



# STAT011 Statistical Methods I

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## Lecture 15 One-Sample $t$ Procedures

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# Review - Statistical inference

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By **CLT**,  $\bar{x} \overset{approx.}{\sim} N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$ .

- ▶ Level  $C$  **confidence interval** for population mean  $\mu$ :  $\bar{x} \pm z^* \frac{\sigma}{\sqrt{n}}$
- ▶ Level  $\alpha$   $z$  test for a population mean  $\mu$ :
  - $H_0 : \mu = \mu_0$ ;  $H_a : \mu > \mu_0$  or  $\mu < \mu_0$  or  $\mu \neq \mu_0$
  - $z = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \overset{approx.}{\sim} N(0, 1)$
  - $P$ -value is computed based on  $H_a$
  - $P \leq \alpha$ , reject  $H_0$ ;  $P > \alpha$ , fail to reject  $H_0$ .
- ▶ For both, we assume **unknown population mean  $\mu$**  and **known population standard deviation  $\sigma$** .
- ▶ What if  $\sigma$  is unknown?

# Outline

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- ▶ Sample standard deviation (SD)
- ▶ Degree of freedom
- ▶ Standard error (SE)
- ▶  $t$  distribution
- ▶ One-sample  $t$  procedures: statistical inference for a population mean based on  $t$  distribution
  - One-sample  $t$  confidence interval
  - One-sample  $t$  test
- ▶ Examples

# Sample standard deviation (SD)

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- ▶ When population standard deviation  $\sigma$  is unknown, we use the sample standard deviation  $s$  to estimate  $\sigma$ .

$$s = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n - 1}}$$

- ▶  $\sigma$  is a population parameter;  $s$  is a sample statistic.
- ▶ Why  $n - 1$ ?

# Sample standard deviation (SD)

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- ▶ Ultimately, we want  $s$  to be an **unbiased estimator** of  $\sigma$ .
- ▶ Let's use simulation to compare three possible ways of calculating sample standard deviation:

1. 
$$s_1 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n - 1}}$$

2. 
$$s_2 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n}}$$

3. 
$$s_3 = \sqrt{\frac{(x_1 - \mu)^2 + (x_2 - \mu)^2 + \cdots + (x_n - \mu)^2}{n}}$$

- ▶ **Note:**  $s_3$  is the formula for calculating SD when population mean  $\mu$  is known.

# Sample standard deviation (SD)

```
set.seed(10)
n <- 25; s1 <- s2 <- s3 <- NULL
for(i in 1:1000){
  x <- rnorm(n) # mu = 0, sigma = 1
  s1[i] <- sd(x) # sqrt(sum((x-mean(x))^2)/(n-1))
  s2[i] <- sqrt(sum((x-mean(x))^2)/n)
  s3[i] <- sqrt(sum((x-0)^2)/n) # mu=0
}
mean(s1)
```

```
## [1] 0.9957356
```

```
mean(s2)
```

```
## [1] 0.9756177
```

```
mean(s3)
```

```
## [1] 0.9965334
```

# Sample standard deviation (SD)

---

1. 
$$s_1 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n - 1}}$$

2. 
$$s_2 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n}}$$

3. 
$$s_3 = \sqrt{\frac{(x_1 - \mu)^2 + (x_2 - \mu)^2 + \cdots + (x_n - \mu)^2}{n}}$$

- ▶ By simulation, mean of  $s_3$  is the closest to the true population SD  $\sigma = 1$ .
- ▶ In reality, since we do not know population mean  $\mu$ , we cannot apply the formula  $s_3$ . We use sample mean  $\bar{x}$ , which is an unbiased estimator of  $\mu$ , to compute the sample SD.
- ▶ Using  $\bar{x}$  brings more uncertainty ( $\mu$  is fixed and  $\bar{x}$  changes from sample to sample) into the estimation.  $s_2$  turns out to be a biased estimator of  $\sigma$ .

# Sample standard deviation (SD)

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$$1. \quad s_1 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n - 1}}$$

$$2. \quad s_2 = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n}}$$

$$3. \quad s_3 = \sqrt{\frac{(x_1 - \mu)^2 + (x_2 - \mu)^2 + \cdots + (x_n - \mu)^2}{n}}$$

- ▶ SD measures the variability of  $n$  **random values**  $x_1, x_2, \dots, x_n$ . However, once  $\bar{x}$  is used ( $s_2$ ), knowing  $x_1, x_2, \dots, x_{n-1}$  and  $\bar{x}$ , we will know  $x_n$  for sure. It measures the variability of only  $n - 1$  **random values** that are free to vary.
- ▶ Therefore, in the formula of sample standard deviation ( $s_1$ ), the denominator is  $n - 1$ , which results in an unbiased estimator of  $\sigma$ .



# Degree of freedom

**Degree of freedom** is the number of values in the final calculation of a statistic that are **free to vary**.

- ▶ It is calculated as the **difference between**
  - Number of independent values that go into the estimate:  $n$
  - Number of statistics used as intermediate steps: **1**
- ▶ For

$$s = \sqrt{\frac{(x_1 - \bar{x})^2 + (x_2 - \bar{x})^2 + \cdots + (x_n - \bar{x})^2}{n - 1}}$$

The degree of freedom for sample SD  $s$  is  $n - 1$ .

# Standard error (SE)

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By CLT,

$$\bar{x} \overset{\text{approx}}{\sim} N \left( \mu, \frac{\sigma}{\sqrt{n}} \right)$$

- ▶ When population SD  $\sigma$  is unknown, we use sample SD  $s$  to replace it.
- ▶ The SD of  $\bar{x}$  becomes

$$\frac{s}{\sqrt{n}}$$

- ▶ This is called the **standard error (SE)** of  $\bar{x}$ .

# Standard error (SE)

When the standard deviation of a **statistic** is **estimated from the data**, the result is called the **standard error (SE)** of the statistic.

- ▶ Population **SD** of a variable:  $\sigma$
- ▶ Sample **SD** of a variable:  $s$
- ▶ **SD** of  $\bar{x}$  (when  $\sigma$  is known):

$$\frac{\sigma}{\sqrt{n}}$$

- ▶ **SE** of  $\bar{x}$  (when  $\sigma$  is unknown):

$$\frac{s}{\sqrt{n}}$$

# Distribution of sample mean

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When  $\sigma$  is known,

$$\bar{x} \stackrel{\text{approx}}{\sim} N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$

And by standardization of Normal distribution,

$$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}} \stackrel{\text{approx}}{\sim} N(0, 1)$$

When  $\sigma$  is unknown and estimated by  $s$ ,

$$t = \frac{\bar{x} - \mu}{s / \sqrt{n}} \stackrel{\text{approx}}{\sim} t(n - 1)$$

# The $t$ distribution

Suppose that an SRS of size  $n$  is drawn from an  $N(\mu, \sigma)$  population. Then the one-sample  $t$  statistic

$$t = \frac{\bar{x} - \mu}{s/\sqrt{n}} \sim t(n - 1)$$

has the  **$t$  distribution** with  **$n - 1$  degrees of freedom**.

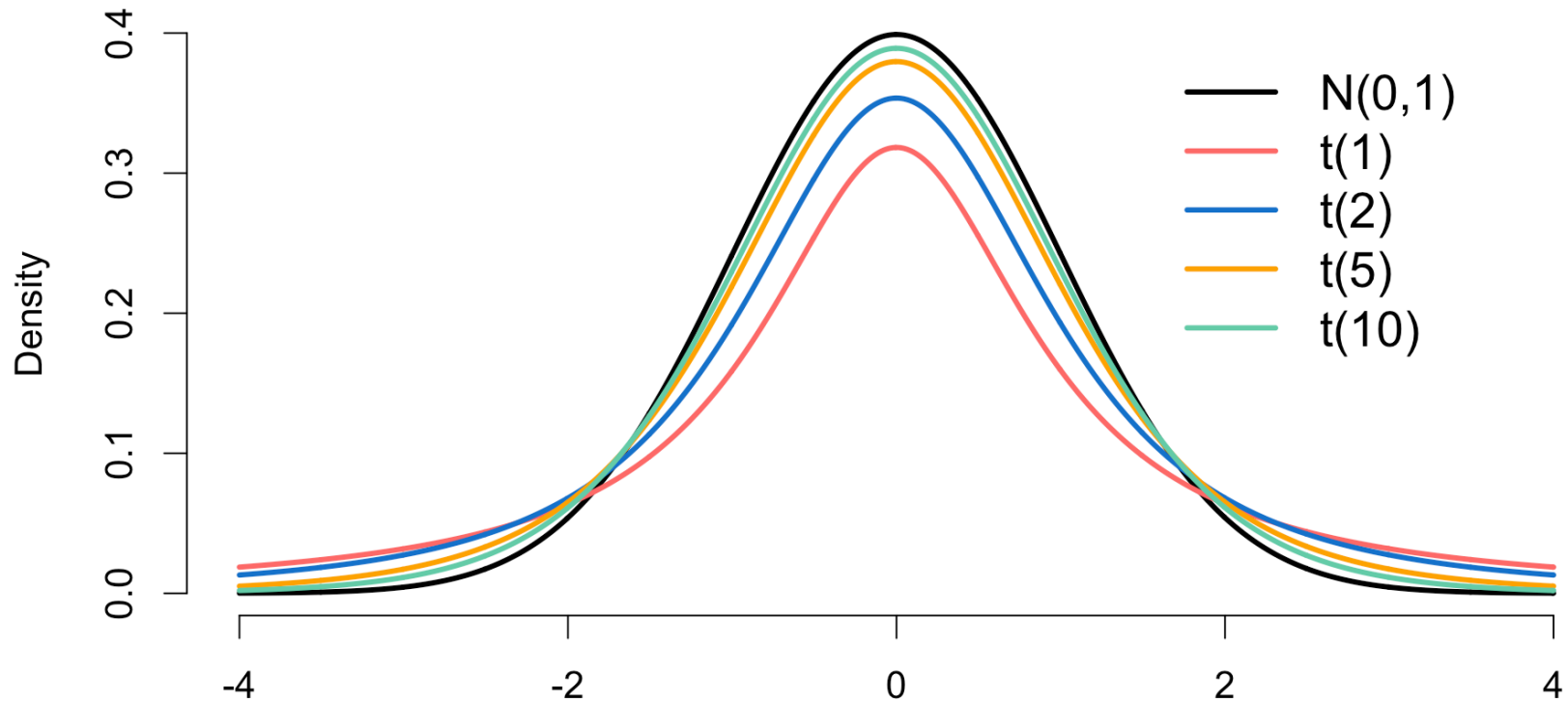
When the population distribution is not Normal,

$$t = \frac{\bar{x} - \mu}{s/\sqrt{n}} \overset{\text{approx}}{\sim} t(n - 1)$$

has an approximate  **$t$  distribution** with  **$n - 1$  degrees of freedom**.

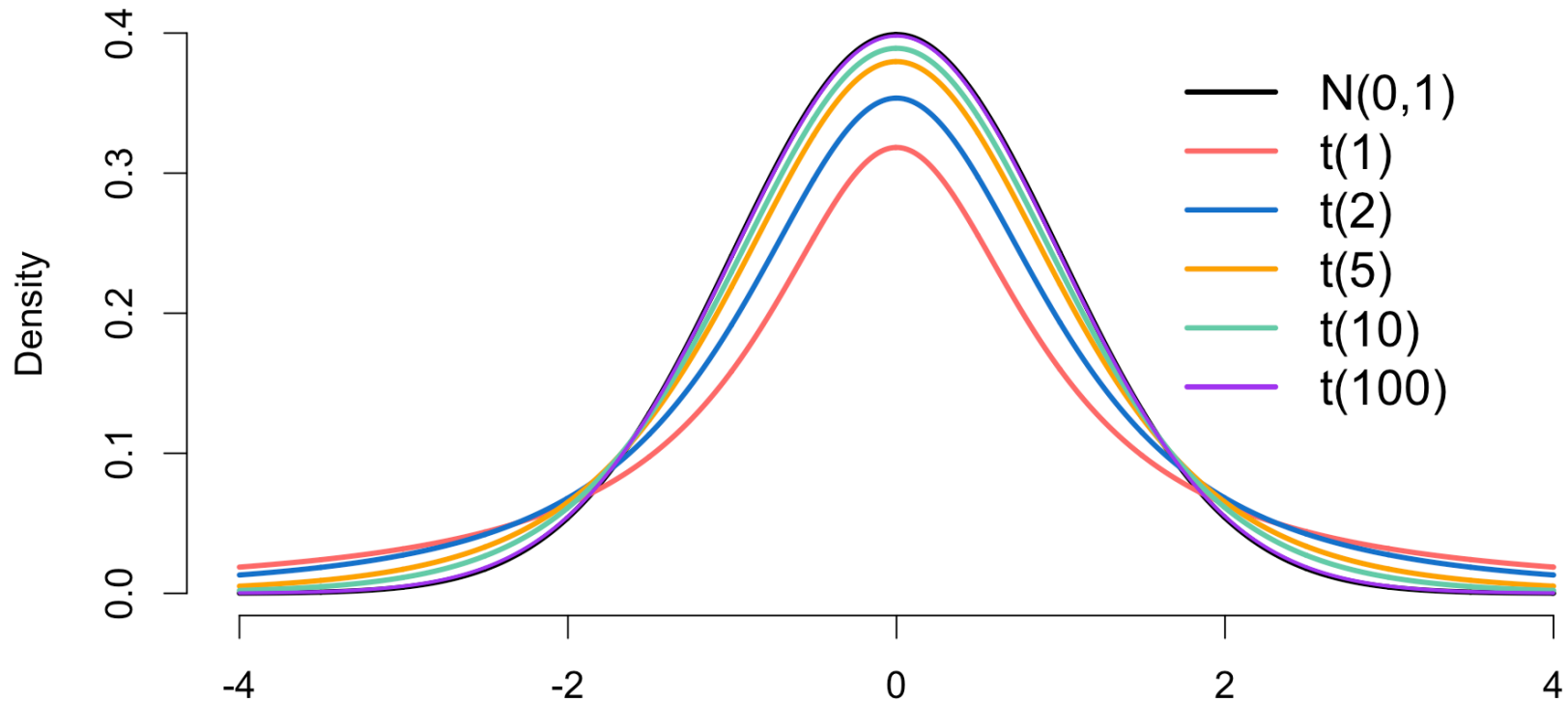
# The $t$ distribution

Density Curves of Normal and  $t$  Distributions



# The $t$ distribution

Density Curves of Normal and  $t$  Distributions



# The $t$ distribution

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- ▶ Symmetric, unimodal, bell-shaped.
- ▶ Approximates the Normal distribution when  $n$  is large.
- ▶ Has heavier tails than the Normal distribution
  - Using  $s$  instead of  $\sigma$  introduces more variability to  $\frac{\bar{x}-\mu}{s/\sqrt{n}}$
  - Using  $t$  distribution results in wider C.I. and larger  $P$ -value than Normal distribution.
  - We are less sure about the inference of population mean when population SD is unknown.



# $t$ distribution in R

```
# dnorm() and dt( , df = n-1)  
dnorm(0); dt(0, df=5); dt(0, df=100)
```

```
## [1] 0.3989423
```

```
## [1] 0.3796067
```

```
## [1] 0.3979462
```

```
# pnorm() and pt( , df = n-1)  
pnorm(0); pt(0, df=5); pt(0, df=100)
```

```
## [1] 0.5
```

```
## [1] 0.5
```

```
## [1] 0.5
```

# $t$ distribution in R

```
# pnorm() and pt( , df = n-1)
```

```
pnorm(-1.96); pt(-1.96, df=5); pt(-1.96, df=100)
```

```
## [1] 0.0249979
```

```
## [1] 0.05364398
```

```
## [1] 0.02638945
```

```
# qnorm() and qt( , df = n-1)
```

```
qnorm(0.975); qt(0.975, df=5); qt(0.975, df=100)
```

```
## [1] 1.959964
```

```
## [1] 2.570582
```

```
## [1] 1.983972
```

# One-sample $t$ confidence interval

Suppose that an SRS of size  $n$  is drawn from a population having unknown mean  $\mu$ .  
A **level  $C$  confidence interval** for  $\mu$  is

$$\bar{x} \pm t^* \frac{s}{\sqrt{n}}$$

where  $t^*$  is the value for the  $t(n - 1)$  density curve with area  $C$  between  $-t^*$  and  $t^*$ .  
The quantity

$$t^* \frac{s}{\sqrt{n}}$$

is the **margin of error**. The confidence level is exactly  $C$  when the population distribution is Normal and is approximately correct for large  $n$  in other cases.

# One-sample $t$ test

Suppose that an SRS of size  $n$  is drawn from a population having unknown mean  $\mu$ . To test the hypothesis  $H_0 : \mu = \mu_0$ , compute the **one-sample  $t$  statistic**

$$t = \frac{\bar{x} - \mu_0}{s / \sqrt{n}}$$

In terms of a random variable  $T$  having the  $t(n - 1)$  distribution, the  $P$ -value for a test of  $H_0$  against

$$H_a : \mu > \mu_0 \text{ is } P(T \geq t)$$

$$H_a : \mu < \mu_0 \text{ is } P(T \leq t)$$

$$H_a : \mu \neq \mu_0 \text{ is } 2P(T \geq |t|)$$

These  $P$ -values are exact if the population distribution is Normal and are approximately correct for large  $n$  in other cases.

# Guidelines for one-sample $t$ procedures

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For sample size  $n$ ,

- ▶  $n < 15$ : Use  $t$  procedures if the data are close to Normal. If the data are clearly non-Normal or if outliers are present, do not use  $t$ .
- ▶  $15 \leq n < 40$ : The  $t$  procedures can be used except in the presence of outliers or strong skewness.
- ▶  $n \geq 40$ : The  $t$  procedures can be used even for clearly skewed distributions when the sample is large.

The  $t$  procedures are quite **robust**.































- ▶ A statistical inference procedure is called **robust** if it is insensitive to violations of the assumptions made.

# Comparing $z$ and $t$ procedures

	$z$ procedures	$t$ procedures
<b>Population SD <math>\sigma</math></b>	Known	Unknown, use sample SD $s$
<b>Level <math>C</math> C.I.</b>	$\bar{x} \pm z^* \frac{\sigma}{\sqrt{n}}$ $z^* = \text{qnorm}(1 - (1 - C) / 2)$	$\bar{x} \pm t^* \frac{s}{\sqrt{n}}$ $t^* = \text{qt}(1 - (1 - C) / 2, \text{df} = n - 1)$
<b>Level <math>\alpha</math> significance test</b>	$H_0 : \mu = \mu_0$ $H_a : \mu > \mu_0, \mu < \mu_0, \mu \neq \mu_0$ $z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}} \overset{\text{approx.}}{\sim} N(0, 1)$ $P(Z \leq z), \text{pnorm}(z)$ $P(Z \geq z), 1 - \text{pnorm}(z)$ $2P(Z \geq  z ), 2 * (1 - \text{pnorm}(\text{abs}(z)))$	$H_0 : \mu = \mu_0$ $H_a : \mu > \mu_0, \mu < \mu_0, \mu \neq \mu_0$ $t = \frac{\bar{x} - \mu_0}{s / \sqrt{n}} \overset{\text{approx.}}{\sim} t(n - 1)$ $P(T \leq t), \text{pt}(t, \text{df} = n - 1)$ $P(T \geq t), 1 - \text{pt}(t, \text{df} = n - 1)$ $2P(T \geq  t ), 2 * (1 - \text{pt}(\text{abs}(t), \text{df} = n - 1))$

# Example 1

Within-platform score of mis-communication (25 emoji for each platform)

	Apple	Google	Microsoft	Samsung	LG
Top 3	 3.64	 3.26	 4.40	 3.69	 2.59
	 3.50	 2.66	 2.94	 2.36	 2.53
	 2.72	 2.61	 2.35	 2.29	 2.51
...	...				
Bottom 3	 1.25	 1.13	 1.12	 1.23	 1.30
	 0.65	 1.06	 1.08	 1.09	 1.26
	 0.45	 0.62	 0.66	 1.08	 0.63

Google, MS, Samsung and LG together

- ▶ Average score of mis-communication: 1.84
- ▶ Number of emoji's: 100
- ▶ Population standard deviation: 0.50
- In fact, this is sample SD.

# Example 1

Assume population SD is known.

$\bar{x} = 1.84$ ,  $\sigma = 0.5$ ,  $n = 100$ ,  $C = 0.95$

95% confidence interval

▶  $C = 0.95$ ,  $z^* = 1.96$  `qnorm(0.975)`

▶ Margin of error

$$m = z^* \frac{\sigma}{\sqrt{n}} = 1.96 \frac{0.5}{\sqrt{100}} = 0.098$$

▶ 95% confidence interval  $\bar{x} \pm m = 1.84 \pm 0.098$

▶ We are 95% confident (about the method) that the population mean score of mis-communication for the four platforms will be within [1.742, 1.938]





# Example 1

Population SD is in fact unknown.

$\bar{x} = 1.84$ ,  $s = 0.5$ ,  $n = 100$ ,  $C = 0.95$

95% confidence interval

▶  $C = 0.95$ ,  $t^* = 1.98$  `qt(0.975, df = 99)`

▶ Margin of error

$$m = t^* \frac{s}{\sqrt{n}} = 1.98 \frac{0.5}{\sqrt{100}} = 0.099$$

▶ 95% confidence interval  $\bar{x} \pm m = 1.84 \pm 0.099$

▶ We are 95% confident (about the method) that the population mean score of mis-communication for the four platforms will be within [1.741, 1.939]



# Example 1

Is the average score of mis-communication of the four from 2, which is the mean score of Apple emoji?

Assume population SD is known.



$\bar{x} = 1.84$ ,  $\sigma = 0.5$ ,  $n = 100$ ,  $\alpha = 0.05$

VS.



►  $H_0 : \mu = 2; H_a : \mu \neq 2$

► 
$$z = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} = \frac{1.84 - 2}{0.5/\sqrt{100}} = \frac{-0.12}{0.05} = -3.2$$

►  $2P(Z \geq |z|) = 2P(Z \geq 3.2) = 0.0014 < 0.05$  `2*(1-pnorm(3.2))`

► The test is significant at level 0.05 and we reject  $H_0$ . The mean score of mis-communication of the four platforms is significantly different from 2.

# Example 1

Is the average score of mis-communication of the four from 2, which is the mean score of Apple emoji?

Population SD is in fact unknown.



$\bar{x} = 1.84, s = 0.5, n = 100, \alpha = 0.05$

VS.



▶  $H_0 : \mu = 2; H_a : \mu \neq 2$

▶ 
$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} = \frac{1.84 - 2}{0.5/\sqrt{100}} = \frac{-0.12}{0.05} = -3.2$$

▶  $2P(T \geq |t|) = 2P(T \geq 3.2) = 0.0018 < 0.05$  `2*(1-pt(3.2, df=99))`

▶ The test is significant at level 0.05 and we reject  $H_0$ . The mean score of mis-communication of the four platforms is significantly different from 2.

## Example 2

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The mean percentage of dialogue spoken by men for the 62 screenplays in 2015 is 0.668. The SD of the 62 screenplays is 0.241.

### 95% confidence interval

$$\bar{x} \pm t^* \frac{s}{\sqrt{n}} = 0.668 \pm 2.00 \times \frac{0.241}{\sqrt{62}} = 0.668 \pm 0.061$$

$$t^* = \text{qt}(0.975, \text{df} = 61) = 1.999624$$

We are 95% confident (about the method) that the population mean percentage of dialogue spoken by men is within [0.607, 0.729]

### Level 0.05 significance test whether population mean greater than 0.5

$$\text{▶ } H_0 : \mu = 0.5, H_a : \mu > 0.5. t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} = \frac{0.668 - 0.5}{0.241/\sqrt{62}} = 5.49.$$

$$t > t^* = 1.7 \text{ qt}(0.95, \text{df}=61) \text{ or } P = 4 \times 10^{-7} < 0.05 \text{ 1-pt}(5.49, \text{df}=61)$$

The test is highly significant at level 0.05. We reject  $H_0$  and conclude that the population mean percentage of dialogue spoken by men is significantly greater than 0.5.

# One-sample $t$ procedures in R

`percent_men` # 62 percentage values of dialogue spoken by men for 2015 movies

```
## [1] 0.98457660 0.97053407 0.32418830 0.27857449 0.77425697 0.84956568
## [7] 0.53103976 0.10586256 0.50441158 0.25436772 0.35793946 0.73917869
## [13] 0.83809736 0.15171331 0.89037260 0.38880671 1.00000000 0.76046885
## [19] 0.39227316 0.28032892 0.91113709 0.54406303 0.84768212 1.00000000
## [25] 0.78186381 0.62622438 0.56980907 0.78833910 0.72210815 0.79612088
## [31] 0.82716454 0.64336662 0.89454643 0.80753437 0.50219759 0.63798364
## [37] 0.25218825 0.74342258 0.79141282 0.93589744 0.80880134 0.68960030
## [43] 0.81827042 0.89763325 0.46960452 0.59691068 0.80664427 0.49614112
## [49] 0.53015726 0.83555121 0.93586918 0.76235198 0.72157216 0.45707300
## [55] 0.65388303 0.80549821 1.00000000 0.90663453 0.74396939 0.71404924
## [61] 0.71968288 0.03445006
```

# One-sample $t$ procedures in R

```
t.test(percent_men, conf.level = 0.95) # 95% confidence interval
```

```
##  
## One Sample t-test  
##  
## data: percent_men  
## t = 21.841, df = 61, p-value < 2.2e-16  
## alternative hypothesis: true mean is not equal to 0  
## 95 percent confidence interval:  
## 0.6066667 0.7289451  
## sample estimates:  
## mean of x  
## 0.6678059
```

- ▶ The 95% confidence interval for the true mean percentage of dialogue spoken by men in 2015 movies is [0.607, 0.729].
- ▶ Here the `t.test()` function automatically runs a two-sided test, but we are interested in a one-side test.

# One-sample $t$ procedures in R

```
t.test(percent_men, alternative = "greater", mu = 0.5)
```

```
##                                ## alternative = "greater", "less" or "two.sided"
## One Sample t-test
##
## data:  percent_men
## t = 5.4883, df = 61, p-value = 4.15e-07
## alternative hypothesis: true mean is greater than 0.5
## 95 percent confidence interval:
##  0.6167384      Inf
## sample estimates:
## mean of x
## 0.6678059
```

- ▶  $H_0 : \mu = 0.5; H_a : \mu > 0.5; t = 5.49; P = 4.15 \times 10^{-7} < 0.05$ . We reject  $H_0$  at level 0.05. Men speak significantly more dialogue than women in 2015 movies.
- ▶ When `t.test()` function is run for a one-sided test, it generates a "one-sided" confidence interval at the same time, which is NOT the correct confidence interval - so ignore it.

# One-sample $t$ procedures in R

```
t.test(percent_men, alternative = "two.sided", mu = 0.5)
```

```
##                                ## alternative = "greater", "less" or "two.sided"
## One Sample t-test
##
## data:  percent_men
## t = 5.4883, df = 61, p-value = 8.299e-07
## alternative hypothesis: true mean is not equal to 0.5
## 95 percent confidence interval:
##  0.6066667 0.7289451
## sample estimates:
## mean of x
## 0.6678059
```

- ▶  $H_0 : \mu = 0.5; H_a : \mu > 0.5; t = 5.49; P = 4.15 \times 10^{-7} < 0.05$ . We reject  $H_0$  at level 0.05. Men speak significantly more dialogue than women in 2015 movies. The 95% confidence interval for  $\mu$  is [0.607, 0.729].
- ▶ When `t.test()` function is run for a two-sided test, it gives the results for the test as well as the confidence interval.



# About homework

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- ▶ Some questions may ask you to calculate the confidence interval and conduct a  $t$  test "**using R**". You should use the `t.test()` function to do the analysis and write down the four steps of the test and report the CI as in Slide 30~32.
- ▶ Some other questions may ask you to do the analysis "**by hand**". Then you should apply the formulas in the definitions of the confidence interval and the test and write everything down in math mode. You may still use R as a calculator and to compute  $t^*$  values and  $P$ -values.
- ▶ This guidance applies to all the subsequent problem sets (Homework 6 to 10).

# Summary

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- ▶ Sample standard deviation (SD)
- ▶ Degree of freedom
- ▶ Standard error (SE)
  - *SD of a statistic estimated from sample data*
- ▶  $t$  distribution `dt( , df = )`, `pt( , df = )`, `qt( , df = )`
- ▶ Statistical inference for a population mean based on  $t$  distribution
  - One-sample  $t$  confidence interval  $\bar{x} \pm t^* \frac{s}{\sqrt{n}}$
  - One-sample  $t$  test  $H_0 : \mu = \mu_0, t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \overset{approx.}{\sim} t(n - 1)$
- ▶ Examples  
`t.test( , conf.level = )`, `t.test( , alternative = , mu = )`