

## 1 k-Nearest Neighbour

$$\mathcal{P}(k|\lambda) = \frac{\lambda^k e^{-\lambda}}{k!} \quad (1)$$

$$\mathcal{P}(k=0|D) = e^{-2nD}$$

$$\mathcal{P}(k=1|\Delta D) = 2n e^{-2n\Delta D} \Delta D$$

$$\mathcal{P}(D|n) \Delta D = e^{-2nD} \cdot 2n e^{-2n\Delta D} \Delta D$$

$$\lim_{\Delta D \rightarrow 0} (\mathcal{P}(D|n) \Delta D) = \lim_{\Delta D \rightarrow 0} (e^{-2nD} \cdot 2n e^{-2n\Delta D} \Delta D)$$

$$\mathcal{P}(D|n) = 2n e^{-2nD} \quad (2)$$

## 2 MCMC Analysis

$$n(\theta) = \frac{1}{a} L(x|\theta) \quad (3)$$

$$\mathcal{P}(x) = \int d\theta L(x|\theta) \pi(\theta) \quad (4)$$

Since we assume a uniform prior,

$$\mathcal{P}(x) = \frac{a}{\Delta\theta} \int d\theta n(\theta)$$

We have samples from  $n$ , with density,

$$\hat{n}(\theta) = \sum_{\alpha=1}^S \delta(\theta - \theta_{\alpha}) \quad (5)$$

The integral over  $\theta$  is simply  $S$ , that is,

$$\hat{n}(\theta) = S$$

Hence, the estimate of the Bayesian Evidence is

$$\hat{\mathcal{P}}(x) = \frac{aS}{\Delta\theta} \quad (6)$$

### 3 Posterior of Local Density $n$

We use a Jeffreys prior on the scale parameter  $n$ , that is,  $\mathcal{P}(n) \propto \frac{1}{n}$

$$\mathcal{P}(n|D) \propto \mathcal{P}(D|n) \mathcal{P}(n)$$

$$\mathcal{P}(n|D) \propto 2n, e^{-2nD} \frac{1}{n}$$

$$\mathcal{P}(n|D) \propto e^{-2nD} \quad (7)$$

Assumption -  $\mathcal{P}(D)$  is a constant.

### 4 Posterior Probability of $a$

We again use a Jeffreys prior on the scale parameter  $a$ , that is,  $\mathcal{P}(a|L) \propto \frac{1}{a}$

$$\mathcal{P}(a|D, L) \propto \mathcal{P}(D|a, L) \mathcal{P}(a|L)$$

$$\mathcal{P}(a|D, L) \propto \int \mathcal{P}(D, n|a, L) \mathcal{P}(a|L) \mathcal{P}(n|a, L) dn$$

We assume  $\mathcal{P}(n|a, L) = \delta\left(n - \frac{L}{a}\right)$ . Therefore,

$$\mathcal{P}(a|D, L) \propto \int \mathcal{P}(D, n|a, L) \frac{1}{a} \delta\left(n - \frac{L}{a}\right) dn$$

But  $\mathcal{P}(D, n|a, L) \propto \mathcal{P}(n|D)$

$$\mathcal{P}(a|D, L) \propto \int e^{-2nD} \frac{1}{a} \delta\left(n - \frac{L}{a}\right) dn$$

$$\mathcal{P}(a|D, L) \propto \frac{1}{a^2} \exp\left(-\frac{2DL}{a}\right)$$

## 5 Considering All Samples

$$\mathcal{P}(a|D, L) = k \left( \prod_{\alpha=1}^S \frac{1}{a} \exp\left(-\frac{2D_{\alpha}L_{\alpha}}{a}\right) \right) \frac{1}{a}$$

$$\log \mathcal{P}(a|D, L) = \log k + \log \left( \prod_{\alpha=1}^S \frac{1}{a} \right) + \log \left( \prod_{\alpha=1}^S \exp\left(-\frac{2D_{\alpha}L_{\alpha}}{a}\right) \right) + \log \left( \frac{1}{a} \right)$$

$$\log \mathcal{P}(a|D, L) = \log k - S \log a - \frac{2}{a} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} - \log a$$

$$\log \mathcal{P}(a|D, L) = \text{const} - (S+1) \log a - \frac{2}{a} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} \quad (8)$$

### 5.1 Most Probable Value of $a$

$$\frac{d}{da} \left[ \text{const} - (S+1) \log a - \frac{2}{a} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} \right] = -\frac{(S+1)}{a} + \frac{2}{a^2} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha}$$

$$\frac{2}{a^2} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} = \frac{S+1}{a}$$

$$a = \frac{2}{S+1} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} \quad (9)$$

### 5.2 Estimate of the Bayesian Evidence

$$\hat{\mathcal{P}}(x) = \frac{aS}{\Delta\theta}$$

$$\hat{\mathcal{P}}(x) = \frac{2S}{\Delta\theta(S+1)} \sum_{\alpha=1}^S D_{\alpha}L_{\alpha} \quad (10)$$

## 6 Bayesian Evidence

We have  $\mathbf{x} = \begin{pmatrix} x_1^{(i)} \\ x_2^{(i)} \end{pmatrix}$  and we want to find  $\boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$ . We will assume uniform priors on  $\mu_1$  and  $\mu_2$  such that  $\mathcal{P}(\mu_1) = \frac{1}{\Delta\mu_1}$  and  $\mathcal{P}(\mu_2) = \frac{1}{\Delta\mu_2}$ . Using Bayes' Theorem,

$$\mathcal{P}(\boldsymbol{\mu}|\mathbf{x}) = \frac{\mathcal{P}(\mathbf{x}|\boldsymbol{\mu}) \mathcal{P}(\boldsymbol{\mu})}{\mathcal{P}(\mathbf{x})} \quad (11)$$

where

$$\mathcal{P}(\mathbf{x}|\boldsymbol{\mu}) = \frac{1}{|2\pi\Sigma|^{\frac{N}{2}}} \exp \left[ -\frac{1}{2} \sum_i \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right)^T \Sigma^{-1} \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right) \right]$$

The Bayesian Evidence is given by

$$\mathcal{P}(\mathbf{x}) = \int \mathcal{P}(\mathbf{x}|\boldsymbol{\mu}) \mathcal{P}(\boldsymbol{\mu}) d\boldsymbol{\mu} \quad (12)$$

Assuming the priors to be independent,

$$\mathcal{P}(\mathbf{x}) = \int \mathcal{P}(\mathbf{x}|\boldsymbol{\mu}) \mathcal{P}(\mu_1) \mathcal{P}(\mu_2) d\boldsymbol{\mu}$$

$$\mathcal{P}(\mathbf{x}) = \frac{1}{\Delta\mu_1 \Delta\mu_2} \int \mathcal{P}(\mathbf{x}|\boldsymbol{\mu}) d\boldsymbol{\mu}$$

$$\mathcal{P}(\mathbf{x}) = \frac{1}{\Delta\mu_1 \Delta\mu_2 |2\pi\Sigma|^{\frac{N}{2}}} \int \exp \left[ -\frac{1}{2} \sum_i \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right)^T \Sigma^{-1} \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right) \right] d\boldsymbol{\mu}$$

$$\mathcal{P}(\mathbf{x}) = \frac{| \frac{2\pi}{N} \Sigma |^{\frac{1}{2}}}{\Delta\mu_1 \Delta\mu_2 |2\pi\Sigma|^{\frac{N}{2}}} \quad (13)$$

We will also assume that

$$\Sigma = \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix}$$

## 7 Gibbs Sampling

$$\mathcal{P}(\boldsymbol{\mu} | \mathbf{x}) \propto \exp \left[ -\frac{1}{2} \sum_i \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right)^T \boldsymbol{\Sigma}^{-1} \left( \mathbf{x}^{(i)} - \boldsymbol{\mu} \right) \right]$$

$$\mathcal{P}(\boldsymbol{\mu} | \mathbf{x}) \propto \exp \left[ -\frac{1}{2\sigma^2} \sum_i \left( \mathbf{x}_1^{(i)} - \mu_1 \right)^2 \right] \exp \left[ -\frac{1}{2\sigma^2} \sum_i \left( \mathbf{x}_2^{(i)} - \mu_2 \right)^2 \right]$$

$$\mathcal{P}(\mu_1 | \mathbf{x}, \mu_2) = \frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp \left[ -\frac{1}{2(\sigma^2/N)} (\mu_1 - \bar{\mathbf{x}}_1)^2 \right]$$

$$\mathcal{P}(\mu_2 | \mathbf{x}, \mu_1) = \frac{1}{\sqrt{2\pi\frac{\sigma^2}{N}}} \exp \left[ -\frac{1}{2(\sigma^2/N)} (\mu_2 - \bar{\mathbf{x}}_2)^2 \right]$$