

Topic: Expectation of random variables

In today's problem class we will be computing expectations of various discrete and continuous random variables.

1. Compute the mean of the following random variables.

(a) $X \sim \text{Poi}(\lambda)$,

(b) $X \sim \text{Exp}(\lambda)$,

Solution: Recall that we have

$$E(X) = \begin{cases} \sum_n nP(X = n), & \text{if } X \text{ is discrete,} \\ \int x f_X(x) dx, & \text{if } X \text{ is continuous.} \end{cases}$$

(a) In the case when $X \sim \text{Poi}(\lambda)$, we have

$$\begin{aligned} E(X) &= \sum_{n=0}^{\infty} nP(X = n) = \sum_{n=0}^{\infty} n \frac{\lambda^n}{n!} e^{-\lambda} = e^{-\lambda} \sum_{n=1}^{\infty} \frac{\lambda^n}{(n-1)!} \\ &= e^{-\lambda} \lambda \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} = e^{-\lambda} e^{\lambda} \lambda = \lambda. \end{aligned}$$

(b) In the case when $X \sim \text{Exp}(\lambda)$, we have

$$E(X) = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \frac{1}{\lambda} \int_0^{\infty} z e^{-z} dz = \Gamma(2) \frac{1}{\lambda} = \frac{1}{\lambda}.$$

2. Let X be a continuous random variable with the following p.d.f.:

$$f_X(x) = \begin{cases} \theta \lambda e^{-\lambda x}, & x \geq 0; \\ (1 - \theta) \lambda e^{\lambda x}, & x < 0. \end{cases}$$

where λ and θ are constants such that $\lambda > 0$ and $0 \leq \theta \leq 1$.

(a) Show that $f_X(x)$ is a valid p.d.f..

(b) Find $E(X)$. *Hint: You might find it useful to look up the definition of the Gamma function.*

(c) Find $\text{Var}(X)$

Solution:

(a) $f_X(x) \geq 0$ for all $x \in \mathbb{R}$, since $\lambda > 0, \theta \in (0, 1)$ and the exponential function is always positive, and

$$\begin{aligned} \int_{-\infty}^{\infty} f_X(x) dx &= \int_{-\infty}^0 (1 - \theta) \lambda e^{\lambda x} dx + \int_0^{\infty} \theta \lambda e^{-\lambda x} dx \\ &= (1 - \theta) e^{\lambda x} \Big|_{-\infty}^0 - \theta e^{-\lambda x} \Big|_0^{\infty} = 1 - \theta + \theta = 1. \end{aligned}$$

(b) For the expectation, we have

$$E(X) = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^0 (1 - \theta) \lambda x e^{\lambda x} dx + \int_0^{\infty} \theta \lambda x e^{-\lambda x} dx.$$

We have, using the transformation $z = -\lambda x$, $dz = -\lambda dx$:

$$\begin{aligned}\int_{-\infty}^0 (1-\theta)\lambda x e^{\lambda x} dx &= (1-\theta) \int_{\infty}^0 (-z)e^{-z}(-1)dz/\lambda = (1-\theta)(-1) \int_0^{\infty} z e^{-z} dz/\lambda \\ &= -(1-\theta)/\lambda \Gamma(1) = -(1-\theta)/\lambda.\end{aligned}$$

Similarly, using the transformation $z = \lambda x$, $dz = \lambda dx$:

$$\begin{aligned}\int_0^{\infty} \theta \lambda x e^{-\lambda x} dx &= \int_0^{\infty} \theta z e^{-z} dz \\ &= \theta/\lambda \Gamma(1) = \theta/\lambda.\end{aligned}$$

Hence, $E(X) = -(1-\theta)/\lambda + \theta/\lambda = (2\theta - 1)/\lambda$.

(c) We compute the second moment first:

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_{-\infty}^0 (1-\theta)\lambda x^2 e^{\lambda x} dx + \int_0^{\infty} \theta \lambda x^2 e^{-\lambda x} dx.$$

For the first integral, using the same transformation as above $z = -\lambda x$, $dz = -\lambda dx$, we get:

$$\begin{aligned}\int_{-\infty}^0 (1-\theta) \frac{1}{\lambda} \lambda^2 x^2 e^{\lambda x} dx &= (1-\theta) \frac{1}{\lambda} \int_{\infty}^0 z^2 e^{-z} (-1) dz/\lambda = \frac{(1-\theta)}{\lambda^2} \int_0^{\infty} z^2 e^{-z} dz \\ &= \frac{(1-\theta)}{\lambda^2} \Gamma(3) = \frac{2(1-\theta)}{\lambda^2}.\end{aligned}$$

Similarly, for the second integral, using the transformation $z = \lambda x$, $dz = \lambda dx$:

$$\begin{aligned}\int_0^{\infty} \theta \lambda x^2 e^{-\lambda x} dx &= \frac{\theta}{\lambda^2} \int_0^{\infty} z^2 e^{-z} dz \\ &= \frac{\theta}{\lambda^2} \Gamma(3) = \frac{2\theta}{\lambda^2}.\end{aligned}$$

So altogether we have

$$E(X^2) = \frac{2}{\lambda^2}$$

and hence

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{2}{\lambda^2} - \frac{(2\theta - 1)^2}{\lambda^2} = \frac{1 - 4\theta^2 + 4\theta}{\lambda^2}.$$