

# STAT011 Statistical Methods I

### Lecture 16 Two-Sample t Procedures I

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### Review

- ▶ Sample standard deviation (SD)
- Degree of freedom
- Standard error (SE)
  - SD of a statistic estimated from sample data
- t distribution dt( , df = ), pt( , df = ), qt( , df = )
- ▶ Statistical inference for a population mean based on *t* distribution
  - One-sample t confidence interval  $\bar{x} \pm t^* \frac{s}{\sqrt{n}}$
  - One-sample t test  $H_0: \mu = \mu_0$ ,  $t = \frac{\bar{x} \mu_0}{s/\sqrt{n}} \stackrel{approx.}{\sim} t(n-1)$
- Examples

```
t.test( , conf.level = ), t.test( , alternative = , mu = )
```

# Review - Comparing z and t procedures

	z procedures	t procedures
Population SD $\sigma$	Known	Unknown, use sample SD s
Level C C.I.	$\bar{x} \pm z^* \frac{\sigma}{\sqrt{n}}$ $z^* = \operatorname{qnorm}(1-(1-C)/2)$	$\bar{x} \pm t^* \frac{s}{\sqrt{n}}$ $t^* = qt(1-(1-C)/2, df=n-1)$
Level α significance test	$H_0: \mu = \mu_0$ $H_a: \mu > \mu_0, \mu < \mu_0, \mu \neq \mu_0$ $z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}} \stackrel{approx.}{\sim} N(0, 1)$ $P(Z \leq z), \text{pnorm(z)}$ $P(Z \geq z), 1 - \text{pnorm(z)}$ $2P(Z \geq  z ), 2 * (1 - \text{pnorm(abs(z))})$	$H_0: \mu = \mu_0$ $H_a: \mu > \mu_0, \mu < \mu_0, \mu \neq \mu_0$ $t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \stackrel{approx.}{\sim} t(n-1)$ $P(T \leq t), \text{pt(t,df=n-1)}$ $P(T \geq t), \text{1-pt(t,df=n-1)}$ $2P(T \geq  t ), \text{2*(1-pt(abs(t),df=n-1)})$

### Outline

- ▶ Matched-pairs two-sample *t* procedures
- Two-sample *t* procedures
  - $\blacksquare$  Two-sample t confidence interval
  - $\blacksquare$  Two-sample t test
- ▶ Pooled two-sample *t* procedures (Lecture 17)
  - Pooled two-sample t confidence interval
  - Pooled two-sample *t* test

- Example: compare mean height of female students and mean height of male students.
- Although called "two-sample", it is not necessarily two data samples. It is usually a single data sample with a quantitative variable (*Height*) and a binary variable that has two categories (*Gender*: female or male).
  - We compare of the mean of *Height* for the two categories of *Gender*.
- In two-sample problems, we evaluate the relationship between a quantitative response variable and a categorical explanatory variable.
  - Are different levels of the categorical variable associated with different means of the quantitative variable? Do female and male students have different height?
  - The relationship is evaluated by **comparing the means of the response** variable for the two groups of the categorical variable.

**Example 1** Homework 7 Q7: compare *Weight* for the two groups of values, before and after the study.

**Example 2** STAT 11 Survey: compare *Height* for male and female students.

```
head(Survey[, c("Height", "Gender")], 4)
```

```
## Height Gender
## 1 61 Female
## 2 66 Male
## 3 70 Male
## 4 63 Female
```

▶ What is the difference in data between the two examples?

62.3

**Example 1** Homework 7 Q7: compare *Weight* for the two groups of values, before and after the study.

```
## ID WeightBefore WeightAfter
## 1 1 55.7 61.7
## 2 2 54.9 58.8
## 3 3 59.6 66.0
## 4 4 4 66.2
```

**Example 2** STAT 11 Survey: compare *Height* for male and female students.

66.2

```
head(Survey[, c("Height", "Gender")], 4)
```

```
## Height Gender
## 1 61 Female
## 2 66 Male
## 3 70 Male
## 4 63 Female
```

## 4

- ▶ The values from the two groups are independent from each other.
- ▶ This is a regular two-sample problem.

- Matched-pairs two-sample problem
  - Values are matached as a pair
  - Values of one group depends on the values of the other
    - A subject with a high *Weight* before the study will (generally) have higher *Weight* after the study.
  - The sample sizes for the two groups are the same
- ▶ **Matched-pairs two-sample problems** can be solved by taking the difference between the paried values and using **one-sample** *t* **prcedures**.
- Regular two-sample problems
  - The values in one group are **independent** of the values in the other group
  - The the sample sizes for the two groups are *not necessarily* the same
- ▶ **Regular two-sample problems** are solved by **two-sample** *t* **prcedures**.

# Matched-pairs two-sample problems

- Matched-pairs two-sample problem
  - Values are matached as a pair
  - Values of one group depends on the values of the other
    - A subject with a high *Weight* before the study will (generally) have higher *Weight* after the study.
  - The sample sizes for the two groups are the same
- ▶ **Matched-pairs two-sample problems** can be solved by taking the difference between the paried values and using **one-sample** *t* **prcedures**.
  - Evaluating whether  $\mu_{WeightBefore} = \mu_{WeightAfter}$  is equivalent to whether  $\mu_{WeightBefore} \mu_{WeightAfter}$  or whether  $\mu_{WeightChange} = 0$ 
    - Construct a CI for  $\mu_{WeightChange}$  and see whether 0 falls into the interval
    - Conduct a test of  $H_0: \mu_{WeightChange} = 0$

- Question of interest:
  - Is population mean of group 1 the same as the population mean of group 2?

Group	Population Mean	Population SD	Sample Mean	Sample SD
1.	$\mu_1$	$\sigma_1$	$ar{x}_1$	$s_1$
2.	$\mu_2$	$\sigma_2$	$\bar{x}_2$	$s_2$

- We are interested the difference between the population means,  $\mu_1 \mu_2$ .
- Confidence interval for  $\mu_1 \mu_2$ , Significance test for  $H_0: \mu_1 = \mu_2$  vs.  $H_a: \mu_1 > \mu_2$  or  $\mu_1 < \mu_2$  or  $\mu_1 \neq \mu_2$ .
- We use the observed sample means  $\bar{x}_1$  and  $\bar{x}_2$  to make inference about the population means  $\mu_1$  and  $\mu_2$ .

#### Suppose population SDs $\sigma_1$ and $\sigma_2$ are known.

- Estimate  $\mu_1 \mu_2$  from sample data:  $\bar{x}_1 \bar{x}_2$
- ▶ By CLT

$$\bar{x}_1 \sim N\left(\mu_1, \frac{\sigma_1}{\sqrt{n_1}}\right) \text{ and } \bar{x}_2 \sim N\left(\mu_2, \frac{\sigma_2}{\sqrt{n_2}}\right)$$

• What is the distribution of  $\bar{x}_1 - \bar{x}_2$ ?

```
set.seed(15)
\# x1 \sim N(10, 3), x2 \sim N(5, 4)
x1 <- rnorm(1000, mean = 10, sd = 3)
x2 <- rnorm(1000, mean = 5, sd = 4)
c(mean(x1), mean(x2), sd(x1), sd(x2), var(x1), var(x2))
## [1] 10.110928 4.873001 3.072772 4.022100 9.441929 16.177286
# Mean and SD of x1 + x2
c(mean(x1 + x2), sd(x1 + x2), var(x1+x2))
## [1] 14.983929 5.114448 26.157578
# Mean and SD of x1 - x2
c(mean(x1 - x2), sd(x1 - x2), var(x1-x2))
## [1] 5.237927 5.008079 25.080851
 \mu_{X_1 \pm X_2} = \mu_{X_1} \pm \mu_{X_1}
Var_{X_1 \pm X_2} = Var_{X_1} + Var_{X_1} \Rightarrow \sigma_{X_1 \pm X_2} = \sqrt{\sigma_{X_1}^2 + \sigma_{X_1}^2}
```

$$\bar{x}_1 \sim N\left(\mu_1, \frac{\sigma_1}{\sqrt{n_1}}\right) \text{ and } \bar{x}_2 \sim N\left(\mu_2, \frac{\sigma_2}{\sqrt{n_2}}\right)$$

Mean of  $\bar{x}_1 - \bar{x}_2$ :

$$\mu_{\bar{x}_1 - \bar{x}_2} = \mu_{\bar{x}_1} - \mu_{\bar{x}_2} = \mu_1 - \mu_2$$

▶ SD of  $\bar{x}_1 - \bar{x}_2$ :

$$\sigma_{\bar{x}_1 - \bar{x}_2} = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

Therefore

$$\bar{x}_1 - \bar{x}_2 \sim N\left(\mu_1 - \mu_2, \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$$

Standardize  $\bar{x}_1 - \bar{x}_2$ ,

$$\frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1)$$

#### Two-sample z statistic

Suppose that  $\bar{x}_1$  is the mean of an SRS of size  $n_1$  drawn from an  $N(\mu_1, \sigma_1)$  population and that  $\bar{x}_2$  is the mean of an independent SRS of size  $n_2$  drawn from an  $N(\mu_2, \sigma_2)$  population. Then the two-sample z statistic

$$z = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

has the standard Normal N(0, 1) sampling distribution.

▶ What if the population SDs are **unknown**?

**Two-sample** *t* **statistic**: Replace  $\sigma_1$  and  $\sigma_2$  in the *z* statistic by  $s_1$  and  $s_2$ 

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t(k)$$

The two-sample *t* statistic approximates a *t*(*k*) distribution with degree of freedom

$$k \approx \frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{1}{n_1 - 1} \left(\frac{s_1^2}{n_1}\right)^2 + \frac{1}{n_2 - 1} \left(\frac{s_2^2}{n_2}\right)^2}$$
 (Welch-Satterthwaite formula) or

 $k \approx \min(n_1 - 1, n_2 - 1)$  (the smaller of  $n_1 - 1$  and  $n_2 - 1$ )

# Two-sample t confidence interval

Suppose that an SRS of size  $n_1$  is drawn from a Normal population with unknown mean  $\mu_1$  and that an independent SRS of size  $n_2$  is drawn from another Normal population with unknown mean  $\mu_2$ . The confidence interval for  $\mu_1 - \mu_2$  given by

$$(\bar{x}_1 - \bar{x}_2) \pm t^* \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

has confidence level at least C. Here,  $t^*$  is the value for the t(k) density curve with area C between  $-t^*$  and  $t^*$ , where k is approximated by either the Welch-Satterthwaite formula or the smaller of  $n_1 - 1$  and  $n_2 - 1$ .

# Two-sample t test

To test the hypothesis  $H_0: \mu_1 - \mu_2 = 0$ , compute the two-sample t statistic

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - 0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \stackrel{approx.}{\sim} t(k)$$

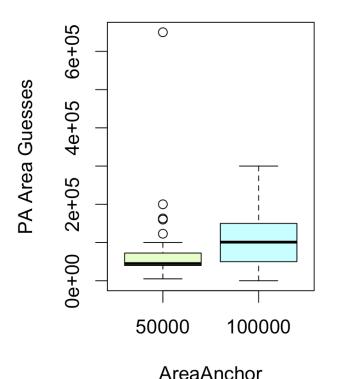
In terms of a random variable T having the t(k) distribution (k is approximated by either the Welch-Satterthwaite formula or the smaller of  $n_1 - 1$  and  $n_2 - 1$ ), the P-value for a test of  $H_0$  against

 $H_a: \mu_1 > \mu_2 \text{ is } P(T \ge t)$ 

 $H_a: \mu_1 < \mu_2 \text{ is } P(T \leq t)$ 

 $H_a: \mu_1 \neq \mu_2 \text{ is } 2P(T \geq |t|)$ 

#### **Boxplot of AreaGuess**



#### **Version 1:**

- Is the area of Pennsylvania more or less than 50,000 square miles?
- Give your best guess at the area of Pennsylvania in square miles.

#### Version 2:

- Is the area of Pennsylvania more or less than 100,000 square miles?
- Give your best guess at the area of Pennsylvania in square miles.

 $\bar{x}_2 = 109.7$ ,  $s_2 = 74.6$ ,  $n_2 = 21$ 

```
mysummary <- function(x){
    c(mean=mean(x), sd=sd(x), n=length(x))
}
# Transform Area in miles to thousand square miles
Survey$AreaGuess <- Survey$AreaGuess/1000
aggregate(AreaGuess ~ AreaAnchor, data=Survey, FUN=mysummary)

## AreaAnchor AreaGuess.mean AreaGuess.sd AreaGuess.n
## 1 50000 62.85715 70.18477 91.00000
## 2 100000 109.70252 74.57255 21.00000

\bar{x}_1 = 62.9, s_1 = 70.2, n_1 = 91
```

▶ A level 0.05 test for whether the two population means are the same or not.

#### 95% confidence interval (by hand)

$$(\bar{x}_1 - \bar{x}_2) \pm t^* \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

$$= (62.9 - 109.7) \pm 2.09 \sqrt{\frac{70.2^2}{91} + \frac{74.6^2}{21}}$$

$$= -46.8 \pm 37.3$$

$$t^* = 2.09 = \text{qt}(0.975, \text{ df}=20)$$
 (the smaller of  $n_1 - 1$  and  $n_2 - 1$  is  $21 - 1 = 20$ )

- ▶ We are 95% confidence that the interval [−84.1, −9.5] will contain the true population mean difference in *AreaGuess*.
- The interval does not contain 0. So the difference in *AreaGuess* between the two groups is significantly different from 0 wording significantly affected the area of PA guessed by the students.

#### Level 0.05 test (by hand)

 $H_0: \mu_1 = \mu_2 \text{ vs. } H_a: \mu_1 \neq \mu_2$ 

 $t = \frac{(\bar{x}_1 - \bar{x}_2) - 0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} = \frac{(62.9 - 109.7) - 0}{\sqrt{\frac{70.2^2}{91} + \frac{74.6^2}{21}}}$ = -2.63

df= 20, 
$$t < t^* = -2.09$$
 (qt(0.975, df=20)) or  $P = 0.016 < 0.05$ , 2\*(1-pt(2.63,df=20))

We reject  $H_0$  at level 0.05. There is significant difference in *AreaGuess* between the two groups. Wording significantly affected the area of PA guessed by the students.

# Two-sample t procedure in R

#### 95% confidence interval and level 0.05 test using R

```
\# t.test(x = , y = , alternative = , mu = , conf.level = )
t.test(Survey$AreaGuess[Survey$AreaAnchor=="50000"],
       Survey$AreaGuess[Survey$AreaAnchor=="100000"])
##
##
    Welch Two Sample t-test
##
## data: Survey$AreaGuess[Survey$AreaAnchor == "50000"] and Survey$AreaGuess[Survey$
## t = -2.6231, df = 28.745, p-value = 0.0138
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -83.38516 -10.30558
## sample estimates:
                              ▶ 95% CI: [-83.4, -10.3]
## mean of x mean of y
## 62.85715 109.70252
                              • Level 0.05 test: H_0: \mu_1 = \mu_2 \text{ vs. } H_a: \mu_1 \neq \mu_2
                                t = -2.62, df = 28.75 and P = 0.014 < 0.05
```

# Two-sample t procedure in R

95% confidence interval and level 0.05 test using R (simpler coding)

```
# t.test(Response ~ Explanatory, data = , alternative = , mu = , conf.level = )
t.test(AreaGuess ~ AreaAnchor, data=Survey)
##
##
   Welch Two Sample t-test
##
## data: AreaGuess by AreaAnchor
## t = -2.6231, df = 28.745, p-value = 0.0138
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -83,38516 -10,30558
## sample estimates:
  mean in group 50000 mean in group 100000
##
##
               62.85715
                                   109.70252
```

Note: the results are slightly different from the calculations by hand, where we use the smaller of  $n_1 - 1$  and  $n_2 - 1$  as the degree of freedom. Here, the R function uses the Welch-Satterthwaite formula for df.

# Summary

- ► Matched-pairs two-sample t procedures
  - Use one-sample t procedures
- Two-sample t procedures
  - Two-sample t confidence interval  $(\bar{x}_1 \bar{x}_2) \pm t^* \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$
  - Two-sample t test  $t = \frac{(\bar{x}_1 \bar{x}_2) 0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t(k)$ 
    - k is approximated by either the Welch-Satterthwaite formula or the smaller of  $n_1 1$  and  $n_2 1$
  - t.test(x = , y = ) or t.test(Reponse ~ Explanatory, data = )