# UC Irvine ISI-BUDS 2023 Day 09: GLM

Zhaoxia Yu

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### UC Irvine ISI-BUDS 2023 Day 09: GLM

Zhaoxia Yu

Learning Objectives

Review of LM

Logistic Regression

Poisson Regression

## **Learning Objectives**

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## **Learning Objectives**

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- Review of LM
- ► GI M
  - Logistic Regression
  - Poisson Regression
  - Multinomial Regression
- The assumption of independent observations

### Review of LM

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The Assumption of Independence

Suppose

$$Y = \beta_0 + x_1 \times \beta_1 + \ldots + x_p \times \beta_p + \epsilon,$$

### where

- the regressand Y is the response / outcome / dependent / endogenous variable
- ▶ the regressors  $(x_1, \dots, x_p)$  are the p covariates / independent / explanatory variables
- the random term  $\epsilon$  has a zero mean and variance  $\sigma^2>0$
- the intercept is  $\beta_0$ , the other p coefficients are  $\beta_1, \dots, \beta_p$

Consider the *i*th observation:

$$Y_i = \beta_0 + x_{i1} \times \beta_1 + \ldots + x_{ip} \times \beta_p + \epsilon_i, i = 1, \ldots, n$$

- Basic assumptions
  - ►  $E(\epsilon_i) = 0$ , which is equivalent to  $E(Y_i|X_i) = \beta_0 + x_{i1} \times \beta_1 + ... + x_{ip} \times \beta_p$
  - Var $(\epsilon_i) = \sigma^2$ . Note, this is equivalent to say  $Var(Y_i|X_i) = \sigma^2$ .
  - $ightharpoonup (\epsilon_1, \cdots, \epsilon_n)$  are mutually independent
- ▶ If  $(\epsilon_1, \dots, \epsilon_n)$  are i.i.d.  $N(0, \sigma^2)$ , we can derive t-tests and F-tests
- Question: what if the assumptions are violated?

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## **Logistic Regression**

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- ▶ A motivating example: Consider a binary response variable, i.e., *Y<sub>i</sub>* takes values of 0 or 1.
- ▶ Is LM a good choice for this problem?

## A Motivating Example of GLM (continued)

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```
Consider the Alzheimer data
```

We create a binary variable

```
alzheimer=read.csv("alzheimer_data.csv", header = TRUE)
#dim(alzheimer)
#names(alzheimer)
attach(alzheimer)
#length(unique(id))
alzh=(diagnosis>0)*1 #"*1" to create a 0-1 variable
```

## A Motivating Example of GLM (continued)

alzh.lm = lm(alzh ~ age + female + educ+lhippo + rhippo)

plot(alzh, predict(alzh.lm)); abline(h=c(0,1), col=2)

par(mar = c(4, 4, 0.5, 0.5))

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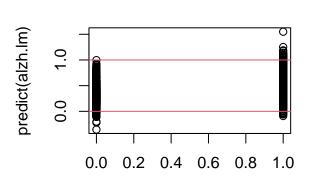
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alzh

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- ► Is alzh.lm a good model for alzh?
- Several assumptions of the LM have been violated, and
- ▶ The predicted values using LM are not between 0 and 1!
- Let  $X_i = (x_{i1}, \dots, x_{ip})^T$ , i.e., the vector of covariates for the *i*th subject.
- Let  $\pi_i = E(Y_i|X_i)$ , the expected probability. We would like to make sure that  $\pi_i \in [0,1]$
- ► How?

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▶ Consider the a special transformation of  $\pi_i$ :

$$logit(\pi_i) = log \frac{\pi_i}{1 - \pi_i} \in (-\infty, \infty)$$

- ► This is the so-called "logit" link!
- $\pi_i = E[Y_i|X_i]$ : probability of having AD for a subject with covariates  $X_i$ .
- $ightharpoonup \frac{\pi_i}{1-\pi_i}$ : odds

We connect  $\pi_i$  and a linear function of the covariates  $X_i$  by assuming

$$\log \frac{\pi_i}{1 - \pi_i} = \beta_0 + x_{i1} \times \beta_1 + \ldots + x_{ip} \times \beta_p$$

- Essentially, we model the log-odds.
- $\triangleright$  But  $Y_i$  is a random variable. We need a distribution. A natural choice is the Bernoulli distribution

$$Y_i|X_i \sim Bernoulli(\pi_i)$$

pmf, mean, variance:

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### Regression

- Estimation of is typically conducted by maximizing the corresponding likelihood function
- How to obtain the likelihood function

$$E(Y_i|X_i) = \pi_i = \frac{\exp\{\beta_0 + x_{i1} \times \beta_1 + \dots + x_{ip} \times \beta_p\}}{1 + \exp\{\beta_0 + x_{i1} \times \beta_1 + \dots + x_{ip} \times \beta_p\}}$$

$$f(Y_i|X_i) = \pi_i^{Y_i} (1-\pi_i)^{1-Y_i}$$
, i.e.,

$$f(Y_i|X_i) = \pi_i \text{ if } Y_i = 1$$

• 
$$f(Y_i|X_i) = 1 - \pi_i \text{ if } Y_i = 0$$

▶ independence: 
$$f(Y|X) = \prod_{i=1}^n f(Y_i|X_i)$$

$$L(\beta_0, \beta_1, \cdots, \beta_p) = f(Y|X)$$

- In the previous slide we model  $E(Y_i|X_i)$ . Because  $Y_i$  is binary, we have  $E(Y_i|X_i) = Pr(Y_i = 1|X_i)$ .
- Retrospective studies are often considered because a prospective study might take many years and is costly.
- In a retrospective study, subjects are recruited based on their disease status. Let z=1 denote being sampled and z=0 otherwise. Let

$$Pr(z = 1|y = 0) = p_0$$
  
 $Pr(z = 1|y = 1) = p_1$ 

For a retrospective study, the logistic regression models Pr(y = 1|z = 1, x)

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# Logistic Regression for Retrospective Studies

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- Does this affect the interpretation of the parameters?
- Let  $\theta = Pr(y=1|x)$  and  $\phi = Pr(y=1|z=1,x)$ . By Bayes' theorem and assuming that z does not dependent on x,

$$\begin{split} \phi &= Pr(y=1|z=1,x) \\ &= \frac{Pr(z=1|y=1,x)Pr(y=1|x)}{Pr(z=1|y=1,x)Pr(y=1|x) + Pr(z=1|y=0,x)Pr(y=0|x)} \\ &= \frac{\rho_1 \theta}{\rho_1 \theta + \rho_0 (1-\theta)} \end{split}$$

# Logistic Regression for Retrospective Studies

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► Therefore

$$log(rac{\phi}{1-\phi}) = log(rac{p_1 heta}{p_0(1- heta)}) = log(p_1/p_0) + log(rac{ heta}{1- heta})$$

- ► The result suggests that, when using logistic regression,
  - the only difference between a prospective study and a retrospective study would be the intercept.
  - the inference for the other parameters is still valid even though the subjects were recruited based on their disease status (such as a retrospective case-control study)

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## **Logistic Regression**

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- ▶ How to obtain the maximum likelihood estimates (MLE) of the parameters  $(\beta_0, \dots, \beta_p)$ ?
  - ► Iteratively re-weighted least squares (IRLS): the default method used by R
  - ► The Newton-Raphson algorithm

# The Motivating Example of Logistic Regression

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```
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```

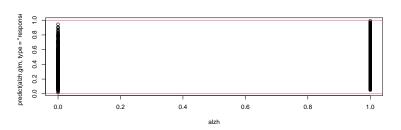
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```
alzh.glm = glm(alzh ~ age + female + educ+lhippo, family=binomial)
par(mar = c(4, 4, 0.5, 0.5))
plot(alzh, predict(alzh.glm, type="response")); abline(h=c(0,1), col=2)
```



#More visualizaitons #https://blogs.uoregon.edu/rclub/2016/04/05/plotting-your-logistic-regression-models/

### Interpreting a logistic regression

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```
summary(alzh.glm)$coefficients[-1,]
```

```
## Estimate Std. Error z value Pr(>|z|)
## age 0.01813761 0.004246088 4.271605 1.940715e-05
## female -1.32020475 0.096534651 -13.675968 1.413151e-42
## educ -0.05640342 0.013279326 -4.247461 2.162067e-05
## lhippo -1.98502114 0.114028821 -17.408065 7.166544e-68
```

### Interpreting a logistic regression

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• What is the  $\beta_0$  in a logistic regression?

- ► Consider the age variable. The estimated coefficient is 0.018138. What information does it provide?
- ► The estimated log-odds AD for subject i is (or add a constant determined by study design, as the previous slides)

$$loigt(\hat{\pi}_i) = \hat{\beta}_0 + \hat{\beta}_{age}age_i + \hat{\beta}_2 female_i + \hat{\beta}_3 educ_i + \hat{\beta}_4 lhippo_i$$

Let  $\tilde{\pi}_i$  denote estimated log-odds after one year

$$loigt(\tilde{\pi}_i) = \hat{\beta}_0 + \hat{\beta}_{age}(age_i + 1) + \hat{\beta}_2 female_i + \hat{\beta}_3 educ_i + \hat{\beta}_4 lhippo_i$$

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The estimated change in log-odds

$$logit(\tilde{\pi}_i) - logit(\hat{\pi}_i) = log \frac{\tilde{\pi}_i}{1 - \tilde{\pi}_i} - log \frac{\hat{\pi}_i}{1 - \hat{\pi}_i} = 0.018138$$

Take exponential of both sides, we have

$$\frac{\frac{\tilde{\pi}_i}{1-\tilde{\pi}_i}}{\frac{\hat{\pi}_i}{1-\hat{\pi}_i}} = exp(0.018138)$$

- The odds of AD in one year later is exp(0.018138) = 1.018303 times of the current odds.
- The estimated increase in odds of AD in a year is  $e^{0.018138} 1 = 1.8303\%$

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### Interpreting a logistic regression

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- A 95% confidence interval
  - First, obtain a 95% C.I. for the difference in log-odds: (0.018138-1.96\*0.004246, 0.018138+1.96\*0.004246) = (0.00982, 0.0265)
  - Then, we transform them to increase in odds:  $(e^{0.00982} 1, e^{0.0265} 1) = (0.99\%, 2.69\%)$

- What if we are interested in the increase in odds of AD in ten years (everything else is fixed)?
- ▶ The estimated increase in odds of AD in 10 years is

$$e^{10*0.018138} - 1 = 19.89\%$$

► A 95% C.I. for 10-year increase in odds:

```
exp(10*c(0.018138-1.96*0.004246, 0.018138+1.96*0.004246))-1
```

## [1] 0.1031375 0.3029118

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- Very often, we also want to know the significance of a variable after adjusting for other important covariates?
- Does age show a significant effect after adjusting for gender, education, and hippocampus volume?
- $\blacktriangleright$  A test for  $H_0: \beta_{age} = 0$  using the Wald test (a type of large-sample test)

```
summary(alzh.glm)$coefficients["age",]
```

```
Estimate
               Std. Error
                                            Pr(>|z|)
1.813761e-02.4.246088e-03.4.271605e+00.1.940715e-05
```

Other tests, such as likelihood ratio test, can also be used

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The Assumption of Independence

 Recall that we used the <u>logit</u> link in the logistic regression

$$g(\pi_i) = logit(\pi_i) = \frac{\pi_i}{1 - \pi_i},$$

where  $\pi_i = E(Y_i|X_i)$ .

- ▶ How about LM?  $g(\mu_i) = \mu_i$ , i.e., LM uses the identity link
- ▶ Poisson  $g(\lambda_i) = log(\lambda_i)$ .  $Y_i | X_i \sim Poisson(\lambda_i)$ .

# Poisson Regression

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## Poisson Regression: The Model

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- ▶ Poisson regression is often used to model count data
- Why are count data special?
  - Count data are non-negative
  - Count data take integer values
- Count data often violate the assumption of "constant variance"
  - Count data often follow a Poisson distribution
  - ► Consider  $K \sim Poisson(\lambda)$ . E(K) = ?, Var(K) = , pmf?

## Poisson Regression: Motivating Example

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- Neurons may <u>fire</u> selectively for particular types of stimuli
- ➤ To understand whether a neuron is a visual-selective neuron, 20 trials were run for each of the five image categories:
  - ▶ animal, fruit, kids, military, space
- ► In each trial, the number of spikes (the number of times that the neuron fired) within a 1-second window was recorded

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```
library(tidyverse)
#https://www.ics.uci.edu/-zhaoxia/Data/chosen_neuron_data.csv
chosen_neuron_data <- read_csv(
    "https://www.ics.uci.edu/-zhaoxia/Data/chosen_neuron_data.csv")
chosen_neuron_data <- chosen_neuron_data[, c(2:4)]
dim(chosen_neuron_data)
```

## [1] 100

names(chosen\_neuron\_data)

```
## [1] "trial_number" "n_spikes" "image_categ"
```

## **Poisson Regression**

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```
attach(chosen_neuron_data)
# Even split of image categories among trials
table(image_categ)
```

## image\_categ

## Animal Fruit Kids Military Space ## 20 20 20 20 20

sapply(split(n\_spikes, image\_categ), mean)

##	Animal	Fruit	Kids Mi	ilitary	Space
##	0.05	3.60	0.15	0.25	0.05

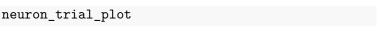
# Poisson Regression: Visualize the count data (by image category)

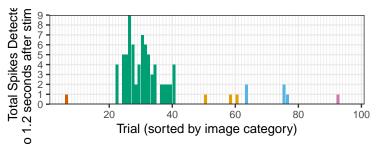
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Animal

Fruit

Kids

Military

Image Category

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## **Poission Regression: Model Summary**

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```
poisson_neuron_table
```

```
Pr(>|z|)
##
            Estimate Std. Error
                                  z value
## Animal
            -2.995732 0.9999998 -2.995733 2.737861e-03
## Fruit
            1.280934 0.1178511 10.869084 1.618171e-27
## Kids
            -1.897120
                      0.5773503 -3.285908 1.016541e-03
## Military -1.386294
                      0.4472132 -3.099851 1.936181e-03
## Space
            -2.995732
                      0.9999998 -2.995733 2.737861e-03
```

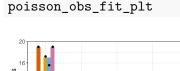
## Poisson Regression: Visualize Observed v.s. **Fitted**

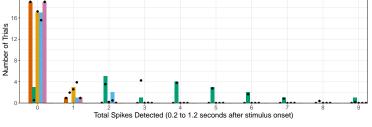
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 Poisson Distribution Fit Image Category Animal

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- Next, we would like to discuss the significance of the image\_categ variable. To do so, we first look at the deviance of a GLM object
- ▶ The deviance of a GLM object obj is

$$2[log(L_{saturated}) - log(L_{obj})]$$

What is the null hypothesis of no visual-selection?

$$H_0: \beta_{Animal} = \beta_{Fruit} = \beta_{Kids} = \beta_{Military} = \beta_{Space}$$

▶ What is the d.f. in a likelihood ratio test (LRT)?

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 Consider two <u>nested</u> models obj1 and obj2, the difference in their deviances is

$$2[log(L_{obj2}) - log(L_{obj1})],$$

which is the LRT statistic.

- What is the saturated model?
  - ▶ Logistic:  $\pi_i = y_i$  and  $L_{saturated} = 1$
  - **Poisson:**  $\lambda_i = y_i$  and  $L_{saturated} = \prod_i \frac{y_i^{y_i} e^{-y_i}}{y_i!}$ .

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```
## Analysis of Deviance Table
##
## Model 1: n_spikes ~ 1
## Model 2: n_spikes ~ i mage_categ - 1
## Resid. Df Resid. Dev Df Deviance Pr(>Chi)
## 1 99 260.985
## 2 95 81.213 4 179.77 < 2.2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1</pre>
```

poisson fit0 = glm(n spikes ~1. data=chosen neuron data, family=poisson(link="log"))

# Test for visual selectivity: Likelihood Ratio Test

anova(poisson\_fit0, poisson\_fit, test = "LRT")

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```
# Test for visual selectivity: Rao's score test
anova(poisson_fit0, poisson_fit, test = "Rao")
```

```
## Analysis of Deviance Table
##
## Model 1: n_spikes - 1
## Model 2: n_spikes - 1 image_categ - 1
## Resid. Dev Df Deviance Rao Pr(>Chi)
## 1 99 260.985
## 2 95 81.213 4 179.77 236.3 < 2.2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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```
## [,1]
## [1,] 0
```

## Poission Regression: Model Interpretation

```
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```

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```

```
\verb|summary(poisson_fit)$| coefficients|
```

- $\hat{\beta}_{Fruit} = 1.2809$ : What does it tell us?
- Recall that we used the log link

# Poission Regression: Model Interpretation

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- Note that the model poisson\_fit does not include  $\beta_0$ .
- ► That's why we can estimate  $\beta_{Animal}$ ,  $\beta_{Fruit}$ ,  $\beta_{Kids}$ ,  $\beta_{Military}$ ,  $\beta_{space}$ .
- Question: how should we interprete the estimated coefficients if the intercept term was included?
  - Try poisson\_fit\_repara <- glm(n\_spikes ~ image\_categ, data = chosen\_neuron\_data, family = poisson(link="log"))</p>
  - Are the two models equivalent? (Lab activity)

# Poisson Regression: Model Interpretation

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- ► Re-parameterization
- ▶ Parameters:
  - poisson\_fit:
  - poisson\_fit\_repara:
- Compare the summary of the two models:
- ► Are they different models?

### Parameterization 1: without intercept

```
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```

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```
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## The Assumption

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The Assumption of Independence
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```
summary(poisson fit)
##
## Call:
## glm(formula = n_spikes ~ image_categ - 1, family = poisson(link = "log"),
      data = chosen neuron data)
##
##
## Deviance Residuals:
      Min
##
                10
                     Median
                                  30
                                          Max
## -2 6833 -0 5876 -0 3162 -0 3162
                                       2.3861
##
## Coefficients:
##
                      Estimate Std. Error z value Pr(>|z|)
## image_categAnimal
                       -2.9957
                                  1.0000 -2.996 0.00274 **
## image_categFruit
                       1.2809
                                 0.1179 10.869 < 2e-16 ***
## image categKids
                       -1.8971 0.5774 -3.286 0.00102 **
## image categMilitary -1.3863
                                  0.4472 -3.100 0.00194 **
## image_categSpace
                       -2.9957
                                   1.0000 -2.996 0.00274 **
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
  (Dispersion parameter for poisson family taken to be 1)
##
      Null deviance: 264.439 on 100 degrees of freedom
  Residual deviance: 81.213 on 95 degrees of freedom
## ATC: 163.18
##
```

## Number of Figher Cooring itemsticas, 6

### Parameterization 2: with intercept

## glm(formula = n\_spikes ~ image\_categ, family = poisson(link = "log"),

summary(poisson fit repara)

data = chosen neuron data)

## ## Call:

##

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```
##
## Deviance Residuals:
      Min
##
                10
                     Median
                                 30
                                         Max
## -2 6833 -0 5876 -0 3162 -0 3162
                                      2.3861
##
## Coefficients:
##
                        Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                      -2.996e+00 1.000e+00 -2.996 0.00274 **
## image_categFruit
                       4.277e+00 1.007e+00 4.247 2.16e-05 ***
## image categKids
                     1.099e+00 1.155e+00 0.951 0.34139
## image categMilitary 1.609e+00 1.095e+00 1.469 0.14178
## image_categSpace
                      -3.140e-16 1.414e+00
                                             0.000 1.00000
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
##
  (Dispersion parameter for poisson family taken to be 1)
##
      Null deviance: 260.985 on 99 degrees of freedom
## Residual deviance: 81.213 on 95 degrees of freedom
## ATC: 163.18
##
## Number of Figher Cooring itemsticas, 6
```

Review of LM

- ▶ The Poisson regression we fit provides estimates of  $\beta_{Animal}$ ,  $\beta_{Fruit}$ ,  $\beta_{Kids}$ ,  $\beta_{Military}$ ,  $\beta_{space}$ , which are the log of the Poisson rates
- What if we are interested in difference between specific groups? e.g.,
  - $\begin{array}{l} & \beta_{\textit{Fruit}} \beta_{\textit{Animal}} \\ & \underline{\beta_{\textit{Fruit}} + \beta_{\textit{Animal}} + \beta_{\textit{Kids}} + \beta_{\textit{Military}} + \beta_{\textit{Space}}} \end{array}$

  - $\beta_{Fruit} \frac{\beta_{Animal} + \beta_{Kids} + \beta_{Military} + \beta_{Space}}{4}$
- ▶ They are linear functions of the coefficients, i.e., in the form of  $a^T \beta$ , where a is a 5-by-1 vector.

### **Linear Functions of Parameters**

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► LM/GLM provides not only estimated coefficients but also the variance-covariance of the estimated covariates

$$\blacktriangleright \text{ Let } \hat{\beta} = c(\hat{\beta}_1, \cdots, \hat{\beta}_p)^T$$

- Let  $\hat{\Sigma}$  denote the estimated variance-covariance of  $\hat{\beta}$
- Let a be linear coefficients

Learning Objectives

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### Inference of Linear Functions of Parameters

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# Learning Objectives

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Poisson Regressio

The Assumption of Independence

- ► Consider a linear function :  $a^T \beta$
- $\triangleright$  Estimate:  $a^T \hat{\beta}$
- ► Variance of the estimate:  $Var(a^T\hat{\beta}) = a^T\hat{\Sigma}a$
- Standard Error (SE): s.e. $(a^T\hat{\beta}) = \sqrt{a^T\hat{\Sigma}a}$
- ► A 95% confidence interval:

$$(a^{T}\hat{\beta} - 1.96 * s.e.(a^{T}\hat{\beta}), a^{T}\hat{\beta} + 1.96 * s.e.(a^{T}\hat{\beta}))$$

► Z-value:  $\frac{a^T \hat{\beta} - a^T \beta}{s.e.(a^T \hat{\beta})}$ 

# Inference of Linear Functions of Parameters: Example

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► Parameter of interest:

$$\frac{\beta_{\textit{Fruit}} \! + \! \beta_{\textit{Animal}} \! + \! \beta_{\textit{Kids}} \! + \! \beta_{\textit{Military}} \! + \! \beta_{\textit{Space}}}{5}$$

```
a=matrix(rep(1/5,5), 1)
a%*%poisson_fit$coefficients #estimate
```

```
## [,1]
## [1,] -1.598789
```

```
## [,1]
## [1,] 0.3192003
```

Learning Objectives

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Logistic Regression

Poisson Regression

Logistic Regression

Poisson Regression

- Linear contrasts are a special family of linear functions
- We say  $a^T \beta = \sum_i a_i \beta_i$  is a linear contrast if  $\sum a_i = 0$ , where  $a = (a_1, \dots, a_p)^T$ .
- Often, we are interested in whether a linear contrast is zero, i.e.,  $H_0: a^T \beta = 0$
- ightharpoonup z-value:  $\frac{a^T \hat{\beta} 0}{s.e.(a^T \hat{\beta})}$

# **Linear Contrast (e.g., Fruit vs Animal)**

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Regressio

```
a <- matrix(c(-1, 1, 0, 0, 0), 1)
#estimate
fruit_animal_est = a%*%poisson_fit$coefficients
#wartance
fruit_animal_var = a%*%summary(poisson_fit)$cov.unscaled%*%t(a)
#z value
print(fruit_animal_est/sqrt(fruit_animal_var))</pre>
```

```
## [,1]
## [1,] 4.247274
```

# Linear Contrast (e.g., Fruit vs Animal) with the multcomp package

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```
library(multcomp)
a <- matrix(c(-1, 1, 0, 0, 0), 1)
t <- glht(poisson_fit, linfct = a)
summary(t)</pre>
```

```
##
## Simultaneous Tests for General Linear Hypotheses
##
## Fit: glm(formula = n_spikes ~ image_categ ~ 1, family = poisson(link = "log"),
## data = chosen_neuron_data)
##
## Linear Hypotheses:
## Estimate Std. Error z value Pr(>|z|)
## 1 == 0    4.277    1.007    4.247   2.16e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Adjusted p values reported -- single-step method)
```

Learning Objectives

Review of LM

gistic

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### **Multinomial Logistic Regression**

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```
library(nnet)
multinom(diagnosis ~ age + female + educ + lhippo + rhippo)
## # weights: 21 (12 variable)
## initial value 2966 253179
## iter 10 value 2372 326777
## final value 2288.461323
## converged
## Call:
## multinom(formula = diagnosis ~ age + female + educ + lhippo +
       rhippo)
##
##
## Coefficients:
     (Intercept)
                                female
                                               educ
                                                       lhippo
                                                                  rhippo
                         age
        3.671844 0.026773068 -1.237127 -0.04694669 -1.251316 -0.4056775
## 2
        8 147569 0 005473649 -1 473799 -0 06379482 -1 794730 -0 7967240
## Residual Deviance: 4576.923
## ATC: 4600 923
```

Learning Objectives

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### Other concerns

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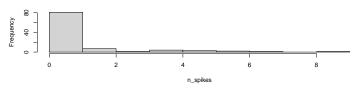
Regressio

The Assumption of Independence

- Dispersion: under- or over-dispersion
- ► Zero-inflated Poisson Regression
- ► Model selection . . .

hist(n\_spikes)

#### Histogram of n spikes



#Interested in how to fit a zero-inflated Poisson regression? See the link #https://www.rdocumentation.org/packages/pscl/versions/1.5.5/topics/zeroinfl

# The Assumption of Independence

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## **Independent Observations**

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► The common assumption we have made in LM and GLM is that the observations are independent with each other

- This is not always the case
- Examples:

Learning Objectives

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# **Independent Observations**

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Learning Objectives

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Logistic Regression

Poisson Regression

- What is the consequence of ignoring data independence?
  - The damage is probably worse than violations of distributions
  - Fortunately, tools have been developed to account for data dependence
  - Linear Mixed-Effects Model (LME)
  - Generalized Linear Mixed-Effects Model (GLMM)