15.458 Prject C: Portfolio and Risk Management

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- 1. **Beta:** Use the data from "crispy04" to examine the CAPM beta between "the market" and your favorite Dow stock for the period 1/1/2008 to 12/31/2009. Pick a stock that was in the DJIA for the entire period and then create and compare the following:
 - (a) A time series plot of the 3-month rolling beta of the stock vs. the S&P 500. (Use the stored, pre-computed values.) The primary tables to use are {dailyvalue, benchvalue}. Tables with supplementary attributes that may be useful include {dim_equity, bench, indexmember, and dim_numtype}.

Answer. We picked Goldman Sachs as our favorite Dow stock. Below is the 3 month-rolling beta of the stock vs the S&P 500, \Box

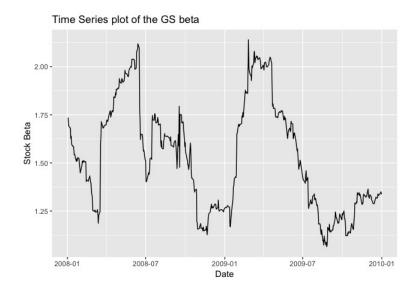


Figure 1: Times Series Plot of the GS Beta

(b) A scatter plot of 1-day stock returns vs. the SPTR (S&P 500 total return index) 1-day returns along with a line of best fit (OLS regression) and its slope equation. How does this result relate to, and compare with, the time series plot above?

Answer. Below is the summary of OLS regression followed by the scatter plot. We see that the 1-day stock returns and the SPTR are positively correlated.

	Estimate	Std. Error	t value	$\Pr(> t)$
Intercept	0.0002	0.0013	0.19	0.8525
Slope	1.5052	0.0588	25.61	0.0000



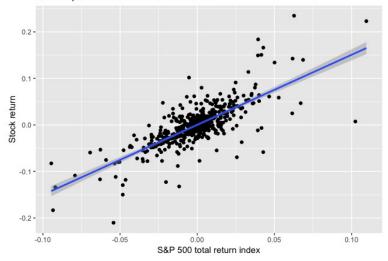


Figure 2: GS return vs. SPTR

(c) A scatter plot of 1-day stock returns vs. the VIX (CBOE implied volatility index) 1-day returns along with regression line and its slope equation. Should the VIX be considered as a "factor" in addition the "the market" as part of a linear regression model for stock returns?

Answer. Below is the summary of OLS regression followed by the scatter plot. We see that the 1-day stock returns and the VIX are negatively correlated.

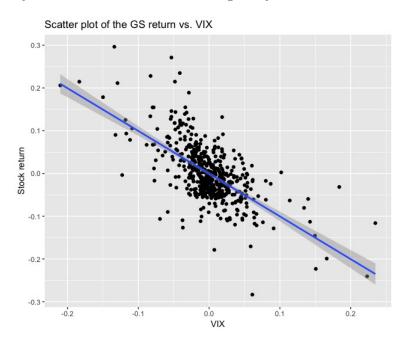


Figure 3: GS return vs. VIX

	Estimate	Std. Error	t value	$\Pr(> t)$
Intercept	-0.0005	0.0023	-0.22	0.8255
Slope	-0.9995	0.0527	-18.96	0.0000

- 2. **Performance modeling:** You are an evaluating a potential investment in a hedge fund. It is the end of 2004, and you have in-depth access to the funds daily position data for the period 2001-2004. Your goal is to characterize the past return profile in order to estimate its future success. Use the data for the equity long/short strategy called "Contra01" in the OLAP database **StatArb03** using the cube **Positions03**. For performance results, use the weighted returns net of transaction costs, given by the fields "Weighted Return Tc." Ignore hedge fund management fees.
 - (a) Give the annualized return, volatility, and Sharpe ratio (using $r_f = 0$) for the strategy.

Answer. Assume log returns and $r_f = 0$, we summarize the statistics of the hedge fund performance as follows:

Annualized Return	Volatility	Sharpe Ratio
110.7%	24.5%	4.53

(b) Use linear regression to report on the CAPM alpha, beta, and R-squared for the strategy, including standard errors of your estimates and t-statistics. Is the alpha statistically different from a CAPM prediction? How would you characterize the market-neutrality of the fund in light of its beta?

Answer. Recall that the CAPM model suggests that

$$r - r_f = \alpha + \beta (r_M - r_f).$$

Below is the summary for linear regression between hedge fund return and excess market return.

	Estimate	Std. Error	t value	Pr(> t)
α	0.0043	0.0005	8.90	0.0000
β	0.0971	0.0402	2.42	0.0159

Observations	Residual Std. Error	R^2 Adjusted	R^2
1004	0.01536	0.005791	0.004799

Notice that the P-value of α is close to zero, so the α is significantly different from CAPM prediction. Similarly, P-value of β is less than 0.05, β is significant, hence the hedge fund is not market-netural.

(c) Extend this regression model to include Fama-French factors, and repeat the analysis in the previous question. (Data is available in crispy04.dbo.famafrench.)

Answer. Recall that the Fama-French model suggests that

$$r - r_f = \alpha + \beta_M(r_m - r_f) + \beta_{SMB} \cdot SMB + \beta_{HML} \cdot HML + \beta_{UMD} \cdot UMD$$

Below is the summary for multi-linear regression between hedge fund return against the Fama-French factors. Notice that α is significant since P-value of α is close to zero. All the beta's are insignificant as the P-value for every beta is greater than 0.05.

	Estimate	Std. Error	t value	$\Pr(> t)$
α	0.0044	0.0005	8.98	0.0000
β_M	0.0498	0.0511	0.97	0.3302
β_{SMB}	-0.1514	0.0874	-1.73	0.0837
β_{HML}	0.0483	0.1132	0.43	0.6700
β_{UMD}	-0.1139	0.0661	-1.72	0.0852

Observations	Residual Std. Error	R^2 Adjusted	R^2
1004	0.01533	0.01288	0.00893

(d) Plot the strategy one-day returns in order, from lowest to highest. What fraction of days were winners and what fraction were losers? What was the median return of winners and what was the median return of the losers?

Answer. See below

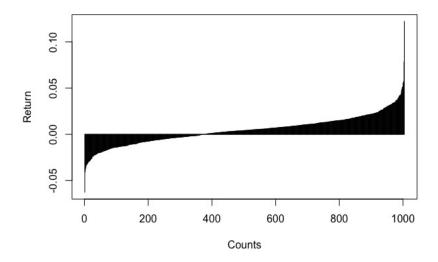


Figure 4: Strategy one-day ordered return

	Winners	Losers
Fracation	63.75%	37.25%
Median	1.02%	-0.80%

(e) Based on your analysis in 2004, how do you expect the fund to perform for the five-year period ahead?

Answer. Based on the Fama-French model, we see that the fund is netural to all Fama-French factors and has a significant alpha. Thus, we would expect the fund to perform very well in the upcoming five-year period.

3. **Performance attribution:** Analyze the performance of the contrarian equity long/short strategy "Contra01" over the period 2005-2009 using the data in the OLAP database **StatArb03** using the cube **Positions03**. For performance results, use the weighted returns net of transaction costs, given by the fields "Weighted Return Tc"

(a) Give the annualized return, volatility, and Sharpe ratio (using $r_f = 0$) for the strategy overall and for the long and short halves of the strategy. (That is, do the same calculations just using the long positions and then just the short positions.)

Answer. Assume log returns and $r_f = 0$, we summarize the statistics of the hedge fund's long strat, short strat and combined weighted returns as follows:

	Annualized Return	Volatility	Sharpe Ratio
Long	-0.21	0.32	-0.65
Short	0.58	0.29	2.03
Combined	0.37	0.28	1.31

(b) Use linear regression to report on the CAPM alpha, beta, and R-squared for the strategy, including standard errors of your estimates and t-statistics. Is the alpha statistically different from a CAPM prediction? How would you characterize the market-neutrality of the fund in light of its beta?

Answer. Below is the summary for linear regression result. We see that the α is significant, β is not significant hence the hedge fund is market neutral.

	Estimate	Std. Error	t value	$\Pr(> t)$
α	0.0014	0.0005	2.70	0.0071
β	0.0105	0.0332	0.32	0.7512

Observations	Residual Std. Error	R^2 Adjusted	R^2
1004	0.01533	0.01288	0.00893

(c) Plot the strategy one-day returns in order, from lowest to highest. What fraction of days were winners and what fraction were losers? What was the median return of winners and what was the median return of the losers?

Answer. See next page for the summary of winners and losers.

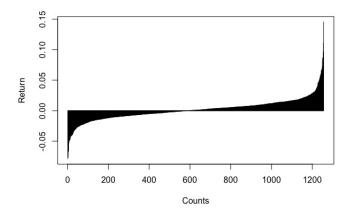


Figure 5: Strategy one-day ordered return

	Winners	Losers
Fracation	53.5%	46.5%
Median	0.83%	-0.82%

(d) How do these results compare with your predictions from the period 2001-2004 from Question #2 above? Do you believe the portfolio manager's style changed or remained constant over time? Why or why not?

Answer. Comparing the manager's performance from 2001-2004 to his performance from 2005-2009, based on CAMP regression, we conclude that the manager is not market neutral from 2001-2004 but market neutral from 2005-2009. Further, the change of winners and losers break down of these two periods corroborate the change of investment style.

(e) Which SIC sector contributed the largest total return, in absolute value, for calendar year 2006? What was the return, and what fraction of strategy total return did it contribute? (A plot or table of total weighted return, aggregated by sector and filtered on 2006 tradedates would be helpful.)

Answer. We performed the analysis using Excel with a query table linked to the OLAP database **StatArb03** using the cube **Positions03**. We filtered with "Domain = Contra01" and "Time = Calendar 2006", presented in the following table:

SIC Sectors	Weighted Return	Weighted Return	Fraction of Strategy
SIC Sectors	Tc Abs	Tc	Total Return
Agriculture Forestry,	0.028	0.003	0.6%
and Fishing	0.028	0.003	0.070
Construction	0.101	-0.005	-1.1%
Finance, Insurance, and Real Estate	0.969	0.093	21.3%
Manufacturing	8.018	0.212	48.2%
Mining	0.537	0.063	14.3%
Retail Trade	0.667	-0.011	-2.6%
Services	3.813	0.056	12.8%
Transportation, Communications,	1.123	0.008	1.8%
Electric, Gas, and Sanitary Services	1.125	0.008	1.870
Wholesale Trade	0.919	0.020	4.7%
Grand Total	16.175	0.439	100.0%

The manufacturing sector contributed the largest total return in 2006 in absolute value, with return of 21.2%, or 8.018 in absolute value, accounting for 48.2% of strategy total return.

(f) Same as previous question using GICS sectors.

Answer. Again, using Excel query with filters, "Domain = Contra01" and "Time = Calendar 2006", we derive the following table:

GICS Sectors	Weighted Return	Weighted Return	Fraction of Strategy
GICS Sectors	Tc Abs	Tc	Total Return
NA	7.219	0.194	44.3%
Energy	0.515	0.042	9.5%
Materials	0.310	-0.013	-3.0%
Industrials	1.121	0.079	18.1%
Consumer Discretionary	1.191	0.006	1.4%
Consumer Staples	0.382	0.017	3.9%
Health Care	2.098	-0.002	-0.4%
Financials	0.564	0.058	13.3%
Information Technology	2.667	0.059	13.5%
Telecommunication Services	0.074	-0.014	-3.1%
Utilities	0.033	0.011	2.5%
Grand Total	16.175	0.439	100.0%

The sector "NA" (meaning stocks that have no specific GICS sector classification) contributed the largest total return in 2006 in absolute value, with return of 19.4%, or 7.219 in absolute value, accounting for 44.3% of strategy total return.

If we consider "NA" not as a sector and exclude it from analysis, then the Information Technology sector contributed the largest total return in 2006 in absolute value, with return of 5.9%, or 2.667 in absolute value, accounting for 13.5% of strategy total return.

- 4. **Risk measurement:** Analyze the exposures of the contrarian equity long/short strategy "Contra01" using the fund data for 2005-2009 in the OLAP database **StatArb03** using the cube **Positions03**. This strategy focused on security selection and mean-reversion dynamics at the individual stock level. While the overall longs and shorts were constrained to be balanced, the individual sectors and other factors were not. Use the data on the history of portfolio exposures, measured primarily by aggregating security weights, to answer the following.
 - (a) For each GICS sector, give its highest and lowest net exposure over the period 2005-2009. Which sectors stayed within +/- 10% of portfolio weight throughout?

Answer. Using the Excel linked query table with filters "Domain = Contra01" and "Time = Calendar 2005-2009", we obtained the daily exposure for each GICS sector. A simple analysis in Excel shows:

GICS Sectors	Maximum Weight	Minimum Weight
NA	0.310	-0.222
Energy	0.310	-0.233
Materials	0.169	-0.127
Industrials	0.163	-0.190
Consumer Discretionary	0.194	-0.240
Consumer Staples	0.092	-0.080
Health Care	0.241	-0.222
Financials	0.333	-0.487
Information Technology	0.241	-0.236
Telecommunication Services	0.042	-0.051
Utilities	0.035	-0.048
Grand Total	0.030	-0.054

The Consumer Staples Sector, the Telecommunication Service Sector, and the Utilities Sector stayed within +/- 10% of portfolio weight throughout.

(b) On 9/15/2008, which GICS sector was the most unbalanced? Give the total portfolio weight long, total weight short, and net portfolio weight for that sector. What was the day's portfolio return?

Answer. Using the Excel linked query table with filters "Domain = Contra01" and "Date = 09/15/2008", we obtain:

GICS Sectors	Total Weight Long	Total Weight Short	Net Portfolio Weight
NA	0.366	-0.254	0.113
Energy	0.014	-0.085	-0.070
Materials	0.014	-0.014	0.000
Industrials	0.056	-0.085	-0.028
Consumer Discretionary	0.099	-0.183	-0.085
Consumer Staples	0.000	-0.014	-0.014
Health Care	0.211	-0.141	0.070
Financials	0.085	-0.113	-0.028
Information Technology	0.155	-0.099	0.056
Grand Total	1.000	-0.986	0.014

On 09/15/2008, the "NA" sector was the most unbalanced. Its total portfolio weight long was 36.6%, total portfolio weight short was -25.4%, and net portfolio weight was 11.3%.

If we consider "NA" not as a sector and exclude it from analysis, then the Consumer Discretionary sector was the most unbalanced. Its total portfolio weight long was 9.9%, total portfolio weight short was -18.3%, and net portfolio weight was -8.5%.

That day's portfolio return was -4.65% (before transaction costs) and -4.80% (after transaction costs).

(c) Same as previous question for 2/27/2007.

Answer. Using the Excel linked query table with filters "Domain = Contra01" and "Date = 02/27/2007", we obtain:

GICS Sectors	Total Weight Long	Total Weight Short	Net Portfolio Weight
NA	0.341	-0.280	0.061
Energy	0.012	-0.024	-0.012
Materials	0.037	-0.037	0.000
Industrials	0.110	-0.061	0.049
Consumer Discretionary	0.110	-0.098	0.012
Consumer Staples	0.037	-0.012	0.024
Health Care	0.122	-0.122	0.000
Financials	0.073	-0.061	0.012
Information Technology	0.134	-0.280	-0.146
Telecommunication Services	0.024	-0.012	0.012
Grand Total	1.000	-0.988	0.012

On 07/27/2007, the Information Technology sector was the most unbalanced. Its total portfolio weight long was 13.4%, total portfolio weight short was -28.0%, and net portfolio weight was -14.6%.

That day's portfolio return was 0.59% (before transaction costs) and 0.44% (after transaction costs).

(d) Suppose a sub-strategy were defined by restricting the "Contra01" portfolio weights and returns to the GICS Financials sector only. Is this sub-strategy "market-neutral" on its own? How do the sub-strategy exposures evolve over time in the absence of constraints? Plot and the net long/short exposure as well as the gross long vs. gross short exposures.

Answer. Using Excel linked query table, we extracted time-series data of weighted return (before transaction costs) of this sub-strategy. We then run a regression on this time series against market return in 2005-2009, applying the CAPM model. Regression result as below:

	Estimate	Std. Error	t value	$\Pr(> t)$
α	0.0003044	0.0001737	1.753	0.0799
β	-0.0035573	0.0114361	-0.311	0.7558
			- 0	_
watic	ms Residual	Std Error	R^2 Adim	etad

0.006162

The beta estimate has a t-stat of -0.311, which is insignificant at the 5% significance level, and we fail to reject the null hypothesis that beta is zero. With a zero beta, the sub-strategy can be considered as "market-neutral" on its own.

0.00007697

-0.0007185

Besides, a time-series plot of the sub-strategy exposure is shown below.

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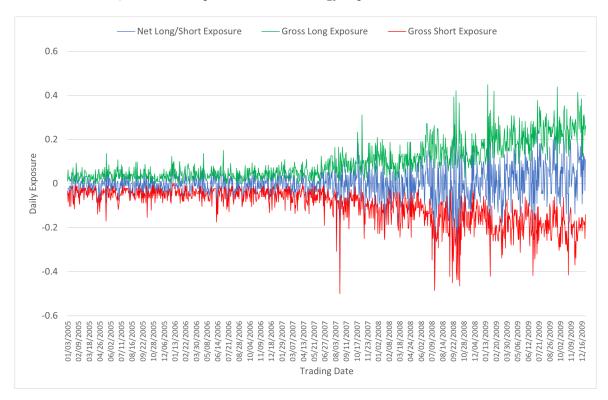


Figure 6: Daily Exposure of Financial-Sector-Only Sub-Strategy

As seen from the plot, the exposure in absolute value stayed low before mid-2007 and started increasing steadily since then.

5. Correlation dynamics: Analyze the behavior of the average correlation among stocks in the Dow Jones Industrial Average for the period 2008-2009 using the data in **crispy04**. Create two plots: one for a 1-month correlation window and the second using a 3-month window. On each graph, plot lines for the following daily time series:

(a) Average ("implied") correlation, as discussed in class,

$$\overline{\rho} \equiv \frac{\sigma_p^2 - \sum_{i=1}^{N} w_i^2 \sigma_i^2}{2 \sum_{i < j} w_i w_j \sigma_i \sigma_j} = \frac{\sigma_p^2 - \sigma_{(0)}^2}{\sigma_{(1)}^2 - \sigma_{(0)}^2}$$

- (b) Index realized variance
- (c) Index variance₀, the variance the index would have if all member correlations were zero.
- (d) Index variance₁, the variance the index would have if all member correlations were 100%. Data Cleaning: We gained raw data from Crispy04 for both index and 35 individual components that was at least once consisted in DJIA index during the period 2008 2009. Then, by adjusting the portfolio to correspond to the three rebalancing attempts for DJIA composition, we got the proper data for further analysis.

Data Processing and Computing: we then assigned the weights for each individual stock everyday and calculate the index realized variance, index variance if all member correlation were zero, and index variance if all member correlation were one. Lastly, we computed the implied volatility via the formula.

Average (Implied) Portfolio Correlation (1-Month Window)

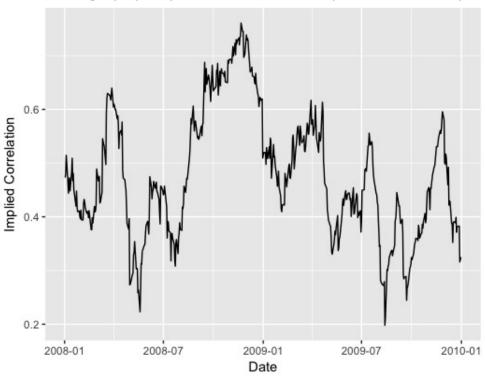


Figure 7: Times Series Plot of Average correlation (1-month Window)

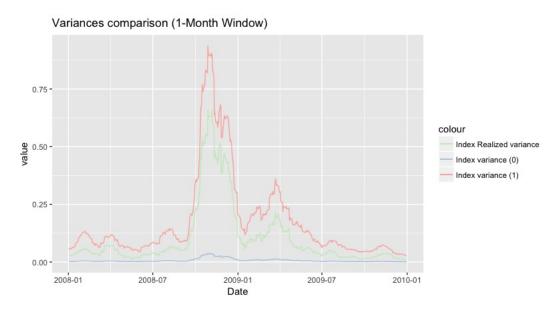


Figure 8: Times Series Plot of Variances under different assumptions (1-month Window)

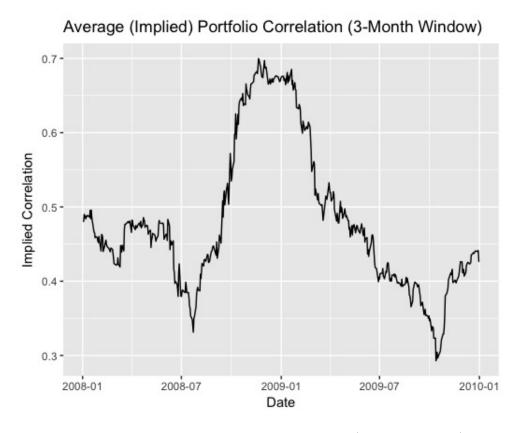


Figure 9: Times Series Plot of Average correlation (3-month Window)



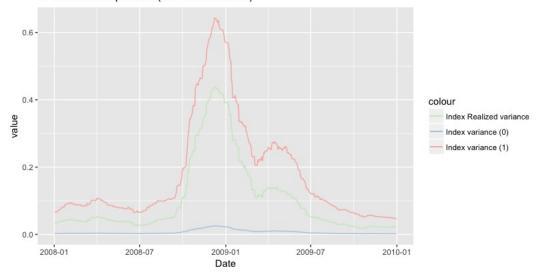


Figure 10: Times Series Plot of Variances under different assumptions (3-month Window)

(e) For each window size, what is the maximum value of the average correlation and when does it occur?

Answer. The maximum correlation is 0.76 under 1-month window size and 0.70 under 3-month window size. Maximum correlation happens at 2008-11-21 under 1-month window size and 2008-11-20 under 3-month window size. As we can see in the graphs for both window size, implied volatility tends to increase as the variance of index increases. Also, because 3-month correlation incorporates more past movements, as a result, the time series is more smooth compared to 1-month window size where everything is more noisy. However, even though the 3-month model may react to big changes a little slower, both time horizon models generate similar results and show similar relationship with the variances.

(f) What constraints should be satisfied by the four values? Do they hold for the period in question?

Answer. First of all, implied correlation should be in the range between -1 and 1 to be meaningful. Secondly, value of Index variance₀ should always be lower than the value of Index variance₁ because Index variance₁ takes positive covariance terms into account. Lastly, the realized variance should be lower than Index variance₁ and higher than Index variance₀ since the implied correlation stays positive over the period. All of the constraints hold for the period 2008-2009.

(g) What could cause the constraints to be violated?

Answer. Because the implied correlation is calculated by using correct formula that fits in every situation, correlation will always stays in the range as long as the variances are in their proper range. However, if the variances data fails to make sense, for example, if variance happens to be smaller than 0, then everything else may consequently break the constraint.

П

Appendix

Q1 R code

```
library('csv')
library('dplyr')
library('readxl')
library('aTSA')
library('outliers')
library('quantmod')
library('ggplot2')
library('xtable')
setwd("~/Desktop/15.458_Financial_Data_Science/PS3")
dt <- read.csv("problem1.csv",header=TRUE)</pre>
colnames(dt) = c('Date', 'SPX', 'SPTR', 'VIX', 'GS', 'beta')
dt$Date = as.Date(dt$Date,"%Y/%m/%d")
#check data integrity
sum(is.na(dt$GS))
sum(is.na(dt$SPX))
sum(is.na(dt$SPTR))
sum(is.na(dt$VIX))
sum(is.na(dt$beta))
#-----#
#1a) Time series of beta vs. SPX
scatterplot_a = ggplot(dt, aes(x=Date, y=beta)) + geom_line()
scatterplot\_a = scatterplot\_a + labs(x = "Date", y='Stock\_Beta', title='Time\_Series\_plot\_of\_therefore the scatterplot is a scatterplot in the scatterplot is a scatterplot in the scatterplot is a scatterplot in the scatterplot in the scatterplot is a scatterplot in the scatterplot is a scatterplot in the scatterplot is a scatterplot in the scatterplot in the scatterplot is a scatterplot in the scatterplot in the scatterplot is a scatterplot in the scatterplot in the scatterplot is a scatterplot in the scatterplot in the
scatterplot_a
#-----#
#1b)scatter plot of GS return vs. SPTR
scatterplot_b = ggplot(dt, aes(x=SPTR, y=GS)) + geom_point()
scatterplot_b
# Add the regression line
scatterplot_b + geom_smooth(method=lm)
#linear regression equation
ols_b = lm(dt\$GS \sim dt\$SPTR)
summary(ols_b)
#-----#
#1c)scatter plot of GS return vs. VIX
scatterplot_c = ggplot(dt, aes(x=GS, y=VIX)) + geom_point()
scatterplot_c = scatterplot_c +labs(x = "VIX",y='Stock_return', title='Scatter_plot_of_the_G
scatterplot_c
# Add the regression line
scatterplot_c + geom_smooth(method=lm)
#linear regression equation
ols_c = lm(dt\$VIX \sim dt\$GS)
summary(ols_c)
Q2 R code
library('csv')
```

```
library('dplyr')
library('readxl')
library('aTSA')
library('outliers')
library('quantmod')
library('ggplot2')
library('scales')
library('xtable')
setwd("~/Desktop/15.458_Financial_Data_Science/PS3")
dt <- read.csv("Q2return.csv",header=TRUE)</pre>
colnames(dt) = c('Date', 'Return')
dt$Date = as.Date(dt$Date,"%m/%d/%Y")
df.fama = read.table("Q2FamaFrench(01-04).txt", sep = ',', header = TRUE)
{\tt colnames(df.fama) = c('Date', 'Risk_{\sqcup}Free_{\sqcup}Rate', 'Market_{\sqcup}Excess_{\sqcup}Return',}
                      'SMB', 'HML', 'UMD')
df.fama$Date = as.Date(df.fama$Date,"%Y/%m/%d")
#-----#
#1a) annualized return, volatility and SR, assume log return
ann_return = mean(dt$Return)*252
vol = sqrt(252)*sd(dt$Return)
sr = ann_return/vol
df.a = data.frame(ann_return, vol, sr)
colnames(df.a) = c('Annualized_Return', 'Volatility', 'Sharpe_Ratio')
#2b) CAPM
answer2b = summary(lm((dt$'Return'-df.fama$'Risk Free Rate')~df.fama$'Market Excess Return')
#2c)
answer2c = summary(lm((dt$'Return'-df.fama$'Risk Free Rate')~df.fama$'Market Excess Return'+df
#2d)
plot(sort(dt$Return), type = 'h',ylab = 'Return', xlab = 'Counts')
frac_winners = length(which(dt$Return>0))/length(dt$Return)
frac_losers = length(which(dt$Return<0))/length(dt$Return)</pre>
median_winners = median( dt$Return[which(dt$Return>0)])
median_losers = median( dt$Return[which(dt$Return<0)])</pre>
df.d = data.frame(c(frac_winners, median_winners), c(frac_losers, median_losers))
colnames(df.d) = c('Winners', 'Losers')
rownames(df.d) = c('Fracation', 'Median')
Q3 R code
library('csv')
library('dplyr')
library('readxl')
library('aTSA')
library('outliers')
library('quantmod')
library('ggplot2')
library('scales')
library('xtable')
```

```
setwd("~/Desktop/15.458_Financial_Data_Science/PS3")
dt <- read.csv("Q3return.csv",header=TRUE)</pre>
colnames(dt) = c('Date', 'Long', 'Short', 'Combined')
dt$Date = as.Date(dt$Date,"%m/%d/%Y")
df.fama = read.table("Q3FamaFrench(05-09).txt", sep = ',', header = TRUE)
'SMB', 'HML', 'UMD')
df.fama$Date = as.Date(df.fama$Date,"%Y-%m-%d")
#-----#
#1a) annualized return, volatility and SR, assume log return for long-short strategy
f <- function(x){</pre>
   ann_return = mean(x)*252
   vol = sqrt(252)*sd(x)
   sr = ann_return/vol
   c(ann_return, vol, sr)
df.a = apply(dt[,-1],2,f)
rownames(df.a) = c('Annualized_Return', 'Volatility', 'Sharpe_Ratio')
#2b) CAPM
df = df= merge(dt,df.fama, by = "Date")
answer2b = summary(lm((df$'Combined'-df$'Risk Free Rate')~df$'Market Excess Return'))
#2d)
plot(sort(dt$Combined), type = 'h',ylab = 'Return', xlab = 'Counts')
frac_winners = length(which(dt$Combined>0))/length(dt$Combined)
frac_losers = length(which(dt$Combined<0))/length(dt$Combined)</pre>
median_winners = median( dt$Combined[which(dt$Combined>0)])
median_losers = median( dt$Combined[which(dt$Combined<0)])</pre>
df.d = data.frame(c(frac_winners, median_winners), c(frac_losers, median_losers))
colnames(df.d) = c('Winners', 'Losers')
rownames(df.d) = c('Fracation', 'Median')
Q4 R code
#import data regression
dt4_e_r <- read.csv("Q4d1_-_Contra01-GICSfinancial-Weights&Returns.csv",header=TRUE)
\texttt{dt4\_e\_mkt} \  \  \, \leftarrow \  \  \, \texttt{read\_csv}(\texttt{"}^{\texttt{Desktop}/\texttt{MIT}/\texttt{Fall}} \sqcup 2018/\texttt{Data} \sqcup \texttt{Science}/\texttt{Project} \sqcup \texttt{C}/\texttt{Q3} \sqcup \neg \sqcup \texttt{FamaFrench} \sqcup (05-09).
#clean the data
dt4_e_mkt$d = as.Date(dt4_e_mkt$d,"%Y-%m-%d")
dt4_e_r$Trade.Date = as.Date(dt4_e_r$Trade.Date,"%m/%d/%Y")
ols4_e = lm((dt4_e_r$Weighted.Return-dt4_e_mkt$rf_str) ~ (dt4_e_mkt$mkt_str-dt4_e_mkt$rf_str))
summary(ols4_e)
Q5 R code
#-----#
#5 correlation
#-----#
#import the data
dt5_comp <- read.csv("~/Desktop/MIT/Fall_2018/Data_Science/Project_C/Q5_query_-_Comp_Stocks.cs
\tt dt5\_djia <- read.csv("~/Desktop/MIT/Fall\_2018/Data\_Science/Project\_C/Q5\_query\_-\_DJIA.csv", headstarted to the contract of 
dt5_comp = as.data.frame(dt5_comp)
dt5\_comp$d = as.Date(dt5\_comp$d,"%Y-%m-%d")
```

```
dt5_comp = dt5_comp[order(dt5_comp$d),]
#check integrity (505 days in total)
test_group = group_by(dt5_comp, ticker)
test_statistis = summarise(test_group,
                                                                                 count = n()
)
#-----#
#clean the data to match djia composition
row_to_keep = which((dt5_comp$d < 2018-02-18 & !(dt5_comp$ticker %in% c('BAC','CVX','KFT','CSC
                                                                  (dt5_{comp}$d > 2018-02-18 \& dt5_{comp}$d < 2018-09-22 \& !(dt5_{comp}$ticker % dt5_{comp}$ticker % dt5_{
                                                                   (dt5_{comp}$d > 2018-09-22 \& dt5_{comp}$d < 2019-06-08 \& !(dt5_{comp}$ticker % dt5_{comp}$ticker % dt5_{
                                                                   (dt5_comp$d > 2019-06-08 & !(dt5_comp$ticker %in% c('MO', 'HON', 'AIG', 'C'
)
dt5_comp = dt5_comp[row_to_keep,]
grouped_data <- group_by(dt5_comp,d)</pre>
statistics <- summarise(grouped_data,</pre>
                                                                        count = n(),
                                                                         total_p = sum(price)#,
                                                                         #weight_sum = sum(weight)
statistics = as.data.frame(statistics)
dt5_comp$weight = NA
#-----#
#assign weight
for (i in 1:nrow(dt5_comp)) {
      for (j in 1:30) {
            dt5_comp$weight[i] = dt5_comp$price[i] / statistics$total_p[ceiling(i/30)]
#calculate index variance (sigma_p^2)
DJIA_one_month = dt5_djia[dt5_djia$nm_numtype == 'vol_021',]
DJIA_three_month = dt5_djia[dt5_djia$nm_numtype == 'vol_063',]
DJIA_one_month$index_variance = DJIA_one_month$value^2
DJIA_three_month$index_variance = DJIA_three_month$value^2
#clean DJIA data
df_DJIA_1m = as.data.frame(DJIA_one_month)%>%
      select('d','index_variance')
df_DJIA_3m = as.data.frame(DJIA_three_month)%>%
     select('d','index_variance')
#-----#
#prepare and calculate grouped data for results
#for one month window
group_1m <- group_by(dt5_comp,d)</pre>
statistics_1m <- summarise(group_1m,</pre>
                                                                                  count = n(),
                                                                                  sigma_zero = sum(vol_021^2 * weight^2)
statistics_1m = data.frame(statistics_1m)
#for three month window
```

```
group_3m <- group_by(dt5_comp,d)</pre>
statistics_3m <- summarise(group_3m,
                          count = n(),
                          sigma_zero = sum(vol_063^2 * weight^2)
statistics_3m = as.data.frame(statistics_3m)
#assign value of variance when correlation equals 0 (sigma_0^2)
df_DJIA_1m$sigma_zero = statistics_1m$sigma_zero
df_DJIA_3m$sigma_zero = statistics_3m$sigma_zero
#-----#
#calculate variance when correlation equals 1 (sigma_1^2)
#deno is denominator in the formula (=sigma(1)^2 - sigma(0)^2)
df_DJIA_1m$deno = 0
df_DJIA_3m$deno = 0
for (i in (30*(0:504)+1)) {
                              #i is start point of each day
 for (j in (i:(i+28))) {
                                #j is start point of calculation of each day
   for (k in (j:(i+29))) {
                                  #k is end point of calculation of each day
     df_DJIA_1m$deno[ceiling(i/30)] = df_DJIA_1m$deno[ceiling(i/30)] +
       2 * dt5\_comp$vol\_021[j] * dt5\_comp$vol\_021[k] * dt5\_comp$weight[j] * dt5\_comp$weight[k] 
   }
 }
}
for (i in (30*(0:504)+1)) {
                            #i is start point of each day
                                #j is start point of calculation of each day
 for (j in (i:(i+28))) {
   for (k in (j:(i+29))) {
                                 #k is end point of calculation of each day
     df_DJIA_3m$deno[ceiling(i/30)] = df_DJIA_3m$deno[ceiling(i/30)] +
       2 * dt5 = comp$vol_063[j] * dt5 = comp$vol_063[k] * dt5 = comp$weight[j] * dt5 = comp$weight[k]
 }
df_DJIA_1m$sigma_one = df_DJIA_1m$sigma_zero + df_DJIA_1m$deno
#-----#
#calculate average volatility
df_DJIA_1m$Rho = (df_DJIA_1m$index_variance - df_DJIA_1m$sigma_zero) / df_DJIA_1m$deno
df_DJIA_3m$Rho = (df_DJIA_3m$index_variance - df_DJIA_3m$sigma_zero) / df_DJIA_3m$deno
#plot the result
df_DJIA_1m$d = as.Date(df_DJIA_1m$d,'%Y-%m-%d')
ggplot(df_DJIA_1m,aes(d,group=1)) +
 geom_line(aes(y=Rho, colour = 'Implied_Correlation')) +
 \texttt{geom\_line(aes(y=index\_variance, colour = 'Index\_Realized\_variance')) +}
 geom_line(aes(y=sigma_zero, colour = 'Index_variance_(0)')) +
 \texttt{geom\_line}(\texttt{aes}(\texttt{y=sigma\_one}, \texttt{colour} = \texttt{'Index}_{\sqcup} \texttt{variance}_{\sqcup}(\texttt{1})\texttt{'})) +\\
 labs(x="Date",y="value",title = 'Average_(Implied)_Portfolio_Correlation_(1-Month_Window)')
 scale_colour_brewer(palette ="Pastel1", direction = -1)
df_DJIA_3m$d = as.Date(df_DJIA_3m$d,'%Y-%m-%d')
ggplot(df_DJIA_3m,aes(d,group=1)) +
 geom_line(aes(y=Rho, colour = 'Implied Correlation')) +
 {\tt geom\_line(aes(y=index\_variance, colour = 'Index_{\sqcup}Realized_{\sqcup}variance'))} \ + \\
 \texttt{geom\_line(aes(y=sigma\_zero, colour = 'Index_{\sqcup} variance_{\sqcup}(0)')) +}
  geom_line(aes(y=sigma_one, colour = 'Indexuvarianceu(1)')) +
```