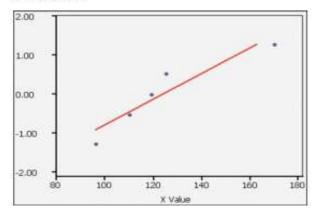
[2] Triola, 11th edition, example 2

Movie Lengths Data Set 9 in Appendix B includes lengths (in minutes) of randomly selected movies. Let's consider only the first 5 movie lengths: 110, 96, 170, 125, 119. With only 5 values, a histogram will not be very helpful in revealing the distribution of the data. Instead, construct a normal quantile plot for these 5 values and determine whether they appear to come from a population that is normally distributed.

The following steps correspond to those listed in the above procedure for constructing a normal quantile plot.

Step 1. First, sort the data by arranging them in order. We get 96, 110, 119, 125, 170.

STATDISK



Step 2. With a sample of size n = 5, each value represents a proportion of 1/5 of the sample, so we proceed to identify the cumulative areas to the left of the corresponding sample values. The cumulative left areas, which are expressed in general as 1/2n, 3/2n, 5/2n, 7/2n, and so on, become these specific areas for this example with n = 5: 1/10, 3/10, 5/10, 7/10, and 9/10. The cumulative left areas expressed in decimal form are 0.1, 0.3, 0.5, 0.7, and 0.9.

Step 3. We now search in the body of Table A-2 for the cumulative left areas of 0.1000, 0.3000, 0.5000, 0.7000, and 0.9000 to find these corresponding z scores: -1.28, -0.52, 0, 0.52, and 1.28.

Step 4. We now pair the original sorted movie lengths with their corresponding z scores. We get these (x, y) coordinates which are plotted in the accompanying STATDISK display: (96, -1.28), (110, -0.52), (119, 0), (125, 0.52), and (170, 1.28).

We examine the normal quantile plot in the STATDISK display. Because the points appear to lie reasonably close to a straight line and there does not appear to be a systematic pattern that is not a straight-line pattern, we conclude that the sample of five movie lengths appears to come from a normally distributed population.

Example 4, Chp8 Triola p.416

Finding the Value of the Test Statistic Let's again consider the claim that the XSORT method of gender selection increases the likelihood of having a baby girl. Preliminary results from a test of the XSORT method of gender selection involved 14 couples who gave birth to 13 girls and 1 boy. Use the given claim and the preliminary results to calculate the value of the test statistic. Use the format of the test statistic given above, so that a normal distribution is used to approximate a binomial distribution. (There are other exact methods that do not use the normal approximation.)

From Figure 8-2 and the example displayed next to it, the claim that the XSORT method of gender selection increases the likelihood of having a baby girl results in the following null and alternative hypotheses: H_0 : p = 0.5 and H_1 : p > 0.5. We work under the assumption that the null hypothesis is true with p = 0.5. The sample proportion of 13 girls in 14 births results in $\hat{p} = 13/14 = 0.929$. Using p = 0.5, $\hat{p} = 0.929$, and p = 14, we find the value of the test statistic as follows:

$$z = \frac{\hat{p} - p}{\sqrt{\frac{pq}{n}}} = \frac{0.929 - 0.5}{\sqrt{\frac{(0.5)(0.5)}{14}}} = 3.21$$

"unusual" (because it is greater than 2). It appears that in addition to being greater than 0.5, the sample proportion of 13/14 or 0.929 is *significantly* greater than 0.5. Figure 8-3 shows that the sample proportion of 0.929 does fall within the range of values considered to be significant because they are so far above 0.5 that they are not likely to occur by chance (assuming that the population proportion is p = 0.5).

Figure 8-3 shows the test statistic of z = 3.21, and other components in Figure 8-3 are described as follows.

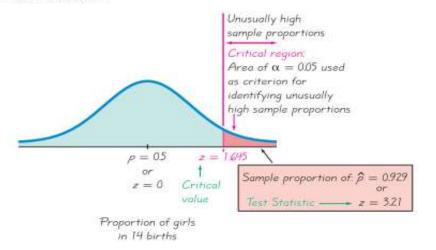


Table A.3 (continued) Areas under the No

z	.00	.01	.02	.03
0.0	0.5000	0.5040	0.5080	0.5120
0.1	0.5398	0.5438	0.5478	0.5517
0.2	0.5793	0.5832	0.5871	0.5910
0.3	0.6179	0.6217	0.6255	0.6293
0.4	0.6554	0.6591	0.6628	0.6664
0.5	0.6915	0.6950	0.6985	0.7019
0.6	0.7257	0.7291	0.7324	0.7357
0.7	0.7580	0.7611	0.7642	0.7673
0.8	0.7881	0.7910	0.7939	0.7967
0.9	0.8159	0.8186	0.8212	0.8238
1.0	0.8413	0.8438	0.8461	0.8485
1.1	0.8643	0.8665	0.8686	0.8708
1.2	0.8849	0.8869	0.8888	0.8907
1.3	0.9032	0.9049	0.9066	0.9082
1.4	0.9192	0.9207	0.9222	0.9236
1.5	0.9332	0.9345	0.9357	0.9370
1.6	0.9452	0.9463	0.9474	0.9484
1.7	0.9554	0.9564	0.9573	0.9582
1.8	0.9641	0.9649	0.9656	0.9664
1.9	0.9713	0.9719	0.9726	0.9732

between 0.9495 and 0. this gives z = 1.645

EXAMPLE 8

Finding a P-Value for a Critical Region in Two Tails Consider the claim that with the XSORT method of gender selection, the likelihood of having a baby girl is different from p = 0.5, and use the test statistic z = 3.21 found from 13 girls in 14 births. First determine whether the given conditions result in a critical region in the right tail, left tail, or two tails, then use Figure 8-5 to find the P-value. Interpret the P-value.

SOLUTION The claim that the likelihood of having a baby girl is different from p = 0.5 can be expressed as $p \neq 0.5$, so the critical region is in two tails (as in Figure 8-4(a)). Using Figure 8-5 to find the P-value for a two-tailed test, we see that the P-value is twice the area to the right of the test statistic z = 3.21. We refer to Table A-2 (or use technology) to find that the area to the right of z = 3.21 is 0.0007. In this case, the P-value is twice the area to the right of the test statistic, so we have:

$$P$$
-value = $2 \times 0.0007 = 0.0014$

Table A.3 Areas under the Normal Curve

z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9989	0.9990	0.9990
3.1	0.9990	0.9991	0.9991	0.9991	0.9992	0.9992	0.9992	0.9992	0.9993	0.9993
3.2	0.9993	0.9993	0.9994	0.9994	0.9994	0.9994	0.9994	0.9995	0.9995	0.9995
3.3	0.9995	0.9995	0.9995	0.9996	0.9996	0.9996	0.9996	0.9996	0.9996	0.9997
3.4	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9998

INTERPRETATION

The P-value is 0.0014 (or 0.0013 if greater precision is used for the calculations). The small P-value of 0.0014 shows that there is a very small chance of getting the sample results that led to a test statistic of z = 3.21. This suggests that with the XSORT method of gender selection, the likelihood of having a baby girl is different from 0.5.

[2] p.419

A1	В	С	D	E	F	G	Н		J	<u> </u>	<u> </u>	M	N	0	P	Q	R
								Row-		Row-sum-	$ \overline{V} $	A softwar	e house	wished to	study the	effect of fo	our
2	1	2	3	4	5	() 6	Row-sum	ave	Row-stdv	squared	•				——————————————————————————————————————	roficiency	
3	18	17.9	18.7	12.9	18.1	17.9	103.50	17.25	2.15	10712.25	5 =H3*H3	individual			· .	•	
4	19	18.2	18.9	12.5	16.9	18.2	103.70	17.28	3 2.46	10753.69)	make the		•	•	•	
5	20	21.7	18.6	12.7	17.2	21.7	111.90	18.65	3.40	12521.61			•		_	a) verbal c	
6	18.8	18.4	19.1	11.5	17.7	18.4	103.90	17.32	2.89	10795.21		• •				c) video	,
7	Data Extrac					1011	grand-ave	17.63						•		ther verba	•
	tolerance	0.05	(4)	=SUM(B3:G6)	X =	423.00	=AVERAGE(B3:G6		Row-stdv) =K7/C10	nor video	•		•		
9	I(row)=	4.00	\smile	=G8*G8	sqr(X) =	178929.00			=STDEV(B3:G	<mark>3</mark>)			•	•		ort), was r	
10	- ()	6.00	,	corr factor =	sqr(X)/IJ =	7455.38						•			•	each. Ea	
		mu2 = mu3		all means are eq	ual (2)		_		=K8-G10		SSTr=SST-SSE		_	•	-		
12	H1: At leas	t one of the i	means is different f	rom the others.					=K11/(C9-1)	2.81	MSTr=SSTr/(I-1)	of subject					
13	324.00	320.41	349.69			320.41										oject was t	
14	361.00	331.24	357.21			331.24						while crea	ating a ce	rtain tax r	eturn repo	rt (use 5%	Ď
15	400.00	470.89	345.96			470.89					=H18-K11	tolerance). The da	ata is in th	e following	g table (in	minutes).
16	353.44	338.56	364.81			338.56				SSE	153.07	Does the	data indi	cate that o	different te	aching me	ethods
17				=SUM(B13:G16)		Sum (Xij)*(Xij)	7616.86				-					naking the	
18				=H17-G10		SST=H17-G10	161.49		f calc	=K12/H19		Also, plea		_			roport :
19				=L16/(C9*(C10-1))	MSE=SSE/I(J-1)	7.65		f=MSTr/MSE	0.37		Also, piec		ille givell	emply lab	IC.	
20																	
21				1	_		1										
22	1	2	3	4	5	6			•								
23	18.3	21.1	18.7	19.5	18.1	17.9	F(table)	3.1			Should one continu	ue to Tukey t	test? Why	? Do it Or V	√hy not ?		

F(alpha, I-1, I(J-1))

18.2 F(0.05,3,20)

(14)

Hence f calc >=

21.7

18.4

	Source		d.f.		Sum of		Mean square	f(calc)
28					squares		-	, ,
29	Methods	I-1=	3	SSTr=	8.42	MSTr=	2.81	0.37
30	Error	I(J-1)=	20	SSE=	153.07	MSE=	7.65	
31	Total		23	SST=	161.49			

20

19.7

18.8

16.9

17.2

17.7

Proposition: Let \overline{X}_i and S_i^2 (i = 1, ..., I) denote the sample mean and variance of the *i*th sample. Define the between-samples estimator $\hat{\sigma}_B^2$ by

$$\hat{\sigma}_B^2 = JS_{\overline{X}}^2 = \frac{J_{i=1}^I (\overline{X}_{i\cdot} - \overline{X}..)^2}{I - 1} = \frac{\sum_{i=1}^I \sum_{j=1}^J (\overline{X}_{i\cdot} - \overline{X}..)^2}{I - 1}$$
(10.1)

18.9

18.6

19.1

and the within-sample estimator $\hat{\sigma}_{W}^{2}$ by

17.6

20.3

18.9

24

25

26

27

34 35

38

17.5

17.5

$$\hat{\sigma}_{W}^{2} = \frac{\sum_{i=1}^{I} S_{i}^{2}}{I} = \frac{1}{I} \left[\sum_{i=1}^{I} \frac{1}{J-1} \sum_{j=1}^{J} (X_{ij} - \overline{X}_{i}.)^{2} \right] = \frac{\sum_{i=1}^{I} \sum_{j=1}^{J} (X_{ij} - \overline{X}_{i}.)^{2}}{I(J-1)}$$
(10.2)

Then $\hat{\sigma}_B^2$ is an unbiased estimator of σ^2 when H_0 is true, but $E(\hat{\sigma}_B^2) > \sigma^2$ when H_0 is false, while $\hat{\sigma}_W^2$ is unbiased for σ^2 whether or not H_0 is true.

(10)Ha is rejected in other words, ways of teaching had 3.1 is untrue

Since Ha is rejected no method is significantly different

no effect on making the reports

$$F = \frac{\hat{\sigma}_B^2}{\hat{\sigma}_W^2} = \frac{J \sum (\overline{X}_i . - \overline{X} ..)^2 / (I - 1)}{\sum \sum (X_{ij} - \overline{X}_i .)^2 / I (J - 1)}$$
(10.3)

Definition: The total sum of squares (SST), treatment sum of squares (SSTr), and error sum of squares (SSE) are given by

$$SST = \sum_{i=1}^{I} \sum_{j=1}^{J} (X_{ij} - \overline{X}_{..})^{2} = \sum_{i=1}^{I} \sum_{j=1}^{J} X_{ij}^{2} - \frac{1}{IJ} X^{2}_{..}$$

$$SSTr = \sum_{i=1}^{I} \sum_{j=1}^{J} (\overline{X}_{i}. - \overline{X}_{..})^{2} = \frac{1}{J} \sum_{i=1}^{I} X_{i}^{2}. - \frac{1}{IJ} X^{2}_{..}$$

$$SSE = \sum_{i=1}^{I} \sum_{j=1}^{J} (X_{ij} - \overline{X}_{i}.)^{2}, \text{ where } X_{i}. = \sum_{j=1}^{J} X_{ij}, X_{..} = \sum_{i=1}^{I} \sum_{j=1}^{J} X_{ij}$$

The fundamental identity of single-factor ANOVA:

$$SST = SSTr + SSE$$

$$MSTr = \frac{SSTr}{I - 1}, \quad MSE = \frac{SSE}{I(J - 1)}, \quad F = \frac{MSTr}{MSE}$$
(10.6)

39 40 41 42 43 Empty Table

	Empty rabi	•			
44	Source	d.f.	Sum of squares	Mean square	F
	Methods				
45					
46	Error				
47	Total				

Table A.7 Critical Values F_{α,ν_1,ν_2} for the F Distribution $\alpha=.05$

 $\alpha = .05$

120	1	2	3	4	5	6	7	8	9	10	12	15	20	24	30	40	60	120	30
1	1614	199.5	2157	2246	230.2	2340	2368	238 9	2405	2419	2439	2459	2480	249 1	250.1	2511	252.2	2533	254 3
2	18.51	19,00	19.16	19 25	19 30	19 33	19 35	19 37	19 38	19 40		19 43		19 45			19 48	19 49	19 50
3	10.13	9.55	9 28	9 12	901	894	8 89	885		8 79	8 74	8 70	8 66	8 64	8 62	8 59			
4	7.71	694	6 59	6 39	6.26	6 16	6 09	6 0 4	6 00		591	5 86	5 80	5 77	5 75	5 72	5 69	8 55 5 66	
5	6.61	5 79	541	5 19	5 0 5	495	4 88	482	4 77	474	4 68	4 62	4 56	4 5 3	4 50	4 46	4 4 3	4 40	4 36
6	5 99	5 14	4 76	4 53	4 39	4 28	421	4 15	4 10	4 06	4 00	394	387	384	381	3 7 7	3 74	3 70	3 6
7	5 59	474	4 35	4 12	397	387	3 79	3 73	3 68	364	357	351	3 4 4	3 4 1	3 38	3 34	3 30	327	3 23
8	5 32	4 46	407	384	3 69	3 58	3.50	3 44	3 39	3 35	3 28	3 22	3 15	3 12	3 08	3 04	301		
9	5 12	4 26	386	3 63	3 48	3 37	3 29	3 23	3 18	3 14	307	301	294	290	2 86	283	279	275	293
10	4 96	4 10	371	3 48	3 33	3 22	3 14	307	302	298	291	285	277	274	2 70	2 66	2.62	2.58	254

В Ε G Example from Devore 2 The Table next (10.4 p425) [1] is a SAS ANOVA table. The last column gives the P-value as Anova example with Tukey 0.0001. Using a significance level of .05, we reject the null hypothesis A biologist wished to study the effects of ethanol on sleep time. A sample of 20 rats, matched for age and other characteristics, was selected, and each rat was given an oral injection having a particular concentration of ethanol per body weight. The rapid H₀: $\mu_1 = \mu_2 = \mu_3 = \mu_4$, since P-value = 0.0001 < .05 = α . True average REM sleep time does appear eye movement (REM) sleep time for each rat was then recorded for a 24-hour period, with the following results: (Example 10to depend on concentration level. 6 [1]). Does the data indicate that the true average REM sleep time depends on the concentration of ethanol? (This example is based on an experiment reported in "Relationship of Ethanol Blood Level to REM and Non-REM Sleep Time and Distribution in the Rat, "Life Sciences", 1978: 839–846.). Use 95% confidence interval. If Ho is rejected then Use Tukey to find the method which has the most impact on REM. 11 12 They \bar{x}_i s differ rather substantially from one another, but there is also a great deal of variability within each sample, so to answer the question precisely we must carry out the ANOVA. 15 Row-sum-Treatment 16 \overline{x}_{i} Row-sum Row-ave Row-stdv squared Type x_{i} 17 (concentration of ethanol) **18** 0 (control) 88.6 157132.96 =J18*J18 73.2 91.4 68 75.2 396.4 79.28 396.40 79.28 10.18 19 1 g/kg 63 53.9 69.2 50.1 71.5 307.7 61.54 9.34 307.70 61.54 94679.29 20 2 g/kg44.9 59.5 40.2 56.3 38.7 239.6 47.92 239.60 57408.16 47.92 9.46 21 4 g/kg 31 39.6 45.3 32.76 25.2 22.7 163.8 26830.44 163.80 32.76 9.56 22 336050.85 =SUM(M18:M21) 55.375 grand-ave 55.38 1107.5 23 z = =AVERAGE(C18:G21) 67210.17 =M22/C27 x..= Row-stdv 24 **Data Extracted** =STDEV(C18:G21) $\Sigma\Sigma x^2ii$ =SUM(C18:G21) X.. = 1107.50 25 tolerance 0.05 19.69472 26 I(row)= 4.00 =G25*G25 =K8-G10 5882.36 **SSTr=SST-SSE** sqr(X..) = 1226556.25 27 =G26/(C26*C27) = 61327.81 1960.79 MSTr=SSTr/(I-1) J(column)= corr factor = sqr(X..)/IJ ==L26/(C26-1) 28 29 7849.96 5358.24 8353.96 4624 5655.04 =H18-K11 30 3969 2905.21 4788.64 SSE 2510.01 5112.25 1487.40 31 2016.01 3540.25 1616.04 3169.69 1497.69 32 961 1568.16 2052.09 635.04 515.29 33 34 =SUM(C29:G32) Sum (Xij)*(Xij) 68697.57 =L27/G36 35 =H17-G10 SST=H17-G10 7369.76 f=MSTr/MSE 36 =L16/(C9*(C10-1)) MSE=SSE/I(J-1) 92.96 37 F(table) 3.24 38 F(0.05,3,16) F(alpha, I-1, I(J-1)) 39 Hence f calc >= 3.24 is True 40 3.24 is True Should one continue to Tukey test? YES W some effect on the REM 13.63 $w = Q_{(\alpha + 1/(1-1))} \sqrt{MSE/J} \rightarrow \text{and get } Q_{(\alpha + 1/(1-1))} \text{ from table}$ 41 Since Ha is NOT rejected thus a/some methods are significantly different 42 Ha is NOT rejected in other words, ways of ethanol treatment HAVE Tukey Method & Formula -----> 43 some effect on the REM. There are I = 4 treatments and 16 df for error, from which $w = Q_{(\alpha, 1, |(I-1)|)} \sqrt{MSE/J} \rightarrow giving$

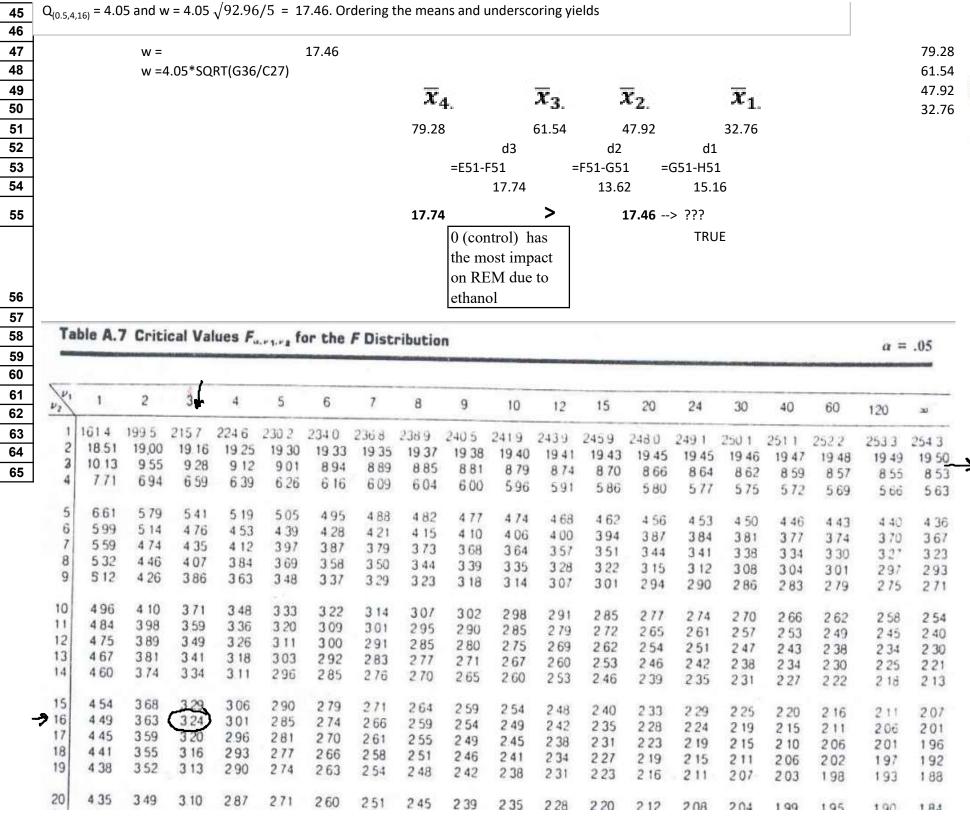


Table A.8 Critical Values Qo,m, for the Studentized Range Distribution

				ماء		171					
ν	α	2	3	4	5	6	7	8	9	10	11
5	.05	3.64 5.70	4.60 6.98	5.22 7.80	5.67 8.42	6.03 8.91	6.33 9.32	6.58 9.67	6.80 9.97	6.99 10.24	7.17 10.48
6	.05	3.46	4.34	4.90	5.30	5.63	5.90	6.12	6.32	6.49	6.65
	.01	5.24	6.33	7.03	7.56	7.97	8.32	8.61	8.87	9.10	9.30
7	.05	3.34	4.16	4.68	5.06	5.36	5.61	5.82	6.00	6.16	6.30
	.01	4.95	5.92	6.54	7.01	7.37	7.68	7.94	8.17	8.37	8.55
8	.05	3.26	4.04	4.53	4.89	5.17	5.40	5.60	5.77	5.92	6.05
	.01	4.75	5.64	6.20	6.62	6.96	7.24	7.47	7.68	7.86	8.03
9	.05	3.20	3.95	4.41	4.76	5.02	5.24	5.43	5.59	5.74	5.87
	.01	4.60	5.43	5.96	6.35	6.66	6.91	7.13	7.33	7.49	7.65
10	.05	3.15	3.88	4.33	4.65	4.91	5.12	5.30	5.46	5.60	5.72
	.01	4.48	5.27	5.77	6.14	6.43	6.67	6.87	7.05	7.21	7.36
11	.05	3.11	3.82	4.26	4.57	4.82	5.03	5.20	5.35	5.49	5.61
	.01	4.39	5.15	5.62	5.97	6.25	6.48	6.67	6.84	6.99	7.13
12	.05	3.08	3.77	4.20	4.51	4.75	4.95	5.12	5.27	5.39	5.51
	.01	4.32	5.05	5.50	5.84	6.10	6.32	6.51	6.67	6.81	6.94
13	.05	3.06	3.73	4.15	4.45	4.69	4.88	5.05	5.19	5.32	5.43
	.01	4.26	4.96	5.40	5.73	5.98	6.19	6.37	6.53	6.67	6.79
14	.05	3.03 4.21	3.70 4.89	4.11 5.32	4.41 5.63	4.64 5.88	4.83 6.08	4.99 6.26	5.13 6.41	5.25 6.54	5.36 6.66
15	.05	3.01	3.67	4.08	4.37	4.59	4.78	4.94	5.08	5.20	5.31
	.01	4.17	4.84	5.25	5.56	5.80	5.99	6.16	6.31	6.44	6.55
16	.05	3.00 4.13	3.65 4.79	4.05 5.19	4.33 5.49	4.56 5.72	4.74 5.92	4.90 6.08	5.03 6.22	5.15 6.35	5.26 6.46

۸1	В				F		ا ں			1			ı	NA.	1	NI .		<u>, </u>		
A1 2	Example from De	vore	D	<u> </u>	Г	G	Н	1	J		<u>K</u>	<u> </u>		M		N	1 ()	<u> </u>	Q
	Anova example		<i>I</i>																	
	1	•	, the effects of ethar	nol on sleep time	. A sample of 20 ra	ts, matched fo	or age and other	characteristics,												
	1 ~	-	was given an oral in	•	•		•													
	7		p time for each rat v																	
	6 [1]). Does the	e data indic	ate that the true ave	erage REM sleep i	time depends on t	ne concentrat	on of ethanol (F	la)? (This												
8	example is base	ed on an ex	periment reported i	n "Relationship of	f Ethanol Blood Le	vel to REM and	d Non-REM Slee	o Time and												
9			fe Sciences", 1978: 8		5% confidence of ir	nterval. If Ho is	rejected then l	Jse Tukey to												
10	find the metho	d which has	the most impact or	ı REM.																
12	·																			
13			substantially from																	
14	within each s	ample, so	to answer the quest	tion precisely w	e must carry out f	he ANOVA.														
15			1119.1	430	200															
16	Туре			Treatment			x_{i}	\overline{x}_{i}							•					
17		00 -		ncentration of etha													7			
	0 (control)	88.6			68	75.2	396.4	79.28												
	1 g/kg	63	53.9		50.1	71.5	307.7	61.54												
	2 g/kg	44.9	59.5 39.6	+	56.3 25.2	38.7 22.7	239.6 163.8	47.92 32.76												
22	4 g/kg	31	39.0	43.3	25.2	22.1	1107.5	55.375												
23	-							x =												
	Data Extracted							77												
25	tolerance									Table	4 8 C	itical Va	lune ()	for t	ho Stu	dontino	d Done	no Diet	albudla	
	I(row)=						1			-		Tolour V		m, v 101 C	m m	ucituize	u nang	ge Disc	ribucio	
27	J(column)=									ν	α	2	3 4	5	6	7		9	10	11
28																	- 8			7-24-
30	-									5	.05	3.64 4	60 5.2	2 567	6.03	633	6 59	6.80	6.00	7.17
31										5	.05 .01	5.70 6	60 5.2 98 7.8	8.42					6.99 10.24	7.17 10.48
	1									5		5.70 6 3.46 4	60 5.2 98 7.8 34 4.9 33 7.0	0 8.42 0 5.30	6.03 8.91 5.63 7.97	6.33 9.32 5.90 8.32	6.58 9.67 6.12 8.61	6.80 9.97 6.32 8.87	10.24 6.49	
32										5 6 7	.01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4	98 7.8 34 4.9 33 7.0 16 4.6	0 8.42 0 5.30 7.56 8 5.06	8.91 5.63 7.97 5.36	9.32 5.90 8.32 5.61	9.67 6.12 8.61	9.97 6.32 8.87 6.00	10.24 6.49 9.10 6.16	10.48 6.65 9.30 6.30
33										5 6 7 8	.01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4	98 7.8 34 4.9 33 7.0 16 4.6 92 6.5 04 4.5	0 8.42 0 5.30 7.56 8 5.06 4 7.01 3 4.89	8.91 5.63 7.97 5.36 7.37 5.17	9.32 5.90 8.32 5.61 7.68 5.40	9.67 6.12 8.61 5.82 7.94 5.60	9.97 6.32 8.87 6.00 8.17 5.77	10.24 6.49 9.10 6.16 8.37 5.92	10.48 6.65 9.30 6.30 8.55 6.05
										7	.01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5	98 7.8 34 4.9 33 7.0 16 4.6 92 6.5 04 4.5 64 6.2	0 8.42 0 5.30 7.56 8 5.06 4 7.01 3 4.89 0 6.62	8.91 5.63 7.97 5.36 7.37 5.17 6.96	9.32 5.90 8.32 5.61 7.68 5.40 7.24	9.67 6.12 8.61 5.82 7.94 5.60 7.47	9.97 6.32 8.87 6.00 8.17 5.77 7.68	10.24 6.49 9.10 6.16 8.37 5.92 7.86	10.48 6.65 9.30 6.30 8.55 6.05 8.03
33 34 35										7 8 9	.01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3	98 7.8 34 4.9 33 7.0 16 4.6 92 6.5 04 4.5	0 8.42 0 5.30 7.56 8 5.06 7.01 8 4.89 0 6.62 1 4.76	8.91 5.63 7.97 5.36 7.37 5.17	9.32 5.90 8.32 5.61 7.68 5.40	9.67 6.12 8.61 5.82 7.94 5.60	9.97 6.32 8.87 6.00 8.17 5.77	10.24 6.49 9.10 6.16 8.37 5.92	10.48 6.65 9.30 6.30 8.55 6.05
33 34 35 36										7	.01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.15 3	98 7.8 34 4.9 33 7.0 16 4.6 92 6.5 04 4.5 64 6.2 95 4.4 43 5.9 88 4.3	0 8.42 0 5.30 7.56 8 5.06 4 7.01 3 4.89 0 6.62 1 4.76 6.35 4.65	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65
33 34 35 36 37										7 8 9	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.15 3 4.48 5 3.11 3	7.8 7.8 7.8 7.8 7.9 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0	0 8.42 0 5.30 7.56 8 5.06 4 7.01 3 4.89 0 6.62 1 4.76 6.35 3 4.65 7 6.14 4.57	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36
33 34 35 36 37 38										7 8 9	.01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.15 3 4.48 5 3.11 3 4.39 5 3.08 3	7.8 7.8 7.8 7.8 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9	0 8.42 0 5.30 7.56 8 5.06 4 7.01 8 4.89 0 6.62 1 4.76 6.35 6.14 4.57 5.97 0 4.51	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05 5.35 6.84	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61
33 34 35 36 37 38 39										7 8 9 10 11	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.11 3 4.39 5 3.08 3 4.32 5	98 7.8 34 4.9 33 7.0 16 4.6 92 6.5 04 4.5 64 6.2 95 4.4 5.9 88 4.3 5.7 82 4.2 15 5.6 77 4.2 5.5	0 8.42 0 5.30 7.56 8 5.06 7.01 3 4.89 6.62 1 4.76 6.35 4.65 6.14 4.57 5.97 4.51 5.84	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25 4.75 6.10	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48 4.95 6.32	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67 5.12 6.51	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05 5.35 6.84 5.27 6.67	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99 5.39 6.81	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61 7.13 5.51 6.94
33 34 35 36 37 38			$w = Q_{[}$	$_{\alpha, I, I(J-1)]} \sqrt{MSE/J} \rightarrow$	→ and get Q _[α, Ι, Ι(J-1)] f	rom table				7 8 9 10	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.15 3 4.48 5 3.11 3 4.39 5 3.08 3 4.32 5 3.06 3	7.8 7.8 7.8 7.8 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9	0 8.42 0 5.30 7.56 8 5.06 4 7.01 8 4.89 6.62 1 4.76 6.35 3 4.65 6.14 4.57 5.97 4.51 5.84 4.45	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25 4.75	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48 4.95	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67 5.12	9.97 6.32 8.87 6.00 8.17 7.68 5.59 7.33 5.46 7.05 5.35 6.84 5.27	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99 5.39	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61 7.13
33 34 35 36 37 38 39 40 41	Tukey Method &	ւ Formula		$_{\alpha, \text{ I, I(J-1)]}}\sqrt{\text{MSE}/J} \rightarrow$, and get $Q_{[\alpha, I, I(J-1)]}$ f	rom table				7 8 9 10 11	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.11 3 4.39 5 3.08 3 4.32 5 3.06 3 4.26 4 3.03 3	7.8 7.8 7.8 7.8 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9	0 8.42 0 5.30 7.56 8 5.06 7.01 3 4.89 6.62 1 4.76 6.35 4.65 6.35 4.65 6.14 6.57 5.97 0 4.51 5.84 4.45 5.73 1 4.41	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25 4.75 6.10 4.69 5.98 4.64	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48 4.95 6.32 4.88 6.19 4.83	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67 5.12 6.51 5.05 6.37 4.99	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05 5.35 6.84 5.27 6.67 5.19 6.53 5.13	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99 5.39 6.81 5.32 6.67 5.25	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61 7.13 5.51 6.94 5.43 6.79 5.36
33 34 35 36 37 38 39 40 41 42 43	Tukey Method &	ı Formula	>			rom table				7 8 9 10 11 12	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.11 3 4.39 5 3.08 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.31 3 4.32 5 3.03 3 4.31 3 4.31 3 4.32 5 3.03 3 4.31 3 4.31 3 4.32 5 3.03 3 4.31 3	7.8 7.8 7.8 7.8 7.8 7.8 7.8 7.8 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9	0 8.42 0 5.30 7.56 8 5.06 7.01 3 4.89 6.62 1 4.76 6.35 4.65 6.14 4.57 5.97 0 4.51 5.84 4.45 5.73 1 4.41 5.63 8 4.37	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25 4.75 6.10 4.69 5.98 4.64 5.88 4.59	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48 4.95 6.32 4.88 6.19 4.83 6.08 4.78	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67 5.12 6.51 5.05 6.37 4.99 6.26 4.94	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05 5.35 6.84 5.27 6.67 5.19 6.53 5.13 6.41 5.08	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99 5.39 6.81 5.32 6.67 5.25 6.54 5.20	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61 7.13 5.51 6.94 5.43 6.79 5.36 6.66 5.31
33 34 35 36 37 38 39 40 41 42	Tukey Method &	. Formula				rom table				7 8 9 10 11 12 13	.01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01 .05 .01	5.70 6 3.46 4 5.24 6 3.34 4 4.95 5 3.26 4 4.75 5 3.20 3 4.60 5 3.11 3 4.39 5 3.08 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.32 5 3.06 3 4.31 3 4.32 5 3.03 3 4.31 3 4.31 3 4.32 5 3.03 3 4.31 3 4.31 3 4.32 5 3.03 3 4.31 3	7.8 7.8 7.8 7.8 7.8 7.8 7.8 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9 7.9	0 8.42 0 5.30 7.56 8 5.06 7.01 3 4.89 0 6.62 1 4.76 6.35 4.65 6.14 4.57 5.97 4.51 5.84 4.45 5.73 4.41 5.63 4.37 5.56	8.91 5.63 7.97 5.36 7.37 5.17 6.96 5.02 6.66 4.91 6.43 4.82 6.25 4.75 6.10 4.69 5.98 4.64 5.88	9.32 5.90 8.32 5.61 7.68 5.40 7.24 5.24 6.91 5.12 6.67 5.03 6.48 4.95 6.32 4.88 6.19 4.83 6.08	9.67 6.12 8.61 5.82 7.94 5.60 7.47 5.43 7.13 5.30 6.87 5.20 6.67 5.12 6.51 5.05 6.37 4.99 6.26	9.97 6.32 8.87 6.00 8.17 5.77 7.68 5.59 7.33 5.46 7.05 5.35 6.84 5.27 6.67 5.19 6.53 5.13 6.41	10.24 6.49 9.10 6.16 8.37 5.92 7.86 5.74 7.49 5.60 7.21 5.49 6.99 5.39 6.81 5.32 6.67 5.25 6.54	10.48 6.65 9.30 6.30 8.55 6.05 8.03 5.87 7.65 5.72 7.36 5.61 7.13 5.51 6.94 5.43 6.79 5.36 6.66

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			\overline{x}	4.	\overline{x}	3.	$\overline{x}_{2.}$		$\overline{x}_{1.}$										
Ta	ble A.	7 Criti	cal Va	lues <i>F</i> .,	. + 1. + 2 fe	or the	F Dist	ributio	n	******								α =	.05
N2 1	1	2	3	4	5	6	7	8	9	10	12	15	20	24	30	40	60	120	30
1 2 3	161 4 18.51 10 13 7.71	199 5 19,00 9 55 6 94	215 7 19 16 9 28 6 59		230 2 19 30 9 01 6 26	234 0 19 33 8 94	236 8 19 35 8 89	238 9 19 37 8 85	240 5 19 38 8 81	2419 1940 879	2439 1941 874	8 70	8 66	249 1 19 45 8 64	250 1 19 46 8 62	8 59	252 2 19 48 8 57	253 3 19 49 8 55	
5	6.61 5.99	5 79 5 14	5 4 1 4 76	5 19 4 53	5 0 5 4 3 9	6 16 4 95 4 28	6 09 4 88 4 21	6 0 4 4 8 2 4 1 5	6 00 4 77 4 10	5 9 6 4 7 4 4 0 6	5 9 1 4 68 4 00	5 86 4 62 3 94	5 80 4 56 3 87	5 77 4 53 3 84	5 75 4 50 3 81	5 72 4 46 3 77	5 69 4 43 3 74	5 66 4 40 3 70	5 63 4 36 3 67
8 9	5 59 5 32 5 12	4 74 4 46 4 26	4 35 4 07 3 86	4 12 3 84 3 63	3 97 3 69 3 48	3 87 3 58 3 37	3 79 3 50 3 29	3 73 3 44 3 23	3 68 3 39 3 18	3 64 3 35 3 14	357 328 307	351 322 301	3 44 3 15 2 94	3 41 3 12 2 90	3 38 3 08 2 86	3 34 3 04 2 83	3 30 3 01 2 79	327 297 275	3 23 2 93 2 71
10 11 12 13	4 96 4 84 4 75 4 67 4 60	4 10 3 98 3 89 3 81 3 74	3 71 3 59 3 49 3 41 3 34	3 48 3 36 3 26 3 18 3 11	3 33 3 20 3 11 3 03 2 96	3 22 3 09 3 00 2 92 2 85	3 14 3 01 2 91 2 83 2 76	307 295 285 277 270	3 02 2 90 2 80 2 71 2 65	298 285 275 267 260	291 279 269 260 253	2 85 2 72 2 62 2 53 2 46	277 265 254 246 239	274 261 251 242 235	270 257 247 238 231	2 66 2 53 2 43 2 34 2 27	2 62 2 49 2 38 2 30	2 58 2 45 2 34 2 25	2 54 2 40 2 30 2 21
15 16 17 18	4 54 4 49 4 45 4 41 4 38	3 68 3 63 3 59 3 55 3 52	3 29 3 24 3 20 3 16 3 13	306 301 296 293 290	290 285 281 277 274	2 79 2 74 2 70 2 66 2 63	271 266 261 258 254	264 259 255 251 248	2 59 2 54 2 49 2 46 2 42	2 54 2 49 2 45 2 41 2 38	2 48 2 42 2 38 2 34 2 31	2 40 2 35 2 31 2 27	2 33 2 28 2 23 2 19	2 29 2 24 2 19 2 15	2 25 2 19 2 15 2 11	2 20 2 15 2 10 2 06	2 22 2 16 2 11 2 06 2 02	2 18 2 11 2 06 2 01 1 97	2 13 2 07 2 01 1 96 1 92
20	4 35	3 49	3.10	287	2.71	2 60	251	2 45	2 39	2 35	2 28	2 23	2 16	211	207	199	198	193	188