

STAT021 Statistical Methods II

Lecture 18 Transforming Predictors

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Outline

The purpose of transforming the predictors is to achieve better model fitting, i.e., to explain as much variability in the reponse variable as possible while at the same time to keep the model simple.

There are many different ways to transform the predictors, such as

- ▶ Applying a function on the predictor (e.g., natrual logarithm function)
- Categorizing a quantitative predictor
- ▶ Treating a categorical predictor quantitative
- ▶ Including the polynomial terms of a quantitative predictor
- ...

Categorization is common

head(HappyPlanet)

```
##
                Country Happiness
                                       GDPpc
                                                   HDI HDI2
                                                                HDI4
## 1
            Philippines
                         59.17430 1678.8520 0.6682183 Low
                                                              Medium
## 2
                 Rwanda 28.34747 398.2085 0.4832405 Low
                                                                 Low
## 3
                Hungary 37.63759 13842.6055 0.8283505 High VeryHigh
## 4
                 Cyprus 45.99012 31386.6326 0.8497454 High VeryHigh
## 5 Trinidad and Tobago 51.86844 16530.1804 0.7718938 High
                                                                High
## 6
                         51.12862
                                   2312.1925 0.6791644 Low
                                                              Medium
               Paraguay
```

To categorize a quantitative variable, the cutoffs may be chosen by

- quantiles
- practical meaning
- model fitting (try different cutoffs and choose the set that results in best model fitting)
- convenience

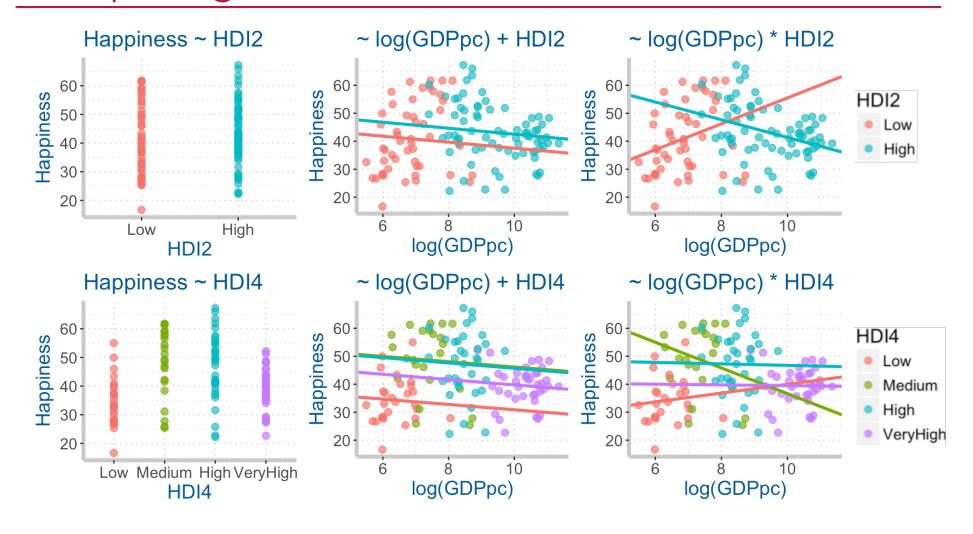
Categorization is common

Example: the Breast Cancer Risk Assessment Tool for breast cancer prediction has the following predictors

- Age $(< 50, \ge 50)$
- Age at first period $(7~11, 12~13, \ge 14 \text{ or unknown})$
- Age at first live birth (< 20 or unknown, $20\sim24$, $25\sim29$ or never, ≥ 30)
- Number of first-degree relatives with breast cancer (0 or unknown, 1, > 1)
- Number of past breast biopsies (0 or unknown, 1 or unknown but with positive, > 1)
- Race/ethnicity
- ▶ Breast density (0~0.25, 0.25~0.5, 0.5~0.75, 0.75~1)

Why categorization?

Comparing the six models - best model?



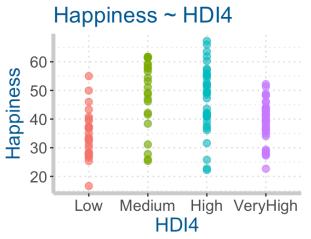
Comparing the six models

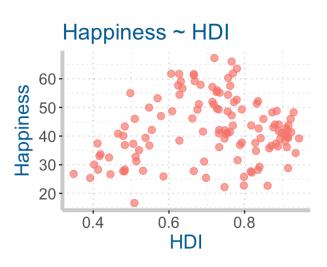
Happiness ~	F test	R^2	R^2_{adj}
1. <i>HDI2</i>	F = 1.2, P = 0.281	0.0095	0.0014
2. log(GDPpc) + HDI2	F = 1.1, P = 0.328	0.0183	0.0021
3. log(GDPpc) * HDI2	F = 4.9, P = 0.003	0.1088	0.0865
4. <i>HDI4</i>	$F = 13.5, P = 1.21 \times 10^{-7}$	0.2522	0.2335
5. log(GDPpc) + HDI4	$F = 10.2, P = 4.13 \times 10^{-7}$	0.2546	0.2295
6. log(GDPpc) * HDI4	$F = 6.1, P = 4.82 \times 10^{-7}$	0.2680	0.2238

Note: In model 5, log(GDPpc) is not significant; In model 6, log(GDPpc) and the interaction terms are not significant.

Model 4 (*Happiness* ~ *HDI4*) is the best one based on adjusted R^2 and nested F tests.

Why categorizing predictors?

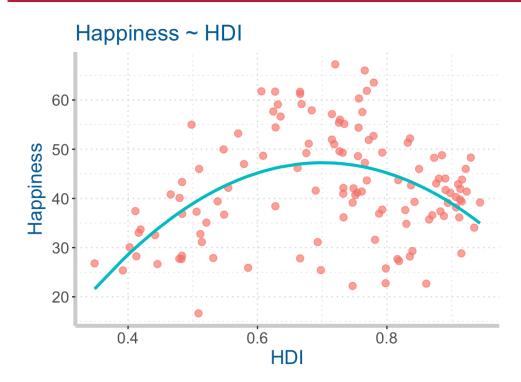




Happiness ~	F test	R^2	R^2_{adj}
HDI4	$F = 13.5, P = 1.21 \times 10^{-7}$	0.2522	0.2335
HDI	F = 2.68, P = 0.10	0.0215	0.0135

- Linear regression of *Happiness* and *HDI* assumes linear relationship, which is in fact a very strong assumption.
- However, coding *HDI* as a categorical variable with four categories assumes means of *Happiness* are different for the four *HDI* levels.
- Categorizing a predictor allows more flexibility in model assumptions about the relationship between the response variable and the predictor but adds more complexity (more parameters) to the model.

Another way to model Happiness ~ HDI



- The relationship between *Happiness* and *HDI*: as *HDI* increases, *Happiness* first goes up and then goes down non-linear.
- ▶ Polynomial Regression

56 perches were caught in a lake in Finland and three variables were measured.

Weight (in grams)

head(perch, 3)

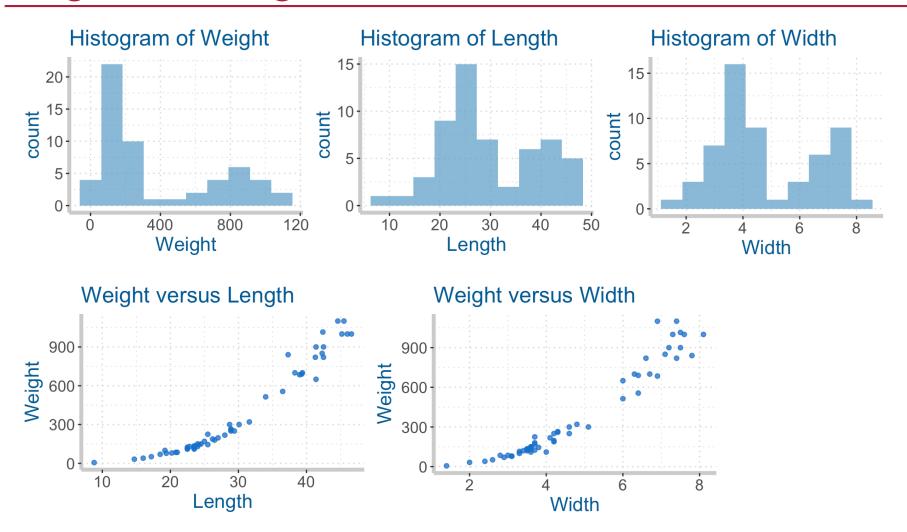
- *Length* (in centimeters)
- Width (in centimeters)

Weight 56 382.24 347.62 ## Length 56 29.57 9.53 ## Width 56 4.74 1.78

```
## Weight Length Width
## 1 5.9 8.8 1.4
## 2 32.0 14.7 2.0
## 3 40.0 16.0 2.4

library(psych) # package for the describe() function
describe(perch)[,2:4] # summary statistics

## n mean sd
```



For a single quantitative predictor *X*, a **polynomial regression model** of degree *k* has the form

$$Y = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_k X^k + \epsilon, \epsilon \stackrel{iid}{\sim} N(0, \sigma)$$

For a single quantitative predictor X, a **quadratic regression model** has the form

$$Y = \beta_0 + \beta_1 X + \beta_2 X^2 + \epsilon, \epsilon \stackrel{iid}{\sim} N(0, \sigma)$$

- We do not consider X and X^2 as two different predictors. Instead, we assume the relationship between Y and the single predictor X is nonlinear and quadratic.
- Usually, if the t test for the slope of the quadratic term X^2 is NOT significant, we may remove the quadratic term from the model (or try a higher order term).
- ▶ However, as long as the highest order term is significant, we keep all the lower order terms.

Response variable: Weight

Predictors: Length (L), Width (W)

We will consider 12 different models that involve *Length* and *Width*.

Complete second-order model is the model that includes linear and quadratic terms for both predictors along with the interaction term.

- 1. Weight $\sim L$
- 2. Weight $\sim L + L^2$
- 3. Weight $\sim W$
- 4. Weight $\sim W + W^2$
- 5. Weight $\sim L + W$
- 6. Weight $\sim L + L^2 + W$
- 7. Weight $\sim L + W + W^2$
- 8. Weight $\sim L + L^2 + W + W^2$
- 9. Weight $\sim L + W + LW$
- 10. Weight $\sim L + L^2 + W + LW$
- 11. Weight $\sim L + W + W^2 + LW$
- 12. Weight $\sim L + L^2 + W + W^2 + LW$

```
summary(m1 <- lm(Weight ~ Length, data=perch))</pre>
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -652.787 43.407 -15.04 <2e-16 ***
## Length 35.001 1.398 25.03 <2e-16 ***
##
## Residual standard error: 98.82 on 54 degrees of freedom
## Multiple R-squared: 0.9207, Adjusted R-squared: 0.9192
## F-statistic: 626.5 on 1 and 54 DF, p-value: < 2.2e-16
summary(m2 <- lm(Weight ~ Length + I(Length^2), data=perch))</pre>
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 128.34533 78.77870 1.629 0.10920
## Length -21.02388 5.41770 -3.881 0.00029 ***
## I(Length^2) 0.90862 0.08689 10.458 1.72e-14 ***
##
## Residual standard error: 56.99 on 53 degrees of freedom
## Multiple R-squared: 0.9741, Adjusted R-squared: 0.9731
## F-statistic: 996.6 on 2 and 53 DF, p-value: < 2.2e-16
```

- ▶ The cubic term is not significant and thus not necessary in the model.
- ▶ The model

$$\widehat{Weight} = 128.3 - 21.0 \times Length + 0.9 \times Length^2$$

seems the best among the three.

```
## (Intercept) -31.735 91.908 -0.345 0.731
## Width -25.635 39.487 -0.649 0.519
## I(Width^2) 20.934 3.827 5.470 1.24e-06 ***
##
## Residual standard error: 75.05 on 53 degrees of freedom
## Multiple R-squared: 0.9551, Adjusted R-squared: 0.9534
## F-statistic: 563.5 on 2 and 53 DF, p-value: < 2.2e-16
```

• Because the quadratic term is significant, although the linear term is not, we still keep both terms in the model.

```
## Length 14.307 5.659 2.528 0.014475 *
## Width 113.500 30.265 3.750 0.000439 ***
##
## Residual standard error: 88.68 on 53 degrees of freedom
## Multiple R-squared: 0.9373, Adjusted R-squared: 0.9349
## F-statistic: 396.1 on 2 and 53 DF, p-value: < 2.2e-16</pre>
```

- Length and Width each has a significant quadratic relationship with Weight. Let's now evaluate models with both predictors, starting from the linear terms.
- Given Width (Length) is held constant, Weight and Length (Width) has a significant linear relationship.
- But this model explains slightly less variability ($R^2 = 0.94$) than the previous quadratic models ($R^2 = 0.97$ and 0.96).

```
summary(m7 <- lm(Weight ~ Length + Width + I(Width^2), data=perch))

## Estimate Std. Error t value Pr(>|t|)

## (Intercept) 36.348 59.832 0.607 0.546

## Length 29.176 3.363 8.674 1.11e-11 ***

## Width -271.668 38.130 -7.125 3.13e-09 ***

## I(Width^2) 30.128 2.688 11.210 1.74e-15 ***

## Residual standard error: 48.43 on 52 degrees of freedom

## Multiple R-squared: 0.9816, Adjusted R-squared: 0.9806

## F-statistic: 927 on 3 and 52 DF, p-value: < 2.2e-16</pre>
```

```
summary(m8 <- lm(Weight ~ Length+I(Length^2)+Width+I(Width^2), data=perch))
##
Estimate Std. Error t value Pr(>|t|)
```

```
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 138.0015 60.4435 2.283 0.026621 *

## Length -15.2436 12.4688 -1.223 0.227124

## I(Length^2) 0.6065 0.1652 3.672 0.000577 ***

## Width -31.0365 73.9416 -0.420 0.676436

## I(Width^2) 10.0718 5.9717 1.687 0.097793 .

##

## Residual standard error: 43.49 on 51 degrees of freedom

## Multiple R-squared: 0.9855, Adjusted R-squared: 0.9843

## F-statistic: 865.5 on 4 and 51 DF, p-value: < 2.2e-16
```

- From model 6 to 8, the added term $Width^2$ term is not/marginally significant. R^2 increases only a little bit. It is probably not necessary to include $Width^2$.
- From model 7 to 8, the added term *Length*² is significant. We probably will keep it.

Weight ~	F	R^2	R^2_{adj}
1. <i>L</i>	626.5	0.9207	0.9192
2. $L + L^2$	996.6	0.9741	0.9731
3. W	714.5	0.9297	0.9284
4. $W + W^2$	563.5	0.9551	0.9534
5.L + W	396.1	0.9373	0.9349
6. $L + L^2 + W$	1114	0.9847	0.9838
$7.L + W + W^2$	927	0.9816	0.9806
8. $L + L^2 + W + W^2$	865.5	0.9855	0.9843

F tests of all the eight models have $P < 2.2 \times 10^{-16}$.

- Model 8 has the largest R^2 and adjusted R^2 .
- ▶ Comparing model 6 and 8, the *Width*² is not significant suggesting that model 8 is not significantly better than model 6.
- Therefore, we choose model 6 as the best model among the 8 models.
- Note model 7 also has comparable assessment as model 6.

summary(m8)\$coefficents

```
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 138.0015 60.4435 2.283 0.026621 *
## Length -15.2436 12.4688 -1.223 0.227124
## I(Length<sup>2</sup>) 0.6065 0.1652 3.672 0.000577 ***
## Width -31.0365 73.9416 -0.420 0.676436
## I(Width<sup>2</sup>) 10.0718 5.9717 1.687 0.097793 .
anova(m6, m8)
## Analysis of Variance Table
##
## Model 1: Weight ~ Length + I(Length^2) + Width
## Model 2: Weight ~ Length + I(Length^2) + Width + I(Width^2)
##
    Res.Df RSS Df Sum of Sq F Pr(>F)
## 1 52 101863
## 2 51 96482 1 5381.3 2.8445 0.09779 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The interaction term $Length \times Width$ is highly significant. Keep it.

Residual standard error: 44.24 on 52 degrees of freedom ## Multiple R-squared: 0.9847, Adjusted R-squared: 0.9838

F-statistic: 1115 on 3 and 52 DF, p-value: < 2.2e-16

```
summary(m10 <- lm(Weight ~ Length * Width + I(Length^2), data=perch))

## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 136.1045 61.8038 2.202 0.0322 *
## Length -19.2196 14.2406 -1.350 0.1831
## Width -3.5967 83.3663 -0.043 0.9658
## I(Length^2) 0.4300 0.3795 1.133 0.2625
## Length:Width 2.6672 2.3089 1.155 0.2534
##
## Residual standard error: 44.12 on 51 degrees of freedom
## Multiple R-squared: 0.9851, Adjusted R-squared: 0.9839
## F-statistic: 840.9 on 4 and 51 DF, p-value: < 2.2e-16</pre>
```

The added quadratic term *Length*² is not significant when the interaction term is already in the model. We may not need it.

The added quadratic term $Width^2$ is not significant, either, when the interaction term is already in the model. We may not need it.

```
summary(m12 <- lm(Weight ~ Length*Width+I(Length^2)+I(Width^2), data=perch))</pre>
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 156.3486
                          61.4152
                                   2.546
                                           0.0140 *
## Length -25.0007 14.2729 -1.752 0.0860 .
        20.9772 82.5877 0.254 0.8005
## Width
## I(Length^2) 1.5719 0.7244 2.170 0.0348 *
## I(Width<sup>2</sup>) 34.4058 18.7455 1.835 0.0724 .
## Length: Width -9.7763 7.1455 -1.368
                                          0.1774
##
## Residual standard error: 43.13 on 50 degrees of freedom
## Multiple R-squared: 0.986, Adjusted R-squared: 0.9846
## F-statistic: 704.6 on 5 and 50 DF, p-value: < 2.2e-16
```

- With both quadratic terms (*Length*² and *Width*²) added, the interaction term becomes insignificant.
- The inferences of model 10, 11 and 12 suggest that we may only need one of the three terms: $Length^2$, $Width^2$ and $Length \times Width$

Weight ~	F	R^2	R^2_{adj}
5. $L + W$	396.1	0.9373	0.9349
6. $L + L^2 + W$	1114	0.9847	0.9838
$7.L + W + W^2$	927	0.9816	0.9806
8. $L + L^2 + W + W^2$	865.5	0.9855	0.9843
9.L + W + LW	1115	0.9847	0.9838
10. $L + L^2 + W + LW$	840.9	0.9851	0.9839
11. $L + W + W^2 + LW$	820	0.9847	0.9835
12. $L + L^2 + W + W^2 + LW$	704.6	0.9860	0.9846

F tests of all the eight models have $P < 2.2 \times 10^{-16}$.

- The complete secondorder model (model 12) has the largest R^2 and adjusted R^2 . Is it the best model?
- Model 10 and 11 are not significantly better than model 9.
- Model 9 has similar R²
 and adjusted R² as model
 6 and model 12. Let's
 compare them using
 nested F test.

```
anova(m6, m12)
## Analysis of Variance Table
##
## Model 1: Weight ~ Length + I(Length^2) + Width
## Model 2: Weight ~ Length * Width + I(Length^2) + I(Width^2)
    Res.Df RSS Df Sum of Sq F Pr(>F)
##
                                            ▶ Model 12 is NOT significantly
## 1 52 101863
## 2 50 93000 2 8863.1 2.3825 0.1027
                                              better than model 6.
anova (m9, m12)
                                            ▶ Model 12 is NOT significantly
## Analysis of Variance Table
##
                                              better than model 9.
## Model 1: Weight ~ Length * Width
## Model 2: Weight ~ Length * Width + I(Length^2) + I(Width^2)
##
    Res.Df
           RSS Df Sum of Sq F Pr(>F)
## 1
        52 101765
## 2 50 93000 2 8764.6 2.3561 0.1052
```

```
anova(m6, m9)

## Analysis of Variance Table

##

## Model 1: Weight ~ Length + I(Length^2) + Width

## Model 2: Weight ~ Length * Width

## Res.Df RSS Df Sum of Sq F Pr(>F)

## 1 52 101863

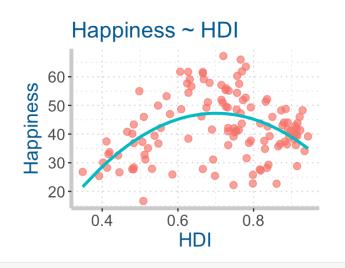
## 2 52 101765 0 98.527
```

- Nested *F* test cannot be used to compare two models with the same number of predictors.
- ▶ Model 6 and model 9 have very close R^2 and adjusted R^2 . Both models are good.
- This output gives the RSS (residual sum of squares, SSE) values of the two models. A model with smaller SSE explains more variablity in the response variable than the model with larger SSE.
- ▶ Therefore, model 9 is slightly-slightly better than model 6.

Happiness ~ HDI

```
summary(m.quan <- lm(Happiness ~ HDI, data=HappyPlanet))</pre>
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 34.834 4.617 7.545 9.01e-12 ***
## HDI 10.379 6.340 1.637 0.104
##
## Residual standard error: 10.98 on 122 degrees of freedom
## Multiple R-squared: 0.02149, Adjusted R-squared: 0.01347
## F-statistic: 2.679 on 1 and 122 DF, p-value: 0.1042
summary(m.cate <- lm(Happiness ~ HDI4, data=HappyPlanet))</pre>
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 34.315 1.862 18.429 < 2e-16 ***
## HDI4Medium 14.337 2.745 5.222 7.53e-07 ***
## HDI4High 12.838 2.478 5.180 9.05e-07 ***
## HDI4VeryHigh 5.146 2.422 2.124 0.0357 *
##
## Residual standard error: 9.675 on 120 degrees of freedom
## Multiple R-squared: 0.2522, Adjusted R-squared: 0.2335
## F-statistic: 13.49 on 3 and 120 DF, p-value: 1.213e-07
```

Happiness ~ HDI



$\widehat{Happiness} = -54.3 + 290.2 \times HDI - 207.2 \times HDI^2$

Happiness ~	F test	R^2	R^2_{adj}
HDI	F = 2.68, P = 0.10	0.0215	0.0135
HDI4	$F = 13.5, P = 1.21 \times 10^{-7}$	0.2522	0.2335
HDI+HDI^2	$F = 16.6, P = 4.37 \times 10^{-7}$	0.215	0.202

summary(m.poly <- lm(Happiness ~ HDI + I(HDI^2), data=HappyPlanet))</pre>

Summary

- ▶ Transforming the predictors:
 - Applying a function on the predictor
 - Categorizing a quantitative predictor
 - Treating a categorical predictor quantitative
 - Including the polynomial terms of a quantitative predictor
 - **...**
- Nested F test and adjusted R^2 are used together to select best model.
- Note: the model selecting process in this lecture did not consider model assumptions. In a full analysis, one should always check whether a model's assumptions are satisfied when selecting models.