

# **AMATH 561** Assignment 1

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**Exercise 1.1**

Let  $\mathcal{F}$  be a  $\sigma$ -algebra of  $\Omega$ . Suppose  $B \in \mathcal{F}$ . Show that  $\mathcal{G} := \{A \cap B : A \in \mathcal{F}\}$  is a  $\sigma$ -algebra of  $B$ .

**Solution**

Let  $\mathcal{F}$  be a  $\sigma$ -algebra of  $\Omega$ . Suppose  $B \in \mathcal{F}$  and define  $\mathcal{G} := \{A \cap B : A \in \mathcal{F}\}$ .

- (i) Since  $\mathcal{F}$  is a  $\sigma$ -algebra, and  $B \in \mathcal{F}$  then  $B^c \in \mathcal{F}$ . Thus,  $\emptyset = B^c \cap B \in \mathcal{G}$ .
- (ii) Suppose  $B_1, B_2, B_3, \dots \in \mathcal{G}$ . Then for each  $i$ ,  $B_i = A_i \cap B$  for some  $A_i \in \mathcal{F}$ . Since  $\mathcal{F}$  is a  $\sigma$ -algebra then  $\bigcup_i A_i \in \mathcal{F}$ . Thus  $\bigcup_i B_i = \bigcup_i (A_i \cap B) = \{x : \forall i, (x \in A_i) \wedge (x \in B)\} = \{x : (\forall i, x \in A_i) \wedge (x \in B)\} = (\bigcup_i A_i) \cap B \in \mathcal{G}$ .
- (iii) Suppose  $B_0 \in \mathcal{G}$ . Then  $B_0 = A_0 \cap B$  for some  $A_0 \in \mathcal{F}$ . Consider the compliment of  $B_0$  in  $B$ . We have,  $B_0^c = (A_0 \cap B)^c = A_0^c \cup B^c = \{x \in B : x \in A_0^c \cup B^c\} = \{x \in B : (x \in A_0^c) \vee (x \in B^c)\} = \{x \in B : (x \in A_0^c)\} = \{x \in B : (x \in A_0^c) \wedge (x \in B)\} = A_0^c \cap B$ . Since  $\mathcal{F}$  is a  $\sigma$ -algebra then  $A_0^c \in \mathcal{F}$  so  $B_0^c = A_0^c \cap B \in \mathcal{G}$ .

This proves  $\mathcal{G}$  is a  $\sigma$ -algebra. □

**Exercise 1.2**

Let  $\mathcal{F}$  and  $\mathcal{G}$  be  $\sigma$  algebras of  $\Omega$ .

- (a) Show that  $\mathcal{F} \cap \mathcal{G}$  is a  $\sigma$ -algebra of  $\Omega$ .
- (b) Show that  $\mathcal{F} \cup \mathcal{G}$  is not necessarily a  $\sigma$ -algebra of  $\Omega$ .

**Solution**

Let  $\mathcal{F}$  and  $\mathcal{G}$  be  $\sigma$  algebras of  $\Omega$ .

- (a) Consider  $\mathcal{F} \cap \mathcal{G}$ .
  - (i) Since each  $\mathcal{F}$  and  $\mathcal{G}$  are  $\sigma$ -algebras,  $\emptyset \in \mathcal{F}$  and  $\emptyset \in \mathcal{G}$ . Therefore  $\emptyset \in \mathcal{F} \cap \mathcal{G}$ .
  - (ii) Suppose  $A_1, A_2, A_3, \dots \in \mathcal{F} \cap \mathcal{G}$ . Then  $A_1, A_2, A_3, \dots \in \mathcal{F}$  and  $A_1, A_2, A_3, \dots \in \mathcal{G}$ . These are each  $\sigma$ -algebras, so  $\bigcup_i A_i \in \mathcal{F}$  and  $\bigcup_i A_i \in \mathcal{G}$ . Therefore  $\bigcup_i A_i \in \mathcal{F} \cap \mathcal{G}$ .
  - (iii) Suppose  $A \in \mathcal{F} \cap \mathcal{G}$ . Then  $A \in \mathcal{F}$  and  $A \in \mathcal{G}$ . These are each  $\sigma$ -algebras, so  $A^c \in \mathcal{F}$  and  $A^c \in \mathcal{G}$ . Therefore  $A^c \in \mathcal{F} \cap \mathcal{G}$ .

This proves  $\mathcal{F} \cap \mathcal{G}$  is a  $\sigma$ -algebra. □

- (b) Let  $\Omega = (0, 3)$ ,  $A = (0, 1)$ ,  $B = (2, 3)$ . Define,

$$\mathcal{F} = \{\emptyset, A, A^c, \Omega\} = \{\emptyset, (0, 1), [1, 3), (0, 3)\} \quad \mathcal{G} = \{\emptyset, B, B^c, \Omega\} = \{\emptyset, (2, 3), (0, 2], (0, 3)\}$$

Then,

$$\mathcal{F} \cup \mathcal{G} = \{\emptyset, A, B, A^c, B^c, \Omega\} = \{\emptyset, (0, 1), (2, 3), [1, 3), (0, 2], (0, 3)\}$$

Observe that  $A \cup B = (0, 1) \cup (2, 3) \notin \mathcal{F} \cup \mathcal{G}$ .

This proves that for  $\sigma$ -algebras  $\mathcal{F}, \mathcal{G}$  of  $\Omega$ , their union  $\mathcal{F} \cup \mathcal{G}$  is not necessarily a  $\sigma$ -algebra of  $\Omega$ . □

**Exercise 1.3**

Describe the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  for the following three experiments:

- (a) a biased coin is tossed three times
- (b) two balls are drawn without replacement from an urn which originally contained two blue and two red balls
- (c) a biased coin is tossed repeatedly until a head turns up

**Solution**

- (a)  $\Omega = \{w_1 w_2 w_3 : w_i \in \{H, T\}\} = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}$ .

$\mathcal{F}$  contains sets which correspond to to questions we could ask about the three coin tosses. For instance, “was the second coin toss a heads?” ( $\{HHH, HHT, THH, THT\}$ ), “were there two heads?” ( $\{HHT, HTH, THH\}$ ), “were the final two coin tosses tails?” ( $\{HTT, TTT\}$ ), etc.

Explicitly,  $\mathcal{F} = \mathcal{P}(\Omega)$ , the power set of  $\Omega$ .

Suppose the coin is biased such that the probability of heads is  $p$  (and the probability of tails is  $1 - p$ ).

Define  $g(h_1 h_2 h_3) = p^k (1 - p)^{3-k}$ , where  $k$  is the number of  $h_1, h_2, h_3$  which are heads.

Define  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  as  $\mathbb{P}(A) = \sum_{d \in A} g(d)$ .

As a quick test, let  $A = \{HHH, HHT\}$ , the set asking the question “were the first two coin tosses heads?”. We have  $\mathbb{P}(A) = g(HHH) + g(HHT) = ppp + pp(1 - p) = p^2$  which corresponds to our natural understanding of the answer to this question.

Suppose  $A, B \subseteq \Omega$  are disjoint. Then,

$$\mathbb{P}(A \cup B) = \sum_{d \in A \cup B} g(d) = \sum_{d \in A} g(d) + \sum_{d \in B} g(d) = \mathbb{P}(A) + \mathbb{P}(B)$$

Clearly  $\mathbb{P}(\emptyset) = 0$  and  $\mathbb{P}(\Omega) = 1$ . Therefore  $\mathbb{P}$  is a well defined probability measure on  $(\Omega, \mathcal{F})$ .

- (b)  $\Omega = \{b_1 b_2 : b_i \in \{R, B\}\} = \{RR, RB, BR, BB\}$

$\mathcal{F}$  contain sets which correspond to questions we could ask about the color of the two balls. For instance, “was the first ball red?” ( $\{RB, RR\}$ ), “was the second ball blue?” ( $\{RB, BB\}$ ), “was the first ball red and the second ball blue?” ( $\{RB\}$ ), etc.

Explicitly,  $\mathcal{F} = \mathcal{P}(\Omega)$ , the power set of  $\Omega$ .

$$\text{Define } g(b_1 b_2) = \begin{cases} 1/6 & b_1 = b_2 \\ 1/3 & b_1 \neq b_2 \end{cases}$$

Then define  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  as  $\mathbb{P}(A) = \sum_{d \in A} g(d)$ .

As a quick test, let  $A = \{RB, RR\}$ , the set asking the question “was the first ball red?”. We have  $\mathbb{P}(A) = g(RB) + g(RR) = 1/3 + 1/6 = 1/2$  which corresponds to our natural understanding of the answer to this question.

Suppose  $A, B \subseteq \Omega$  are disjoint.

$$\mathbb{P}(A \cup B) = \sum_{d \in A \cup B} g(d) = \sum_{d \in A} g(d) + \sum_{d \in B} g(d) = \mathbb{P}(A) + \mathbb{P}(B)$$

Clearly  $\mathbb{P}(\emptyset) = 0$  and  $\mathbb{P}(\Omega) = g(RR) + g(RB) + g(BR) + g(BB) = 1/6 + 1/3 + 1/3 + 1/6 = 1$ . Therefore  $\mathbb{P}$  is a well defined probability measure on  $(\Omega, \mathcal{F})$ .

(c)  $\Omega = \{w_1 w_2 \dots w_{n-1} w_n : w_1, \dots, w_{n-1} = T, w_n = H\}$

Again,  $\mathcal{F}$  contains sets which correspond to questions we could ask about the coin tosses. For instance, “was the first toss a tails?” ( $\{TH, TTH, TTTH, \dots\}$ ), “was the last toss a heads?” ( $\{H, TH, TTH, TTTH, \dots\} = \Omega$ ), etc.

Explicitly,  $\mathcal{F} = \mathcal{P}(\Omega)$ , the power set of  $\Omega$ .

Suppose the coin is biased such that the probability of heads is  $(1 - q)$  (and the probability of tails is  $q$ ).

Define  $g(w_1 w_2 \dots w_n) = q^{n-1}(1 - q)$ , where  $g() = 0$ .

Define  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  as  $\mathbb{P}(A) = \sum_{d \in A} g(d)$ .

As a quick test, let  $A = \{TTTTH, TTH\}$ . Then  $\mathbb{P}(A) = g(TTTTH) + g(TTH) = q^4(1 - q) + q^2(1 - q)$  as expected.

Suppose  $A, B \subseteq \Omega$  are disjoint. Then, since  $A, B$  are countable,

$$\mathbb{P}(A \cup B) = \sum_{d \in A \cup B} g(d) = \sum_{d \in A} g(d) + \sum_{d \in B} g(d) = \mathbb{P}(A) + \mathbb{P}(B)$$

By definition  $\mathbb{P}(\emptyset) = 0$ . Finally  $\mathbb{P}(\Omega) = \sum_{n=1}^{\infty} q^{n-1}(1 - q)$  which we verify is equal to 1 using Mathematica (`Sum[p^(n-1) (1-p), {n, 1, Infinity}]/FullSimplify`). Therefore  $\mathbb{P}$  is a well defined probability measure on  $(\Omega, \mathcal{F})$ .

**Exercise 1.4**

Suppose  $X$  is a continuous random variable with distribution  $F_X$ . Let  $g$  be a strictly increasing continuous function. Define  $Y = g(X)$ .

- (a) What is  $F_Y$ , the distribution of  $Y$ ?
- (b) What is  $f_Y$ , the density of  $Y$ ?

**Solution**

Suppose  $X$  is a continuous random variable with distribution  $F_X$ . Let  $g : D \rightarrow \mathbb{R}$  be a strictly increasing continuous function.

- (a) Since  $g$  is strictly increasing and continuous, then it has an inverse  $g^{-1}$  on its range  $R := g(D)$ . Since  $g$  is strictly increasing, by the intermediate value theorem, we know  $R$  is an interval. Let  $a = \inf R$  and  $b = \sup R$ . Then, for  $y \leq a$  we have  $F_Y(y) = 0$  and for  $y \geq b$  we have  $F_Y(y) = 1$ . For  $y \in R$ , assuming  $g^{-1}$  is differentiable, we have,

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(g(X) \leq y) = \mathbb{P}(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$$

Thus,

$$F_Y(y) = \begin{cases} 1 & y \geq b \\ F_X(g^{-1}(y)) & a < y < b \\ 0 & y \leq a \end{cases}$$

Since  $\{a, b\}$  is a set of measure zero it doesn't really matter how we define  $F_Y$  on these points.

- (b) Recall  $F_Y(y) = \int_{-\infty}^y f_Y(u) du$  for a continuous random variable  $Y$ . Thus for  $a < y < b$ , observing that  $dF_X(y)/dy = f_X(y)$  by the above result, and assuming  $g^{-1}$  is differentiable,

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(g^{-1}(y)) = f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

For  $y < a$  and  $a > b$  we have  $F_Y(y)$  constant so  $f_Y(y)$  is zero.

Thus,

$$f_Y(y) = \begin{cases} 0 & y > b \\ f_X(g^{-1}(y)) \frac{d}{dx} g^{-1}(y) & a < y < b \\ 0 & y < a \end{cases}$$

Since  $\{a, b\}$  is a set of measure zero it doesn't really matter how we define  $F_Y$  on these points.

**Exercise 1.5**

Suppose  $X$  is a continuous random variable with distribution  $F_X$ . Find  $F_Y$  where  $Y$  is given by:

- (a)  $X^2$
- (b)  $\sqrt{|X|}$
- (c)  $\sin(X)$
- (d)  $F_X(X)$

**Solution**

Since  $X$  is continuous we can write  $F_X(x) = \mathbb{P}(X \leq x) = \int_{-\infty}^x f_X(u)du$  for some density  $f_X : \mathbb{R} \rightarrow [0, \infty)$ .

In general,  $\mathbb{P}(a < X < b) = \int_a^b f_X(u)du = F_X(b) - F_X(a)$  by the FTC.

- (a) Let  $Y = X^2$ . We then have,

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(X^2 \leq y)$$

If  $y \leq 0$  then  $F_Y(y) = \mathbb{P}(X^2 \leq y \leq 0) = 0$  since  $X^2 = 0$  on a set of measure 0).

If  $y > 0$  then,

$$F_Y(y) = \mathbb{P}(X^2 \leq y) = \mathbb{P}(|X| \leq \sqrt{y}) = \mathbb{P}(-\sqrt{y} \leq X \leq \sqrt{y}) = F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

Finally,

$$F_Y(y) = \begin{cases} 0 & y \leq 0 \\ F_X(\sqrt{y}) - F_X(-\sqrt{y}) & y > 0 \end{cases}$$

- (b) Let  $Y = \sqrt{|X|}$ . We then have,

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(\sqrt{|X|} \leq y)$$

If  $y \leq 0$  then  $F_Y(y) = \mathbb{P}(\sqrt{|X|} \leq y \leq 0) = 0$  (since  $\sqrt{|X|} = 0$  on a set of measure 0).

If  $y > 0$

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(\sqrt{|X|} \leq y) = \mathbb{P}(|X| \leq y^2) = \mathbb{P}(-y^2 \leq X \leq y^2) = F_X(y^2) - F_X(-y^2)$$

Finally,

$$F_Y(y) = \begin{cases} 0 & y \leq 0 \\ F_X(y^2) - F_X(-y^2) & y > 0 \end{cases}$$

- (c) Let  $Y = \sin(X)$ . We then have,

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(\sin(X) \leq y)$$

If  $y \leq -1$  then  $F_Y(y) = \mathbb{P}(\sin(X) \leq y \leq 0) = 0$  since  $\sin(X) = 1$  on a set of measure 0.

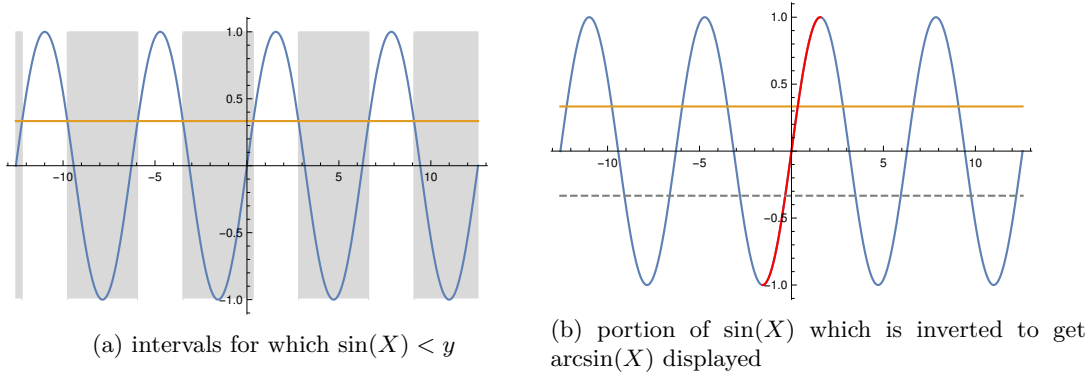


Figure 1: Exercise 1.5(c)

Similarly, if  $y \geq 1$  then  $F_Y(y) = \mathbb{P}(\sin(X) \leq 1 \leq y) = 1$ .

Figure 1a shows the intervals for which  $\sin(X) < y$  is grey, for some  $y$  drawn in orange. Figure 1b shows which part of  $\sin$  the inverse is defined on in red. We see that we can find the intervals for which  $\sin(X) < y$  by finding the intersection of  $\sin(X)$  and  $y$  as well as the intersection of  $\sin(X)$  and  $-y$  and then appropriately shifting these endpoints.

If  $-1 < y < 1$  then,  $\sin(X) \leq y$  if and only if for some integer  $k$ ,

$$\arcsin(y) + 2k\pi < X < \pi - \arcsin(y) + 2k\pi = \pi - \arcsin(y) + 2k\pi$$

Thus, since  $\mathbb{Z}$  is countable,

$$\begin{aligned} F_Y(y) &= \mathbb{P}(\arcsin(y) + 2k\pi < X < \pi - \arcsin(y) + 2k\pi, \text{ for any } k \in \mathbb{Z}) \\ &= \mathbb{P}\left(x \in \bigcup_{k \in \mathbb{Z}} (\arcsin(y) + 2k\pi < X < \pi - \arcsin(y) + 2k\pi)\right) \\ &= \sum_{k \in \mathbb{Z}} \mathbb{P}(\arcsin(y) + 2k\pi < X < \pi - \arcsin(y) + 2k\pi) \\ &= \sum_{k \in \mathbb{Z}} [F_X(\arcsin(y) + 2k\pi) - F_X(\pi - \arcsin(y) + 2k\pi)] \end{aligned}$$

Therefore,

$$F_Y(y) = \begin{cases} 0 & y \leq -1 \\ \sum_{k \in \mathbb{Z}} [F_X(\arcsin(y) + 2k\pi) - F_X(\pi - \arcsin(y) + 2k\pi)] & -1 < y < 1 \\ 1 & y \geq 1 \end{cases}$$

(d) Let  $Y = F_X(X)$ . We then have,

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(F_X(X) \leq y)$$

Recall  $F_X$  is a (not necessarily strictly) increasing function from  $\mathbb{R}$  to  $[0, 1]$ . We deal with this in the following way: Define  $g : \mathbb{R} \rightarrow \mathbb{R}$  by  $g(y) = \sup\{a : F_X(a) \leq y\}$ . If  $F_X$  is strictly increasing then  $g(y) = y$  for all  $y$  as desired. Now define  $\hat{F}_X : \mathbb{R} \rightarrow [0, 1]$  as  $\hat{F}_X(y) = F_X(g(y))$ . This function is strictly increasing and injective so therefore invertible. Denote the inverse by  $F_X^{-1}$ .

Recall  $F_X$  goes from 0 to 1. If  $y < 0$  then  $F_Y(y) = \mathbb{P}(F_X(X) \leq y < 0) = 0$ . If  $y > 1$  then  $F_Y(y) = \mathbb{P}(F_X(X) \leq 1 < y) = 1$ .



For  $0 < y < a$ , we have,

$$F_Y(y) = \mathbb{P}(F_X(X) \leq y) = \mathbb{P}(X \leq \hat{F}_X^{-1}(y)) = F_X(\hat{F}_X^{-1}(y)) = y$$

Thus,

$$F_Y(y) = \begin{cases} 0 & y < 0 \\ y & 0 < y < 1 \\ 1 & y > 1 \end{cases}$$

Again since the points  $\{0, 1\}$  is a set of measure zero it doesn't really matter how we define  $F_Y$  on these points.

**Exercise 1.6**

Suppose  $X$  is a continuous random variable defined on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . Let  $f$  be the density of  $X$  under  $\mathbb{P}$  and assume  $f > 0$ . Let  $g$  be the density function of a random variable. Define  $Z := g(X)/f(X)$ .

- (a) Show that  $Z \equiv d\tilde{\mathbb{P}}/d\mathbb{P}$  defines a Radon-Nikodym derivative.
- (b) What is the density of  $X$  under  $\tilde{\mathbb{P}}$ ?

**Solution**

Suppose  $X$  is a continuous random variable defined on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . Let  $f$  be the density of  $X$  under  $\mathbb{P}$  and assume  $f > 0$ . Let  $g$  be the density function of a random variable. Define  $Z := g(X)/f(X)$ .

- (a) Since  $g$  is a density function,  $g \geq 0$ . Thus,  $Z = g(X)/f(X) \geq 0$ . Moreover, since  $g$  is a density,  $\int_{\Omega} g(x)dx = 1$ . Thus,

$$\mathbb{E}Z = \int_{\Omega} \frac{g(x)}{f(x)} f(x) dx = \int_{\Omega} g(x) dx = 1$$

Then define  $\tilde{\mathbb{P}}(A) = \mathbb{E}Z \mathbb{1}_A$ .

Then  $Z$  is a Radon-Nikodym derivative of  $\tilde{\mathbb{P}}$  with respect to  $\mathbb{P}$ .

- (b) We first compute the distribution of  $X$  under  $\tilde{\mathbb{P}}$  assuming  $\Omega = \mathbb{R}$ .

$$F_X(y) = \tilde{\mathbb{P}}(X \leq y) = \mathbb{E}Z \mathbb{1}_{\{X \leq y\}} = \int_{\Omega} \frac{g(x)}{f(x)} \mathbb{1}_{\{X \leq y\}} f(x) dx = \int_{-\infty}^y g(x) dx$$

Thus,

$$f_X(y) = \frac{d}{dy} \int_{-\infty}^y g(x) dx = g(y)$$