

Lecture 7: Limiting Mean Excess Time, Branching Processes, Delayed Renewal Process

Parimal Parag

1 Renewal thmry Contd. - Key Renewal thm-rem and Applications

1.1 Example:

Consider the number of commodities desired by customers at a store follows a distribution G . The ordering policy of the store is as follows: For some fixed s, S , if the inventory level after serving a customer is x , then the amount ordered is

$$\begin{cases} S - x & \text{if } x < s \\ 0 & \text{if } x \geq s \end{cases}$$

$X(t)$ denote the inventory level at time t . We are interested in finding $\lim_{t \rightarrow \infty} \mathbb{P}(X(t) \geq x)$. From alternating renewal process thmry, we have

$$\begin{aligned} \lim_{t \rightarrow \infty} \mathbb{P}(X(t) \geq x) &= \mathbb{E}[\text{ON time}] \\ &= \frac{\mathbb{E}[\sum_{i=1}^{N_x} X_i]}{\mathbb{E}[\sum_{i=1}^{N_s} X_i]} = \frac{\mathbb{E}[N_x]}{\mathbb{E}[N_s]}. \end{aligned}$$

where $N_x = \min\{n \in \mathbb{N} : \sum_{i=1}^n Y_i > s - x\}$ and Y_1, Y_2, \dots denote the successive customer demands. Since Y_i s are iid, we can interpret $N_x - 1$ as the number of renewals till time $S - x$. Y_i s are the inter arrival times of the processes. Thus

$$\lim_{t \rightarrow \infty} \mathbb{P}(X(t) \geq x) = \frac{m_G(S - x) + 1}{m_G(S - s) + 1}, s \leq x \leq S.$$

1.2 Limiting Mean Excess Time

Consider a nonlattice renewal process and we are interested in computing the mean excess time of the process. We start by writing the renewal equation of mean excess life time, $\mathbb{E}[Y(t)]$.

$$\begin{aligned} \mathbb{E}[Y(t)] &= \mathbb{E}[Y(t) | S_{N(t)} = 0] F^c(t) + \int_0^t \mathbb{E}[Y(t) | S_{N(t)} = y] F^c(t - y) dm(y) \\ &= \mathbb{E}[X - t | X > t] F^c(t) + \int_0^t \mathbb{E}[X - (t - y) | X > t - y] F^c(t - y) dm(y). \end{aligned}$$

From Key Renewal thmrem, we have

$$\begin{aligned}
\lim_{t \rightarrow \infty} \mathbb{E}[Y(t)] &= \frac{1}{\mu} \int_0^\infty \mathbb{E}[X - t | X - t > 0] F^c(t) dt \\
&= \frac{1}{\mu} \int_{t=0}^\infty \int_{x=t}^\infty x dF(x) F^c(t) dt \\
&= \frac{1}{\mu} \int_{x=0}^\infty \int_{t=0}^x x dF(x) F^c(t) dt \\
&= \frac{\mathbb{E}[X^2]}{2\mu}.
\end{aligned}$$

Proposition 1.1. *If the inter arrival time is nonlattice and $\mathbb{E}[X^2] < \infty$, by corollary , we have $\mu(m(t) + 1) = t + \mathbb{E}[Y(t)]$*

$$\lim_{t \rightarrow \infty} (m(t) - \frac{t}{\mu}) = \frac{\mathbb{E}[X^2]}{2\mu^2} - 1.$$

1.3 Age-dependent Branching Process

Suppose an organism lives upto a time period of $X \sim F$ and produces $N \sim P$ number of offspring. Let $X(t)$ denote the number of organisms alive at time t . The stochastic process $\{X(t), t \geq 0\}$ is called an age-dependent branching process. We are interested in computing $M(t) = \mathbb{E}[X(t)]$ when $m = \mathbb{E}[N] = \sum_{j \in \mathbb{N}} j P_j$.

Theorem 1.2. *If $X(0) = 1$, $m > 1$ and F is non lattice, then*

$$\lim_{t \rightarrow \infty} M(t) = \frac{m - 1}{m^2 \alpha \int_0^\infty x e^{-\alpha x} dF(x)},$$

where $\alpha > 0$ is unique such that $\int_0^\infty x e^{-\alpha x} dF(x) = \frac{1}{m}$.

Proof. Condition on T_1 , the life time of first organism,

$$\begin{aligned}
M(t) &= \int_0^\infty \mathbb{E}[X(t) | T_1 = y] dF(y) \\
&\stackrel{(a)}{=} \int_{y=0}^t 1 dF(y) + \int_{y=t}^\infty m M(t - y) dF(y).
\end{aligned}$$

Thus we get

$$M(t) = F^c(t) + m \int_0^t M(t - y) dF(y) \quad (1)$$

Let α denote the unique positive number such that $\int_0^\infty x e^{-\alpha x} dF(x) = \frac{1}{m}$ and $G(y) = m \int_0^y e^{-\alpha y} dF(y)$. Upon multiplying both sides of equation (1) by $e^{-\alpha t}$ and defining $f(t) = e^{-\alpha t} M(t)$, $h(t) = e^{-\alpha t} F^c(t)$,

$$\begin{aligned}
f &= h + f * G \\
&= h + G * (h + f * G) \\
&\vdots = h + h * \sum_{i=1}^{\infty} G_i \\
&= h + h * m_G.
\end{aligned}$$

Or, $f(t) = h(t) + \int_0^t h(t-s)dm_G(s)$. It can be shown that $h(t)$ is dRi and hence by Key Renewal thmrem,

$$f(t) \rightarrow \frac{\int_0^{\infty} e^{-\alpha t} F^c(t) dt}{\int_0^{\infty} x dG(x)}.$$

$$\begin{aligned}
\int_0^{\infty} e^{-\alpha t} F^c(t) dt &= \int_0^{\infty} e^{-\alpha t} \int_t^{\infty} dF(x) dt \\
&= \int_0^{\infty} \int_0^x e^{-\alpha t} dt dF(x) \\
&= \int_0^{\infty} (1 - e^{-\alpha x}) dF(x) \\
&= \frac{1}{\alpha} (1 - \frac{1}{m}) \quad (\text{by the definition of } \alpha).
\end{aligned}$$

Also $\int_0^{\infty} x dG(x) = m \int_0^{\infty} x e^{-\alpha x} dF(x)$. Hence the result follows. \square

1.4 Delayed Renewal Process

Let $\{X_n : n \in \mathbb{N}\}$ be independent but $X_1 \sim G$ and $X_i \sim F$, $i \geq 2$ then the counting process $\{N_D(t) : t \geq 0\}$ is called general renewal process or delayed renewal process. Let $S_0 = 0$ and $S_n = \sum_{i=1}^n X_i$. We have

$$\begin{aligned}
N_D(t) &= \sup\{n \in \mathbb{N} : S_n \leq t\}, \\
P(N_D(t) = n) &= P(S_n \leq t) - P(S_{n+1} \leq t) \\
&= G * F^{n-1}(t) - G * F^n(t), \\
m_D(t) &= \mathbb{E}[N_D(t)] = \sum_{n \in \mathbb{N}} G * F^{n-1}(t).
\end{aligned}$$

Taking the Laplace transform of $m_D(t)$, denoted as $\tilde{m}_D(s) = \frac{\tilde{G}(s)}{1-\tilde{F}(s)}$.

Proposition 1.3. *The following holds:*

1. $\lim_{t \rightarrow \infty} \frac{N_D(t)}{t} = \frac{1}{\mu}$.
2. $\lim_{t \rightarrow \infty} \frac{m_D(t)}{t} = \frac{1}{\mu}$.
3. If F is non-lattice, $\lim_{t \rightarrow \infty} m_D(t+a) - m_D(t) = \frac{a}{\mu_F}$.

4. If F and G are lattice with period d , $\mathbb{E}[\text{\# of renewals at } nd] = \frac{d}{\mu_F}$.

5. If F is nonlattice, $\mu < \infty$ and $h \in L^1$, then

$$\lim_{t \rightarrow \infty} \int_0^t h(t-x) dm_D(x) = \frac{\int_0^\infty h(t) dt}{\mu}.$$

1.4.1 Example:

Let $\{X_n : n \in \mathbb{N}\}$ be iid discrete observed. A pattern x_1, x_2, \dots, x_k is said to occur at time n if $X_n = x_k, X_{n-1} = x_{k-1}, \dots, X_{n-k+1} = x_1$. If we have iid tosses and consider $N(n)$ as the number of times pattern 0, 1, 0, 1 appear in n tosses, with $P(H) = p = 1 - q$, the process is a delayed renewal processes. To find the mean number of tosses for the first time the pattern 0, 1, 0, 1 appear,

$$\begin{aligned} \mathbb{E}[\text{first time pattern 0, 1, 0, 1 appears}] &= \mathbb{E}[\text{first time pattern 0, 1 appears}] \\ &\quad + \mathbb{E}[\text{time between patterns 0, 1, 0, 1}] \\ &= p^{-1}q^{-1} + p^{-2}q^{-2}. \end{aligned}$$

Similarly we can show that $\mathbb{E}[\text{first time } k \text{ heads}] = \sum_{i=1}^n p^{-i}$.