



Statistical Intervals for a Single Sample

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LEARNING OBJECTIVES

After careful study of this chapter, you should be able to do the following:

- Construct confidence intervals on the mean of a normal distribution, using either the normal distribution or the t distribution method
- 2. Construct confidence intervals on the variance and standard deviation of a normal distribution
- 3. Construct confidence intervals on a population proportion
- 4. Use a general method for constructing an approximate confidence interval on a parameter



8-1 Introduction

- In the previous chapter we illustrated how a parameter can be estimated from sample data. However, it is important to understand how good is the estimate obtained.
- Bounds that represent an interval of plausible values for a parameter are an example of an interval estimate.
- Three types of intervals will be presented:
 - Confidence intervals
 - Prediction intervals
 - Tolerance intervals



8-2.1 Development of the Confidence Interval and its Basic Properties

Suppose that $X_1, X_2, ..., X_n$ is a random sample from a normal distribution with unknown mean $\underline{\mu}$ and known variance σ^2 . From the results of Chapter 5 we know that the sample mean \overline{X} is normally distributed with mean $\underline{\mu}$ and variance σ^2/n . We may **standardize** \overline{X} by subtracting the mean and dividing by the standard deviation, which results in the variable

$$Z = \frac{\overline{X} - \mu}{\sigma / \sqrt{n}} \tag{8-3}$$

Now Z has a standard normal distribution.



8-2.1 Development of the Confidence Interval and its Basic Properties

A confidence interval estimate for μ is an interval of the form $l \le \mu \le u$, where the endpoints l and u are computed from the sample data. Because different samples will produce different values of l and u, these end-points are values of random variables L and U, respectively. Suppose that we can determine values of L and U such that the following probability statement is true:

$$P\{L \le \mu \le U\} = 1 - \alpha \tag{8-4}$$

where $0 \le \alpha \le 1$. There is a probability of $1 - \alpha$ of selecting a sample for which the CI will contain the true value of μ . Once we have selected the sample, so that $X_1 = x_1, X_2 = x_2, \ldots, X_n = x_n$, and computed l and u, the resulting **confidence interval** for μ is

$$l \le \mu \le u \tag{8-5}$$



8-2.1 Development of the Confidence Interval and its Basic Properties

The endpoints or bounds l and u are called lower- and upper-confidence limits, respectively.

• Since Z follows a standard normal distribution, we can write:

$$P\left\{-z_{\alpha/2} \le \frac{\overline{X} - \mu}{\sigma/\sqrt{n}} \le z_{\alpha/2}\right\} = 1 - \alpha$$

Now manipulate the quantities inside the brackets by (1) multiplying through by σ/\sqrt{n} , (2) subtracting \overline{X} from each term, and (3) multiplying through by -1. This results in

$$P\left\{\overline{X} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \le \mu \le \overline{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}\right\} = 1 - \alpha \tag{8-6}$$



8-2.1 Development of the Confidence Interval and its Basic Properties

Definition

If \bar{x} is the sample mean of a random sample of size n from a normal population with known variance σ^2 , a $100(1 - \alpha)\%$ CI on μ is given by

$$\overline{x} - z_{\alpha/2}\sigma/\sqrt{n} \le \mu \le \overline{x} + z_{\alpha/2}\sigma/\sqrt{n}$$
 (8-7)

where $z_{\alpha/2}$ is the upper $100\alpha/2$ percentage point of the standard normal distribution.



Example 8-1

ASTM Standard E23 defines standard test methods for notched bar impact testing of metallic materials. The Charpy V-notch (CVN) technique measures impact energy and is often used to determine whether or not a material experiences a ductile-to-brittle transition with decreasing temperature. Ten measurements of impact energy (J) on specimens of A238 steel cut at 60°C are as follows: 64.1, 64.7, 64.5, 64.6, 64.5, 64.3, 64.6, 64.8, 64.2, and 64.3. Assume that impact energy is normally distributed with $\sigma = 1J$. We want to find a 95% CI for μ , the mean impact energy. The required quantities are $z_{\alpha/2} = z_{0.025} = 1.96$, n = 10, $\sigma = 1$, and $\overline{x} = 64.46$. The resulting 95% CI is found from Equation 8-7 as follows:

$$\overline{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \le \mu \le \overline{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

$$64.46 - 1.96 \frac{1}{\sqrt{10}} \le \mu \le 64.46 + 1.96 \frac{1}{\sqrt{10}}$$

$$63.84 \le \mu \le 65.08$$

That is, based on the sample data, a range of highly plausible vaules for mean impact energy for A238 steel at 60° C is $63.84J \le \mu \le 65.08J$.



Confidence Level and Precision of Error

The length of a confidence interval is a measure of the precision of estimation.

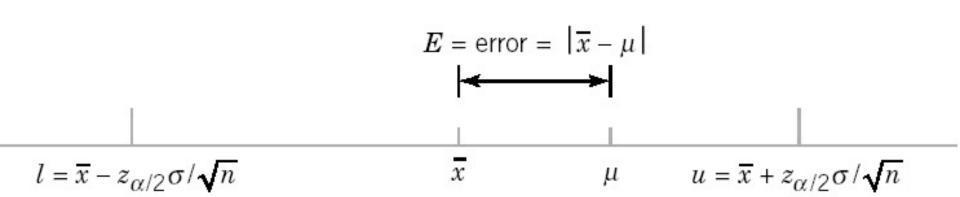


Figure 8-2 Error in estimating μ with $\overline{\chi}$.



8-2.2 Choice of Sample Size

Definition

If \bar{x} is used as an estimate of μ , we can be $100(1 - \alpha)\%$ confident that the error $|\bar{x} - \mu|$ will not exceed a specified amount E when the sample size is

$$n = \left(\frac{z_{\alpha/2}\sigma}{E}\right)^2 \tag{8-8}$$



Example 8-2

To illustrate the use of this procedure, consider the CVN test described in Example 8-1, and suppose that we wanted to determine how many specimens must be tested to ensure that the 95% CI on μ for A238 steel cut at 60°C has a length of at most 1.0*J*. Since the bound on error in estimation *E* is one-half of the length of the CI, to determine *n* we use Equation 8-8 with E = 0.5, $\sigma = 1$, and $z_{\alpha/2} = 0.025$. The required sample size is 16

$$n = \left(\frac{z_{\alpha/2}\sigma}{E}\right)^2 = \left[\frac{(1.96)1}{0.5}\right]^2 = 15.37$$

and because n must be an integer, the required sample size is n = 16.



8-2.3 One-Sided Confidence Bounds

Definition

A $100(1 - \alpha)\%$ upper-confidence bound for μ is

$$\mu \le u = \overline{x} + z_{\alpha} \sigma / \sqrt{n} \tag{8-9}$$

and a $100(1 - \alpha)\%$ lower-confidence bound for μ is

$$\overline{x} - z_{\alpha} \sigma / \sqrt{n} = l \le \mu$$
 (8-10)



8-2.5 A Large-Sample Confidence Interval for μ

Definition

When n is large, the quantity

$$\frac{\overline{X} - \mu}{S/\sqrt{n}}$$

has an approximate standard normal distribution. Consequently,

$$\overline{x} - z_{\alpha/2} \frac{s}{\sqrt{n}} \le \mu \le \overline{x} + z_{\alpha/2} \frac{s}{\sqrt{n}}$$
 (8-13)

is a large sample confidence interval for μ , with confidence level of approximately $100(1-\alpha)\%$.



Example 8-4

An article in the 1993 volume of the *Transactions of the American Fisheries Society* reports the results of a study to investigate the mercury contamination in largemouth bass. A sample of fish was selected from 53 Florida lakes and mercury concentration in the muscle tissue was measured (ppm). The mercury concentration values are

1.230	0.490	0.490	1.080	0.590	0.280	0.180	0.100	0.940
1.330	0.190	1.160	0.980	0.340	0.340	0.190	0.210	0.400
0.040	0.830	0.050	0.630	0.340	0.750	0.040	0.860	0.430
0.044	0.810	0.150	0.560	0.840	0.870	0.490	0.520	0.250
1.200	0.710	0.190	0.410	0.500	0.560	1.100	0.650	0.270
0.270	0.500	0.770	0.730	0.340	0.170	0.160	0.270	



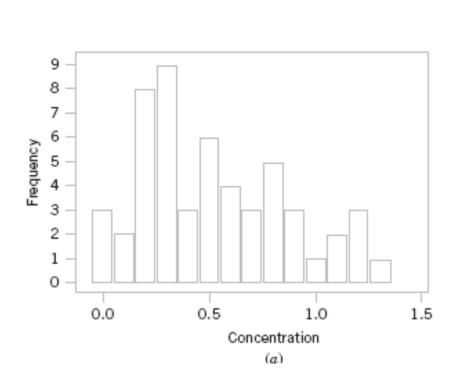
Example 8-4 (continued)

The summary statistics from Minitab are displayed below:

Variable	N	Mean	Median	TrMean	StDev	SE Mean
Concentration	53	0.5250	0.4900	0.5094	0.3486	0.0479
Variable	Minimum	Maximum	Q1	Q3		
Concentration	0.0400	1.3300	0.2300	0.7900		



Example 8-4 (continued)



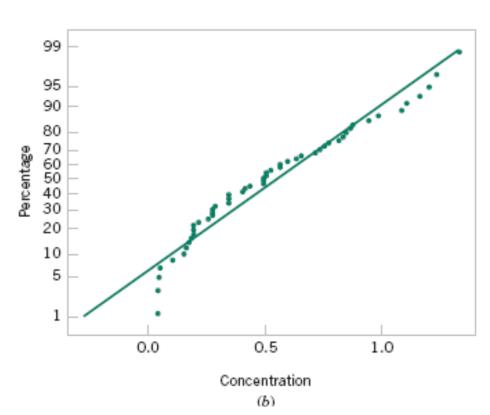


Figure 8-3 Mercury concentration in largemouth bass (a) Histogram. (b) Normal probability plot



Example 8-4 (continued)

Figure 8-3(a) and (b) presents the histogram and normal probability plot of the mercury concentration data. Both plots indicate that the distribution of mercury concentration is not normal and is positively skewed. We want to find an approximate 95% CI on μ . Because n > 40, the assumption of normality is not necessary to use Equation 8-13. The required quantities are n = 53, $\bar{x} = 0.5250$, s = 0.3486, and $z_{0.025} = 1.96$. The approximate 95% CI on μ is

$$\overline{x} - z_{0.025} \frac{s}{\sqrt{n}} \le \mu \le \overline{x} + z_{0.025} \frac{s}{\sqrt{n}}$$

$$0.5250 - 1.96 \frac{0.3486}{\sqrt{53}} \le \mu \le 0.5250 + 1.96 \frac{0.3486}{\sqrt{53}}$$

$$0.4311 \le \mu \le 0.6189$$

This interval is fairly wide because there is a lot of variability in the mercury concentration measurements.



8-3.1 The t distribution

Let X_1, X_2, \ldots, X_n be a random sample from a normal distribution with unknown mean μ and unknown variance σ^2 . The random variable

$$T = \frac{\overline{X} - \mu}{S/\sqrt{n}}$$
(8-15)

has a t distribution with n-1 degrees of freedom.



8-3.1 The t distribution

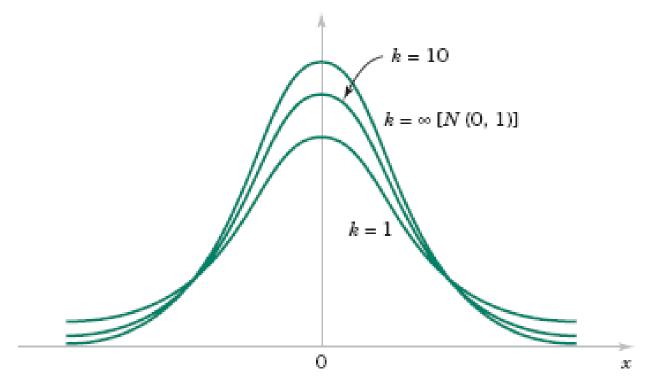


Figure 8-4 Probability density functions of several *t* distributions.



8-3.1 The t distribution

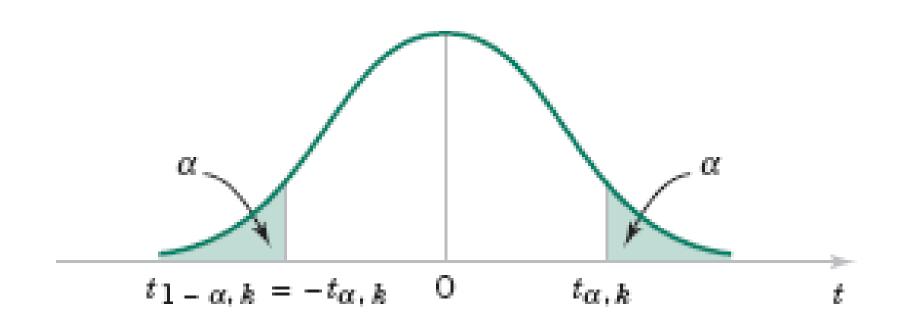


Figure 8-5 Percentage points of the *t* distribution.



8-3.2 The t Confidence Interval on μ

If \bar{x} and s are the mean and standard deviation of a random sample from a normal distribution with unknown variance σ^2 , a $100(1 - \alpha)$ percent confidence interval on μ is given by

$$\overline{x} - t_{\alpha/2, n-1} s / \sqrt{n} \le \mu \le \overline{x} + t_{\alpha/2, n-1} s / \sqrt{n}$$
 (8-18)

where $t_{\alpha/2,n-1}$ is the upper $100\alpha/2$ percentage point of the t distribution with n-1 degrees of freedom.

One-sided confidence bounds on the mean are found by replacing $t_{\alpha/2,n-1}$ in Equation 8-18 with $t_{\alpha,n-1}$.



Example 8-5

An article in the journal *Materials Engineering* (1989, Vol. II, No. 4, pp. 275–281) describes the results of tensile adhesion tests on 22 U-700 alloy specimens. The load at specimen failure is as follows (in megapascals):

19.8	10.1	14.9	7.5	15.4	15.4
15.4	18.5	7.9	12.7	11.9	11.4
11.4	14.1	17.6	16.7	15.8	
19.5	8.8	13.6	11.9	11.4	

The sample mean is $\bar{x} = 13.71$, and the sample standard deviation is s = 3.55. Figures 8-6 and 8-7 show a box plot and a normal probability plot of the tensile adhesion test data, respectively. These displays provide good support for the assumption that the population is normally distributed. We want to find a 95% CI on μ . Since n = 22, we have n - 1 = 21 degrees of freedom for t, so $t_{0.025,21} = 2.080$. The resulting CI is

$$\overline{x} - t_{\alpha/2, n-1} s / \sqrt{n} \le \mu \le \overline{x} + t_{\alpha/2, n-1} s / \sqrt{n}$$

$$13.71 - 2.080(3.55) / \sqrt{22} \le \mu \le 13.71 + 2.080(3.55) / \sqrt{22}$$

$$13.71 - 1.57 \le \mu \le 13.71 + 1.57$$

$$12.14 \le \mu \le 15.28$$

The CI is fairly wide because there is a lot of variability in the tensile adhesion test measurements.



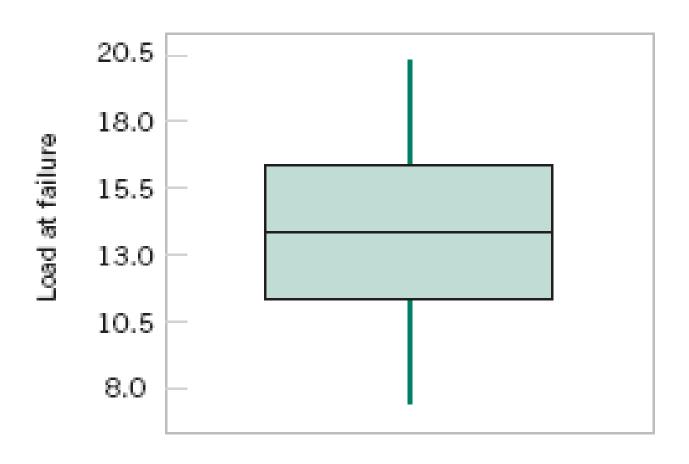


Figure 8-6 Box and Whisker plot for the load at failure data in Example 8-5.



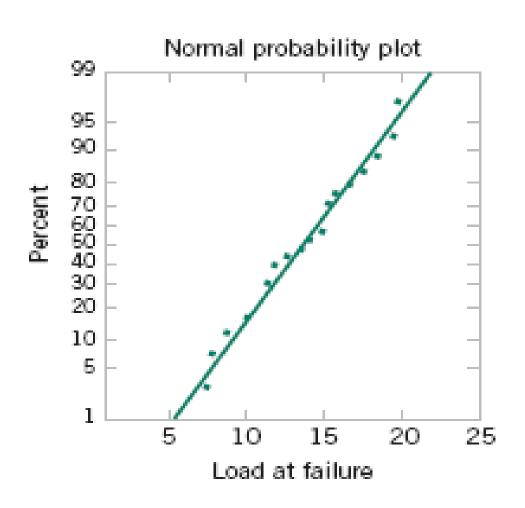


Figure 8-7 Normal probability plot of the load at failure data in Example 8-5.



Definition

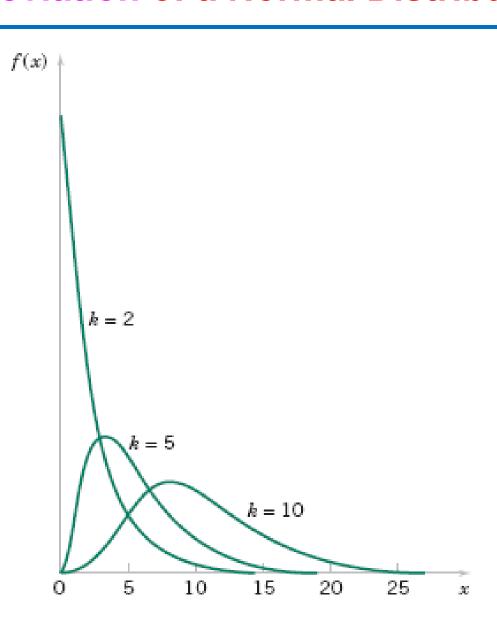
Let $X_1, X_2, ..., X_n$ be a random sample from a normal distribution with mean μ and variance σ^2 , and let S^2 be the sample variance. Then the random variable

$$X^2 = \frac{(n-1)S^2}{\sigma^2} \tag{8-19}$$

has a chi-square (χ^2) distribution with n-1 degrees of freedom.



Figure 8-8 Probability density functions of several χ^2 distributions.





Definition

If s^2 is the sample variance from a random sample of *n* observations from a normal distribution with unknown variance σ^2 , then a $100(1 - \alpha)\%$ confidence interval on σ^2 is

$$\frac{(n-1)s^2}{\chi^2_{\alpha/2,n-1}} \le \sigma^2 \le \frac{(n-1)s^2}{\chi^2_{1-\alpha/2,n-1}}$$
(8-21)

where $\chi^2_{\alpha/2,n-1}$ and $\chi^2_{1-\alpha/2,n-1}$ are the upper and lower $100\alpha/2$ percentage points of the chi-square distribution with n-1 degrees of freedom, respectively. A confidence interval for σ has lower and upper limits that are the square roots of the corresponding limits in Equation 8-21.



One-Sided Confidence Bounds

The $100(1 - \alpha)\%$ lower and upper confidence bounds on σ^2 are

$$\frac{(n-1)s^2}{\chi^2_{\alpha,n-1}} \le \sigma^2$$
 and $\sigma^2 \le \frac{(n-1)s^2}{\chi^2_{1-\alpha,n-1}}$ (8-22)

respectively.



Example 8-6

An automatic filling machine is used to fill bottles with liquid detergent. A random sample of 20 bottles results in a sample variance of fill volume of $s^2 = 0.0153$ (fluid ounces)². If the variance of fill volume is too large, an unacceptable proportion of bottles will be under- or overfilled. We will assume that the fill volume is approximately normally distributed. A 95% upper-confidence interval is found from Equation 8-22 as follows:

$$\sigma^2 \le \frac{(n-1)s^2}{\chi_{0.95,19}^2}$$

or

$$\sigma^2 \le \frac{(19)0.0153}{10.117} = 0.0287 \text{ (fluid ounce)}^2$$

This last expression may be converted into a confidence interval on the standard deviation σ by taking the square root of both sides, resulting in

$$\sigma \leq 0.17$$

Therefore, at the 95% level of confidence, the data indicate that the process standard deviation could be as large as 0.17 fluid ounce.



Normal Approximation for Binomial Proportion

If n is large, the distribution of

$$Z = \frac{X - np}{\sqrt{np(1-p)}} = \frac{\hat{P} - p}{\sqrt{\frac{p(1-p)}{n}}}$$

is approximately standard normal.

The quantity $\sqrt{p(1-p)/n}$ is called the standard error of the point estimator \hat{P} .

If \hat{p} is the proportion of observations in a random sample of size n that belongs to a class of interest, an approximate $100(1 - \alpha)\%$ confidence interval on the proportion p of the population that belongs to this class is

$$\hat{p} - z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \le p \le \hat{p} + z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$
 (8-25)

where $z_{\alpha/2}$ is the upper $\alpha/2$ percentage point of the standard normal distribution.



Example 8-7

In a random sample of 85 automobile engine crankshaft bearings, 10 have a surface finish that is rougher than the specifications allow. Therefore, a point estimate of the proportion of bearings in the population that exceeds the roughness specification is $\hat{p} = x/n = 10/85 = 0.12$. A 95% two-sided confidence interval for p is computed from Equation 8-25 as

$$\hat{p} - z_{0.025} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \le p \le \hat{p} + z_{0.025} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

or

$$0.12 - 1.96\sqrt{\frac{0.12(0.88)}{85}} \le p \le 0.12 + 1.96\sqrt{\frac{0.12(0.88)}{85}}$$

which simplifies to

$$0.05 \le p \le 0.19$$



Choice of Sample Size

The sample size for a specified value E is given by

$$n = \left(\frac{z_{\alpha/2}}{E}\right)^2 p(1 - p) \tag{8-26}$$

$$E = z_{\alpha/2} \sqrt{p(1-p)/n}$$

An upper bound on *n* is given by

$$n = \left(\frac{z_{\alpha/2}}{E}\right)^2 (0.25) \tag{8-27}$$



Example 8-8

Consider the situation in Example 8-7. How large a sample is required if we want to be 95% confident that the error in using \hat{p} to estimate p is less than 0.05? Using $\hat{p} = 0.12$ as an initial estimate of p, we find from Equation 8-26 that the required sample size is

$$n = \left(\frac{z_{0.025}}{E}\right)^2 \hat{p}(1-\hat{p}) = \left(\frac{1.96}{0.05}\right)^2 0.12(0.88) \approx 163$$

If we wanted to be at least 95% confident that our estimate \hat{p} of the true proportion p was within 0.05 regardless of the value of p, we would use Equation 8-27 to find the sample size

$$n = \left(\frac{z_{0.025}}{E}\right)^2 (0.25) = \left(\frac{1.96}{0.05}\right)^2 (0.25) \approx 385$$



One-Sided Confidence Bounds

The approximate $100(1 - \alpha)\%$ lower and upper confidence bounds are

$$\hat{p} - z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \le p$$
 and $p \le \hat{p} + z_{\alpha} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$ (8-28)

respectively.



Example

A manufacturer of electric calculators takes a random sample of 1200 calculators and finds that there are eight defective units. Construct a 96% lower confident bound for the population proportion about defective units.



IMPORTANT TERMS AND CONCEPTS

Confidence coefficient
Confidence interval
Confidence interval
for a population
proportion
Chi-squared
distribution

Confidence intervals
for the mean of a
normal distribution
Confidence interval for
the variance of a
normal distribution
Confidence level

Error in estimation
Large sample confidence
interval
One-sided confidence
bounds
Precision of parameter
estimation

Prediction interval
Tolerance interval
Two-sided confidence
interval
t distribution