

The generalized front-door criterion for estimation of indirect causal effects of a confounded treatment

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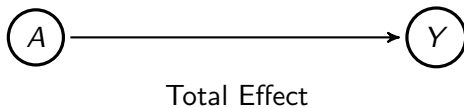
Overview

- Background and motivation
- Nonparametric identification
- Robust estimation
- Simulation study

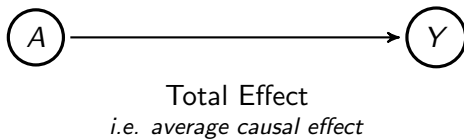
Background: the main goal

The goal of this work is to develop methodology to identify and estimate a novel causal effect, the population intervention indirect effect, in the presence of exposure-outcome unmeasured confounding.

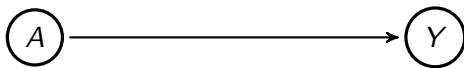
Background: mediation analysis



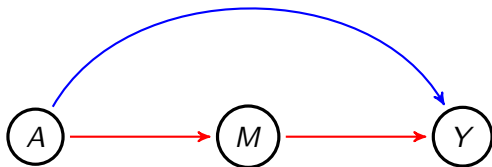
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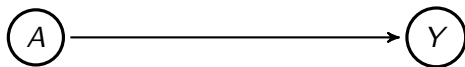
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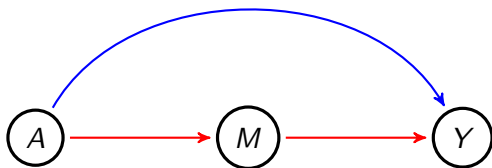
Total Effect
i.e. average causal effect



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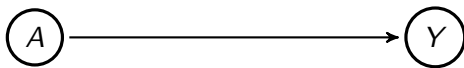


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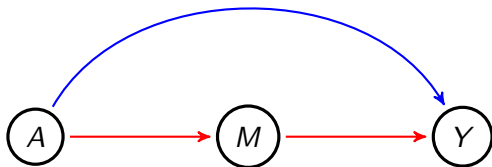


Direct and Indirect Effects

Background: mediation analysis



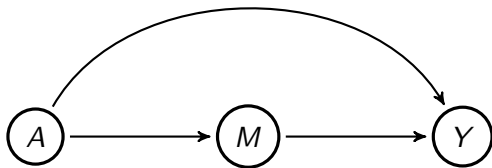
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Direct and Indirect Effects
*i.e. natural **direct** and **indirect** effects*

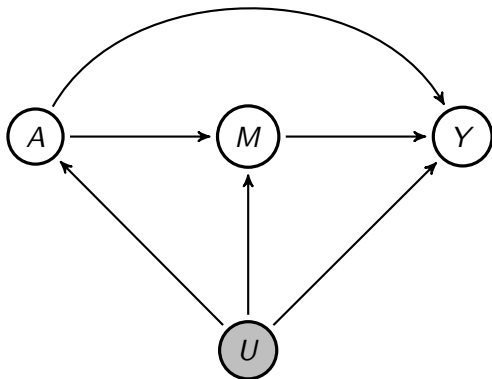
Background: classic mediation setting

Before discussing estimation of these effects, we need *identification*...



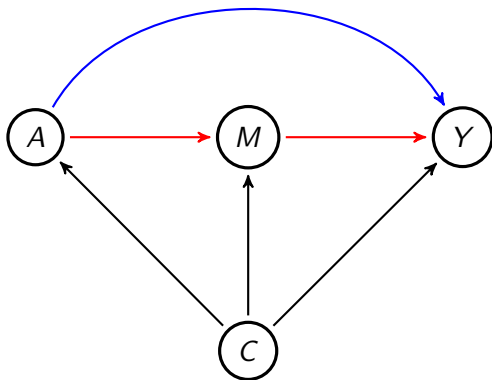
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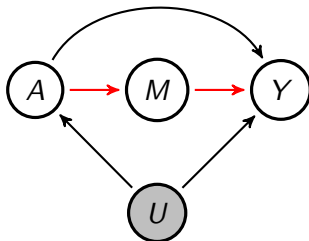
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- What if the exposure is not randomized?

We provide nonparametric identification for an **indirect effect** in the presence of unmeasured confounding of the exposure-outcome relation.



Motivating example: birth surveillance study

- Interest lies in quantifying the extent to which the effect of HIV infection on adverse birth outcomes (AE) operates through maternal medication use during pregnancy (ART)

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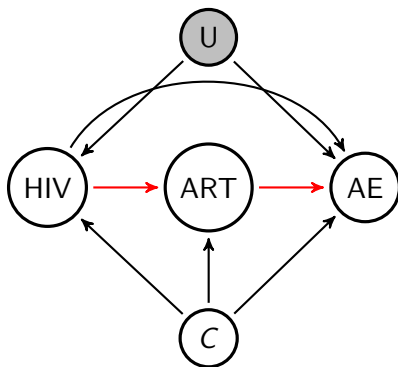
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As the exposure is not randomized, we choose to decompose a total effect known as the **Population Intervention Effect** (PIE) into direct and indirect components (Hubbard and Van Der Laan, 2008).

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The PIE measures the expected change of an outcome from its observed value in a population to that in the same observed population had contrary to fact no one been exposed

Notation and definitions

Define the following counterfactual variables,

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Decomposition of the PIE interest:

$$\begin{aligned} PIE &= E[Y] - E[Y(0, M(0))] \\ &= \underbrace{E[Y] - E[Y(A, M(0))]}_{\text{Population Intervention Indirect Effect}} + \underbrace{E[Y(A, M(0))] - E[Y(0, M(0))]}_{\text{Population Intervention Direct Effect}} \end{aligned}$$

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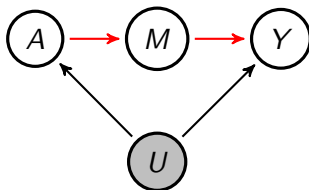
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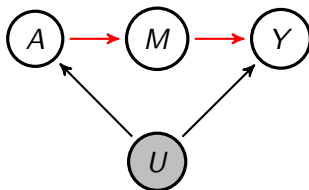
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Connection to Pearl's front-door criterion



- ① A1. $M(0) \perp A \mid C = c \quad \forall c$
- ② A2. $Y(a, m) \perp M \mid A = a, C = c \quad \forall m, a, c$
- ③ F1. $Y(a, m) = Y(0, m) = Y(m) \quad \forall a, m$

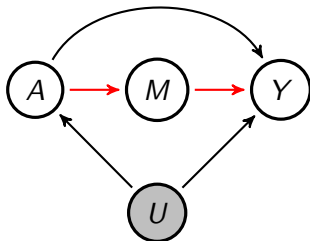
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$$PIIE = E(Y) - E(Y(0, M(0))) = PIE$$

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- ② A2. $Y(a, m) \perp M \mid A = a, C = c \quad \forall m, a, c$
- ③ A3. $Y(a, m) \perp M(0) \mid A = a, C = c \quad \forall m, a, c$

$$PIIE = E(Y) - E(Y(A, M(0)))$$

Nonparametric identification of the PIIE

Theorem (1)

If assumptions A1-A3 are satisfied, the population intervention indirect effect (PIIE) is nonparametrically identified. That is,

$$PIIE = E[Y] - E[Y(A, M(0))] = E[Y] - \Psi$$

$$\begin{aligned}\Psi = & \sum_{m,c} Pr(M = m|A = 0, C = c) \\ & \times \sum_a E(Y|A = a, M = m, C = c) \\ & \times Pr(A = a|C = c)Pr(C = c)\end{aligned}$$

Estimation of the PIIIE

Type	Model	Estimator
Parametric (<i>mle</i>)	$\mathcal{M}_{y,m,a}$	$\hat{\Psi}_{mle} = \frac{1}{n} \sum_{i=1}^n \left\{ \sum_m \hat{P}r(M_i = m A_i = a^*, C_i = c) \right. \\ \left. \times \sum_a \hat{E}(Y A_i = a, M_i = m, C_i = c) \hat{P}r(A_i = a C_i = c) \right\}$
Semiparametric (<i>sp1</i>)	\mathcal{M}_m	$\hat{\Psi}_{sp1} = \frac{1}{n} \sum_{i=1}^n Y_i \frac{\hat{f}(M_i 0, C_i)}{\hat{f}(M_i A_i, C_i)}$
Semiparametric (<i>sp1</i>)	$\mathcal{M}_{y,a}$	$\hat{\Psi}_{sp2} = \frac{1}{n} \sum_{i=1}^n \frac{I(A_i = 0)}{\hat{P}r(A_i = 0 C_i)} E(E\{Y A_i, M_i, C_i\} C_i)$
Doubly-robust semiparametric (<i>dr</i>)	$\mathcal{M}_{m \cup y, a}$	$\hat{\Psi}_{dr} = \frac{1}{n} \sum_{i=1}^n [Y - \hat{E}(Y A_i, M_i, C_i)] \frac{\hat{f}(M_i 0, C_i)}{\hat{f}(M_i A_i, C_i)} \\ + \frac{I(A_i = 0)}{\hat{P}r(A_i = 0 C_i)} \left(\sum_a \hat{E}(Y a, M_i, C_i) \hat{f}(a C_i) - \sum_{a, \bar{m}} \hat{E}(Y a, \bar{m}, C_i) \hat{f}(\bar{m} A_i, C_i) \hat{f}(a C_i) \right) \\ + \sum_m \hat{E}(Y A_i, m, C_i) \hat{f}(m 0, C_i)$

Simulation study setup

Goal of simulation study is two-fold:

- ① Verify robustness to exposure-outcome confounding for all confounders
- ② Assess performance under various forms of model misspecification

Simulation study setup

- Set of binary confounders C

$$C_1 \sim \text{Ber}(.6), \quad C_2 \mid C_1 \sim \text{Ber}(\text{expit}(1 + .5c_1)), \quad C_3 \sim \text{Ber}(.3)$$

- Binary exposure A

$$A \mid C_1, C_2 \sim \text{Ber}(\text{expit}(.5 + .2c_1 + .4c_2 + .5c_1c_2 + .2c_3))$$

- Continuous mediator M

$$M \mid A, C_1, C_2 \sim N(1 + a - 2c_1 + 2c_2 + 8c_1c_2, 4)$$

- Continuous outcome Y

$$Y \mid A, M, C_1, C_2 \sim N(1 + 2a + 2m - 8am + 3c_1 + c_2 + c_1c_2 + c_3, 1)$$

Simulation study setup

Simulations were run 10,000 times on a sample size of 1,000 under the following scenarios:

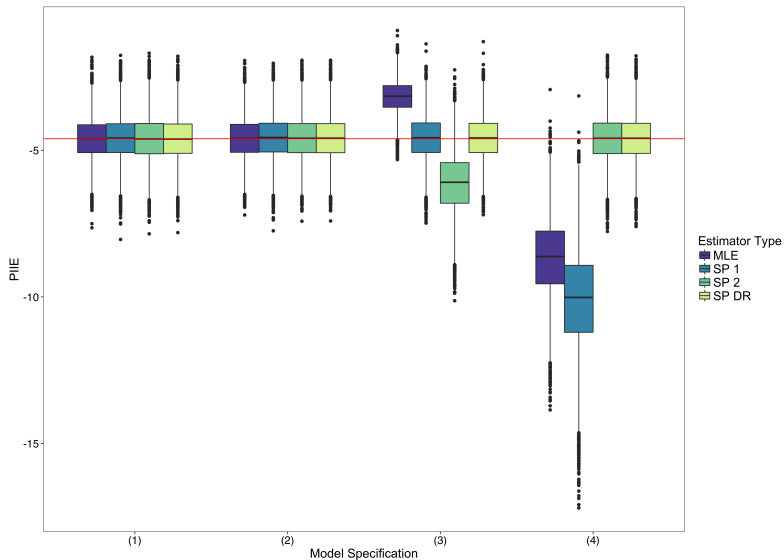
- (1) $\mathcal{M}_{y,m,a}$: models for A, M, Y correct
- (2) $\mathcal{M}'_{y,m,a}$: same as above with confounder of $A - Y$ relationship excluded from the models
- (3) \mathcal{M}_m : models for M correct^a
- (4) $\mathcal{M}_{y,a}$: models for A, Y correct^b

^a Y misspecified by leaving out AM interaction

^b M misspecified by leaving out interaction term between confounders

Simulation study results

Population Intervention Indirect Effect estimate by model and estimator



<http://bit.ly/enar-frontdoor>

Key references

- [1] Hubbard, A. E., & Van Der Laan, M. J. (2008). Population intervention models in causal inference. *Biometrika*, 95(1), 35-47.
- [2] Pearl, J. (2009). *Causality*. Cambridge University Press.
- [3] VanderWeele, T. (2015). *Explanation in causal inference: methods for mediation and interaction*. Oxford University Press.