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Question: toss a fair coin $n = 10000$ times. How many heads?

$X = \sum_{i=1}^n X_i$, $X_i \sim \text{Bern}(1/2)$. $\mathbb{E}[X] = 5000$. But $\mathbb{P}(X = 5000) = \binom{10^4}{5000} \cdot 2^{-10^4} \approx 0.008$.

Theorem 0.1 (Weak Law of Large Numbers) Let X_1, \dots, X_n be IID RVs with mean $\mathbb{E}[X_1] = \mu$. Then, for all $\varepsilon > 0$,

$$\mathbb{P}\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - \mu\right| > \varepsilon\right) \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

So $\mathbb{P}(X \in [5000 - n\varepsilon, 5000 + n\varepsilon]) \approx 1$.

Theorem 0.2 (Central Limit Theorem) Let X_1, \dots, X_n be IID RVs with mean $\mathbb{E}[X_1] = \mu$. Let $\text{Var}(X_1) = \sigma^2 < \infty$. Then $\frac{1}{\sigma\sqrt{n}} \sum_{i=1}^n (X_i - \mu) \xrightarrow{D} N(0, 1)$, i.e.

$$\mathbb{P}\left(\frac{1}{\sigma\sqrt{n}} \sum_{i=1}^n (X_i - \mu) \in A\right) \rightarrow \int_A \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

for all A .

So $\mathbb{P}\left(X \in \left[\frac{n}{2} - \frac{\sqrt{n}}{2} Q^{-1}(\delta), \frac{n}{2} + \frac{\sqrt{n}}{2} Q^{-1}(\delta)\right]\right) \geq 1 - \delta$, for n large enough, where $Q(\delta) = \int_{\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$. We have $Q^{-1}(x) \propto \sqrt{\log \frac{1}{x}}$. So interval has length $\propto \sqrt{n} \sqrt{\log \frac{1}{\delta}}$.

Theorem 0.3 (Chebyshev's Inequality) $\mathbb{P}(|X - \mu| \geq \varepsilon) \leq \frac{\text{Var}(X)}{\varepsilon^2}$ for all $\varepsilon > 0$.

Corollary 0.4 $\mathbb{P}\left(\left|\sum_{i=1}^n (X_i) - \frac{n}{2}\right| \geq t\right) \leq \frac{\text{Var}(\sum_{i=1}^n X_i)}{t^2} = n \frac{\sigma^2}{t^2} \leq \delta$ where $t = \sqrt{n}\sigma/\sqrt{\delta}$.

So $\mathbb{P}(X \in [\frac{n}{2} - t, \frac{n}{2} + t]) \geq 1 - \delta$.

Question 2: we have N coupons. Each day receive one uniformly at random independent of the past. How many days until all coupons received?

We have $X = \sum_{i=1}^n X_i$, $X_i \sim \text{Geom}(\frac{1}{n})$. $\mathbb{E}[X] = \sum_i \mathbb{E}[X_i] \approx n \log n$ (verify this).

Question 3: Let $(X_1, \dots, X_n), (Y_1, \dots, Y_n)$ be IID. What is the longest common subsequence, i.e. $f(X_1, \dots, X_n, Y_1, \dots, Y_n) = \max\{k : \exists i_1, \dots, i_k, j_1, \dots, j_k \text{ s.t. } X_{i_j} = Y_{j_j} \forall j \in [k]\}$. Computing f is NP-hard. f is smooth.

Principle: a smooth function of many independent random variables concentrates around its mean.

Tower property of conditional expectation: $\mathbb{E}(\mathbb{E}(Z | X, Y) | Y) = \mathbb{E}(Z | Y)$.

1. The Chernoff-Cramer method

Theorem 1.1 (Markov's Inequality) Let Y be a non-negative RV. For any $t \geq 0$,

$$\mathbb{P}(Y \geq t) \leq \frac{\mathbb{E}[Y]}{t}.$$

Proof. We have $Y = Y \mathbb{I}_{\{Y \geq t\}} + Y \mathbb{I}_{\{Y < t\}} \geq t \cdot \mathbb{I}_{\{Y \geq t\}}$. Taking expectations gives the result. □

Corollary 1.2 (Chebyshev's Inequality) $\mathbb{P}(|Y - \mathbb{E}[Y]| \geq t) \leq \frac{\mathbb{E}[Y - \mathbb{E}[Y]]^2}{t^2}$

Proof. Take $Z = Y - \mathbb{E}[Y]$ and use Markov's. □

Let $\varphi : \mathbb{R} \rightarrow \mathbb{R}_+$ be non-decreasing, then $\mathbb{P}(\varphi(Y) \geq \varphi(t)) \leq \frac{\mathbb{E}[\varphi(Y)]}{\varphi(t)}$. For $\varphi(t) = t^2$, we can use $\text{Var}\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n \text{Var}(X_i)$.

Exercise 1.3 Prove WLLN, assuming that $\text{Var}(X_1) < \infty$, using Chebyshev's inequality.

Notation 1.4 For $\lambda > 0$, let $\varphi_\lambda(t) = e^{\lambda t}$.

Note $\mathbb{P}(Z \geq t) = \mathbb{P}(e^{\lambda Z} \geq e^{\lambda t}) \leq \frac{\mathbb{E}[e^{\lambda Z}]}{\varphi_\lambda(t)} = e^{-(\lambda t - \log(\mathbb{E}[e^{\lambda Z}]))}$. So if $\mathbb{E}[e^{\lambda Z}] < \infty$, then we have exponential concentration.

$F(\lambda) := \mathbb{E}[e^{\lambda Z}] = \sum_{i=0}^{\infty} \frac{\lambda^i \mathbb{E}[Z^i]}{i!}$. We have $\varphi_Z(\lambda) = \log(F(\lambda))$ is additive: if $Z = \sum_{i=1}^n Z_i$, Z_i independent, then $\varphi_Z(\lambda) = \log(\mathbb{E}[e^{\lambda Z}]) = \sum_i \log \mathbb{E}[e^{\lambda Z_i}]$.

So $\mathbb{P}(Z \geq t) \leq \inf_{\lambda > 0} e^{-(\lambda t - \varphi_\lambda(Z))} = e^{-\sup_{\lambda > 0} (\lambda t - \varphi_Z(\lambda))}$. This is the Chernoff bound. We denote $\varphi_Z^*(t) = \sup_{\lambda > 0} \lambda t - \varphi_Z(\lambda)$. This is Cramer's transform of Z .

Goal is to obtain upper bound on $\varphi_Z(\lambda)$, as this will give concentration. The function $\varphi_{Z - \mathbb{E}[Z]}(\lambda)$ gives upper bounds on $\mathbb{P}(Z - \mathbb{E}[Z] \geq t)$, the function $\varphi_{-Z + \mathbb{E}[Z]}(\lambda)$ gives upper bounds on $\mathbb{P}(Z - \mathbb{E}[Z] \leq -t)$.

Proposition 1.5 Properties of $\varphi_Z(\lambda)$:

1. $\varphi_Z(\lambda)$ is convex and infinitely differentiable on (a, b) , where $b = \sup_{\lambda > 0} \{\mathbb{E}[e^{\lambda Z}] < \infty\}$.
2. $\varphi_Z^*(t) \geq 0$ and convex.
3. If $t > \mathbb{E}[Z]$, then $\varphi_Z^*(t) = \sup_{\lambda \in \mathbb{R}} (\lambda t - \varphi_Z(\lambda))$, the Fenchel-Legendre dual.