

Introduction

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Summary

## Chapter 7: Point Estimation of Parameters

### Learning objectives

1. Introduction
2. Sampling Distributions
3. General Concepts Of Point
4. Estimation Methods Of Point Estimation

## Introduction

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## Summary

Let  $X$  is a random variable with probability distribution  $f(x)$ , which is characterized by the unknown  $\theta = (\theta_1, \theta_2, \dots, \theta_k)$

For example,  $X \sim N(\mu, \sigma^2)$  then  $\theta = (\mu, \sigma^2)$ .

How to “determine” the values of  $\theta$ ?

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In engineering, we often need to estimate

1. The mean  $\mu$  of a single population.
2. The variance  $\sigma^2$  (or standard deviation ) of a single population.
3. The proportion  $p$  of items in a population that belong to a class of interest.
4. The difference in means of two populations,  $\mu_1 - \mu_2$ .
5. The difference in two population proportions,  $p_1 - p_2$ .

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## The important results on point estimation

1. For  $\mu$  , the estimate is the  $\hat{\mu} = \bar{x}$ , **sample mean.**
2. For  $\sigma^2$  , the estimate is  $\hat{\sigma}^2 = s^2$  , **the sample variance.**
3. For  $p$ , the estimate is  $\hat{p} = x/n$  , **the sample proportion.**
4. For  $\mu_1 - \mu_2$  , the estimate is  $\hat{\mu}_1 - \hat{\mu}_2 = \bar{x}_1 - \bar{x}_2$
5. For  $p_1 - p_2$  , the estimate is  $\hat{p}_1 - \hat{p}_2$

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### Definition

### Random Sample

The random variables  $X_1, \dots, X_n$  are called a **random sample** of size  $n$  if

- The  $X_i$ 's are independent
- Every  $X_i$  has the same probability distribution

### Definition

### Statistic

- A **statistic**  $\hat{\Theta}$  is any function of the observations  $X_1, \dots, X_n$ :

$$\hat{\Theta} = h(X_1, \dots, X_n)$$

- The probability distribution of a statistic is called a **sampling distribution**.

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### Example

### Two important statistic

- Sample mean  $\bar{X}$

$$\bar{X} = \frac{X_1 + \dots + X_n}{n}$$

- Sample variance  $S^2$

$$S^2 = \frac{(X_1 - \bar{X})^2 + \dots + (X_n - \bar{X})^2}{n-1}$$

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## Theorem

In above example, if  $(X_1, \dots, X_n)$  is a random sample of size  $n$  take from a normal distribution  $N(\mu, \sigma^2)$  then

- $\bar{X}$  has a normal distribution  $N(\mu, \sigma^2/n)$

- $\frac{(n-1)S^2}{\sigma^2}$  has a chi-square distribution with  $n-1$  degrees of freedom (see pages 273-274).

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## Theorem

## Central Limit Theorem

Let  $(X_1, \dots, X_n)$  is a random sample of size  $n$  take from a population with mean  $\mu$  and finite variance  $\sigma^2$ , and if  $\bar{X}$  is the sample mean, the limiting form of the distribution of

$$Z = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}}$$

As  $n \rightarrow \infty$ , is the standard normal distribution.

**Remark:** The normal approximation for  $\bar{X}$  depends on the sample size  $n$ .



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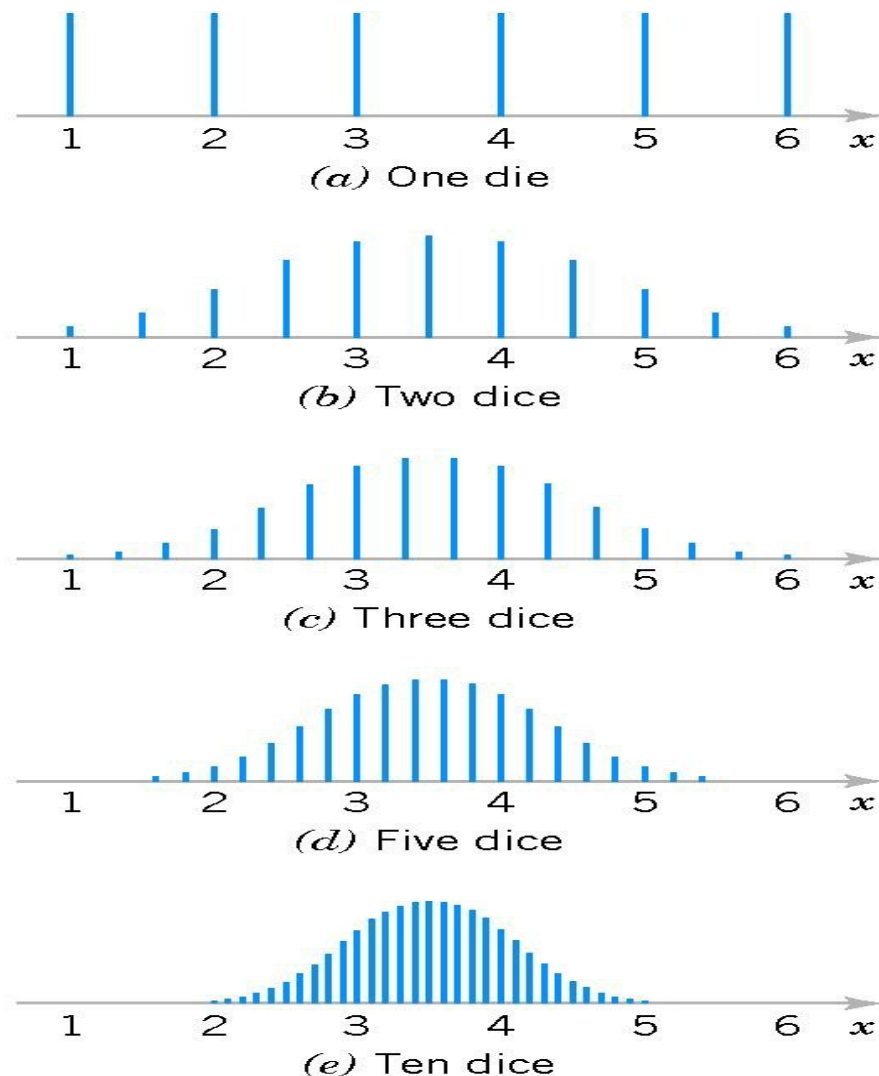
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**Figure 7-1**  
Distributions of  
average scores  
from throwing  
dice.



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### Definition

### Point Estimate

- A **point estimate** of some population parameter  $\theta$  is a single numerical value  $\hat{\theta}$  of a statistic  $\hat{\Theta}$ .
- The statistic  $\hat{\Theta}$  is called the **point estimator**.

## Two steps to find point estimation:

Step 1. Determine  $\hat{\Theta}$  by using the theoretical results.

Steps 2. Calculate  $\hat{\theta}$  from the experimental data.

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### Example

Market researchers use the number of sentences per advertisement as a measure of readability for magazine advertisements. The following represents a random sample of the number of sentences found in 15 advertisements. Find a point estimate of the population mean,  $\mu$ .

9 20 18 16 9 9 11 13 22 16 5 18 6 6 5

$$\text{Step 1. } \hat{\Theta} = \frac{X_1 + \dots + X_{15}}{15}$$

Step 2. A point estimate for  $\mu$  is

$$\hat{\mu} = \frac{9 + 20 + \dots + 6 + 5}{15} = \frac{183}{15} = 12.2$$

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## Definition

## Unbiased Estimator

The point estimator  $\hat{\Theta}$  is an **unbiased estimator** for the parameter  $\theta$  if

$$E(\hat{\Theta}) = \theta$$

If the estimator is not unbiased, then the difference

$$E(\hat{\Theta} - \theta)$$

is called the **bias** of the estimator  $\hat{\Theta}$ .

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## Example

Suppose that  $X$  is a random variable with mean  $\mu$  and variance  $\sigma^2$ . Let  $X_1, \dots, X_n$  be a random sample of size  $n$  from the population represented by  $X$ . Show that the sample mean  $\bar{X}$  and sample variance  $S^2$  are unbiased estimators of  $\mu$  and  $\sigma^2$ , respectively.

For the sample mean.

$$E(\bar{X}) = E \frac{X_1 + \dots + X_n}{n} = \frac{EX_1 + \dots + EX_n}{n} = \mu$$

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Now consider the **sample variance**. We have

$$\begin{aligned}
 E(S^2) &= E \left[ \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1} \right] = \frac{1}{n-1} E \sum_{i=1}^n (X_i - \bar{X})^2 \\
 &= \frac{1}{n-1} E \sum_{i=1}^n (X_i^2 + \bar{X}^2 - 2\bar{X}X_i) = \frac{1}{n-1} E \left( \sum_{i=1}^n X_i^2 - n\bar{X}^2 \right) \\
 &= \frac{1}{n-1} \left[ \sum_{i=1}^n E(X_i^2) - nE(\bar{X}^2) \right]
 \end{aligned}$$

Hence  $E(S^2) = \sigma^2$

# MINIMUM VARIANCE UNBIASED ESTIMATOR

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## Definition

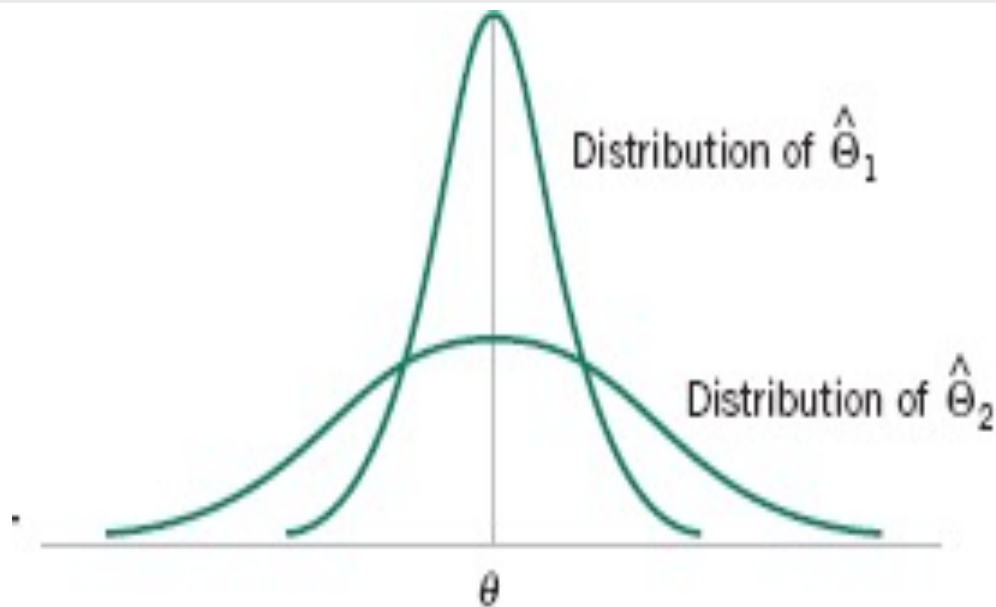
## Minimum Variance Unbiased Estimator

If we consider all unbiased estimators of  $\theta$ , the one with the smallest variance is called the **minimum variance unbiased estimator** (MVUE).

**Figure 7-5**

The sampling  
distributions of  
two unbiased  
estimators

$\hat{\Theta}_1$  and  $\hat{\Theta}_2$ .



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## Theorem

Let  $(X_1, \dots, X_n)$  is a random sample of size  $n$  take from a normal distribution  $N(\mu, \sigma^2)$ , the sample mean  $\bar{X}$  is the MVUE for  $\mu$ .



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## Definition Minimum Variance Unbiased Estimator

The **standard error** of an estimator  $\hat{\Theta}$  is its standard deviation, given by

$$\sigma_{\hat{\Theta}} = \sqrt{V(\hat{\Theta})}$$

If the standard error involves unknown parameters that can be estimated, substitution of those values into  $\sigma_{\hat{\Theta}}$  produces an **estimated standard error**, denoted by  $\hat{\sigma}_{\hat{\Theta}}$

## Remark

If the random sample of size  $n$  take from a normal distribution  $N(\mu, \sigma^2)$  then  $\bar{X}$  has a normal distribution  $N(\mu, \sigma^2/n)$ .

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Standard error of  $\bar{X}$  is

$$\sigma_{\hat{X}} = \frac{\sigma}{\sqrt{n}}$$

If we did not know  $\sigma$  then the estimated standard error of  $\bar{X}$  would be

$$\hat{\sigma}_{\hat{X}} = \frac{s}{\sqrt{n}}$$

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## Exercise

A new method of measuring the thermal conductivity of Armco iron: Using a temperature of  $100^{\circ}\text{F}$  and a power input of 550 watts, the following 10 measurements of thermal conductivity (in  $\text{Btu/hr-ft-}^{\circ}\text{F}$ ) were obtained:

41.60, 41.48, 42.34, 41.95, 41.86,  
42.18, 41.72, 42.26, 41.81, 42.04

a, Find a point estimate of the mean thermal conductivity at  $100^{\circ}\text{F}$  and 550 watts.

b, Find the estimated standard error of  $\bar{X}$

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## The $k$ th moments

Let  $(X_1, \dots, X_n)$  be a random sample from random variable  $X$  with the probability distribution  $f(x)$

- The  $k$ th moment of  $X$  are denoted by  $M_k := E(X^k)$
- The corresponding  $k$ th **sample moment** is

$$\bar{X}^k = \frac{X_1^k + \dots + X_n^k}{n}$$

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## The $k$ th moments

Let  $(X_1, \dots, X_n)$  be a random sample from either a probability mass function or probability density function with  $m$  unknown parameters  $(\theta_1, \theta_2, \dots, \theta_m)$ . The moment estimators  $\hat{\Theta}_1, \dots, \hat{\Theta}_m$  are found the following system

$$\begin{cases} M_1 = \bar{X}^1 \\ \dots\dots\dots \dots \\ M_m = \bar{X}^m \end{cases}$$

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## Example

Suppose that  $X_1, \dots, X_n$  be a random sample of size  $n$  from a normal distribution  $N(\mu, \sigma^2)$ . Find the moment estimators of  $\mu$  and  $\sigma^2$ .

We have to solve following system

$$\begin{cases} M_1 = \bar{X}^1 \\ M_2 = \bar{X}^2 \end{cases} \quad \text{or} \quad \begin{cases} EX = \frac{X_1 + \dots + X_n}{n} \\ EX^2 = \frac{X_1^2 + \dots + X_n^2}{n} \end{cases}$$

Here  $E(X) = \mu$  and  $E(X^2) = V(X) + (EX)^2 = \sigma^2 + \mu^2$

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Hence, by solving the above system we have

$$\mu = \bar{X}, \quad \sigma^2 = \frac{\sum_{i=1}^n X_i^2 - n(\bar{X})^2}{n} = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n}$$

Finally, the moment estimators are

$$\hat{\mu} = \bar{X},$$

Unbiased  
estimator

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n}$$

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## Maximum likelihood Estimator

Suppose that  $X$  is a random variable with probability distribution  $f(x, \theta)$ , where  $\theta$  is a single unknown parameter. Let  $x_1, x_2, \dots, x_n$  be the observed values in a random sample of size  $n$ . Then the **likelihood function** of the sample is

$$L(\theta) = f(x_1, \theta) f(x_2, \theta) \dots f(x_n, \theta)$$

The maximum likelihood estimator (MLE) of  $\theta$  is the value of  $\theta$  that maximizes the  $L(\theta)$ .



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## Example

Let  $X$  is a Bernoulli random variable. The probability mass function is

$$f(x, p) = \begin{cases} p^x (1-p)^{1-x}, & x = 0, 1 \\ 0 & \text{otherwise} \end{cases}$$

Find the maximum likelihood estimator of  $p$ .

The likelihood function of a random sample of size  $n$  is

$$\begin{aligned} L(p) &= p^{x_1} (1-p)^{1-x_1} p^{x_2} (1-p)^{1-x_2} \cdots p^{x_n} (1-p)^{1-x_n} \\ &= \prod_{i=1}^n p^{x_i} (1-p)^{1-x_i} = p^{\sum_{i=1}^n x_i} (1-p)^{n - \sum_{i=1}^n x_i} \end{aligned}$$

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$$\ln L(p) = \left( \sum_{i=1}^n x_i \right) \ln p + \left( n - \sum_{i=1}^n x_i \right) \ln(1 - p)$$

$$\frac{d \ln L(p)}{dp} = \frac{\sum_{i=1}^n x_i}{p} - \frac{\left( n - \sum_{i=1}^n x_i \right)}{1 - p}$$

The maximum likelihood estimator of  $p$  is

$$\hat{P} = \frac{X_1 + \dots + X_n}{n}$$

Is this a unbiased estimator?

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## Exercise

Let  $X$  is a random variable, which has probability mass function is

$$f(x) = \begin{cases} p & , \quad x = -1 \\ 1 - p & , \quad x = 1 \\ 0 & , \quad otherwise \end{cases}$$

Find the maximum likelihood estimator of  $p$ .

Solution:

$$f(x) = \begin{cases} p^{\frac{1-x}{2}} (1-p)^{\frac{1+x}{2}} & , \quad x = -1, 1 \\ 0 & , \quad otherwise \end{cases}$$

$$\hat{p} = \frac{1}{2} - \frac{X_1 + \dots + X_n}{2n}$$

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## Invariance Property

Let  $\hat{\theta}_1, \dots, \hat{\theta}_k$  be the maximum likelihood estimators of the parameters  $\theta_1, \dots, \theta_k$ . Then the maximum likelihood estimator of any function  $h(\theta_1, \dots, \theta_k)$  of these parameters is the same function  $h(\hat{\theta}_1, \dots, \hat{\theta}_k)$  of the estimators  $\hat{\theta}_1, \dots, \hat{\theta}_k$ .

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## Exercise

Let  $X$  be normally distributed  $N(\mu, \sigma^2)$ , where both  $\mu$  and  $\sigma^2$  are unknown.

- (a) Find the maximum likelihood estimator of  $\mu$  and  $\sigma^2$ .
- (b) Find the maximum likelihood estimator of  $\theta = \mu^2\sigma^2$ .

Solution:

$$\hat{\mu} = \bar{X}, \quad \hat{\sigma}^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n}$$

$$\hat{\theta} = (\bar{X})^2 \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n}$$

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We have studied:

1. Sampling distributions
2. General concepts of point estimation: Unbiased Estimator
3. Methods of point estimation
  - Method of Moments
  - Method of Maximum Likelihood

Homework: Read slides of the next lecture.