

# **AMATH 561** Assignment 2

Tyler Chen

**Exercise 2.1**

Let  $\Omega = \{a, b, c, d\}$  and  $\mathcal{F} = 2^\Omega$  (the set of all subsets of  $\Omega$ ). We define a probability measure  $\mathbb{P}$  as follows

$$\mathbb{P}(a) = 1/6, \quad \mathbb{P}(b) = 1/3, \quad \mathbb{P}(c) = 1/4, \quad \mathbb{P}(d) = 1/4$$

Next, define three random variables,

$$\begin{array}{llll} X(a) = 1, & X(b) = 1, & X(c) = -1, & X(d) = -1 \\ Y(a) = 1, & Y(b) = -1, & Y(c) = 1, & Y(d) = -1, \end{array}$$

and  $Z = X + Y$ .

- List the sets in  $\sigma(X)$ .
- What are the values of  $\mathbb{E}[Y|X]$  for  $\{a, b, c, d\}$ ? Verify the partial averaging property:  $\mathbb{E}[\mathbb{1}_A \mathbb{E}[Y|X]] = \mathbb{E}[\mathbb{1}_A Y]$ .
- What are the values of  $\mathbb{E}[Z|X]$  for  $\{a, b, c, d\}$ ? Verify the partial averaging property.

**Solution**

- Recall that  $\sigma(X) = \{\{X \in A\} \subseteq \Omega : A \in \mathcal{B}(\mathbb{R})\} = \{\{w : X(w) \in A\} : A \in \mathcal{B}(\mathbb{R})\}$ . Therefore,

$$\sigma(X) = \{\emptyset, \{a, b\}, \{c, d\}, \Omega\}$$

- We compute,

$$\begin{aligned} \mathbb{E}[Y|X](a) &= \mathbb{E}[Y|X = X(a)] = \mathbb{E}[Y|X = 1] = \frac{1\mathbb{P}(a) - 1\mathbb{P}(b)}{\mathbb{P}(a) + \mathbb{P}(b)} = -\frac{1}{3} \\ \mathbb{E}[Y|X](b) &= \mathbb{E}[Y|X = X(b)] = \mathbb{E}[Y|X = 1] = \frac{1\mathbb{P}(a) - 1\mathbb{P}(b)}{\mathbb{P}(a) + \mathbb{P}(b)} = -\frac{1}{3} \\ \mathbb{E}[Y|X](c) &= \mathbb{E}[Y|X = X(c)] = \mathbb{E}[Y|X = -1] = \frac{1\mathbb{P}(c) - 1\mathbb{P}(d)}{\mathbb{P}(c) + \mathbb{P}(d)} = 0 \\ \mathbb{E}[Y|X](d) &= \mathbb{E}[Y|X = X(d)] = \mathbb{E}[Y|X = -1] = \frac{1\mathbb{P}(c) - 1\mathbb{P}(d)}{\mathbb{P}(c) + \mathbb{P}(d)} = 0 \end{aligned}$$

For each set  $A \in \sigma(X)$  we verify that  $\mathbb{E}[\mathbb{1}_A \mathbb{E}[Y|X]] = \mathbb{E}[\mathbb{1}_A Y]$  as follows,

$A$	$\mathbb{E}[\mathbb{1}_A \mathbb{E}[Y X]]$	$\mathbb{E}[\mathbb{1}_A Y]$
$\emptyset$	0	0
$\{a, b\}$	$-\frac{1}{3}\mathbb{P}(a) - \frac{1}{3}\mathbb{P}(b) = -\frac{1}{6}$	$1\mathbb{P}(a) - 1\mathbb{P}(b) = -\frac{1}{6}$
$\{c, d\}$	$0\mathbb{P}(c) + 0\mathbb{P}(d) = 0$	$1\mathbb{P}(c) - 1\mathbb{P}(d) = 0$
$\Omega$	$-\frac{1}{3}\mathbb{P}(a) - \frac{1}{3}\mathbb{P}(b) + 0\mathbb{P}(c) + 0\mathbb{P}(d) = -\frac{1}{6}$	$1\mathbb{P}(a) - 1\mathbb{P}(b) + 1\mathbb{P}(c) - 1\mathbb{P}(d) = -\frac{1}{6}$

- Write,

$$Z(a) = 2, \quad Z(b) = 0, \quad Z(c) = 0, \quad Z(d) = -2$$

We compute,

$$\begin{aligned}\mathbb{E}[Z|X](a) &= \mathbb{E}[Z|X = X(a)] = \mathbb{E}[Z|X = 1] = \frac{2\mathbb{P}(a) + 0\mathbb{P}(b)}{\mathbb{P}(a) + \mathbb{P}(b)} = \frac{2}{3} \\ \mathbb{E}[Z|X](b) &= \mathbb{E}[Z|X = X(b)] = \mathbb{E}[Z|X = 1] = \frac{2\mathbb{P}(a) + 0\mathbb{P}(b)}{\mathbb{P}(a) + \mathbb{P}(b)} = \frac{2}{3} \\ \mathbb{E}[Z|X](c) &= \mathbb{E}[Z|X = X(c)] = \mathbb{E}[Z|X = -1] = \frac{0\mathbb{P}(c) + 2\mathbb{P}(d)}{\mathbb{P}(c) + \mathbb{P}(d)} = -1 \\ \mathbb{E}[Z|X](d) &= \mathbb{E}[Z|X = X(d)] = \mathbb{E}[Z|X = -1] = \frac{0\mathbb{P}(c) + 2\mathbb{P}(d)}{\mathbb{P}(c) + \mathbb{P}(d)} = -1\end{aligned}$$

For each set  $A \in \sigma(X)$  we verify that  $\mathbb{E}[\mathbb{1}_A \mathbb{E}[Z|X]] = \mathbb{E}[\mathbb{1}_A Z]$  as follows,

$A$	$\mathbb{E}[\mathbb{1}_A \mathbb{E}[Z X]]$	$\mathbb{E}[\mathbb{1}_A Z]$
$\emptyset$	0	0
$\{a, b\}$	$\frac{2}{3}\mathbb{P}(a) + \frac{2}{3}\mathbb{P}(b) = \frac{1}{3}$	$2\mathbb{P}(a) + 0\mathbb{P}(b) = \frac{1}{3}$
$\{c, d\}$	$-1\mathbb{P}(c) + -1\mathbb{P}(d) = -\frac{1}{2}$	$0\mathbb{P}(c) - 2\mathbb{P}(d) = -\frac{1}{2}$
$\Omega$	$\frac{2}{3}\mathbb{P}(a) + \frac{2}{3}\mathbb{P}(b) - 1\mathbb{P}(c) - 1\mathbb{P}(d) = -\frac{1}{6}$	$2\mathbb{P}(a) + 0\mathbb{P}(b) + 0\mathbb{P}(c) - 2\mathbb{P}(d) = -\frac{1}{6}$

**Exercise 2.2**

Fix a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . Let  $Y$  be a square integrable random variable:  $\mathbb{E}Y^2 < \infty$  and let  $\mathcal{G}$  be a sub- $\sigma$ -algebra of  $\mathcal{F}$ . Show that

$$\mathbb{V}(Y - \mathbb{E}[Y|\mathcal{G}]) \leq \mathbb{V}(Y - X) \quad \forall X \in \mathcal{G}$$

**Solution**

Suppose further  $\mathbb{E}[(Y - X)^2] < \infty$  (we make this assumption so that  $\mathbb{V}[Y - X]$  exists).

Clearly  $(\mathbb{E}[Y|\mathcal{G}] - X) \in \mathcal{G}$ . Then, by partial averaging,  $\mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)\mathbb{E}[Y|\mathcal{G}]] = \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)Y]$ . Therefore,

$$\begin{aligned} \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)(Y - \mathbb{E}[Y|\mathcal{G}])] &= \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)Y] - \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)\mathbb{E}[Y|\mathcal{G}]] \\ &= \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)Y] - \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)Y] \\ &= 0 \end{aligned}$$

Then, since  $\mathbb{E}[(Y - X)^2]$ , exists,

$$\begin{aligned} \mathbb{E}[(Y - X)^2] &= \mathbb{E}[(Y - \mathbb{E}[Y|\mathcal{G}]) + (\mathbb{E}[Y|\mathcal{G}] - X)^2] \\ &= \mathbb{E}[(Y - \mathbb{E}[Y|\mathcal{G}])^2] + \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)^2] + 2\mathbb{E}[(Y - \mathbb{E}[Y|\mathcal{G}])(\mathbb{E}[Y|\mathcal{G}] - X)] \\ &= \mathbb{E}[(Y - \mathbb{E}[Y|\mathcal{G}])^2] + \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)^2] \end{aligned}$$

Again by partial averaging,  $\mathbb{E}[\mathbb{E}[Y|\mathcal{G}]] = \mathbb{E}[\mathbb{1}_\Omega \mathbb{E}[Y|\mathcal{G}]] = \mathbb{E}[\mathbb{1}_\Omega Y] = \mathbb{E}[Y]$  so that  $\mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]] = 0$ . Then,

$$\begin{aligned} \mathbb{E}[Y - X]^2 &= \mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}] + \mathbb{E}[Y|\mathcal{G}] - X]^2 \\ &= (\mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]] + \mathbb{E}[\mathbb{E}[Y|\mathcal{G}] - X])^2 \\ &= \mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]]^2 + \mathbb{E}[\mathbb{E}[Y|\mathcal{G}] - X]^2 + 2\mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]]\mathbb{E}[\mathbb{E}[Y|\mathcal{G}] - X] \\ &= \mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]]^2 + \mathbb{E}[\mathbb{E}[Y|\mathcal{G}] - X]^2 \end{aligned}$$

Thus, subtracting this result from the first,

$$\begin{aligned} \mathbb{E}[(Y - X)^2] - \mathbb{E}[Y - X]^2 &= \mathbb{E}[(Y - \mathbb{E}[Y|\mathcal{G}])^2] - \mathbb{E}[Y - \mathbb{E}[Y|\mathcal{G}]]^2 + \mathbb{E}[(\mathbb{E}[Y|\mathcal{G}] - X)^2] - \mathbb{E}[\mathbb{E}[Y|\mathcal{G}] - X]^2 \\ \mathbb{V}[Y - X] &= \mathbb{V}[Y - \mathbb{E}[Y|\mathcal{G}]] + \mathbb{V}[\mathbb{E}[Y|\mathcal{G}] - X] \end{aligned}$$

Therefore, since  $\mathbb{V}[\mathbb{E}[Y|\mathcal{G}] - X] \geq 0$ , for any  $X \in \mathcal{G}$ ,

$$\mathbb{V}[Y - X] \geq \mathbb{V}[Y - \mathbb{E}[Y|\mathcal{G}]] \quad \square$$

Exercise 2.3 Give an example of a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , a random variable  $X$  and a function  $f$  such that  $\sigma(f(X))$  is strictly smaller than  $\sigma(X)$  but  $\sigma(f(X)) \neq \{\emptyset, \Omega\}$ . Give a function  $g$  such that  $\sigma(g(X)) = \{\emptyset, \Omega\}$ .

Let  $\Omega = \{a, b, c\}$  and  $\mathcal{F} = 2^\Omega$ . Define  $\mathbb{P}(a) = \mathbb{P}(b) = \mathbb{P}(c) = 1/3$ .

Define  $X$  as  $X(a) = 0$ ,  $X(b) = -1$ ,  $X(c) = 1$ .

Thus,  $\sigma(X) = \{\emptyset, \{a\}, \{b\}, \{c\}, \{b, c\}, \{a, c\}, \{a, b\}, \Omega\}$

Since  $X(\Omega) \subset \mathbb{R}$ , define  $f, g : \mathbb{R} \rightarrow \mathbb{R}$  by  $f(x) = x^2$  and  $g(x) = 0$ . Then  $f(X(a)) = 0$ ,  $f(X(b)) = f(X(c)) = 1$  and  $g(X(a)) = g(X(b)) = g(X(c)) = 0$ .

Therefore  $\sigma(f(X)) = \{\emptyset, \{a\}, \{b, c\}, \Omega\}$  so  $\sigma(f(X)) \subsetneq \sigma(X)$  so  $\sigma(f(X))$  is strictly smaller than  $\sigma(X)$ .

Similarly,  $\sigma(g(X)) = \{\emptyset, \Omega\}$ .

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**Exercise 2.4**

On a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  define random variables  $X$  and  $Y_0, Y_1, Y_2, \dots$  and suppose  $\mathbb{E}[X] < \infty$ . Define  $F_n := \sigma(Y_0, Y_1, \dots, Y_n)$  and  $X_n = \mathbb{E}[X|F_n]$ . Show that the sequence  $X_0, X_1, X_2, \dots$  is a martingale under  $\mathbb{P}$  with respect to the filtration  $(F_n)_{n \geq 0}$ .

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**Solution**

Since  $(F_n)_{n \geq 0}$  is a filtration, then  $F_n$  is a sub  $\sigma$ -algebra of  $F_{n+1}$ . Therefore, by iterated conditioning,

$$\mathbb{E}[X_{n+1}|F_n] = \mathbb{E}[\mathbb{E}[X|F_{n+1}]|F_n] = \mathbb{E}[X|F_n] = X_n$$

This proves the sequence  $X_0, X_1, X_2, \dots$  is a martingale under  $\mathbb{P}$  with respect to the filtration  $(F_n)_{n \geq 0}$ .  
 $\square$

**Exercise 2.5**

Let  $X_0, X_1, \dots$  be i.i.d Bernoulli random variables with parameter  $p$  (i.e.,  $P(X_i = 1) = p$ ). Define  $S_n = \sum_{i=1}^n X_i$  where  $S_0 = 0$ . Define

$$Z_n := \left( \frac{1-p}{p} \right)^{2S_n - n} \quad n = 0, 1, 2, \dots$$

Let  $\mathcal{F}_n := \sigma(X_0, X_1, \dots, X_n)$ . Show that  $Z_n$  is a martingale with respect to this filtration.

**Solution**

Observe,

$$Z_{n+1} = \left( \frac{1-p}{p} \right)^{2S_{n+1} - (n+1)} = \left( \frac{1-p}{p} \right)^{2S_n - n} \left( \frac{1-p}{p} \right)^{2X_{n+1} - 1} = Z_n \left( \frac{1-p}{p} \right)^{2X_{n+1} - 1}$$

Then, since  $X_{n+1}$  is independent of all other  $X_j$ ,  $X_{n+1}$  is independent of  $F_n$ . Thus, using the definition of expectation of a discrete random variable,

$$\begin{aligned} \mathbb{E} \left[ \left( \frac{1-p}{p} \right)^{2X_{n+1} - 1} \middle| F_n \right] &= \mathbb{E} \left[ \left( \frac{1-p}{p} \right)^{2X_{n+1} - 1} \right] \\ &= (p) \left( \frac{1-p}{p} \right)^{2 \cdot 1 - 1} + (1-p) \left( \frac{1-p}{p} \right)^{2 \cdot 0 - 1} \\ &= (1-p) + p \\ &= 1 \end{aligned}$$

Therefore, by taking out what is known,

$$\mathbb{E}[Z_{n+1} | F_n] = \mathbb{E} [Z_n ((1-p)/p)^{2X_{n+1} - 1} | F_n] = Z_n \mathbb{E} [((1-p)/p)^{2X_{n+1} - 1} | F_n] = Z_n$$

This proves  $(Z_n)_{n \geq 0}$  is a martingale with respect to this filtration. □