# Models for Stationary Time Series:

### MA Processes

Week V: Video 15

STAT 485/685, Fall 2020, SFU

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# Video 15 Learning Objectives

By the end of this video, we should be able to:

- Define a moving average process of order q, i.e. MA(q)
- Derive the mean function, autocovariance function and autocorrelation function for MA(1) and MA(2) processes
- Recognize some key properties of the general MA(q) process, including the behaviour of its autocorrelation function

# Moving Average Processes

Definition: A moving average process of order q (i.e. MA(q)) is defined as:

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2} - \theta_3 e_{t-3} - \dots - \theta_q e_{t-q}$$

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<u>Also notice:</u> The coefficients now have negative signs in front of them! This is to make some expressions easier to interpret (we will see these soon).

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What's In a Name? It is a (weighted) average of white noise process terms, where the weights get shifted one spot over each time:

$$\begin{split} Y_{t+1} &= e_{t+1} - \theta_1 e_t - \theta_2 e_{t-1} - \theta_3 e_{t-2} - \dots - \theta_q e_{t-q+1} \\ Y_{t+2} &= e_{t+2} - \theta_1 e_{t+1} - \theta_2 e_t - \theta_3 e_{t-1} - \dots - \theta_q e_{t-q+2} \\ Y_{t+3} &= e_{t+3} - \theta_1 e_{t+2} - \theta_2 e_{t+1} - \theta_3 e_t - \dots - \theta_q e_{t-q+3} \\ &\vdots \end{split}$$

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- Autocovariance function:

$$\begin{aligned} & \textit{Cov}(Y_{t}, Y_{t-k}) \\ &= \textit{Cov}(e_{t} - \theta e_{t-1}, e_{t-k} - \theta e_{t-k-1}) \\ &= \textit{Cov}(e_{t}, e_{t-k}) - \theta \textit{Cov}(e_{t}, e_{t-k-1}) - \theta \textit{Cov}(e_{t-1}, e_{t-k}) + \theta^{2} \textit{Cov}(e_{t-1}, e_{t-k-1}) \end{aligned}$$

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$$\begin{aligned} & Cov(Y_t, Y_{t-k}) \\ &= Cov(e_t - \theta e_{t-1}, e_{t-k} - \theta e_{t-k-1}) \\ &= Cov(e_t, e_{t-k}) - \theta Cov(e_t, e_{t-k-1}) - \theta Cov(e_{t-1}, e_{t-k}) + \theta^2 Cov(e_{t-1}, e_{t-k-1}) \end{aligned}$$
For  $k = 0$ :  $\sigma_a^2 + \theta^2 \sigma_a^2 = \sigma_a^2 (1 + \theta^2)$ 

For  $k = \pm 1$ :  $-\sigma_e^2 \theta$ 

Otherwise: 0

### The First-Order MA Process: Properties

$$Y_t = e_t - \theta e_{t-1}$$

Properties of the MA(1) process:

$$\mu_t = E(Y_t) = 0$$
 for all  $t$  
$$\gamma_k = Cov(Y_t, Y_{t-k}) = \begin{cases} \sigma_e^2 (1 + \theta^2) & \text{for } k = 0 \\ \sigma_e^2 (-\theta) & \text{for } k = 1 \\ 0 & \text{for } k > 1 \end{cases}$$

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Notice: The correlation cuts off after lag 1! This will be very important later.

### The First-Order MA Process: About $\rho_1$

We've seen that, for the MA(1) process:

$$\rho_1 = -\frac{\theta}{1+\theta^2}$$

$$ho_k=0 ext{ for } k>1$$

The correlations in the process are fully determined by the lag-1 autocorrelation.

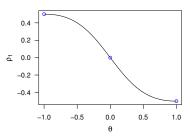
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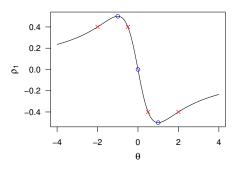


So,  $\rho_1$  can be positive or negative, depending on the value of  $\theta$ . The closer  $\theta$  is to  $\pm 1$ , the closer  $\rho_1$  is to  $\mp 0.5$ .

# The First-Order MA Process: About $\rho_1$ (cont'd)

$$ho_1 = -rac{ heta}{1+ heta^2}$$

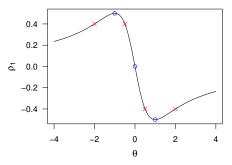
What if we consider a wider range of possible  $\theta$ -values?



# The First-Order MA Process: About $\rho_1$ (cont'd)

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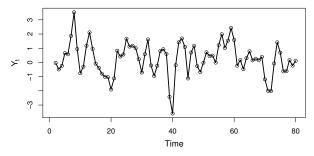


Notice:  $\rho_1$  is the same for any  $\theta$  vs.  $1/\theta$ .

This raises the question of **invertibility**: When given a value (or estimate) of  $\rho_1$ , we can't determine the true value of  $\theta$ , unless we place some restrictions on  $\theta$ .

# The First-Order MA Process: Example 1

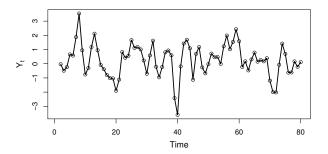
Example: MA(1) process with  $\theta = -0.9$ : (i.e.,  $\rho_1 = 0.4972$ )



```
> e.vec <- rnorm(n=80, mean=0, sd=1)
> y.vec <- rep(NA, times=80)
> for (t in 2:80)
{
     y.vec[t] <- e.vec[t] + 0.9*e.vec[t-1]
}
> plot(c(1:80), y.vec, type='o', xlab='Time', ylab=expression(Y[t]))
```

### The First-Order MA Process: Example 1 (cont'd)

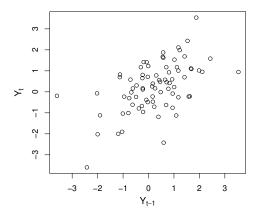
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Due to the positive lag-1 autocorrelation, observations "hang together" quite a bit.

### The First-Order MA Process: Example 1 (cont'd)

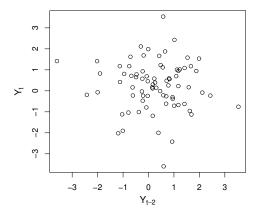
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What do we see in the plot of  $Y_t$  vs.  $Y_{t-1}$ ?

### The First-Order MA Process: Example 1 (cont'd)

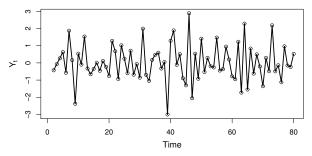
Example: MA(1) process with  $\theta = -0.9$ : (i.e.,  $\rho_1 = 0.4972$ )



What do we see in the plot of  $Y_t$  vs.  $Y_{t-2}$ ?

# The First-Order MA Process: Example 2

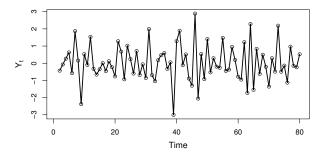
Example: MA(1) process with  $\theta = +0.9$ : (i.e.,  $\rho_1 = -0.4972$ )



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> e.vec <- rnorm(n=80, mean=0, sd=1)
> y.vec <- rep(NA, times=80)
> for (t in 2:80)
{
      y.vec[t] <- e.vec[t] - 0.9*e.vec[t-1]
}
> plot(c(1:80), y.vec, type='o', xlab='Time', ylab=expression(Y[t]))
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### The First-Order MA Process: Example 2 (cont'd)

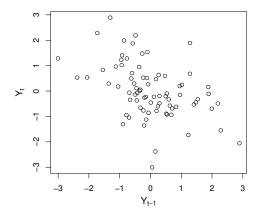
Example: MA(1) process with  $\theta = +0.9$ : (i.e.,  $\rho_1 = -0.4972$ )



Due to the negative lag-1 autocorrelation, the plot is quite "jagged" over time.

### The First-Order MA Process: Example 2 (cont'd)

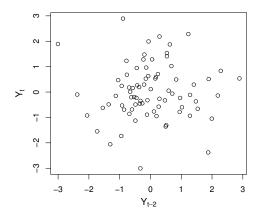
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What do we see in the plot of  $Y_t$  vs.  $Y_{t-2}$ ?

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Autocovariance function:

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$$\begin{aligned} &Cov(Y_{t}, Y_{t-k}) \\ &= Cov(e_{t} - \theta_{1}e_{t-1} - \theta_{2}e_{t-2}, e_{t-k} - \theta_{1}e_{t-k-1} - \theta_{2}e_{t-k-2}) \\ &= Cov(e_{t}, e_{t-k}) - \theta_{1}Cov(e_{t}, e_{t-k-1}) - \theta_{2}Cov(e_{t}, e_{t-k-2}) \\ &- \theta_{1}Cov(e_{t-1}, e_{t-k}) + \theta_{1}^{2}Cov(e_{t-1}, e_{t-k-1}) + \theta_{1}\theta_{2}Cov(e_{t-1}, e_{t-k-2}) \\ &- \theta_{2}Cov(e_{t-2}, e_{t-k}) + \theta_{1}\theta_{2}Cov(e_{t-2}, e_{t-k-1}) + \theta_{2}^{2}Cov(e_{t-2}, e_{t-k-2}) \end{aligned}$$

For 
$$k=0$$
:  $\sigma_e^2 + \theta_1^2 \sigma_e^2 + \theta_2^2 \sigma_e^2 = \sigma_e^2 \left(1 + \theta_1^2 + \theta_2^2\right)$ 

For 
$$k = \pm 1$$
:  $-\theta_1 \sigma_e^2 + \theta_1 \theta_2 \sigma_e^2 = \sigma_e^2 (-\theta_1 + \theta_1 \theta_2)$ 

For 
$$k = \pm 2$$
:  $\sigma_e^2(-\theta_2)$ 

Otherwise: 0

### The Second-Order MA Process: Properties

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2}$$

Properties of the MA(2) process:

$$\mu_t = E(Y_t) = 0 \quad \text{for all } t$$

$$\gamma_k = Cov(Y_t, Y_{t-k}) = \begin{cases} \sigma_e^2 (1 + \theta_1^2 + \theta_2^2) & \text{for } k = 0 \\ \sigma_e^2 (-\theta_1 + \theta_1 \theta_2) & \text{for } k = 1 \\ \sigma_e^2 (-\theta_2) & \text{for } k = 2 \\ 0 & \text{for } k > 2 \end{cases}$$

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$$\rho_k = \textit{Corr}(Y_t, Y_{t-k}) = \begin{cases} 1 & \text{for } k = 0 \\ (-\theta_1 + \theta_1 \theta_2) / (1 + \theta_1^2 + \theta_2^2) & \text{for } k = 1 \\ -\theta_2 / (1 + \theta_1^2 + \theta_2^2) & \text{for } k = 2 \\ 0 & \text{for } k > 2 \end{cases}$$

Notice: The correlation cuts off after lag 2!

### The Second-Order MA Process: Example

Example: MA(2) process:

$$Y_t = e_t - 0.9e_{t-1} + 0.7e_{t-2}$$

i.e., 
$$\theta_1 = 0.9 \ \& \ \theta_2 = -0.7.$$

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$$\rho_1 = \frac{-\theta_1 + \theta_1 \theta_2}{1 + \theta_1^2 + \theta_2^2} = \frac{-0.9 + (0.9)(-0.7)}{1 + (0.9)^2 + (-0.7)^2} = -0.67$$

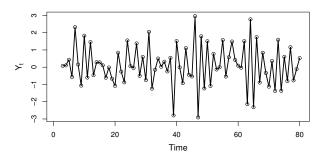
$$\rho_2 = \frac{-\theta_2}{1 + \theta_1^2 + \theta_2^2} = \frac{0.7}{1 + (0.9)^2 + (-0.7)^2} = 0.30$$

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Then:  $\rho_1 = -0.67 \& \rho_2 = 0.30$ .

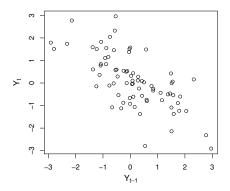


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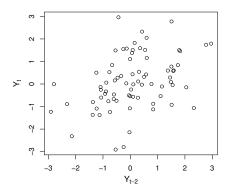
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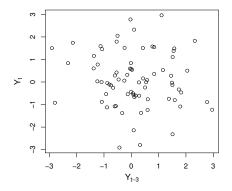
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Try 
$$q=1$$
:

$$\gamma_k = \begin{cases} \sigma_e^2 \left( 1 + \theta^2 \right) & \text{for } k = 0 \\ \sigma_e^2 \left( -\theta \right) & \text{for } k = 1 \\ 0 & \text{for } k > 1 \end{cases}$$

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Try 
$$q = 2$$
:

$$\gamma_{k} = \begin{cases} \sigma_{e}^{2} \left( 1 + \theta_{1}^{2} + \theta_{2}^{2} \right) & \text{for } k = 0 \\ \sigma_{e}^{2} \left( -\theta_{1} + \sum_{j=2}^{2} \theta_{j-1} \theta_{j} \right) = \sigma_{e}^{2} \left( -\theta_{1} + \theta_{1} \theta_{2} \right) & \text{for } k = 1 \\ \sigma_{e}^{2} \left( -\theta_{2} \right) & \text{for } k = 2 \\ 0 & \text{for } k > 2 \end{cases}$$

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# The General MA(q) Process (cont'd)

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$$\mu_t = E(Y_t) = 0$$
 for all  $t$ 

$$\gamma_k = \textit{Cov}(Y_t, Y_{t-k}) = \begin{cases} \sigma_e^2 \left(1 + \sum_{j=1}^q \theta_j^2\right) & \text{for } k = 0\\ \sigma_e^2 \left(-\theta_k + \sum_{j=k+1}^q \theta_{j-k} \theta_j\right) & \text{for } k = 1, 2, \dots, q\\ 0 & \text{for } k > q \end{cases}$$

$$\rho_k = \textit{Corr}(Y_t, Y_{t-k}) = \begin{cases} 1 & \text{for } k = 0 \\ \left(-\theta_k + \sum_{j=k+1}^q \theta_{j-k} \theta_j\right) / \left(1 + \sum_{j=1}^q \theta_j^2\right) & \text{for } k = 1, 2, \dots, q \\ 0 & \text{for } k > q \end{cases}$$

Notice: The correlation cuts off after lag q!

#### Final Comments

That's all for now!

In this video, we've learned about the moving average process of order q, i.e.  $\mathsf{MA}(q)$ .

We derived the properties for the special cases MA(1) and MA(2), and we learned how the autocorrelation function for an MA(q) process cuts off after lag q.

**Next Week in STAT 485/685:** Autoregressive (AR) processes, ARMA processes, and more.

# Thank you!

#### References:

- Cryer, J. D., & Chan, K. S. (2008). Time series analysis: with applications in R. Springer Science and Business Media.
- [2] Chan, K. S., & Ripley, B. (2020). TSA: Time Series Analysis. R package version 1.2.1.