# An Interval-Valued Neural Network Approach for Uncertainty Quantification in Short-Term Wind Speed Prediction

Ronay Ak, Valeria Vitelli, and Enrico Zio, Senior Member, IEEE

Abstract—We consider the task of performing prediction with neural networks (NNs) on the basis of uncertain input data expressed in the form of intervals. We aim at quantifying the uncertainty in the prediction arising from both the input data and the prediction model. A multilayer perceptron NN is trained to map interval-valued input data onto interval outputs, representing the prediction intervals (PIs) of the real target values. The NN training is performed by nondominated sorting genetic algorithm-II, so that the PIs are optimized both in terms of accuracy (coverage probability) and dimension (width). Demonstration of the proposed method is given in two case studies: 1) a synthetic case study, in which the data have been generated with a 5-min time frequency from an autoregressive moving average model with either Gaussian or Chi-squared innovation distribution and 2) a real case study, in which experimental data consist of wind speed measurements with a time step of 1 h. Comparisons are given with a crisp (single-valued) approach. The results show that the crisp approach is less reliable than the interval-valued input approach in terms of capturing the variability in input.

Index Terms—Interval-valued neural networks (NNs), multi-objective genetic algorithm (MOGA), prediction intervals (PIs), short-term wind speed forecasting, uncertainty.

#### I. INTRODUCTION

PREDICTION plays a crucial role in every decision-making process, and for this reason, it should consider any source of uncertainty that may affect its outcome. Prediction uncertainty can arise due to measurement errors, lack of knowledge in input data, and model approximation errors (e.g., due to imperfections in the model formulation) [1]–[3]. For practical purposes, uncertainties can be classified into two distinct types [3]: 1) epistemic and 2) aleatory. The former derives from imprecise model representation of the system behavior, in terms of uncertainty in both the hypotheses

Manuscript received March 25, 2014; revised July 21, 2014 and November 10, 2014; accepted January 9, 2015. Date of publication February 26, 2015; date of current version October 16, 2015.

R. Ak is with the Chair on Systems Science and the Energetic Challenge, European Foundation for New Energy—Electricité de France, École Centrale Paris, Châtenay-Malabry 92290, France (e-mail: ronay.ak@supelec.fr).

V. Vitelli is with the Oslo Center for Biostatistics and Epidemiology, Department of Biostatistics, University of Oslo, Oslo, Norway (e-mail: valeria.vitelli@medisin.uio.no).

E. Zio is with the Chair on Systems Science and the Energetic Challenge, European Foundation for New Energy—Electricité de France, École Centrale Paris, Châtenay-Malabry 92290, France, and also with the Department of Energy, Politecnico di Milano, Milan 20133, Italy (e-mail: enrico.zio@ecp.fr).

Color versions of one or more of the figures in this paper are available online at http://ieeexplore.ieee.org.

Digital Object Identifier 10.1109/TNNLS.2015.2396933

assumed (structural uncertainty) and the values of the model parameters (parameter uncertainty) [4]. The latter describes the inherent variability of the observed physical phenomenon, and it is, therefore, also named stochastic uncertainty, irreducible uncertainty, or inherent uncertainty [5].

Uncertainty quantification is the process of representing the uncertainty in the system inputs and parameters, propagating it through the model, and then revealing the resulting uncertainty in the model outcomes [2].

In the literature, methods such as evidence theory [5], probability modeling [6], estimation of neural network (NN)-based prediction intervals (PIs) [7]-[11], conformal prediction [12], [13], interval analysis [14]-[16], fuzzy set theory [17], and, in particular, type-2 fuzzy sets and systems [18]-[21], as well as extensions of fuzzy mathematical morphology [22], [23], Monte Carlo simulation [24], and Latin hypercube sampling [25], have been used to efficiently represent, aggregate, and propagate different types of uncertainty through computational models. Interval analysis is a powerful technique for bounding solutions under uncertainty. The uncertain model parameters are described by upper and lower bounds, and the corresponding bounds in the model output are computed using interval functions and interval arithmetic [26]. These bounds contain the true target value with a certain confidence level. The interval-valued representation can also be used to reflect the variability in the inputs (e.g., extreme wind speeds in a given area, minimum and maximum of daily temperature, and so on.), or their associated uncertainty (e.g., strongly skewed wind speed distributions), that is, to express the uncertain information associated with the input parameters [14]-[16], [27].

In this paper, we present an interval-valued time series prediction modeling framework based on a data-driven learning approach, more specifically a multilayer perceptron NN. Demonstration of the proposed method is given in two case studies: 1) a synthetic case study, with 5-min simulated data and 2) a real case study, involving hourly wind speed measurements. In both the cases, short-term prediction (1-h and day-ahead, respectively) is performed considering both the uncertainty in the model structure and the variability (within-hour and within-day, respectively) in the inputs.

The wind speed prediction case study has been chosen for its relevance to wind power production. Among the various renewable energy candidates, wind energy has received fast-growing attention throughout the world, and the utilization of wind power has increased dramatically over the past decade: the worldwide wind capacity has reached 296 GW by the end of June 2013, out of which 13980 MW have been added in the first half of 2013 [28]. This increasing integration of wind energy into power grid leads to additional uncertainty in the system due to the stochastic characteristics of wind itself. Wind power variations in short-term time scales have significant effects on power system operations, such as regulation, load following, balancing, unit commitment, and scheduling [8], [29], [30]. Thus, accurate prediction of wind speed and its uncertainty is critical for the safe, reliable, and economical operation of power systems [29], [30]. In other words, optimal integration of wind power into the grid requires highly accurate predictions with a reliable assessment of the uncertainties associated with the system. For this, PIs are a simple way to communicate a measure of the uncertainty in the predictions. PIs are preferable results of the prediction, rather than point estimates, because they provide information on the confidence in the prediction [7]-[11], considering the underlying uncertainties.

In this paper, an interval representation has been given to the hourly and daily inputs using two different approaches (see Section IV), which quantify both the within-hour and within-day variability in two different ways. The network maps interval-valued input data into an interval output, providing the estimated PIs for the real target. PIs are composed of lower and upper bounds within which the actual target is expected to lie with a predetermined probability [7]–[11]. The NN prediction model is trained by a multi-objective genetic algorithm (MOGA) (the powerful nondominated sorting genetic algorithm-II (NSGA-II)), so that the PIs are optimal both in terms of accuracy (coverage probability) and dimension (width).

The PI coverage probability (PICP) represents the probability that the set of estimated PI values will contain a certain percentage of the true output values. PI width (PIW) simply measures the extension of the interval as the difference of the estimated upper and lower bound values. The network uses interval-valued data, but its weights and biases are crisp (i.e., single-valued). The NSGA-II training procedure generates Pareto-optimal solution sets, which include nondominated solutions for the two objectives (PICP and PIW).

The originality of this paper appears in two aspects: 1) while the existing papers on short-term wind speed/power prediction use single-valued data as inputs, obtained as a within-hour [11], [29] or within-day average [31], [32], we give an interval representation to hourly/daily inputs using two approaches (Section IV), which properly account (in two different ways) for the within-hour/day variability and 2) we handle the PIs problem in a multi-objective framework [9], [11], [33], whereas the existing relevant methods for wind speed/power prediction [8] consider only one objective for optimization. In addition, the proposed approach integrates the estimation of the PIs in its learning procedure, while several methods construct PIs in two steps (first doing point prediction and then constructing PIs). In short, with this paper, we are able to account for both aleatory uncertainty in wind speed and epistemic uncertainty due to model parameters and to demonstrate an indication of how the uncertainties in input affect the output quantities, by the interval representation of the input variables.

It is worth recalling that in [9], we have compared the multi-objective genetic algorithm (MOGA) and singleobjective simulated annealing (SOSA) methods on a synthetic case study. The SOSA has been proposed in support of the lower and upper bound estimation (LUBE) method in [7]. The comparison results show that the PIs produced by NSGA-II compare well with those obtained by LUBE and are satisfactory in both the objectives of high coverage and small width. In [33], we have implemented the NSGA-II to train an NN to provide the PIs of the scale deposition rate. We have performed k-fold cross-validation to guide the choice of the NN structure (i.e., the number of hidden neurons) with good generalization performance. We have used a hypervolume indicator metric to compare the Pareto fronts obtained in each cross-validation fold. All these analyses have been performed with single-valued inputs for both works. For exemplification, in [11] single-valued historical wind speed values  $W_{t-1}$  and  $W_{t-2}, \ldots, W_{t-k}$  have been selected as input variables for predicting  $W_t$  in output. In [33], the case study concerns the scale (deposition) rate on the metal surfaces of equipment used in offshore oil wells. The output variable is the scale rate (y), and it has been predicted using the single-valued influencing input variables: temperature (T) and pressure (P), water composition (W), and fluid velocity (V) near the metal surfaces.

This paper is organized as follows. Section II introduces the basic concepts of interval-valued NNs for PIs estimation. In Section III, basic principles of multi-objective optimization are briefly recalled, and the use of NSGA-II for training an NN to estimate PIs is illustrated. Experimental results of the synthetic case study and of the real case study concerning wind speed prediction are given in Section IV. Finally, Section V concludes the paper with a critical analysis of the results and some ideas for future studies.

#### II. NEURAL NETWORKS AND PREDICTION INTERVALS

NNs are a class of nonlinear statistical models inspired by brain architecture, capable of learning complex nonlinear relationships among variables from the observed data [34]. This is done by a process of parameter tuning called training. It is common to think of an NN model as a way of solving a nonlinear regression problem of the kind [35], [36]:

$$y = f(x; w^*) + \varepsilon(x), \quad \varepsilon(x) \sim N(0, \sigma_{\varepsilon}^2(x))$$
 (1)

where x and y are the input and output vectors of the regression, respectively, and  $w^*$  represents the vector of values of the parameters of the model function f. The term  $\varepsilon(x)$  is the error associated with the regression model f, and it is assumed normally distributed with zero mean. For simplicity of illustration, in the following, we assume that y is 1-D. An estimate  $\hat{w}$  of  $w^*$  can be obtained by a training procedure aimed at minimizing the quadratic error function on a training set of input/output

values  $D = \{(x_i, y_i), i = 1, 2, ..., n_p\},\$ 

$$E(w) = \sum_{i=1}^{n_p} (\hat{y}_i - y_i)^2$$
 (2)

where  $\hat{y}_i = f(x_i; \hat{w})$  represents the output provided by the NN corresponding to the input  $x_i$ , and  $n_p$  is the total number of training samples.

A PI is a statistical estimator composed of upper and lower bounds that include a future unknown value of the target y(x) with a predetermined probability, called confidence level in [7]–[11].

To evaluate the quality of the PIs, we take the PICP and the PIW [7], [10] as measures: the former represents the probability that the set of estimated PIs will contain the true output values y(x) (to be maximized), and the latter simply measures the extension of the interval as the difference of the estimated upper bound and lower bound values (to be minimized). In general, these two measures are conflicting (i.e., wider intervals give larger coverage), but in practice it is important to have narrow PIs with high coverage probability (CP) [7].

When interval-valued data [26] are used as input, each input pattern  $x_i$  is represented as an interval  $x_i = [x_i^-, x_i^+]$ , where  $x_i^- \le x_i^+$  are the lower and upper bounds (real values) of the input interval, respectively. Each estimated output value  $\hat{y}_i$  corresponding to the *i*th sample  $x_i$  is, then, described by an interval as well,  $\hat{y}_i = [\hat{y}_i^-, \hat{y}_i^+]$ , where  $\hat{y}_i^-$  and  $\hat{y}_i^+$ ,  $(\hat{y}_i^- \leq \hat{y}_i^+)$ , are the estimated lower and upper bounds of the PI in output, respectively.

The mathematical formulation of the PICP and PIW measures given by [7] has been modified for interval-valued input and output data as follows:

$$PICP = \frac{1}{n_p} \sum_{i=1}^{n_p} c_i \tag{3}$$

where  $n_p$  is the number of training samples in the considered input data set, and

$$c_{i} = \begin{cases} \frac{1}{\operatorname{diam}(y_{i} \cap \hat{y}_{i})} & y_{i} \subseteq [\hat{y}_{i}^{-}, \hat{y}_{i}^{+}] \\ \frac{\operatorname{diam}(y_{i} \cap \hat{y}_{i})}{0} & y_{i} \not\subset [\hat{y}_{i}^{-}, \hat{y}_{i}^{+}] \wedge y_{i} \cap \hat{y}_{i} \neq \emptyset \end{cases}$$
(4)

where  $y_i = [y_i^-, y_i^+], y_i^- \le y_i^+$  are the lower and upper bounds (true values) of the output interval, respectively, and diam() indicates the width of the interval. For exemplification, (4) means that if the interval-valued real target is covered by the estimated PI, i.e., if the target is a subinterval of the estimated PI, then  $c_i$  is equal to 1. If the estimated PI does not cover the entire real target, but the intersection of the two is not empty, then  $c_i$  is equal to the ratio of diam $(y_i \cap \hat{y}_i)$  and the width of the interval  $y_i$ , and in that case,  $c_i$  takes a value smaller than 1. Finally, if the estimated PI does not cover the entire real target and the intersection of the two is empty, then the coverage  $c_i$  of the *i*th sample is 0. Due to lack of further information, this calculation corresponds to the probabilistic assumption that the target  $y_i$  can take any value in  $[y_i^-, y_i^+]$ 

with uniform probability, i.e., each point in  $[y_i^-, y_i^+]$  is equally likely to be a possible value of y.

For PIW, we consider the normalized quantity

NMPIW = 
$$\frac{1}{n_p} \frac{\sum_{i=1}^{n_p} (\hat{y}_i^+ - \hat{y}_i^-)}{y_{\text{max}} - y_{\text{min}}}$$
(5)

where NMPIW stands for normalized mean PIW, and  $y_{min}$  and  $y_{max}$  represent the minimum and maximum values of the true targets (i.e., the bounds of the range in which the true values fall). Normalization of the PIW by the range of targets makes it possible to objectively compare the PIs, regardless of the techniques used for their estimation or the magnitudes of the true targets.

Note that (3) is an empirical version of PICP, which yields an estimate of PICP according to the frequentist interpretation of probability theory. Similarly, (5) yields an estimate for NMPIW.

# III. NONDOMINATED SORTING GENETIC ALGORITHM-II (NSGA-II) MULTI-OBJECTIVE OPTIMIZATION FOR NEURAL NETWORK TRAINING

The problem of finding PIs optimal both in terms of CP and width can be formulated in a multi-objective optimization framework considering the two conflicting objectives PICP and NMPIW.

#### A. Multi-Objective Optimization by NSGA-II

In all generality, a multi-objective optimization problem considers a number of objectives,  $f_m, m = 1, 2, ..., M$ ,  $1, 2, \ldots, J$ , and equality  $h_k$ , inequality  $g_i$ , j= k = 1, 2, ..., K constraints, and bounds on the decision variables  $x_i$ , i = 1, 2, ..., I. Mathematically the problem can be written as follows [37]:

Minimise/Maximise 
$$f_m(x)$$
,  $m = 1, 2, ..., M$  (6)

subject to 
$$g_j(x) \ge 0, \quad j = 1, 2, ..., J$$
 (7)

$$h_k(x) = 0, \quad k = 1, 2, \dots, K$$
 (8)

$$h_k(x) = 0, \quad k = 1, 2, ..., K$$
 (8)  
 $x_i^{(l)} \le x_i \le x_i^{(u)}, \quad i = 1, 2, ..., I.$  (9)

A solution,  $x = \{x_1, x_2, \dots, x_I\}$ , is an *I*-dimensional decision variable vector in the solution space  $R^{I}$ , restricted by the constraints (7) and (8) and by the bounds on the decision variables (9).

The search for optimality requires that the M objective functions  $f_m(x)$ , m = 1, 2, ..., M are evaluated in correspondence to the decision variable vector x in the search space. The comparison of solutions during the search is performed in terms of the concept of dominance [37]. Precisely, in the case of a minimization problem, solution  $x_a$  is regarded to dominate solution  $x_b(x_a > x_b)$  if the following conditions are satisfied:

$$\forall i \in \{1, 2, \dots, M\}: f_i(x_a) \le f_i(x_b) \land$$
 (10)

$$\exists j \in \{1, 2, \dots, M\}: f_j(x_a) < f_j(x_b).$$
 (11)

If any of the above two conditions is violated, the solution  $x_a$  does not dominate the solution  $x_b$ , and  $x_b$  is said to be nondominated by  $x_a$ . Eventually, the search aims at identifying a set of optimal solutions  $x^* \in R^I$  which are superior to any other solution in the search space with respect to all objective functions, and which do not dominate each other. This set of optimal solutions is called Pareto-optimal set; the corresponding values of the objective functions form the so-called Pareto-optimal front in the objective functions space.

In this paper, we use GA for the multi-objective optimization. GA is a population-based metaheuristics inspired by the principles of genetics and natural selection [38]. It can be used for solving multi-objective optimization problems [39], [40]. Among the several options for MOGA, we adopt NSGA-II, as the comparative studies show that it is very efficient [38], [40], [41].

# B. Implementation of NSGA-II for Training an NN for Estimating PIs

In this paper, we extend the method described in [7] to a multi-objective framework for estimating output PIs from interval-valued inputs. More specifically, we use NSGA-II for finding the values of the parameters of the NN, which optimize two objective functions PICP (3) and NMPIW (5) in a Pareto-optimality sense (for ease of implementation, the maximization of PICP is converted to minimization by subtracting from one, i.e., the objective of the minimization is 1-PICP).

The practical implementation of NSGA-II on our specific problem involves two phases: 1) initialization and 2) evolution. These can be summarized as follows (for more details on the NSGA-II implementation see [33]).

## 1) Initialization Phase:

Step 1: Split the input data into training  $(D_{\text{train}})$  and testing  $(D_{\text{test}})$  subsets.

Step 2: Fix the maximum number of generations and the number of chromosomes (individuals)  $N_c$  in each population; each chromosome codes a solution by G real-valued genes, where G is the total number of parameters (weights) in the NN. Set the generation number n=1. Initialize the first population  $P_n$  of size  $N_c$ , by randomly generating  $N_c$  chromosomes

Step 3: For each input vector x in the training set, compute the lower and upper bound outputs of the  $N_c$  NNs, each one with G parameters.

Step 4: Evaluate the two objectives PICP and NMPIW for the  $N_c$  NNs (one pair of values 1-PICP and NMPIW for each of the  $N_c$  chromosomes in the population  $P_n$ ).

Step 5: Rank the chromosomes (vectors of G values) in the population  $P_n$  by running the fast nondominated sorting algorithm [41] with respect to the pairs of objective values, and identify the ranked nondominated fronts  $F_1, F_2, \ldots, F_k$  where  $F_1$  is the best front,  $F_2$  is the second best front, and  $F_k$  is the least good front.

Step 6: Apply to  $P_n$  a binary tournament selection based on the crowding distance [41], for generating an intermediate population  $S_n$  of size  $N_c$ .

Step 7: Apply the crossover and mutation operators to  $S_n$ , to create the offspring population  $Q_n$  of size  $N_c$ .

Step 8: Apply Step 3 onto  $Q_n$  and obtain the lower and upper bound outputs.

Step 9: Evaluate the two objectives corresponding to the solutions in  $Q_n$ , as in Step 4.

#### 2) Evolution Phase:

Step 10: If the maximum number of generations is reached, stop and return  $P_n$ . Select the first Pareto front  $F_1$  as the optimal solution set. Otherwise, go to Step 11.

Step 11: Combine  $P_n$  and  $Q_n$  to obtain a union population  $R_n = P_n \cup Q_n$ .

Step 12: Apply Steps 3–5 to  $R_n$  and obtain a sorted union population.

Step 13: Select the  $N_c$  best solutions from the sorted union to create the next parent population  $P_{n+1}$ .

Step 14: Apply Steps 6–9 onto  $P_{n+1}$  to obtain  $Q_{n+1}$ . Set n = n + 1; and go to Step 10.

Finally, the best front in terms of nondominance and diversity of the individual solutions is chosen. Once the best front is chosen, the testing step is performed on the trained NN with optimal weight values.

Note that, herein, the diversity corresponds to crowding distance [41]. Each solution i in the population has two attributes: 1) nondomination rank  $i_{rank}$  and 2) crowding distance  $i_{\text{distance}}$ . For a solution pair, i and j, we have  $i \prec_n j$ if  $i_{\text{rank}} < j_{\text{rank}}$  or  $(i_{\text{rank}} = j_{\text{rank}})$  and  $i_{\text{distance}} > j_{\text{distance}}$ . That is, if there are two solutions under consideration with different nondomination ranks, we prefer the one with the lower (better) rank. Otherwise, if both solutions have same ranking, i.e., belong to the same nondominated front, we select the solution which locates in the region with the smaller number of points. Note that the nondominant solutions are found by performing the fast nondominated sorting algorithm: the chromosomes (vectors of G values) in the population  $P_n$  are ranked by running the fast nondominated sorting algorithm [41] with respect to the pairs of objective values, and then we identify the ranked nondominated fronts  $F_1, F_2, \ldots, F_k$  where  $F_1$  is the best front,  $F_2$  is the second best front, and  $F_k$  is the least good front. Finally, in order to obtain the optimal Pareto front to be used in practice, we take the first 50 nondominated solutions in the first front  $F_1$ . Of course, one can select more solutions. For further explanations, we refer the readers to [41].

The total computational complexity of the proposed algorithm depends on two suboperations: 1) nondominated sorting and 2) fitness evaluation. The time complexity of nondominated sorting is  $O(M(N_c)^2)$ , where M is the number of objectives and  $N_c$  is the population size [41]. In the fitness evaluation phase, NSGA-II is used to train an NN which has  $n_p$  input samples. Since for each individual of the population a fitness value is obtained, this process is repeated  $N_c \times n_p$  times. Hence, time complexity of this phase is  $O(N_c \times n_p)$ . In conclusion, the computational complexity of one generation is  $O(MN_c)^2 + N_c \times n_p$ .

#### IV. EXPERIMENTS AND RESULTS

Two case studies have been considered: 1) a synthetic case study, consisting of four time series data sets generated according to different input variability scenarios and 2) a real case study concerning time series of wind speed data. The synthetic time series data sets have been generated with

a 5-min time frequency from an autoregressive moving average (ARMA) model with either Gaussian or Chi-squared innovation distribution. For what concerns the real case study, hourly measurements of wind speed for a period of 3 years (from 2010 to 2012) related to Regina, a region of Canada, have been used [42].

The synthetic case study is aimed at considering hourly data and the effects of within-hour variability. Hourly interval input data is obtained from the 5-min time series data by two different approaches, which we refer to as min-max and mean: the former obtains hourly intervals by taking the minimum and the maximum values of the 5-min time series data within each hour; the latter, instead, obtains one-standard deviation intervals  $[\overline{x_i} - s_i, \overline{x_i} + s_i]$  by computing the sample mean  $(\overline{x_i})$  and standard deviation  $(s_i)$  of each 12 within-hour 5-min data sample. Single-valued (crisp) hourly input have also been obtained as a within-hour average, i.e., by taking the mean of each 12 within-hour 5-min data sample, for comparison. The wind speed case study considers the effect of within-day variability, and min-max and mean approaches are applied to the 24 within-day hourly data samples.

The architecture of the NN model consists of one input, one hidden, and one output layer. The number of input neurons is set to 2 for both case studies, since an autocorrelation analysis [43] has shown that the historical past values  $x_{t-1}$ and  $x_{t-2}$  should be used as input variables for predicting  $x_t$ in output. The number of hidden neurons is set to 10 for the synthetic case study and to 15 for the real case study, after a trial-and-error process. The number of output neurons is 1 in the input-interval case, since in this case a single neuron provides an interval in output; conversely, in order to estimate PIs starting from crisp input data, the number of output neurons must be set equal to 2, to provide the lower and upper bounds. As activation functions, the hyperbolic tangent function is used in the hidden layer and the logarithmic sigmoid function is used at the output layer. We remark that all arithmetic calculations throughout the estimation process of the interval-valued NN have been performed according to interval arithmetic (interval product, sum, and so on).

To account for the inherent randomness of NSGA-II, five different runs of this algorithm have been performed and an overall best nondominated Pareto front has been obtained from the five individual fronts. To construct such best nondominated front, the first (best) front of each of the 5 runs is collected, and the resulting set of solutions is subjected to the fast nondominated sorting algorithm [41] with respect to the two objective functions. Then, the ranked nondominated fronts  $F_1, F_2, \ldots, F_k$  are identified, where  $F_1$  is the best front,  $F_2$  is the second best front, and  $F_k$  is the worst front. Solutions in the first (best) front  $F_1$  are then retained as the overall best front solutions. This procedure gives us the overall best nondominated Pareto front for the training set. After we have obtained this overall best front, we have performed testing using each solution included on the front.

For the first case study, the first 80% of the input data have been used for training and the rest for testing. For the second, a validation process has been performed. Thus, the

TABLE I

NSGA-II AND SOSA PARAMETERS USED
IN THE EXPERIMENTS

Parameter	Numerical value
MaxGen	300
$N_c$	50
$P_{m\_int}$	0.06
$P_c$	0.8
μ	0.9
$\eta$	50
T <sub>init</sub>	200
$T_{min}$	10 <sup>-50</sup>
CWC <sub>int</sub>	$10^{80}$
Geometric cooling schedule of SA	$T_{k+1} = T_k \times 0.95$

data set has been divided into three parts: 1) the first 60% is used for training; 2) 20% for validation; and 3) the remaining 20% for testing. All data have been normalized within the range [0.1, 0.9].

Table I contains the parameters of the NSGA-II for training the NN. MaxGen indicates the maximum number of generations which is used as a termination condition and  $N_c$  indicates the total number of individuals per population.  $P_c$  indicates the crossover probability and is fixed during the run.  $P_{m_{\text{int}}}$  is the initial mutation probability and it decreases at each iteration (generation) by

$$P_{m_{\text{int}}} \times e^{\left(-\frac{\text{gen}}{\text{MaxGen}}\right)}$$
. (12)

The average CPU times for both training and testing of NN have been recorded using MATLAB on a PC with 4 GB of RAM and a 2.53-GHz processor. The average CPU time for the entire training process with 300 generations takes  $\sim$ 5 h; whereas the construction of testing PIs, i.e., for the online prediction of PIs, is very fast, of the order of 1 min. It is needless to say that from the user point of view, the computational burden of the training phase is relatively less important [7], [8], since the training phase is, usually, only performed once. Note that the computational load is dependent on the complexity of the structure of the model (e.g., number of input neurons, hidden layers, and hidden neurons), the size of the data set, and the performance of the learning algorithm. Moreover, as we have performed interval arithmetic calculations (using the MATLAB INTLAB Version 6 toolbox), the total computation time for training the NN with intervalvalued inputs has significantly increased compared to the CPU times of the case study performed in [11] with single-valued inputs.

#### A. Synthetic Case Study

Four synthetic data sets have been generated according to the following model:

$$y(t) = f(t) + \delta(t) \tag{13}$$

where f(t) is the deterministic component and  $\delta(t)$  is the stochastic one, and the time horizon is 50 days which

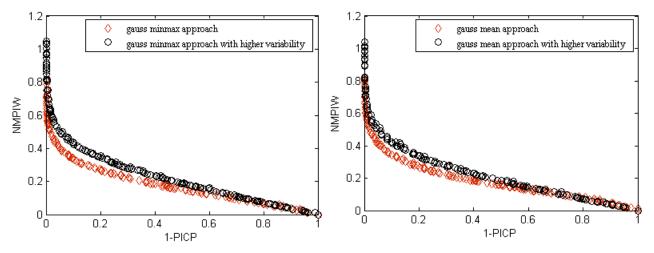


Fig. 1. Testing solutions for the Gaussian time series. Left: min-max approach. Right: mean approach.

makes 1200 h. The deterministic component has the following expression:

$$f(t) = 10 + 1.5 * \sin\left(\frac{2\pi t}{T_1}\right) + \sin\left(\frac{2\pi t}{T_2}\right)$$
 (14)

where the period  $T_1$  of the first periodic component has been set equal to 1 week, while  $T_2$  is 1 day. The stochastic component  $\delta(t)$  of the generating model in (13) is given by an ARMA(p,q) model [43], with p=2 autoregressive terms, with same coefficients  $\phi_1=\phi_2=0.1$ , and q=1 innovation term with coefficient given by  $\varphi_1=0.05$ . Four different scenarios are then considered, which differ in the distribution chosen for the innovation term, and in the higher or lower innovation variability: in two of the four scenarios the innovation is Gaussian, and has variance equal to 1 and 9, respectively, while in the other two scenarios the innovation has a Chi-squared distribution, with 2 or 5 degrees of freedom (corresponding to a variance equal to 4 and 10, respectively). We thus generate four different 5-min time series data sets, from which we will obtain either crisp or interval hourly data.

Fig. 1 shows the testing solutions corresponding to the first (best) Pareto front found after training the NN on interval data constructed by the min-max approach (left) and mean approach (right). The plots show the solutions for the data generated from a Gaussian distribution. On each plot, two testing fronts are illustrated: the ones where solutions are marked as circles have been obtained after training the NN on the interval data showing higher variability, while the ones with solutions marked as diamonds have been obtained after training the NN on the interval data having lower variability. Testing solutions obtained with data showing a lower variability are better than the ones with higher variability; hence, we can conclude that a higher variability in the input data may cause less reliable prediction results, and should thus be properly considered. Pareto fronts of solutions obtained for the data generated from a Chi-squared distribution are similar, and the results robust with respect to the choice of the innovation distribution.

Given the overall best Pareto set of optimal model solutions (i.e., optimal NN weights), it is necessary to select one

NN model for use. For exemplification purposes, a solution is here subjectively chosen as a good compromise in terms of high PICP and low NMPIW. The selected solution is characterized by 95% CP and a NMPIW equal to 0.420 for the min-max approach applied to lower variability Gaussian data. The results on the testing set give a CP of 95.5% and an interval width of 0.412. Fig. 2 shows 1-h ahead PIs for the selected Pareto solution, estimated on the testing set by the trained NN; the interval-valued targets included in the testing set are also shown in the figure.

Moreover, we also plot in Fig. 3 the 5-min original time series data (testing set), corresponding to the generating scenario with Gaussian distribution and low variability, together with the estimated PIs corresponding to the selected solution: the solid line shows the 5-min original time series data, while the dashed lines are the PIs, estimated starting from interval input data constructed with the min-max approach within each hour. Since the time step for the estimated PIs is 1 h, in order to compare them with the 5-min original time series data, we have shown the same lower and upper bounds within each hour in Fig. 3; thus, the PIs appear as a step function if compared with the original 5-min data.

In order to compare the Pareto front optimal solutions obtained with crisp and interval-valued inputs, a new normalized measure of the mean PIW, named NMPIW\*, has been *a posteriori* calculated as follows:

$$NMPIW^* = \frac{RT}{RRT} \times \frac{NRT}{0.8} \times NMPIW$$
 (15)

where RT, RRT, and NRT represent, respectively, the range of target (i.e., the range of the nonnormalized hourly training data in input), the range of real target (i.e., the range of the nonnormalized 5-min original time series data over the training set), and the range of normalized target (i.e., the range of the normalized hourly training data in input,  $y_{\text{max}} - y_{\text{min}}$ ). Note that, unless the synthetic scenario changes, RRT takes the same value for min–max, mean, and crisp approaches. The idea behind renormalization is to be able to compare PIs estimated from both interval and crisp approaches with respect to 5-min original time series data. As NMPIW for each

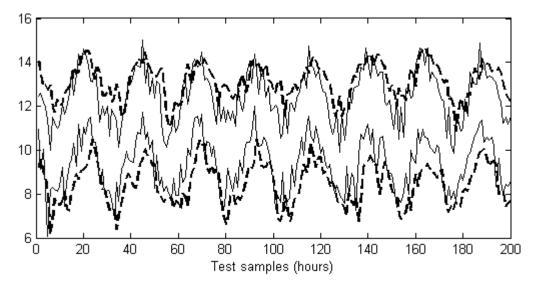


Fig. 2. Estimated PIs for 1-h ahead prediction on the testing set (dashed lines), and interval-valued input data (target) constructed by the min-max approach from the Gaussian distribution scenario with lower variability on the testing set (solid lines).

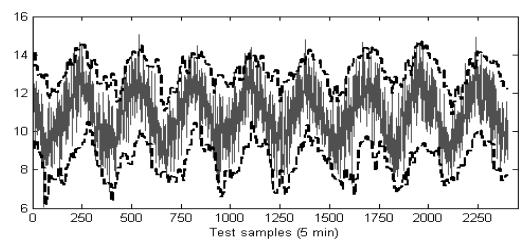


Fig. 3. Estimated PIs for 1-h ahead prediction on the testing set (dashed lines), and the original 5-min time series data on the testing set (solid line) obtained in the Gaussian distribution scenario with lower variability.

solution on the Pareto front has been calculated by dividing the mean PIW (MPIW) by the range of the training set in question, which is different for the two approaches, the Pareto fronts corresponding to the two approaches are not comparable. In order to analyze the performance of each approach with respect to 5-min original time series data, one should carry out a renormalization process that considers the range of the data set involved in the comparison, and which leads the estimated PIs to a common unit of measure. As a numerical example for the calculation of NMPIW\*, we have considered a testing solution, obtained on the synthetic data generated from the Gaussian distribution with lower variability and with the crisp approach, shown in Fig. 4. The selected solution results in a CP of 91% and an interval width of 0.328 on the testing. The values of RT, RRT, and NRT are 6.87, 11.383, and 0.647, respectively. Thus, using (16), we have obtained NMPIW\* as follows:

NMPIW\* = 
$$\frac{6.87}{11.383} \times \frac{0.647}{0.8} \times 0.328 = 0.16.$$
 (16)

Moreover, for each solution on each Pareto front, a PICP\* value has been a posteriori calculated. Equations (3) and (4) have been used with  $y_i$  representing nonnormalized 5-min original time series data, and with  $c_i = 1$ , if  $y_i \in [L(x_i), U(x_i)]$  and otherwise  $c_i = 0$ , where  $L(x_i)$  and  $U(x_i)$  indicate denormalized lower and upper bounds of the estimated PIs. Since estimated PIs have been obtained with hourly input data, while original data have a 5-min time frequency, in order to a posteriori calculate PICP\* with respect to the original data we have assumed the same lower and upper bounds,  $[L(x_i), U(x_i)]$ , for each 5-min time step within each hour. Renormalization allows us to convert current Pareto fronts to new ones whose CP and interval size are calculated according to the 5-min data set, and are comparable across different (crisp and interval) approaches.

In Fig. 4, a comparison between the testing fronts obtained with interval-valued and crisp inputs are shown. Solutions have been plotted according to the renormalized measures, i.e., the axes of the plots correspond to the new quantities NMPIW\* and 1-PICP\*, so that they can be compared. It can be

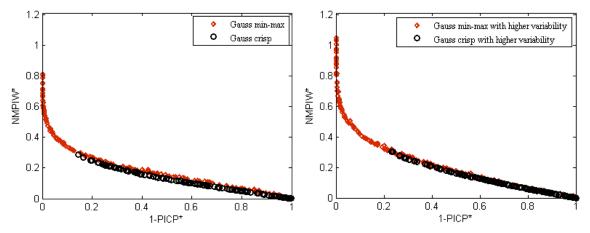


Fig. 4. Testing solutions obtained in the synthetic case study with interval-valued (min-max approach) and crisp approaches. Data have been generated from the Gaussian distribution with low (left) and higher (right) variability.

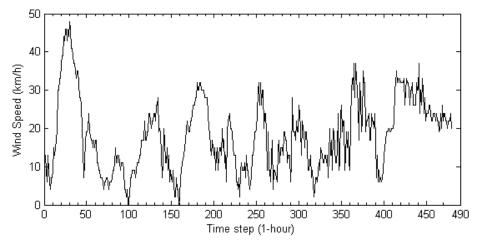


Fig. 5. Raw hourly wind speed data set used in this paper: first 20 days.

appreciated that the solutions obtained with a crisp approach never result in CPs >90% with respect to the original data. Furthermore, when the variability in the original data increases (right plots), the crisp approach gives less reliable results in terms of CP, which is <80%. However, a model should take the within-hour variability (high or low) into account and be capable of properly capturing it. Predictions resulting in a CP lower than expected show the poor prediction power of the crisp approach, which cannot be considered a reliable support to decision making in the presence of high variability.

# B. Real Case Study: Short-Term Wind Speed Prediction

In this section, results of the application of the proposed method to short-term wind speed forecasting with interval-input data are detailed. The data set considered for the analysis consists in hourly wind speed data measured in Regina, Saskatchewan, a region of central Canada. Canada finished 2014 with nearly 9,700 MW of total installed wind energy capacity, producing enough to power over 2 million homes and equivalent to  $\sim\!4\%$  of the total electricity demand in Canada [44]. The actual situation in Saskatchewan is characterized by the presence of four large wind farms located throughout the region, with a total capacity of  $\sim\!198$  MW [45].

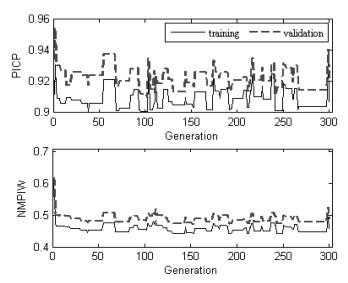


Fig. 6. Evaluation of PICP (top) and NMPIW (bottom) for min-max approach along MOGA iterations using training and validation sets.

The wind speed data set, covering the period from January 1, 2010, to December 30, 2012, has been downloaded from the website [42]. Since hourly data have been collected, 24 wind speed values are available for each day. Fig. 5 shows

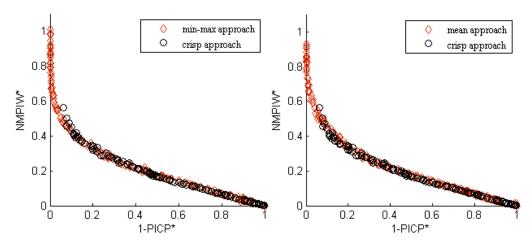


Fig. 7. Comparison between crisp and interval-valued approaches testing solutions, after renormalization, for day-ahead wind speed prediction. Left: min-max with respect to crisp approach comparison. Right: mean with respect to crisp approach comparison.

TABLE II
TRAINING, VALIDATION, AND TESTING RESULTS
OBTAINED BY NSGA-II

Tra	aining	Val	idation	T	esting
PICP (%)	NMPIW	PICP (%)	NMPIW	PICP (%)	NMPIW
90.1	0.440	91.0	0.470	91.4	0.452
90.3	0.446	92.0	0.474	92.6	0.461
90.6	0.452	91.6	0.482	92.6	0.466
91.6	0.456	92.5	0.486	93.3	0.471
92.1	0.466	93.1	0.494	93.9	0.480
93.2	0.487	94.3	0.514	95.1	0.500
93.6	0.493	94.4	0.526	94.8	0.506
94.3	0.529	96.0	0.562	96.6	0.546
96.9	0.578	97.7	0.606	98.3	0.595
97.9	0.636	98.9	0.673	98.9	0.651
98.5	0.662	99.2	0.692	99.3	0.679
99.2	0.721	99.5	0.757	99.7	0.739

the behavior of hourly wind speed values only in the first 20 days, for the sake of clarity: one can appreciate the within-day variability in each individual day. The wind speed changes from 0 to 72 km/h with an unstable behavior. From this raw hourly wind speed data, one can obtain daily interval wind speed data with the min-max and mean approach described at the beginning of Section IV. The so obtained data sets include 1095 intervals among which the first 60% is used for training, 20% for validation, and the remaining 20% for testing.

The procedure described in Sections II and III has been applied for day-ahead wind speed prediction, both with interval and crisp inputs. Crisp results are reported for comparison, in terms of daily averages of the raw hourly data, with the same data splitting for training, validation, and testing sets. For both interval and crisp cases, the historical wind speed values at previous time steps  $W_{t-1}$  and  $W_{t-2}$  have been used as inputs to estimate the output  $W_t$ .

When an optimal solution is selected from the front obtained by optimizing the NN on the basis of the training data, it is possible that the CP resulting from the application of this optimal NN to unseen data is lower than the one obtained on the training data. Thus, a validation set has been also selected, to test the generalization power of the proposed method. In other words, the aim is to test whether the selection of the solution with the required CP on the training data will result in well-calibrated PIs on the validation data or not. Fig. 6 shows the values of PICP and NMPIW obtained on the validation set along the iterations of the MOGA (for the min-max approach). To obtain these graphs, at each iteration of the training process, we have selected the solution from the training front which either results in 90% PICP or is closest to (and, if possible, above) 90%. Then, the selected solution has been used on the validation set, and the corresponding PICP and NMPIW values have been recorded. The motivation behind these plots is to show the capability of the MOGA algorithm to generate reliable predictions on unseen

Table II reports the PICP and NMPIW values of the selected training and validation solutions corresponding to those having CP between 90% and 100% on the overall best nondominated Pareto front. These solutions are obtained by the min–max approach. From inspection both of Table II and the profiles of both objectives on the training and validation sets shown in Fig. 6, we can observe that the training, validation, and testing results do not show significant difference. The PICP evaluation is coherent with NMPIW; hence, we can conclude that the proposed method results in well-calibrated PIs not only on the training set but also on the validation set.

In Fig. 7, the testing solutions obtained with the intervalvalued min–max and mean approaches, and with crisp inputs, are shown. The figure has been plotted according to the renormalized solutions, as explained in Section IV-A, i.e., the axes of the plot correspond to the new quantities NMPIW\* and 1-PICP\*. As already appreciated in the synthetic case study, one can notice that the solutions obtained with a crisp approach do not result in a CP >95% with respect to the original data. Furthermore, looking at the solutions in Fig. 7 which show a

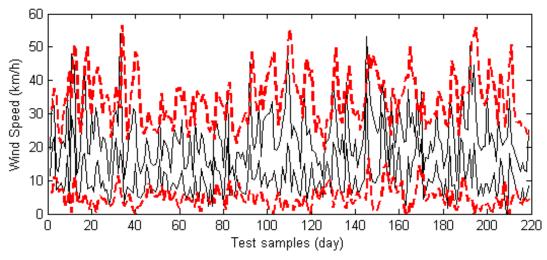


Fig. 8. Estimated PIs with interval inputs for day-ahead wind speed prediction on the testing set (dashed lines), and interval-valued wind speed data (constructed by the mean approach) included in the testing set (solid lines).

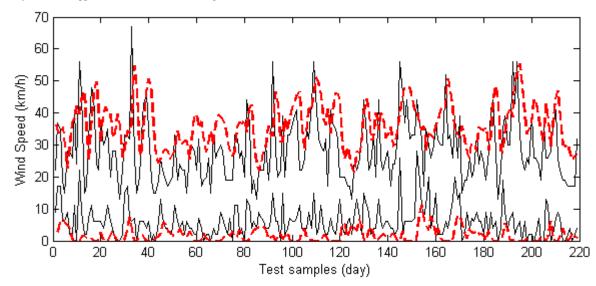


Fig. 9. Estimated PIs (dashed lines) with interval inputs for day-ahead wind speed prediction on the testing set and interval-valued wind speed data (constructed by the min-max approach) included in the testing set (solid lines).

CP >90%, the ones corresponding to the crisp approach give larger interval size. Since in practice it is important to have narrow PIs with high CP, an interval-inputs approach is more suited to reliable decision making.

From the overall best Pareto set of optimal solutions (i.e., optimal NN weights) obtained after training the network on the interval input data constructed with the min–max and mean approaches, a solution must be chosen. The selection of the solution might be accomplished by setting a constraint on one of the objective and choosing the optimal value for the other one, or by considering some other methods to weigh the two objectives [46]. In general, the selection should represent the preferences of the decision makers. Here, for simplicity's sake, we do not introduce any specific formal method of preference assignment, but subjectively choose a good compromise solution: for the min–max approach, the results give a CP of 92.1% and interval width of 0.466 on the training, and a CP of 93.9% and interval width of 0.480 on the testing. For the

mean approach, the selected solution results in a CP of 91.7% and interval width of 0.424 on the training, and a CP of 93% and interval width of 0.437 on the testing.

Figs. 8 and 9 report day-ahead PIs (dashed lines) for the selected Pareto solutions, with respect to the mean and min-max approaches, respectively, estimated on the testing set by the trained NN. The interval-valued targets (solid lines) included in the testing set are also shown in the figures. As wind speed cannot be negative, to reflect the real physical phenomena the negative lower bounds of the PIs have been replaced with zeros. From inspection of the figures, we observe that the target profile of the mean approach is more accurate than that of the min-max approach. However, the peak points have been covered relatively better by the min-max approach than by the mean approach. Hence, which one would be preferably chosen depends on the application. The mean approach might be considered more similar to classical methods for short-term wind speed/power prediction using single-valued

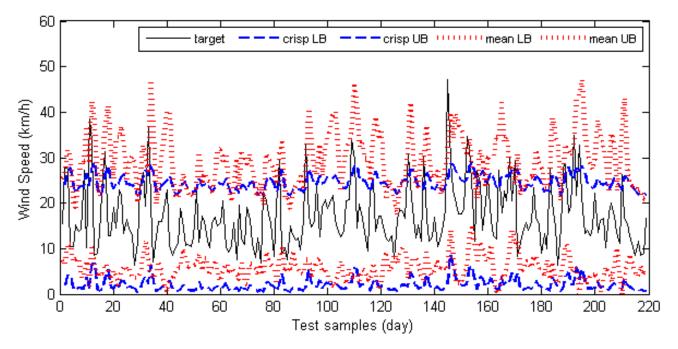


Fig. 10. Estimated PIs with interval (dotted lines) and crisp (dashed lines) inputs for day-ahead wind speed prediction on the testing set, and single-valued (crisp) daily wind speed data included in the testing set (solid line).

data as inputs, obtained as a within-hour or within-day average. By this approach, we can add information to the single-valued averages, and thus we can include in the model the potential uncertainty caused by the data itself showing a within-hour/day variability. Hence, the mean approach is a well-suited interval inputs alternative to the classical crisp inputs one, and it might be considered more feasible in practice.

In order to compare the interval-valued and crisp approaches in a clear way, we have shown the PIs obtained by both approaches in one figure. In Fig. 10, we have shown the estimated day-ahead PIs corresponding to the selected solutions obtained by mean and crisp approaches, respectively, on the daily crisp wind speed testing set by the trained NN. The solutions have been selected from the overall best Pareto set of optimal solutions obtained by mean and crisp approaches. These solutions result in 91.8% CP\* and 0.483 NMPIW\* for the mean approach, and in 91.3% CP and 0.495 NMPIW for the crisp approach, on the testing data set. It is clear that the solution obtained by the mean approach dominates the one obtained by the crisp approach. Note that PICP\* and NMPIW\* values have been a posteriori calculated only for the mean approach; as the crisp approach has been trained with the crisp daily wind speed training set, it is not necessary to convert PICP and NMPIW to PICP\* and NMPIW\* values.

Similarly, Fig. 11 has been plotted by considering *a posteriori* calculated PICP\* and NMPIW\* values (Fig. 7) corresponding to the two solutions selected from the overall best Pareto fronts of min-max and crisp approaches, respectively. These solutions result in 91.4% CP\* and 0.452 NMPIW\* for min-max approach, and 91.2% CP\* with 0.472 NMPIW\* for crisp approach, on the testing data set (raw hourly wind speed data). It is obvious that the solution obtained by min-max approach is superior to the one obtained

by crisp approach. In other words, we have obtained higher quality PIs with interval-valued input approach. Note that this comparison is done on the raw hourly wind speed data set. Since the time step for the estimated PIs is 1 day, in order to compare them with the hourly original time series data, we have shown in Fig. 11 the same lower and upper bounds within each day; thus, the PIs appear as a step function if compared with the original 1-h data. Due to space limitations, we have only plotted the estimated PIs obtained by min–max approach.

From the results shown in Figs. 10 and 11, one might comment that the PIs obtained with the interval inputs approach are capable of capturing the peak points (highest and lowest) of the target of interest (hourly data). Although there are some highly extreme values dropping out of the estimated PIs, the interval approach leads to better coverage of the intermittent characteristic of wind speed than the crisp approach. In other words, the interval approach manages to describe more efficiently the short-term variability of wind speed.

# C. Comparison With Single-Objective Simulated Annealing (SOSA) Method

In this section, we present the results from a comparison with a method called LUBE proposed by Khosravi *et al.* in [7] to estimate PIs with single-valued (crisp) inputs. In their paper, SOSA algorithm was used to train the NN and adopted the cost function defined in (17), which combines PICP and NMPIW, to be minimized.

The cost function proposed in [7] is called coverage width-based criterion (CWC) and defined as follows:

$$CWC = NMPIW(1 + \gamma (PICP)e^{-\eta (PICP - \mu)})$$
 (17)

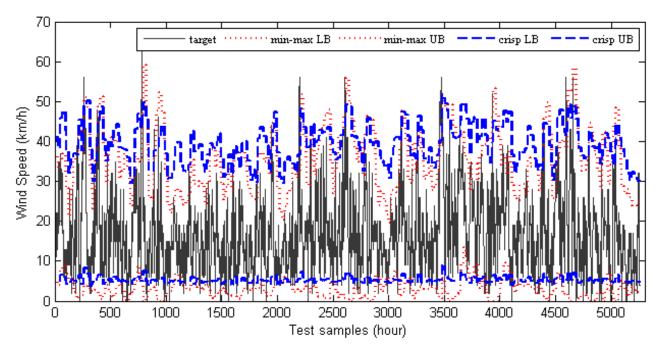


Fig. 11. Estimated PIs with interval (dotted lines) and crisp (dashed lines) inputs for day-ahead wind speed prediction on the testing set, and single-valued (crisp) raw hourly wind speed data (solid line).

where  $\eta$  and  $\mu$  are constants. The role of  $\eta$  is to magnify any small difference between  $\mu$  and PICP. The value of  $\mu$  gives the nominal confidence level, which is set to 90% in our experiments (Table I). Then,  $\eta$  and  $\mu$  are the two parameters determining how much penalty is paid by the PIs with low CP. The function  $\gamma$  (PICP) is equal to 1 during training, whereas in the testing of the NN is given by the following step function:

$$\gamma (PICP) = \begin{cases} 0, & PICP \ge \mu \\ 1, & PICP < \mu. \end{cases}$$
 (18)

To compare SOSA and the proposed MOGA method, we ran the SOSA using the same interval-valued wind speed training data. For SOSA, the initial temperature has been determined after a trial-and-error procedure. It has been tried with values of 5, 200, and 500: it turns out that the SOSA with an initial temperature of 200 gives best performance. Table I contains the parameters of the SOSA; the maximum number of generation has been set to 500.

The training process has been repeated five times. Training and testing results in each run have been reported in Table III. Due to space limitation, we have put only the min-max approach results.

According to the results reported in Table III, the training and corresponding testing solutions do not show high consistency in terms of CP and interval size among the five runs performed. In other words, there is a high difference among the results: SOSA gives high CP value in one run, whereas it generates less accurate PIs in another one; 3 out of 5 runs give CP values smaller than the predetermined nominal confidence level, that is, 90% in our experiments. Thus, CWC values are quite high for those runs. Although the existing works done based on the SOSA LUBE method with single-valued inputs show promising results for the construction of PIs [7]–[10],

TABLE III
PICP AND NMPIW VALUES OBTAINED BY SOSA
WITH RESPECT TO WIND SPEED DATA
SET (TRAINING/TESTING)

SOSA METHOD	PICP (%)	NMPIW	CWC
1	93.8 / 95.6	0.567 / 0.578	0.649 / 0.578
2	71.7 / 73.8	0.300 / 0.312	2897 / 1032
3	72.0 / 75.2	0.297 / 0.310	2425 / 519.6
4	75.5 / 76.3	0.317 / 0.328	439.0 / 311.3
5	92.1 / 95.1	0.725 / 0.752	0.978 / 0.752

the results reported in this paper (Table III) demonstrate a drawback about SOSA method's robustness on this specific problem.

For comparison purpose, we have selected the run giving the smallest CWC value on the training set, which is 0.649. Note that in previous works of [7] and [47], the mean or median value of several runs has been used as final prediction result. The selected run results in 93.8% CP and 0.567 NMPIW on the training set, and a CP of 95.6% and interval width of 0.578 on the testing. By comparison, we have selected a solution from the overall best Pareto front obtained by MOGA min-max approach. This selected solution gives a CP of 94.3% and interval width of 0.529 on the training, and a CP of 96.6% and interval width of 0.546 on the testing. For what concerns the mean approach, we have observed similar results: 2 out of the 5 runs have given CP less than 90% both on training and testing sets. The runs resulting in CP of >90% have quite large interval widths (>50%). We have selected a run that has the smallest CWC value: it has a CP of 93.1% with 0.520 NMPIW on the training and 94.7% CP and interval width of 0.531 on the testing data sets. In contrast, the MOGA method has given

a solution with 93.3% CP with 0.440 interval size on the training, and 94.4% CP with 0.453 interval size on the testing set.

It is clear that the solutions obtained by MOGA dominate the best ones obtained by SOSA. It is worth pointing out that as both solutions obtained by min-max method give large interval sizes (~50%), they cannot provide useful information in practice because the uncertainty level is too high to support a reliable and informed decision in typical application contexts. However, with the MOGA approach one can select a solution from the Pareto front giving tight PIW with a high CP, which satisfies the predetermined nominal confidence level. In short, from the results reported in Table III, one can conclude that the SOSA method does not give high quality PIs with respect to the interval-valued time series forecasting case study considered in this paper.

### V. CONCLUSION

The goal of the research presented in this paper is to quantitatively represent the uncertainty in NNs predictions of time series data, originating both from variability in the input and in the prediction model itself. The application focus has been on wind speed, whose forecasting is crucial for the energy market, system adequacy, and service quality in power grid with integrated wind energy systems. Accuracy of predictions of power supply and quantitative information on the related uncertainty is relevant both for the power providers and the system operators.

Specifically, we have presented two approaches that can be used to process interval-valued inputs with multilayer perceptron NNs. The method has been applied on a synthetic case study and on a real case study, in which the data show a high short-term variability (within-hour and within-day). The results obtained reveal that the interval-valued input approach is capable of capturing the variability in the input data with the required coverage. The results enable different strategies to be planned according to the range of possible outcomes within the interval forecast.

As for future research, the use of an ensemble of different NNs will be considered to further increase the accuracy of the predictions, and type-2 fuzzy sets can be integrated into the proposed model as an alternative way to represent the input uncertainty.

## REFERENCES

- [1] J. C. Refsgaard, J. P. van der Sluijs, J. Brown, and P. van der Keur, "A framework for dealing with uncertainty due to model structure error," *Adv. Water Resour.*, vol. 29, no. 11, pp. 1586–1597, 2006.
- [2] H. Cheng, "Uncertainty quantification and uncertainty reduction techniques for large-scale simulations," Ph.D. dissertation, Dept. Comput. Sci. Appl., Virginia Polytechnic Inst. State Univ., Blacksburg, VA, USA, 2009.
- [3] E. Zio and T. Aven, "Uncertainties in smart grids behavior and modeling: What are the risks and vulnerabilities? How to analyze them?" *Energy Policy*, vol. 39, no. 10, pp. 6308–6320, 2011.
- [4] N. Pedroni, E. Zio, and G. E. Apostolakis, "Comparison of bootstrapped artificial neural networks and quadratic response surfaces for the estimation of the functional failure probability of a thermal-hydraulic passive system," *Rel. Eng. Syst. Safety*, vol. 95, no. 4, pp. 386–395, 2010.
- [5] H. Agarwal, J. E. Renaud, E. L. Preston, and D. Padmanabhan, "Uncertainty quantification using evidence theory in multidisciplinary design optimization," *Rel. Eng. Syst. Safety*, vol. 85, nos. 1–3, pp. 281–294, 2004.

- [6] J. C. Helton, "Uncertainty and sensitivity analysis in the presence of stochastic and subjective uncertainty," J. Statist. Comput. Simul., vol. 57, nos. 1–4, pp. 3–76, 1997.
- [7] A. Khosravi, S. Nahavandi, D. Creighton, and A. F. Atiya, "Lower upper bound estimation method for construction of neural networkbased prediction intervals," *IEEE Trans. Neural Netw.*, vol. 22, no. 3, pp. 337–346, Mar. 2011.
- [8] H. Quan, D. Srinivasan, and A. Khosravi, "Short-term load and wind power forecasting using neural network-based prediction intervals," *IEEE Trans. Neural Netw. Learn. Syst.*, vol. 25, no. 2, pp. 303–315, Feb. 2014.
- [9] R. Ak, Y.-F. Li, and E. Zio, "Estimation of prediction intervals of neural network models by a multi-objective genetic algorithm," in Proc. 10th Int. FLINS Conf. Uncertainty Model. Knowl. Eng. Decision Making (FLINS), Istanbul, Turkey, Aug. 2012, pp. 1036–1041.
- [10] A. Khosravi, S. Nahavandi, D. Creighton, and A. F. Atiya, "Comprehensive review of neural network-based prediction intervals and new advances," *IEEE Trans. Neural Netw.*, vol. 22, no. 9, pp. 1341–1356, Sep. 2011.
- [11] R. Ak, V. Vitelli, and E. Zio, "Uncertainty modeling in wind power generation prediction by neural networks and bootstrapping," in *Proc. Esrel Conf.*, Amsterdam, The Netherlands, Sep./Oct. 2013, pp. 3191–3196.
- [12] H. Papadopoulos, V. Vovk, and A. Gammerman, "Regression conformal prediction with nearest neighbours," *J. Artif. Intell. Res.*, vol. 40, no. 1, pp. 815–840, 2011.
- [13] H. Papadopoulos and H. Haralambous, "Reliable prediction intervals with regression neural networks," *Neural Netw.*, vol. 24, no. 8, pp. 842–851, 2011.
- [14] L. V. Barboza, G. P. Dimuro, and R. H. S. Reiser, "Towards interval analysis of the load uncertainty in power electric systems," in *Proc. Int. Conf. Probab. Methods Appl. Power Syst.*, Ames, IA, USA, Sep. 2004, pp. 538–544.
- [15] A. M. San Roque, C. Maté, J. Arroyo, and Á. Sarabia, "iMLP: Applying multi-layer perceptrons to interval-valued data," *Neural Process. Lett.*, vol. 25, no. 2, pp. 157–169, 2007.
- [16] A. L. S. Maia, F. de A. T. de Carvalho, and T. B. Ludermir, "Forecasting models for interval-valued time series," *Neurocomputing*, vol. 71, nos. 16–18, pp. 3344–3352, 2008.
- [17] H.-J. Zimmermann, Fuzzy Set Theory and Its Applications. Norwell, MA, USA: Kluwer, 2001.
- [18] D. Zhai and J. M. Mendel, "Uncertainty measures for general type-2 fuzzy sets," *Inf. Sci.*, vol. 181, no. 3, pp. 503–518, 2011.
- [19] Q. Liang and J. M. Mendel, "Interval type-2 fuzzy logic systems: Theory and design," *IEEE Trans. Fuzzy Syst.*, vol. 8, no. 5, pp. 535–550, Oct. 2000
- [20] F. Gaxiola, P. Melin, F. Valdez, and O. Castillo, "Interval type-2 fuzzy weight adjustment for backpropagation neural networks with application in time series prediction," *Inf. Sci.*, vol. 260, pp. 1–14, Mar. 2014.
- [21] J. A. Sanz, A. Fernandez, H. Bustince, and F. Herrera, "IVTURS: A linguistic fuzzy rule-based classification system based on a new interval-valued fuzzy reasoning method with tuning and rule selection," *IEEE Trans. Fuzzy Syst.*, vol. 21, no. 3, pp. 399–411, Jun. 2013.
- [22] P. Sussner, M. Nachtegael, T. Mélange, G. Deschrijver, E. Esmi, and E. Kerre, "Interval-valued and intuitionistic fuzzy mathematical morphologies as special cases of L-fuzzy mathematical morphology," *J. Math. Imag. Vis.*, vol. 43, no. 1, pp. 50–71, 2012.
- [23] M. Nachtegael, P. Sussner, T. Mélange, and E. E. Kerre, "On the role of complete lattices in mathematical morphology: From tool to uncertainty model," *Inf. Sci.*, vol. 181, no. 10, pp. 1971–1988, 2011.
- [24] E. Hofer, M. Kloos, B. Krzykacz-Hausmann, J. Peschke, and M. Woltereck, "An approximate epistemic uncertainty analysis approach in the presence of epistemic and aleatory uncertainties," *Rel. Eng. Syst. Safety*, vol. 77, no. 3, pp. 229–238, 2002.
- [25] J. C. Helton and F. J. Davis, "Latin hypercube sampling and the propagation of uncertainty in analyses of complex systems," *Rel. Eng. Syst. Safety*, vol. 81, no. 1, pp. 23–69, 2003.
- [26] R. E. Moore, R. B. Kearfott, and M. J. Cloud, *Introduction to Interval Analysis*. Philadelphia, PA, USA: SIAM, 2009.
- [27] J. Arroyo and C. Maté, "Introducing interval time series: Accuracy measures," in *Proc. 17th Comput. Statist. Symp.*, Rome, Italy, Sep. 2006, pp. 1139–1146.
- [28] World Wind Energy Association. Half-Year Report 2013. [Online]. Available: http://www.wwindea.org/wwea-publishes-half-year-report-2013-2/, accessed Oct. 2013.

- [29] A. U. Haque, P. Mandal, M. E. Kaye, J. Meng, L. Chang, and T. Senjyu, "A new strategy for predicting short-term wind speed using soft computing models," *Renew. Sustain. Energy Rev.*, vol. 16, no. 7, pp. 4563–4573, 2012.
- [30] J. Jung and R. P. Broadwater, "Current status and future advances for wind speed and power forecasting," *Renew. Sustain. Energy Rev.*, vol. 31, pp. 762–777, Mar. 2014.
- [31] H. Bouzgou and N. Benoudjit, "Multiple architecture system for wind speed prediction," *Appl. Energy*, vol. 88, no. 7, pp. 2463–2471, 2011.
- [32] A. More and M. C. Deo, "Forecasting wind with neural networks," *Marine Struct.*, vol. 16, no. 1, pp. 35–49, 2003.
- [33] R. Ak et al., "NSGA-II-trained neural network approach to the estimation of prediction intervals of scale deposition rate in oil & gas equipment," Expert Syst. Appl., vol. 40, no. 4, pp. 1205–1212, 2013.
- [34] M. Pulido, P. Melin, and O. Castillo, "Genetic optimization of ensemble neural networks for complex time series prediction," in *Proc. Int. Joint Conf. Neural Netw.*, San Jose, CA, USA, Jul./Aug. 2011, pp. 202–206.
- [35] E. Zio, "A study of the bootstrap method for estimating the accuracy of artificial neural networks in predicting nuclear transient processes," *IEEE Trans. Nucl. Sci.*, vol. 53, no. 3, pp. 1460–1478, Jun. 2006.
- [36] D. L. Shrestha and D. P. Solomatine, "Machine learning approaches for estimation of prediction interval for the model output," *Neural Netw.*, vol. 19, no. 2, pp. 225–235, 2006.
- [37] Y. Sawaragi, H. Nakayama, and T. Tanino, "Theory of multiobjective optimization," in *Mathematics in Science and Engineering*, R. Bellman, Ed. Orlando, FL, USA: Academic, 1985.
- [38] A. Konak, D. W. Coit, and A. E. Smith, "Multi-objective optimization using genetic algorithms: A tutorial," *Rel. Eng. Syst. Safety*, vol. 91, no. 9, pp. 992–1007, 2006.
- [39] E. Zitzler and L. Thiele, "Multiobjective evolutionary algorithms: A comparative case study and the strength Pareto approach," *IEEE Trans. Evol. Comput.*, vol. 3, no. 4, pp. 257–271, Nov. 1999.
- [40] N. Srinivas and K. Deb, "Muiltiobjective optimization using nondominated sorting in genetic algorithms," *Evol. Comput.*, vol. 2, no. 3, pp. 221–248, 1994.
- [41] K. Deb, A. Pratap, S. Agarwal, and T. Meyarivan, "A fast and elitist multiobjective genetic algorithm: NSGA-II," *IEEE Trans. Evol. Comput.*, vol. 6, no. 2, pp. 182–197, Apr. 2002.
- [42] Canadian Weather Office. *Data Download for Regina*. [Online]. Available: http://regina.weatherstats.ca/download.html, accessed Apr. 2013.
- [43] G. E. P. Box, G. M. Jenkins, and G. C. Reinsel, *Time Series Analysis: Forecasting and Control*. Hoboken, NJ, USA: Wiley, 2008.
- [44] Canadian Wind Energy Association (Canwea). [Online]. Available: http://canwea.ca/wind-energy/installed-capacity/, accessed Jan. 2015.
- [45] SaskPower. Annual Report 2012. [Online]. Available: http://www.saskpower.com/about-us/current-reports/past-annual-reports/?linkid= MM\_past\_annual\_reports, accessed Apr. 2013.

- [46] E. Zio, P. Baraldi, and N. Pedroni, "Optimal power system generation scheduling by multi-objective genetic algorithms with preferences," *Rel. Eng. Syst. Safety*, vol. 94, no. 2, pp. 432–444, 2009.
- [47] A. Khosravi and S. Nahavandi, "Combined nonparametric prediction intervals for wind power generation," *IEEE Trans. Sustain. Energy*, vol. 4, no. 4, pp. 849–856, Oct. 2013.

Ronay Ak received the Ph.D. degree in Energy and Power Systems (Engineering) discipline from the Chair on Systems Science and the Energetic Challenge, European Foundation for New Energy—Électricité de France, École Superieure d'Électricité, Paris, France, in 2014.

She is currently a Guest Researcher at the National Institute of Standards and Technology (NIST), Gaithersburg, MD, USA. Her current research interests include uncertainty quantification, big data analytics, machine learning algorithms, reliability analysis of wind-integrated power networks, and multi-objective optimization.

Valeria Vitelli received the Ph.D. degree in Mathematical Models and Methods for Engineering, with a focus on statistical models for classification of high-dimensional data, from Politecnico di Milano, Milan, Italy, in 2012.

She was a Post-Doctoral Researcher with the Chair on Systems Science and the Energetic Challenge, European Foundation for New Energy—Électricité de France, École Centrale Paris, Châtenay-Malabry, France, and the École Supérieure d'Électricité, Paris, France, from 2012 to 2013. She is currently a Post-Doctoral Researcher with the Department of Biostatistics, University of Oslo, Oslo, Norway. Her current research interests include prediction and uncertainty quantification methods for complex data.

Enrico Zio (SM'06) received the Ph.D. degree in nuclear engineering from the Politecnico di Milano, Milan, Italy, in 1995, and the Massachusetts Institute of Technology, Cambridge, MA, USA, in 1998.

He is currently the Director of the Chair on Systems Science and the Energetic Challenge, European Foundation for New Energy—Électricité de France, École Centrale Paris, Châtenay-Malabry, France, and the École Superieure d'Électricité, Paris, France, and a Full Professor with the Politecnico di Milano. His current research interests include the characterization and modeling of the failure/repair/maintenance behavior of components, complex systems and their reliability, maintainability, prognostics, safety, vulnerability, and security, Monte Carlo simulation methods, soft computing techniques, and optimization heuristics.