Random Variables

- A random variable X on a sample space Ω is a function that assigns to each sample point $\omega \in \Omega$ a real number $X(\omega)$.
- Discrete r.v.'s take values in a range that is finite or countably
- Since a r.v. is defined on a probability space, we can calculate these probabilities given the probabilities of the sample points.
- Let a be any number in the range of a r.v. X. Then the set $\{\omega \in \Omega : X(\omega) = a\}$ is an event in the sample space.
- The distribution of a discrete r.v. X is the collection of values $\{(a, Pr[X = a]) : a \in \mathcal{A}\}$
- The collection of events $X = a, a \in \mathcal{A}$, satisfy two important properties:
 - any two events $X = a_1$ and $X = a_2$ with $a_1 \neq a_2$ are
 - the union of all these events is equal to the entire sample space Ω

Binomial Distribution

- $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$ where i = 0, 1, ..., n
- $$\begin{split} &Pr[X \geq n] = \sum_{i=n}^{n+k} \binom{n+k}{i} (1-p)^i p^{n+k-i} \\ \bullet & \text{The expectation (or } mean \text{ or } average) \text{ of a discrete random} \end{split}$$
 variable X is defined as $E(X) = \sum_{a \in \mathcal{A}} a \times Pr[X = a]$

Linearity of Expectation

- E(X + Y) = E(X) + E(Y)
- \bullet E(cX) = cE(X)

Variance

- "Spread" of distances from the mean.
- $Var(X) = E((X \mu)^2) = E(X^2) \mu^2$
- Standard deviation, $\sigma = \sqrt{Var(X)}$
- $Var(cX) = c^2 Var(X)$
- Var(X + Y) = Var(X) + Var(Y) (independent)
- E(XY) = E(X)E(Y) (independent)
- $\bullet Var(X+Y) = Var(X) + Var(Y) + 2(E(XY) E(X)E(Y))$

If X is uniform random variable $\{1...n\}$ with prob $\frac{1}{n}$:

- $Var(X) = \frac{n^2 1}{12}$
- $\sigma(X) = \sqrt{\frac{b^2 1}{12}}$ $Var(X_i) = E(X_i^2) E(X_i)^2$

Continuous Probability

- integral = 1
- $E[X] = \int_{-\infty}^{\infty} XP(X)dx = \sum XP(X)$
- $P[X_0 < X < X_1] = \int_{X_0}^{X_1} XP(X)dx$ $Var(x) = E[(X \mu)^2] = E[X^2] \mu^2 = \int_{-\infty}^{\infty} (x \mu)^2 f(x)dx$

Central Limit Theorem

- $\begin{array}{l} \bullet \ \ A_n = \frac{\sqrt{n}(A_n \mu)}{\sigma} = \frac{\sum_{i=1}^n X_i n\mu}{\sigma\sqrt{n}} \\ \bullet \ \ \sigma_\mu = \frac{\sigma}{\sqrt{n}}. \ \ \text{plug in to normal pdf and intergate to get intervals.} \end{array}$

Geometric Distribution

- Flip n coins, stop after first heads. P(heads) = p.
- $P[X = i] = (1 p)^i p$

- $E(X) = \frac{1}{p}$ $Var(X) = \frac{1-p}{p^2}$ $Pr[X \ge i] = (1-p)^{i-1}$ for i = 1, 2, ...

Poisson Distribution

- P[X] = 0 iff x < 0 AND $P[x] = \lambda e^{-\lambda x}$ for x > 0
- $E[X] = \lambda$
- $P[X=i] = \frac{\lambda^i e^{-\lambda}}{}$
- $Var(x) = \lambda$

Exponential Distribution

- $f(x) = \lambda e^{-\lambda x}$ for x > 0, otherwise 0.

- $Var(x) = \frac{1}{\lambda^2}$ $P[X \ge C] = \int_C^\infty \lambda e^{-\lambda x} dx$

Normal Distribution

- Parameters σ and μ , centered at $x = \mu$, standard deviation is σ
- standard normal: $\mu = 0$ and $\sigma = 1$ $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$
- \bullet $E(x) = \overset{\mathsf{v}}{\mu}$
- $P[X \ge C] = \int_C^\infty \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2} dx$ (just integrate the
- Note: 68% of the data lies within 1σ from the mean, 95% within 2σ , and 99.7% within 3σ

Probability Density Function

- f(x) = 0 for x < 0 and $x > \ell$
- $f(x) = \frac{1}{\ell}$ for $0 < x < \ell$
- $E(X) = \int_{-\infty}^{\infty} u f(x) dx = \frac{\ell}{2}$ $Var(X) = \frac{\ell^2}{12}$

Formulas

- Chebyshev (random var X with $E(X) = \mu$, $\alpha > 0$):
- $Pr[|X \mu| \ge \alpha] \le \frac{Var(X)}{\alpha^2}$ Chebyshev part 2: (random var X with $E(X) = \mu$, $\sigma = \sqrt{Var(X)}$:
 - $Pr[|X \mu| \ge \beta \sigma] \le \frac{1}{\beta 2}$
- Markov's inequality $(X > 0, E(X) = \mu, \alpha > 0)$: $Pr[X \ge \alpha] \le \frac{E(X)}{\alpha}$ • Joint Density function:

$$Pr[a \le X \le b, c \le Y \le d] = \int_c^d \int_a^b f(x, y) dx dy$$

- probability per unit area:
- $Pr[x \le X \le x + \delta, y \le Y \le y + \delta] = \int_{y}^{y+\delta} \int_{x}^{x+\delta} f(u, v) du dv \approx$
- Independent variables x, y, then f(x,y) = f(x)f(y)

Simpsons paradox: A paradox in which a trend that appears in different groups of data disappears when these groups are combined. Cardinality: In order to determine whether two sets have the same cardinality, we need to demonstrate a bijection f between the two

We say that a set S is **countable** if there is a bijection between S and \mathbb{N} or some subset of \mathbb{N} . Thus any finite set S is countable (since there is a bijection between S and the subset $\{0, 1, 2, \dots, m-1\}$, where m = |S| is the size of S).

If there is a one-to-one function $f:A\to B$, then the cardinality of A is less than or equal to that of B. Now to show that the cardinality of A and B are the same we can show that $|A| \le |B|$ and $|B| \le |A|$. Cantors Diagonalization: $\mathcal{P}(S) = \{T : T \subseteq S\}.$

Sets

- \mathbb{Z} integers (countable. f(x)=2x, f(-x)=-2x+1)
- N natural (countable)
- O rational (countable N/N. also spiral method)
- \mathbb{R} = real (uncountable always can find avg(a,b))
- Union of countable and uncountable must yield uncountable set.

The Halting Problem:

Given the description of a program and its input, we would like to know if the program ever halts when it is executed on the given input.

$$TestHalt(P, I) = \begin{cases} \text{"yes"} & \text{if program P halts on input I} \\ \text{"no"} & \text{if program P loops on input I} \end{cases}$$

Proof: Define the program

Turing(P)

$$if \ TestHalt(P,P) = "yes" \ then \ loop \ for ever$$
 else halt

So if the program P when given P as input halts, then Turing loops forever; otherwise, Turing halts.

TestHalt cannot exist, so it is impossible for a program to check if any general program halts.