Lab 10: Multinomial and Ordered Probit Models

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Code has been adapted frm Simon Hoellerbauer.

Load necessary packages:

Ordered Probit

Data and resources for this section from http://stats.idre.ucla.edu/r/dae/ordinal-logistic-regression/.

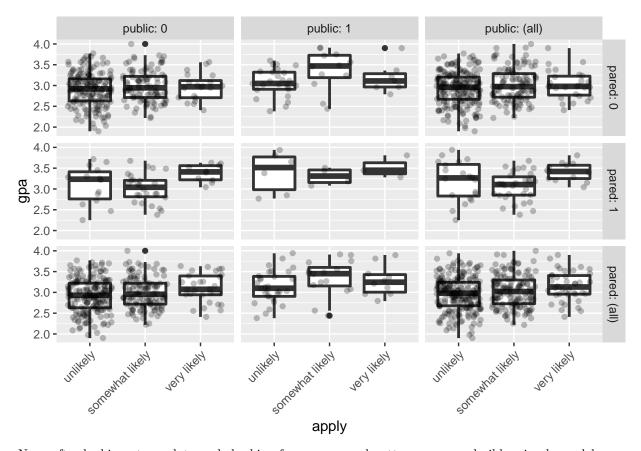
We are going to expand an earlier example that we used for logistic regression: the data about application/acceptance into graduate school.

These data include our DV, 'a three level variable called apply, with levels unlikely, somewhat likely, and very likely, and three predictor variables: pared (dichotomous, does either parent have a graduate degree?) public (dichotomous, 0 or 1 indicating whether or not the students' undergrad institution is public) and gpa (continuous).

```
grad <- read_dta("https://stats.idre.ucla.edu/stat/data/ologit.dta")
#when using read_dta, and we want to turn labelled vectors into factors,
#we have to use to_factor from labelled packaged
grad$apply <- to_factor(grad$apply)</pre>
```

Remember that before we even ran any tests for our logit model, we made a plot to understand the distribution of our variables of interest? We should do that again. Let's try a boxplot- or rather, a series of boxplots.

```
ggplot(grad, aes(x = apply, y = gpa)) +
  geom_boxplot(size = 1) +
  geom_jitter(alpha = .25) +
  facet_grid(pared~public, margins = TRUE, labeller=label_both) +
  theme(axis.text.x = element_text(angle = 45, hjust = 1, vjust = 1))
```



Now, after looking at our data and checking for any unusual patterns, we can build a simple model.

For ordered outcomes, we need to use the MASS package (rather than base R). Otherwise, our syntax looks much the same.

```
#First, a few checks to make sure our variables are in order:
is.factor(grad$apply)
## [1] TRUE
print(levels(grad$apply))
## [1] "unlikely"
                          "somewhat likely" "very likely"
#Now, our model:
m1 <- polr(apply ~ pared + public + gpa,</pre>
           data = grad, method = "probit", Hess = TRUE)
summary(m1)
## polr(formula = apply ~ pared + public + gpa, data = grad, Hess = TRUE,
       method = "probit")
##
##
## Coefficients:
##
            Value Std. Error t value
## pared 0.59811
                      0.1579 3.78881
                      0.1728 0.05878
## public 0.01016
## gpa
          0.35815
                      0.1568 2.28479
##
## Intercepts:
```

```
## wnlikely|somewhat likely 1.2968 0.4675 2.7738
## somewhat likely|very likely 2.5028 0.4766 5.2517
##
## Residual Deviance: 717.4951
## AIC: 727.4951
#What is the effect of increasing each predictor by one unit?
#How do the odds of moving from a lower to adjacent higher category change?
coefs <- m1$coefficients[1:3]
#odds ratios
exp(coefs)</pre>
```

```
## pared public gpa
## 1.818675 1.010211 1.430684
```

We can interpret these exponentiated coefficients in the same way as odds ratios in a logistic regression, but for the odds of each outcome category. Here, we can say that for pared, a one unit increase in parental education, i.e., going from 0 (Low) to 1 (High), the odds of very likely applying versus somewhat likely or unlikely applying combined are 1.81x greater, given that all of the other variables in the model are held constant.

We can (and should) also informally test the proportional odds assumption. What's this?

```
m1_prop_1 <- glm(I(as.numeric(apply) >= 2) ~ pared, family = "binomial",
                  data = grad)
m1_prop_2 <- glm(I(as.numeric(apply) >= 3) ~ pared, family = "binomial",
                  data = grad)
#Pared = 0
check_1 <- m1_prop_1$coef[1] - m1_prop_2$coef[1]</pre>
check_1
## (Intercept)
      2.062399
#Pared = 1
check_2 <- (m1_prop_1$coef[1] + m1_prop_1$coef[2]) -</pre>
  (m1_prop_2$coef[1] + m1_prop_2$coef[2])
check_2
## (Intercept)
      2.112541
##
```

When pared is equal to no (0) the difference between the predicted value for apply greater than or equal to two and apply greater than or equal to three is roughly 2 (-0.378 - -2.440 = 2.062). For pared equal to yes the difference in predicted values for apply greater than or equal to two and apply greater than or equal to three is also roughly 2 (0.765 - -1.347 = 2.112).

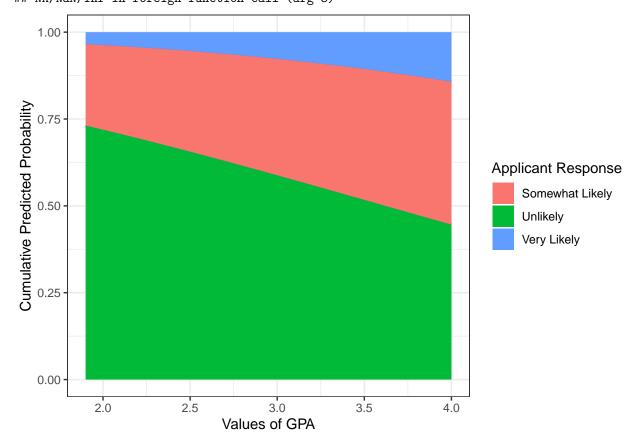
We will also want to generate predicted probabilities and plot them. Let's first do so with a plot of the cumulative probability (so the probabilities of being in each subcategory sum to 1).

We can also check this first:

```
tau1 <- m1$zeta[1] #extract latent thresholds (calculated w polr) from our model!
tau2 <- m1$zeta[2]
p1 <- invlogit(tau1) # 1</pre>
```

```
p2 <- invlogit(tau2) - invlogit(tau1) # 2</pre>
p3 <- 1-invlogit(tau2)
p1 + p2 + p3
## unlikely|somewhat likely
##
And now move on to plotting.
# if we wanted to plot the range of gpa for all four combinations of pared and
#public, we could do the following:
\#qrad\ 2 \leftarrow expand.qrid(pared = c(0:1),
                       public = c(0,1),
                        qpa = seq(from = 1.9, to = 4, length.out = 1000))
#We would then create an indicator variable for the four combinations, and then
#use facet_wrap in ggplot to create 4 different plots
#however, if we just want one easily interpretable plot, we can do the following:
table(grad$pared); table(grad$public)
##
##
     Ω
        1
## 337 63
##
##
     0
         1
## 343 57
grad_2 <- expand.grid(pared = 0, #modal value</pre>
                      public = 0, #modal value
                       gpa = seq(from = 1.9, to = 4, length.out = 1000))
colnames(grad_2)<- c("pared", "public", "gpa")</pre>
probs <- predict(m1, grad_2, type = "probs") #type = is different than for binomial model
head(rowSums(probs)) #just to double check
## 1 2 3 4 5 6
## 1 1 1 1 1 1
cumul_probs <- t(apply(probs, 1, cumsum)) #get the cumulative sum of the rows</pre>
#transpose into columns for plotting and merging with values of gpa
head(cumul_probs)
      unlikely somewhat likely very likely
##
                      0.9657982
## 1 0.7311624
## 2 0.7309140
                      0.9657411
                                          1
## 3 0.7306654
                      0.9656839
                                          1
## 4 0.7304167
                      0.9656266
                                          1
## 5 0.7301679
                      0.9655692
## 6 0.7299190
                      0.9655118
plot_data <- data.frame(grad_2$gpa, cumul_probs)</pre>
ggplot(data=plot_data, aes(x=grad_2.gpa))+
    geom_smooth(data=plot_data, aes(x=grad_2.gpa, y=unlikely), se=FALSE)+
    geom_smooth(data=plot_data, aes(x=grad_2.gpa, y=somewhat.likely), se=FALSE)+
```

```
## `geom_smooth()` using method = 'gam' and formula 'y ~ s(x, bs = "cs")'
## `geom_smooth()` using method = 'gam' and formula 'y ~ s(x, bs = "cs")'
## `geom_smooth()` using method = 'gam' and formula 'y ~ s(x, bs = "cs")'
## Warning: Computation failed in `stat_smooth()`:
## NA/NaN/Inf in foreign function call (arg 3)
```



And plotting a non-cumulative probability example.

```
#data frame
grad_3 <- expand.grid(
  pared = 1, #assume parents education =1
  public = c(1,0), #vary values of public
  gpa = mean(grad$gpa)) #average gpa

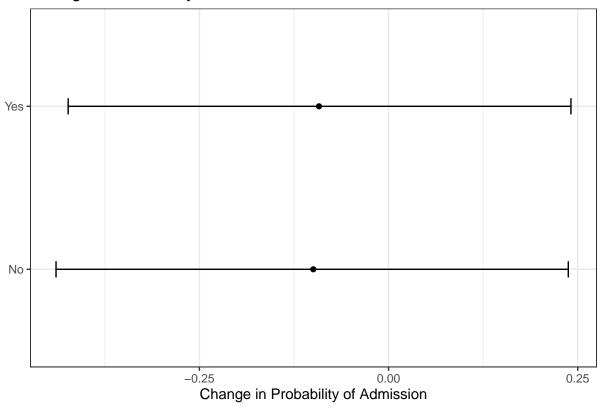
library(MASS)
set.seed(831213)</pre>
```

```
coef_samp <- mvrnorm(1e3,coef(m1)</pre>
                      ,vcov(m1)[1:3,1:3]) #keep relevant vals of vcov matrix
## Get linear predictors
hyp_X <- model.matrix(~pared+public+gpa</pre>
                       ,data = grad_3)[,-1] #no intercept
eta <- hyp_X %*% t(coef_samp)</pre>
summary(t(eta))
##
## Min.
          :0.1523 Min.
                            :0.134
## 1st Qu.:1.3838 1st Qu.:1.349
## Median :1.6981 Median :1.682
## Mean :1.6885
                     Mean :1.670
## 3rd Qu.:2.0149 3rd Qu.:1.983
## Max.
           :3.3749 Max.
                             :3.318
## Get pred probs samples
taus <- m1$zeta #extract thresholds</pre>
pred_prob_diffs <- apply(eta,c(1,2)</pre>
                     ,function(x, J, tau){
                      prob_vec <- array(NA,J) #J categories - array() is generalizable</pre>
                      prob_vec[1] <- plogis(tau[1]-x) #prob of being in 1st category</pre>
                       for(i in 2:(J-1)) #move to next category
                         prob_vec[i] <- plogis(tau[i]-x) - plogis(tau[i-1]-x)</pre>
                         #probability of being in middle category
                      prob_vec[J] <- 1 - plogis(tau[J-1]-x)</pre>
                         #probability of being in top category (sums to 1)
                      return(prob_vec[J]-prob_vec[1])
                       # Diff. in prob from top to bottom cat.
                     J = 3 #3 possible outcomes
                     ,tau = taus)
## Get effect on probability difference - if looking at more than one quantity
#fx_on_prob_diff <- apply(pred_prob_diffs, 2</pre>
                         #, function(x){
                          \#return(c(public=x[2]-x[1]))
                          #})
## Obtain 90% confidence interval on probabilities
fx_on_prob_ci <- t(apply(pred_prob_diffs,1</pre>
                        ,quantile,probs=c(0.05,0.5,0.95)))
#find difference in predicted probabilities and confidence around those differences
public <- c("Yes", "No")</pre>
#for easy plotting
fx_on_prob_ci <- as.data.frame(cbind(fx_on_prob_ci, public))</pre>
colnames(fx_on_prob_ci)<-c("low", "pe", "high", "public")</pre>
```

```
fx_on_prob_ci$low <- as.numeric(fx_on_prob_ci$low) #fix issues with factors
fx_on_prob_ci$pe <- as.numeric(fx_on_prob_ci$pe)
fx_on_prob_ci$high <- as.numeric(fx_on_prob_ci$high)

## Produce plot
library(ggplot2)
ggplot(fx_on_prob_ci, aes(x=pe, y=as.factor(public))) +
  labs(x="Change in Probability of Admission", y="") +
  ggtitle("Change in Probability for Public Vs Non Public")+
  geom_point(aes(x=pe)) +
  geom_errorbarh(aes(xmin=low, xmax=high), height=0.1) + #errorbar for confidence
  theme_bw()</pre>
```

Change in Probability for Public Vs Non Public



#not much of an effect, as our model told us originally!

Multinomial and Conditional Logit

```
#Loading additional necessary packages
library(nnet) #for multinom()
library(mlogit) #for conditional mnl
#library(pscl) #for example data
library(cowplot) #we will use it to put plots together (an extension to ggplot2)
library(tidyverse)
```

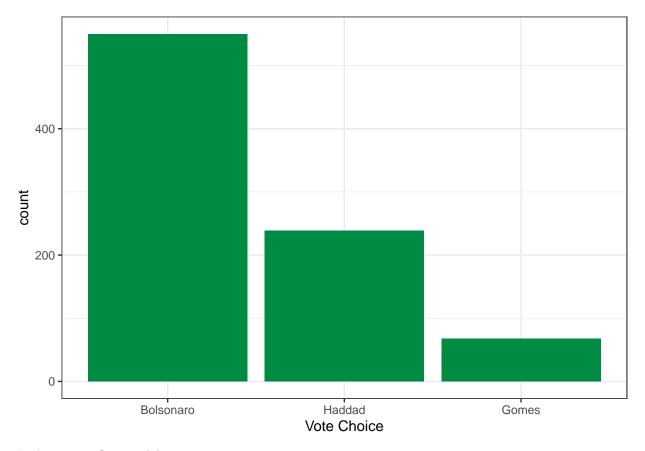
For this example, we'll use data from the 2017/18 round of LAPOP in Brazil. The data has a variable called "vote_choice" which is a categorical variable which we'll treat as unordered. In particular, we'll examine who

individuals voted for in Brazil in the 2018 national election (note there were MANY candidates, we'll only examine three - Bolsonaro, Haddad, and Gomes).

```
brazil_data <- readRDS("BrazilLAPOP_2018.rds")</pre>
#A bit of housekeeping:
#vote choice variable
brazil_data <- subset(brazil_data, vote_choice == "1501" |</pre>
                         vote choice == "1502" |
                         vote_choice == "1503")
brazil_data$vote_choice <- as.factor(brazil_data$vote_choice)</pre>
recode(brazil_data$vote_choice, "1501" = "Bolsonaro", "1502" = "Haddad",
       "1503" = "Gomes")
levels(brazil_data$vote_choice) <- c("Bolsonaro", "Haddad", "Gomes")</pre>
#capital punishment variable
brazil_data$capital_pun <- ifelse(brazil_data$capital_pun == 1, 1,</pre>
                                    ifelse(brazil_data$capital_pun == 2, 0, NA))
#sex variable
#male is 1, female is 0
brazil_data$sex <- ifelse(brazil_data$sex == 1, 1,</pre>
                                    ifelse(brazil_data$sex == 2, 0, NA))
```

Let's make a descriptive plot of our DV. This is good practice - something to consider for your replication papers!

```
ggplot(data=brazil_data)+
  geom_bar(aes(x=vote_choice), fill="springgreen4")+
  scale_x_discrete(name="Vote Choice")+
  theme_bw()
```



Let's run our first model.

We can change the reference category if we wish, and re-run:

```
## # weights: 15 (8 variable)
## initial value 909.650975
## iter 10 value 678.683284
## final value 678.338150
## converged
## Call:
## multinom(formula = vote_choice ~ sex + age + capital_pun, data = brazil_data,
## Hess = TRUE)
##
## Coefficients:
## (Intercept) sex age capital_pun
```

```
## Gomes
             -0.62025531 0.3802102 -0.019085614 -0.3093388
##
## Std. Errors:
##
             (Intercept)
                                sex
                                             age capital_pun
               0.2639588 0.1597405 0.005448676
## Bolsonaro
                                                   0.1602376
               0.4609828 0.2863508 0.010459885
## Gomes
                                                   0.2879088
##
## Residual Deviance: 1356.676
## AIC: 1372.676
We can also, as you may recall, run this as a series of binary logits (after some minor data recoding):
#filter data
brazil_data_2<-subset(brazil_data[which(brazil_data$vote_choice=="Bolsonaro" |</pre>
                                            brazil_data$vote_choice=="Haddad"),])
brazil_data_3<-subset(brazil_data[which(brazil_data$vote_choice=="Haddad" |
                                            brazil_data$vote_choice=="Gomes"),])
#recode values
brazil_data_2$vote_choice<-ifelse(brazil_data_2$vote_choice=="Haddad", 0,1)
#votes for Bolsonaro
brazil_data_3$vote_choice<-ifelse(brazil_data_3$vote_choice=="Haddad", 0,1)
#votes for Gomes
#check it out
head(brazil_data_2$vote_choice)
head(brazil_data_3$vote_choice)
#models!
glm_1<-glm(vote_choice~sex + age + capital_pun,</pre>
           data=brazil_data_2,
           family=binomial (link="logit"))
glm 2<-glm(vote choice~sex + age + capital pun,
           data=brazil_data_3,
           family=binomial (link="logit"))
coefs<-rbind(coef(glm_1), coef(glm_2), coef(multi_2))</pre>
coefs
```

0.2222934

What assumptions are we relaxing when we take this strategy?

Bolsonaro 0.07747397 0.5880453 0.008399133

Finally, we can run this as a slightly different model using a multinomial choice model and another package. To do so, we need to take a few extra steps. This allows us to change our data from a single observation per individual that records their vote intent to multiple observations for an individual that encompass all of their choices.

This means we can have information about alternatives that would be constant across individuals (ie policy position for each of these candidates) and information about individuals that is constant across alternatives (like gender or age of an individual).

In these particular data we only have individual-level information. However, we can set up these data in the same way we would a model that allows us alternative choices.

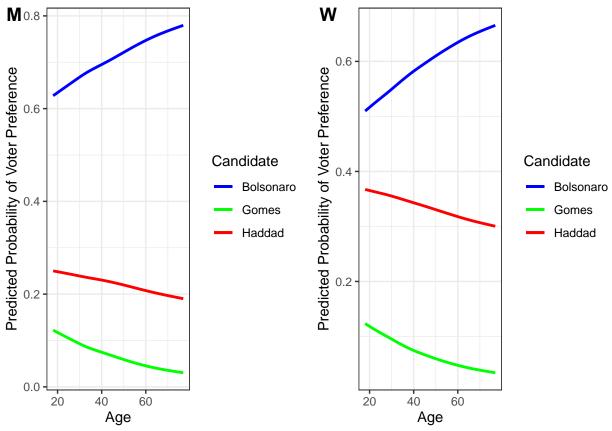
```
#First, a look at our data
brazil data[1:10,]
levels(brazil_data$vote_choice)
#Reshape- we can do this with a function in the mlogit package
brazil data <- as.data.frame(na.omit(brazil data))</pre>
brazil_long <- mlogit.data(brazil_data,</pre>
                              shape = "wide",
                              choice = "vote_choice",
                              alt.levels = c("Haddad", "Bolsonaro", "Gomes"))
                             #These match the levels above
#Check our data now
head(brazil_long)
brazil_long[1:10,]
#Replicate our multinomial logit and check with the 2nd model, above:
multi_3 <- mlogit(vote_choice ~ 0|sex+age+capital_pun,</pre>
                  data=brazil_long, reflevel="Haddad")
# 0/ is used in this case because we have no 'alternative-level variables'
head(model.matrix(multi_3))
check<-list(summary(multi_3), summary(multi_2))</pre>
check #look at the intercept for multi_3 and compare to multi_2
```

Let's return to our multinomial model (multi_2) and generate some QIs (predicted probabilities). We'll take a random sample of our data and then use our expand.grid() to see the effect of different combinations of our explanatory variables.

```
\#c(M=mean(x), quantile(x, c(0.25, 0.75)))
\#\}))
```

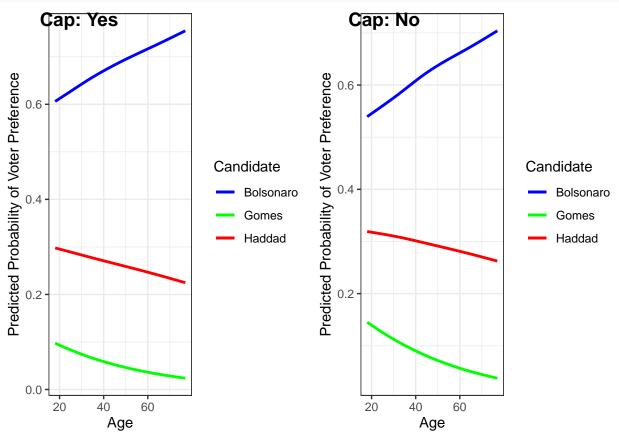
And we can plot, with a bit more detail:

```
#Let's separate out our predicted probabilities for each candidate by gender:
m pred <-plotdat[which(plotdat$V4==1),]</pre>
m_pred <- m_pred[sample(nrow(m_pred), 1000), ]</pre>
f_pred <-plotdat[which(plotdat$V4==0),]</pre>
f_pred <- f_pred[sample(nrow(f_pred), 1000), ]</pre>
mplot<-ggplot(data=m_pred)+</pre>
  geom_smooth(data=m_pred, aes(x=V5, y=Haddad, colour="Haddad"), se=FALSE)+
  geom_smooth(data=m_pred, aes(x=V5, y=Bolsonaro, colour="Bolsonaro"), se=FALSE)+
  geom_smooth(data=m_pred, aes(x=V5, y=Gomes, colour="Gomes"), se=FALSE)+
  labs(x="Age", y="Predicted Probability of Voter Preference")+
  scale colour manual(name="Candidate",
    values=c(Haddad="red", Bolsonaro="blue", Gomes="green"))+
  theme_bw()
fplot<-ggplot(data=f_pred)+</pre>
  geom_smooth(data=f_pred, aes(x=V5, y=Haddad, colour="Haddad"), se=FALSE)+
  geom_smooth(data=f_pred, aes(x=V5, y=Bolsonaro, colour="Bolsonaro"), se=FALSE)+
  geom_smooth(data=f_pred, aes(x=V5, y=Gomes, colour="Gomes"), se=FALSE)+
  labs(x="Age", y="Predicted Probability of Voter Preference")+
  scale_colour_manual(name="Candidate",
    values=c(Haddad="red", Bolsonaro="blue", Gomes="green"))+
  theme bw()
both <- plot_grid(mplot, fplot, labels = c('M', 'W'))</pre>
both
```



#Now let's separate out by support for capital punishment cap_pred <-plotdat[which(plotdat\$V6==1),]</pre> cap_pred <- cap_pred[sample(nrow(cap_pred), 1000),]</pre> nocap pred <-plotdat[which(plotdat\$V6==0),]</pre> nocap pred <- nocap pred[sample(nrow(nocap pred), 1000),]</pre> cap_plot<-ggplot(data=cap_pred)+</pre> geom_smooth(data=cap_pred, aes(x=V5, y=Haddad, colour="Haddad"), se=FALSE)+ geom_smooth(data=cap_pred, aes(x=V5, y=Bolsonaro, colour="Bolsonaro"), se=FALSE)+ geom_smooth(data=cap_pred, aes(x=V5, y=Gomes, colour="Gomes"), se=FALSE)+ labs(x="Age", y="Predicted Probability of Voter Preference")+ scale_colour_manual(name="Candidate", values=c(Haddad="red", Bolsonaro="blue", Gomes="green"))+ theme_bw() nocap_plot<-ggplot(data=nocap_pred)+</pre> geom_smooth(data=nocap_pred, aes(x=V5, y=Haddad, colour="Haddad"), se=FALSE)+ geom_smooth(data=nocap_pred, aes(x=V5, y=Bolsonaro, colour="Bolsonaro"), se=FALSE)+ geom_smooth(data=nocap_pred, aes(x=V5, y=Gomes, colour="Gomes"), se=FALSE)+ labs(x="Age", y="Predicted Probability of Voter Preference")+ scale_colour_manual(name="Candidate", values=c(Haddad="red", Bolsonaro="blue", Gomes="green"))+ theme_bw()





Lab Assignment

Ordered Outcome Practice

Use the data below. Estimate an ordered logit with 'warm' as the DV, using male, white, age, education, and prestige as independent variables.

Check the proportional odds assumption for at least one of your variables - does it hold?

Generate a cumulative predicted probability plot.

Data and description from Christopher Adolph: Example data from 1977, 1989 GSS: Attitudes towards working mothers: "A working mother can establish just as warm and secure of a relationship with her child as a mother who does not work." SD, D, A, SA Covariates: male respondent; white respondent; age of respondent; years of education of respondent; prestige of respondent's occupation (% considering prestigious)

hw <- read.csv(url("http://faculty.washington.edu/cadolph/mle/ordwarm2.csv"))</pre>

Multinomial Logit Practice

Get the data below, which contains information about high school students' college program plans based on a number of predictors (school type, gender, test scores). 'prog' is your dependent variable.

Produce a meaningful plot of your DV, with appropriate labels.

Model 'prog' as an outcome based on factors you think may matter for high school students' choices. Use a meaningful reference level for your discussion.

Modify this model in two ways: as a set of logits, and as a choice model.

Produce predicted probabilities from your model and plot them. To do so, you may build a hypothetical variable as in the past or plot average predicted probabilities. Make these plots well-labeled and informative.