

Bayesian multimodeling: Variational inference-2

2024

Local variational optimization, idea

Consider a problem of approximation of $f(x) = \exp(-x)$ by linear function.

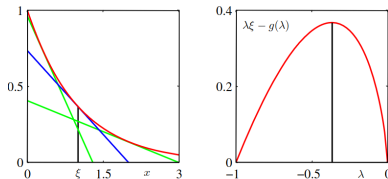
Any linear function will be a lower bound on $f(x)$ if it corresponds to a tangent. Use Taylor series:

$$y(x) = f(x_0) + f'(x_0)(x - x_0)$$

or, using $\lambda = -f(x_0)$

$$y(x) = \lambda x - \lambda + \lambda \log(-\lambda).$$

Where is the variational optimization here?



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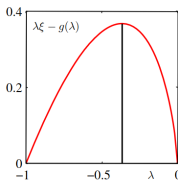
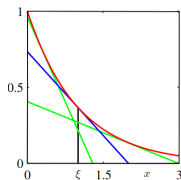
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or, using $\lambda = -f(x_0)$

$$y(x) = \lambda x - \lambda + \lambda \log(-\lambda).$$

We must find the tightest bound:

$$\max_{\lambda} \lambda x - \lambda + \lambda \log(-\lambda)$$



Local variational optimization and Evidence

Using similar approach we can approximate more interesting functions, for example sigmoid:

$$\log \sigma(x) = -\frac{x}{2} - \log(e^{\frac{x}{2}} + e^{\frac{-x}{2}}).$$

Note that $f(x) = -\log(e^{\frac{x}{2}} + e^{\frac{-x}{2}})$ is convex by x^2 .

Optimal value gives:

$$\sigma(x) \geq \sigma(x_0) \exp \left((x - x_0) / 2 - \lambda(x_0) (x_0^2 - x^2) \right).$$

The evidence integral becomes quadratic \Rightarrow we can use an approximation by Gaussian, similar to Laplace approximation.

Model selection: coherent Bayesian inference

First level: find optimal parameters:

$$\mathbf{w} = \arg \max \frac{p(\mathcal{D}|\mathbf{w})p(\mathbf{w}|\mathbf{h})}{p(\mathcal{D}|\mathbf{h})},$$

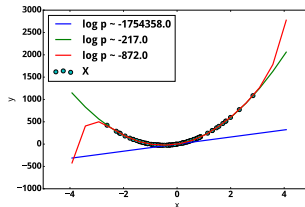
Second level: find optimal model:

Evidence:

$$p(\mathcal{D}|\mathbf{h}) = \int_{\mathbf{w}} p(\mathcal{D}|\mathbf{w})p(\mathbf{w}|\mathbf{h})d\mathbf{w}.$$



Model selection scheme



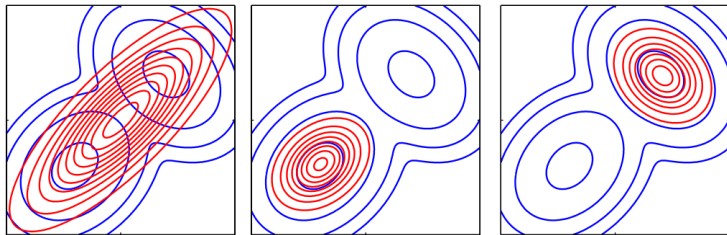
Polynomial regression example

Evidence lower bound, ELBO

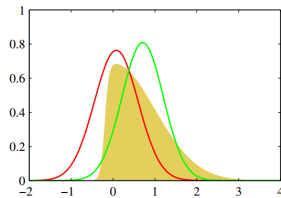
Evidence lower bound is a method of approximation of intractable distribution $p(\mathbf{w}|\mathcal{D}, \mathbf{h})$ with a distribution $q(\mathbf{w}) \in \mathcal{Q}$.

Evidence lower bound estimation often reduces to optimization problem

$$\begin{aligned} \log p(\mathcal{D}|\mathbf{h}) &\geq \\ &\geq - \int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathbf{w}|\mathcal{D})}{q(\mathbf{w})} d\mathbf{w} = \mathbb{E}_{\mathbf{w}} \log p(\mathcal{D}|\mathbf{w}) - \text{KL}(q(\mathbf{w})||p(\mathbf{w}|\mathbf{h})). \end{aligned}$$



Variational inference vs. expectation propagation (Bishop)



Laplace Approximation vs
Variational inference

ELBO estimation

ELBO maximization

$$\int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathbf{y}, \mathbf{w} | \mathbf{X}, \mathbf{h})}{q(\mathbf{w})} d\mathbf{w}$$

is equivalent to KL-divergence minimization between $q(\mathbf{w}) \in \mathfrak{Q}$ and posterior distribution $p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h})$:

$$\hat{q} = \arg \max_{q \in \mathfrak{Q}} \int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathbf{y}, \mathbf{w} | \mathbf{X}, \mathbf{h})}{q(\mathbf{w})} d\mathbf{w} \Leftrightarrow$$

$$\hat{q} = \arg \min_{q \in \mathfrak{Q}} D_{\text{KL}}(q(\mathbf{w}) || p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h})),$$

$$D_{\text{KL}}(q(\mathbf{w}) || p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h})) = \int_{\mathbf{w}} q(\mathbf{w}) \log \left(\frac{q(\mathbf{w})}{p(\mathbf{w} | \mathbf{y}, \mathbf{X}, \mathbf{h})} \right) d\mathbf{w}.$$

Outline, global variational methods

- Can we use something except Gaussian distribution?
 - ▶ Yes, we can
- Does it need to have an analytical form?
 - ▶ No
- Does it need to have some specific properties except continuity?
 - ▶ No
- Do we need a parametric distribution to approximate posterior?
 - ▶ No
- Is it always about Evidence approximation?
 - ▶ In general, no. We can use other probability distances.

Reparametrization trick

Reparamterization idea:

$$\varepsilon = S_{\theta}(\mathbf{w}), \quad \mathbf{w} = S_{\theta}^{-1}(\varepsilon).$$

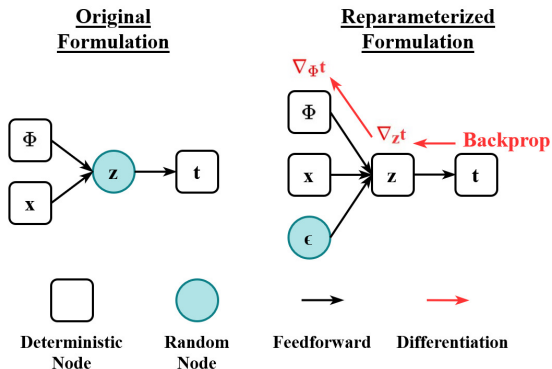
Then:

$$\nabla_{\theta} E_q f(\mathbf{w}) = E_q \nabla_{\theta} f(S_{\theta}^{-1}(\varepsilon)).$$

Example:

$$w \sim \mathcal{N}(\mu, \sigma^2) \rightarrow S(w) = \frac{w - \mu}{\sigma} \sim \mathcal{N}(0, 1).$$

Challenge: calculation of S^{-1} is an expensive operation.



Source: wikipedia

Normalizing Flows

Given an invertible smooth mapping \mathbf{g} (flow) and a distribution $\mathbf{z} \sim q$.
Then $q(\mathbf{g}(\mathbf{z}))$ is a distribution:

$$q(\mathbf{g}(\mathbf{z})) = q(\mathbf{z}) \left(\det \frac{\partial \mathbf{g}}{\partial \mathbf{z}} \right)^{-1}.$$

Example: planar flow:

$$\mathbf{g}(\mathbf{z}) = \mathbf{z} + \mathbf{w}_1 \sigma(\mathbf{w}_2^T \mathbf{x}).$$

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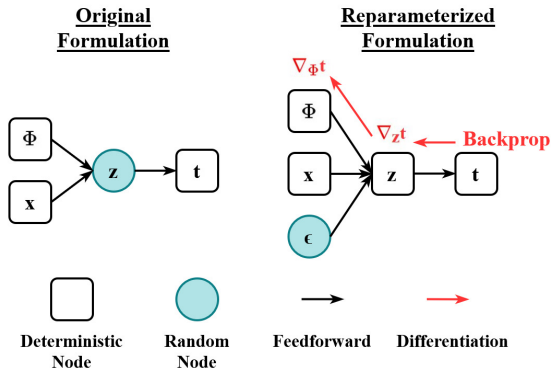
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Source: wikipedia

Implicit reparametrization trick

$$\nabla_{\theta} E_q f(\mathbf{w}) = E_q \nabla_{\mathbf{w}} f(\mathbf{w}) \nabla_{\theta} \mathbf{w}.$$

Use a total gradient formula for $\varepsilon = S_{\theta}(\mathbf{w})$:

$$\nabla_{\mathbf{w}} S_{\theta}(\mathbf{w}) \nabla_{\theta} \mathbf{w} + \nabla_{\theta} S_{\theta}(\mathbf{w}) = 0 \rightarrow$$

$$\rightarrow \nabla_{\theta} \mathbf{w} = -(\nabla_{\mathbf{w}} S_{\theta}(\mathbf{w}))^{-1} \nabla_{\theta} S_{\theta}.$$

Obtain an expression without inverse function for S .

For 1d samples we can use, for example:

$$S(\mathbf{w}) = F(\mathbf{w}|\theta) \sim \mathcal{U}(0, 1).$$

Table 4: Test negative log-likelihood (lower is better) for VAE on MNIST. Mean \pm standard deviation over 5 runs. The von Mises-Fisher results are from [9].

Prior	Variational posterior	$D = 2$	$D = 5$	$D = 10$	$D = 20$	$D = 40$
$\mathcal{N}(0, 1)$	$\mathcal{N}(\mu, \sigma^2)$	131.1 ± 0.6	107.9 ± 0.4	92.5 ± 0.2	88.1 ± 0.2	88.1 ± 0.0
Gamma(0.3, 0.3)	Gamma(α, β)	132.4 ± 0.3	108.0 ± 0.3	94.0 ± 0.3	90.3 ± 0.2	90.6 ± 0.2
Gamma(10, 10)	Gamma(α, β)	135.0 ± 0.2	107.0 ± 0.2	92.3 ± 0.2	88.3 ± 0.2	88.3 ± 0.1
Uniform(0, 1)	Beta(α, β)	128.3 ± 0.2	107.4 ± 0.2	94.1 ± 0.1	88.9 ± 0.1	88.6 ± 0.1
Beta(10, 10)	Beta(α, β)	131.1 ± 0.4	106.7 ± 0.1	92.1 ± 0.2	87.8 ± 0.1	87.7 ± 0.1
Uniform($-\pi, \pi$)	vonMises(μ, κ)	127.6 ± 0.4	107.5 ± 0.4	94.4 ± 0.5	90.9 ± 0.1	91.5 ± 0.4
vonMises(0, 10)	vonMises(μ, κ)	130.7 ± 0.8	107.5 ± 0.5	92.3 ± 0.2	87.8 ± 0.2	87.9 ± 0.3
Uniform(S^D)	vonMisesFisher($\boldsymbol{\mu}, \kappa$)	132.5 ± 0.7	108.4 ± 0.1	93.2 ± 0.1	89.0 ± 0.3	90.9 ± 0.3

MCMC and variational inference

MCMC idea: Sample from the simple distribution and accpet them, if the ratio is greater than some threshold:

$$\min \left(1, \frac{p(\mathbf{w}^\tau | \mathbf{y}, \mathbf{X}, \mathbf{h})}{p(\mathbf{w}^{\tau-1} | \mathbf{y}, \mathbf{X}, \mathbf{h})} \right),$$

where \mathbf{w}^τ is set based on the previous sample:

$$\mathbf{w}^\tau = T(\mathbf{w}^{\tau-1}).$$

Salimans et al., 2014: let's interperete the sequence of some operator T application as a variational optimization:

$$T^1 \circ \dots T^\eta(\mathbf{w}) \rightarrow p(\mathbf{w}^\tau | \mathbf{y}, \mathbf{X}, \mathbf{h}).$$

Maclaurin et. al, 2015: use gradient descent as such operator. Do not reject samples at all.

Optimization operator, Maclaurin et. al, 2015

Definition

Let T be an algorithm of changing model parameters \mathbf{w}' using previous parameter values \mathbf{w} :

$$\mathbf{w}' = T(\mathbf{w}).$$

Definition

Let L be a continuous loss function.

Define a gradient descent operator in the following way:

$$T(\mathbf{w}) = \mathbf{w} - \beta \nabla L(\mathbf{w}, \mathbf{y}, \mathcal{D}).$$

Gradient descent for evidence estimation

Consider posterior probability maximization:

$$L = -\log p(\mathcal{D}, \mathbf{w}|\mathbf{h}) = - \sum_{\mathcal{D} \in \mathcal{D}} \log p(\mathcal{D}|\mathbf{w}, \mathbf{h})p(\mathbf{w}|\mathbf{h})$$

Optimize neural network in a multi-start regime with r initial parameter values $\mathbf{w}_1, \dots, \mathbf{w}_r$ using (stochastic) gradient descent:

$$\mathbf{w}' = T(\mathbf{w}).$$

The parameter vectors $\mathbf{w}_1, \dots, \mathbf{w}_r$ are from some latent distribution $q(\mathbf{w})$.

Entropy

We can rewrite variational inference using differential entropy term:

$$\begin{aligned}\log p(\mathcal{D}|\mathbf{f}) &\geq \int_{\mathbf{w}} q(\mathbf{w}) \log \frac{p(\mathcal{D}, \mathbf{w}|\mathbf{h})}{q(\mathbf{w})} d\mathbf{w} = \\ &\quad \mathbb{E}_{q(\mathbf{w})}[\log p(\mathcal{D}, \mathbf{w}|\mathbf{h})] + S(q(\mathbf{w})),\end{aligned}$$

where $S(q(\mathbf{w}))$ is a differential entropy:

$$S(q(\mathbf{w})) = - \int_{\mathbf{w}} q(\mathbf{w}) \log q(\mathbf{w}) d\mathbf{w}.$$

Gradient descent for evidence estimation

Statement

Let L be a Lipschitz function, and optimization operator be a bijection. Then entropy difference for two steps is:

$$S(q'(\mathbf{w})) - S(q(\mathbf{w})) \simeq \frac{1}{r} \sum_{g=1}^r (-\beta \text{Tr}[\mathbf{H}(\mathbf{w}'^g)] - \beta^2 \text{Tr}[\mathbf{H}(\mathbf{w}'^g)\mathbf{H}(\mathbf{w}'^g)]).$$

Final estimation for the τ optimization step:

$$\begin{aligned} \log \hat{p}(\mathbf{Y}|\mathcal{D}, \mathbf{h}) &\sim \frac{1}{r} \sum_{g=1}^r L(\mathbf{w}_\tau^g, \mathcal{D}, \mathbf{Y}) + S(q^0(\mathbf{w})) + \\ &+ \frac{1}{r} \sum_{b=1}^{\tau} \sum_{g=1}^r (-\beta \text{Tr}[\mathbf{H}(\mathbf{w}_b^g)] - \beta^2 \text{Tr}[\mathbf{H}(\mathbf{w}_b^g)\mathbf{H}(\mathbf{w}_b^g)]), \end{aligned}$$

\mathbf{w}_b^g is a parameter vector for optimization g on the step b , $S(q^0(\mathbf{w}))$ is an initial entropy.

How to calculate Hessian trace?

Problem

$$\text{Tr}[\mathbf{H}(\mathbf{w}_b^g)]$$

Statement

Let \mathbf{U} be a symmetric matrix and \mathbf{v} be the random vector with the following properties:

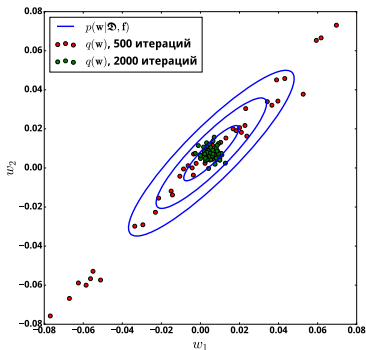
- ① $E v_i = 0$;
- ② $\text{Var}(v_i) = 1$.

Then

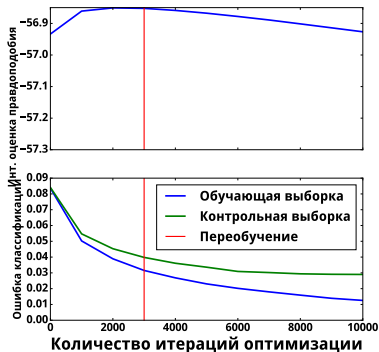
$$E \mathbf{v}^T \mathbf{U} \mathbf{v} = \text{Tr}[\mathbf{U}].$$

Overfitting, Maclaurin et. al, 2015

Gradient descent does not optimize KL-divergence $KL(q(\mathbf{w})||p(\mathbf{w}|\mathcal{D}, \mathbf{h}))$. Evidence estimation gets worse while optimization tends to the optimal parameter values. This can be considered as a overfitting start.



Convergence



Overfitting start

Stochastic gradient Langevin dynamics

A modification of SGD:

$$T = \mathbf{w} - \beta \nabla L + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \frac{\beta}{2})$$

where β changes with a number of iterations:

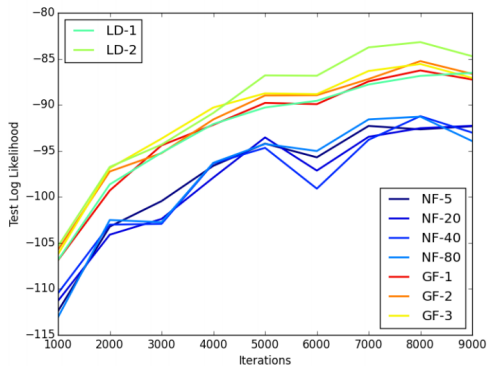
$$\sum_{\tau=1}^{\infty} \beta_{\tau} = \infty, \quad \sum_{\tau=1}^{\infty} \beta_{\tau}^2 < \infty.$$

Statement [Welling, 2011]. Distribution $q^{\tau}(\mathbf{w})$ converges to posterior distribution $p(\mathbf{w}|\mathbf{X}, \mathbf{f})$.
Entropy adjustment:

$$\hat{S}(q^{\tau}(\mathbf{w})) \geq \frac{1}{2} |\mathbf{w}| \log \left(\exp \left(\frac{2S(q^{\tau}(\mathbf{w}))}{|\mathbf{w}|} \right) + \exp \left(\frac{2S(\epsilon)}{|\mathbf{w}|} \right) \right).$$

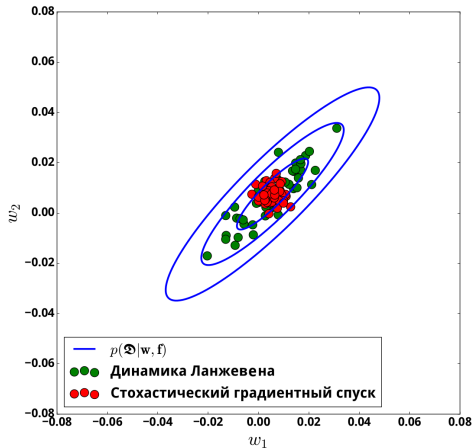
Stochastic gradient Langevin dynamics for generative models

Altieri et al., 2015: sample latent variable \mathbf{z} and use SGLD as a normalizing flow.



SGLD vs SGD

Parameter distribution after 2000 iterations:



Stein operator

Given a smooth probability function p and a smooth vector function ϕ . Define a Stein operator as the following:

$$\mathcal{A}_p \phi(\mathbf{x}) = \nabla_{\mathbf{x}} \log p(\mathbf{x}) \phi^T + \nabla_{\phi} \phi(\mathbf{x}).$$

Stein's identity:

$$\mathbb{E}_{\mathbf{x} \sim p} \mathcal{A}_p \phi(\mathbf{x}) = 0.$$

If we use q instead of p in the \mathcal{A}_p we get a non-zero result, but close to zero as soon as p is close to q .

Let $T(\mathbf{x}) = \mathbf{x} + \varepsilon \phi(\mathbf{x})$. Then:

$$\nabla_{\varepsilon} KL(q||p)|_{\varepsilon=0} = \mathbb{E}_{\mathbf{x} \sim q} \text{trace} \mathcal{A}_p \phi.$$

Given a kernel \mathbf{K} , the optimal ϕ for minimizing KL is:

$$\phi^*(\mathbf{x}') = \mathbb{E}_{\mathbf{x} \sim q} \nabla_{\mathbf{x}} \log p(\mathbf{x}) \mathbf{K}(\mathbf{x}, \mathbf{x}') + \nabla_{\mathbf{x}} \mathbf{K}(\mathbf{x}, \mathbf{x}').$$

Stein operator: algorithm

Algorithm 1 Bayesian Inference via Variational Gradient Descent

Input: A target distribution with density function $p(x)$ and a set of initial particles $\{x_i^0\}_{i=1}^n$.

Output: A set of particles $\{x_i\}_{i=1}^n$ that approximates the target distribution $p(x)$.

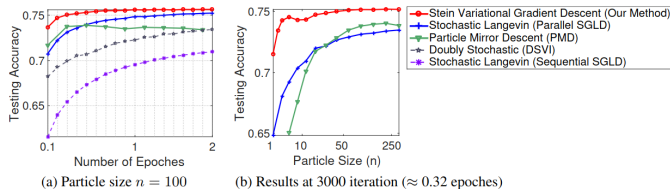
for iteration ℓ **do**

$$x_i^{\ell+1} \leftarrow x_i^\ell + \epsilon_\ell \hat{\phi}^*(x_i^\ell) \quad \text{where} \quad \hat{\phi}^*(x) = \frac{1}{n} \sum_{j=1}^n [k(x_j^\ell, x) \nabla_{x_j^\ell} \log p(x_j^\ell) + \nabla_{x_j^\ell} k(x_j^\ell, x)], \quad (8)$$

where ϵ_ℓ is the step size at the ℓ -th iteration.

end for

Stein operator: results



Dataset	Avg. Test RMSE		Avg. Test LL		Avg. Time (Secs)	
	PBP	Our Method	PBP	Our Method	PBP	Ours
Boston	2.977 ± 0.093	2.957 ± 0.099	-2.579 ± 0.052	-2.504 ± 0.029	18	16
Concrete	5.506 ± 0.103	5.324 ± 0.104	-3.137 ± 0.021	-3.082 ± 0.018	33	24
Energy	1.734 ± 0.051	1.374 ± 0.045	-1.981 ± 0.028	-1.767 ± 0.024	25	21
Kin8nm	0.098 ± 0.001	0.090 ± 0.001	0.901 ± 0.010	0.984 ± 0.008	118	41
Naval	0.006 ± 0.000	0.004 ± 0.000	3.735 ± 0.004	4.089 ± 0.012	173	49
Combined	4.052 ± 0.031	4.033 ± 0.033	-2.819 ± 0.008	-2.815 ± 0.008	136	51
Protein	4.623 ± 0.009	4.606 ± 0.013	-2.950 ± 0.002	-2.947 ± 0.003	682	68
Wine	0.614 ± 0.008	0.609 ± 0.010	-0.931 ± 0.014	-0.925 ± 0.014	26	22
Yacht	0.778 ± 0.042	0.864 ± 0.052	-1.211 ± 0.044	-1.225 ± 0.042	25	25
Year	$8.733 \pm \text{NA}$	$8.684 \pm \text{NA}$	$-3.586 \pm \text{NA}$	$-3.580 \pm \text{NA}$	7777	684

Rényi Divergence Variational Inference

Rényi's α -divergence

$$D_\alpha[p||q] = \frac{1}{\alpha - 1} \log \int p(\theta)^\alpha q(\theta)^{1-\alpha} d\theta \quad (1)$$

- ① continuous and non-decreasing on $\alpha \in \{\alpha : |D_\alpha| < +\infty\}$
- ② for $\alpha \notin \{0, 1\}$, $D_\alpha[p||q] = \frac{\alpha}{1-\alpha} D_{1-\alpha}[p||q] \rightarrow D_\alpha[p||q] \leq 0, \alpha < 0$

Different divergence functions

α	Definition	Notes
$\alpha \rightarrow 1$	$\int p(\boldsymbol{\theta}) \log \frac{p(\boldsymbol{\theta})}{q(\boldsymbol{\theta})} d\boldsymbol{\theta}$	<i>Kullback-Leibler (KL) divergence</i> , used in VI (KL[q p]) and EP (KL[p q])
$\alpha = 0.5$	$-2 \log(1 - \text{Hel}^2[p q])$	function of the square <i>Hellinger distance</i> zero when $\text{supp}(q) \subseteq \text{supp}(p)$
$\alpha \rightarrow 0$	$-\log \int_{p(\boldsymbol{\theta}) > 0} q(\boldsymbol{\theta}) d\boldsymbol{\theta}$	(not a divergence)
$\alpha = 2$	$-\log(1 - \chi^2[p q])$	proportional to the χ^2 -divergence
$\alpha \rightarrow +\infty$	$\log \max_{\boldsymbol{\theta} \in \Theta} \frac{p(\boldsymbol{\theta})}{q(\boldsymbol{\theta})}$	<i>worst-case regret</i> in <i>minimum description length principle</i> [24]

- When $\alpha = 0$, we get an approximate Evidence (like in IWAE).
- When $\alpha = 0.5$, we get ELBO.
- When $\alpha \rightarrow \infty$ we get mode-seeking (also called zero-forcing) optimization.
- When $\alpha \rightarrow -\infty$ we get mass-preserving optimization.

Gaussian example

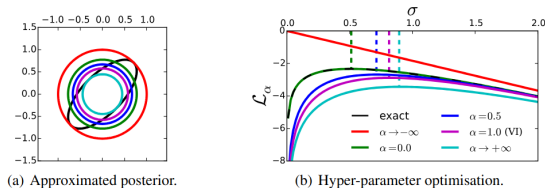


Figure 1: Mean-Field approximation for Bayesian linear regression. In this case $\varphi = \sigma$ the observation noise variance. The bound is tight as $\sigma \rightarrow +\infty$, biasing the VI solution to large σ values.

Note, $\alpha \rightarrow \infty$ works similar to MAP, but still can give some non-degenerate probabilistic estimations.

Regression example

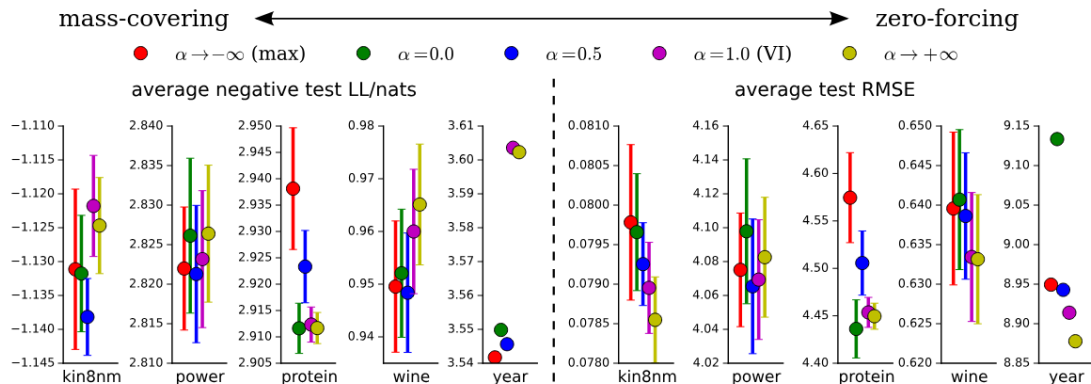


Figure 4: Test LL and RMSE results for Bayesian neural network regression. The lower the better.

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Organizational issues

- 29th of October: no classes, moving to the 31st?
- Technical meeting: On 5th of November?
- Format: each team shows the version of basic code, the draft version of blog-post and draft version of the documentation (deployed on the server or in stand-alone mode).
- All the presented materials must be stored at the github
 - ▶ for blog-post, you can put read-only link for the overleaf if you write the post here.
- Talks: please make them.