

Lecture 15: Examples of Queueing Networks

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1 Network of Queues Contd.

The last lecture ended with a conjecture on the network of queues. First we state the following theorem.

Theorem 1.1. *Assuming that $\lambda_i < \mu_i$, for all i , in steady state, the number of customers at service i are independent and the limiting probabilities are given by*

$$Pr(n_1, n_2, \dots, n_k) = \prod_{i=1}^k \left(\frac{\lambda_i}{\mu_i} \right)^{n_i} \left(1 - \frac{\lambda_i}{\mu_i} \right).$$

Also, from the reversed chain, we have the following.

Corollary 1.2. *The process of customers departing the system from the server i , $i = 1, 2, \dots, k$, are independent Poisson processes having respective rates $\lambda_i(1 - \sum_j P_{ij})$.*

Proof. We have already shown that in the reverse process, customers arrive to server i from outside the system according to independent Poisson processes having rates $\lambda_i(1 - \sum_j P_{ij})$, $i \geq 1$. Since an arrival from outside corresponds to a departure out of the system from server i in the forward process, the result follows. \square

1.1 The Erlang Loss Formula

consider a queueing system in which there are k servers and customers arrive according to a Poisson process with rate λ . If an arriving customer finds all the k servers busy, the customer is lost (does not enter the system). The service times of server are assumed to be distributed according to some general distribution G . Assume that G has a density g . Let $\lambda(t)$ denote the hazard rate function. That is,

$$\lambda(t) = \frac{g(t)}{\bar{G}(t)}$$

is the instantaneous probability density that a t unit old service will end. Assume that the states are ordered. i.e. $\underline{x} = (x_1, x_2, \dots, x_n)$, $x_1 \leq x_2 \leq \dots \leq x_n$, where x_n denote the service time of n^{th} customer ($n \leq k$). The process of successive states will be a Markov process in the sense that the conditional distribution of any future state, given the present state and all the past states, will depend only on the present. Even though the process is not a continuous-time Markov chain, we can extend and use the theory so far developed to analyse the process.

Corollary 1.3. *The reverse process is also a k -server loss system with service distribution G in which arrivals occur according to a Poisson process with rate λ . The state at any time represents the ordered residual service times of customers in service currently.*

Proof. We shall prove the above conjecture and obtain the limiting distribution. For any state $\underline{x} = (x_1, x_2 \dots x_i, \dots x_n)$ and $e_i(\underline{x}) = (x_1, x_2 \dots x_{i-1}, x_{i+1} \dots x_n)$. In the original process when the state is \underline{x} it will instantaneously go to $e_i(\underline{x})$ with a probability density equal to $\lambda(x_i)$. Similarly, in the reversed process, we see that if the state is $e_i(\underline{x})$, then it will instantaneously go to \underline{x} if a customer having service time x_i instantaneously arrives. So,

Forward : $\underline{x} \rightarrow e_i(\underline{x})$ w.p. intensity $\lambda(x_i)$;

Reverse : $e_i(\underline{x}) \rightarrow \underline{x}$ with joint prob. intensity $\lambda g(x_i)$.

Hence if $p(\underline{x})$ represents the limiting density, in accordance with Theorem 1.5 (Lecture 14), we would need that

$$p(\underline{x})\lambda(x_i) = P(e_i(\underline{x}))\lambda g(x_i),$$

or, since $\lambda(x_i) = g(x_i)/\bar{G}(x_i)$,

$$p(\underline{x}) = p(e_i(\underline{x}))\lambda G(x_i).$$

Letting $i = 1$ and iterating the above yields,

$$\begin{aligned} p(\underline{x}) &= \lambda G(x_1)p(e_1(\underline{x})) \\ &= \lambda G(x_1)\lambda G(x_2)p(e_1(e_1(\underline{x}))) \\ &\vdots \\ &= \Pi_{i=1}^n \lambda \bar{G}(x_i)P(\phi), \end{aligned}$$

where $P(\phi)$ is the limiting probability that the system is empty. Integrating over vector \underline{x} yields

$$\begin{aligned} Pr(n \text{ in the system}) &= P(\phi)\lambda^n \int \dots \int_{x_1 \leq x_2 \dots x_n} \Pi_{i=1}^n \bar{G}(x_i) dx_1 dx_2 \dots dx_n \\ &= P(\phi) \frac{\lambda^n}{n!} \int \dots \int_{x_1, x_2, \dots x_n} \Pi_{i=1}^n \bar{G}(x_i) dx_1 dx_2 \dots dx_n \\ &= P(\phi) \frac{(\lambda E[S])^n}{n!}, n=1, 2 \dots k, \end{aligned}$$

where $E[S] = \int \bar{G}(x)dx$ is the mean service time. Upon using the fact that

$$P(\phi) + \sum_{n=1}^k Pr(n \text{ in the system}) = 1.$$

we obtain

$$Pr(n \text{ in the system}) = \frac{(\lambda E[S])^n / (n!)}{\sum_{i=0}^n (\lambda E[S])^i / (i!)}. \quad (1)$$

This can be written as

$$Pr(\underline{x}) = \frac{\lambda^n \prod_{i=1}^n \bar{G}(x_i)}{\sum_{i=0}^n (\lambda E[S])^i / (i!)}.$$

Observe that the conditional distribution of the ordered ages given that there are n customers in the system is

$$Pr(\underline{x} | n \text{ in the system}) = \frac{p(\underline{x})}{Pr(n \text{ in the system})} = \frac{n! \prod_{i=1}^n \frac{\bar{G}(x_i)}{E[S]}}{(\lambda E[S])^n / n!}.$$

As $\bar{G}(x)/E[S]$ is just the density of the equilibrium distribution of G , if the conjecture is valid, the limiting distribution of the number of customers in the system depends on G only through its mean and given that there are n customers in the system the ages are independent and identically distributed according to the equilibrium distribution of G . To complete the proof of the conjecture, we must consider the transitions of the forward process from \underline{x} to $(0, \underline{x})$ when $n < k$. Now

Forward: $\underline{x} \rightarrow (0, \underline{x})$ with instantaneous density λ ;
Reverse: $(0, \underline{x}) \rightarrow \underline{x}$ with probability 1.

Hence in conjunction with Theorem 1.5 (Lecture 14) we must verify that

$$p(\underline{x})\lambda = p(0, \underline{x}),$$

which follows from refEquilibriumDistribution since $\bar{G}(0) = 1$. □

We have thus proven the following:

Theorem 1.4. *The limiting distribution of the number of customers in the system is given by*

$$Pr(n \text{ in the system}) = Pr(n \text{ in the system}) = \frac{(\lambda E[S])^n / (n!)}{\sum_{i=0}^n (\lambda E[S])^i / (i!)} \quad (2)$$

and given that there are n in the system the ages (or the residual times) of these n are independent and identically distributed according to the equilibrium distribution of G .

The model considered is often called the Erlang loss system and 1 is called is called the Erlang loss formula. By using the reversed process, we also have the following corollary.

Corollary 1.5. *In Erlang loss model the departure process (including both customers completing service and those that are lost) is a Poisson process at rate λ .*

Proof. The corollary follows as in the reversed process arrivals of all customers (including that are lost) constitutes a Poisson process. □