

Compulsory Assignment 1

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Problem 1

Let $\mu = E(X) = \begin{pmatrix} 0 \\ 2 \end{pmatrix}$ and $\Sigma = cov(X) = \begin{pmatrix} 3 & 1 \\ 1 & 3 \end{pmatrix}$ s.t.

$$Y = \begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} = \begin{pmatrix} \frac{1}{\sqrt{2}} & -\frac{1}{\sqrt{2}} \\ \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \end{pmatrix} X$$

1a

(i) We want to find the mean vector and the covariance vector of Y .

$$E(Y) = E(AX) = AE(X) = \begin{pmatrix} -\frac{2}{\sqrt{2}} \\ \frac{2}{\sqrt{2}} \end{pmatrix}$$

$$\begin{aligned} cov(Y) &= cov(AX) = Acov(X)A^T \\ &= \begin{pmatrix} 2 & 0 \\ 0 & 4 \end{pmatrix} \end{aligned}$$

(ii) The distribution of Y is a bivariate normal distribution, where

$$Y \sim N(E(Y), cov(Y))$$

(iii) We can observe that Y_1 and Y_2 is independent since

$$cov(Y_1, Y_2) = 0$$

1b

Let the pdf be given as the equation of an ellipse s.t.

$$\begin{aligned} f(x) &= a, \quad a > 0 \\ (x - \mu)^T \Sigma^{-1} (x - \mu) &= b. \end{aligned}$$

The relation of b and a can be derived as follows,

$$\begin{aligned} f(x) &= k \cdot \exp\left(- (x - \mu)^T \Sigma (x - \mu)\right) = a \\ \ln k - \ln a &= (x - \mu)^T \Sigma (x - \mu) \end{aligned}$$

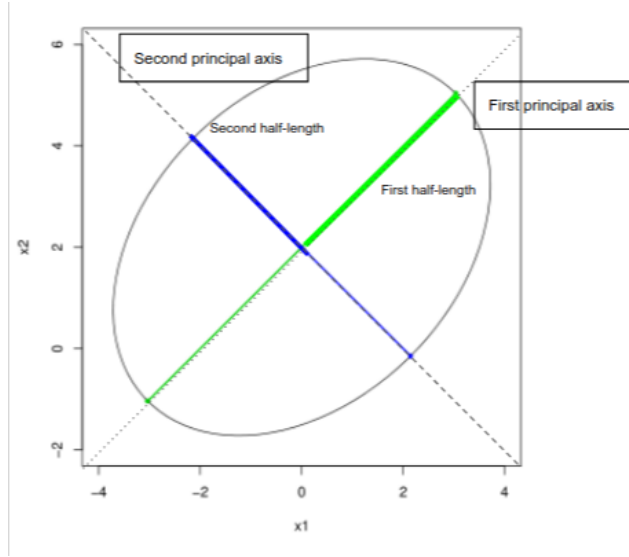
Thus $b = \ln k - \ln a$, where $k = \frac{1}{2\pi |\Sigma|}$. Clearly, we can observe that the alignment of the ellipse is oriented along the eigenvectors of Σ . Furthermore, the half lengths are described by the scalar b and eigenvalues

$$l_1 = \sqrt{b} \sqrt{\lambda_1} \quad \text{and} \quad \sqrt{b} \sqrt{\lambda_2}.$$

Since $(x - \mu) \Sigma^{-1} (x - \mu)$ is a sum of normal distributed variables can we compute the probability a random variable being inside the ellipse α by using the fact that

$$(x - \mu)^{-1} \Sigma (x - \mu) \sim \chi_2^2.$$

Hence, the probability can be computed using $\chi_2^2(\alpha) \leq b \iff \alpha \approx 0.9$.



Problem 2

2a

Let $X = [X_1, X_2, X_3, \dots, X_n]^T$ be a stochastic vector and a vector of ones $\mathbf{1} = \mathbf{1}_{n \times 1}$.

(i)

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} \mathbf{1}^T [X_1, \dots, X_n]^T = \frac{1}{n} \mathbf{1}^T X$$

(ii)

$$\begin{aligned} S^2 &= \frac{1}{(n-1)} X^T C X = \frac{1}{(n-1)} X^T C C X \\ &= \frac{1}{(n-1)} (C X)^T (C X) \\ &= \frac{1}{(n-1)} (X - \mathbf{1} \bar{X})^T (X - \mathbf{1} \bar{X}) \\ &= \frac{1}{(n-1)} \sum_{i=1}^n (X_i - \bar{X}) (X_i - \bar{X}) \end{aligned}$$

2b

We want to show the independence of \bar{X} and S^2 . Firstly, let us emphasize the result that

$$\frac{1}{n} \mathbf{1}^T (C) = \frac{1}{n} \mathbf{1}^T \left(I - \frac{\mathbf{1} \mathbf{1}^T}{n} \right) = \frac{1}{n} \mathbf{1}^T - \frac{1}{n} \mathbf{1}^T = 0.$$

And utilize the fact that

$$\text{cov}(\bar{X}, S^2) = \text{cov} \left(\frac{1}{n} \mathbf{1}^T X, C X \right) = \frac{1}{n} \mathbf{1}^T \sigma I C = \sigma \cdot 0.$$

Hence, \bar{X} and S^2 are independent.

2c

Let $Y = \Sigma^{-1} (X - \mu)$ and

$$\begin{aligned} YCY^T &= (\Sigma^{-1} (X - \mu)) C (\Sigma^{-1} (X - \mu))^T = \frac{1}{\sigma^2} (X - \mu) C (X - \mu)^T \\ &= \frac{1}{\sigma^2} X^T CX \end{aligned}$$

The last step comes from that $(C\mu)_i = 0$.

1 References