

应用数据分析与建模介绍(GE14208)



Lecture6 计数数据分析

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目录

- 1. 泊松分布
- 2. 泊松回归
- 3. 广义线性模型

变量分类:

- 定量变量(quantitative): 有具体数值 含义
- 定性变量(qualitative, categorical): 分类数据,两分类与多分类

有些数据既不是传统的定量数据,也不是定性数据。比如交通事故的死亡人数,每小时共享单车被租用的次数,它有什么特点?

- 非负
- 整数

这类数据,我们称之为计数数据(Count Data)。用线性回归模型分析是否合适?

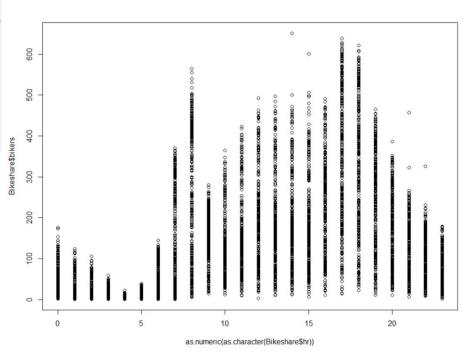
```
```{r libraries}
library(ggplot2)
library(ISLR2)
philly<- read.csv("Philly.csv")</pre>
data("Bikeshare")
```{r Check the data}
### 1. 查看数据
## 查看Injury分布
print(table(philly$Injury))
```

对Bikeshare数据中的bikers建立线性回归模型

> print(summary(lm0\$fitted.values))
 Min. 1st Qu. Median Mean 3rd Qu. Max.
-143.74 61.89 142.38 143.79 221.90 441.36

$$Y = \beta_0 + \beta_1 * X_1$$
, 出现的问题?

- 预测结果有负值
- 预测结果不为整数
- bikers方差显然随小时有明显分布差异



如果 $\log(Y) = \beta_0 + \beta_1 * X_1$? 也有问题:结果解释,以及Y = 0的情况

假设随机变量Y为非负整数, $Y \in \{0,1,2,...\}$ 。如果Y服从泊松分布,则

$$P(Y=k) = \frac{e^{-\lambda}\lambda^k}{k!}$$

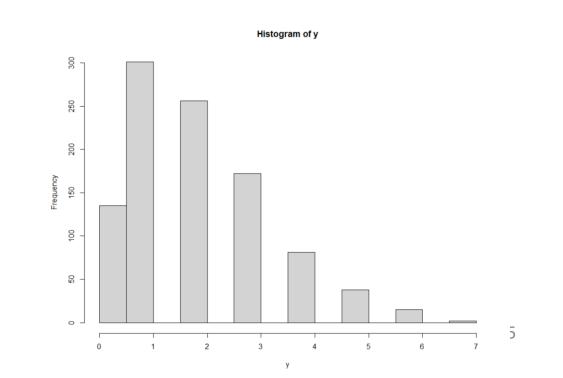
for k = 0, 1, 2, ...

式中: $\lambda > 0$

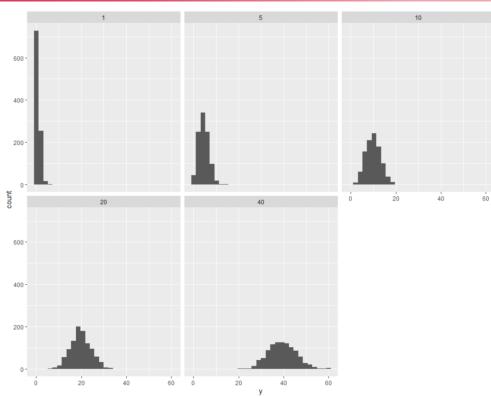
对于泊松分布:

$$\lambda = E(Y) = Var(Y)$$

1. 随机产生一个lambda为2的泊松数据集 y<- rpois(n=1000, lambda = 2) print(hist(y))



```
### 2. 创造不同1ambda的数据集
lambdas<- c(1,5,10,20,40)
for(i in 1:length(lambdas)){
  y < -rpois(n=1000, lambda = lambdas[i])
  dat<- data.frame(y)</pre>
  dat$lambda<- lambdas[i]
  if(i==1){
    dats<- dat
  }else{
    dats<- rbind(dats,dat)</pre>
  print(i)
print(head(dats))
  画出不同lambda对应的直方图
q1<- ggplot(data = dats)+
  geom_histogram(aes(x=y))+
  facet_wrap(~lambda)
print(g1)
```



随着λ增大,泊松分布会趋向于正态分布。

泊松分布:
$$P(y) = \frac{e^{-\lambda}\lambda^y}{k!}$$

假设每次交通事故受伤符合一个 $\lambda = 1$ 的泊松分布,那么每次交通事故受伤人数为0, 1, 2, 不多于2人的概率分别为:

$$P(Y=0) = \frac{e^{-1} * 1^0}{0!} = 0.368$$

$$P(Y=1) = \frac{e^{-1} * 1^{1}}{1!} = 0.368$$

$$P(Y=2) = \frac{e^{-1} * 1^2}{2!} = 0.184$$

$$P(Y \le 2) = P(Y = 0) + P(Y = 1) + P(Y = 2)$$

计算概率 #概率密度函数 p1<- dpois(x=0, lambda = 1) print(p1)

print(p3)

分布函数F(x) p4<- ppois(q=2,lambda = 1) print(p4)

假设我们有n个观察值 (y_i) 来自于泊松回归模型,则其似然公式为

$$l(\lambda) = \prod_{i=1}^{n} \frac{e^{-\lambda} \lambda^{y_i}}{y_i!}, \ \lambda = \frac{\sum_{i=1}^{n} y_i}{n}$$

对于泊松回归, 我们对平均值λ进行建模, 即

$$P(Y = k) = \frac{e^{-\lambda} \lambda^k}{k!}$$
$$log(\lambda) = \beta_0 + \beta_1 * X_1 + \dots + \beta_p * X_p$$

也就是
$$\lambda = e^{(\beta_0 + \beta_1 * X_1 + \dots + \beta_p * X_p)}$$
。

式中:
$$\beta_0$$
, β_1 , ..., β_p 为需要估计的回归系数。

假设我们有n个观察值 (x_i, y_i) 来自于泊松回归模型,则其似然公式为

$$l(\beta_0, \beta_1, \dots, \beta_p) = \prod_{i=1}^{n} \frac{e^{-\lambda_i \lambda_i^{y_i}}}{y_i!}$$
$$\lambda_i = e^{(\beta_0 + \beta_1 * x_{i1} + \dots + \beta_p * x_{ip})}$$
$$x_i = (x_{i1}, x_{i2}, \dots, x_{ip})$$

对Philly数据中的Injury分别建立线性回归模型与泊松回归模型

```
## 线性回归
lm1<- lm(Injury~Intersection+Daytime+Weather,</pre>
           data = philly)
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
           (Intercept)
Intersection 0.237813 0.007134 33.336 < 2e-16 ***
Daytime 0.088326 0.007392 11.949 < 2e-16 ***
WeatherRain -0.076644 0.009981 -7.679 1.62e-14 ***
                    0.025402 -7.662 1.84e-14 ***
WeatherSnow -0.194640
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1.032 on 85705 degrees of freedom
 (2 observations deleted due to missingness)
Multiple R-squared: 0.01639. Adjusted R-squared: 0.01635
F-statistic: 357.1 on 4 and 85705 DF, p-value: < 2.2e-16
```

Residual deviance: 77658 on 85705 degrees of freedom

(2 observations deleted due to missingness)

估计模型:

• 线性回归模型:

Y = 0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnow = 0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnow = 0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnow = 0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnow = 0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnow = 0.914053 + 0.914054 + 0.91

AIC: 222347

• 泊松回归模型:

```
\lambda = \rho^{0.914053 + 0.237813 * Intersection + 0.088326 * Daytime - 0.076644 * WeatherRain - 0.194640 * WeatherSnown}
```

对Philly数据中的Injury分别建立线性回归模型与泊松回归模型

```
lm1<- lm(Injury~Intersection+Daytime+Weather,</pre>
           data = philly)
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
            0.914053 0.007343 124.472 < 2e-16 ***
(Intercept)
Intersection 0.237813 0.007134 33.336 < 2e-16 ***
Daytime 0.088326 0.007392 11.949 < 2e-16 ***
WeatherRain -0.076644 0.009981 -7.679 1.62e-14 ***
                     0.025402 -7.662 1.84e-14 ***
WeatherSnow -0.194640
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Residual standard error: 1.032 on 85705 degrees of freedom
 (2 observations deleted due to missingness)
Multiple R-squared: 0.01639. Adjusted R-squared: 0.01635
F-statistic: 357.1 on 4 and 85705 DF, p-value: < 2.2e-16
参数解释:
                                                                 AIC: 222347
```

```
## 泊松回归
poisson1<- glm(Injury~Intersection+Daytime+Weather,
          data = philly,
          family = poisson
Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) -0.085487  0.007181 -11.905 < 2e-16 ***
Intersection 0.222348 0.006768 32.855 < 2e-16 ***
Daytime
            WeatherRain -0.072198  0.009503  -7.597  3.03e-14 ***
WeatherSnow -0.198012  0.026045  -7.603  2.90e-14 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for poisson family taken to be 1)
   Null deviance: 79072 on 85709 degrees of freedom
Residual deviance: 77658 on 85705 degrees of freedom
  (2 observations deleted due to missingness)
```

Intersection:

线性回归

- ▶ 线性回归:发生在交叉口的交通事故平均受伤人数比非交叉口多0.237813人。
- ▶ 泊松回归: 发生在交叉口的交通事故平均受伤人数是非交叉口的exp(0.222348) = 1.249006倍

• Daytime:

- ▶ 线性回归:发生在交叉口的交通事故平均受伤人数比非交叉口多_____人?
- 泊松回归:发生在交叉口的交通事故平均受伤人数是非交叉口的_____倍?

对Philly数据中的Injury分别建立线性回归模型与泊松回归模型

```
## 线性回归
lm1<- lm(Injury~Intersection+Daytime+Weather,</pre>
          data = philly)
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.914053 0.007343 124.472 < 2e-16 ***
```

```
Intersection 0.237813 0.007134 33.336 < 2e-16 ***
Daytime 0.088326 0.007392 11.949 < 2e-16 ***
```

WeatherRain -0.076644 0.009981 -7.679 1.62e-14 *** 0.025402 -7.662 1.84e-14 *** WeatherSnow -0.194640

```
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

Residual standard error: 1.032 on 85705 degrees of freedom (2 observations deleted due to missingness) Multiple R-squared: 0.01639. Adjusted R-squared: 0.01635 F-statistic: 357.1 on 4 and 85705 DF, p-value: < 2.2e-16

平均值-方差关系:

- 线件回归: 方差恒定, σ^2 .
- 泊松回归: 方差等于平均值, 也就是说方差是会变化的, $Var(Y) = \lambda = e^{\left(\beta_0 + \beta_1 * X_1 + \cdots + \beta_p * X_p\right)}$ 。

平均值-方差关系:

- 线性回归:可能为负值。
- 泊松回归: 总是非负值。

```
## 泊松回归
poisson1<- glm(Injury~Intersection+Daytime+Weather,</pre>
         data = philly,
         family = poisson)
Coefficients:
```

Estimate Std. Error z value Pr(>|z|)(Intercept) -0.085487 0.007181 -11.905 < 2e-16 *** Intersection 0.222348 0.006768 32.855 < 2e-16 *** Davtime WeatherRain -0.072198 0.009503 -7.597 3.03e-14 *** WeatherSnow -0.198012 0.026045 -7.603 2.90e-14 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 1) Null deviance: 79072 on 85709 degrees of freedom Residual deviance: 77658 on 85705 degrees of freedom

(2 observations deleted due to missingness) AIC: 222347

3. 广义线性模型(Generalized Linear Model)

• 线性回归模型(Linear Regression):

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

$$E(Y|X_1, \dots, X_p) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

• 逻辑回归模型(Logistic Regression):

$$Y \sim Ber(p)$$

$$\log\left(\frac{E(Y|X_1,...,X_p)}{1 - E(Y|X_1,...,X_p)}\right) = \log(\frac{p}{1 - p}) = \beta_0 + \beta_1 * X_1 + \dots + \beta_p * X_p$$

• 泊松回归模型(Poisson Regression):

$$Y \sim Poi(\lambda)$$

$$\log(E(Y|X_1,...,X_p)) = \log(\lambda(X_1,...,X_p)) = \beta_0 + \beta_1 * X_1 + \dots + \beta_p * X_p$$

3. 广义线性模型(Generalized Linear Model)

上面均对 $E(Y|X_1,...,X_n)$ 进行一些变换,使得其成为预测变量的线性函数

- 线性回归: $E(Y|X_1,...,X_p) = \beta_0 + \beta_1 X_1 + \cdots + \beta_p X_p$
- 逻辑回归: $\log\left(\frac{E(Y|X_1,...,X_p)}{1-E(Y|X_1,...,X_p)}\right) = \beta_0 + \beta_1 * X_1 + \cdots + \beta_p * X_p$
- 泊松回归: $\log(E(Y|X_1,...,X_p)) = \log(\lambda(X_1,...,X_p)) = \beta_0 + \beta_1 * X_1 + \cdots + \beta_p * X_p$

这个变换函数我们称之为连接函数(Link Function), η 。

$$\eta(E(Y|X_1,...,X_p) = \mu) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

线性回归:
$$\eta(\mu) = \mu$$

逻辑回归:
$$\eta(\mu) = \log(\frac{\mu}{1-\mu})$$

线性回归: $\eta(\mu) = \log(\mu)$

3. 广义线性模型(Generalized Linear Model)

指数模型家族(Exponential Family):

- 高斯分布(Gaussian Distribution)
- 伯努利分布(Bernoulli Distribution)
- 泊松分布(Poisson Distribution)
- 负二项分布(Negative Distribution)
- 指数分布(Exponential Distribution)
- 伽马分布(Gamma Distribution)
- ...

• An exponential family distribution has the following form,

$$p(x \mid \eta) = h(x) \exp\{\eta^{\top} t(x) - a(\eta)\}$$
(1)

- The different parts of this equation are
 - The natural parameter η
 - The sufficient statistic t(x)
 - The underlying measure h(x), e.g., counting measure or Lebesgue measure
 - The log normalizer $a(\eta)$,

$$a(\eta) = \log \int h(x) \exp\{\eta^{\mathsf{T}} t(x)\}. \tag{2}$$

Here we integrate the unnormalized density over the sample space. This ensures that the density integrates to one.

Source: Prof. David M. Blei

广义线性模型(Generalized Linear Model, GLM):如果因变量Y符合指数模型家族的某种分布,且可以将其平均值表示为自变量的线性组合,那么这个回归模型就被称为广义线性模型。

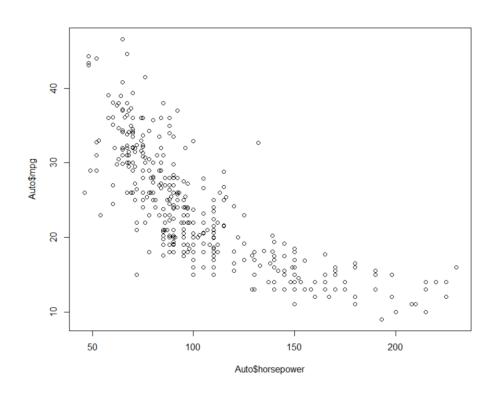
用Auto数据集,回顾线性回归模型: Im vs. glm。

```
## 线性回归: lm vs. glm
lm1<- lm(mpg~horsepower,data = Auto)
print(lm1)
glm1<- glm(mpg~horsepower,data = Auto)
print(glm1)
```

注意: glm中的family没有做任何设置!

小知识1: lm与glm结果完全一致!

用Auto数据集, 查看多项式(Polynomial)数据:



根据散点图,mpg 与horsepower是 否呈现线性关系, 或者是其他关系?

用Auto数据集, 查看多项式(Polynomial)数据:

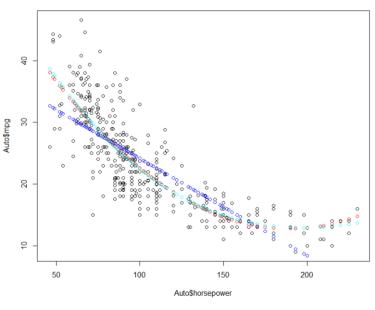
```
## 多项式(polynomial)数据创造: poly()和I()
#一次方
a1 < -poly(x=3, degree = 1, raw = TRUE)
print(a1)
# 二次方(Quadratic)
a2 < -poly(x=3, degree = 2, raw = TRUE)
print(a2)
# 立方(Cubic)
a3 < -poly(x=3, degree = 3, raw = TRUE)
print(a3)
# 四次方(Quartic)
a4<- poly(x=Auto$horsepower, degree = 4,raw = TRUE)
print(head(a4))
a5<- I(Auto$horsepower^4)
print(head(a5))
```

```
> print(a2)
                               > print(a3)
       1 2
                                     1 2 3
 [1,] 3 9
                               [1,] 3 9 27
attr(,"degree")
                               attr(,"degree")
 \lceil 1 \rceil 1 2
                               [1] 1 2 3
attr(,"class")
                               attr(,"class")
 [1] "poly" "matrix"
                               [1] "poly"
                                               "matrix"
> head(a4)
            2
[1,] 130 16900 2197000 285610000
[2,] 165 27225 4492125 741200625
[3,] 150 22500 3375000 506250000
[4,] 150 22500 3375000 506250000
[5,] 140 19600 2744000 384160000
[6,] 198 39204 7762392 1536953616
> head(a5)
   285610000 741200625 506250000
                                  506250000 384160000 1536953616
```

用Auto数据集,查看多项式(Polynomial)数据:

```
## 1.0 多项式公式分析mpg影响因素
# 模型1: 一次方
glm1<- glm(mpg~horsepower,data = Auto)</pre>
# 模型2: 二次方: I()
glm2<-glm(mpg\sim horsepower+I(horsepower^2), data = Auto)
# 模型3: 三次方: poly()
glm3<- glm(mpg~poly(horsepower, 3, raw = TRUE), data = Auto)
# 模型评价: AIC和MSE(Mean squared error)
aics<- c(glm1$aic,glm2$aic,glm3$aic)</pre>
mses<- c(mean((glm1\$y - glm1\$fitted.values)^2),
         mean((glm2$y - glm2$fitted.values)^2),
         mean((glm3$y - glm3$fitted.values)^2))
print(aics)
                        > print(aics)
[1] 2363.324 2274.354 2275.531
print(mses)
                        > print(mses)
                        [1] 23.94366 18.98477 18.94499
# 三个模型fit结果图示
plot1<- plot(x=Auto$horsepower,y=Auto$mpg)+
  points(x=Auto$horsepower,y=qlm1$fitted.values,col="blue")+
  points(x=Auto$horsepower,y=glm2$fitted.values,col="red")+
  points(x=Auto$horsepower,y=glm3$fitted.values,col="cyan")
print(plot1)
```

根据散点图,mpg与horsepower 的二次项拟合良好!



谢谢!