

# MATH 142: Mathematical Modeling, Quiz 3

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## Question 1

### Part A

Similar to what we did in class and on the homework:

$$P_n(t + \Delta t) = \sigma_{n+1}P_{n+1}(t) + \nu_n\alpha_nP_n(t) + \beta_{n-1}P_{n-1}(t),$$

where  $\sigma_{n+1}$  is the probability that a death occurred from a population of  $n+1$  individuals,  $\nu_n$  is the probability that no deaths occurred from a population of  $n$  individuals,  $\alpha_n$  is the probability that no births occurred from a population of  $n$  individuals, and  $\beta_{n-1}$  is the probability that a birth occurred from a population of  $n-1$  individuals

Since  $\Delta t$  is small:

$$\begin{aligned}\nu_n &= (1 - \mu\Delta t)^n \approx 1 - n\mu\Delta t \\ \sigma_{n+1} &= 1 - (1 - \mu\Delta t)^{n+1} \approx 1 - (1 - (n+1)\mu\Delta t) = (n+1)\mu\Delta t \\ \alpha_n &= (1 - \lambda\Delta t)^n \approx 1 - n\lambda\Delta t \\ \beta_{n-1} &= 1 - (1 - \lambda\Delta t)^{n-1} \approx 1 - (1 - (n-1)\lambda\Delta t) = (n-1)\lambda\Delta t\end{aligned}$$

Plugging in:

$$P_n(t + \Delta t) \approx (n+1)\mu\Delta tP_{n+1}(t) + (1 - n\mu\Delta t)(1 - n\lambda\Delta t)P_n(t) + (n-1)\lambda\Delta tP_{n-1}(t)$$

Expanding:

$$P_n(t + \Delta t) \approx (n+1)\mu\Delta tP_{n+1}(t) + (1 - n\lambda\Delta t - n\mu\Delta t + n^2\mu\lambda\Delta t^2)P_n(t) + (n-1)\lambda\Delta tP_{n-1}(t)$$

Dividing both sides by  $\Delta t$ :

$$\frac{P_n(t + \Delta t) - P_n(t)}{\Delta t} \approx (n+1)\mu P_{n+1}(t) + (-n\lambda - n\mu + n^2\mu\lambda\Delta t)P_n(t) + (n-1)\lambda P_{n-1}(t)$$

Rearranging:

$$\frac{P_n(t + \Delta t) - P_n(t)}{\Delta t} \approx (n+1)\mu P_{n+1}(t) + (-n\lambda - n\mu + n^2\mu\lambda\Delta t)P_n(t) + (n-1)\lambda P_{n-1}(t)$$

Taking the limit and dropping the  $\approx$ :

$$\frac{dP_n}{dt} = (n+1)\mu P_{n+1} + (-n\lambda - n\mu)P_n(t) + (n-1)\lambda P_{n-1}$$

Further simplifying:

$$\frac{dP_n}{dt} = (n-1)\lambda P_{n-1} + (n+1)\mu P_{n+1} - (\lambda + \mu)nP_n, \text{ for } n = 2, 3, 4, \dots \quad (1)$$

From (1), we derive  $\frac{dP_0}{dt}$ :

$$\frac{dP_0}{dt} = \mu P_1, \quad (2)$$

noting that  $P_{-1}$  does not make sense in this context and  $(\lambda + \mu) * 0 * P_0 = 0$ .

From (1), we derive  $\frac{dP_1}{dt}$ :

$$\frac{dP_1}{dt} = 2\mu P_2 - (\lambda + \mu)P_1, \quad (3)$$

noting that  $(1 - 1)\lambda P_{1-1} = 0$ .

## Part B

(i)

Let  $\lambda < \mu$ , and taking the limit:

$$\lim_{t \rightarrow \infty} P_0(t) = \lim_{t \rightarrow \infty} \left( \frac{\mu e^{(\mu-\lambda)t} - \mu}{\mu e^{(\mu-\lambda)t} - \lambda} \right)^{N_0}$$

Factor out  $e^{(\mu-\lambda)t}$ :

$$\lim_{t \rightarrow \infty} \left( \frac{e^{(\mu-\lambda)t} \left( \mu - \frac{\mu}{e^{(\mu-\lambda)t}} \right)}{e^{(\mu-\lambda)t} \left( \mu - \frac{\lambda}{e^{(\mu-\lambda)t}} \right)} \right)^{N_0}$$

Canceling out  $e^{(\mu-\lambda)t}$ :

$$\lim_{t \rightarrow \infty} \left( \frac{\mu - \frac{\mu}{e^{(\mu-\lambda)t}}}{\mu - \frac{\lambda}{e^{(\mu-\lambda)t}}} \right)^{N_0}$$

Using property of limits:

$$\frac{\lim_{t \rightarrow \infty} \left( \mu - \frac{\mu}{e^{(\mu-\lambda)t}} \right)^{N_0}}{\lim_{t \rightarrow \infty} \left( \mu - \frac{\lambda}{e^{(\mu-\lambda)t}} \right)^{N_0}}$$

Taking the limit of the numerator and denominator:

$$1$$

Thus, when  $\lambda < \mu$ , the probability that the population will eventually go extinct is 1.

(ii)

Let  $\lambda > \mu$ , which means  $\mu - \lambda < 0$ . We take the limit of  $P_0$

$$\left( \lim_{t \rightarrow \infty} \left( \frac{\mu e^{(\mu-\lambda)t} - \mu}{\mu e^{(\mu-\lambda)t} - \lambda} \right) \right)^{N_0}$$

We factor out  $e^{(\mu-\lambda)t}$  and it cancels out to get:

$$\left( \lim_{t \rightarrow \infty} \left( \frac{\mu - \frac{\mu}{e^{(\mu-\lambda)t}}}{\mu - \frac{\lambda}{e^{(\mu-\lambda)t}}} \right) \right)^{N_0}$$

Using exponential property:

$$\left( \lim_{t \rightarrow \infty} \left( \frac{\mu - \mu e^{(\lambda-\mu)t}}{\mu - \lambda e^{(\lambda-\mu)t}} \right) \right)^{N_0}$$

Using L'Hopital's Rule:

$$\left(\lim_{t \rightarrow \infty} \left( \frac{-\mu(\lambda - \mu)e^{(\lambda - \mu)t}}{-\lambda(\lambda - \mu)e^{(\lambda - \mu)t}} \right) \right)^{N_0}$$

Taking the limit:

$$\left(\frac{\mu}{\lambda}\right)^{N_0}$$

Thus, when  $\lambda > \mu$ , the probability that the population will eventually go extinct is  $\left(\frac{\mu}{\lambda}\right)^{N_0}$ .

## Question 2

### Part A

(i)

$d_{k+1}$  is the distance traveled by the bacterium during its  $(k+1)^{st}$  run.

$$d_{k+1} = \begin{cases} 10 \mu m & , \text{ with probability } \frac{1}{4} \\ -10 \mu m & , \text{ with probability } \frac{1}{4} \\ 0 \mu m & , \text{ with probability } \frac{1}{2} \end{cases}$$

(ii)

Let L = moving  $10 \mu m$  to the left, R = moving  $10 \mu m$  to the right, and T = tumbling. We can find the probability that a bacterium is at the location of the food ( $x = 10 \mu m$ ) at  $t = 3$  seconds through brute force:

2a ii.

	Position ( $\mu m$ )						
	-30	-20	-10	0	10	20	30
1			L	T	R		
2		LL	LT TL	LR TT RL	RT TR	RR	
3	LLL	LLT LTL TLL	LLR LTT TLT LRL TTL RLL	LTR TLR LRT TTT RLT RTL TRL	LRR TRR RLR RTT TET PRL	RTR TRR RET	RRR

  

Prob of LRR =  $\frac{1}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{64}$

TRR =  $\frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{4} = \frac{1}{16}$

RLR =  $\frac{1}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{64}$

RTT =  $\frac{1}{4} \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{16}$

TRT =  $\frac{1}{2} \cdot \frac{1}{4} \cdot \frac{1}{2} = \frac{1}{16}$

RRL =  $\frac{1}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{64}$

15

64

### Part B

We are starting from  $x = 0$ . After  $k_1$  time steps, the average position is  $k_1(\frac{2}{3}l - \frac{1}{3}l) = \frac{k_1}{3}l$ . After  $k_1 + k_2$  time steps, the average position is  $\frac{k_1}{3}l + k_2(\frac{1}{2}l - \frac{1}{2}l) = \frac{k_1}{3}l$ . Finally, after  $k_1 + k_2 + k_3$  time steps, the average position is  $\frac{k_1}{3}l + k_3(\frac{1}{5}l - \frac{4}{5}l) = \frac{k_1}{3}l - \frac{3k_3}{5}l = (\frac{k_1}{3} - \frac{3k_3}{5})l$ .