# Boston University Department of Electrical and Computer Engineering

## ENG EC 414 Introduction to Machine Learning

## **HW 1 Solution**

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**Important:** Before you proceed, please read the documents pertaining to *Homework formatting and submission guidelines* in the Homeworks section of Blackboard.

**Note:** Problem difficulty = number of coffee cups **...** 

**Problem 1.1** (*Linear Algebra*) [22pts] Let  $\mathbf{v}_1 = [1, 1, 0]^{\mathsf{T}}$ ,  $\mathbf{v}_2 = [0, 1, 1]^{\mathsf{T}}$ , and  $\mathbf{v}_3 = [1, 1, 1]^{\mathsf{T}}$ , be three column vectors. Note:  $^{\mathsf{T}}$  means transpose.

(a) [1pt] The dimension d of the Euclidean space  $\mathbb{R}^d$  containing  $\mathbf{v}_1$  is:

**Solution:** 3

(b) [1pt] The length, i.e., norm  $\|\mathbf{v}_1\|$ , of  $\mathbf{v}_1$  is:

**Solution:**  $\sqrt{1^2 + 1^2 + 0^2} = \sqrt{2}$ .

(c) [1pt] The dot product, i.e., inner product  $\langle \mathbf{v}_1, \mathbf{v}_2 \rangle = \mathbf{v}_2^{\mathsf{T}} \mathbf{v}_1$ , of  $\mathbf{v}_1$  and  $\mathbf{v}_2$  is:

**Solution:**  $1 \cdot 0 + 1 \cdot 1 + 0 \cdot 1 = 1$ .

(d) [1pt] Are  $\mathbf{v}_1$  and  $\mathbf{v}_2$  perpendicular (orthogonal)? Yes/No, Why?

**Solution:** 

No, because their inner product is not zero.

(e) [1pt] Are  $\mathbf{v}_1$  and  $\mathbf{v}_2$  linearly independent? Yes/No, Why?

**Solution:** 

Yes: Let  $V = [\mathbf{v}_1, \mathbf{v}_2]$  and  $\mathbf{a} = [a_1, a_2]^{\mathsf{T}}$ . The only solution to the equation  $V\mathbf{a} = 0$ , which tests the linear dependence of  $\mathbf{v}_1$  and  $\mathbf{v}_2$ , is  $\mathbf{a} = (V^{\mathsf{T}}V)^{-1}\mathbf{0} = \mathbf{0}$ . Note:  $(V^{\mathsf{T}}V)$  is an invertible matrix (its determinant is not zero).

(f) [5pts]  $\clubsuit$  If  $\operatorname{Proj}_{\mathcal{S}}(\mathbf{v}_3) = a_1\mathbf{v}_1 + a_2\mathbf{v}_2$ , where  $a_1, a_2$  are scalars, denotes the orthogonal projection of  $\mathbf{v}_3$  onto the subspace  $\mathcal{S}$  spanned by  $\mathbf{v}_1$  and  $\mathbf{v}_2$ , then  $\mathbf{a} = (a_1, a_2)^{\mathsf{T}} =$ 

**Solution:**  $(2/3, 2/3)^{T}$ :

Let  $V = [\mathbf{v}_1, \mathbf{v}_2]$ . Then  $\operatorname{Proj}_{\mathcal{S}}(\mathbf{v}_3) = V\mathbf{a}$ . According to the orthogonality principle,  $\mathbf{v}_3 - \operatorname{Proj}_{\mathcal{S}}(\mathbf{v}_3)$  is orthogonal to all vectors in  $\mathcal{S}$ , in particular, both  $\mathbf{v}_1$  and  $\mathbf{v}_2$ . Thus,  $V^{\mathsf{T}}(\mathbf{v}_3 - V\mathbf{a}) = 0$ . Solving, we get  $\mathbf{a} = (V^{\mathsf{T}}V)^{-1}V^{\mathsf{T}}\mathbf{v}_3 = (2/3, 2/3)^{\mathsf{T}}$ , and  $\operatorname{Proj}_{\mathcal{S}}(\mathbf{v}_3) = (2/3, 4/3, 2/3)^{\mathsf{T}}$ .

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(g) [5pts] Consider the following subset of  $\mathbb{R}^3$ :  $S := \{\mathbf{x} : \mathbf{v}_3^\top \mathbf{x} - 3 = 0\}$ . Compute the Euclidean distance of  $\mathbf{v}_1$  from S and the point in S which is closest to  $\mathbf{v}_1$ .

## **Solution:**

The point

$$\mathbf{v} := \mathbf{v}_1 - (\mathbf{v}_3^{\mathsf{T}} \mathbf{v}_1 - 3) \frac{\mathbf{v}_3}{\|\mathbf{v}_3\|^2} = (1^{1/3}, 1^{1/3}, 1/3)^{\mathsf{T}}$$

lies in S since  $\mathbf{v}_3^{\mathsf{T}}\mathbf{v} - 3 = 0$ . If  $\mathbf{u}$  is any other point in S, then we have  $\mathbf{v}_3^{\mathsf{T}}\mathbf{u} - 3 = 0$  and therefore  $\mathbf{v}_3^{\mathsf{T}}(\mathbf{v} - \mathbf{u}) = 0$ . Thus,

$$\|\mathbf{v}_{1} - \mathbf{u}\|^{2} = \|(\mathbf{v}_{1} - \mathbf{v}) + (\mathbf{v} - \mathbf{u})\|^{2} = \|\mathbf{v}_{1} - \mathbf{v}\|^{2} + \|\mathbf{v} - \mathbf{u}\|^{2} + 2\underbrace{(\mathbf{v}_{1} - \mathbf{v})^{\mathsf{T}}(\mathbf{v} - \mathbf{u})}^{0} \ge \|\mathbf{v}_{1} - \mathbf{v}\|^{2}.$$

This shows that  $\mathbf{v}$  is the point in S that is closest to  $\mathbf{v}_1$ . The minimum distance is equal to  $\|\mathbf{v}_1 - \mathbf{v}\| = \|\mathbf{v}_3^\top \mathbf{v}_1 - 3\| \|\mathbf{v}_3\| = 1/\sqrt{3}$ .

(h) [4pts] Let  $B = \begin{bmatrix} 4 & 3 \\ 3 & 4 \end{bmatrix}$ . Compute: (i) its eigenvalues and (ii) a set of orthonormal eigenvectors.

## **Solution:**

- (i) Eigenvalues:  $\lambda = 1,7$  obtained as the solutions to the quadratic equation  $\det(B \lambda I) = 0$  where I is the  $2 \times 2$  identity matrix. (ii) One set of orthonormal eigenvectors (not unique):  $\mathbf{u}_1 = \frac{1}{\sqrt{2}}(1,1)^{\top}$  and  $\mathbf{u}_2 = \frac{1}{\sqrt{2}}(1,-1)^{\top}$ . Note: orthonormal means, mutually orthogonal (zero inner product) and unit norm (length).
- (i) [3pts] The trace tr(D) of a square matrix D is the sum of all its elements along the main diagonal. Let D = ABC, where the dimensions of A, B, and C are, respectively,  $p \times q$ ,  $q \times r$ , and  $r \times p$ . What is the relationship between: tr(ABC), tr(BCA), and tr(CAB)? Explain.

#### Solution:

They are all equal! Proof:  $tr(D) = \sum_i D_{ii}$ . Also,  $D_{ij} = \sum_{k,l} A_{ik} B_{kl} C_{lj}$ . Thus,  $tr(D) = \sum_{i,k,l} A_{ik} B_{kl} C_{li}$  which is symmetric with respect to circular re-orderings of A - B - C (in that order).

**Problem 1.2** (Multivariate Calculus) [9pts] Let A be a  $d \times d$  matrix and  $\mathbf{b}, \mathbf{x} \in \mathbb{R}^d$  be two  $d \times 1$  column vectors. Let  $f(\mathbf{x})$  denote a real-valued function of d variables (d components of  $\mathbf{x}$ ).

(a) [2pts] Compute the gradient vector  $\nabla f(\mathbf{x}) = \left(\frac{\partial f}{\partial x_1}(\mathbf{x}), \dots, \frac{\partial f}{\partial x_d}(\mathbf{x})\right)^{\top}$  when  $f(\mathbf{x}) = \mathbf{b}^{\top} \mathbf{x}$ . Solution:

$$f(\mathbf{x}) = \sum_{i} b_{i} x_{i} \Rightarrow \frac{\partial f}{\partial x_{i}}(\mathbf{x}) = b_{i} \Rightarrow \nabla f(\mathbf{x}) = \mathbf{b}.$$

(b) [3pts] Compute the gradient vector  $\nabla f(\mathbf{x})$  when  $f(\mathbf{x}) = \mathbf{x}^{\mathsf{T}} A \mathbf{x}$ .

#### **Solution:**

Let  $f(\mathbf{x}) = \mathbf{x}^{\mathsf{T}} \mathbf{b}$  where  $\mathbf{b} = A\mathbf{x}$  is a function of  $\mathbf{x}$ . Then  $f(\mathbf{x}) = \sum_{i} x_{i} b_{i}$ . By the chain rule,

$$\frac{\partial f}{\partial x_i} = \sum_i \left[ b_j \frac{\partial x_j}{\partial x_i} + x_j \frac{\partial b_j}{\partial x_i} \right] = b_i + \sum_i x_j A_{ji} = \sum_i (A_{ij} + A_{ji}) x_j = \sum_i (A + A^\top)_{ij} x_j.$$

Hence,  $\nabla f(\mathbf{x}) = (A + A^{\mathsf{T}})\mathbf{x}$ .

(c) [4pts]  $\clubsuit$  Let A be symmetric and invertible. If  $f(\mathbf{x}) = \frac{1}{2}\mathbf{x}^{\top}A\mathbf{x} + \mathbf{b}^{\top}\mathbf{x}$ , then find  $\mathbf{x}$ 's for which  $f(\mathbf{x})$  is minimum or maximum.

## **Solution:**

 $f(\mathbf{x})$  is differentiable everywhere and is therefore minimized or maximized when  $\nabla f(\mathbf{x}) = 0$  or at infinity. Using parts (a) and (b) and the fact that  $A = A^{\top}$  (symmetric matrix) we get  $\nabla f(\mathbf{x}) = A\mathbf{x} + \mathbf{b} = 0 \Rightarrow \mathbf{x} = -A^{-1}\mathbf{b}$  is the unique solution (since A is invertible). Thus  $f(\mathbf{x})$  is minimum or maximum at  $\mathbf{x} = -A^{-1}\mathbf{b}$  or at infinity. If A was positive definite, then  $f(\mathbf{x})$  would be minimum at  $\mathbf{x} = -A^{-1}\mathbf{b}$  and maximum (in fact,  $+\infty$ ) at infinity. If A was negative definite, then  $f(\mathbf{x})$  would be maximum at  $\mathbf{x} = -A^{-1}\mathbf{b}$  and minimum (in fact,  $-\infty$ ) at infinity. If A was neither positive nor negative definite, then the maximum is  $+\infty$  and the minimum is  $-\infty$  and both occur at infinity (along different directions).

**Problem 1.3** (*Probability*) [21pts] Let W = Y + U, where label Y and additive noise U are independent random variables,  $P(Y = +1) = \frac{3}{4}$ ,  $P(Y = -1) = \frac{1}{4}$ , and U is continuous with  $U \sim \text{Uniform}[-3, 3]$ .

Let observation (feature)  $X = 2 \times 1(W > 0) - 1$ . where 1(event) is the indicator function of the *event* and equals one if the *event* is true and equals zero if the *event* is false.

(a) [2pts] Sketch the graph of p(w|Y = +1). Proper labeling of axes and key points is needed to receive full credit.

**Solution:** The graph of p(w|Y = +1) is piecewise constant. It is equal to the constant 1/6 from w = -2 to w = 4 and the constant 0 for all other values of w. Key break points to indicate: w = -2, 4. Key values to indicate 1/6. Label w for horizontal axis.

(b) [2pts] Compute the joint pmf p(x, y) = P(X = x, Y = y) for all  $x, y \in \{-1, +1\}$ .

**Solution:** Note:  $2 \times 1(W > 0) - 1 = \text{sign}(W)$ . Thus both X and Y are discrete and take only the two values:  $\pm 1$  with nonzero probability.  $P(X = +1, Y = +1) = P(W > 0, Y = +1) = P(W > 0|Y = +1) \times P(Y = +1) = \frac{4}{6} \times \frac{3}{4} = \frac{1}{2}$ . Similarly, for other x, y:

$$p(x,y) = \begin{vmatrix} x = -1 & x = +1 \\ y = +1 & \frac{1}{4} & \frac{1}{2} \\ y = -1 & \frac{1}{6} & \frac{1}{12} \end{vmatrix}$$

(c) [1pt] Compute the marginal pmf p(x) = P(X = x), for x = -1, 1.

**Solution:**  $P(X = -1) = P(X = -1, Y = +1) + P(X = -1, Y = -1) = \frac{1}{4} + \frac{1}{6} = \frac{5}{12}$ . Also,  $P(X = +1) = 1 - P(X = -1) = \frac{7}{12}$ 

(d) [2pts] Compute the mean/expectation:  $\mu_X = E[X]$  and  $\mu_Y = E[Y]$ .

**Solution:** We have,  $\mu_X = E[X] = -1 \times \frac{5}{12} + 1 \times \frac{7}{12} = \frac{1}{6}$ . Similarly,  $\mu_Y = E[Y] = -1 \times \frac{1}{4} + 1 \times \frac{3}{4} = \frac{1}{2}$ .

(e) [2pts] Compute the variance:  $\sigma_X^2 = \text{var}(X)$ .

## **Solution:**

We have 
$$\sigma_X^2 = \text{var}(X) = E[X^2] - (E[X])^2 = 1 - (1/6)^2 = 35/36$$
.

(f) [2pts] Compute the correlation E[XY]. Are X and Y orthogonal? Explanation is needed to receive credit.

## **Solution:**

 $E[XY] = 1 \times P(X = 1, Y = 1) + 1 \times P(X = -1, Y = -1) - 1 \times P(X = -1, Y = +1) - 1 \times P(X = +1, Y = -1) = \frac{1}{2} + \frac{1}{6} - \frac{1}{4} - \frac{1}{12} = \frac{1}{3}$ . No, X and Y are not orthogonal because their correlation is not equal to zero.

(g) [2pts] Compute the covariance:  $cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$ . Are *X* and *Y* uncorrelated? Explanation is needed to receive credit.

**Solution:**  $cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - \mu_X \mu_Y = 1/3 - 1/6 \times 1/2 = 1/4 \neq 0$ . No, X and Y are not uncorrelated because  $cov(X, Y) \neq 0$ .

(h) [2pts] Compute the conditional pmf p(x|y) = P(X = x|Y = y) for all  $x, y \in \{-1, +1\}$ .

**Solution:** *Note:* both *X* and *Y* are discrete and take only the two values:  $\pm 1$  with nonzero probability.  $P(X = +1|Y = +1) = P(W > 0|Y = +1) = \frac{4}{6} = \frac{2}{3}$ . Similarly, for other *x*, *y*:

		x = -1	x = +1
p(x y) =	y = +1	1/3	2/3
	y = -1	2/3	1/3

(i) [2pts] Are X and Y independent? Explanation is needed to receive credit.

**Solution:** X, Y not independent:  $P(X = +1|Y = +1) = \frac{2}{3} \neq P(X = +1|Y = -1) = \frac{1}{3}$ .

(j) [4pts] Compute the conditional expectation/mean E[Y|X=x] as a function of x.

**Solution:**  $E[Y|X=x] = 1 \cdot P(Y=+1|X=x) - 1 \cdot P(Y=-1|X=x) = 2P(Y=1|X=x) - 1 = 2P(X=x|Y=1) \cdot P(Y=1)/P(X=x) - 1 = \begin{cases} 2 \cdot \frac{2/3 \cdot 3/4}{2/3 \cdot 3/4 + 1/3 \cdot 1/4} - 1 = 5/7 & \text{if } x = +1, \\ 2 \cdot \frac{1/3 \cdot 3/4}{1/3 \cdot 3/4 + 2/3 \cdot 1/4} - 1 = 1/5 & \text{if } x = -1. \end{cases}$ 

**Problem 1.4** (Random Vectors) [10pts] Let  $Z_1, Z_2$  be i.i.d. (scalar) standard Gaussians (normals)  $\mathcal{N}(0, 1)$ , i.e., independent and identically Gaussians with zero mean and unit variance. Let

$$\underbrace{\begin{bmatrix} X \\ Y \end{bmatrix}}_{\mathbf{U}} = \underbrace{\begin{bmatrix} 1 & 0 \\ 1 & 1 \end{bmatrix}}_{A} \underbrace{\begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix}}_{\mathbf{Z}} + \underbrace{\begin{bmatrix} -3 \\ 2 \end{bmatrix}}_{\mathbf{v}}$$

- (a) [2pts] Compute the  $2 \times 1$  mean vector  $E[\mathbf{U}]$ .
- (b) [4pts] Compute the  $2 \times 2$  (auto- or self-) covariance matrix Cov(U) of the random vector  $\mathbf{U} = (X, Y)^{\mathsf{T}}$ .
- (c) [4pts] Compute the  $2 \times 2$  cross-covariance matrix  $Cov(\mathbf{U}, \mathbf{Z})$  of the random vector  $\mathbf{U}$  and the random vector  $\mathbf{Z} = (Z_1, Z_2)^{\mathsf{T}}$ .

## **Solution:**

- (a) [2pts] Since  $Z_1$  and  $Z_2$  have zero means, we have  $E[\mathbf{U}] = A E[\mathbf{Z}] + \mathbf{v} = \mathbf{v} = (-3, 2)^{\mathsf{T}}$ .
- (b) [4pts] Since  $Z_1$  and  $Z_2$  have zero means, are independent (therefore uncorrelated), and have and unit variances,  $Cov(\mathbf{Z}) = I_2$ , the 2×2 identity matrix. Thus,  $Cov(\mathbf{U}) = A \cdot Cov(\mathbf{Z}) \cdot A^{\top} = A \cdot I_2 \cdot A^{\top} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \end{bmatrix}$ .
- (c) [4pts]  $Cov(\mathbf{U}, \mathbf{Z}) = Cov(A\mathbf{Z} + \mathbf{v}, \mathbf{Z}) = A \cdot Cov(\mathbf{Z}) = A \cdot I_2 = \begin{bmatrix} 1 & 0 \\ 1 & 1 \end{bmatrix}$ .