PC3 = -0.11), indicating individuals with less severe liver disease.
- Cluster 2: This group reflects advanced liver dysfunction (PC1 = 1.49), higher liver enzyme activity (PC2 = 0.51), and pronounced physical symptoms (PC3 = 1.34), suggesting individuals with more severe liver disease.
- Cluster 3: This group displays moderate liver dysfunction (PC1 = 0.93), higher liver enzyme activity (PC2 = 1.23), and fewer physical symptoms (PC3 = -1.20), indicating individuals with moderate liver disease severity.
- Cluster 4: This group shows severe liver dysfunction (PC1 = 2.74), low liver enzyme activity (PC2 = -1.75), and few physical symptoms (PC3 = -0.82), suggesting individuals with advanced liver dysfunction but fewer physical symptoms.

5.2.2 K-means clustering

In this case, the non-hierarchical K-means method was applied to the data to form clusters. This algorithm iteratively assigns data points to the nearest cluster centroid, then updates the centroids based on the mean of the points in each cluster. This process continues until the centroids no longer change with the aims to minimize the within-cluster variance, resulting in compact and well-separated clusters. In order to decide the number of clusters, we used elbow method and the gap statistic. Both visual inspection and gap statistic gave rise to similar conclusion: the best number of clusters is four. In fact, the maximum gap is observed at k=4 with a value of 0.727.

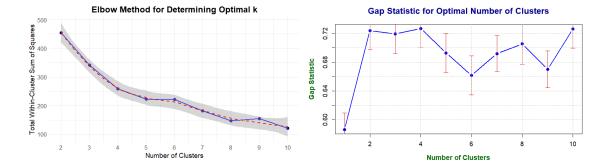


Figure 11: Elbow method

Figure 12: Gap Statistic for different K.

Looking at the following table, similar outcomes were obtained to those when 4 clusters were specified in the previous section. Therefore, we can conclude that the created clusters are stable. In this case, the second cluster (C2) is very significant to the first principal component. The interpretation is really similar to the previous part. In this case, Cluster 1 and Cluster 3 represent moderate liver dysfunction, while Cluster 2 reflects more severe liver dysfunction with biochemical markers. Cluster 4 represents the least severe cases, with a large group size. Visual representation of the clusters found via K-means can be found in Figure 16 at Apendix C.

Cluster	PC1	PC2	PC3	Size
1	0.602	0.469	1.619	21
2	2.942	-1.117	-0.12	24
3	1.16	1.411	-0.898	22
4	-1.236	-0.159	-0.129	88

Table 3: Principal Component Values, Clusters, and Group Sizes

Finally, we analyze the distribution of mortality status across the different clusters. A clear bijective correspondence can be established between the clusters in both the hierarchical and K-means models. Therefore, we will focus on interpreting the results from one of these methods, specifically the survival status in the Ward's model.

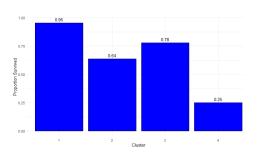


Figure 13: Survival status in Ward's method

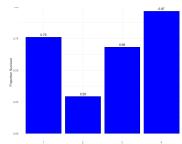


Figure 14: Survival status in K-Means

Cluster 1 is associated with less severe liver disease and, as a result, exhibits a lower mortality rate among its members. Clusters 2 and 3 are linked to moderate liver disease, with Cluster 2 patients experiencing more pronounced physical symptoms. Consequently, the survival rate is relatively high in these clusters, particularly in Cluster 3, where physical symptoms are less common. However, the survival rate in both clusters is still lower than that of Cluster 1. Cluster 4, on the other hand, shows the highest mortality rate, reflecting the advanced liver dysfunction and significantly reduced enzyme activity in these patients.

6 Conclusion

This project allowed us to explore various model selection techniques, including LASSO and Bayesian model averaging, in the context of a prediction problem. For regression tasks, we employed different approaches based on the specific objectives. LASSO and Adaptive LASSO were utilized for predictive analysis, with the optimal regularization parameter selected through cross-validation. These methods were particularly suited for improving predictive accuracy by shrinking coefficients and selecting relevant features. In contrast, Bayesian model averaging was employed to enhance interpretability by accounting for model uncertainty and combining insights from multiple models.

The results indicate that LASSO outperformed Adaptive LASSO in terms of predictive accuracy. This suggests that LASSO's regularization approach may better balance complexity and accuracy in this specific context, while Adaptive LASSO, despite its potential for refinement, showed relatively weaker performance in prediction tasks. Consistent model selection methods like BIC, EBIC and Bayessian Model Averaging let us identify the most significant variables for the survival output as well as consider some important interactions. Posterior inclusion probabilities and coefficient posterior mean estimates give us an idea of how changes on most important covariates impact the survival probability.

Given more time and better computational resources, additional methodologies could be explored to enhance the effectiveness of the models. In the context of logistic regression, incorporating higher-order interaction terms could be an avenue for improvement. Thus far, the analysis has been limited to second-order interactions, leaving the potential impact of more complex relationships between variables unexplored. Additionally, other regularization techniques, such as Elastic Net, could be explored to strike a balance between the L1 and L2 penalties, potentially improving performance when dealing with correlated features.

Regardind unsupervised learning, due to the non-normality of the variables, we applied Principal Component Analysis (PCA) for dimensionality reduction, which facilitated the interpretation of clinical patterns. We then used the reduced-dimensional data to identify clusters using both K-means and hierarchical clustering models. Both algorithms yielded similar results and interpretations. With more time, it would have been valuable to explore transforming the variables to achieve normality, which could have allowed us to apply Probabilistic PCA. Additionally, for a more probabilistic approach to clustering, we could have implemented Gaussian Mixture Models.

A Hepatitis overview

Hepatitis is an infectious, immunological or toxic disease that causes inflammation of the liver. The general symptoms of most types of hepatitis include weakness, generalized fatigue, loss of appetite, fever, chills, nausea, jaundice and darkened urine. Hepatitis caused by viruses is an endemic disease and can affect people of all ages. There are different types of viral hepatitis: A, B, C, D, E, F and G [12].

The diagnosis of this disease is performed through blood tests and/or liver biopsy. Blood tests measure the levels of various molecules produced in the liver, as well as other parameters related to liver function. These include bilirubin and albumin levels, hepatic enzyme counts such as GOT, alkaline phosphatase and blood coagulation time.

The bilirubin level is directly proportional to liver damage, while low albumin levels indicate liver failure. On the other hand, although alkaline phosphatase and GOT enzyme levels alone are not highly specific, deviations in both are indicative of liver damage. Finally, since most coagulation factors are synthesized in the liver, measuring blood coagulation time helps detect potential liver problems. All these parameters are also used to assess the severity of the disease. A liver biopsy is performed to detect liver damage at the cellular level.

In the initial phase of the disease, the liver enlarges (hepatomegaly). As the disease progresses, the liver begins to harden, leading to cirrhosis. Although these two conditions rarely occur simultaneously, they can appear in some patients. When cirrhosis develops in the liver, complications such as portal hypertension arise.

The portal vein is a blood vessel that carries blood from the intestines to the liver for filtration, after which the blood is transported to the heart. When the liver becomes fibrotic, there is increased pressure in the blood vessels of the liver, which impairs blood transport. Consequently, blood accumulates in blood vessels located upstream of the liver, such as the spleen, causing it to enlarge (splenomegaly). Thus, we can conclude that splenomegaly is a clinical sign of portal hypertension in the context of hepatitis. In addition to splenomegaly, other clinical manifestations of portal hypertension include capillary dilation (telangiectasia), the presence of free fluid in the abdominal cavity (ascites) and abdominal varices, among others.

Regarding treatment, corticosteroids, which are anti-inflammatory drugs, are often administered to counteract the liver's inflammatory response. Since the disease can sometimes be caused by a virus, antivirals may also be used. However, antivirals are generally ineffective for most types of viral hepatitis and are usually only administered in cases of hepatitis B and C. The best results are observed in hepatitis C, where many patients achieve recovery.

B BMA Inclusion Probabilities and Posterior Mean

Predictor	Estimate	2.5% CI	97.5% CI	MargPP
(Intercept)	-4.614	-8.897	0.938	0.993
protime	0.036	0.000	0.062	0.947
albumin	1.240	0.000	2.178	0.899
bilirubin	-0.564	-1.040	0.000	0.871
sexmale	-0.012	-1.523	1.531	0.861
spidersTrue	-0.654	-1.605	0.256	0.722
anorexiaTrue	0.699	-0.211	1.712	0.690
steroidTrue	0.454	-0.519	1.427	0.630
ascitesTrue	-0.385	-1.491	0.809	0.617
liver_firmTrue	0.450	0.000	1.468	0.551
malaiseTrue	-0.063	-1.032	0.911	0.360
varicesTrue	-0.040	-1.031	0.883	0.326
$spleen_palpableTrue$	-0.078	-1.096	0.779	0.299
fatigueTrue	0.127	-0.649	1.174	0.290
antiviralsTrue	0.055	-0.720	1.060	0.249
age	-0.007	-0.066	0.000	0.237
histologyTrue	0.013	-0.575	0.705	0.171
alk_phosphate	0.000	-0.001	0.006	0.111
liver_bigTrue	0.030	0.000	0.668	0.109
sgot	-0.000	0.000	0.000	0.042

Table 4: Coefficient Estimate and margin probabilities of each covariate

Predictor	Inclusion Probability
Intercept	1.000
protime	0.992
sexmale	0.978
spidersTrue	0.969
alk_phosphate	0.967
albumin	0.958
age	0.926
spidersTrue : alk_phosphate	0.925
liver_firmTrue	0.901
bilirubin	0.848
anorexiaTrue	0.759
liver_firmTrue : alk_phosphate	0.548
alk_phosphate : protime	0.347
fatigueTrue	0.302
malaiseTrue	0.219
ascitesTrue	0.146
varicesTrue	0.172
liver_bigTrue	0.114
$spleen_palpableTrue$	0.084
steroidTrue	0.084
histologyTrue	0.077
sgot	0.074
antiviralsTrue	0.080

Table 5: Inclusion Probabilities of each principal covariate and some significant interactions

C Clustering figures

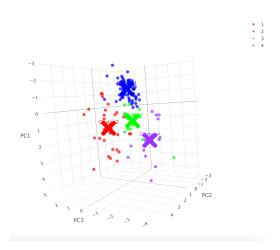


Figure 15: Representation through Ward's method of 4 clusters, with centroids marked with a cross.

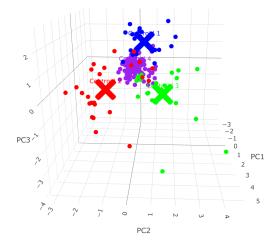


Figure 16: Representation through K-means algorithm of 4 clusters, with centroids marked with a cross.

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E Code

```
#PART1: LOGISTIC REGRESSION, NON LINEAR MODELS AND LASSO
set.seed(142)

#Read database

db<-read.csv(file="hepatitis.csv",header = T, sep = ",", na.strings = "NA", dec = ".")

#Columns with no information will be given NA

db[db == ""] <- NA

#Edit class variable
db$class[db$class=="live"] <- "1"
db$class[db$class=="die"] <- "0"
db$class[db$class=="die"] <- "0"
db$class[db$class=="die"] <- "0"
db$class(- as.numeric(db$class)

#Identify columns of class character
char_cols <- sapply(db, is.character)

#Convert them to factor
db[, char_cols] <- lapply(db[, char_cols], as.factor)
summary(db)

#Handle NAN
#Install mice package
#install.packages("mice")
library(mice)
```

```
methods <- make.method(db)</pre>
         #Impute missing values with NAN choose m=3 since our database is not very
imputed_data <- mice(db, m = 3, method = methods, maxit = 50, seed = 123)
summary(imputed_data)</pre>
 27
 29
 31
         imputed_data$method
db <- complete(imputed_data)
summary(db)</pre>
 33
 35
        db$class[db$class=="live"]<- "1"
db$class[db$class=="die"]<- "0"
db$class<- as.numeric(db$class)
plot(db$class,ylab = "")
 37
 39
 40
         boxplot(db$age)
boxplot(db$bilirubin)
boxplot(db$alk_phosphate)
boxplot(db$sgot)
 42
 43
         boxplot(db$albumin)
boxplot(db$protime)
 46
 48
 50
51
         #install.packages("glmnet")
library(glmnet)
 52
         # Use colnam colnames(db)
 56
        #Load neccesary libraries library(glmnet)
 58
 59
 60
         # Prepare data
db$class <- as.factor(db$class)</pre>
 62
 64
         #Function to compute AIC, BIC and EBIC
calculate_criteria <- function(y, y_hat, coef_matrix, X, gamma = 1) {</pre>
 66
          carculate_criteria '= function(y, y_nat, coef_matri)
n <- length(y)
p <- ncol(X)
rss <- sum((y - y_hat)^2)
df <- sum(coef_matrix != 0)
log_lik <- -n/2 * (log(2 * pi) + log(rss/n) + 1)</pre>
 67
 68
 70
 71
72
           #information criterion
aic <- -2 * log_lik + 2 * df
bic <- -2 * log_lik + log(n) * df
ebic <- bic + 2 * gamma * log(p) * df
 75
76
            return(list(AIC = aic, BIC = bic, EBIC = ebic))
 79
 80
 81
         modelo1 <- glm(class ~ age + sex + steroid + antivirals + fatigue + malaise + anorexia + liver_big + liver_firm + spleen_palpable + spiders + ascites + varices + bilirubin + alk_phosphate + sgot + albumin + protime + histology, family = binomial(link = "logit"), data = db)
 83
84
 85
 87
          summary(modelo1)
 89
         #take into account second order ineractions
model_formula <- as.formula(paste(
    "class ~ ("</pre>
 91
           paste(names(db)[-length(names(db))], collapse = " + "),
 93
         ))
 95
         #ordinary logistic regression with interactions(more parameters than data)
modelo2 <- glm(model_formula, family = binomial(link = "logit"), data = db)
 96
97
          summary (modelo2)
 99
100
        set.seed(142)
X <- model.matrix(model_formula,data=db)[,-1]</pre>
104
         y <- db$class
105
        #Fit Lasso ith cross validation
```

```
cv_lasso <- cv.glmnet(X, y, family = "binomial", alpha = 1)
109
          cv_lasso$lambda.min
111
          coef_lasso<-coef(cv_lasso, s = "lambda.min")[,1]</pre>
         coef_df <- as.data.frame(as.matrix(coef_lasso))
coef_df$Variable <- rownames(coef_df)</pre>
116
         #Filter non cero coeficients
coef_df_non_zero <- coef_df[coef_df$'V1' != 0, ]</pre>
119
120
121
         #Compute AIC,BIC and EBIC for this model
y<-as.numeric(db$class)-1</pre>
         fitted_probs <- predict(cv_lasso, newx = X, s = cv_lasso$lambda.min, type = "response")
fitted_probs <- as.numeric(as.character(fitted_probs))
coef_matrix <- as.matrix(coef(cv_lasso, s = cv_lasso$lambda.min))
matrix <- as.matrix(db[, -i])  # Asumiendo que la primera columna es la variable respue:
criteria<-calculate_criteria(y,fitted_probs,coef_matrix,matrix)</pre>
124
125
126
128
         print(criteria)
129
130
         set.seed(142)
coef_initial <- as.vector(coef(cv_lasso, s = "lambda.min"))[-1]
weights <- 1 / abs(coef_initial)</pre>
132
134
          weights[is.infinite(weights)] <- 1e10 #How to handle null coefficients</pre>
136
         #Fix adaptive Lasso
adaptive_lasso <- glmnet(X, y, alpha = 1, family = "binomial", penalty.factor = weights)</pre>
138
140
         cv_adaptive_lasso <- cv.glmnet(X, y, alpha = 1, family = "binomial", penalty.factor = weights)
142
         best_lambda <- cv_adaptive_lasso$lambda.min
coef_adaptative_ccoef(cv_adaptive_lasso, s = best_lambda)
coef_adaptative_df<- as.data.frame(as.matrix(coef_adaptative))
coef_adaptative_df$Variable <- rownames(coef_adaptative_df)
coef_adaptative_df_non_zero <- coef_adaptative_df[coef_adaptative_df$'s1' != 0, ]</pre>
144
146
148
        #Compute AIC,BIC and EBIC for this model
y<-as.numeric(db$class)-1
fitted_probs_adapt <- predict(cv_adaptive_lasso, newx = X, s = cv_adaptive_lasso$lambda.min, type = "response")
fitted_probs_adapt <- as.numeric(as.character(fitted_probs_adapt))
coef_matrix_adapt <- as.matrix(coef(cv_adaptive_lasso, s = cv_adaptive_lasso$lambda.min))
matrix_adapt <- as.matrix(db[, -1])  # Asuniendo que la primera columna es la variable respuesta
criteria_adapt<-calculate_criteria(y,fitted_probs_adapt,coef_matrix_adapt,matrix_adapt)
print(criteria_adapt)</pre>
150
154
155
156
157
158
159
160
161
162
163
         library(pROC)
roc_el<-roc(db$class,fitted_probs, quiet=TRUE)</pre>
165
         roc_el$auc
167
         plot(roc_el,print.thres=TRUE,print.thres.col="dark green")
169
         db$probs<-ifelse (fitted_probs < 0.757, 0, 1)
.Tablew1 <- xtabs(~db$class+db$probs)</pre>
170
          prop.table(.Tablew1,1)
          roc_el2<-roc(db$class,fitted_probs_adapt, quiet=TRUE)</pre>
         roc_e12 auc
plot(roc_e12, print.thres=TRUE, print.thres.col="dark green")
         db$probs_adapt<-ifelse (fitted_probs_adapt < 0.665, 0, 1)
.Tablew1 <- xtabs("db$class+db$probs_adapt)
179
          .Tablew1
          prop.table(.Tablew1,1)
181
182
183
          colnames (db)
         db<-db[,-c(21,22,23)]
185
186
187
        #Model Inference
```

```
#Install mombf packag for model selection and sparseMatrixStats as it is a necessary dependance install.packages("sparseMatrixStats") install.packages("mombf")
193
      #load the packages into environment
library("sparseMatrixStats")
library("mombf")
197
      199
200
201
202
203
204
      fitbic
205
      #display the best model obtained
summary(fitbic)
206
207
208
      209
210
211
                                   sgot + albumin + protime + histology),data=db, family='binomial')
212
214
      fitebic
      #display the best model obtained
summary(fitebic)
216
217
218
220
      fit1 <- modelSelection((class ~ age + sex + steroid + antivirals + fatigue + malaise + anorexia + liver_big + liver_firm + spleen_palpable + spiders + ascites + varices + bilirubin + alk_phosphate
222
224
                                        sgot + albumin + protime + histology),data=db, family='binomial')
      coefbma <- coef(fit1)</pre>
225
      # Convert matrix to a data frame for
coefbma_df <- as.data.frame(coefbma)</pre>
227
      # Sort by the 'margpp' column (descending order)
229
230
      coefbma_sorted <- coefbma_df[order(coefbma_df$margpp, decreasing = TRUE), ]</pre>
231
232
      print(coefbma_sorted)
233
235
236
237
      library(BAS)
238
239
240
241
242
      **Identify numerical and categorical columns
numerical_cols <- sapply(db, is.numeric)
categorical_cols <- !numerical_cols
243
244
246
      preProc <- preProcess(db[, numerical_cols], method = c("center", "scale"))
db_normalized <- predict(preProc, db[, numerical_cols])</pre>
248
249
      #combine back the db
db_combined <- bind_cols(db[, categorical_cols], as.data.frame(db_normalized))</pre>
252
      #Run imposing force.heredity TRUE for hierarchical interactions
#The run took several hours
bas_model<-bas.glm(</pre>
254
255
256
         model_formula,
family = binomial(link = "logit"),
258
259
         data=db_combined,
260
         method =
         force.heredity = TRUE)
261
262
      # We compute inclusion probabilities
inclusion_probabilities <- bas_model$probne0.MCMC
names <- bas_model$namesx</pre>
263
264
265
266
      #store it in a df for
results <- data.frame(</pre>
267
268
269
     Variable = names,
```

```
Prob = as.numeric(inclusion_probabilities)
270
        print(results)
272
               can check now the highest posterior inclusion probability covariates/interactions
274
        results <- results[order(-results$Prob),]
print(results)
275
276
        #we can also check de highest probability model HPM
HPM <- predict(bas_model, estimator = "HPM")
variable.names(HPM)</pre>
278
279
280
281
282
283
284
285
286
        coef_output <- coef(bas_model)</pre>
        # Extract posterior means and inclusion probabilities
posterior_means <- as.numeric(coef_output$postmean) # Ensure it's numeric
inclusion_probs <- as.numeric(coef_output$probne0) # Ensure it's numeric</pre>
287
288
289
290
       # tompine into a data frame
coef_table <- data.frame(
  Predictor = coef_output$namesx, # Predictor names
  PosteriorMean = posterior_means, # Numeric posterior means
  InclusionProb = inclusion_probs # Numeric inclusion probabilities
)</pre>
291
292
293
295
297
       298
299
301
        # We reorder by inclusion probability
coef_table <- coef_table[order(-coef_table$InclusionProb), ]</pre>
303
304
305
        # Print the results
print(coef_table)
```

Listing 1: Lasso and model inference Code

```
3
                       #Model dataset definition with a decrease with the matter of the matter 
                       #Columns with no is
db[db == ""] <- NA</pre>
                       #Edit class variable
db$class[db$class=="live"]<- "1"
db$class[db$class=="die"]<- "0"</pre>
                       db$class<- as.numeric(db$class)
                       #Identify columns of class character char_cols <- sapply(db, is.character)
 13
14
15
                       #Convert them to factor
db[, char_cols] <- lapply(db[, char_cols], as.factor)</pre>
16
17
                         summary(db)
 19
20
                      #Install mice package
#install.packages("mice")
library(mice)
21
23
                        methods <- make.method(db)</pre>
                       #Impute missing values with NAN choose m=3 since our database is not very big imputed_data <- mice(db, m = 3, method = methods, maxit = 50, seed = 123) summary(imputed_data)
 29
 30
                        imputed_data$method
31
                       db <- complete(imputed_data)
summary(db)</pre>
36
                      library(dplyr)
 38
                   #columns to transform
```

```
41
43
     #Get numerical features in order to get dummies db <- db \% > \%
45
       47
      summary (db)
49
     #Not take gender into account
db_numeric <- db[,c(1,3,4,5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20)]
db_numeric[] <- lapply(db_numeric, function(col) {
    if (is.factor(col)) {
50
51
54
55
          return(as.numeric(as.character(col)))
     return(col)
))
#Turn into numeric to compute covariance matrix
56
57
     library(ggplot2)
library(reshape2)
59
60
      # Calculate the correlation matrix
62
63
      a <- cor(db_numeric)
64
66
      a_long <- melt(a)
68
     69
70
       theme(axis.text.x = element_text(angle = 45, hjust = 1)) +
ggtitle("Covariance Matrix")
 74
76
78
     filtered_cov - ifelse(abs(a) > 0.3, a, NA)
print(filtered_cov)
80
      library(ggplot2)
82
      library(reshape2)
84
      filtered_cov_long <- melt(filtered_cov, na.rm = TRUE)</pre>
86
      #Create heatmap
ggplot(data = filtered_cov_long, aes(x = Var1, y = Var2, fill = value)) +
88
       geom_tile(color = "white") +
scale_fill_gradient2(low = "blue", high = "red", mid = "white", midpoint = 0,
limit = c(-1, 1), space = "Lab") +
89
90
91
       theme_minimal() +
theme(axis.text.x = element_text(angle = 45, hjust = 1))
92
93
95
96
     mat<- db_numeric[,c(9,10,11,12,13,14,15,16,17,19)]
y<-mat[,c(10)]</pre>
97
      mat1 <-mat[,c(1,2,3,4,5,6,7,8,9)]
99
      mat1<-scale(mat1)</pre>
101
      mat1 <- as.data.frame(mat1)
sapply(mat, class)</pre>
      library(nortest)
107
      shapiro.test(mat1$bilirubin)
      shapiro.test(mat1$alk_phosphate)
shapiro.test(mat1$sgot)
shapiro.test(mat1$albumin)
109
111
      shapiro.test(mat1$albumin)
shapiro.test(mat1$protime)
shapiro.test(mat1$spleen_palpable)
shapiro.test(mat1$spiders)
shapiro.test(mat1$ascites)
113
      shapiro.test(mat1$varices)
119
      cor1 <- cor(mat1)
121
```

```
eigen(cor1)
eig <- eigen
123
                 eig <- eigen(cor1)
eigenvalues <- eig$values
125
127
                 ggplot(data = data.frame(x = 1:length(eigenvalues), y = eigenvalues), aes(x = x, y = y)) +
geom_point(color = "blue", size = 3) +
geom_line(color = "blue", linewidth = 1) +
labs(
129
                         labs(
  title = "Scree Plot",
  x = "Number of Components",
  y = "Eigenvalue"
132
134
136
                         theme_minimal()
                                                                kes place when we have only one eigenvalue. As said before, it is senseless so we take three.
137
 138
                  library(psych)
                 #PCA with the number of components specification pca_result <- prcomp(mat1, scale. = TRUE)
140
                  #Plot the variance explained by each component
explained_variance <- summary(pca_result)$importance[2, ]</pre>
141
142
143
                  explained_variance
144
145
                 variance_df <- data.frame(
Component = 1:length(explained_variance),</pre>
146
                    Variance_Explained = explained_variance
148
150
                ggplot(variance_df, aes(x = as.factor(Component), y = Variance_Explained)) +
geom_bar(stat = "identity", fill = "steelblue") +
geom_lone(aes(group = 1, y = Variance_Explained), color = "red", linewidth = 1) +
geom_point(color = "red", size = 2) +
theme_minimal() +
151
 153
154
                   theme_minimal() +
labs(title = "Proportion of Explained Variance",
    x = "Principal Component",
    y = "Explained Variance") +
scale_x_discrete(labels = as.character(variance_df$Component))
156
 157
158
 159
160
162
                loadings <-pca_result$rotation loadings[,1:3]
164
165
                  pca_rotated_scores <- pca_result$x[, 1:3]</pre>
166
167
                #Save loadings in a data frame
loadings_df <- as.data.frame(loadings[,1:3])
loadings_df$Variable <- rownames(loadings_df)</pre>
168
169
170
173
174
                  library(tidyr)
                  loadings_long <- gather(loadings_df, key = "Component", value = "Loading", -Variable)
                  #Ballion | description | 
176
                     geom_bar(stat =
coord_flip() +
178
                       theme_minimal() +
labs(title = "PCA Loadings",
    x = "Variables",
    y = "Loadings") +
179
181
                     scale_fill_brewer(palette = "Set2")
183
185
                #Save each component for cluster a
mat1$PC1 <- pca_rotated_scores[,1]
mat1$PC2 <- pca_rotated_scores[,2]
mat1$PC3 <- pca_rotated_scores[,3]
mat1$class<-y</pre>
187
189
                 #plot the results
library(car)
191
                 # Set up the plotting area for three plots in one row par(mfrow = c(1, 3), mar = c(5, 5, 3, 1))
193
195
                # Plot 1: PC2 vs PC1
plot(PC2 PC1, data = subset(mat1, class == "1"),
    col = "blue", pch = 16, cex = 1.5,
    xlab = "PC1", ylab = "PC2",
    main = "PC2 vs PC1",
    xlim = c(min(mat1$PC1) - 1, max(mat1$PC1) + 1),
    ylim = c(min(mat1$PC2) - 1, max(mat1$PC2) + 1),
    cex.lab = 1.2, cex.main = 1.5, cex.axis = 1.1)
197
198
199
200
201
202
203
204
```

```
207
209
210
       abline(h = 0, v = 0, col = "gray", lty = 2, lwd = 1)
211
       212
213
214
215
216
217
       218
219
220
221
              main = "PC3 vs PC1",
xlim = c(min(mat1$PC1) - 1, max(mat1$PC1) + 1),
ylim = c(min(mat1$PC3) - 1, max(mat1$PC3) + 1),
cex.lab = 1.2, cex.main = 1.5, cex.axis = 1.1)
222
223
224
225
226
       227
228
230
       # Add reference lines at 0
abline(h = 0, v = 0, col = "gray", lty = 2, lwd = 1)
232
233
234
       legend("topright", legend = c("Alive", "Death"),
    pch = c(16, 17), col = c("blue", "red"),
    cex = 1.2, bg = "white", box.lwd = 1.5)
236
238
       # Plot 3: PC3 vs PC2
plot(PC3 ~ PC2, data = subset(mat1, class == "1"),
    col = "blue", pch = 16, cex = 1.5, #
    xlab = "PC2", ylab = "PC3",
    main = "PC3 vs PC2",
240
242
244
              main = res vs rez,
xlim = c(min(mat1$PC2) - 1, max(mat1$PC2) + 1),
ylim = c(min(mat1$PC3) - 1, max(mat1$PC3) + 1),
cex.lab = 1.2, cex.main = 1.5, cex.axis = 1.1)
245
246
247
248
       249
250
251
252
253 \\ 254
       # Add reference lines at 0
abline(h = 0, v = 0, col = "gray", lty = 2, lwd = 1)
255
       256
257
258
259
260
       # Reset the plot lay
par(mfrow = c(1, 1))
261
262
263
265
266
       mat2<-mat1
       datuak <- mat2[,-c(1:9)]  # Remove unnecessary columns from mat2 datuak2 <- mat2[,-c(1:9)]  #Create a copy for second analysis  # Get the distance matrix with Euclidean distance (squared)
267
       # Get the distance matrix with Euclidean distance distantrix <- dist(datuak, method = "euclidean") distmatrix2 <- distmatrix^2 distmatrix2 # Display the squared distance matrix #set seed for reproducibility set.seed(123)
269
271
273
275
       # Use Ward method for hierarchical clustering based on the factors F1 and F2
HClust.1 <- hclust(dist(model.matrix(~-1 + PC1 + PC2+ PC3, datuak))^2, method = "ward.D")</pre>
277
279
       plot(HClust.1, main = "Cluster Dendrogram", xlab = "Individual number in the dataset", ylab = "", sub = "")
281
       # Set 4 clusters from the dendrogram
# Improved hierarchical clustering plot
283
284
285
      par(mar = c(5, 4, 4, 2) + 0.1)
```

```
# Create the basic dendrogram plot
plot(HClust.1,
    main = "Hierarchical Clustering: 4 Clusters", # Title
    xlab ="", # X-axis label
    ylab = "Height", # Subtitl
    col.main = "darkblue", # Title c
    col.lab = "darkgreen", # Axis la
    cex.main = 1.5, # Title c
    cex.lab = 1.2, # Axis st
    font.lab = 2, # Bold ax
    font.main = 2.
289
291
294
295
297
298
299
                       font.lab = 2,
font.main = 2,
lwd = 2,
hang = -1,
col.axis = "black")
300
301
302
303
304
            # Add colored rectangle around the 4 clusters
rect.hclust(HClust.1, k = 4, border = "blue")
305
306
307
            # Reset plotting parameters (optional)
par(mfrow = c(1, 1))
308
309
310
312
            # Print the number of individuals in each cluster summary(as.factor(cutree(HClust.1, k =4 )))
313
314
315
            by(model.matrix(~-1 + PC1 + PC2+ PC3, datuak), as.factor(cutree(HClust.1, k = 4)), colMeans)
316
317
318
320
            321
322
324
            # Ensure centroids is a list of numeric vectors centroids <- lapply(centroids, unlist) # Convert each element to a vector if needed
325
326
328
329
            centroids_df <- do.call(rbind, centroids)</pre>
330
            # Convert the result to a data frame if necessary
centroids_df <- as.data.frame(centroids_df)</pre>
331
332
334
            colnames(centroids_df) <- c("PC1", "PC2", "PC3")</pre>
336
            # Add cluster identifiers (assuming there are 4 clusters)
centroids_df$cluster <- factor(1:4)  # Add cluster identifiers</pre>
337
338
339
340
341
            library(plotly)
342
            # Get the cluster assignments
clusters <- cutree(HClust.1, k = 4)</pre>
343
345
             # Create the 3D scatter plot for the data points
346
           # Create the 3D scatter plot for the data points
fig <- plot_ly(
    x = mat1$PC1,  # First principal component (PC1)
    y = mat1$PC2,  # Second principal component (PC2)
    z = mat1$PC3,  # Third principal component (PC3)
    color = factor(clusters),  # Cluster colors
    colors = c("blue", "red1", "green1", "purple1"),  # Custom colors for clusters
    type = "scatter3d",  # 3D scatter plot
    mode = "markers",  # Display points as markers
    marker = list(size = 5, opacity = 0.7)  # Point size and opacity for data point</pre>
347
348
349
350
351
353
               mode = "markers", # Display points as markers
marker = list(size = 5, opacity = 0.7) # Point size and opacity for data points
354
355
356
357
           # Add centroids to the plot as a separate trace with larger and distinct markers
fig <- fig %>% add_trace(
    x = centroids_df$PC1,  # Centroid positions for PC1
    y = centroids_df$PC2,  # Centroid positions for PC2
    z = centroids_df$PC3,  # Centroid positions for PC3
    color = factor(centroids_df$C1,  # Color the centroids by cluster
    colors = c("blue4", "red4", "green4", "purple4"),  # Matching centroid colors
    type = "scatter3d",
    mode = "markers+text",  # Display markers and text labels
    text = paste("Cluster", centroids_df$cluster),  # Cluster label
    marker = list(
358
359
361
362
363
364
365
366
367
368
                 marker = list(
```

```
size = 10,  # Larger marker size for centroids
symbol = "x",  # Use "x" symbol for centroids
line = list(width = 3),  # Thicker border for centroid markers
opacity = 2  # Full opacity for centroids
369
370
371
373
374
          showlegend = FALSE # Hide the legend for centroids
375
376
377
       378
379
380
381
382
383
      (2,)
384
385
386
387
388
       fig
389
390
391
       # Add the cluster assignment to the 'datuak'
datuak$hclus.label <- cutree(HClust.1, k = 4)</pre>
392
393
394
395
       head(datuak)
396
397
398
400
402
403
404
       # Create the feature matrix
data_matrix <- model.matrix(~-1 + PC1 + PC2 +PC3, datuak2)</pre>
406
407
       # Test k-means for different numbers of clusters \mathtt{set.seed} (123) # Ensures reproducibility
408
409
410
411
       total_withinss <- numeric()</pre>
412
413
       for (k in 2:10) {
   kmeans_result <- kmeans(data_matrix, centers = k, iter.max = 10) # Ejecuta k-means para cada k
   total_withinss[k] <- kmeans_result$tot.withinss # Guarda el within-cluster sum of squares
414
415
416
417
418
       elbow_data <- data.frame(
   k = 2:10,</pre>
419
420
421
          total_withinss = total_withinss[2:10]
422
423
       # Create the elbow plot with ggplot2
ggplot(elbow_data, aes(x = k, y = total_withinss)) +
  geom_point(color = "darkblue", size = 3) + # Points for each k
geom_line(color = "blue", linewidth = 1) + # Line connecting the points
geom_smooth(method = "loess", color = "red", linetype = "dashed", size = 1) + # Smoothed line
424
425
426
427
428
429
          labs(
           title = "Elbow Method for Determining Optimal k",
430
          x = "Number of Clusters",
y = "Total Within-Cluster Sum of Squares"
) +
431
432
433
          theme_minimal(base_size = 14) +
435
436
437
439
440
          scale_x_continuous(breaks = 2:10) # Ensure x-axis ticks correspond to the
441
443
444
445
       if (!require(cluster)) install.packages("cluster")
library(cluster)
446
447
448
       data_matrix <- model.matrix(~-1 + PC1 + PC2+ PC3, datuak2)</pre>
449
```

```
# Set parameters for the gap statistic
set.seed(123) # For reproducibility
max_clusters <- 10 # Maximum number of clusters to test</pre>
452
453
454
455
456
457
        K.max = max_clusters,
B = 50, # Number of iter.max = 10)
458
                                                                                bootstrap samples for reference data
459
460
461
        # Display the gap statistic results
print(gap_stat)
462
463
464
         # Improved gap statistic plot
465
       # Improved gap statistic plot
plot(gap_stat,
    main = "Gap Statistic for Optimal Number of Clusters",  # Title with a larger font
    xlab = "Number of Clusters",  # X-axis label
    ylab = "Gap Statistic",  # Y-axis label
    col = "blue",  # Line color
    lwd = 2,  # Line width for better vis:
    pch = 16,  # Point type (filled circles
    cex = 1.2,  # Point size
    col main = "darkblue"  # Title color
\frac{466}{467}
468
469
470
471
472
473
                cex = 1.2,
col.main = "darkblue",
col.lab = "darkgreen",
cex.main = 1.5,
cex.lab = 1.2,
474
476
                cex.axis = 1.1,
font.lab = 2,
478
                 font.main = 2,
xlim = c(1, length(gap_stat$Tab[,1]))
480
       )
482
484
        # Add gridlines for better readability
grid(col = "gray", lty = "dotted")
486
        #Add horizontal line at y = 0 to better highlight the gaps
abline(h = 0, col = "red", lty = 2)
488
489
490
492
493
494
495
         set.seed(142)
         kmeans_result <- kmeans(model.matrix(~-1 +PC1 + PC2+ PC3, datuak2), centers = 4, iter.max = 1000000)
496
497
498
499
         kmeans_result$size
500
        # Calculate Cluster Centroids from kmeans_result$centers
centroids2_df <- as.data.frame(kmeans_result$centers)</pre>
501
503
        # Assign column names based on your principal com
colnames(centroids2_df) <- c("PC1", "PC2", "PC3")</pre>
504
505
506
         # Add cluster identifiers
centroids2_df$cluster <- factor(1:nrow(centroids2_df))</pre>
507
508
509
510
511
         print(centroids2_df)
512
513
         kmeans_result$withinss # Within Cluster Sum of Squares
515
        kmeans_result$tot.withinss # Total Within Sum of Squares
kmeans_result$betweenss # Between Cluster Sum of Squares
517
        # Add cluster assignments to the original dataset
datuak2$cluster <- kmeans_result$cluster</pre>
519
521
        assignCluster <- function(kmeans_result, data) {
   data$cluster <- kmeans_result$cluster
   return(data)
}</pre>
526
         assignCluster(kmeans_result, datuak2)
         datuak2 <- assignCluster(kmeans_result, datuak2)</pre>
529
530
        library(plotly)
531
```

```
#Plot with colours each individual
fig <- plot_ly(
   data = datuak2,
   x = "PC1,
   y = "PC2,
   z = "PC3,
   type = "scatter3d"</pre>
             z = TPC3,
type = "scatter3d",
mode = "markers",
marker = list(size = 5),
color = "factor(cluster), # Mapear clusters como factores
colors = c("blue", "red", "green", "purple"), # Definir colores para 4 clusters
showlegend = TRUE # Asegurar que la leyenda est visible
540
541
542
543
544
545
546
          #Add centroids
fig <- fig %>% add_trace(
    data = centroids2_df,
    x = "PC1,
    y = "PC2,
    z = "PC3,
    type = "scatter3d",
    mode = "markers+text",
    text = paste("Centroid", centroids2_df$cluster),
    marker = list(
        size = 10,
        symbol = "x",
        line = list(width = 3),
        color = "factor(cluster),
        colors = c("blue", "red", "green", "purple")
),
547
548
549
550
551
553
554
555
556
558
560
561
562
563
              showlegend = FALSE
564
          #Legend
fig <- fig %>% layout(
    title = "3D K-Means Clustering with Centroids (4 Clusters)",
    scene = list(
566
567
568
569
                 xaxis = list(title = "PC1"),
yaxis = list(title = "PC2"),
zaxis = list(title = "PC3")
571
572
          )
573
574
575
           # Mostrar el gr fico
576
          fig
577
578
579
580
          #FINAL INTERPRETATION#
#Add the corresponding variable
datuak$class<-y</pre>
581
582
583
           datuak2$class
584
585
           library(ggplot2)
586
587
588
           datuak2$survived=datuak2$class==1
           # Calculate the proportion of survivors within each cluster
survival_proportion <- tapply(datuak2$survived, datuak2$cluster, function(x) mean(x, na.rm = TRUE))</pre>
589
591
592
         593
594
595
597
599
600
          # Add text labels to the bars showing the protext(x = seq_along(survival_proportion),
    y = survival_proportion + 0.05,
    labels = round(survival_proportion, 2),
    col = "black")
601
603
604
605
606
607
           Survival_data <- data.frame(

Cluster = factor(1:length(survival_proportion)), # Cluster labels

Proportion = survival_proportion # Survival proportions
608
609
610
611
612
613
        ggplot(survival_data, aes(x = Cluster, y = Proportion)) +
```

```
616
617
619
621
623
624
         datuak$survived=datuak$class==1
         # Calculate the proportion of survivors within each cluster survival_proportion <- tapply(datuak$survived, datuak$hclus.label, function(x) mean(x, na.rm = TRUE))
625
627
         628
629
630
631
632
633
634
635
636
637
         text(x = seq_along(survival_proportion),
    y = survival_proportion + 0.05, # Position the text slightly above the bars
    labels = round(survival_proportion, 2),
    col = "black")
638
640
642
         # Using ggplot2 for better visualization
survival_data <- data.frame(
   Cluster = factor(1:length(survival_proportion)), # Cluster labels
   Proportion = survival_proportion # Survival proportions</pre>
644
646
648
         # Plot using ggplot2
ggplot(survival_data, aes(x = Cluster, y = Proportion)) +
   geom_bar(stat = "identity", fill = "blue", color = "black") +
   geom_text(aes(label = round(Proportion, 2)), vjust = -0.5, color = "black") +
   labs(title = "Proportion of Survivors in Each Cluster",
        x = "Cluster",
        y = "Proportion Survived") +
   theme_minimal()
649
650
651
652
653
654
655
656
```

Listing 2: PCA and Clustering Code