

# Logistic Regression

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# Discriminative Linear Models – Logistic Regression

Logistic regression, despite its name, is a discriminative approach for classification

Rather than modeling the distribution of observed samples  $X|C$ , we directly model the class posterior distribution  $C|X$

We need to define a model for the class posterior distribution  $P(C = c|X = \mathbf{x})$

# Discriminative Linear Models – Logistic Regression

For a 2-class problem, we have seen that the Gaussian model with tied covariance provides log-likelihood ratios that are linear functions of our data

$$l(\mathbf{x}) = \log \frac{f_{X|C}(\mathbf{x}|h_1)}{f_{X|C}(\mathbf{x}|h_0)} = \mathbf{w}^T \mathbf{x} + c$$

and the class log-posterior probability ratio is

$$\log \frac{P(C = h_1|\mathbf{x})}{P(C = h_0|\mathbf{x})} = \log \frac{f_{X|C}(\mathbf{x}|h_1)}{f_{X|C}(\mathbf{x}|h_0)} + \log \frac{\pi}{1 - \pi} = \mathbf{w}^T \mathbf{x} + b$$

The prior information (including possible non-uniform costs) has been absorbed in the bias term  $b$ .

Classification rules take the form of linear functions

# Logistic Regression

We now adopt a complementary approach: we **assume** that our decision rule should be a **linear rule** represented by  $(\mathbf{w}, b)$

Given  $\mathbf{w}$  and  $b$ , we can compute the expression for the posterior class probability as

$$\begin{aligned}P(C = h_1|\mathbf{x}, \mathbf{w}, b) &= e^{(\mathbf{w}^T \mathbf{x} + b)} P(C = h_0|\mathbf{x}, \mathbf{w}, b) \\&= e^{(\mathbf{w}^T \mathbf{x} + b)} (1 - P(C = h_1|\mathbf{x}, \mathbf{w}, b))\end{aligned}$$

Solving for  $P(C = h_1|\mathbf{x}, \mathbf{w}, b)$  we obtain

$$P(C = h_1|\mathbf{x}, \mathbf{w}, b) = \frac{e^{(\mathbf{w}^T \mathbf{x} + b)}}{1 + e^{\mathbf{w}^T \mathbf{x} + b}} = \frac{1}{1 + e^{-(\mathbf{w}^T \mathbf{x} + b)}} = \sigma(\mathbf{w}^T \mathbf{x} + b)$$

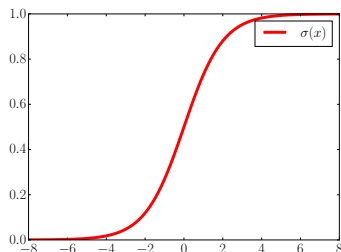
where

$$\sigma(x) = \frac{1}{1 + e^{-x}}$$

is called *sigmoid* function (or *logistic* function)

# Logistic Regression

Sigmoid function:



Some properties of  $\sigma(x)$  that will come useful later

- $1 - \sigma(x) = \sigma(-x)$
- $\frac{d\sigma(x)}{dx} = \sigma(x) (1 - \sigma(x))$

# Logistic Regression

The expression  $P(C = h_1 | \mathbf{x}, \mathbf{w}, b) = \sigma(\mathbf{w}^T \mathbf{x} + b)$  provides a **model** that allows computing the posterior probabilities for  $h_1$  and  $h_0$

The model **assumes** that decision rules are **linear surfaces** (hyperplanes) orthogonal to  $\mathbf{w}$ . The model parameters are  $(\mathbf{w}, b)$

If we knew  $(\mathbf{w}, b)$  then we could compute the predictive distribution for the class labels  $P(C = h_1 | \mathbf{x}, \mathbf{w}, b)$

We have seen an **indirect** way of computing  $(\mathbf{w}, b)$  with a generative model

- The tied covariance Gaussian feature model, combined with application class prior, can be cast as a linear decision rule of the form  $\mathbf{w}^T \mathbf{x} + b \leq 0$

# Logistic Regression

However, in the following we are interested to find an alternative way to **estimate an effective classification rule** that **does not require** an explicit model of the feature vectors distribution

We thus ignore generative models for  $X$  and concentrate directly on the form of the class posterior probabilities

Again, we follow a frequentist approach, i.e. compute an estimate for  $w$  and  $b$  from a set of training samples

# Logistic Regression

We assume we have a labeled dataset

$$\mathcal{D} = [(\mathbf{x}_1, c_1), \dots (\mathbf{x}_n, c_n)]$$

We also assume that feature vectors and corresponding class labels are **independent and identically distributed** (i.i.d.). given the model parameters (we made the same assumption for generative models):

$$[(\mathbf{X}_i, C_i) \perp\!\!\!\perp (\mathbf{X}_j, C_j)] \mid \boldsymbol{\theta}$$

Here  $\boldsymbol{\theta}$  are the model parameters  $\boldsymbol{\theta} = (\mathbf{w}, b)$

As for generative models, the **complete-data likelihood** for  $\boldsymbol{\theta}$  consists of the joint density of the observed training set variables, given the parameter vector  $\boldsymbol{\theta}$ :

$$\mathcal{L}(\boldsymbol{\theta}) = f_{\mathbf{X}_1 \dots \mathbf{X}_n, C_1 \dots C_n \mid \boldsymbol{\theta}}(\mathbf{x}_1 \dots \mathbf{x}_n, c_1 \dots c_n \mid \boldsymbol{\theta})$$



# Logistic Regression

Since we assume i.i.d. observations, we can factorize the likelihood as

$$\begin{aligned}\mathcal{L}(\theta) &= f_{X_1 \dots X_n, C_1 \dots C_n | \theta}(\mathbf{x}_1 \dots \mathbf{x}_n, c_1 \dots c_n | \theta) \\ &= \prod_{i=1}^n f_{X, C | \theta}(\mathbf{x}_i, c_i | \theta)\end{aligned}$$

The difference with respect to generative models is in the way we represent the joint density  $f_{X, C | \theta}$ :

$$f_{X, C | \theta}(\mathbf{x}_i, c_i) = P(C = c_i | \mathbf{X} = \mathbf{x}_i, \theta) f_X(\mathbf{x}_i)$$

Note that, since classification requires computing posterior class probabilities, and we are defining an explicit model for such terms, we do not need to explicitly provide an expression for the **marginal feature density**  $f_X(\mathbf{X})$

# Logistic Regression

The complete-data log-likelihood becomes

$$\begin{aligned}\log \mathcal{L}(\boldsymbol{\theta}) &= \log \prod_{i=1}^n f_{X,C|\boldsymbol{\theta}}(\mathbf{x}_i, c_i|\boldsymbol{\theta}) \\ &= \sum_{i=1}^n \log P(C = c_i | \mathbf{X}_i = \mathbf{x}_i, \boldsymbol{\theta}) + \sum_{i=1}^n \log f_X(\mathbf{x}_i)\end{aligned}$$

As for generative models, we can estimate the model parameters by following a Maximum Likelihood approach

$$\boldsymbol{\theta}_{ML} = \arg \max_{\boldsymbol{\theta}} \log \mathcal{L}(\boldsymbol{\theta})$$

Since the model parameters  $\boldsymbol{\theta}$  influence only the first sum, optimization of the log-likelihood corresponds to the **maximization of the conditional probability** of the **observed dataset labels**, given the observed feature vectors and the model parameters

$$\ell(\boldsymbol{\theta}) = \sum_{i=1}^n \log P(C = c_i | \mathbf{X}_i = \mathbf{x}_i, \boldsymbol{\theta})$$

# Logistic Regression

We now need to explicit our model in the expression for  $\ell$

We assume that the label for class  $h_1$  is 1, and the label for class  $h_0$  is 0

Our model specifies the probabilities for observed class labels in terms of  $\mathbf{w}$  and  $b$ :

$$y_i = P(C_i = 1|\mathbf{x}_i, \mathbf{w}, b) = \sigma(\mathbf{w}^T \mathbf{x}_i + b)$$

It follows that

$$P(C_i = 0|\mathbf{x}_i, \mathbf{w}, b) = 1 - y_i = 1 - \sigma(\mathbf{w}^T \mathbf{x}_i + b) = \sigma(-\mathbf{w}^T \mathbf{x} - b)$$

# Logistic Regression

We note that  $C_i|\mathbf{x}_i, \mathbf{w}, b$  follows a Bernoulli distribution

$$C_i|\mathbf{x}_i, \mathbf{w}, b \sim \text{Ber}(\sigma(\mathbf{w}^T \mathbf{x}_i + b)) = \text{Ber}(y_i)$$

The conditional probability of the class labels, i.e. our **objective function**, is thus given by

$$\begin{aligned}\ell(\mathbf{w}, b) &= \log \prod_{i=1}^n P(C_i = c_i | \mathbf{x}_i, \mathbf{w}, b) \\ &= \log \prod_i y_i^{c_i} (1 - y_i)^{(1-c_i)} \\ &= \sum_{i=1}^n [c_i \log y_i + (1 - c_i) \log(1 - y_i)]\end{aligned}$$

# Logistic Regression

Our goal is the maximization of  $\ell$ , which corresponds to the **maximization of the likelihood function**, i.e., a Maximum Likelihood solution. We thus seek  $\mathbf{w}^*, b^*$  that maximize  $\ell(\mathbf{w}, b)$ :

$$\mathbf{w}^*, b^* = \arg \max_{\mathbf{w}, b} \ell(\mathbf{w}, b) = \arg \max_{\mathbf{w}, b} \sum_{i=1}^n [c_i \log y_i + (1 - c_i) \log(1 - y_i)]$$

We can show that the ML solution is also the solution that minimizes the average **cross-entropy** between the distribution of observed and predicted labels

Rather than maximizing  $\ell(\mathbf{w}, b)$ , we can **minimize**

$$J(\mathbf{w}, b) = -\ell(\mathbf{w}, b) = \sum_{i=1}^n -[c_i \log y_i + (1 - c_i) \log(1 - y_i)]$$

The expression

$$H(c_i, y_i) = -[c_i \log y_i + (1 - c_i) \log (1 - y_i)]$$

represents the binary *cross-entropy* between the distribution of observed and predicted labels for the  $i$ -th sample

More in general, let  $P$  and  $Q$  be two distributions over the same domain

The cross-entropy between the two distributions is defined as

$$H(P, Q) = -\mathbb{E}_{P(x)} [\log Q(x)]$$

For discrete distributions, this can be expressed as

$$H(P, Q) = -\sum_{x \in \mathcal{S}} P(x) \log Q(x)$$

In our case,  $P$  is the empirical distribution of class labels, **from the point of view of an observer  $\mathcal{E}$  who knows the actual label:**

$$P(C_i = 1 | \mathbf{X}_i = \mathbf{x}_i, \mathcal{E}) = \begin{cases} 1 & \text{if } c_i = 1 \\ 0 & \text{if } c_i = 0 \end{cases}$$

$$P(C_i = 0 | \mathbf{X}_i = \mathbf{x}_i, \mathcal{E}) = \begin{cases} 0 & \text{if } c_i = 1 \\ 1 & \text{if } c_i = 0 \end{cases}$$

or, equivalently

$$P(C_i = 1 | \mathbf{X}_i = \mathbf{x}_i, \mathcal{E}) = c_i, \quad P(C_i = 0 | \mathbf{X}_i = \mathbf{x}_i, \mathcal{E}) = 1 - c_i$$

i.e., a Bernoulli distribution with parameter  $c_i$

# Logistic Regression

Distribution  $Q$  is the distribution for the **predicted labels according to our recognizer  $\mathcal{R}$**

$$Q(c) = P(C_i = c | X_i = x_i, \mathcal{R}(\mathbf{w}, b))$$

i.e.

$$Q(1) = P(C_i = 1 | X_i = x_i, \mathcal{R}(\mathbf{w}, b)) = y_i = \sigma(\mathbf{w}^T \mathbf{x}_i + b)$$

$$Q(0) = P(C_i = 0 | X_i = x_i, \mathcal{R}(\mathbf{w}, b)) = 1 - y_i = 1 - \sigma(\mathbf{w}^T \mathbf{x}_i + b)$$

Logistic regression looks for the **minimizer of the average cross-entropy** between the distribution for the training set labels of an evaluator  $\mathcal{E}$  **who knows the real label** and the distribution for the training set labels as **predicted by the model  $\mathcal{R}(\mathbf{w}, b)$**  itself

The cross-entropy is a measure of goodness of the predictions, and the evaluation is performed over the training data itself



The cross-entropy, as a function of  $Q$ , is minimized when  $Q = P$

The cross-entropy can also be interpreted as a measure of the difference between  $P$  and  $Q$

In our case, it measures how different is the predicted distribution  $\text{Ber}(y_i)$  from the empirical label distribution  $\text{Ber}(c_i)$  (the distribution of the evaluator  $\mathcal{E}$ )

Minimization of the average cross-entropy means we are looking for conditional label distributions as similar (on average) as possible to the empirical one, given the model constraints (i.e., a linear classification rule)

Alternatively, as we have seen, we can regard the process as maximization of the likelihood for the observed labels

Another interesting interpretation of the logistic regression objective can be obtained by rewriting the cross-entropy in terms of  $z_i = 2c_i - 1$

The terms  $z_i$  still represent class labels, however for samples of class  $h_1$  we have  $z_i = 1$ , whereas for samples of class  $h_0$  we have  $z_i = -1$ :

$$z_i = \begin{cases} 1 & \text{if } c_i = 1 \\ -1 & \text{if } c_i = 0 \end{cases}$$

# Logistic Regression

The objective function that we want to minimize corresponds to the sum of  $n$  terms

$$J(\mathbf{w}, b) = \sum_i H(c_i, y_i)$$

where

$$H(c_i, y_i) = -[c_i \log y_i + (1 - c_i) \log(1 - y_i)]$$

Note that  $H(c_i, y_i)$  is a function of  $c_i$ , but also of  $\mathbf{w}$ ,  $b$  and  $\mathbf{x}_i$ , since

$$y_i = \sigma(\mathbf{w}^T \mathbf{x}_i + b)$$

Let  $s_i = \mathbf{w}^T \mathbf{x}_i + b$ . In terms of  $z_i$  we can rewrite  $H$  as

$$\begin{aligned} H(c_i, y_i) &= -[c_i \log y_i + (1 - c_i) \log(1 - y_i)] \\ &= \begin{cases} -\log \sigma(s_i) & \text{if } c_i = 1 \ (z_i = 1) \\ -\log(1 - \sigma(s_i)) = -\log \sigma(-s_i) & \text{if } c_i = 0 \ (z_i = -1) \end{cases} \\ &= -\log \sigma(z_i s_i) \\ &= -\log \sigma(z_i(\mathbf{w}^T \mathbf{x}_i + b)) \\ &= \log \left( 1 + e^{-z_i(\mathbf{w}^T \mathbf{x}_i + b)} \right) \end{aligned}$$

The objective function can thus be rewritten as

$$\begin{aligned} J(\mathbf{w}, b) &= \sum_{i=1}^n H(c_i, y_i) \\ &= \sum_{i=1}^n \log \left( 1 + e^{-z_i(\mathbf{w}^T \mathbf{x}_i + b)} \right) \\ &= \sum_{i=1}^n l(z_i(\mathbf{w}^T \mathbf{x}_i + b)) \end{aligned}$$

where

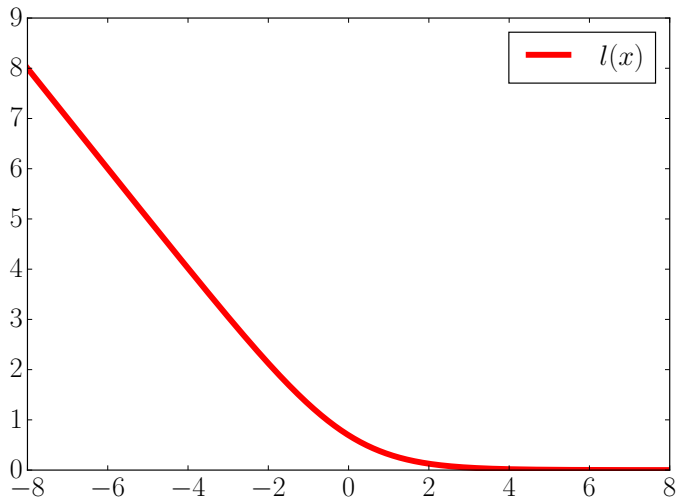
$$l(x) = \log(1 + e^{-x})$$

is the **logistic loss** function.

Our goal is to find the **minimizer** of  $J(\mathbf{w}, b)$

# Logistic Regression

## Logistic loss



# Logistic Regression

We can interpret the function as the **cost** of the prediction made with model  $(\mathbf{w}, b)$  for each sample

Remember that the log-posterior class probability ratio for sample  $\mathbf{x}_i$  is

$$\log \frac{P(C_i = 1 | \mathbf{X}_i = \mathbf{x}_i)}{P(C_i = 0 | \mathbf{X}_i = \mathbf{x}_i)} = \mathbf{w}^T \mathbf{x}_i + b = s_i$$

The decision rule takes the form  $s_i \lessgtr 0$

Since  $s_i = \mathbf{w}^T \mathbf{x}_i + b$ , decision rules are linear hyperplane orthogonal to the vector  $\mathbf{w}$

$s_i$  is related to the distance of the sample  $\mathbf{x}_i$  from the separating surface

# Logistic Regression

When  $s_i$  is positive, our classifier is favoring class  $h_1$ , whereas negative  $s_i$  means we are classifying the sample as belonging to class  $h_0$ .

The cost we pay for each sample is  $l(z_i s_i)$

- The prediction and the actual class agree:  $z_i = 1, s_i > 0$  or  $z_i = -1, s_i < 0$ . Then  $z_i s_i > 0$ , and we pay a low cost. The cost becomes exponentially smaller (asymptotically) as the absolute value of  $s_i$  increases (we move away from the separation surface)
- The prediction and the actual class disagree:  $z_i = 1, s_i < 0$  or  $z_i = -1, s_i > 0$ . Then  $z_i s_i < 0$ , and we pay a cost that increases (asymptotically) linearly with  $s_i$



# Logistic Regression

We can thus interpret the logistic regression objective as a measure of an **empirical risk**<sup>1</sup>. Our goal is minimizing the empirical risk

More in general, empirical risk minimization is a framework for the estimation of classification models which aims at minimizing an **empirical risk function** over our training data

Generalized risk minimization problem: minimize the risk  $R(\theta)$

$$R(\theta) = \sum_i l(\mathbf{w}, \mathbf{x}_i, z_i)$$

where  $l$  is called loss (or cost) function, and  $\theta$  are the parameters of the classification model, e.g.  $\theta = (\mathbf{w}, b)$  in our case.

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<sup>1</sup>Empirical because it's computed on the observed samples

# Logistic Regression

Logistic regression solutions cannot be computed in closed form

We will resort to numerical solvers

A numerical solver iteratively looks for the minimizer of a function

We will use the L-BFGS algorithm

The algorithm requires a function that computes the loss and its gradient with respect to  $w$  and  $b$

In the laboratory we will see how to implement the minimization

# Logistic Regression

If classes are linearly separable, the logistic regression solution is not defined

**Linearly separable classes:** there exist  $\mathbf{w}$  and  $b$  such that all training samples lie on the correct side of the corresponding separation surface ( $z_i > 0 \iff s_i > 0$ )

In this case, we can make the values of  $s_i$  arbitrarily high by simply increasing the norm of  $\mathbf{w}$  (and changing accordingly the value of  $b$ )

As we increase  $\|\mathbf{w}\|$ , the loss becomes lower, thus we are decreasing the objective function

The function does not have a minimum, but has an infimum  $\inf J(\mathbf{w}, b) = 0$ , corresponding to  $\|\mathbf{w}\| \rightarrow \infty$

# Logistic Regression

To make the problem solvable again, we can look for solutions with small norm by introducing a norm penalty to the objective function

The penalty is called **regularization** term

The objective function that we minimize is

$$\tilde{R}(\mathbf{w}, b) = \frac{\tilde{\lambda}}{2} \|\mathbf{w}\|^2 + \sum_{i=1}^n \log \left( 1 + e^{-z_i(\mathbf{w}^T \mathbf{x}_i + b)} \right)$$

where  $\tilde{\lambda}$  is a hyper-parameter that allows specifying the relative weight of the regularization term

Alternatively, we look for the minimizer of

$$R(\mathbf{w}, b) = \frac{\lambda}{2} \|\mathbf{w}\|^2 + \frac{1}{n} \sum_{i=1}^n \log \left( 1 + e^{-z_i(\mathbf{w}^T \mathbf{x}_i + b)} \right)$$

where the risk is *averaged* over all samples, and  $\lambda$  is the regularization coefficient.

# Logistic Regression

$\lambda$  is a hyper-parameter, and should be selected as to optimize the performance of the classifier

Note that  $\lambda$  cannot be computed by minimizing  $R$  with respect to  $\lambda$ , as we would obtain the trivial solution  $\lambda = 0$

The selection of good values for  $\lambda$  should thus be based on other approaches, such as **cross-validation**

The model is called **regularized** Logistic Regression, and is an example of a **regularized risk minimization** problem

$$R(\mathbf{w}, b) = \Omega(\mathbf{w}, b) + \frac{1}{n} \sum_{i=1}^n l(\mathbf{x}_i, z_i, \mathbf{w}, b)$$

The regularization term  $\Omega$  (in our case  $\frac{\lambda}{2} \|\mathbf{w}\|^2$ ) can be interpreted as a term that favors simpler solutions (we will see explicitly why small norm of  $\mathbf{w}$  can be interpreted as a simpler solution when discussing Support Vector Machines)

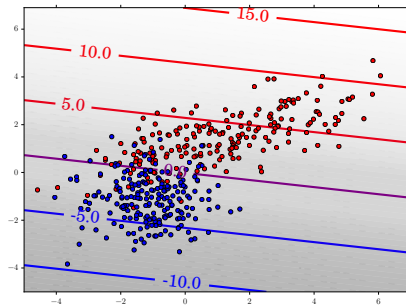
Regularization allows reducing the risk of over-fitting the training data

Of course, if  $\lambda$  is too large, we will obtain a solution that has small norm, but is not able to well separate the classes

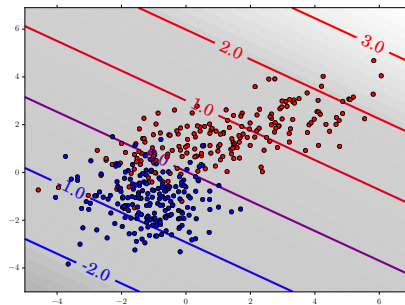
On the other hand, if  $\lambda$  is too small, we will get a solution that has good separation on the training set, but may have poor classification accuracy for unseen data (i.e. poor **generalization**)

# Logistic Regression

$\lambda = 0$



$\lambda = 1$



Some considerations:

- The non-regularized model is invariant to linear transformations of the feature vectors.
- The regularized version of the model, on the other hand, is not invariant.
- It is therefore useful, in some cases, to pre-process data so that dynamic ranges of different features are similar
- Cross-validation can help in identifying good pre-processing strategies



Common **preprocessing strategies** that may be worth trying:

- Center the data (either using the training set, or a weighted mean — e.g. the average of class means):  $\mathbf{x}'_i = \mathbf{x}_i - \boldsymbol{\mu}$
- Standardize variances (e.g. divide each feature by its own standard deviation computed over the training set):  $\mathbf{x}'_{i,[j]} = \mathbf{x}_{i,[j]} / \sigma_{[j]}$
- Whiten the covariance matrix (i.e. normalize variances while making features uncorrelated):  $\mathbf{x}'_i = \mathbf{A}\mathbf{x}_i$ , where  $\mathbf{A} = \boldsymbol{\Sigma}^{-\frac{1}{2}}$  and  $\boldsymbol{\Sigma}$  is the training set covariance (a variation consists in replacing  $\boldsymbol{\Sigma}$  with the within-class covariance)
- L2 (or length) normalization:  $\mathbf{x}'_i = \frac{\mathbf{x}_i}{\|\mathbf{x}_i\|}$  (often after centering and whitening)

# Logistic Regression

The logistic regression score can be interpreted as the logarithm of the ratio between class posterior probabilities

The model parameters have been trained as to optimize the probability of the **training set labels**

Therefore, the model will implicitly reflect the **empirical class prior** of the **training set**

The model posterior probabilities are thus suited for applications whose effective prior is close to the empirical training set prior, but may provide poor performance for different applications

To soften this issue, we can recover a score that behaves like a log-likelihood ratio by subtracting from the score  $s$  the empirical prior log-odds of the training set  $\log \frac{n_T}{n_F}$

$$s_{llr} = \mathbf{w}^T \mathbf{x} + b - \log \frac{n_T}{n_F} = \mathbf{w}^T \mathbf{x} + b - \log \frac{\pi_{emp}^{tr}}{1 - \pi_{emp}^{tr}}$$

We can then use  $s_{llr}$  as a log-likelihood ratio

For a given application  $(\pi_T, 1, 1)$  we can compute decisions by comparing  $s_{llr} \lessgtr -\log \frac{\pi_T}{1-\pi_T}$

# Logistic Regression

While this allows for effective decisions for the model  $(\mathbf{w}, b)$ , the orientation of  $\mathbf{w}$  is still optimized for the training set empirical prior

Changing the threshold allows us to select a better decision rule, but only among hyperplanes that are orthogonal to  $\mathbf{w}$

If we know the application prior  $\pi_T$  before training the model, then it may be preferable to directly optimize the separation rule for the target application prior

The model is sometimes called prior-weighted logistic regression:

$$R(\mathbf{w}) = \frac{\lambda}{2} \|\mathbf{w}\|^2 + \frac{\pi_T}{n_T} \sum_{i|z_i=1} l(z_i s_i) + \frac{1 - \pi_T}{n_F} \sum_{i|z_i=-1} l(z_i s_i)$$

# Logistic Regression

The standard logistic regression model corresponds to the prior-weighted logistic regression model trained with the **training set empirical prior**  $\pi_{emp}^{tr}$

Also in this case, we can compute a llr-like score

$$s_{llr} = \mathbf{w}^T \mathbf{x} + b - \log \frac{\pi_T}{1 - \pi_T}$$

that can then be employed to estimate optimal decisions for other applications, should the need arise

# Logistic Regression

We test the model on MNIST digit pairs (e.g. 0 vs 1, 0 vs 2, ...)

MNIST — Average Pairwise EER for Logistic Regression

DimRed	$\lambda = 0$	$\lambda = 0.00001$	$\lambda = 0.001$	$\lambda = 0.1$	Tied Gau
RAW [768]	1.7%	1.4%	1.2%	2.0%	—
PCA [50]	1.4%	1.4%	1.4%	2.1%	1.7%
PCA [100]	1.3%	1.2%	1.2%	2.0%	1.5%

LogReg obtains better performance than the Gaussian model

Regularization is important, especially when we do not reduce the dimensionality (we have more parameters to estimate, so over-fitting is more severe)

If we regularize too much the model performs poorly again

# Multiclass Logistic Regression

We now consider a problem with  $K$  classes, labeled from 1 to  $K$

To extend the Logistic Regression model to multiclass tasks we start again from the form of the posterior probabilities of the Linear Gaussian classifier with uniform priors

$$\log P(C = j|\mathbf{x}) = \mathbf{w}_j^T \mathbf{x} + \mathbf{b}_j + k(\mathbf{x})$$

where  $k(\mathbf{x})$  collects terms that depend on the sample  $\mathbf{x}$ , but not on the class  $j$ .

It follows that, as a function of the class,

$$P(C = j|\mathbf{x}) \propto e^{\mathbf{w}_j^T \mathbf{x} + \mathbf{b}_j}$$

# Multiclass Logistic Regression

Since we have a closed-set classification problem, then

$$\sum_{j=1}^K P(C = j|\mathbf{x}) = 1$$

The normalization factor for  $P(C = j|\mathbf{x})$  is thus

$$\sum_{j=1}^K e^{\mathbf{w}_j^T \mathbf{x} + b_j}$$

and the posterior probability corresponds to

$$P(C = k|\mathbf{x}) = \frac{e^{\mathbf{w}_k^T \mathbf{x} + b_k}}{\sum_j e^{\mathbf{w}_j^T \mathbf{x} + b_j}}$$

Function  $f(\mathbf{s}) = \left[ \frac{e^{s_1}}{\sum_j e^{s_j}}, \dots, \frac{e^{s_K}}{\sum_j e^{s_j}} \right]$  is called **softmax**



# Multiclass Logistic Regression

Given the model parameters

$$\mathbf{W} = [\mathbf{w}_1 \quad \dots \quad \mathbf{w}_K] , \quad \mathbf{b} = \begin{bmatrix} b_1 \\ \vdots \\ b_K \end{bmatrix}$$

the logistic regression model allows computing the probability of each class

$$P(C = k | \mathbf{W}, \mathbf{b}, \mathbf{x}) = \frac{e^{\mathbf{w}_k^T \mathbf{x} + b_k}}{\sum_j e^{\mathbf{w}_j^T \mathbf{x} + b_j}}$$

If we consider sample  $\mathbf{x}_i$ , its class posterior distribution is thus a **categorical** distribution

$$C_i | \mathbf{W}, \mathbf{b}, \mathbf{X}_i = \mathbf{x}_i \sim \text{Cat}(\mathbf{y}_i)$$

where

$$\mathbf{y}_{ik} = \frac{e^{\mathbf{w}_k^T \mathbf{x}_i + b_k}}{\sum_j e^{\mathbf{w}_j^T \mathbf{x}_i + b_j}}$$

# Multiclass Logistic Regression

As for the binary case, we can express the log-probability of the training class labels as

$$\ell(\mathbf{W}, \mathbf{b}) = \sum_{i=1}^n \log P(C_i = c_i | \mathbf{X}_i = \mathbf{x}_i, \mathbf{W}, \mathbf{b})$$

Remember that the categorical density  $P(C_i = c_i | \mathbf{X}_i = \mathbf{x}_i, \mathbf{W}, \mathbf{b})$  can be expressed using a 1-of-K encoding, as

$$\log P(C_i = c_i | \mathbf{X}_i = \mathbf{x}_i, \mathbf{W}, \mathbf{b}) = \log P(C_i = c_i | \mathbf{y}_i) = \sum_{k=1}^K \mathbf{z}_{ik} \log \mathbf{y}_{ik}$$

where  $\mathbf{z}_i$  is a vectors that has all component equal to 0, except for the index  $c_i$  which is equal to 1

$$\mathbf{z}_i = [0 \dots 0, 1, 0 \dots 0] \text{ , } \mathbf{z}_{ik} = \begin{cases} 1 & \text{if } c_i = k \\ 0 & \text{otherwise} \end{cases}$$

# Multiclass Logistic Regression

The labels log-probability can thus be expressed as

$$\ell(\mathbf{W}, \mathbf{b}) = \sum_{i=1}^n \sum_{k=1}^K z_{ik} \log y_{ik}$$

The terms  $y_{ik}$  are, again,

$$y_{ik} = \frac{e^{\mathbf{w}_k^T \mathbf{x}_i + b_k}}{\sum_j e^{\mathbf{w}_j^T \mathbf{x}_i + b_j}}$$

and represent the distribution for the class labels according to the Logistic Regression model

# Multiclass Logistic Regression

As for the binary case, the expression

$$H(\mathbf{z}_i, \mathbf{y}_i) = - \sum_{k=1}^K z_{ik} \log y_{ik}$$

represents the (multiclass) cross-entropy between the observed and predicted label distributions for sample  $\mathbf{x}_i$

As for the binary case, we estimate  $\mathbf{W}$  and  $\mathbf{b}$  as to maximize the likelihood for the training labels

The ML solution is again the solution that minimizes the (average) cross-entropy:

$$\arg \max_{\mathbf{W}, \mathbf{b}} \ell(\mathbf{W}, \mathbf{b}) = \arg \max_{\mathbf{W}, \mathbf{b}} \left[ - \sum_{i=1}^n H(\mathbf{z}_i, \mathbf{y}_i) \right] = \arg \min_{\mathbf{W}, \mathbf{b}} \sum_{i=1}^n H(\mathbf{z}_i, \mathbf{y}_i)$$

# Multiclass Logistic Regression

Compared to the binary case, the model is over-parametrized (i.e., we can add a constant vector to all terms  $w_i$  without changing the model)

In particular, for a 2-class problem, if we subtract  $w_2$  from both  $w_1$  and  $w_2$ , we recover exactly the binary logistic regression objective.

# Multiclass Logistic Regression

Finally, as for the binary class, we can cast the problem as a minimization of a loss function

We rewrite the objective in terms of class labels  $c$  as

$$\begin{aligned} J(\mathbf{W}, \mathbf{b}) &= - \sum_{i=1}^n \sum_{k=1}^K z_{ik} \log y_{ik} \\ &= - \sum_{i=1}^n \log \frac{e^{\mathbf{w}_{c_i}^T \mathbf{x}_i + b_{c_i}}}{\sum_{c'=1}^K e^{\mathbf{w}_{c'}^T \mathbf{x}_i + b_{c'}}} \\ &= \sum_{i=1}^n \left[ \log \left( \sum_{c'=1}^K e^{\mathbf{w}_{c'}^T \mathbf{x}_i + b_{c'}} \right) - \mathbf{w}_{c_i}^T \mathbf{x}_i - b_{c_i} \right] \\ &= \sum_{i=1}^n l(\mathbf{x}_i, c_i, \mathbf{W}, \mathbf{b}) \end{aligned}$$

$l$  is also called softmax loss

# Multiclass Logistic Regression

Again, we can add a regularization term to reduce over-fitting. We thus look for the minimizer of

$$R(\mathbf{W}, \mathbf{b}) = \Omega(\mathbf{W}) + \frac{1}{n}J(\mathbf{W}, \mathbf{b})$$

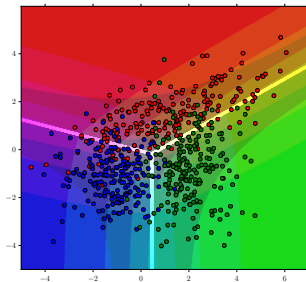
Different regularizers can be used, for example

$$\Omega(\mathbf{w}_1, \dots, \mathbf{w}_N) = \frac{1}{2} \sum_i \|\mathbf{w}_i\|^2$$

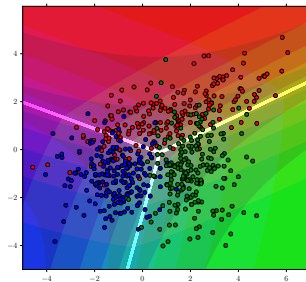
Again, we can replace the average cross-entropy with prior-weighted average cross entropy to account for priors that are different from the empirical training set prior

# Multiclass Logistic Regression

Logistic Regression



Gaussian





# Multiclass Logistic Regression

MNIST — Error rates for Logistic Regression

DimRed	$\lambda = 0$	$\lambda = 0.00001$	$\lambda = 0.001$	$\lambda = 0.1$	Tied Gau
RAW [768]	8.0%	7.4%	7.9%	12.9%	—
PCA [50]	8.8%	8.8%	8.9%	13.3%	12.6 %
PCA [100]	7.8%	7.8%	8.2%	12.9%	12.3%
PCA+LDA [9]	10.9%	10.9%	11.0%	12.4%	12.3 %

The multiclass logistic regression performs better than the Gaussian model — indeed, the Gaussian assumption is not very accurate for the features we are considering. LogReg **assumes linear separation**, but **does not assume a specific distribution for the features**

Again, regularization is important, especially when the feature space is large

# Multiclass Logistic Regression

Linear logistic regression on MNIST performs better than our Tied-Covariance Gaussian classifier, however it's far worse than our non-linear Gaussian classifier

Remember that, for binary LR, we assumed linear separation surfaces

$$\log \frac{P(C = h_1|\mathbf{x})}{P(C = h_0|\mathbf{x})} = \mathbf{w}^T \mathbf{x} + b$$

which has the same form as the Gaussian classifier with tied covariance

For Gaussian classifier with non-tied covariances we have

$$\log \frac{P(C = h_1|\mathbf{x})}{P(C = h_0|\mathbf{x})} = \mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{b}^T \mathbf{x} + c = s(\mathbf{x}, \mathbf{A}, \mathbf{b}, c)$$

# Multiclass Logistic Regression

The expression

$$s(\mathbf{x}, \mathbf{A}, \mathbf{b}, c) = \mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{b}^T \mathbf{x} + c$$

is quadratic in  $\mathbf{x}$ , however it is **linear** in  $\mathbf{A}$  and  $\mathbf{b}$

Indeed, we can rewrite  $s(\mathbf{x}, \mathbf{A}, \mathbf{b}, c)$  as

$$s(\mathbf{x}, \mathbf{A}, \mathbf{b}, c) = \langle \mathbf{x} \mathbf{x}^T, \mathbf{A} \rangle + \mathbf{b}^T \mathbf{x} + c$$

where  $\langle \cdot, \cdot \rangle$  denotes the inner product

$$\langle \mathbf{A}, \mathbf{B} \rangle = \sum_i \sum_j A_{ij} B_{ij}$$

We can further express  $\langle \mathbf{A}, \mathbf{x} \mathbf{x}^T \rangle$  as

$$\langle \mathbf{A}, \mathbf{x} \mathbf{x}^T \rangle = \text{vec}(\mathbf{x} \mathbf{x}^T)^T \text{vec}(\mathbf{A})$$

$\text{vec}(\mathbf{M})$  is the operator that stacks the columns of matrix  $\mathbf{M}$

# Multiclass Logistic Regression

If we define

$$\phi(\mathbf{x}) = \begin{bmatrix} \text{vec}(\mathbf{x}\mathbf{x}^T) \\ \mathbf{x} \end{bmatrix}$$

and

$$\mathbf{w} = \begin{bmatrix} \text{vec}(\mathbf{A}) \\ \mathbf{b} \end{bmatrix}$$

then the class log-posterior ratio can be expressed as

$$s(\mathbf{x}, \mathbf{w}, c) = \mathbf{w}^T \phi(\mathbf{x}) + c$$

We can thus train a LR model using feature vectors  $\phi(\mathbf{x})$  rather than  $\mathbf{x}$

We will obtain a model that has linear separation surface in the space defined by the mapping  $\phi$

This space is also called **expanded feature space**

# Multiclass Logistic Regression

The LR model (both binary and multiclass) allows computing linear separation rules for the transformed features  $\phi(x)$

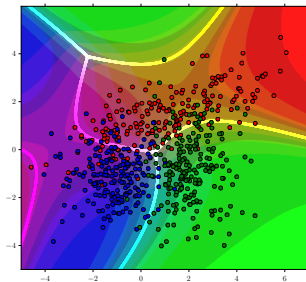
Since expressions  $w^T \phi(x) + c$  correspond to quadratic forms in the original feature space, we are actually estimating **quadratic separation surfaces** in the original space

In general, we can consider a transformation  $\phi(x)$  of our feature space such that our classes are (approximately) linearly separable in the expanded feature space

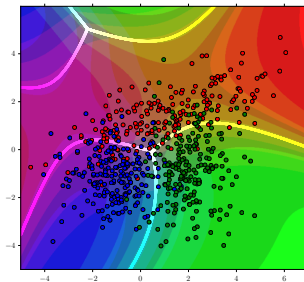
We have to pay attention that the dimensionality of the expanded feature space can grow very quickly — For example, polynomial expansions of degree  $d$  result in a feature space of dimensionality  $O(M^d)$ , where  $M$  is the dimensionality of  $x$ .

# Multiclass Logistic Regression

Logistic Regression



Gaussian



# Multiclass Logistic Regression

MNIST — Average pairwise EER for LR with quadratic feature expansion

DimRed	$\lambda = 0$	$\lambda = 1e^{-5}$	$\lambda = 1e^{-3}$	$\lambda = 1e^{-1}$	Gaussian
PCA [50]	1.0%	1.0%	0.9%	1.5%	0.8%

MNIST — Multiclass error rates for LR with quadratic feature expansion

DimRed	$\lambda = 0$	$\lambda = 1e^{-5}$	$\lambda = 1e^{-3}$	$\lambda = 1e^{-1}$	Gaussian
PCA [50]	2.3%	1.9%	1.7%	3.1%	3.6%

# Multiclass Logistic Regression

Among possible non-linear methods we will consider, depending on the course (01URTOV / 01HERUU):

- Support Vector Machines, that allow defining linear models in an **implicitly** expanded feature space
- Neural networks, that allow jointly estimating a parametric transformation  $\phi(\mathbf{x}, \mathbf{\Pi})$  and a classification rule  $(\mathbf{w}, b)$  in the expanded feature space