

Lecture 2. Math Review (2)

Sim, Min Kyu, Ph.D.
mksim@seoultech.ac.kr

Math Review 17

$$\begin{cases} \wedge : \min \\ \vee : \max \end{cases} \quad \begin{cases} X^+ : \max(1, 0), \text{ positive part } \geq 0 \\ X^- : \min(1, 0), \text{ negative part } \leq 0 \end{cases}$$

■ For $X \sim U(20, 40)$, Evaluate $\mathbb{E}[X \wedge 25]$ and $\mathbb{E}[(25 - X)^+]$

$$\text{pdf } f(x) = \begin{cases} \frac{1}{20} & 20 \leq x \leq 40 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \mathbb{E}[X \wedge 25] &= \int_{-\infty}^{\infty} \min(x, 25) f(x) dx = \int_{20}^{40} \min(x, 25) \cdot \frac{1}{20} dx = \frac{1}{20} \left\{ \int_{20}^{25} x dx + \int_{25}^{40} 25 dx \right\} \\ &= \frac{1}{20} \left(\frac{1}{2} x^2 \Big|_{20}^{25} + 25x \Big|_{25}^{40} \right) = 24.375 \end{aligned}$$

$$\begin{aligned} \mathbb{E}[(25 - X)^+] &= \int_{-\infty}^{\infty} \max(25 - x, 0) f(x) dx = \int_{20}^{40} \max(25 - x) \cdot \frac{1}{20} dx \\ &= \frac{1}{20} \left\{ \int_{20}^{25} (25 - x) dx + \int_{25}^{40} 0 dx \right\} = \frac{1}{20} \left[25x - \frac{1}{2} x^2 \right]_{20}^{25} = \frac{1}{20} \cdot 12.5 = \frac{5}{8} \end{aligned}$$

Math Review 18 Poisson Distribution: $p(x) = P(X=x) = \frac{e^{-\lambda} \lambda^x}{x!}$ for $x=1, 2, 3, 4, \dots$

■ For $X \sim Poi(8)$, $\lambda=8$.

$$\bullet P(X=0) = \frac{e^{-8} \cdot 8^0}{0!} = e^{-8}$$

$$\bullet P(2 \leq X \leq 4) = \frac{e^{-8} \cdot 8^2}{2!} + \frac{e^{-8} \cdot 8^3}{3!} + \frac{e^{-8} \cdot 8^4}{4!} = e^{-8} \times 288$$

$$\begin{aligned} \bullet P(X > 2) &= 1 - P(X \leq 2) \\ &= 1 - \frac{e^{-8} \cdot 8^2}{2!} - \frac{e^{-8} \cdot 8^1}{1!} - \frac{e^{-8} \cdot 8^0}{0!} = 1 - e^{-8} \times 41 \end{aligned}$$

$$\exp f(x) = \lambda e^{-\lambda x}$$

$$\int uv' = uv - \int u'v$$

- For $X \sim \exp(7)$, Evaluate $\mathbb{E}[\max(X, 7)]$

$$\begin{aligned} f(x) &= 7e^{-7x}, \quad \mathbb{E}[\max(X, 7)] = \int_{-\infty}^{\infty} \max(x, 7) f(x) dx = \int_0^7 7 f(x) dx + \int_7^{\infty} x f(x) dx \\ &= 7[-e^{-7x}]_0^7 + [-xe^{-7x}]_7^{\infty} - \int_7^{\infty} e^{-7x} dx = 7[-e^{-7x}]_0^7 + [-xe^{-7x}]_7^{\infty} - \left[-\frac{1}{7}e^{-7x}\right]_7^{\infty} \\ &= -7e^{-49} + 7e^0 + 0 + 7e^{-49} + 0 + \frac{1}{7}e^{-49} = 7 + \frac{1}{7}e^{-49}. \end{aligned}$$

$$f(x) = \lambda e^{-\lambda x}, F(x) = 1 - e^{-\lambda x} \quad (x \geq 0)$$

- For $X \sim \exp(8)$, find x^* such that $F(x^*) = 0.6$

$$f(x) = 8e^{-8x}, F(x) = 1 - e^{-8x}$$

$$F(x^*) = 0.6 \Rightarrow 0.4 = e^{-8x^*} \Rightarrow \ln 0.4 = -8x^*, x^* = -\frac{1}{8} \ln 0.4$$

Math Review 21

- For $X \sim U(10, 20)$, find x^* such that $F(x^*) = 0.7$

$$f(x) = \begin{cases} 1/10 & 10 \leq x \leq 20 \\ 0 & \text{o.w.} \end{cases} \quad F(x) = \begin{cases} 0 & x < 10 \\ \frac{x-10}{10} & 10 \leq x \leq 20 \\ 1 & x > 20 \end{cases}$$

$$F(x^*) = \frac{x^* - 10}{10} = 0.7 \Rightarrow x^* = 17.$$

blank

Math Review 22 - Matrix Algebra

■ Matrix Multiplication

$1 \times 2 \cdot 2 \times 2$

$$\begin{matrix} 1 \times 2 & 2 \times 2 & 1 \times 2 \\ (.6 & .4) \begin{pmatrix} .7 & .3 \\ .5 & .5 \end{pmatrix} = \begin{pmatrix} (.6 \times .7 + .4 \times .5) & (.6 \times .3 + .4 \times .5) \end{pmatrix} \\ & & = (.62 & .38) \end{matrix}$$

■ Solve

$$A \times b = b \quad \begin{array}{l} \# \text{ eq: } m \\ \# \text{ var: } n \end{array} \quad \begin{cases} m > n \\ m = n \\ m < n \end{cases}$$

$m \times n \quad n \times 1 \quad m \times 1$

$$\begin{array}{c} \text{same!} \\ (\pi_1 \quad \pi_2) \begin{pmatrix} .7 & .3 \\ .5 & .5 \end{pmatrix} = (\pi_1 \quad \pi_2) \\ \pi_1 + \pi_2 = 1 \end{array} \left\{ \begin{array}{l} 0.7\pi_1 + 0.5\pi_2 = \pi_1 \\ 0.3\pi_1 + 0.5\pi_2 = \pi_2 \\ \pi_1 + \pi_2 = 1 \end{array} \right. \Rightarrow$$

$$0.7\pi_1 + 0.5(1 - \pi_1) = \pi_1$$

$$0.5 = 0.8\pi_1 \Rightarrow \pi_1 = \frac{5}{8}, \pi_2 = \frac{3}{8}$$

stationary distribution

$$(\pi P = \pi \quad \& \quad \sum \pi_i = 1)$$

- Solve the following system of equations

$$x = y$$

$$y = 0.5z$$

$$z = 0.6 + 0.4x$$

$$x + y + z = 1$$

no sol...?

■ Solve stationary distribution.

$$(\pi_0 \quad \pi_1 \quad \pi_2) \begin{pmatrix} -2 & 2 \\ 3 & -5 & 2 \\ & 3 & -3 \end{pmatrix} = (0 \quad 0 \quad 0)$$

$$\pi_0 + \pi_1 + \pi_2 = 1$$

$$-2\pi_0 + 3\pi_1 = 0$$

$$2\pi_0 - 5\pi_1 + 3\pi_2 = 0$$

$$2\pi_1 - 3\pi_2 = 0$$

$$\Rightarrow 2\pi_0 = 3\pi_1$$

$$2\pi_1 = 3\pi_2$$

$$\Rightarrow \pi_0 = \frac{3}{2}\pi_1$$

$$\pi_2 = \frac{2}{3}\pi_1$$

$$\Rightarrow \frac{3}{2}\pi_1 + \pi_1 + \frac{2}{3}\pi_1 = 1, \quad \pi_0 = \frac{9}{19}$$

$$\pi_1 = \frac{6}{19}$$

$$\pi_2 = \frac{4}{19}$$

$$P = \begin{pmatrix} .7 & .3 \\ .5 & .5 \end{pmatrix}$$

■ What is P^2 ?

$$P^2 = \begin{pmatrix} .7 & .3 \\ .5 & .5 \end{pmatrix} \begin{pmatrix} .7 & .3 \\ .5 & .5 \end{pmatrix} = \begin{pmatrix} .64 & .36 \\ .60 & .40 \end{pmatrix}$$

■ Solve

$$10 \times (\pi_1 \quad \pi_2 \quad \pi_3 \quad \pi_4) \begin{pmatrix} .7 & .3 \\ .5 & .5 \\ & .6 & .4 \\ & .3 & .7 \end{pmatrix} = (\pi_1 \quad \pi_2 \quad \pi_3 \quad \pi_4) \times 10$$

$$\pi_1 + \pi_2 + \pi_3 + \pi_4 = 1$$

$$\left. \begin{array}{l} 7\pi_1 + 5\pi_2 = 10\pi_1 \\ 3\pi_1 + 5\pi_2 = 10\pi_2 \\ 6\pi_3 + 3\pi_4 = 10\pi_3 \\ 4\pi_3 + 7\pi_4 = 10\pi_4 \end{array} \right\} \begin{array}{l} 3\pi_1 = 5\pi_2 \\ 4\pi_3 = 3\pi_4 \end{array} \left\{ \begin{array}{l} \pi_1 + \frac{2}{5}\pi_2 + \pi_3 + \frac{4}{5}\pi_4 = 1 \\ \frac{5}{7}\pi_1 + \frac{7}{3}\pi_3 = 1 \end{array} \right. \quad \dots \text{not unique sol.}$$

blank how to represent solution

$$\pi_1 + \frac{2}{5}\pi_1 + \pi_3 + \frac{4}{5}\pi_3 = \frac{8}{5}\pi_1 + \frac{7}{5}\pi_3 = 1$$

i) Let $p := \pi_1$, $p \in [0, 1]$

$$\pi_1 = p \quad \pi_3 = \frac{3}{7}(1 - \frac{8}{5}p)$$

$$\pi_2 = \frac{2}{5}p \quad \pi_4 = \frac{4}{7}(1 - \frac{8}{5}p)$$

$$\pi = [p, \frac{2}{5}p, \frac{3}{7}(1 - \frac{8}{5}p), \frac{4}{7}(1 - \frac{8}{5}p)]$$

for any $p \in [0, 1]$

ii) Let $p := \pi_1 + \pi_2 = \frac{8}{5}\pi_1$

$$\pi_1 = \frac{5}{8}p \quad \pi_3 = \frac{3}{7}(1 - p)$$

$$\pi_2 = \frac{2}{5}\pi_1 = \frac{2}{8}p \quad \pi_4 = \frac{4}{7}(1 - p)$$

$$\pi = [\frac{5}{8}p, \frac{2}{8}p, \frac{3}{7}(1 - p), \frac{4}{7}(1 - p)]$$

for any $p \in [0, 1]$

Math Review 23 - Express infinite vector math form

- Solve the following and express π_i for $i = 0, 1, 2, \dots$

$$\begin{aligned}\pi_0 + \pi_1 + \pi_2 + \dots &= 1 \\ \underline{0.02\pi_0 + 0.02\pi_1 + 0.02\pi_2 + \dots} &= \pi_0 \\ = 0.02(\pi_0 + \pi_1 + \dots) & \quad 0.98\pi_0 = \pi_1 \\ = 0.02 \times 1 = \pi_0 & \quad 0.98\pi_1 = \pi_2 \\ & \quad 0.98\pi_2 = \pi_3 \\ & \quad \dots = \dots\end{aligned}$$

$$\begin{aligned}\pi_0 &= 0.02 \\ \pi_1 &= 0.02 \times 0.98 \\ \pi_2 &= 0.02 \times 0.98^2 \\ &\vdots \\ \pi_i &= 0.02 \times 0.98^i \\ \text{for } i &= 0, 1, 2, \dots\end{aligned}$$

blank

Math Review 24 - Infinite Geometric Series

■ When $|r| < 1$, $S = a + ar + ar^2 + ar^3 + \dots$ 등비수열의 합.

$$S = a + ar + ar^2 + ar^3 + \dots$$

$$- \quad rS = ar + ar^2 + ar^3 + \dots$$

$$(1-r)S = a \Rightarrow S = \frac{a}{1-r}$$

$$\sum_{n=0}^{\infty} ar^n = \frac{a}{1-r}$$

Math Review 25 - Geometric Series

■ When $r \neq 1$, $S = a + ar + ar^2 + ar^3 + \dots + ar^{n-1}$

$$S = a + ar + ar^2 + \dots + ar^{n-1}$$

$$- \left| \begin{array}{l} rS = ar + ar^2 + \dots + ar^{n-1} + ar^n \end{array} \right.$$

$$(1-r)S = a - ar^n$$

$$\therefore S = \frac{a}{1-r} (1-r^n)$$

Math Review 26 - Power Series

■ When $|r| < 1$, $S = r + 2r^2 + 3r^3 + 4r^4 + \dots$

$$\begin{array}{r} S = r + 2r^2 + 3r^3 + \dots \\ - \quad rS = \quad r^2 + 2r^3 + \dots \\ \hline (1-r)S = r + r^2 + r^3 + \dots \\ = \frac{r}{1-r} \end{array}$$

$$\therefore S = \frac{r}{(1-r)^2}$$

$$\sum_{n=0}^{\infty} ar^n = \frac{a}{1-r}$$

blank

Math Review 27 - Conditional Probabilities

- Show that $\mathbb{P}(A|B \cap C)\mathbb{P}(B|C) = \mathbb{P}(A \cap B|C)$

LHS)

$$\mathbb{P}(A|B \cap C)\mathbb{P}(B|C) = \frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(B \cap C)} \times \frac{\mathbb{P}(B \cap C)}{\mathbb{P}(C)} = \frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(C)}$$

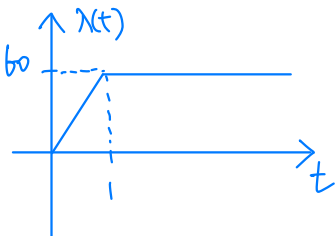
RHS)

$$\mathbb{P}(A \cap B|C) = \frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(C)}$$

blank

Math Review 28 - Formulation of time varying function

- During the first hour ($0 \leq t \leq 1$), $\lambda(t)$ increases linearly from 0 to 60. After the first hour, $\lambda(t)$ is constant at 60. Draw plot for $\lambda(t)$ and express the function in math form.



$$\lambda(t) = \begin{cases} 60t & (0 \leq t \leq 1) \\ 60 & (t > 1) \end{cases}$$

blank

Exercise 1

- For $X \sim U(a, b)$, show pdf \rightarrow cdf

$$\text{pdf } f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{o.w.} \end{cases}$$

$$\text{cdf } F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x \leq b \\ 1 & x > b \end{cases}$$

Exercise 2

- For $X \sim \text{exp}(\lambda)$, what is pdf $f(x)$?

$$f(x) = \lambda e^{-\lambda x}$$

Exercise 3

- For $X \sim \exp(\lambda)$, show pdf \rightarrow cdf.

$$\text{pdf } f(x) = \lambda e^{-\lambda x} \quad (x \geq 0)$$

$$\begin{aligned} \text{cdf } F(x) &= P(X < x) = 1 - P(X \geq x) = 1 - \int_x^{\infty} \lambda e^{-\lambda y} dy = 1 - [-e^{-\lambda y}]_x^{\infty} \\ &= 1 - (0 + e^{-\lambda x}) = 1 - e^{-\lambda x} \quad (x \geq 0) \quad \square \end{aligned}$$

Exercise 4

- For $X \sim \exp(\lambda)$, show that $\mathbb{E}X = 1/\lambda$

$$\lambda > 0, f(x) = \lambda e^{-\lambda x}$$

$$\begin{aligned}\mathbb{E}X &= \int_{-\infty}^{\infty} x f(x) dx = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \left[-x e^{-\lambda x} \right]_0^{\infty} - \int_0^{\infty} -e^{-\lambda x} dx = \int_0^{\infty} e^{-\lambda x} dx \\ &= \left[-\frac{1}{\lambda} e^{-\lambda x} \right]_0^{\infty} = \frac{1}{\lambda}\end{aligned}$$

Exercise 5

- For $X \sim \text{exp}(\lambda)$, show that $\text{Var}(X) = 1/\lambda^2$
- (hint) (You may prove this by continuing the following:)

pf) Since $\text{Var}(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2$ and we know $\mathbb{E}X = 1/\lambda$ from **hw1.#4**, we need to know what $\mathbb{E}X^2$ is.

$$\begin{aligned}\mathbb{E}X^2 &= \int_{-\infty}^{\infty} x^2 f(x) dx = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \lambda \int_0^{\infty} x^2 e^{-\lambda x} dx \\ &= \lambda \left(x^2 \cdot -\frac{1}{\lambda} e^{-\lambda x} \Big|_0^{\infty} - \int_0^{\infty} 2x \cdot -\frac{1}{\lambda} e^{-\lambda x} dx \right) \\ &= \dots\end{aligned}$$

blank

$$\begin{aligned} \mathbb{E}X^2 &= \int_{-\infty}^{\infty} x^2 f(x) dx = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \left[-x^2 e^{-\lambda x} \right]_0^{\infty} + \int_0^{\infty} +2x e^{-\lambda x} dx \\ &= 2\lambda \left\{ -\frac{1}{\lambda} x e^{-\lambda x} \Big|_0^{\infty} + \int_0^{\infty} +\frac{1}{\lambda} e^{-\lambda x} dx \right\} = 2 \left[-\frac{1}{\lambda^2} e^{-\lambda x} \right]_0^{\infty} = 2 \left(0 + \frac{1}{\lambda} \right) = \frac{2}{\lambda^2} \end{aligned}$$

$$\text{Var}X = \mathbb{E}X^2 - (\mathbb{E}X)^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

$$\text{Std}X = \frac{1}{\lambda}$$

■ Solution

$$\begin{aligned} &= \lambda \left((0 - 0) + \frac{2}{\lambda} \int_0^{\infty} x e^{-\lambda x} dx \right) = 2 \int_0^{\infty} x e^{-\lambda x} dx \\ &= 2 \left[x \cdot -\frac{1}{\lambda} e^{-\lambda x} \Big|_0^{\infty} - \int_0^{\infty} -\frac{1}{\lambda} e^{-\lambda x} dx \right] \\ &= 2 \left[\infty \cdot -\frac{1}{\lambda} e^{-\lambda \infty} - \left(0 \cdot -\frac{1}{\lambda} e^{-\lambda 0} \right) + \frac{1}{\lambda} \left[-\frac{1}{\lambda} e^{-\lambda x} \right]_0^{\infty} \right] \\ &= 2 \left[0 - 0 + \frac{1}{\lambda} \left[-\frac{1}{\lambda} e^{-\lambda \infty} + \frac{1}{\lambda} e^{-\lambda 0} \right] \right] \\ &= 2 \cdot \frac{1}{\lambda} \left(0 + \frac{1}{\lambda} \right) = \frac{2}{\lambda^2} \end{aligned}$$

Since $Var(x) = \mathbb{E}X^2 - (\mathbb{E}X)^2$, $Var(x) = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$

Exercise 6

- Show that Exponential distribution is memoryless.

r.v. X is memoryless when $P(X > s+t | X > t) = P(X > s)$
for $(s, t \geq 0)$

$$F(x) = 1 - e^{-\lambda x} = P(X \leq x)$$

$$\begin{aligned} P(X > s+t | X > t) &= \frac{P(X > s+t)}{P(X > t)} = \frac{1 - F(s+t)}{1 - F(t)} \\ &= \frac{e^{-(s+t)\lambda}}{e^{-t\lambda}} = e^{-(s+t-t)\lambda} = e^{-s\lambda} = 1 - F(s) \\ &= 1 - P(X \leq s) = P(X > s) \end{aligned}$$

Exercise 7

■ For $X \sim \text{Poi}(\lambda)$, show that $\mathbb{E}X = \lambda$

$$P(X=k) = \frac{\lambda^k e^{-\lambda}}{k!} \quad (k=0,1,2,\dots)$$

$$\begin{aligned}\mathbb{E}X &= \sum_{k=0}^{\infty} k p(k) = \sum_{k=0}^{\infty} k \cdot \frac{\lambda^k e^{-\lambda}}{k!} = \sum_{k=1}^{\infty} \lambda \frac{\lambda^{k-1} e^{-\lambda}}{(k-1)!} = \sum_{k=0}^{\infty} \lambda \cdot \frac{\lambda^k e^{-\lambda}}{k!} = \lambda \sum_{k=0}^{\infty} \frac{\lambda^k e^{-\lambda}}{k!} \\ &= \lambda \sum_{k=0}^{\infty} p(k) = \lambda\end{aligned}$$

Exercise 8

■ For $X \sim \text{Poi}(\lambda)$, show that $\text{Var}(X) = \lambda$

■ (hint) (you may prove by continuing following.)

pf) Since $\text{Var}(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2$ and we know $\mathbb{E}X = \lambda$ from **hw1.#7**, we need to know what $\mathbb{E}X^2$ is.

$$\begin{aligned}\mathbb{E}X^2 &= \sum_{x=-\infty}^{\infty} x^2 p(x) = \sum_{x=0}^{\infty} x^2 \frac{\lambda^x e^{-\lambda}}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} x^2 \frac{\lambda^x}{x!} \\&= e^{-\lambda} \left(0^2 \frac{\lambda^0}{0!} + \sum_{x=1}^{\infty} x^2 \frac{\lambda^x}{x!} \right) \\&= e^{-\lambda} \left(\sum_{x=1}^{\infty} x \frac{\lambda^x}{(x-1)!} \right) \\&= e^{-\lambda} \left(\sum_{x=1}^{\infty} (x-1+1) \frac{\lambda^x}{(x-1)!} \right) \\&= e^{-\lambda} \left(\sum_{x=1}^{\infty} (x-1) \frac{\lambda^x}{(x-1)!} + \sum_{x=1}^{\infty} \frac{\lambda^x}{(x-1)!} \right) \\&= e^{-\lambda} \left(\sum_{x=2}^{\infty} \frac{\lambda^x}{(x-2)!} + \sum_{x=1}^{\infty} \frac{\lambda^x}{(x-1)!} \right) \\&= \dots\end{aligned}$$

blank

$$p(x) = \frac{\lambda^x e^{-\lambda}}{x!} \quad x=0,1,2,\dots \quad x-1=y$$

$$\begin{aligned} \mathbb{E}X^2 &= \sum_{x=0}^{\infty} x^2 p(x) = \sum_{x=0}^{\infty} x^2 \frac{\lambda^x e^{-\lambda}}{x!} = \sum_{x=1}^{\infty} \lambda x \cdot \frac{\lambda^{x-1} e^{-\lambda}}{(x-1)!} = \sum_{y=0}^{\infty} \lambda(y+1) \frac{\lambda^y e^{-\lambda}}{y!} \\ &= \lambda \left\{ \sum_{y=0}^{\infty} y \frac{\lambda^y e^{-\lambda}}{y!} + 1 \right\} = \lambda \{ \mathbb{E}X + 1 \} = \lambda(\lambda+1) = \lambda^2 + \lambda \end{aligned}$$

$$\text{Var} X = \mathbb{E}X^2 - (\mathbb{E}X)^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

■ Solution

$$\begin{aligned} &= e^{-\lambda} \left[\lambda^2 \sum_{x=2}^{\infty} \frac{\lambda^{x-2}}{(x-2)!} + \lambda \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} \right] \\ &= e^{-\lambda} \left[\lambda^2 \sum_{y=0}^{\infty} \frac{\lambda^y}{y!} + \lambda \sum_{z=0}^{\infty} \frac{\lambda^z}{z!} \right], \text{ where } y := x-2, z := x-1 \\ &= e^{-\lambda} [\lambda^2 \cdot e^{\lambda} + \lambda \cdot e^{\lambda}] \\ &= \lambda^2 + \lambda \end{aligned}$$

$$\text{Var}(x) = \mathbb{E}X^2 - (\mathbb{E}X)^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

(Note that the property of exponential function is used, i.e. $e^x = \sum_{i=0}^{\infty} \frac{x^i}{i!}$)

Exercise 9

■ Let X be a discrete random variable with $\mathbb{P}(X = i) = ci$ for positive, odd integers $i < 8$; otherwise, the probability is zero.

- (a) Compute the value of c . (by using the one of the properties of pmf)
- (b) What is the mean of X ?
- (c) What is the second moment of X ? (i.e. What is $\mathbb{E}X^2$?)
- (d) What is the variance of X ?
- (e) Compute $\mathbb{E}[(X - 3)^+]$

$$a) i = 1, 3, 5, 7 \Rightarrow c(1+3+5+7) = 1 \quad \therefore c = \frac{1}{16}$$

$$b) \mathbb{E}X = \sum_{-\infty}^{\infty} xp(x) = \frac{1}{16}(1^2+3^2+5^2+7^2) = 21/4$$

$$c) \mathbb{E}X^2 = \sum_{-\infty}^{\infty} x^2 p(x) = \frac{1}{16}(1^3+3^3+5^3+7^3) = 496/16 = 31$$

$$d) \text{Var}X = \mathbb{E}X^2 - (\mathbb{E}X)^2 = 31 - \left(\frac{21}{4}\right)^2 = 3.4375$$

blank

$$e) (x-3)^+ = \begin{cases} 0 & x \leq 3 \\ x-3 & x > 3 \end{cases}$$

$$\begin{aligned} E((X-3)^+) &= \sum_{-\infty}^{\infty} (x-3)^+ p(x) = (5-3) \frac{5}{16} + (7-3) \frac{1}{16} \\ &= \frac{1}{16} (10+2) = \frac{12}{16} \end{aligned}$$

■ Solution

(a) Since $\sum_{i=-\infty}^{\infty} p(i) = \sum_{i=-\infty}^{\infty} \mathbb{P}(X = i) = 1,$

$$\mathbb{P}(X = 1) + \mathbb{P}(X = 3) + \mathbb{P}(X = 5) + \mathbb{P}(X = 7) = c + 3c + 5c + 7c = 1$$

$$\therefore c = 1/16$$

(b)

$$\begin{aligned}\mathbb{E}X &= \sum_{i=-\infty}^{\infty} i \cdot p(i) = \sum_{i=-\infty}^{\infty} i \cdot \mathbb{P}(X = i) = 1 \cdot p(1) + 3 \cdot p(3) + 5 \cdot p(5) + 7 \cdot p(7) \\ &= \frac{1}{16}(1^2 + 3^2 + 5^2 + 7^2) = \frac{84}{16} = \frac{21}{4}\end{aligned}$$

(c)

$$\begin{aligned}\mathbb{E}X^2 &= \sum_{i=-\infty}^{\infty} i^2 \cdot p(i) = \sum_{i=-\infty}^{\infty} i^2 \cdot \mathbb{P}(X = i) = 1^2 \cdot p(1) + 3^2 \cdot p(3) + 5^2 \cdot p(5) + 7^2 \cdot p(7) \\ &= \frac{1}{16}(1^3 + 3^3 + 5^3 + 7^3) = \frac{496}{16} = \frac{124}{4} = 31\end{aligned}$$

$$(d) \operatorname{Var}(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2 = 31 - \left(\frac{21}{4}\right)^2 = 3.4375$$

(e)

$$\begin{aligned}\mathbb{E}[(X-3)^+] &= \mathbb{E}[\max(X-3, 0)] = \sum_{i=-\infty}^{\infty} \max(i-3, 0) \cdot \mathbb{P}(X=i) \\&= \max(1-3, 0) \cdot \mathbb{P}(X=1) + \max(3-3, 0) \cdot \mathbb{P}(X=3) \\&\quad + \max(5-3, 0) \cdot \mathbb{P}(X=5) + \max(7-3, 0) \cdot \mathbb{P}(X=7) \\&= 2 \cdot \frac{5}{16} + 4 \cdot \frac{7}{16} = \frac{38}{16} = \frac{19}{8}\end{aligned}$$

Exercise 10

$$X \sim \text{Poi}(5)$$

■ Let X be a Poisson random variable with parameter 5, and let $Y = \min(X, 2)$

- (a) What is the pmf of X ?
- (b) What is the mean of X ?
- (c) What is the variance of X ?
- (d) What is the pmf of Y ? (i.e. Specify $\mathbb{P}(Y = i)$ for $i = 0, 1, 2, \dots$)
- (e) Compute $E[Y]$?

$$a) p(x) = \frac{5^x e^{-5}}{x!}$$

$$b) \lambda = 5$$

$$c) \lambda = 5$$

$$d) \mathbb{P}(Y=y)$$

$$= \mathbb{P}(\min(X, 2) = y)$$

$$\left\{ \begin{array}{l|l} p(0) = e^{-5} & y=0 \\ p(1) = 5e^{-5} & y=1 \\ p(2) = 1 - 6e^{-5} & y=2 \\ p(i) = 0, i \geq 3 & \end{array} \right.$$

$$\begin{aligned} e) \sum_{-\infty}^{\infty} y p(y) &= 0 \cdot e^{-5} + 1 \cdot 5e^{-5} \\ &\quad + 2(1 - 6e^{-5}) \\ &= 2 - 7e^{-5} \end{aligned}$$

■ Solution

(a) $p(x) = \mathbb{P}(X = x) = \frac{e^{-5}5^x}{x!}$

(b) 5 ($= \lambda$)

(c) 5 ($= \lambda$)

(d) We have $p(y) = \mathbb{P}(Y = y) = \mathbb{P}(\min(X, 2) = y)$. It follows:

■ $p(0) = \mathbb{P}(Y = 0) = \mathbb{P}(\min(X, 2) = 0) = \mathbb{P}(X = 0) = \frac{e^{-5}5^0}{0!} = e^{-5}$

■ $p(1) = \mathbb{P}(Y = 1) = \mathbb{P}(\min(X, 2) = 1) = \mathbb{P}(X = 1) = \frac{e^{-5}5^1}{1!} = 5e^{-5}$

■ $p(2) = \mathbb{P}(Y = 2) = \mathbb{P}(\min(X, 2) = 2) = \mathbb{P}(X \geq 2) = 1 - \mathbb{P}(X = 0) - \mathbb{P}(X = 1) = 1 - e^{-5} - 5e^{-5} = 1 - 6e^{-5}$

■ $p(3) = \mathbb{P}(Y = 3) = 0$ ($\because \min(X, 2) \leq 2$ always), also $p(i) = 0$ for all $i \geq 3$

■ Therefore,

$$p(y) = \begin{pmatrix} e^{-5} \text{ for } y = 0 \\ 5e^{-5} \text{ for } y = 1 \\ 1 - 6e^{-5} \text{ for } y \geq 2 \end{pmatrix}$$

(e)

$$\mathbb{E}Y = \sum_{i=-\infty}^{\infty} yp(y) = 0 \cdot p(0) + 1 \cdot p(1) + 2 \cdot p(2) = 5e^{-5} + 2(1 - 6e^{-5}) = 2 - 7e^{-5}$$

Exercise 11

■ Let Y be a random variable with pdf ce^{-3y} for $y > 0$, and zero otherwise.

- (a) Determine c .
- (b) What is the mean, variance, and squared coefficient of variation of Y ?
- (c) Compute $\mathbb{P}(Y > 4)$
- (d) Compute $\mathbb{P}(Y > 7 | Y > 3)$
- (e) Your answer for (c) and answer for (d) should be same. Discuss shortly why.
- (f) What is the point x^* such that $\mathbb{P}(Y > x^*) = 2/3$?

$$a) \int_{-\infty}^{\infty} ce^{-3y} dy = \int_0^{\infty} ce^{-3y} dy = \left[-\frac{1}{3}ce^{-3y} \right]_0^{\infty} = \frac{1}{3}c = 1 \quad \therefore c = 3$$

$$b) f(x) = \begin{cases} 3e^{-3y} & y > 0 \\ 0 & \text{o.w.} \end{cases} \quad Y \sim \exp(3) \quad \mathbb{E}Y = \frac{1}{3} \quad \text{Var}X = \left(\frac{1}{3}\right)^2$$
$$cv^2 = \frac{(\frac{1}{3})^2}{(\frac{1}{3})^2} = 1$$

$$c) \mathbb{P}(Y > 4) = 1 - \mathbb{P}(Y < 4) = 1 - (1 - e^{-3 \cdot 4}) = e^{-12}$$

$$d) \mathbb{P}(Y > 4+3 | Y > 3) = \mathbb{P}(Y > 4) = e^{-12}$$

e) $\exp(\lambda)$ is memoryless

blank

$$f) P(Y > x^*) = 1 - P(Y < x^*) = 1 - F(x^*) = e^{-\lambda x^*} = 2/3$$

$$\Rightarrow -\lambda x^* = \ln(2/3) \quad \therefore x^* = -\frac{1}{\lambda} \ln(2/3)$$

■ Solution

(a)

$$\int_{-\infty}^{\infty} p(y) dy = 1$$

$$\rightarrow \int_0^{\infty} ce^{-3y} dy = c \cdot -\frac{1}{3} e^{-3y} \Big|_0^{\infty} = -\frac{c}{3} [0 - e^{-3 \cdot 0}] = \frac{c}{3} = 1$$

$$\therefore c = 3$$

(b) pdf of $\begin{cases} 3e^{-3y} & y \geq 0 \\ 0 & \text{otherwise} \end{cases}$ implies $Y \sim \exp(3)$. Thus, $\mathbb{E}Y = \frac{1}{3}$, $VarY = \frac{1}{3^2}$,

$$\text{and } \text{sqr-cv} = \frac{Var(Y)}{(\mathbb{E}Y)^2} = \frac{1/3^2}{(1/3)^2} = 1$$

(c) cdf of $\exp(3)$ is $\begin{cases} 1 - e^{-3y} & \text{for } y \geq 0 \\ 0 & \text{otherwise} \end{cases}$ Thus,

$$\mathbb{P}(Y > 4) = 1 - \mathbb{P}(Y \leq 4) = 1 - (1 - e^{-3 \cdot 4}) = e^{-12}$$

$$(d) \mathbb{P}(Y > 7 | Y > 3) = \frac{\mathbb{P}(Y > 7, Y > 3)}{\mathbb{P}(Y > 3)} = \frac{1 - (1 - e^{-3 \cdot 7})}{1 - (1 - e^{-3 \cdot 3})} = \frac{e^{-21}}{e^{-9}} = e^{-12}$$

(e) memoryless property

(f)

$$\mathbb{P}(Y > x^*) = 1 - \mathbb{P}(Y \leq x^*) = 1 - (1 - e^{-3x^*}) = \frac{2}{3}$$

$$\rightarrow e^{-3x^*} = \frac{2}{3}$$

$$\rightarrow -3x^* = \ln(2/3)$$

$$\rightarrow x^* = -\frac{1}{3} \ln(2/3)$$

"Man can learn nothing unless he proceeds from the known to the unknown."
- Claude Bernard"