

# Markov Chains

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## 1 Basics

A **discret-time stochastic process** is defined as a sequence of discrete random variables  $\{X_n\}_{n \in \mathbb{N}_0}$ , each taking values in a countable state space  $E$

A discrete time stochastic process satisfying the **Markov condition** is called a **Markov Chain**

$$\Pr(X_n = j | X_{n-1} = i, X_{n-2} = x_{n-2}, \dots, X_0 = x_0) = \Pr(X_n = j | X_{n-1} = i)$$

for all  $n \in \mathbb{N}$  and for all  $x_0, \dots, x_{n-2}, i, j \in E$ .

The Markov chain is **time-homogenous** if

$$\Pr(X_{n+1} = j | X_n = i) = \Pr(X_1 = j | X_0 = i)$$

for every  $n \in \mathbb{N}_0$  and  $i, j \in E$

### 1.1 Chapman-Kolmogorov (CK) equations

The **n-step** transition probability is

$$p_{ij}(n) = \Pr(X_{m+n} = j | X_m = i)$$

For a time homogenous Markov chain we have the **CK-equations**

$$p_{ij}(m+n) = \sum_{l \in E} p_{il}(m) p_{lj}(n)$$

where  $m \in \mathbb{N} \cup \{0\}$  and  $n \in \mathbb{N}$

The formula for **n-step transition matrix** follows:

$$P_{m+n} = P_m P_n$$

and in particular

$$P_n = P^n$$

## 1.2 First passage and hitting times

The **first passage time** is

$$T_j = \min\{n \in \mathbb{N} : X_n = j\}$$

In other words,  $\{T_j = n\} = \{X_n = j, X_{n-1} \neq j, \dots, X_1 \neq j\}$ , if  $X_n \neq j, \forall n \in \mathbb{N}$ , then  $T_j = \infty$ .

The **first passage probability** is

$$f_{ij}(n) = \Pr(T_j = n | X_0 = i), n \in \mathbb{N}_0$$

from which the hitting probability follows

$$f_{ij} = \Pr(T_j < \infty | X_0 = i) = \sum_{n=0}^{\infty} f_{ij}(n)$$

With the **special case** being  $f_{ij}(0) = 0$ .

Decomposing the **n-step transition probability**

$$p_{ij}(n) = \sum_{l=1}^n f_{ij}(l) p_{jj}(n-l)$$

## 1.3 Generating Functions of Markov Chain

Recall the **probability generating function**

$$G_X(s) = \sum_{x=0}^{\infty} s^x \Pr(X = x)$$

where this holds on the support

$$\mathcal{S}_{\mathcal{X}} = \left\{ s \in \mathbb{R} : \sum_{x=0}^{\infty} |s|^x \Pr(X = x) < \infty \right\}$$

The generating functions here are

$$G_{p_{ij}(n)} = \sum_{n=0}^{\infty} p_{ij}(n) s^n$$

$$G_{f_{ij}(n)} = \sum_{n=0}^{\infty} f_{ij}(n) s^n$$

By arguing using equating coefficients and an identity, we have a **theorem**

$$G_{p_{ij}(n)} = \delta_{ij} + G_{f_{ij}(n)}(s) G_{p_{ij}(n)}$$

The identity used is decomposition of n-step transition probability.

## 2 Recurrence and Transience

A state  $j$  is **recurrent** if and only if

$$\sum_{n=1}^{\infty} p_{jj}(n) = \infty$$

A state  $j$  is **transient** if and only if

$$\sum_{n=1}^{\infty} p_{jj}(n) < \infty$$

**Examples:** Examples of transient, irreducible chains

The **number of periods** that the chain is in state  $j$  (or **number of visits** to  $j$ ) is

$$N_j = \sum_{n=1}^{\infty} I_n(j)$$

where  $I_n(j)$  is the indicator function taking value 1 if  $X_n = j$  and 0 otherwise.

The **expected number of visits** to state  $j$  given  $X_0 = j$  is

$$\mathbb{E}[N_j | X_0 = j] = \sum_{n=0}^{\infty} p_{jj}(n)$$

proof using generating functions:

Taking  $s \rightarrow 1$  and using Abel's theorem, we can deduce...

### 2.1 Properties of recurrent/transient states

**Theorem (Number of visits is geometric for transient states)**

If  $j$  is transient, then

$$\Pr(N_j = n | X_0 = j) = f_{jj}^{n-1}(1 - f_{jj}), n \in \mathbb{N}$$

Let  $i \neq j$ , then

$$\Pr(N_j = n | X_0 = i) = \begin{cases} 1 - f_{ij} & n = 0 \\ f_{ij} f_{jj}^{n-1} (1 - f_{jj}) & n \geq 1 \end{cases}$$

Intuition is that the chain visits  $j$  for the first time and returns to it for  $n - 1$  times, then leaves it.

Therefore, it follows that for  $i \neq j$  by the mean of geometric distribution,

$$\mathbb{E}[N_j | X_0 = i] = \frac{f_{ij}}{1 - f_{jj}}$$

and

$$\mathbb{E}[N_j | X_0 = j] = \frac{1}{1 - f_{jj}}$$

### Theorem(Unlikely to visit a transient state)

If  $j$  is transient, then

$$\lim_{n \rightarrow \infty} p_{ij}(n) = 0, \forall j \in E$$

Similar results hold for null recurrent states.

## 2.2 Mean recurrence time, null and positive recurrence

The **mean recurrence time**  $\mu_j$  is

$$\mu_j = \mathbb{E}[T_j | X_0 = j] = \sum_{n=1}^{\infty} n f_{jj}(n)$$

where we recall that  $\{T_j = n\} = \{X_n = j, X_{n-1} \neq j, \dots, X_1 \neq j\}$ .

Similarly, we can define the **mean first passage time**

$$\mu_{ij} = \mathbb{E}[T_j | X_0 = i] = \sum_{n=1}^{\infty} n f_{ij}(n)$$

those expectations can be finite or infinite; for transient states, they must be infinite.

For a recurrent state  $j$ , it is called **null recurrent** if  $\mu_j = \infty$  and **positive recurrent** if  $\mu_j < \infty$ .

**Theorem (unlikely to visit null recurrent state)** A state  $j$  is null recurrent, if and only if

$$\lim_{n \rightarrow \infty} p_{jj}(n) = 0, \forall j \in E$$

In addition if the above equation holds,

$$\lim_{n \rightarrow \infty} p_{ij}(n) = 0, \forall i \neq j \in E$$

## 2.3 Examples

## 3 Aperiodicity and Ergodicity

The **period** of a state  $j$  is

$$d(j) = \gcd\{n : p_{jj}(n) > 0\}$$

It is not necessarily true that  $p_{jj}(d(j)) > 0$  (cf. Notes Pg. 36).

$$P = \begin{pmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \\ 0.5 & 0 & 0.5 & 0 \end{pmatrix}$$

A state is **ergodic** if it is positive recurrent and aperiodic.

## 4 Communicating classes

- We say that a state  $j$  is **accessible** from state  $i$  if the chain can reach  $j$  at some time, written as  $i \rightarrow j$ .
- Two states  $i$  and  $j$  are **communicating** if there exists a state  $k$  such that  $i \rightarrow k$  and  $k \rightarrow j$ , we write  $i \leftrightarrow j$ ; this is an **equivalence relation**.
- If  $i \neq j$ , then  $i \rightarrow j$  if and only if  $f_{ij} > 0$ .

### 4.1 Properties preserved by Communicating Classes

- Same period
- Same transience/recurrence
- Null recurrence

For a **set of states**  $C$ :

- $C$  is **closed** if  $\forall i \in C, j \notin C, p_{ij} = 0$
- $C$  is **irreducible** if all states in the set communicate with each other

Therefore, an irreducible set of states share the same properties described above.

**Theorem (Recurrence and closed)** If  $C$  is a communicating class of recurrent states, then  $C$  is closed.

**Theorem (Stochastic matrix on closed states)** The stochastic matrix  $P$  restricted to a closed set of closed states  $C$  is still a stochastic matrix.

### 4.2 Decomposition of Chains

The state space can be partitioned into communicating classes.

$$E = T \cup \left( \bigcup_i C_i \right)$$

where  $T$  is the set of transient states and  $C_i$ 's are irreducible closed sets of recurrent states.

### 4.3 Class Properties

The **classes** refer to communicating classes.

**Theorem (Finite Chains have recurrent)** When state space is **finite**, at least one state is *recurrent* and all *recurrent* states are **positive**

**Remark** This combined with later results on stationarity makes a chain with finite state space particularly nice.

**Remark** It follows that there are no null recurrent states in a finite state space.

**Theorem (Finite and closed)** If  $C$  is a finite, closed communicating class, then all states are positive recurrent.

Communicating class properties

Type of Class	Finite	Infinite
Closed	positive recurrent	positive/null recurrent, transient
Not closed	transient	transient

## 5 Gambler's Ruin

- Starting at state  $i \in \{0, 1, \dots, N\}$
- $p$  of winning one unit and  $1 - p$  of losing one unit
- Assume successive games are independent

**Question:** What is the probability of reaching  $N$  before reaching 0?

**Strategy:** Use **first step analysis**:

- Define  $V_i = \min\{n \in \mathbb{N}_0 : X_n = i\}$ , the first time being present at state  $i$ , desired event is  $V_N < V_0$
- Consider the conditional probability of starting at  $i$  and ending up at  $N$

$$h_i = \Pr(V_N < V_0 \mid X_0 = i)$$

- Consider the first step

$$\begin{aligned} h_i &= \Pr(V_N < V_0 \mid X_1 = i + 1, X_0 = i) \Pr(X_1 = i + 1 \mid X_0 = i) \\ &\quad + \Pr(V_N < V_0 \mid X_1 = i - 1, X_0 = i) \Pr(X_1 = i - 1 \mid X_0 = i) \\ &= h_{i+1}p + h_{i-1}(1 - p) \end{aligned}$$

Finish by solving this recurrence relation

$$h_i = \begin{cases} \frac{1 - (q/p)^i}{1 - (q/p)^N} & p \neq \frac{1}{2} \\ \frac{i}{N} & p = \frac{1}{2} \end{cases}$$

# Stationarity

We are interested in the equilibrium states of a chain

- A **distribution** is a row vector  $\lambda$  with  $\sum_j \lambda_j = 1$
- If  $\lambda P = \lambda$  then it is called *invariant*

### 5.1 Stationary distributions of irreducible chains

**Theorem** Every irreducible chain has a **stationary distribution**  $\pi$  if and only if all states are **positive recurrent**

- $\pi$  is unique
- $\pi = \mu_i^{-1}$  the inverse of mean recurrence time

We first have some lemmas:

$$l_{ji}(n) = \Pr(X_n = i, T_j \geq n \mid X_0 = j)$$

being the probability that the chain reaches  $i$  in  $n$  steps without returning to  $j$

**Lemma (Decomposing the first hitting)**

$$f_{jj}(m + n) = \sum_{i \in E, i \neq j} l_{ji}(m) f_{ij}(n)$$

from which  $f_{jj}(m + n) \geq l_{ji}(m) f_{ij}(n)$  follows

**Lemma (Formula for hitting)** We also have the following recurrence relation for  $l_{ji}(n + 1)$

$$l_{ji}(n+1) = \sum_{r \in E, r \neq j} p_{ri} l_{jr}(n)$$

with  $l_{ji}(1) = p_{ji}$ .

**Lemma (Existence):** A positive recurrent chain has a stationary distribution.

**Proof:** (constructive)

- **(Step1 Construction)** Let  $N_i(j)$  be the number of visits to state  $i$  before visiting state  $j$  (again); the sum of such numbers over  $i$  is equal to the hitting time  $T_j$
- Define  $\rho_i(j)$  to be the expected number of visits to the state  $i$  between two successive visits to state  $j$  (in this step the **recurrence** of the chain is used, as the  $T_j$  is finite with probability 1)

$$\begin{aligned} \rho_i(j) &= \mathbb{E}[N_i(j)|X_0 = j] \\ &= \sum_n \Pr(X_n = i, T_j \geq n | X_0 = j) \\ &= \sum_n l_{ji}(n) \end{aligned}$$

- Now the mean hitting time can be computed as

$$\begin{aligned} \mu_j &= \mathbb{E}\left[\sum_i N_i(j) | X_0 = j\right] \\ &= \sum_i \rho_i(j) \end{aligned}$$

- which can be written as sum of  $\rho_i(j)$  by Tonelli and linearity of conditional expectation
- **(Step2 Finiteness)** Use a lemma to bound  $\rho_i(j)$  so it's finite
- Namely write  $\rho_i(j) = \sum_n l_{ji}(n)$  and bound using the fact that the chain is irreducible, so there exists  $f_{ij}(n^*) > 0$ , so  $f_{jj}(m + n^*) \geq l_{ji}(m) f_{ij}(n^*)$
- **(Step3 Stationarity)** Use recurrence to show

$$\begin{aligned} \rho_i(j) &= \sum_n l_{ji}(n) \\ &= p_{ji} + \sum_{n=2} \sum_{r, r \neq j} p_{ri} l_{jr}(n-1) \\ &= p_{ji} \rho_j(j) + \sum_{n=1} \sum_{r \neq j} p_{ri} l_{jr}(n) \\ &= p_{ji} \rho_j(j) + \sum_{r, r \neq j} p_{ri} \sum_{n=1} l_{jr}(n) \\ &= \sum_r \rho_r(j) p_{ri} \end{aligned}$$

- which uses the fact that  $\rho_j(j) = 1$
- This  $\rho_i(j)$  does not necessarily give a probability vector when the chain is not positive recurrent.

- Now if the chain is positive recurrent, we have  $\mu_j$  finite for every  $j$ , we have

$$\pi_i = \frac{\rho_i(j)}{\mu_j}$$

Therefore, we conclude that

Every **irreducible, recurrent** chain has a positive solution to the equation  $\mathbf{x} = \mathbf{x}P$ , which is unique up to a multiplicative constant (see PS2 for proof).

Moreover, the chain is

- positive recurrent if  $\sum_i x_i < \infty$
- null recurrent if  $\sum_i x_i = \infty$

Also, from the proof, we conclude 3 **identities**:

(Sum of expected visits)

$$\mu_j = \sum_i \rho_i(j)$$

(Sum of hitting prob)

$$\rho_i(j) = \sum_n l_{ji}(n)$$

(Sum of number of visits)

$$T_j = \sum_i N_i(j)$$

**Lemma (Alternative expectation)**  $T$  is a nonnegative integer-valued random variable and  $A$  is an event with  $\Pr(A) > 0$ . Then

$$\mathbb{E}[T|A] = \sum_{n=0}^{\infty} \Pr(T \geq n|A)$$

**Lemma** If a stationary distribution exists, then the chain is positive recurrent and the distribution must be given by  $\pi_i = \mu_i^{-1}$

**proof:** ...

In a **reducible chain**, the following results are useful:

- $\pi_i = 0$  for stationary distribution  $\pi$  if  $i$  is transient or null recurrent, so we can only compute the positive recurrent states and set the rest of  $\pi_i$  to 0
- A discrete time Markov chain with **finite** state space always has at least one stationary distribution.
- This distribution is **unique** unless it has two or more closed communicating classes.
- Every stationary distribution is a **linear combination** of the stationary distributions of the closed communicating classes, with coefficients adding up to 1.



## 5.2 Limiting Distribution

A distribution  $\pi$  is a **limiting distribution** of a chain if  $\pi$  satisfies

$$\lim_{n \rightarrow \infty} p_{ij}(n) = \pi_j$$

**Theorem** For an irreducible, aperiodic chain

$$\lim_{n \rightarrow \infty} p_{ij}(n) = \frac{1}{\mu_j}$$

It follows that for an irreducible, aperiodic, and positive recurrent state, the limiting distribution is its unique stationary distribution

$$\lim_{n \rightarrow \infty} p_{ij}(n) = \pi_j = \frac{1}{\mu_j}$$

### Example of Chain with no limiting distribution

Consider the transition matrix of two alternating states

$$P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$$

the even and odd powers differ, but it has stationary distribution  $\pi = (1/2, 1/2)$ .

## 5.3 Ergodic Theorem

The **number of visits to  $i$  before time  $n$**  is defined as

$$V_i(n) = \sum_{k=0}^{n-1} \mathbf{1}_{\{X_k=i\}}$$

**Theorem** If a chain is irreducible,  $V_i(n)/n$  denotes the proportion of time the chain spent in state  $i$  before time  $n$

$$\Pr \left( \frac{V_i(n)}{n} \rightarrow \frac{1}{\mu_j} \text{ as } n \rightarrow \infty \right) = 1$$

## 5.4 Summary of properties of irreducible chains

### 1 Positive Recurrent

- Stationary distribution **exists** and is **unique**
- All mean recurrence times  $\mu_j = \mathbb{E}[T_j | x_0 = j]$  are finite and  $\pi_j = \frac{1}{\mu_j}$
- $V_i(n)/n \rightarrow \pi_i$
- If the chain is aperiodic, the limiting distribution is the stationary distribution

### 2 Null Recurrent

- All mean recurrence times are infinite
- No stationary distribution

- $V_i(n)/n \rightarrow 0$
- The limiting distribution is 0

### 3 Transient

- All mean recurrence times are infinite (any state is eventually never visited)
- No stationary distribution
- $V_i(n)/n \rightarrow 0$
- The limiting distribution is 0

## 6 Time reversibility

In this section, we assume the Markov chains are **irreducible and positive recurrent**, therefore there is a unique stationary distribution  $\pi$ .

The **reversed chain** for some  $N \in \mathbb{N}$  is defined as

$$Y_n = X_{N-n}$$

**Theorem (Reversed still Markov)** The reversed chain is a Markov chain

$$\Pr(Y_{n+1} = j | Y_n = i) = \frac{\pi_j}{\pi_i} p_{ji}$$

A Markov chain  $(X_n)$  is called **time-reversible** if its transition matrix is the same as the transition matrix of its reversed chain.

**Theorem** A Markov chain is time-reversible if and only if

$$\pi_i p_{ij} = \pi_j p_{ji}$$

this condition is called **detailed balance**.

**Theorem (Time reversible and positive recurrence)** For an irreducible chain, if **there is a vector**  $\pi$  such that the condition in the first theorem holds for all  $i, j$ , then the chain is time-reversible and positive recurrent with stationary distribution  $\pi$ .