The Early Career Gender Wage Gap among University Graduates

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Abstract

A large literature has shown that the gender wage gap is small in the beginning of the career and increases gradually with age, mostly because of family decisions, i.e., a penalty caused by child birth. Using a unique dataset that links university graduates with detailed employment records from the German social security register, we find that a significant gender wage gap already exists in the first job after graduation before most young individuals make family decisions. However, the gender wage gap decreases in the first year after the labor market entry before slowly increasing over time. As an explanation for the decrease of the gender wage gap, we find that female graduates have higher returns to firm and occupation changes than their male counterparts. Specifically, women may use firm and occupation changes to correct for a skill mismatch, which is more common for women than men in the first job.

JEL Classification: I23, J16, J31, J71

Keywords: Gender Wage Gap, University Graduates, Early Career

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1 Introduction

Despite advances in women's education and career opportunities in recent decades, a persistent gender wage gap remains prevalent in economically advanced nations (Goldin, 2014; Olivetti and Petrongolo, 2016), which is even larger among individuals with higher levels of education (Blau and Kahn, 2017; OECD, 2019). Many studies have examined the gender wage gap among the highly educated and found that women's lower labor supply and more frequent career interruptions (mainly due to child care) compared to men are the main reasons for this gender wage gap.¹

Less is known about the existence and development of a gender wage gap at the beginning of a career. This lack is surprising given that starting wages and growth in early career wages have long-lasting effects on the future labor market outcomes of university graduates and potentially on the gender wage gap. For example, Oreopoulos et al. (2012); Kahn (2010); Oyer (2006) show that labor market conditions, such as recessions, can also have an impact on entry wages and consequently on wages in the long run. Moreover, prior wages usually determine wage increases due to promotions within the same firm (Graham et al., 2000) and even wage rises as a result of a job change are usually based on previous wages (Hansen and McNichols, 2020). These findings show that entry wages are important in determining future wages over the long run and are therefore important for the origin of a gender wage gap.

Whether a gender wage gap exists in the first years after graduation is theoretically ambiguous. Particularly in the first job, some common reasons for pay differences between men and women, such as family-related decisions (e.g., childbirth or marriage), career-related developments (e.g., promotions), work experience, and firm-specific networks, may not yet be relevant.² Therefore, we expect no or only a small gender wage gap in the initial job, especially when we take into account gender differences in the field of study and the characteristics of the employer in the first job.

However, particularly in the first job, both the applicants and the firms face considerable uncertainty. Firms can only assess the labor market productivity of candidates without prior work experience based on their university grades and interview performance. Given that women tend to have higher GPAs than men nowadays (Becker et al., 2010; Francesconi and Parey, 2018), on the one hand, we may even expect the gender wage gap conditional on differences in field of study choice and employer characteristics to be in favor of women in the first job. On the other hand, existing studies show that female applicants negotiate less in job interviews than male applicants (Babcock and Laschever, 2009; Bertrand, 2011) and they may face

¹For example, see Cortes and Pan (2020), Kleven et al. (2019a), Adda et al. (2017) and Kuziemko et al. (2018).

²The mean age of German mothers at first birth was 30.5 in 2021 (Federal Statistical Office, 2022), while the average age of labor market entry for women in our sample is 27. Moreover, the average age of women at birth is expected to increase with the level of education. Therefore, this issue is not expected to be of high magnitude in the case of women with a master's degree at labor market entry.

statistical discrimination at labor market entry (Altonji and Pierret, 2001; Pinkston, 2006). Furthermore, differences in preferences, such as risk aversion, can be particularly important at the beginning of the career, as women may be more likely to accept lower-paying job offers due to their higher levels of risk aversion (Cortés et al., 2021). As a result, the gender wage gap could be substantial and in favor of men, given differences in field of study and employer characteristics in the first job.

This ambiguity about the gender wage gap may be even greater in the years after labor market entry, when firms have observed the productivity of their employees or when graduates change jobs to increase their wages. If women earn less than men in the first job as a result of discrimination, the gap may narrow over time as women move to less discriminatory firms or as employers learn about employees' true productivity over time (Altonji and Pierret, 2001; Farber and Gibbons, 1996). However, the gender wage gap may also increase over time due to job changes in the early stages of a career, as the literature shows that women generally realize lower returns to job mobility compared to men (Albrecht et al., 2018; Topel and Ward, 1992). In addition, the gap may widen over time as family-related decisions become more important over time. Overall, the gender wage gap at labor market entry and the dynamics of the gender wage gap in the early years of a career remain unclear and important to study.

This study examines the gender wage gap immediately after entering the labor market and its evolution during the initial years of a career for more than 5,000 university graduates with a master's degree or equivalent. We use unique administrative data on graduates of a large German university linked with detailed social security data from the Integrated Employment Biographies (IEB). This linked administrative dataset provides a wide range of information from the two data sources, including sociodemographic characteristics of the graduates, the attained university degree, field of study, the final high-school and university grades, the date of enrollment and the exact timing of graduation, labor market entry, and any occupation or firm changes.

Using these data, we first estimate the gender wage gap at labor market entry among university graduates. Our findings show that males have significantly higher wages than females in their first full-time job immediately after graduation, despite our homogeneous and highly educated sample with high labor market attachment. The estimated unadjusted gender wage gap of about 12.5 log points corresponds to around 10 euros per day or 300 euros per month. The adjusted gender wage gap, conditional on a comprehensive set of personal and pre-graduation controls, is equal to 6.2 log points. Including occupation fixed effects reduces the gender wage gap to 4.7 log points. Other post-graduation characteristics, such as the timing of the first job, firm fixed effects, the share of women in the firm, and the location of the firm do not alter the gender wage gap substantially.

Second, since both career paths and wages vary widely across fields of study (Altonji et al., 2016), we con-

duct a subgroup analysis for four broad groups of fields of study: economics and business; mathematics and natural sciences; humanities and social sciences; and medical studies. The results show that the unadjusted (raw) gender wage gap in the first job is prominent in almost all field groups except medical studies. For mathematics and natural sciences, the gender wage gap disappears when controlling for the major subject within the field of study. The adjusted gender wage gap is the highest in the humanities and social sciences. This field group also has the lowest average daily wage in the first job, the highest variation in wages, and the highest share of females.

Third, as dynamics are very important, particularly in the early years of a career, and have an impact on future wage growth, we examine the dynamics of the gender wage gap over the first years after labor market entry. Our findings reveal a decrease in the estimated gender wage gap in the first three years after labor market entry, followed by an increase in subsequent years. The largest reduction in the wage gap occurs one year after labor market entry. Moreover, we demonstrate that this decrease is observed only among economics and business graduates and humanities and social sciences graduates who change both firms and occupations within one year of entering the labor market. However, this decline does not occur for graduates from other fields of study or for those who remain in the same firm or occupation.

Finally, our analysis focuses on these two field groups; economics and business, humanities and social sciences. It shows that mainly women who change firms and occupations after their first job drive the decline in the gender pay gap, as women benefit more from these changes than men. Our data show that women are more likely than men to work in a mismatched occupation at the first job. By changing both firms and occupations, women move out of the lowest paid occupations and are able to correct this mismatch, leading to a higher increase in wages compared to men. After comparing this empirical finding with several theories in the gender wage gap literature, the one explanation for our finding may be that women immediately after graduation have a high preference for certain job and firm amenities, such as job meaning and relevance, leading them to initially accept mismatched jobs. One year after the labor market entry, their preferences might change, and they correct this mismatch by changing their occupation and firm. However, our data do not allow for a definitive test of this hypothesis.

Our study contributes to the existing literature in three important ways. First, several studies examine the dynamics of the gender wage gap over the life cycle and find evidence that the gender wage gap is smaller at younger ages but increases over time, mainly due to family-related decisions (Bertrand et al., 2010; Albrecht et al., 2018; Manning and Swaffield, 2008). Although these studies provide valuable insight into the dynamics of the gender wage gap in general, they do not focus on the first job after graduation. The

³The effect of the child penalty on females' labor market outcomes is explored in several studies. For example, Kleven et al. (2019b), Dustmann et al. (2009).

few papers that examine the gender wage gap at the beginning of a career rely primarily on survey data. For example, Cortés et al. (2021) find in a survey of about 1,350 business school graduates in the US that women earn 10 percent less than men in their first job. In a related German study, Francesconi and Parey (2018) find an adjusted gap of 5-10 log points among full-time German college graduates after 12-18 months. In contrast to this literature, we are the only study investigating the gender wage gap that uses administrative data and zooms in on the first job after graduation. Administrative data help to avoid reporting bias that can occur in survey-based studies at the beginning of a career due to frequent job changes. With our data, we are able to analyze the origins of the gender wage gap when graduates earn wages for the first time in their careers.⁴

Second, in addition to accurately identifying the first job, this novel linked dataset offers several other advantages. For example, the administrative nature of the data overcomes concerns associated with missing data, response rates, or measurement error due to retrospective questions. In addition, most other data used to study the gender wage gap either lack comprehensive information on graduates' pre-graduation characteristics (field of study, GPA) or are unable to track graduates as they transition into the labor market and lack information on graduates' occupation, industry, and other important employment characteristics. In contrast, our linked data provide access to accurate and comprehensive measures of human capital determinants of productivity, including academic grades and field of study, as well as detailed information on employment, wages, and occupations.

Third, our study provides unique insights into early career job dynamics and their impact on the gender wage gap. At the beginning of careers, high information friction can lead to poor job matches in the labor market for recent graduates (Vesterlund, 1997). Fredriksson et al. (2018) highlights high separation rates among inexperienced employees due to limited information about the labor market. Consequently, graduates may correct these mismatches by changing jobs once they accumulate some experience. Topel and Ward (1992) further emphasize that job mobility is more prevalent during the early years of labor market entry, with young employees typically experiencing an average of seven full-time jobs within the first ten years after entering the labor market. The existing literature has shown that job changes are associated with wage growth (Albrecht et al., 2018; Manning and Swaffield, 2008; Del Bono and Vuri, 2011; Topel and Ward, 1992). Several studies also observe that men are more likely to change jobs⁵ and tend to benefit more from job mobility than women, thereby exacerbating the gender wage gap over time (Albrecht et al., 2018; Manning and Swaffield, 2008; Del Bono and Vuri, 2011; Topel and Ward, 1992). However, these studies do not focus on the first years after labor market entry because it is difficult to observe this crucial early period without

⁴The study by Kunze (2005) uses administrative data but focuses on younger graduates who have completed an apprenticeship. 5 Except for the study by Albrecht et al. (2018) find similar probabilities of job change for men and women.

detailed administrative data.

In contrast, we are able to follow all graduates without attrition over time, which allows us to observe the precise timing of any job changes and job search periods within the dynamic first years after labor market entry. This information allows us to observe the share of female and male graduates from each field of study who change their jobs and to observe the returns to their mobility, which may have long-lasting effects on their future labor market careers. In addition, using information on firms, such as industry sector, firm size and location, we are able to observe the influence of firm and occupation characteristics on the evolution of the wage gap for job changers.

The remainder of this paper is structured as follows. Section 2 describes the dataset, its advantages and shortcomings, characterizes the sample of university graduates used in the analysis, and presents some descriptive statistics. The results of the estimated gender wage gap at labor market entry and the dynamics of the gender wage gap over the first few years of a career are presented in Section 3 and Section 4 respectively. We examine gender differences in firm and occupational mobility in Section 5 and the underlying reasons for this mobility in Section 6, before concluding in Section 7.

2 The Linked Administrative Dataset and University Graduates

2.1 Data

This study is based on a unique administrative dataset of graduates of the University of Regensburg linked with register data from the German Integrated Employment Biographies (IEB) of the Institute for Employment Research (IAB). The linked dataset combines detailed study information on each graduate from the registry of the University of Regensburg with information on individual employment records covering the whole employment biography of jobs subject to social security contributions from the IEB dataset.

The University of Regensburg is a large university located in Bavaria, Germany.⁶ It offered almost all fields of study except engineering degrees during the observed period. The available dataset from the University of Regensburg covers all its graduates from 1995 to 2016. The data are highly reliable, as they are based on administrative records from the university registry. The dataset provides information on the personal characteristics of each graduate, such as year of birth, gender, nationality, as well as the district and grade of the certificate of general qualification for university admission (Abitur), hereafter referred to as the final high-school grade point average (GPA). The dataset also includes study-related characteristics

⁶In 2021, the University of Regensburg was ranked 44th out of approximately 400 universities in Germany (Times Higher Education, 2021). Liste der Hochschulen in Deutschland (2021) provides further details on the size of the University of Regensburg.

at the university, such as the field of study, the type of university degree attained, the final GPA, and the dates of enrollment and graduation.

The IEB is a large administrative dataset on individuals' employment biographies provided by the IAB for the period 1975-2019. The information provided by the dataset is highly reliable, as it is a legal requirement in Germany for all employers to provide information on their employees to the German Social Security Administration. The IEB dataset includes individuals in employment covered by social security, excluding the self-employed and civil servants. Thus, the IEB dataset covers about 80 percent of the total labor force in Germany. In addition to the precise timing of employment and out-of-employment spells, the dataset provides information on gross daily wages, industry, occupation (3-digit), full-time status, and other employment characteristics (Dorner et al., 2010). The data from the University of Regensburg are merged with the IEB dataset using a linkage procedure established at the IAB based on an individual's full name, gender, and date of birth, with a 90 percent match rate (Möller and Rust, 2017).

2.2 Sample Choice

The focus of this study is on individuals with a master's degree or equivalent with available university GPA data. We further focus on graduates working full-time in their first job with a daily wage of at least 10 Euros,⁸ and omit graduates who do not work full-time in their first job after graduation, even if they subsequently switch to full-time employment, because we consider this to be the most policy-relevant group with a higher labor market attachment and the results are easier to interpret for a more homogeneous group.⁹ If an individual has more than one wage spell at a given time, we choose the "main" employment spell as defined by the IAB.

Furthermore, we exclude individuals who are older than 35 years (10 percent or 110 individuals) and keep graduates who only work in regular jobs; for example, we exclude jobs such as internships, student jobs, etc. (10 percent or 293 individuals). We also exclude graduates with a gap of more than 15 months between graduation and their first employment spell (9 percent or 1078 individuals), as these individuals may have spent time abroad or already worked on a self-employed basis (which is not captured by the data). Since our main analysis focuses on the first full-time job after graduation and subsequent years, we keep graduates

⁷Please refer to the study by Möller and Rust (2017) for a more detailed explanation of the matching procedure.

⁸Wages are deflated to 2010 euros using the consumer price index.

⁹Since information on working hours is not available in the linked dataset, we focus on full-time jobs in order to eliminate a potential bias in the gender wage gap induced by differences in working hours. Since we focus on full-time employees in our main analysis, the working hours of men and women should be reasonably comparable. However, even if employees are fairly homogeneous in terms of full-time employment, males may still work more hours than females, allowing them to earn higher wages (e.g., Goldin and Katz, 2016). The study by Francesconi and Parey (2018) documents that differences in hours worked among full-time employees do not significantly explain the gender wage gap among German graduates about 12 to 18 months after graduation. Therefore, we expect that our results are not driven by differences in working hours between full-time employed female and male graduates.

with wage spells at the beginning of the first job and also one year after the first job (9 percent or 812 individuals). After dropping observations with missing values (mainly university GPA with 118 missing), the final sample for the main wage estimations consists of 5,212 individuals.

2.3 Descriptive Statistics

Table 1 presents descriptive statistics at labor market entry for our preferred sample of university graduates in full-time employment, which we use for the wage analysis in the following sections. While panel A of Table 1 documents pre-graduation characteristics, such as university and high-school GPA, duration of study, non-German citizenship, and others, panel B presents post-graduation characteristics, including characteristics of the first job.

Panel A of Table 1 shows that both men and women complete their degrees in about five and a half years on average, although women graduate at a younger age. The majority of graduates at the University of Regensburg (around 70 percent) acquire some form of work experience before graduation, with females more likely to work during their studies than males. In addition, consistent with the literature, the share of female graduates increases over graduation cohorts, with the male-female ratio even reversing in the most recent cohort group (2007-2010). This is also in line with the overall population of German graduates. ¹⁰ Consistent with the literature (see e.g., Becker et al., 2010), female graduates enroll at the University of Regensburg with better final high-school grades and leave the university with slightly better university grades (by around 11 percent of the sample standard deviation) than males. ¹¹ Females are also more than twice as likely to graduate in humanities and social sciences. Mathematics and natural sciences graduates account for nearly a quarter of all male graduates, compared to only eight percent of all female graduates. Nevertheless, economics and business remain the dominant field of choice for both genders. Finally, the share of women studying medicine is around 10 percentage points higher than that of men. Table 1 also shows that female graduates are more likely than males to earn a magister or state examination degree. ¹² The vast majority of graduates have a diploma degree, with a higher proportion of men than women.

Table C.2 in the Appendix compares our estimation sample with official German register data and other representative studies. For 2010, our sample shows a slightly higher female representation (50 percent) compared to the register data (46 percent). The share of females by field of study is also comparable. The

¹⁰According to BMBF (2021), while the female-to-male ratio was around 0.8 in 1995, it increased to 1.1 in 2010. Appendix Figure ?? shows the female-to-male ratio for the underlying sample of Regensburg University graduates who work full-time in their first job after graduation. Among these graduates, the female-to-male ratio increases over time from around 0.3 to 1.1. Compared to the total population of German graduates, the share of females is lower in the first graduation cohorts and converges to the total population over time. Appendix Figure ?? further illustrates that in our data the proportion of women is increasing for all fields of study except for mathematics and natural sciences.

 $^{^{11}}$ In our data, final high-school grades are only available in the data beginning with the 2001 graduation cohort.

¹²Because the Bologna reform was implemented in Germany not earlier than between 2005 and 2010, only a small proportion of graduates in the sample have a master's degree.

Table 1: Descriptive Statistics

	М	Male		emale	D:0
	Mean	(Std. dev.)	Mean	(Std. dev.)	Diff.
Panel A: Pre-graduation and Persona					
Final High-School Grade (Abitur)	2.245	(0.616)	2.077	(0.594)	0.168***
Individuals		1,476		1,200	
Final University Grade	2.058	(0.604)	1.998	(0.568)	0.060***
Non-German Citizenship	0.015	(0.121)	0.033	(0.179)	-0.019***
Graduation Age	27.238	(1.864)	26.579	(1.906)	0.660***
Duration of Study	5.592	(1.308)	5.631	(1.372)	-0.039
Apprenticeship	0.065	(0.247)	0.061	(0.239)	0.004
Worked During Studies	0.673	(0.469)	0.742	(0.438)	-0.069***
Origin is Bavaria	0.876	(0.329)	0.861	(0.346)	0.016
Graduation Year					
- 1995 - 1998	0.283	(0.451)	0.156	(0.363)	0.127***
- 1999 - 2002	0.227	(0.419)	0.172	(0.377)	0.055***
- 2003 - 2006	0.227	(0.419)	0.269	(0.443)	-0.042***
- 2007 - 2010	0.263	(0.440)	0.403	(0.491)	-0.141***
Field of Study					
- Economics and Business	0.469	(0.499)	0.328	(0.469)	0.141***
- Mathematics and Natural Sciences	0.224	(0.417)	0.077	(0.267)	0.147***
- Humanities and Social Sciences	0.111	(0.314)	0.300	(0.458)	-0.189***
- Medical Studies	0.196	(0.397)	0.295	(0.456)	-0.099***
Type of Degree					
- Diploma	0.747	(0.435)	0.576	(0.494)	0.172***
- Magister	0.046	(0.210)	0.114	(0.317)	-0.068***
- Master	0.010	(0.102)	0.015	(0.123)	-0.005
- State Examination (Staatsexamen)	0.196	(0.397)	0.295	(0.456)	-0.099***
Individuals		3,258	1,954		
Panel B: Post-graduation Characteris Left the State	0.196	(0.397)	0.179	(0.383)	0.118
Left the State Left the City	0.683	(0.465)	0.655	(0.476)	0.028*
v		, ,	0.000	,	0.020
Mean of Job Search Duration	3.747	(3.128)	3.817	(3.039)	-0.070
Duration of Job Search					
- less than 1 Month	0.190	(0.392)	0.161	(0.368)	0.028***
- 1-3 Months	0.326	(0.469)	0.319	(0.466)	0.007
- 3-5 Months	0.214	(0.410)	0.247	(0.431)	-0.033***
- more than 5 Months	0.270	(0.444)	0.272	(0.445)	-0.002
Firm Size		(5.45.5)		(= 15.3	
- less than 25 Employees	0.238	(0.426)	0.247	(0.432)	-0.009
- 25-250 Employees	0.273	(0.445)	0.266	(0.442)	0.007
- 250-2000 Employees	0.254	(0.436)	0.273	(0.446)	-0.018
- more than 2000 Employees	0.235	(0.424)	0.214	(0.411)	0.020*
Share of Women in Firm		(0.4==)	0	(0.45.)	
- less than 40%	0.356	(0.479)	0.201	(0.401)	0.155***
- 40%-70%	0.405	(0.491)	0.383	(0.486)	0.022
- more than 70%	0.239	(0.427)	0.417	(0.493)	-0.177***
Individuals		3,258		1,954	

Note: This table shows summary statistics of graduates' pre-graduation and post-graduation characteristics. The sample consists of graduates with a master's degree or equivalent who work in a full-time job as their first job after graduation and also who have a wage spell 1 year after the first job. ***, ** and * denote significance at the 1, 5, and 10% levels.

largest difference is observed in mathematics and natural sciences. In our data the proportion of females in this field is 18 percent, while in the register data it is 35 percent. In our sample, students have a slightly better final high-school GPA than in the register data and similar university grades with the representative sample from the survey data Francesconi and Parey (2018). The graduates in our sample are, on average, about five months younger because we use examination rather than ex-matriculation dates. Notably, 11 percent of our students are non-German, compared to 22 percent in the register data, as we cannot observe individuals in our data if they move to another country after graduation.

Panel B of Table 1 presents post-graduation and employment characteristics, such as mobility, the time between graduation and the first full-time job, ¹³ establishment size, and the share of women in the establishment of the first job. ¹⁴ The table shows that about 70 percent of the graduates find their first full-time job outside of the city of Regensburg, with males being slightly more mobile. On average, female graduates take longer to find their first job than male graduates, around 3.7 and 3.8 months, respectively. A breakdown of the duration of the job search into different categories shows that the share of male graduates with a job search duration of less than a month is higher. Finally, female and male graduates tend to work in establishments of similar size. However, in line with the literature, women are more likely to work in establishments with a higher proportion of female employees. A potential explanation for this might be the sorting of university graduates into specific industries by gender, resulting in female-dominated industries (Hellerstein et al., 2011).

3 Gender Wage Gap at Labor Market Entry

We begin our analysis by looking at gender wage differences at labor market entry.¹⁵ To identify gender differences, we estimate the following regression equation:

$$ln_wage_i = \alpha + \gamma female_i + \beta X_i + \epsilon_i \tag{1}$$

 ln_wage_i is the log real daily wage at the first job. The analysis of the dynamics of the gender wage gap uses log daily wages 1-5 years after the initial job as the dependent variable.

Table 2 presents the first results of our empirical analysis for the 1995 to 2010 graduation year cohorts. Column (1), controlling only for the year of graduation, shows a significant negative coefficient of 12.5 log points for the female dummy. This unadjusted (raw) gender wage gap indicates that female graduates earn

 $^{^{13}}$ Hereafter referred to as "job search duration", even though this time is not necessarily spent searching for a job.

¹⁴The data only include information on establishments not on firms. However, in this paper, we use the terms "establishment" and "firm" interchangeably.

¹⁵By labor market entry, we refer to the first job after university graduation and use these terms interchangeably.

12.5 log points less in daily wages than their male counterparts in their first job after graduation. This gap is smaller than in the study by Francesconi and Parey (2018), which finds a raw gender wage gap of around 20 log points based on survey data collected in a few selected years between 1988 and 2010 and conducted among graduates 12-18 months after graduation. This difference in the raw gender wage gap may reflect the timing of their data (they do not focus on the first job after graduation) as well as our conservative definition of the first job, i.e., we have a more homogeneous group of graduates with higher labor market attachment.

Table 2: Gender Wage Gap at Labor Market Entry

Dependent Variable: Log Daily	Wage						
		Addi	tional				
		Per	sonal and I	Pre-Gradua	tion	Post-Gr	aduation
			Charac	teristics		Charac	teristics
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Female	-0.125*** (0.012)	-0.130*** (0.012)	-0.068*** (0.012)	-0.064*** (0.012)	-0.062*** (0.012)	-0.045*** (0.011)	-0.047*** (0.011)
Graduation year	Yes						
Personal characteristics	No	Yes	Yes	Yes	Yes	Yes	Yes
Field of study FE	No	No	Yes	Yes	Yes	Yes	Yes
Final university grade	No	No	No	Yes	Yes	Yes	Yes
Pre-graduation characteristics	No	No	No	No	Yes	Yes	Yes
Occupation FE	No	No	No	No	No	Yes	Yes
Post-graduation characteristics	No	No	No	No	No	No	Yes
R-squared	0.036	0.038	0.211	0.225	0.235	0.334	0.399
Individuals	5,212	5,212	5,212	5,212	5,212	5,212	5,212

Note: This table shows the gender wage gap at labor market entry, based on the OLS model specified in equation 1. The sample consists of graduates with a master's degree or equivalent who work full-time in their first job after graduation and also have a wage spell 1 year after the first job. The dependent variable is the log gross daily wage in the first job. The control variables are added step-wise. Column (1) shows the results with only the year of graduation as a control. Column (2) adds personal characteristics such as age and (non-)German citizenship, and column (3) adds field of study (17 categories). Column (4) adds the final university grade and column (5) adds pre-graduation characteristics: duration of study, location of the final high-school examination, a dummy for apprenticeship and a dummy for working while studying. Column (6) adds 3-digit occupation fixed effects. Column (7) shows the results after adding post-graduation characteristics. These include job search duration, job location, 1-digit industry fixed effects, firm size (7 categories), the share of women in the firm (3 categories) and the starting month of the first job. Robust standard errors are reported in parentheses. ***, ** and * denote significance at the 1, 5, and 10% levels, respectively.

While we add personal characteristics in column (2), in column (3) we additionally control for 17 fields of study (see Table C.1 in the Appendix for a list of these 17 fields), which leads to a big decline in the gender wage gap to 6.8 log points in the first job after graduation. This striking decrease in the gender wage gap confirms the findings of existing studies (see, e.g., Machin and Puhani, 2003; Black et al., 2008) that female

students sort into fields of study associated with lower wages.

Columns (4) and (5) show the extent to which the results change when we include the final university grades and other pre-graduation characteristics (duration of study, location of the high school, having completed an apprenticeship and having worked while studying) in the regression. The estimated gender wage gap barely changes after controlling for these characteristics, suggesting that neither final university grades nor other pre-graduation characteristics explain a large part of the gender wage gap. Table C.4 in the Appendix presents additional results from an Oaxaca-Blinder decomposition. The decomposition also shows that the most important contributor to the gender wage gap among pre-graduation characteristics is the field of study, accounting for 40 percent of the total gender wage gap at the first job. Since the field of study explains the largest part of the gender wage gap and, therefore, explanations for the pay gap may strongly vary between field of study, in the next section we examine the gender wage gap within a broader set of field of study categories.

Columns (6) and (7) additionally include occupation fixed effects (at the 3-digit level) and other post-graduation characteristics in the estimation. However, it is not clear whether these post-graduation variables should be included in the estimation, as they may themselves be outcomes of the variable of interest, such as choice of location or type of job (or occupation).¹⁷ After adding occupation fixed effects to the estimation, the gender wage gap decreases to 4.5 log points, indicating that, similar to the field of study, the occupation of the first job after graduation explains a large part of the gender wage gap. Finally, column (7) adds all post-graduation controls, which does not further reduce the gender wage gap.

Overall, the gender pay gap remains significant at 6.2 log points for graduates in the same field of study, with similar grades and other pre-graduation and personal characteristics, and at 4.7 log points when we also condition on occupation and other post-graduation characteristics. These gaps are highly significant, as the unadjusted (raw) gender wage gap (12.5 log points) corresponds to 10 euros per day, or around 300 euro per month, less pay for women compared to men in their first job.¹⁸

Field of Study

Since earnings vary by field of study¹⁹ and we have shown that field of study is the main contributor to the gender wage gap at labor market entry, we next examine the gender wage gap across fields of study. The results in Table 3 show that the gender wage gap is high for all fields of study, except medical studies. The raw gender wage gap (controlling only for the year of graduation) is 8.6 log points for economics and business

¹⁶When we add final high school grades to the estimation, the gender wage gap decreases by only 0.003 log points.

¹⁷Angrist and Pischke (2009) define controls that can be dependent variables as "bad controls".

¹⁸We also estimate the same equation including all types of first jobs and document it in Table C.3 of the Appendix.

¹⁹Several studies demonstrate that different fields of study yield varying labor market returns. For example, Deming and Noray (2020); Kirkeboen et al. (2016); Kelly et al. (2010) and Altonji et al. (2012).

graduates. This gap is comparable to the study by Bertrand et al. (2010), which finds a raw gender wage gap of 8.9 log points at the time of graduation for MBA graduates, and to the study by Francesconi and Parey (2018), which finds a raw gender wage gap of 10.3 log points for economics and business graduates.

For the remaining fields of mathematics and natural sciences, humanities and social sciences, and medical studies, the raw gender wage gap is 14.1 log points, 10.2 log points, and 1.5 log points, respectively. The raw gender wage gap for mathematics and natural sciences is similar to the finding of Francesconi and Parey (2018) for the STEM field. The biggest difference with the study of Francesconi and Parey (2018) is in the field of medical studies, where we find no gender wage gap. This is to be expected since the wages of medical graduates (especially doctors) are set by collective bargaining agreements at the beginning of their careers and, therefore, the gender wage gap is very small.

Table 3: Gender Wage Differences at Labor Market Entry by Field of Study

Dependent Variable: Log Daily Wage				
	Economics	Mathematics and	Humanities and	Medical
	and Business	Natural Sciences	Social Sciences	Studies
-	(1)	(2)	(3)	(4)
Add				
Graduation Year	-0.086***	-0.141***	-0.102***	-0.015
	(0.018)	(0.032)	(0.030)	(0.021)
Personal Characteristics	-0.091***	-0.142***	-0.102***	-0.010
	(0.018)	(0.032)	(0.032)	(0.021)
Field of Study	-0.091***	-0.038	-0.124***	-0.030
	(0.018)	(0.031)	(0.032)	(0.019)
Final University Grade	-0.086***	-0.029	-0.121***	-0.029
	(0.018)	(0.032)	(0.032)	(0.019)
Pre-Graduation Characteristics	-0.085***	-0.000	-0.123***	-0.033*
	(0.018)	(0.031)	(0.032)	(0.019)
Occupation FE	-0.063***	-0.005	-0.097***	-0.029*
	(0.018)	(0.030)	(0.032)	(0.017)
Post-Graduation Characteristics	-0.060***	-0.018	-0.097***	-0.029*
	(0.017)	(0.030)	(0.032)	(0.015)
Share of Females	0.295	0.171	0.619	0.475
Average Daily Wage (Euro)	108.4	110.4	84.25	105.18
Individuals	2,167	882	947	1,216

Note: This table shows the gender wage gap at labor market entry, based on OLS model specified in equation 1. The sample consists of graduates with a master's degree or equivalent who work in a full-time job as their first job after graduation and also who have a wage spell 1 year after the first job. The controls variables are added gradually. Row (1) shows the results with only the graduation year as a control. Row (2) adds personal characteristics such as age and (not) having German citizenship and row (3) adds detailed field of study category. Row (4) adds the final university grade and Row (5) adds pre-graduation characteristics: duration of study, location of the final high-school examination, a dummy for apprenticeship and a dummy for working during studying. Row (6) adds 3-digit occupation fixed effects. Row (7) shows the results after adding post-graduation characteristics. These include job search time, job location, 1-digit industry fixed effects, firm size (7 categories), the share of women in firms (3 categories), and the beginning month of the first job. Robust standard errors in parentheses.

***, ** and * denote significance at the 1, 5, and 10% levels.

After controlling for detailed field of study categories, the gender wage gap becomes insignificant for the mathematics and natural sciences. The gap does not exist anymore because women in these fields tend to sort into lower-paid fields such as biology rather than physics. After adding controls, the largest and most significant gender wage gap is observed in those fields typically characterized by a higher female composition and lower earnings.

The finding that the gender wage gap is higher in fields with a higher share of females aligns with the literature on peer effects, which suggests that a higher proportion of females in the classroom may influence females to choose lower-paid occupations (Brenøe and Zölitz, 2020) and positions with a lower wage growth (Zölitz and Feld, 2018), ultimately exacerbating the gender wage gap over time. In addition, there may also be some unobserved labor market characteristics (such as labor demand, discrimination, etc.) that are more relevant for these field groups. For example, if women observe discrimination, they may not choose male-dominated fields (Blau and Kahn, 2017).

To investigate this explanation further, we employ the Oaxaca-Blinder decomposition by field and reveal that the unexplained component of the gender wage gap is most pronounced within the humanities and social sciences field. Specifically, the unexplained part constitutes 82 percent, 23 percent, and 86 percent of the gender wage gap for economics and business, mathematics and natural sciences, and humanities and social sciences, respectively. However, as we mentioned earlier, the unexplained part could also stem from other unobserved characteristics of the labor market that remain beyond the scope of our available data (Blau and Kahn, 2017).

4 Gender Wage Gap in the first Years after Labor Market Entry

After analyzing the gender wage gap at the first job, we now examine how the gender wage gap at the beginning of the career evolves during the first five years after labor market entry. For the analysis, we only include individuals with high labor market attachment who are employed full-time over 5 years after their first job. We decide for a balanced sample to avoid that individuals leave the sample because of family decisions which likely may occur more often for females. We only include full-time employees because our data do not include the exact number of hours worked for part-time employees, which leads to bias in the hourly wage for part-time employees.

This restriction results in a sample size of 2,280 male and 1,205 female graduates, with approximately twothirds of the graduates with high labor market attachment. The blue line in Figure 1 shows the unadjusted gender wage gap for the new sample corresponding to the specification in column (1) of Table 2, with only the year of graduation added as a control. The red line shows the gender wage gap adjusted for pre-graduation

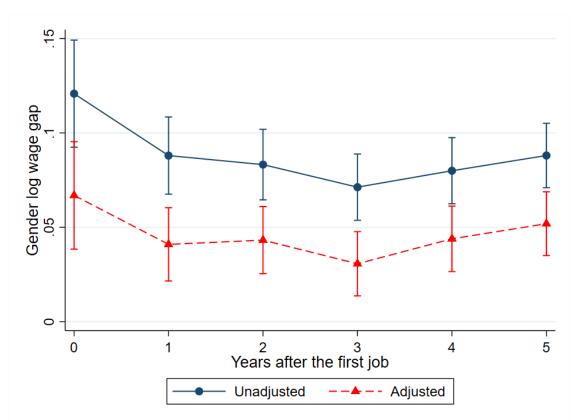


Figure 1: Dynamics of the Gender Wage Gap in the First Years After Labor Market Entry

Note: This figure plots the gender wage gap over the first 5 years after the first job. Each year is estimated separately based on the OLS model specified in equation 1. The sample size is 3,585 (2,280 males, 1,205 females). The sample consists of individuals who have wage spells in a full-time job over the first 5 years after labor market entry. The dependent variable is the log gross daily wage. The unadjusted gender wage gap includes only the year of graduation as a control variable. The adjusted gender wage gap includes personal and pre-graduation characteristics as controls. The personal characteristics include age and German citizenship. The pre-graduation characteristics include duration of study, the location of the final high-school exam, and working while studying. All estimations include the starting month of the first job as a control. e also control for childbirth between years.

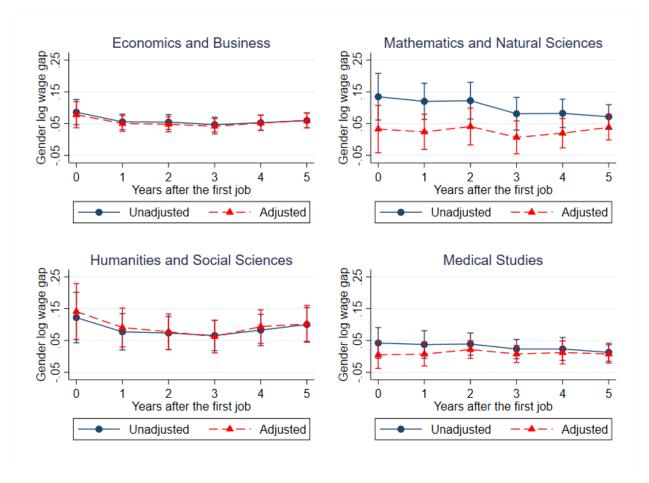
and personal characteristics corresponding to the specification in column (5) of Table 2.

In the first job (0 years after labor market entry in Figure 1), the unadjusted gender wage gap is around 12 log points and drops to around 6.9 log points after including additional controls. Although the sample is somewhat more restricted and the adjusted wage gap is slightly higher, these gaps are consistent with the results presented in Table 2.

Looking at the evolution of the unadjusted gender wage gap over time, we find a sharp decrease of around 3 log points one year after labor market entry from 12 to 9 log points. After this initial drop, the gender wage gap remains relatively stable over the subsequent four years, with a slight increase after year 3. We also observe a similar pattern for the adjusted gender wage gap, which falls by around 4 log points one year after

the first job and remains relatively stable thereafter. These results are robust to focusing on an unbalanced sample of individuals with a first job spell, as documented in Figure B.1 of the Appendix.

Figure 2: Dynamics of the Gender Wage Gap in the First Years After Labor Market Entry by Field of Study



Note: The sample sizes are 1,698, 677, 541 and 624 respectively. The dependent variable is log gross daily wage. The unadjusted gender wage gap includes only graduation year as a control variable. The adjusted gender wage gap contains personal and pre-graduation characteristics as controls. The personal characteristics include age and having German citizenship. The pre-graduation characteristics include duration of study, the place of the high-school final exam and working during studying. All estimations include the beginning month of the first job as a control. Additionally, we control for having a child between the years.

As we have shown that the gender wage gap in the first job varies considerably by field of study, Figure 2 examines whether the evolution of the gender wage gap in the first five years of employment also shows some variation across fields of study. The figure clearly shows that the phenomenon of the gender wage gap narrowing one year after the first job is concentrated among graduates in economics, business, humanities and social sciences, the fields with the largest gender wage gap at labor market entry. Among medical graduates, the gender wage gap is small and does not change significantly over time.

5 Firm and Occupation Mobility in the First Year After Labor Market Entry

After establishing that the gender wage gap narrows in the first 12 months after labor market entry, we want to examine this reduction in more detail. As a starting point, Table 4 analyzes the evolution of the gender wage gap one year after labor market entry for those fields of study with a decrease in the gender wage gap after one year: economics, business, humanities and social sciences (panel A) and those without a decrease: mathematics, natural sciences and medical studies (panel B).

The result in column (1) shows that, controlling for year of graduation and pre-graduation characteristics, women earn 9.8 log points less than men at labor market entry among economics and humanities graduates. Moreover, both female and male wages increase one year after the first job. However, female wages increase by 3.6 log points more on average than male wages. In line with Figure 2, we do not find a decrease in the gender wage gap one year after the first job for graduates in mathematics, natural sciences, and medical studies (panel B).

Previous research shows that firm and/or occupational mobility affects wages and contributes to wage growth (Bartel et al., 2007; Topel and Ward, 1992) and that mobility is especially important in the early stages of a career (Albrecht et al., 2018). For example, Topel and Ward (1992) find that job changes explain more than one-third of wage growth. To analyse whether job changes explain the drop in the gender wage gap, we continue our investigation by separating the sample into graduates who stay in the same firm and/or occupation (column 2 of Table 4), those who change occupation but remain in the same firm (column 3), those who change firm but remain in the same occupation (column 4), and those individuals who change firm and take up a new occupation (column 5).²⁰

Panel A of Table 4, which focuses on graduates in economics, business, humanities and social sciences, demonstrates that the gender wage gap does not decrease significantly for those individuals who either stay in the same firm and/or occupation one year after starting their first job (columns 2-4). In contrast, female firm and occupation changers increase their wages on average by around 19 log points more than their male counterparts (column 5). This increase has to be seen in the context that males who change either firm, occupation, or both benefit from these changes by around 25-27 log points, while the stayers increase their wages by only 10 log points. In addition, the initial gender wage gap is larger for individuals who change firms, occupations, or both than for individuals who remain in their position. The group that changes firm and occupation has the largest initial gap (almost 34 log points). Interestingly, the allocation of men and women to the four groups is relatively similar, so differences in shares do not seem to explain the different

 $^{^{20}\}mathrm{We}$ define an occupation change if the 3-digit occupation code changes.

Table 4: Gender Wage Gap by Job Change Status

Dependent Variable: Log	Daily Wage				
	Pooled	Stayers	Only Firm Changers	Only Occupation Changers	Firm and Occupation Changers
_	(1)	(2)	(3)	(4)	(5)
Panel A: Economics, I	Business, Hur	nanities and	Social Science	es	
1 Year After \times Female	0.036*** (0.012)	-0.000 (0.008)	$0.038 \\ (0.063)$	$0.061 \\ (0.083)$	0.193*** (0.070)
1 Year After	0.130*** (0.007)	0.098*** (0.006)	0.252*** (0.042)	0.248*** (0.063)	0.271*** (0.039)
Female	-0.098*** (0.015)	-0.056*** (0.014)	-0.156** (0.068)	-0.139* (0.074)	-0.340*** (0.069)
Share of female	1	0.737	0.115	0.101	0.131
Share of male R-squared	$\begin{array}{c} 1 \\ 0.240 \end{array}$	$0.796 \\ 0.274$	$0.091 \\ 0.306$	$0.058 \\ 0.323$	$0.105 \\ 0.328$
Individuals	$3{,}114$	$2,\!407$	267	194	312
Panel B: Mathematics	, Natural Sci	ences and M	edical Studie	s	
1 Year After \times Female	-0.001 (0.011)	$0.006 \\ (0.009)$	-0.010 (0.042)	-0.140 (0.132)	-0.088 (0.095)
1 Year After	0.150*** (0.008)	0.127*** (0.007)	0.182*** (0.031)	0.341*** (0.098)	0.386*** (0.054)
Female	-0.019 (0.016)	-0.024 (0.016)	$0.008 \\ (0.034)$	$0.074 \\ (0.171)$	$0.049 \\ (0.096)$
Share of female	1	0.783	0.144	0.032	0.081
Share of male R-squared	$\frac{1}{0.393}$	$0.838 \\ 0.399$	$0.086 \\ 0.624$	$0.037 \\ 0.499$	$0.070 \\ 0.382$
Individuals	2,098	1,718	204	60	137

Note: This table shows the gender wage gap at labor market entry by job change status, based on OLS model specified in equation 1. The sample consists of graduates with a master's degree or equivalent who work in a full-time job as their first job after graduation and also who have a wage spell 1 year after the first job. The dependent variable is the log gross daily wage at the first job. The estimations include personal and pre-graduation characteristics as controls. The personal characteristics include age and having German citizenship. The pre-graduation characteristics include duration of study, the place of the high-school final exam, and working during studying. All estimations include the beginning month of the first job as a control. Robust standard errors in parentheses. ***, ** and * denote significance at the 1, 5, and 10% levels.

evolution of the gender wage gap after one year.

For mathematics, natural sciences, and medical studies, we find no reduction in the gender wage gap in the first 12 months after starting the first job, and we observe no initial gap for any of the changer groups (panel B of Table 4). However, even for these fields, the results show that the movers have the highest wage growth, between 18 and 39 log points. Overall, the table shows that women in economics, business, humanities, and social science benefit more than males from a complete new start after their first job, which includes a change of firm and occupation. This new start drives the observed reduction in the gender wage gap within one year after the first job.

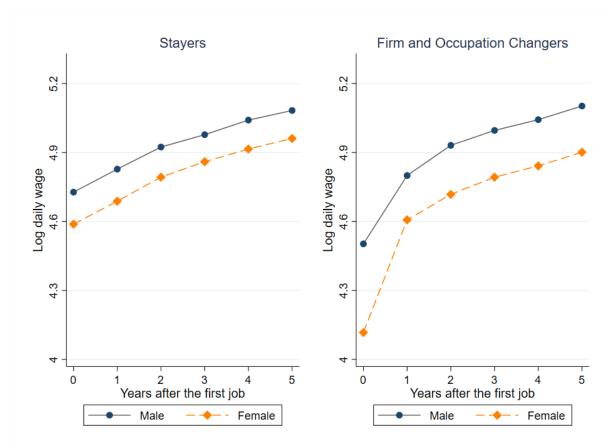


Figure 3: The Dynamic of Wages for Job Stayers and Job and Firm Changers by Gender

Note: The figures plot the dynamic of wages over 5 years after the first job for stayers (left panel) and firm and occupation changers (right panel). The sample sizes are 312, and 137 respectively. The dependent variable is log gross daily wage. The graduation year, personal and pre-graduation characteristics added as controls. The personal characteristics include age and having German citizenship. The pre-graduation characteristics include duration of study, the place of the high-school final exam, and working during studying. All estimations include the beginning month of the first job as a control. Additionally, we control for having a child between the years.

As a next step in our analysis, Figure 3 shows the dynamics of wages over 5 years after the first job for stayers and those graduates who change both occupation and firm. Confirming the result of Table 4, female and male movers initially earn lower wages on average than stayers. However, the wage difference between stayers and movers is higher for females than for males, due to the very low entry wages of those females who later change both occupation and firm. To summarize, women with low entry wages appear to correct their low wages more than men by changing their firm and occupation within one year of entering the labor

6 Why is Switching Firm and Occupation in the Beginning of the Career more Beneficial for Females than Males

In this section, we use our rich administrative data to investigate why women benefit more than men from changing firms and occupations after their first job after graduation. We use two estimation approaches to conduct our analysis. First, we estimate whether women who change firms and occupations differ from men who change firms and occupations in terms of their demographic characteristics, university outcomes, and characteristics of their first job, and whether these gender differences differ from those of stayers. Second, we estimate whether firm and occupation characteristics change differently for males and females after job transitions. In the end of the section, the results of the two estimation approaches are discussed with respect to common theoretical explanations for the gender wage gap.

Table 5 compares gender differences in personal and pre-graduation characteristics (both of which are constant over time), between firm and occupation changers and stayers. In Table 6, Panel A compares gender differences in the characteristics of the first job for individuals who change firms and occupations with those who stay in the same position. Panel B of Table 6 examines gender differences in the characteristics of the first and subsequent job one year later for occupation and firm changers. Table C.5 in the Appendix shows the mean values of all variables from Tables 5 and 6 by gender and the corresponding mean gender differences for stayers and firm and occupation changers.

The first row of Table 5 reports the interaction coefficients between the female variable and a dummy for firm and occupation change. The coefficients in the first row indicate that none of the personal characteristics, pre-graduation characteristics, or job search characteristics differ more between males and females who change firms and occupations than between males and females who stay in the same firm and occupation.

However, Table 6 demonstrates that out of several characteristics, job-education mismatch and occupational rank (Columns 6-11) are two characteristics that differ between males and females who switch occupation and firm in the first job and develop differently after the job change. We separate job-education mismatch into two types: horizontal and vertical mismatch. Horizontal mismatch is the field-occupation mismatch, where the employee's field of study does not match the field required for the job. A vertical mismatch is a skill mismatch where the skill level of the employee's qualification does not match the requirements of the job. Since our sample includes highly skilled university graduates, only jobs for which university graduates are overqualified are defined as vertical mismatches.²¹ Occupation rank is a measure

 $^{^{21}}$ There is a large body of literature showing that both vertical and horizontal mismatches have a negative effect on wages

Table 5: Personal and Pre-Graduation Characteristics of Firm and Occupation Changers

Dependent variables:											
	Per	rsonal		Pre-Graduation							
	Chara	cteristics		Chara	acteristics			Characteristic			
	Age at the first job (1)	non-German (2)	Duration of Study (3)	Working During Studying (4)	Apprent. (5)	Origin is Bayern (6)	Final Uni. Grade (7)	Duration of Job Search (8)			
Female × Firm and Occupation Changers	0.058 (0.269)	-0.021 (0.017)	0.226 (0.171)	-0.012 (0.053)	-0.019 (0.037)	-0.031 (0.041)	0.071 (0.073)	-0.342 (0.366)			
Firm and Occupation Changers	0.444*** (0.156)	0.003 (0.010)	$0.063 \\ (0.096)$	$0.002 \\ (0.037)$	0.054** (0.026)	$0.021 \\ (0.027)$	$0.026 \\ (0.045)$	-0.035 (0.246)			
Female	-0.803*** (0.082)	0.026*** (0.007)	-0.053 (0.053)	0.126*** (0.018)	-0.010 (0.010)	0.034** (0.014)	-0.206*** (0.023)	0.141 (0.131)			
Means of dependent variable	27.309	0.023	5.515	0.719	0.074	0.858	2.115	3.859			
R-squared	0.042	0.006	0.003	0.018	0.004	0.002	0.030	0.001			
Individuals	2,719	2,719	2,719	2,719	2,719	2,719	2,719	2,719			

Note: The table documents gender differences in personal, pre-graduation, and finding first job characteristics, between firm and occupation changers and stayers. The sample size is 2,719, including stayers (column 2, Table 4) and firm and occupation changers (column 5, Table 4). Each column is a different estimation which is time invariant. The estimations include a female dummy, a dummy variable for firm and occupation changer dummy (= 1 if an individual changes her/his firm and occupation within 1 year after the labor market entry, = 0 if an individual stays at the same firm and occupation), and an interaction of these dummies. All estimations include the beginning month of the first job as a control. Robust standard errors in parentheses.

***, ** and * denote significance at the 1, 5, and 10% levels.

that ranks occupations by their average wage.²²

Specifically, columns (6) and (7), panel B of Table 6 show that female job changers are more likely than male job changers to work in horizontally (by 12.7 percentage points) and vertically (by 13.9 percentage points) mismatched first jobs after graduation. However, female job changers reduce the frequency of vertical mismatch one year after the first job by 13.1 percentage points compared to men.

After correcting the vertical mismatch, one might expect that women will receive a higher wage as soon as they correct their mismatch. Although column (9), panel B of Table 6 shows that female job changers work in lower-ranked occupations in their first job after graduation, they do not move to (significantly) higher-paid occupations on average compared to men (first row, column (9), panel B).

However, Figure B.2 in the Appendix shows that the occupational rank distributions of male and female job changers are quite different at the first job, as females are less likely to work in higher-paid occupations and more likely to work in lower-paid occupations than males. After the job change, the distributions of males and females converge, especially in the lower tail, as females predominantly move from lower-paid occupations to higher-paid occupations. In line with the convergence in the lower tail, Panel B, Column (10) of Table 6 shows that after a firm and occupation change, women reduce the probability of being in the bottom decile of ranked occupations relative to men by 11.5 percentage points.

⁽Wolbers, 2003; Robst, 2007; Boudarbat and Montmarquette, 2009; Heijke et al., 2003).

²²Average wages within 3 digit occupation codes are calculated using the SIAB data which is 2% of the ieb data.

Table 6: Job Characteristics of Firm and Occupation Changers

Dependent variables:											
	Median Daily Log Wage of Full-time Employees	Share of Part-time Employees	Share of High Qualified Employees	Share of Women in a Firm	Log Firm Size	Horizontal Mismatch	Vertical Mismatch	Horizontal or Vertical Mismatch	Occupation Rank	Occupation Rank < Quantile 10	Occupation Rank > Quantile 90
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
Panel A: First Job Characte	eristics of Firm and	Occupatio	n Changers								
Female \times Firm and Occupation Changers	-0.086* (0.049)	-0.007 (0.025)	-0.013 (0.032)	0.044 (0.027)	-0.204 (0.252)	0.008 (0.056)	0.117** (0.058)	0.070 (0.053)	-22.371* (11.815)	0.098** (0.044)	-0.025 (0.034)
Firm and Occupation Changers	-0.120*** (0.028)	0.018 (0.016)	-0.063*** (0.021)	-0.006 (0.017)	-0.341** (0.166)	0.119*** (0.035)	0.122*** (0.039)	0.150*** (0.037)	-7.680 (7.488)	0.017 (0.024)	-0.002 (0.026)
Female	-0.041*** (0.014)	-0.035*** (0.008)	-0.016 (0.011)	0.071*** (0.009)	-0.102 (0.090)	0.122*** (0.017)	$0.005 \\ (0.021)$	0.051** (0.021)	-15.238*** (3.602)	0.013 (0.012)	-0.036*** (0.013)
Mean of Dependent Variables	4.617	0.781	0.372	0.482	5.099	0.219	0.506	0.594	226.031	0.104	0.104
R-squared	0.029	0.008	0.009	0.035	0.005	0.032	0.014	0.018	0.016	0.009	0.004
Individuals	2,719	2,719	2,719	2,719	2,719	2,719	2,719	2,719	2,719	2,719	2,719
Panel B: Jobs Characteristic	cs Before and After	the Job C	hange within	Firm and	Occupation	Changers					
Female \times 1 year after	$0.046 \\ (0.053)$	-0.016 (0.029)	0.011 (0.036)	-0.007 (0.032)	-20.260 (483.871)	-0.060 (0.048)	-0.131* (0.079)	-0.153** (0.071)	$21.816 \\ (14.263)$	-0.115** (0.049)	-0.035 (0.049)
1 year after	0.123*** (0.031)	$0.009 \\ (0.018)$	0.033 (0.024)	-0.022 (0.020)	377.790 (364.812)	-0.028 (0.034)	$0.006 \\ (0.055)$	0.028 (0.049)	13.199 (9.661)	-0.040 (0.030)	0.057 (0.040)
Female	-0.161*** (0.047)	-0.018 (0.024)	-0.051* (0.030)	0.119*** (0.027)	-488.739** (235.755)	0.127** (0.055)	0.139** (0.055)	0.122** (0.050)	-38.079*** (11.476)	0.102** (0.044)	-0.061* (0.033)
Means of dependent R-squared	4.558 0.113	0.792 0.093	0.338 0.059	0.484 0.085	4.829 0.035	0.296 0.065	0.619 0.060	$0.721 \\ 0.044$	223.341 0.054	$0.128 \\ 0.047$	0.114 0.053
Individuals	624	624	624	624	624	624	624	624	624	624	624

Note: Panel A documents the gender differences in the first job characteristics among firm and occupation changers and stayers. The sample size is 2,719, including stayers (column 2, Table 4) and firm and occupation changers (column 5, Table 4). Panel B uses firm and occupation characteristics as dependent variables (each column is a different estimation) which are time-variant, i.e., may vary before and after the job change. The sample size is 312 including only firm and occupation changers (column 5, Table 4). The estimations include a female dummy, a dummy variable for one year after the job change (=0 for the first job, =1 for the new job one year after the first job) and an interaction of these dummies. All estimations include the beginning month of the first job as a control. Robust standard errors in parentheses. ***, ** and * denote significance at the 1, 5, and 10% levels.

Overall, the results of Tables 5 and 6 demonstrate that females are more likely than males to start in an occupation which is in the bottom tail of the occupation rank distribution and change to higher-paid occupations if they switch occupation and firm. Moreover, they start in occupations where they are overqualified and correct this vertical mismatch by changing occupation and firm. As our data show that correcting the vertical mismatch at the first job explains the decrease in the gender wage gap, the question arises why women need to change firms as well as occupations to correct the mismatch. We now test common hypotheses in the gender wage gap literature which may explain our finding.

Different Types of Discrimination

A first potential explanation for the decline in vertical mismatch and the gender wage gap within 1 year after the first job is that firms discriminate against women at the hiring stage, when employers do not have sufficient information about the productivity of new hires (Altonji and Blank, 1999). As a first type of discrimination, "screening discrimination", Pinkston (2003) documents that the productivity signals that employers receive from females are noisier than males. Because of this, productivity signals at the hiring stage have a smaller or no effect on women's wages, while they have a larger effect on men's wages. In our case of university graduates, since the employer is able to observe the CVs of the applicants, the final university grades may be the only signal for the employer. We would expect that men with higher grades do not change jobs because they already have a good match in their first job. On the contrary, women with higher grades may receive a mismatch compared to their male counterparts at the beginning of their careers and thus change jobs in order to correct the mismatch. However, our results show that female movers and stayers have better grades on average than their male counterparts, both female and male movers have worse grades than stayers. Nevertheless, the difference in the gender gap between movers and stayers is insignificant, as the interaction term is insignificant (column (7), Table 5).

The literature on the gender wage gap suggests that females may face statistical discrimination in the labor market. This means that employers may expect lower productivity from females and hire them for less suitable jobs. Consequently, conditional on being hired, females would work in more mismatched jobs and receive lower initial wages at the beginning of their careers. However, over time, as employers learn about the actual productivity of new hires ("employer learning"), such mismatches could be corrected, resulting in higher wages for women ((Altonji and Blank, 1999; Altonji and Pierret, 2001; Pinkston, 2003)). We examine whether discrimination could explain the differences in how men and women benefit from changing firms and occupations. If females were statistically discriminated against, we would expect the gender wage gap to likely narrow, not only for those who change firms and occupations but also for stayers. Since we do not find a significant reduction in the gender wage gap for stayers, statistical discrimination is unlikely to explain

the differential returns to changing occupations and firms.

Another form of discrimination suggested by the literature is taste-based discrimination, where employers pay women lower wages to compensate for their (or co-workers') disutility.²³ The higher mismatch and the lower initial wage of female movers relative to male movers (Table 6) may indicate some form of taste-based discrimination (Becker, 1971). However, if firms discriminate against women, switching to non-discriminatory firms should be sufficient for women to improve their wages relative to men, while an additional change in occupation should not be necessary. As Table 4 shows, this is not the case, as the gender wage gap does not narrow significantly for those who only change firms. Furthermore, if taste-based discrimination explains the gender wage gap, we should observe that women who change firms will move to firms with more women, as these firms typically discriminate less. Contrary to this hypothesis, women who change their firm and occupation are more likely to work in firms with a higher share of women in their first job (panel A of Table 6). Moreover, our estimation results show that women do not switch to firms with a higher female share compared to men (column (4), panel B of Table 6).

Risk aversion & Job Searching Time

The literature shows that risk aversion may be an important component of early career job search (Cortés et al., 2021) job transitions, with women typically having higher level of risk aversion compared to men (Niederle and Vesterlund, 2007). More risk-averse women may have lower reservation wages at the beginning of their careers (Pissarides, 1974; Feinberg, 1977; Acemoglu and Shimer, 1999; Cox and Oaxaca, 1992; Pannenberg, 2010) and consequently accept job offers relatively earlier, even if the job pays less and is not a good match (Cortés et al., 2021). However, these women may not be satisfied with lower wages and mismatch and change jobs when they find a higher-paid and better match job. In this case, we would expect women to spend less time searching for a job after graduation than men, and women who find a job more quickly would be more likely to change jobs. Column 8 of Table 5 shows that although job changers find a job slightly earlier than stayers, there is no significant gender difference in job search duration for stayers and job changers in the fields of economics, business, humanities and social sciences. In addition, wages do not decrease significantly with the duration of the job search.

Job Amenities

As an alternative explanation, a growing literature shows that women prefer non-wage job amenities such as flexibility or meaning, relevance, or responsibility of the occupation more than men who have a higher

 $^{^{23}}$ Becker (1964) shows in the model that firms practicing taste-based discrimination cannot survive in the competitive market in the long run.

preference for wages (Goldin and Katz, 2011; Goldin, 2014; Flabbi and Moro, 2012; Brenøe and Zölitz, 2020). Changing preferences for certain job attributes may also be a mechanism for job change. On the one hand, females may change to more flexible jobs in anticipation of having children in the future. However, these changes may not lead to higher wage gains. On the other hand, at the beginning of their careers, women may prefer lower-paying jobs with a vertical mismatch to compensate for some job amenities, but over time they change their preferences and switch to higher-paying and less flexible jobs.

Assuming that larger firms have more flexible work arrangements (Albrecht et al., 2018), we test whether women switch to larger firms to enjoy more flexibility. Our findings do not show that females are more likely to sort into larger or smaller firms compared to males as a result of a job change (column 5, Table 6). However, our data do not cover other proxy for job amenities, such as the meaning of jobs, schedule adaptability, or telecommuting (De Schouwer and Kesternich, 2022). Therefore, we believe that changing preferences for job amenities may still be an important explanation for why women reduce their vertical mismatch and increase their wages when they switch firms and occupations.

Assuming that larger firms have more flexible work arrangements (Albrecht et al., 2018), we test whether women switch to larger firms to enjoy more flexibility. However, we do not find that females are more likely than men to sort into larger or smaller firms compared to males as a result of a job change (column 5, Table 6). However, our data do not cover other proxy for job amenities, such as the meaning of jobs, schedule adaptability, or telecommuting (De Schouwer and Kesternich, 2022). Therefore, we believe that changing preferences for job amenities may still be an important explanation for why women reduce their vertical mismatch and increase their wages when they switch firms and occupations.

7 Conclusion

Although a large number of studies have investigated the gender wage gap, the existence and potential explanations for early career gender wage differences remain unclear. This paper analyzes the gender wage gap among graduates of a German university with a master's degree or equivalent at the beginning of their careers and over the first years after labor market entry. We take advantage of a unique dataset that links administrative data on graduates of a German university with employment registers of the German social security system. This dataset includes extensive information on students' socio-demographic characteristics, their educational and labor market outcomes, and the precise timing of graduation, labor market entry, and any job changes.

We find a significant gender wage gap among university graduates in their first job, which persists even after controlling for an extensive set of controls. The largest gender wage gap is observed among humanities and social sciences graduates, where the share of females is highest and the average daily wage is lowest. We find no significant gender differences in the wages of mathematics, natural sciences, and medical graduates in their first job after graduation. Moreover, in contrast to previous studies, we find an immediate decrease in the gender wage gap one year after labor market entry, which remains almost stable thereafter.

Further analysis shows that the decline in the gender wage gap is concentrated among individuals who change firms and occupations after their first job with a degree in economics, business, humanities, and social sciences. As an explanation for the decrease in the gender wage gap, we also show that female graduates are more likely to start their careers in jobs for which they are overqualified and subsequently correct this skill mismatch, leading to an increase in their wages. Correcting the mismatch is costly for females, which may be an additional explanation for the wage gap later in their career.

Universities have a important opportunity to mitigate the risk of future skills mismatches by implementing counseling interventions. These interventions can provide valuable information on effective job search strategies and potential wage losses due to skill mismatches, particularly for female students. Our study also highlights significant differences in labor market entry and early career paths depending on the chosen field of study. For this reason, counseling programs that help students understand their career prospects should be tailored specifically to each field of study. By implementing such counseling, universities can provide graduates with the insight they need to navigate the dynamic labor market successfully.

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A Appendix: Data and Additional Descriptive Statistics

This section describes all data preparation steps before the main analysis. If there are multiple contemporaneous employment spells for an individual, we use the main employment spell and exclude the remaining employment spells. The main employment spell as defined by the IAB is the spell with the longest job duration and the highest daily wage. Furthermore, in order to eliminate errors in daily wages for full-time employees, we follow the literature and disregard daily wages below 10 Euro from the main sample (Dustmann et al., 2009; Bruns, 2019).

One issue to consider is that wages in the IEB dataset are only reported up to the social contribution threshold, as the information on wages is obtained from the German Social Security report. Thus, wages above the social contribution limit are right-censored. However, since we analyze the gender wage gap at the beginning of the career, there are not many censored wages in our restricted sample; censored wages account for only 1.3 percent for the first job and around 4.7 percent a year after the first job, with a small increase in subsequent years after graduation.

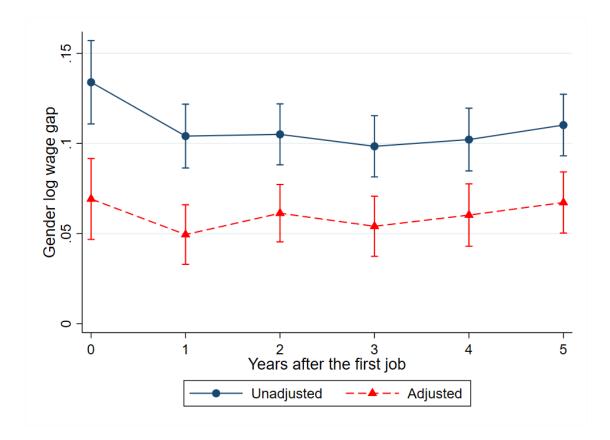
Moreover, working hours are not recorded in the IEB dataset, as only information is only available on whether a person works in a full-time or part-time job (working more or less than 30 hours per week). For this reason, we only focus on graduates who have a full-time job in their first job after graduation. An individual is considered a full-time employee if he or she works more than 35 hours per week.

We include occupational categories using 3-digit occupational codes (KldB 1988) in the estimations. Since the occupational structure has changed over time, the Federal Employment Agency introduced a new classification (KldB 2010) in 2011 that better fits the current German occupational structure. Since the new classification is more detailed (5-digit) than the old one, there is a significant increase in missing values in the occupation variable in 2011 (Antoni et al., 2016). To address this issue, we fill in the missing values in 2011 by keeping the last occupation spell before the change in occupational classification, and replacing it with latter missing spells, if the place of residence and work, industry code and establishment ID does not change. Following this procedure, the number of missing values in the occupation code reduces significantly for 2011.

Moreover, childbirths are not directly observed in our linked data. However, we can identify family-related interruptions in employment based on the IEB data by applying a reliable approach developed by Müller et al. (2017). This method allows us to identify the timing of employment interruptions and (approximately) the timing of childbirths.

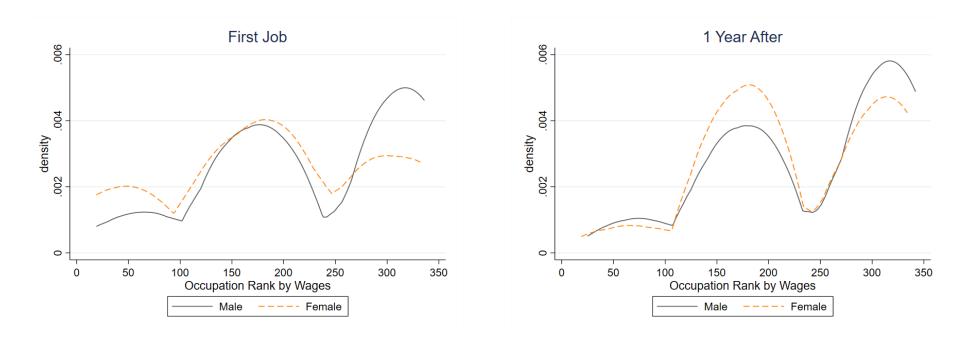
B Appendix: Figures

Figure B.1: Unbalanced Sample: Dynamics of the Gender Wage Gap over Years After Labor Market Entry



Note: The sample size is 3,409 (2,262 male, 1,147 female). The sample consists of graduates who work in a full-time job as their first job after graduation. The dependent variable is log gross daily wage. The unadjusted gender wage gap includes only graduation year as a control variable. The adjusted gender wage gap contains personal and pre-graduation characteristics as controls. The personal characteristics include age and having German citizenship. The pre-graduation characteristics include duration of study, the place of the high-school final exam and working during studying. All estimations include the beginning month of the first job as a control. Additionally, we control for having a child between the years.

Figure B.2: Distribution of Occupations for Female and Male Job Changers



Note: The figure shows the distribution of male and female job changers across occupations. The sample size is 312 including only firm and occupation changers (column 5, Table 4). The plot on the left shows the distribution before the job change (at the first job) and the plot on the right after the job change (1 year after the first job) by gender. The x-axis plots occupations ranked by average earnings from the lowest paying occupation to the highest paying occupation. The ranking of occupations by average earnings is calculated from the SIAB dataset, a 2 percent sample of the entire IEB dataset.

C Appendix: Tables

Table C.1: Field of Study Categorization

Field of Study (Combined)	Field of Study (Detailed)
Economics and Business	Economics Business and Management
Mathematics and Natural Sciences	Information Systems Mathematics and Computer Science Physics Chemistry Biology
Humanities and Social Sciences	Geography Theology History, Archaeology and Humanities Languages, Literature and Culture Philosophy, Sociology and Political Psychology Education and Sport
Medical Studies	Medicine Dental Medicine Pharmacy

Table C.2: Descriptive Statistics - Sample Comparison to Population

	Estimation Sample	Population
Share of Females ¹	0.498	0.457
Economics and Business	0.415	0.484
Mathematics and Natural Sciences	0.182	0.353
Humanities and Social Sciences	0.761	0.713
Medical Studies	0.518	0.581
Final High-school GPA ²	2.20	2.15
Economics and Business	2.36	2.45
Mathematics and Natural Sciences	2.17	1.99
Humanities and Social Sciences	2.31	2.27
Medical Studies	1.81	1.88
University GPA ³	2.04	2.02
Graduation Age ¹	27.37	27.80
Share of Non-German Graduates ³	0.217	0.110

Note: This table presents summary statistics on the characteristics of graduates before and after graduation and compares them with official register data and representative survey data from other studies. Our estimation sample consists of graduates with a master's degree or equivalent working full-time in their first job after graduation.

- 1. The reference year for both data is 2010. Source: Federal Statistical Office, 2011.
- 2. The reference year is WS 2006/2007 in the survey data from Simeaner (2014) and 2007 in our sample.
- 3. The years are pooled for 1993-2009 in the survey data from Francesconi and Parey (2018) and pooled for 1995-2010 in our data.

Table C.3: Gender Wage Gap at Labor Market Entry

Dependent Variable: Log Daily	Wage								
		Addi	tional						
			Personal	and Pre-G	raduation		Post-Graduation		
			С	haracterist	ics		Charac	teristics	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	
Female	-0.271***	-0.201***	-0.212***	-0.109***	-0.101***	-0.097***	-0.045***	-0.045***	
	(0.018)	(0.014)	(0.014)	(0.014)	(0.014)	(0.014)	(0.013)	(0.012)	
Graduation year	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
Full-time employment	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
Personal characteristics	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
Field of study FE	No	No	Yes	Yes	Yes	Yes	Yes	Yes	
Final university grade	No	No	No	Yes	Yes	Yes	Yes	Yes	
Pre-graduation characteristics	No	No	No	No	Yes	Yes	Yes	Yes	
Occupation FE	No	No	No	No	No	Yes	Yes	Yes	
Post-graduation characteristics	No	No	No	No	No	Yes	No	Yes	
R-squared	0.050	0.417	0.424	0.493	0.498	0.501	0.630	0.652	
Individuals	10,149	10,149	10,149	10,149	10,149	10,149	10,149	10,149	

Note: This table shows the gender wage gap at labor market entry, based on OLS model specified in equation 1. The sample consists of all individuals who work in any kind of jobs after the graduation. The dependent variable is the log gross daily wage at the first job. The controls variables are added gradually. Column (1) shows the results with only the graduation year as a control. Column (2) adds a full-time indicator. Column (3) adds personal characteristics such as age and (not) having German citizenship and column (4) adds field of study (17 categories). Column (5) adds the final university grade and column (6) adds pre-graduation characteristics: duration of study, location of the final high-school examination, a dummy for apprenticeship and a dummy for working during studying. Column (7) adds 3-digit occupation fixed effects. Column (8) shows the results after adding post-graduation characteristics. These include job search time, job location, 1-digit industry fixed effects, firm size (7 categories), the share of women in firms (3 categories), and the beginning month of the first job. Robust standard errors in parentheses. ***, ** and * denote significance at the 1, 5, and 10% levels.

Table C.4: Oaxaca-Blinder Decomposition

Dependent Variable: Log D	aily Wage						
	F	irst Job	1 Y	1 Year After			
Mean of male daily wage		4.691		4.832			
Mean of female daily wage		4.570		4.733			
Raw gender wage gap	0.	.121***	0.	099***			
	\log	Percent of	\log	Percent of			
-	points	gap explained	points	gap explained			
Total explained	0.052***	43	0.050***	50			
Total unexplained	0.069***	57	0.049***	50			
Explained by:							
Graduation year	0.005	4	0.006**	6			
Age	0.005**	4	0.004**	4			
Non-German	-0.002**	2	-0.002**	2			
Field of study	0.048***	40	0.044***	44			
Duration of study	-0.000	0	-0.000	0			
Working during studying	-0.001	1	-0.001	1			
Apprenticeship	-0.000	0	-0.000	0			
Grade	-0.006**	5	-0.004***	4			
Place of the final	0.004*	3	0.004	4			
high-school examination							

Note: This table shows Oaxaca-Blinder decomposition results. Decomposition methods allow to split the mean wage gap into an explained component (due to differences in characteristics) and an unexplained component (due to differences in returns to these characteristics). The decomposition model used in this table is the aggregate two-fold decomposition. See Fortin et al. (2011) for detailed information on the methodology, and Jann et al. (2008) for a STATA application.

Table C.5: Descriptive Statistics - Stayers and Firm and Occupation Changers

Dependent variables:							
	Stayers			Firm and Occupation Changers			
	Male	Female	Male-Female	Male	Female	Male-Femal	
Panel A: Economics, Business, Humanities and Social Sciences	1						
Age at the First Job	27.467	26.674	0.793***	27.899	26.990	0.908***	
non-German	0.014	0.038	-0.024***	0.017	0.015	0.003	
Duration of Study	5.527	5.463	0.064	5.576	5.716	-0.140	
Working During Studying	0.659	0.784	-0.124***	0.676	0.787	-0.110**	
Apprenticeship	0.074	0.063	0.011	0.133	0.096	0.037	
Origin is Bayern	0.843	0.876	-0.033**	0.867	0.875	-0.008	
Final Uni. Grade	2.187	1.976	0.210***	2.223	2.082	0.140**	
Duration of Job Search	3.388	3.518	-0.129	3.110	3.010	0.100	
Median Daily Log Wage of Full-time Employees in a Firm	4.662	4.622	0.041***	4.543	4.416	0.126***	
Share of Part-time Employees in a Firm	0.791	0.757	0.035***	0.809	0.767	0.042*	
Share of High Qualified Employees in a Firm	0.395	0.379	0.016	0.332	0.302	0.030	
Share of Women in a Firm	0.452	0.524	-0.071***	0.446	0.561	-0.115***	
Log Firm Size	5.222	5.119	0.102	4.881	4.575	0.306	
Horizontal Mismatch Occupation	0.148	0.270	-0.122***	0.267	0.397	-0.130**	
Vertical Mismatch	0.469	0.475	-0.005	0.591	0.713	-0.122**	
Horizontal or Vertical Mismatch	0.538	0.588	-0.051**	0.688	0.809	-0.121**	
Occupation Rank	236.061	220.823	15.238***	228.381	190.772	37.609***	
Occupation Rank < Quantile 10	0.090	0.118	-0.029**	0.102	0.213	-0.111***	
Occupation Rank > Quantile 90	0.121	0.085	0.036***	0.119	0.059	0.060*	
Observations	1503	904	0.030	176	136	0.000	
Panel B: Mathematics, Natural Sciences and Medical Studies							
Age at the First Job	27.476	26.993	0.482***	28.120	27.027	1.093***	
non-German	0.014	0.018	-0.004	0.024	0.060	-0.036	
Duration of Study	5.692	5.870	-0.004 -0.178**	5.938	6.008	-0.030 -0.070	
Working During Studying	0.660	0.661	-0.178	0.747	0.008 0.720	0.027	
Apprenticeship	0.043	$0.001 \\ 0.041$	0.002	0.133	$0.720 \\ 0.120$	0.027 0.013	
Origin is Bayern	0.043 0.914	0.041 0.855	0.059***	0.133	0.120 0.820	0.013	
Origin is Bayern Final Uni. Grade	0.914 1.858	$\frac{0.855}{2.007}$	-0.149***	0.880 1.835	0.820 1.845	-0.010	
			-0.149**		$\frac{1.845}{3.462}$		
Duration of Job Search Median Daily Log Wage of Full-time Employees in a Firm	$3.330 \\ 4.576$	3.671	0.077***	2.748	$\frac{3.462}{4.236}$	-0.714 0.070	
		4.499		4.306			
Share of Part-time Employees in a Firm	0.689	0.609	0.080*** 0.075***	0.677	0.625	0.053	
Share of High Qualified Employees in a Firm	0.317	0.242		0.219	0.198	0.021	
Share of Women in a Firm	0.552	0.714	-0.162***	0.522	0.755	-0.233***	
Log Firm Size	5.594	5.690	-0.095	3.934	4.046	-0.111	
Horizontal Mismatch Occupation	0.196	0.084	0.112***	0.425	0.260	0.165*	
Vertical Mismatch	0.121	0.100	0.021	0.402	0.460	-0.058	
Horizontal or Vertical Mismatch	0.263	0.139	0.124***	0.575	0.500	0.075	
Occupation Rank	285.685	290.981	-5.296	227.080	199.800	27.280	
Occupation Rank < Quantile 10	0.116	0.098	0.018	0.172	0.180	-0.008	
Occupation Rank > Quantile 90	0.372	0.584	-0.212***	0.172	0.160	0.012	
Observations	1148	570		87	50		

Note: This table shows summary statistics of graduates' personal, pre-graduation, post-graduation and first job characteristics of stayers and firm and occupation changers. The sample consists of graduates with a master's degree or equivalent who work in a full-time job as their first job after graduation and also who have a wage spell 1 year after the first job. ***, ** and * denote significance at the 1, 5, and 10% levels.