

How Do Agents Form Inflation Expectations? Evidence from the Forecast Uncertainty

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Abstract

Density forecasts of macroeconomic variables provide one additional moment restriction, uncertainty, for testing various models of expectation formation deviating from the full-information rational expectations (FIRE). This paper first documents the persistent dispersion and inefficient revisions in inflation uncertainty of professionals and households, and shows how the uncertainty conveys different information from the widely used proxies to uncertainty such as cross-sectional disagreement and forecast errors. Second, I provide additional reduced-form test results and structural estimates for three models deviating from FIRE: sticky expectations (SE), noisy information (NI), and diagnostic expectations (DE) by jointly accounting for its predictions for different moments. The estimates suggest the SE model with an inflation process featuring stochastic volatility accounts for the joint dynamics of inflation and forecast moments better than other competing theories.

Keywords: Inflation, Expectation Formation, Rigidity, Overreaction, Uncertainty, Density Forecast

JEL Codes: D83, E31, E70

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1 Introduction

Theories on how agents form expectations in ways deviating from rational expectations (RE) have been proliferating over the past decade. On one hand, various theories built upon different micro-foundations produce somewhat similar macro patterns. For instance, the information rigidity documented by [Coibion and Gorodnichenko \(2012, 2015\)](#), i.e. the sluggish response in aggregate expectations to new information, can be micro-founded by both Sticky Expectations (SE)¹ and Noisy Information (NI)². On the other hand, there are important and subtle differences in testable predictions from various theories regarding both individual forecasts and aggregate moments. For instance, in contrast with the models featuring information rigidity at the aggregate level, the theory of Diagnostic Expectation (DE) ([Bordalo et al., 2018, 2020](#)) implies overreaction to news at the individual level.³

One of the crucial steps ahead in advancing this literature is to test these theories better using expectation surveys. This paper does a cross-moment estimation of each theory by jointly accounting for its predictions about different moments of the forecasts such as forecast errors (FE), cross-sectional disagreements (Disg), and *uncertainty* (Var). Although reduced-form tests focused on first moments have been sufficient to rejecting the null of the full-information rational expectations (FIRE), identifying the differences among these non-FIRE theories requires more information from second or higher-order moments and needs to rely on more restrictions across moments. Such an idea builds on the seminal work [Coibion and Gorodnichenko \(2012\)](#), which compares different theories by examining if the impulse responses of inflation expectations to externally estimated shocks align with the qualitative predictions of these models. This paper adopts a similar approach,⁴ with the key novelty being utilizing an additional moment of forecasts in the estimation, i.e. the uncertainty.

Existing work that studies expectation formation based on survey data mostly focuses on the individual patterns of the mean forecast, forecast errors, and, to the most, cross-sectional disagreements. But there are additional insights from the forecast uncertainty, which is only available if each survey respondent is asked to assign their own perceived probabilities to a range of values of the variable. Meanwhile, it is worth emphasizing that beyond the usefulness in identifying models of expectation formation, understanding the exact formation mechanism of inflation uncertainty per se, in addition to expectations, is important since inflation uncertainty has both microeconomic implications via precautionary saving motives and portfolio investments, as well as macroeconomic implications on inflation dynamics and asset prices.

Two patterns of the observed uncertainty are in violation of benchmark full-information rational expectation (FIRE). First, the persistent cross-sectional dispersion of uncertainty across individual forecasters (Figure 2 and 3), like that in the average forecasts, is inconsistent with the FIRE as the latter assumes that agents agree on the data

¹[Mankiw and Reis \(2002\)](#); [Carroll \(2003\)](#); [Reis \(2006\)](#).

²[Lucas \(1972\)](#), [Woodford \(2001\)](#).

³Recently, [Kohlhas and Walther \(2021\)](#) proposes an extended model of NI allowing for multiple unobserved components that reconciles the coexistence of underreaction and overreaction.

⁴Some other contemporaneous papers also structurally estimate theories on expectation formation based on single or multiple moments of surveyed expectations, such as [Giacomini et al. \(2020\)](#); [Xie \(2019\)](#); [Farmer et al. \(2021\)](#).

generating process and have the common knowledge of available information. Second, dynamically, across different vintages of the forecast, the revision in uncertainty is a measure of information gain or the degree of forecasting efficiency (Table 4). FIRE implies that information gain is sufficiently large to match the volatility of realized shocks, but information rigidity implies inefficient revisions in uncertainty.

A deeper insight in this paper is that various models of expectation formation have distinctive parameter restrictions about the uncertainty and its relationship with other moments such as forecast errors and disagreement. Utilizing information from uncertainty helps identify model parameters, especially when the commonly used moments are not sufficient. In addition, even for competing theories that have similar qualitative patterns, a cross-moment structural estimation could be useful in evaluating if the indirectly implied model parameters from the survey data are empirically realistic.

In the theoretical front, this paper contributes to the literature of expectation formation by explicitly deriving predictions of uncertainty of three non-FIRE workhorse models of expectation formation and one hybrid version seen in the literature: (1) Sticky Expectations (SE); (2) Noisy Information (NI); (3) Diagnostic Expectations (DE); (4) Diagnostic Expectations/Noisy Information (DENI).⁵

FIRE benchmark predicts the ex-ante uncertainty to be exactly equal to the variance of ex-post forecast errors, and to the size of the conditional volatility of inflation. But other models imply different patterns. For instance, in SE, the extra uncertainty arises compared to in FIRE because of lagged updating of the most recent information, hence a slower resolution of uncertainty. It also predicts a greater uncertainty than the size of ex-post forecast errors (a pattern observed in both SCE and SPF). In NI, the uncertainty endogenously depends on the noisiness of signals and determines agents' degree of reaction to the news in the Kalman filtering problem. The model accommodates a flexible relative size of the ex-ante uncertainty and ex-post forecast errors depending on the size of noisiness of signals. Different from the both, the canonical model of DE assumes the uncertainty to be equal to that in FIRE.

These predictions allow for an empirical parameter identification of each model using survey data. The headline finding from doing such is that using additional information from surveyed inflation uncertainty further confirms the previous empirical finding of information rigidity using low-order moments in the literature (Coibion and Gorodnichenko (2012, 2015); Coibion et al. (2018)). This means qualitatively, SE and NI outperform the DE and DENI in capturing the rigidity in the aggregate forecasting moments, including that in uncertainty. Note, however, that the estimates of DE do suggest a coexistent overreactive mechanism at the individual level, i.e. a non-zero fraction of agents in the economy have a positive degree of overreaction $\hat{\theta}$. This is consistent with the finding in the literature showing the coexistence of rigidity and overreaction (Angeletos et al. (2021); Kohlhas and Walther (2021)).

The key real-value-added by surveyed uncertainty turns out to be differentiating two rigidity models, SE and NI. Despite the seemingly similar qualitative patterns, SE outperforms NI in yielding more sensible and consistent model parameters. In particular, the estimated updating rates λ are consistently around one third (meaning one third fraction of population update each quarter for both households and profession-

⁵For instance, Bordalo et al. (2020) embeds DE in a NI model to jointly account for the heterogeneous and extrapolating expectations.

als), while in contrast, the estimated noisiness of public and private signals σ_ϵ and σ_ξ in NI are unrealistic high and unstable, i.e. at a ballpark value of 3 percentage points or higher. These are arguably too high given an unconditional standard deviation of inflation of 0.8 (for headline CPI) or 0.4 for (Core PCE) in the sample period.

With comparable estimates of different theories using cross-moments restrictions, one can evaluate the model sensitivity of each model of expectations formation in four dimensions: (1) the moments used for estimation, i.e. only the forecast error or higher moments such as disagreement and uncertainty. (2) the specification of the underlying process of the inflation, i.e. an AR(1) with constant volatility or one with different components of the time-varying volatility as that in [Stock and Watson \(2007\)](#). (3) if estimating the underlying process and expectations separately or jointly. The former basically recovers the inflation process only based on inflation data, while the latter let the expectations provide information for estimating the inflation process. (4) whether it accounts for different agents such as households and forecasters' expectations equally well.⁶

Some key findings from the model sensitivity evaluation are as follows, which are also summarized as a scoring card in Table 10.

- **Overall robustness** Sticky expectations (SE) produces the most stable parameter estimate in accounting for the observed patterns of *aggregate* expectations and inflation among competing theories, for both households and professionals, regardless of targeted moments, and for both AR(1) and SV model of inflation.
- **Targeted moments** The parameter estimates of both SE and NI are insensitive to the use of moments in estimation, while in contrast, the estimates of DE and DENI vary greatly depending on the moments used.
- **Mutual-dependence of expectation formation and inflation process** All the theories are somewhat sensitive to whether jointly estimate the inflation process and the survey moments, or separately estimate the two. This reflects the mutual dependence of the underlying inflation process and the degree of deviation from FIRE postulated by a particular theory. For instance, a higher jointly estimated persistence of inflation ρ always come with a lower estimate of updating rate λ (higher rigidity) in SE, and a smaller degree of overreaction $\hat{\theta}$ in DE.⁷ Intuitively speaking, a great degree of reaction in inflation expectations to shocks may come from either a FIRE response to a highly persistent shock, or a lack of information rigidity (or, on the flip side, overreaction).
- **The role of stochastic volatility** SE estimate is most robust to an alternative process of SV. In contrast, DE and DENI are particularly sensitive to alternating the inflation process. Across all theories, however, allowing for two unobserved components and SV significantly improves within-model consistency of each theory. Most notably, modeling NI as a signal-extraction problem of unobserved component as in SV produces much more sensible and stable estimates of noisiness of signals.

⁶[Cornand and Hubert \(2022\)](#) documents systematic differences in forecasts of various types of agents in terms of revision frequency and size of cross-sectional disagreement.

⁷Such interdependence is also discussed in [Afrouzi et al. \(2020\)](#) in the context of DE models.

- **The type of agents** Within each theory, households’ expectations exhibit less consistency across moments compared to professionals. This evidence of inconsistency adds to the evidence rejecting full-information-rationality based on reduced-form tests.

Finally, in addition to the structural estimation with survey data, I also provide additional reduced-form evidence in violation of the benchmark FIRE predictions using the revision of uncertainty. Like autoregression tests of revisions in forecasts, revision in uncertainty can also be used to test the efficiency in expectation formation. The new result from this paper is that the revision in uncertainty has a serial correlation that is not consistent with the level of forecast efficiency predicted by rational expectation. This provides an additional result that rejects the null of the rational expectation hypothesis.

1.1 Related Literature

This paper is related to four strands of literature. First, it is related to a series of empirical work directly testing and evaluating various theories on expectation formation using survey data. For instance, [Mankiw et al. \(2003\)](#), [Carroll \(2003\)](#), [Branch \(2004\)](#). More recent examples include [Coibion and Gorodnichenko \(2012, 2015\)](#); [Coibion et al. \(2018\)](#) that test common implications of various theories with different micro-foundations. In addition to testing particular sets of theories, there are also a number of papers that show people’s expectations are driven by individual heterogeneity such as socioeconomic characteristics, cognitive abilities, experiences of macroeconomic histories ([Malmendier and Nagel \(2015\)](#), [Das et al. \(2017\)](#) and [D’Acunto et al. \(2019\)](#)⁸). In terms of the methodology, this paper is closest to [Giacomini et al. \(2020\)](#), which estimates theories of expectation formation using cross-moment restrictions. However, all of these studies simply rely upon point forecasts instead of density forecast or surveyed uncertainty. This is one theme on which this paper differs from the existing literature.

Second, [Manski \(2004\)](#), [Delavande et al. \(2011\)](#), [Manski \(2018\)](#) and many other papers have advocated long for eliciting probabilistic questions measuring subjective uncertainty in economic surveys. Although the initial suspicion concerning to people’s ability in understanding, using and answering probabilistic questions is understandable, [Bertrand and Mullainathan \(2001\)](#) and other work have shown respondents have the consistent ability and willingness to assign a probability (or “percent chance”) to future events. [Armantier et al. \(2017\)](#) have a thorough discussion on designing, experimenting and implementing the consumer expectation surveys to ensure the quality of the responses.⁹ Broadly speaking, the literature has argued that going beyond the revealed preference approach, availability of survey data provides economists with direct information on agents’ expectations and helps avoids imposing arbitrary assumptions. This insight holds for not only point forecast but also and even more importantly, for uncertainty, because for any economic decision made by a risk-averse agent, not only the expectation but also the perceived risks matter a great deal.

⁸See [D’Acunto et al. \(2023\)](#) for a thorough survey of the empirical evidence of heterogeneous inflation expectations and their drivers.

⁹Other literature includes [Van der Klaauw et al. \(2008\)](#) and [Delavande \(2014\)](#), etc.

Third, by approximating subjective uncertainty directly from density responses, this paper contributes to the literature that develops and uses a variety of measures of uncertainty, especially in the macroeconomic context. There is a long tradition of approximating uncertainty by measures that are more directly available in survey data or that can be estimated by econometric methods. For instance, [Bachmann et al. \(2013\)](#) use ex-ante disagreement and ex-post forecast errors computed from forecasters' surveys as proxies of uncertainty. [Jurado et al. \(2015\)](#) define the time-varying uncertainty as conditional volatility of the non-forecastable component of a variable, and estimate it using multiple macroeconomic series. [Binder \(2017\)](#) approximate uncertainty from rounding in survey data based on the insights from cognitive literature. Besides, the text-based approach such as [Bloom \(2009\)](#) constructs indices of policy uncertainty based on texts of newspaper reporting. Although these proxies are all meant to capture the notion of uncertainty, as shown in Section 2.4, cross-validation seems to suggest they are statistically uncorrelated or even negatively correlated.

Fourth, the literature that has been originally developed under the theme of forecast efficiency¹⁰ provides a framework analyzing the dynamics of uncertainty useful for the purpose of this paper. The focus of the forecasting efficiency literature is evaluating forecasters' performance and improving forecasting methodology, but it can be adapted to test the theories of expectation formation of different types of agents. This is especially relevant to this paper as I focus on the uncertainty.

The paper is organized as followed. Section 2 shows the stylized patterns of different forecasting moments of professional forecasts and households. Section 3 first sets up a unified framework in which testable predictions of different theories can be compared. Also, I derive a various moment conditions from these theories. Section 4 undertakes reduced-form time-series regressions that test the null hypothesis of FIRE and implications of different theories. Section 5 includes results from estimating the theory-specific parameters using simulated method of moments. It also evaluates the sensitivity of the model specification. Section 6 concludes the paper and discusses the future research directions.

2 Data and Facts

2.1 Definition and notation

An agent i is forming expectations about a stochastic macroeconomic variable y_{t+h} ¹¹, the inflation in this paper. Denote $f_{i,t+h|t}$ as agent i 's h -period-ahead density forecast. $f_{i,t+h|t}$ is the conditional density of y_{t+h} given the information set $I_{i,t}$ available at time t .

$$f_{i,t+h|t} \equiv f_{i,t}(y_{t+h}|I_{i,t})$$

¹⁰[Nordhaus \(1987\)](#), [Davies and Lahiri \(1995\)](#), [Clements \(1997\)](#), [Faust and Wright \(2008\)](#), [Patton and Timmermann \(2012\)](#).

¹¹Only in the context of aggregate variable, it makes sense to study the population moments such as average expectations and disagreements. Studying expectations of idiosyncratic variables requires individual panel data, as well as the idiosyncratic realizations of the variable.

The information set could be agent-specific, thus it has subscript i . The specific content contained in I_t varies from different models of expectation. For instance, sticky expectation (SE) and rational inattention¹² literature all assume that agents are not able to update new information instantaneously. So the information set may not contain the most recent realization of the variable of forecast y_t . In contrast, NI assumes that the information set only contains noisy signals of the underlying variables. Different theories may also differ in terms of the mapping from information to conditional density forecasts.¹³ For instance, DE deviates from Bayesian learning by allowing agents to overweight new information that is particularly salient.

Accordingly, h-period-ahead mean forecast at t , denoted as $y_{i,t+h|t}$, is the conditional expectation of y_{t+h} by the agent i .

$$y_{i,t+h|t} \equiv E_{i,t}(y_{t+h}) = \int y_{t+h} f_{i,t+h|t} dy_{t+h}$$

Similarly, individual forecasting variance $Var_{i,t+h|t}$, hereafter termed as individual uncertainty in this paper, is the conditional variance corresponding to the forecast density distribution.

Individual forecast error $FE_{i,t+h|t}$ is the difference of individual forecast at time t and ex post realized value of y_{t+h} . By definition, positive (negative) forecast errors mean overpredict (underpredict) the variables.

$$FE_{i,t+h|t} = y_{i,t+h|t} - y_{t+h}$$

The population analogs of individual mean forecast, uncertainty and forecast errors are simply the average of the individual moments taken across agents. Denote them as $\bar{y}_{t+h|t}$, $\overline{Var}_{t+h|t}$, and $\overline{FE}_{t+h|t}$, respectively. Hereafter, they are termed as the population mean forecast, population uncertainty and population forecast error, respectively. In addition, disagreement is defined as the cross-sectional variance of mean forecasts of individual agents, denoted as $\overline{Disg}_{t+h|t}$.

I refer to the 3 individual indicators and 4 population indicators defined above as moments, and they are listed in Table 1.

Table 1: Definition and Notation of Moments

Individual Moments	Population Moments
Mean forecast: $y_{i,t+h t}$	Average forecast: $\bar{y}_{t+h t}$
Forecast error: $FE_{i,t+h t}$	Average forecast error: $\overline{FE}_{t+h t}$
Uncertainty: $Var_{i,t+h t}$	Average uncertainty: $\overline{Var}_{t+h t}$
	Disagreements: $\overline{Disg}_{t+h t}$

¹²Sims (2003).

¹³There are other classes of models that fall into this category, which assume alternative mapping from the agent's information set to the conditional density. For instance, Patton and Timmermann (2010); Farmer et al. (2021) find that the disagreements are driven by not only the difference in information, but also heterogeneity in prior and models. Macaulay and Moberly (2022) shows survey evidence for heterogeneity in perceived persistence of inflation shocks. More theoretical work includes multi-prior or model uncertainty such as Hansen and Sargent (2001), Hansen and Sargent (2008), etc.

2.2 Benchmark predictions from full-information rational expectation (FIRE)

We start by assuming an underlying process of inflation. In the benchmark scenario, we assume the underlying true process of y_t is $AR(1)$ with persistence parameter $0 < \rho < 1$ and i.i.d. shock ω_t .

$$y_t = \rho y_{t-1} + \omega_t, \quad \omega_t \sim N(0, \sigma_\omega^2) \quad (1)$$

The assumption of i.i.d. shock with constant volatility turns out insufficient to match the time-varying pattern of uncertainty. Therefore, as an extension, I also consider an alternative assumption that inflation consists of two unobservable components with stochastic volatility. I will specify this explicitly later to focus on comparing different models of expectation formation under a simple $AR(1)$ process.

In the FIRE benchmark, it is assumed that all agents perfectly observe y_t at time t and understand the true process of y . Therefore, the individual forecast is $\rho^h y_t$, which is shared by all agents. Therefore, it is also equal to the average forecast.

Both individual and population forecast errors are simply the realized shocks between $t + 1$ and $t + h$.

$$FE_{t+h|t}^* = - \sum_{s=1}^h \rho^{s-1} \omega_{t+h-s} \quad (2)$$

I use superscript of $*$ to denote all the moments according to FIRE. It is easy to see that the forecast error is orthogonal to information available till time t . This provides a well known null hypothesis of FIRE.¹⁴ These FE-based restrictions of FIRE provide the foundations for the tests used in Section 6.

The unconditional variance of the FE, or equivalently, its square (due to its zero unconditional mean), is equal to the following.

$$FE_{\bullet+h|\bullet}^{*2} = \sum_{s=1}^h \rho^{2(s-1)} \sigma_\omega^2 \quad (3)$$

The uncertainty about future y simply comes from uncertainty about unrealized shocks between t and $t + h$. With the same model in mind (Equation 1) and the same information y_t , everyone's uncertainty is equal to the weighted sum of the future volatility before its realization (Equation 4). In FIRE, there is neither disagreement about the mean, nor disagreement about the uncertainty.¹⁵ Furthermore, it is exactly equal to the variance of forecast errors, $FE_{\bullet+h|\bullet}^{*2}$.

¹⁴Another well-known prediction of FIRE is that forecast errors of non-overlapping horizon are not correlated, namely, $Cov(FE_{t+h|t}^*, FE_{t+s+h|t+s}^*) = 0 \quad \forall s \geq h$. This is not the case within h periods, as the realized shocks in overlapping periods enter both forecast errors.

¹⁵This is the same to Jurado et al. (2015)'s terminology.

$$\text{Var}_{\bullet+h|\bullet}^* = \sum_{s=1}^h \rho^{2(s-1)} \sigma_{\omega}^2 \quad (4)$$

The time-series behavior of h-period-ahead uncertainty, i.e. $\text{Var}_{t+h|t}$, $\text{Var}_{t+h+1|t+1}$, etc., depends on the true process of y . Specifically, it depends on whether σ_{ω}^2 is time-varying. If time-invariant, h-period-ahead uncertainty is simply a constant. In the baseline case, I make such an assumption, in general, it may not be true. In the extension, I make alternative assumptions of the inflation process allowing for stochastic volatility.¹⁶

Another testable implication of rationality lies in the revision of uncertainty. Hereafter, we refer to revision (instead of change) as the difference of moments across vintages of the forecast with the fixed terminate date of realization. For instance, the uncertainty revision for from $t - 1$ to t the h-period-ahead forecast is the difference between the uncertainty about y_{t+h} at time t and the uncertainty about y_{t+h} at time $t - 1$.

Moving from t to $t + 1$, for instance, the revision in uncertainty is simply a negative constant independent of the time. There is an unambiguous reduction in uncertainty (or information gain in the forecasting literature) as more and more shocks have realized. In the most intuitive case, from the one-step-ahead forecast at $t - 1$ to that of t , i.e. $h = 1$, the variance drops exactly by the resolution of uncertainty of ω_t , which is σ_{ω}^2 , to zero uncertainty.

$$\text{Var}_{t+h|t+1}^* - \text{Var}_{t+h|t}^* = -\rho^{2(h-1)} \sigma_{\omega}^2 \quad (5)$$

Lastly, FIRE has predictions about the disagreement. As agents perfect update the same information, there is no disagreement at any point of the time.

$$\text{Disg}_{t+h|t}^* = 0 \quad \forall t \quad (6)$$

Table 2 summarizes all other expectation moments of 1-period-ahead inflation ($h = 1$) as predicted by FIRE and a process of AR1. Both variances of forecast errors ($FEVar$) and average uncertainty Var are equal to the size of the shock to inflation σ_{ω}^2 , while in contrast, the disagreement is always zero, hence has zero correlation with $FEVar$ and Var . In Section 2.4, I discuss in greater detail how the data counterparts of the forecast moments are inconsistent with these predictions.

2.3 Data

This paper utilizes one special feature of professional and household inflation expectation survey to test various theories of expectation formation, which is that both surveys contain density forecasts by asking individual respondents to fill a number of externally specified bins representing different ranges of values.

¹⁶For example, [Justiniano and Primiceri \(2008\)](#), [Vavra \(2013\)](#) on time-varying volatility of inflation.

Table 2: Moments of Inflation and Expectations

	SPF	SCE	FIRE+AR	FIRE+SV
InfAV	0	0	0	0
InfVar	0.159	0.653	$\sigma_\omega^2/(1 - \rho^2)$	N/A
InfATV	0.125	0.621	$\rho\sigma_\omega^2/(1 - \rho^2)$	N/A
FE	0.136	1.772	0	0
FEVar	0.133	0.923	σ_ω^2	$\bar{\sigma}_\eta^2 + \bar{\sigma}_z^2$
FEATV	0.097	0.89	0	0
Disg	0.183	2.585	0	0
DisgVar	0.028	0.057	0	0
DisgATV	0.021	0.025	0	0
Var	0.242	1.75	σ_ω^2	$\bar{\sigma}_\eta^2 + \bar{\sigma}_z^2$
VarVar	0.001	0.023	0	>0
VarATV	0.001	0.004	0	>0

This table reports the moments of demeaned inflation and inflation expectations of both SPF and SCE used in the model estimation. Core CPI inflation is used for SPF, and the headline inflation is used for SCE. The sample period is 2007M1-2020M3 for SPF and 2013M1-2020M3 for SCE. For SCE moments, both disagreement (Disg) and uncertainty (Var) are computed using the regression residuals of individual mean forecast and uncertainty after controlling for individual fixed effects.

Survey of Professional Forecasters(SPF) has reported the individual-level density forecasts of core CPI and core PCE inflation since 2007. In addition, for these two inflation measures, both forecasts for current-year inflation, basically nowcast, and one-year-ahead forecasts are elicited via density surveys. This makes it possible to directly test the implications of the revisions in uncertainty across different vintages of forecasts.

The New York Fed Survey of Consumer Expectation(SCE), which started in 2013 also asked households to provide their perceived probabilities about 1-year-ahead and 3-year-ahead inflation for various ranges of values each month¹⁷. This allows for comparing 3-year-ahead forecast at time $t - 3$ with 1-year-ahead forecast at $t - 1$. Since the maximum duration for households to stay in the panel is 12 months (for about one-third of the households), forecast revision can be only examined at the population level. The advantage of SCE compared to SPF is its monthly frequency. This provides an invaluable chance to explore the dynamics of uncertainty. A summary of the two surveys is in Table 3.

Converting expressed probability forecasts based on externally divided bins into an underlying subjective distribution requires a density distribution. I closely follow Engelberg et al. (2009)'s method with a small modification to estimate the density distribution of each individual respondents in SPF. Answers with at least 3 bins with positive probabilities or 2 bins but open-ended from either left or right are fit with a generalized beta distribution. Depending on if there is open-ended bin on either side

¹⁷Most importantly, the survey respondents are guaranteed to assign probabilities to all bins that sum up to one, as a feature of the online survey design.

with positive probability, 2-parameter or 4-parameter beta distribution are estimated, respectively. Those with only two bins with positive probabilities and adjacent are fit with a triangular distribution. Answers with only one bin of positive probability are fit by a uniform distribution.¹⁸ This is the same approach adopted by the New York Fed researchers [Armantier et al. \(2017\)](#).

To avoid the biases introduced by extreme answers or data errors, I undertake some winsorization for both data sets. For SPF, I drop the outliers of mean forecast and uncertainty estimates at both the top and bottom one percentile, as these are typically abnormal that are due to measurement errors or other reasons. For SCE, I drop the top and bottom 5 percent of mean forecasts and uncertainty, as households' mean forecasts are inclined to give extreme values. All the results in this paper are robust to different thresholds, such as 10 and 1 percentile.¹⁹

Table 3: Information of Data

	SCE	SPF
Time period	2013-2021M7	2007-2022Q2
Frequency	Monthly	Quarterly
Sample Size	1,300	30-50
Var in Density	1-yr inflation	1-yr-ahead Core CPI and Core PCE
Panel Structure	stay up to 12 months	average stay for 5 years
Individual Info	Education, Income, Age, Location	Industry

Throughout the paper, I use three measures of inflation: headline CPI, core CPI, and core PCE. Depending on the specific variable of forecast in the survey series, the realization of the corresponding inflation is used to compute moments such as forecast errors. Specifically, SPF has density forecasts for both core CPI and core PCE²⁰. For SCE, as the household responders are asked about the overall inflation, it is the most reasonable to interpret it as headline CPI inflation. To simplify the expression, from now on, core CPI and core PCE are simply referred to as CPI and PCE, respectively.

2.4 Stylized facts

2.4.1 Relationship between the size of forecast errors, disagreement and uncertainty

Despite the significant differences in magnitudes between professionals' forecasting moments and households, both types of agents share common patterns in terms of the relationship across various moments. Figure [1a](#), [1b](#), and [1c](#) plot the population

¹⁸See the [Python program](#) with detailed steps of estimation.

¹⁹For mean forecasts and uncertainty, respectively, this means dropping 6528 and 5096 observations, out of 68887 observations in total.

²⁰SPF also has density forecasts for GDP deflator (for GNP prior to 1992: Q1) going back to 1968. Since the ranges of the values and the definition are not consistent over time, I do not use them in this paper.

uncertainty against expected inflation, forecast errors, and disagreements in the first, second, and third rows, respectively.

It is both documented in the literature ²¹ and cited in popular narratives that high inflation (or expected inflation) is typically associated with high inflation uncertainty. The observable correlation between the realized inflation and the directly measured forecasting uncertainty by both professionals and households is indeed consistent with such a pattern, with such a positive correlation primarily driven by co-movement of the two since the Covid pandemic. Over the entire period between 2007-2021, the correlation coefficients are 0.28 for SPF’s Core CPI forecast, and 0.58 for SCE’s CPI expectation, respectively. But it’s worth noting that such comovement was primarily driven by the dynamics since 2020 with the uprising inflation. ²²

Figure 1b inspects the relationship between the size of the forecast error and uncertainty. According to the benchmark prediction under FIRE, the ex-ante forecast uncertainty is equal to the variance of ex-post forecast errors on average, as shown in Table 2. In the data, the correlation coefficients of the two are 0.25, and 0.42 for SPF Core CPI forecasts and SCE’s forecasts, respectively. Excluding the post-2020 sample yields even smaller correlation coefficients. The fact that the two are at mostly weakly correlated is inconsistent with the FIRE benchmark prediction.

Figure 1c examines the relationship between disagreement and uncertainty. A large body of empirical literature in macroeconomics use disagreement, which is often more available than the uncertainty, as a proxy of latter. ²³ This practice implicitly assumes some form of deviations from the benchmark FIRE, as in FIRE, regardless of the inflation process, disagreement should be always zero, and it is therefore not correlated with the average uncertainty. ²⁴ The empirical correlation between disagreement and uncertainty is indeed weakly positive, which are 0.33 for professionals, and 0.78 for households. The positive correlation between the two was primarily driven by the post-2020 dynamics, which was a period of rapidly rising inflation.

To summarize, compared to the FIRE prediction of a perfect correlation between ex-ante uncertainty and the size of ex-post forecast error in the benchmark FIRE, the empirical patterns of professionals and household expectations exhibit divergent and time-varying behaviors of the two.

2.4.2 Dispersion in forecast uncertainty

A persistent disagreement in expectations has been used as important stylized evidence inconsistent with the assumption of identical expectation embedded in FIRE (Mankiw et al. (2003)). A similar argument can be made with the dispersion in forecasting uncertainty. FIRE predicts individuals share an equal degree of uncertainty. In contrast, SE predicts that the uncertainty of individuals differs in that agents are not equally updated at a point of the time (Equation 8). NI generates a homogeneous

²¹For example Ball et al. (1990); Ball (1992).

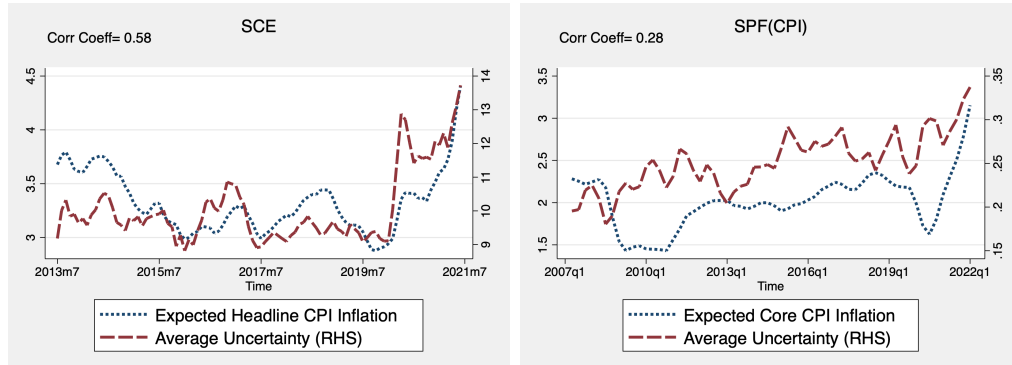
²²The correlation coefficients before 2020 were 0.12 and -0.24, respectively.

²³For instance, Bachmann et al. (2013) used ex-ante disagreement and ex-post forecast errors as two measures of uncertainty and find that both uncertainty indicators lead to a reduction in real economic activity.

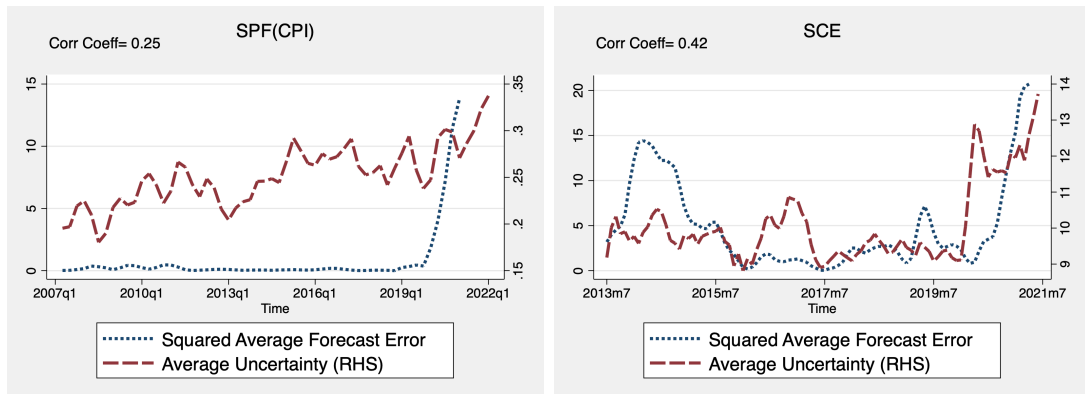
²⁴This point was made very clearly by Zarnowitz and Lambros (1987). Manski (2018) also points out that much empirical work has confused the dispersion with uncertainty.

Figure 1: Uncertainty and Other Moments

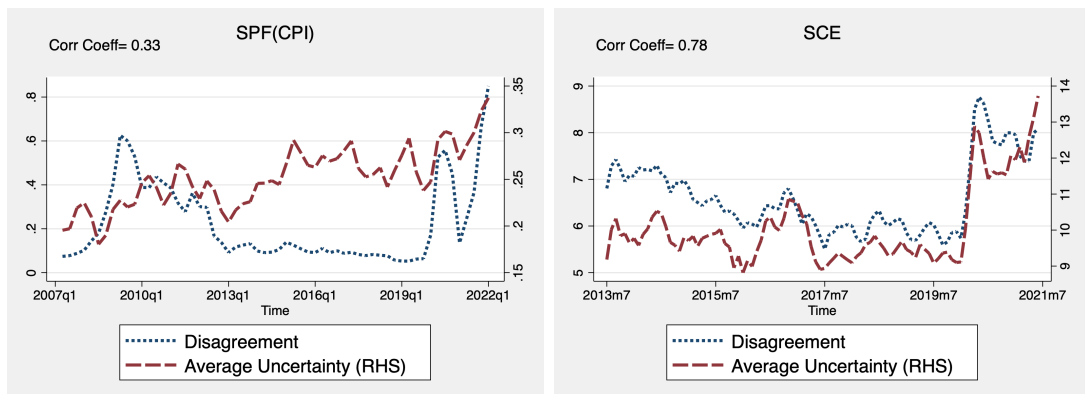
(a) Expected Inflation and Uncertainty



(b) Squared Forecast Errors and Uncertainty



(c) Disagreement and Uncertainty

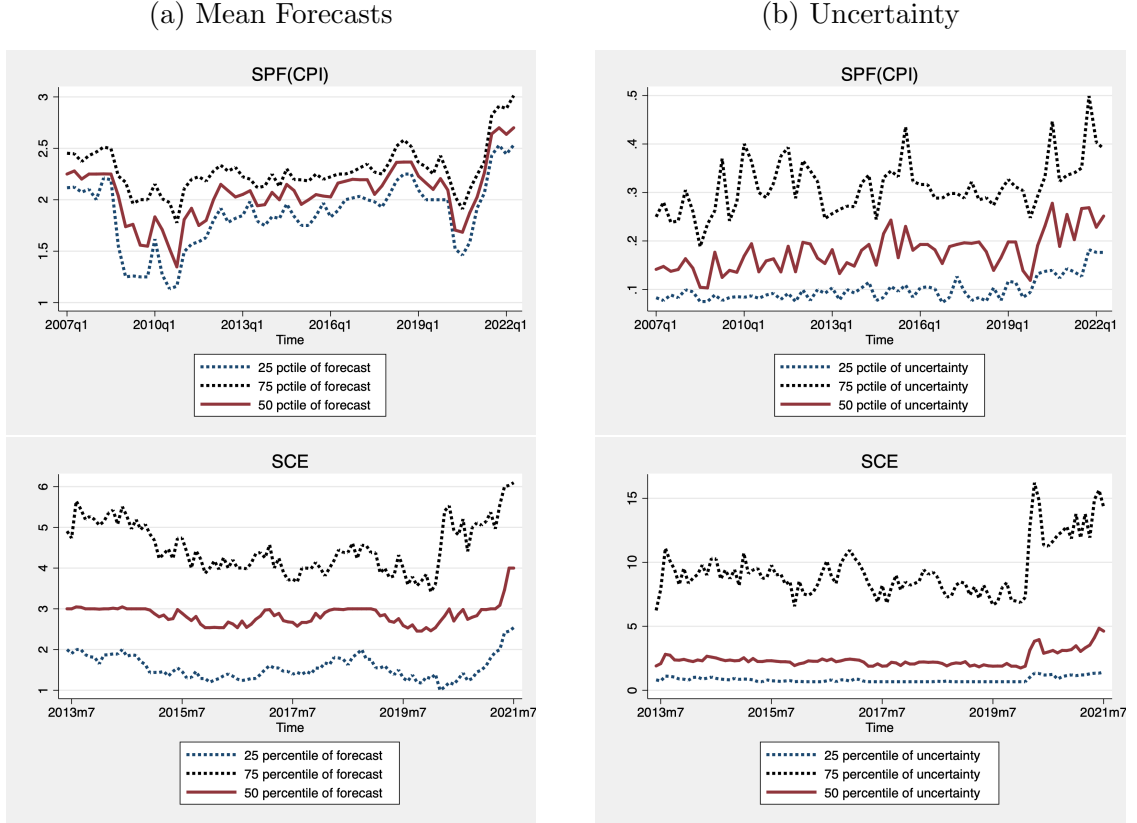


Note: From the left to right: SPF's forecasts of core CPI and SCE's household forecast of headline CPI. From the top to the bottom: uncertainty (dash line) versus expected inflation (dot) with correlation coefficient of 0.58, and 0.28, respectively; uncertainty (dash) versus square of the realized forecast errors (dot) with correlation coefficient of 0.25, 0.42, respectively; uncertainty (dash) versus disagreements (dot) with correlation coefficient of 0.33, and 0.78, respectively.

degree of uncertainty only under the stringent conditions of an equal precision of signals and the same prior for uncertainty (Equation 25). DE predicts an equal degree of uncertainty across agents (Equation 30). Therefore, taken by the safe value, the presence of dispersion of uncertainty across agents is not consistent with predictions of FIRE, the canonical version of NI and DE.

Figure 2 plots the median inflation expectation along with its 25/75 percentiles in the left and its counterpart in uncertainty in the right column. Not only there is long-lasting dispersion in individual forecasts, i.e. disagreement, but also notable heterogeneity in uncertainty across agents. And not surprisingly, the dispersion of both forecasts and the uncertainty of households are both of a much greater magnitude than that of the professionals. The 25/75 interquartile range (IQR) of households point forecasts is 4-5 percentage points, compared to 1 percentage point of professionals. And the IQR of the uncertainty of households is around 150-200 times(12-14 times in standard deviation terms) of that of professional forecasters.

Figure 2: Dispersion of Mean Forecasts and Uncertainty

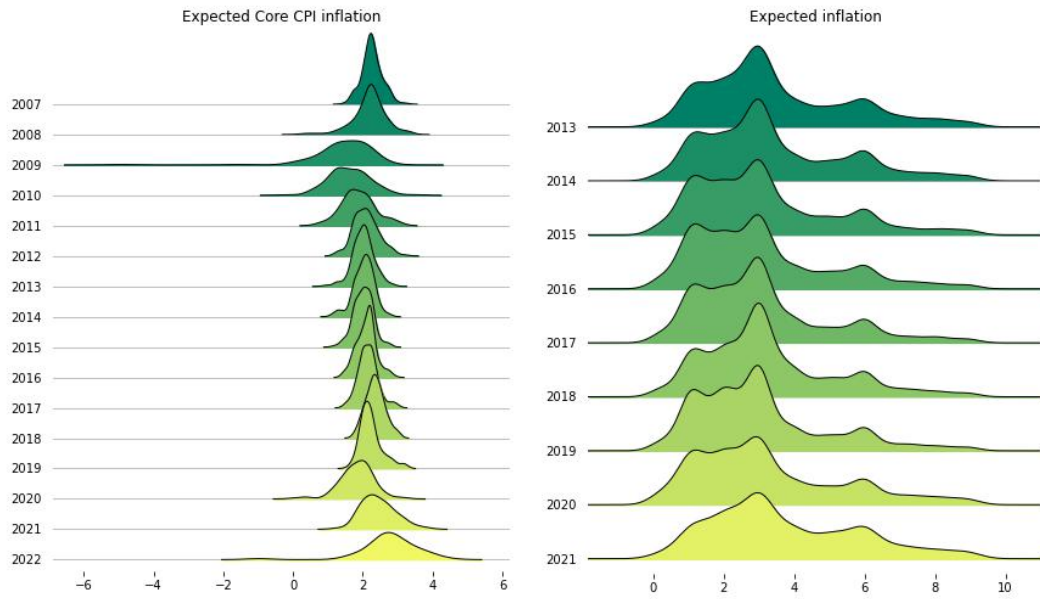


Besides, in terms of the distribution of uncertainty, there are differences and similarities between households and professionals. Households' uncertainty is more skewed toward the right (higher uncertainty), meaning there is a wide dispersion in the high values of uncertainty. This can be also seen in Figure 3, where I plot the kernel density estimated distribution of uncertainty over the sample period. ²⁵

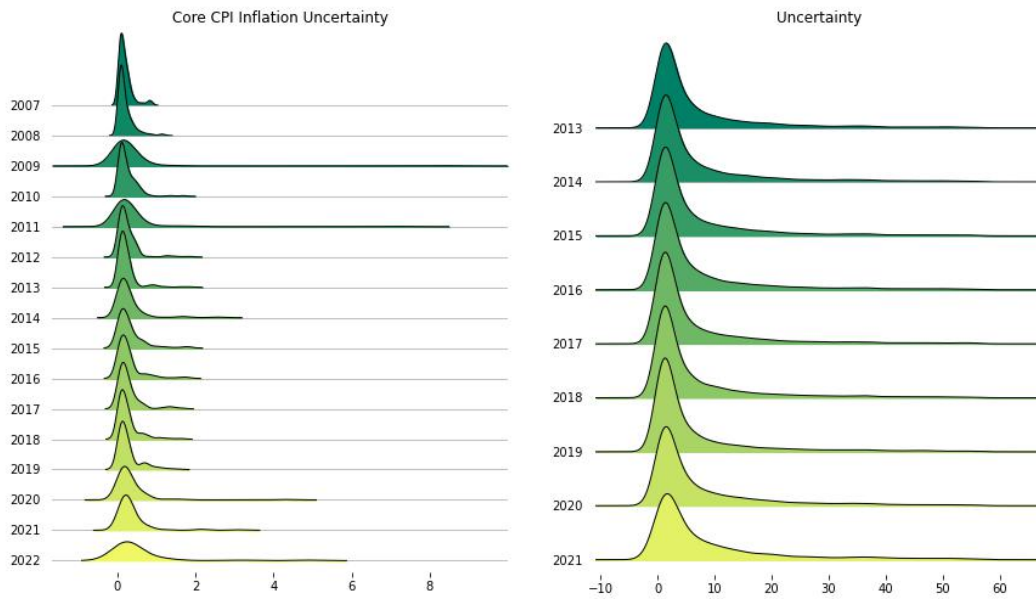
²⁵Kumar et al. (2015) also presents the dispersion in uncertainty using a shorter-period of sample for SCE.

Figure 3: Distribution of Mean Forecast and Uncertainty

(a) Mean Forecasts



(b) Uncertainty



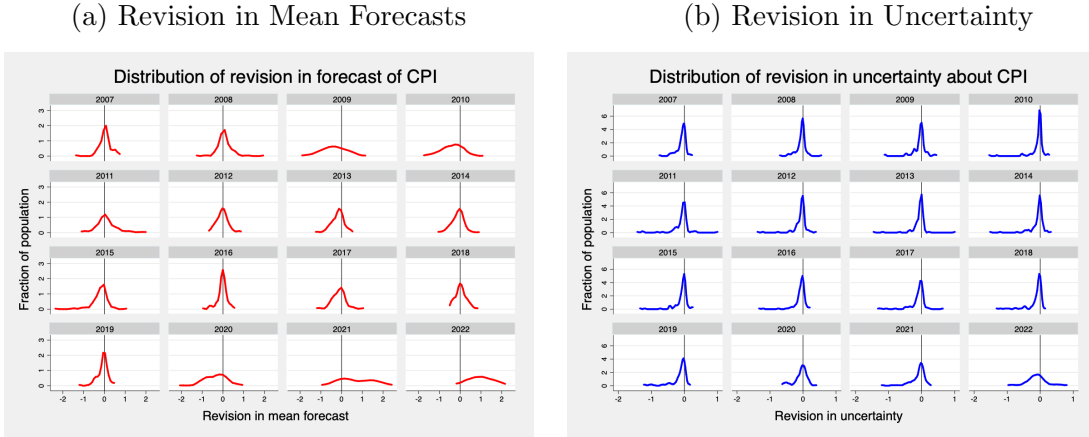
Another pattern worth discussing in Figure 2 is that there is a notable rise in the dispersion of professional forecasts in post-2020 period with a significant rising in inflation, primarily driven by an increase of upper side of the forecast (i.e. 75 percentile forecast increases from 1% to 2%).²⁶ This is consistent with the observation in the top left two graphs of Figure 3 that the distribution of inflation forecasts in recent years have become flattened.

²⁶This should be interpreted with caution, since the disagreements of SPF forecasts shown in Figure 1c actually exhibits a gradual decline.

2.4.3 Revision in uncertainty

FIRE also predicts an unambiguous reduction in uncertainty as one approaches the date of realization, where the drop is exactly equal to the volatility of the realized shocks. I only focus on SPF forecasts here, as the individual-specific revisions are not available in SCE. Figure 4 plots the average revision in mean forecasts and uncertainty from 1-year-ahead forecasts in year $t - 1$ to the current-year nowcast in year t . The more negative range in which the revision lies, the more “rational” of the forecast. The histogram suggests that uncertainty revisions are left-skewed relative to zero. This implies, on average, forecasters feel more certain for her nowcasts relative to her forecast made one year before. However, over the entire sample, there is always a positive fraction of forecasters who revise uncertainty upward, which is inconsistent with the benchmark prediction. A formal test of revision equal to zero or being negative will be carried out in Section 4.

Figure 4: Distribution of Revision in Forecasts and Uncertainty



3 Theories of Expectation Formation

3.1 Sticky expectation (SE)

The theory of sticky expectation (Mankiw and Reis (2002), Carroll (2003) etc.), regardless of various micro-foundations, builds upon the assumption that agents do not update information instantaneously as they do in FIRE. One tractable assumption is that, by a homogenous and time-independent probability, λ , agents update their information up-to-date. Specifically, at any point of time t , each agent learns about the up-to-date realization of y_t with the probability of λ ; otherwise, they form the expectation based on the most recent up-to-date realization of $y_{t-\tau}$, where τ is the time experienced since the last update.

Denote the mean forecast of a non-updater since $t - \tau$ as $y_{i,t+h|t-\tau}$ since her forecast conditions upon the information up till $t - \tau$.

$$y_{i,t+h|t-\tau}^{se} = \rho^{h+\tau} y_{t-\tau} \quad (7)$$

$$\text{Var}_{i,t+h|t-\tau}^{se} = \text{Var}_{t+h|t-\tau}^* = \sum_{s=1}^{h+\tau} \rho^{2(s-1)} \sigma_{\omega}^2 \quad (8)$$

For non-updaters, the uncertainty is essentially dependent upon the time between the most recent update $t - \tau$ and the forecast date $t - h$, which is $\tau + h$. FIRE basically assumes $\tau = 0$ for all the agents and all the time, namely all agents' last update takes place in the previous period. So setting $\tau = 0$ in the above equation gives the uncertainty in FIRE common to everyone.

At the individual level, the key difference between FIRE and SE is that the latter does not reduce uncertainty as efficiently as in the former, primarily because of the rigidity of incorporating new information. But note that the rigidity in updating according to SE cannot be systematically observable at the individual level, both in terms of forecasts errors and uncertainty. This is because the behaviors of each individual forecast specifically depend on if she updates or not in that period.

Relying upon the law of large numbers, one can derive testable predictions about population moments that allow us to conduct tests of sticky expectation and recover rigidity parameter λ .

With the SE mechanism, the average forecast is a weighted average of update-to-date rational expectation and lagged average expectation, as reproduced below.²⁷ It can also be expressed as a weighted average of all the past realizations of y . Setting $\lambda = 1$, then the SE collapses to FIRE, and the average forecast is equal to y 's long-run mean of zero.

$$\bar{y}_{t+h|t}^{se} = \lambda \underbrace{y_{t+h|t}^*}_{\text{rational expectation at } t} + (1 - \lambda) \underbrace{\bar{y}_{t+h|t-1}^{se}}_{\text{average forecast at } t-1} \quad (9)$$

It follows that the average forecast errors are serially correlated, as described in Equation 10.

$$\overline{FE}_{t+h|t}^{se} = (1 - \lambda) \rho \overline{FE}_{t+h-1|t-1}^{se} + \lambda FE_{t+h|t}^{*2} \quad (10)$$

The unconditional variance of the h -period-ahead forecast error is proportional to that of the FIRE model.

$$\overline{FE}_{\bullet+h|\bullet}^{se2} = \frac{\lambda^2}{1 - (1 - \lambda)^2 \rho^2} FE_{\bullet+h|\bullet}^{*2} \quad (11)$$

Regarding the uncertainty, average uncertainty at any point of time is now a weighted average of uncertainty to agents whose last updates have taken place in different periods of the past. Since at any point of the time, there are agents who have not updated the recent realization of the shocks, thus with higher uncertainty, the population uncertainty is unambiguously higher than the case of FIRE. (See Appendix 6 for detailed derivations.)

²⁷See Coibion and Gorodnichenko (2012) for detailed steps.

$$\begin{aligned}
\overline{\text{Var}}_{t+h|t}^{se} &= \sum_{\tau=0}^{+\infty} \underbrace{\lambda(1-\lambda)^\tau}_{\text{fraction of non-updater until } t-\tau} \underbrace{\text{Var}_{t+h|t-\tau}^*}_{\text{uncertainty based on updating by } t-\tau} \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \sum_{s=1}^{h+\tau} \rho^{2(s-1)} \sigma_\omega^2 \\
&= \left(\frac{\lambda \rho^{2h}}{1 - \rho^2 + \lambda \rho^2} - 1 \right) \frac{1}{\rho^2 - 1} \sigma_\omega^2 \\
&\geq \overline{\text{Var}}_{t+h|t}^*
\end{aligned} \tag{12}$$

With respect to revision, the inefficiency of reducing uncertainty in SE takes the following form at the aggregate level. Since not all agents incorporate the recently realized shocks, the revision in average uncertainty exhibits a serial correlation, as described in Equation 13. It is a weighted average of the resolution of uncertainty from the most recent shocks and its lagged counterpart.

$$\overline{\text{Var}}_{t+h|t+1}^{se} - \overline{\text{Var}}_{t+h|t}^{se} = (1-\lambda)(\overline{\text{Var}}_{t+h|t}^{se} - \overline{\text{Var}}_{t+h|t-1}^{se}) - \lambda \rho^{2(h-1)} \sigma_\omega^2 \tag{13}$$

In particular, the second component is the information gain from the most recent realization of the shock, underweighted by $\lambda < 1$. The first component is the inefficiency sourced from the stickiness of updating. The higher rigidity (lower λ), the smaller the efficiency gain or uncertainty reduction compared to in FIRE.

Lastly, SE also predicts non-zero disagreements and sluggish adjustment compared to FIRE. This is because of different lags in updating across populations.

$$\overline{\text{Disg}}_{t+h|t}^{se} = \lambda \sum_{\tau=0}^{\infty} (1-\lambda)^\tau (y_{t+h|t-\tau} - \bar{y}_{t+h|t})^2 \tag{14}$$

From time t to $t+1$, the change in disagreement comes from two sources. One is newly realized shock at time $t+1$. The other component is from people who did not update at time t and update at time $t+1$.²⁸

3.1.1 Summary of predictions of SE

- Average forecast errors are zero across time, but have a positive serial correlation and have a higher variance across time than in FIRE.
- Population disagreement is positive on average across time.
- Population average uncertainty is higher than FIRE, and its revision under-reacts to the realized volatility of the shocks.

²⁸Coibion and Gorodnichenko (2012) derive the impulse response of disagreement at time $t+k$ to a shock that realized at t to be $\rho^{2(h+k)}(1-(1-\lambda)^{k+1})(1-\lambda)^{k+1}\omega_t^2$. Disagreement increases following the shock and gradually returns to its steady-state level.

3.2 Noisy information (NI)

A class of models (Lucas (1972), Sims (2003), Woodford (2001), and Maćkowiak and Wiederholt (2009), etc), noisy information (NI hereafter), describes the expectation formation as a process of extracting or filtering a true variable y_t from a sequence of realized signals. The starting assumption is that the agent cannot observe the true variable perfectly. But unlike SE, NI assumes agents keep track of the realizations of signals instantaneously all the time.

I adopt the identical framework to Coibion and Gorodnichenko (2015) by assuming that an agent i observes two signals, s^{pb} being a public signal common to all agents, and s_i^{pr} , being a private signal specific to the agent i . The generating process of two signals is assumed to be the following.

$$\begin{aligned} s_t^{pb} &= y_t + \epsilon_t, & \epsilon_t &\sim N(0, \sigma_\epsilon^2) \\ s_{i,t}^{pr} &= y_t + \xi_{i,t} & \xi_{i,t} &\sim N(0, \sigma_\xi^2) \end{aligned} \quad (15)$$

Stacking the two signals into one vector $s_{i,t} = [s_t^{pb}, s_{i,t}^{pr}]'$ and $v_{i,t} = [\epsilon_t, \xi_{i,t}]'$, the equations above can be rewritten as

$$\begin{aligned} s_{i,t} &= Hy_t + v_{i,t} \\ \text{where } H &= [1, 1]' \end{aligned} \quad (16)$$

Now any agent trying to forecast the future y has to form her expectation of the contemporaneous y . Denote it as $y_{i,t|t}^{ni}$, which needs to be inferred from the signals particular to the agent i . The agent's best h -period ahead forecast is simply iterated h periods forward based on the AR(1) process, and it is equal to $\rho^h y_{i,t|t}$. This is the same as FIRE.

What is different from FIRE is that the agent makes her best guess of y_t using Kalman filtering at the time t . Specifically, the mean forecast of an individual i is the posterior mean based on her prior and realized signals $s_{i,t}$.

$$\begin{aligned} y_{i,t|t}^{ni} &= \underbrace{y_{i,t|t-1}^{ni}}_{\text{prior}} + P \underbrace{(s_{i,t|t} - s_{i,t|t-1})}_{\text{innovations to signals}} \\ &= (1 - PH)y_{i,t|t-1}^{ni} + PHy_t + Pv_{i,t} \end{aligned} \quad (17)$$

where the Kalman gain P is a vector of size of two that determines the degrees of reaction to signals.

$$P = [P_\epsilon, P_\xi] = \text{Var}_{i,t|t-1}^{ni} H (H' \text{Var}_{i,t|t-1}^{ni} H + \Sigma^v)^{-1} \quad (18)$$

$\text{Var}_{i,t|t-1}^{ni}$ is the forecast uncertainty of y_t based on prior beliefs up to $t-1$. And Σ^v is a 2-by-2 matrix indicating the noisiness of the two signals.

$$\Sigma^v = \begin{bmatrix} \sigma_\epsilon^2 & 0 \\ 0 & \sigma_\xi^2 \end{bmatrix} \quad (19)$$

Individual forecast partially responds to new signals as $PH < 1$. $PH = 1$ is a special case when both signals are perfect, thus $\Sigma^v = 0$, then the formula collapses to FIRE.

A comparable parameter with $1 - \lambda$ in SE that governs rigidity in NI is $1 - PH$. It is a function of previous period uncertainty, $\text{Var}_{t-1|t-1}$, and noisiness of the signals determined by Σ^v . Note that beyond steady state, P is time-varying as the variance is updated by the agent each period (as in Equation 23). With a constant volatility in AR(1), we can focus on the Kalman gain in steady state, corresponding to a constant variance. Therefore, I drop time t from P .

What differentiates average forecast from individual's is the role played by private signals. On average, private signals cancel out across agents, therefore, only public signals enter the average forecast, hence, average forecast errors (Equation 21).

$$\begin{aligned}\bar{y}_{t+h|t}^{ni} &= \rho^h [(1 - PH) \underbrace{\bar{y}_{t+h|t-1}^{ni}}_{\text{Average prior}} + P \underbrace{\bar{s}_t}_{\text{Average signals}}] \\ &= (1 - PH) \bar{y}_{t+h|t-1}^{ni} + P_\epsilon \epsilon_t\end{aligned}\tag{20}$$

The population average forecast error under NI takes a similarly recursive form as in SE.

$$\overline{FE}_{t+h|t}^{ni} = (1 - PH) \rho \overline{FE}_{t+h-1|t-1}^{ni} + \rho^h P_\epsilon \epsilon_t + FE_{t+h|t}^{*}\tag{21}$$

The unconditional variance of the forecast errors is unambiguously greater than $FE_{t+h|t}^{*2}$ in FIRE, as shown in Equation 22.

$$\overline{FE}_{\bullet+1|\bullet}^{ni2} = \frac{\rho^{2h} P_\epsilon^2 \sigma_\epsilon^2 + FE_{\bullet+h|\bullet}^{*2}}{(PH)^2}\tag{22}$$

Kalman filtering also updates the variance recursively according to the rule of normal updating. The posterior variance at time t is a linear function of uncertainty in the previous period and variance of signals.

$$\text{Var}_{i,t|t}^{ni} = \text{Var}_{i,t|t-1}^{ni} - \text{Var}_{i,t|t-1}^{ni} H' (H \text{Var}_{i,t|t-1}^{ni} H' + \Sigma^v)^{-1} H \text{Var}_{i,t|t-1}^{ni}\tag{23}$$

The unconditional nowcasting variance can be solved as the steady-state value of the Equation 23. In the steady-state, there is no heterogeneity across agents in forecasting uncertainty and the nowcasting uncertainty becomes a constant. Thus, we can drop the subscript i . Note that the average uncertainty non-linearly depend on the noisiness of the two signals σ_ϵ^2 and σ_ξ^2 , as well as the volatility of inflation itself. The relationship is not monotonic, either.

Beyond steady-state, the Equation 23 also directly gives the revision in uncertainty from time $t - 1$ to t . The newly arrived information, albeit noisy, still brings about information gains, thus leading to an unambiguous drop in uncertainty. But due to the

signal is not perfect, i.e. $\Sigma^v \neq 0$, there is inefficiency in reducing uncertainty compared to in FIRE.

$$\text{Var}_{i,t|t}^{ni} - \text{Var}_{i,t|t-1}^{ni} = -\text{Var}_{i,t|t-1}^{ni} H' (H \text{Var}_{i,t|t-1}^{ni} H' + \Sigma^v)^{-1} H \text{Var}_{i,t|t-1}^{ni} < 0 \quad (24)$$

The h -period-ahead forecasting uncertainty comes from both nowcasting uncertainty and volatility of unrealized shocks in the future. (Equation 25)

$$\text{Var}_{i,t+h|t}^{ni} = \rho^{2h} \text{Var}_{i,t|t}^{ni} + \sum_{s=1}^h \rho^{2(s-1)} \sigma_\omega^2 \geq \text{Var}_{t+h|t}^* \quad (25)$$

As a result of NI mechanism, the revision in h -period-ahead uncertainty from $t-1$ to t only partially reacts to the resolution of uncertainty from newly realized shock ω_t in the past period.

NI also predicts non-zero disagreement in the presence of private signals. The size of the disagreement increases with the noisiness of the private signals.

$$\overline{\text{Disg}}_{t+h|t}^{ni} > 0 \quad (26)$$

3.2.1 Summary of predictions of NI

- Individual and population forecast errors are zero on average across time, but are positively serially correlated and have a higher variance than that in FIRE.
- The steady-state/unconditional uncertainty of both individual and population are higher than FIRE, and it depends on, but does not monotonically increase with the noisiness of private and public signals. Uncertainty revision is smaller than that in FIRE due to the underreaction to noisy signals.
- Population disagreement is positive on average across time, and it is a non-linear and non-monotonic function of the volatility of private signals σ_ξ^2 .

3.3 Diagnostic expectations (DE)

Different from the previous two theories featuring informational rigidity, diagnostic expectation (Bordalo et al. (2018)) introduces an extrapolation mechanism in expectation formation that results in overreactions to the news (Bordalo et al. (2020)). Both SE and NI deviate from FIRE in terms of the information set available to the agents, while DE deviates from FIRE in terms of the processing of an otherwise fully updated information set.

Skipping over its micro foundation, the following equation captures the essence of DE's mechanism. Each individual i 's h -period-ahead forecast consists of two components. The first component can be considered as a rational forecast based on the fully updated y_t . The second component corresponds to the potential overreaction to the unexpected surprises from $t-1$ to t . The degree of overreaction is governed by the

parameter θ . The premise of DE models is that $\theta \geq 0$, capturing the fact that the agent overly responded to the realized forecast errors. The forecast collapses to the FIRE when $\theta = 0$.

$$\bar{y}_{i,t+h|t}^{de} = \rho^h y_t + \theta_i (\rho^h y_t - \bar{y}_{i,t+h|t-1}^{de}) \quad (27)$$

There is no room for disagreement with a homogenous degree of overreaction. To account for the possibility of a positive disagreement, I assume θ to be different across different agents, therefore, I add the subscript i to the parameter. Since agents are equally informed about the realizations of the variable, the only room for disagreement to be positive is heterogeneous degrees of overreaction. To capture this, I assume θ_i to follow a normal distribution across the population, $N(\hat{\theta}, \sigma_\theta^2)$. So the DE model has two parameters. Disagreement increases with the dispersion of overreaction, σ_θ .

The average forecast takes exactly the same form, with the individual-specific θ_i replaced by the population average $\hat{\theta}$. Therefore, we focus on average forecast errors directly, as written below. (See Appendix 6 for derivations)

$$\overline{FE}_{t+h|t}^{de} = FE_{t+h|t}^* + \hat{\theta}(FE_{t+h-1|t}^* - \rho \overline{FE}_{t+h-1|t-1}^{de}) \quad (28)$$

The formula contains the intuition behind the extrapolation mechanism in DE: moving from $t-1$ to t , the h -period-ahead forecast error exceeds that of FIRE forecast error by exactly the surprise to the expectation formed at $t-1$ with a degree of overreaction $\hat{\theta}$.

The unconditional variance of h -period-ahead forecast errors is equal to the following.

$$\overline{FE}_{\bullet+h|\bullet}^{de2} = \frac{(1 + \hat{\theta})^2}{1 + \hat{\theta}^2 \rho^2} FE_{\bullet+h-1|\bullet}^{*2} + \frac{\sigma_\omega^2}{1 + \hat{\theta}^2 \rho^2} \quad (29)$$

Finally, as to the uncertainty, since the mechanism of extrapolation in DE does not change the agent's perceived distribution of the future shocks, benchmark DE theory predicts the forecast uncertainty to remain the same as in FIRE.

$$\overline{Var}_{t+h|t}^{de} = \overline{Var}_{t+h|t}^* \quad (30)$$

3.3.1 Summary of predictions of DE

- Average forecast errors are zero on average across time, but have **negative** serial correlation and higher variance than FIRE.
- The average uncertainty is equal to that in FIRE.
- Population disagreements are positive on average across time if there is heterogeneity in the degree of overreaction.

3.4 Diagnostic Expectation (DE) augmented with heterogeneous information (DENI)

[Bordalo et al. \(2020\)](#) embeds heterogeneous information in a standard DE model. Their motivation is primarily to generate cross-sectional disagreement in forecasts because the baseline version predicts zero-dispersion unless heterogeneity in the degree of overreaction is introduced in the first place, as we did in the previous section. The framework is essentially a hybrid of the NI and DE. It maintains the assumption regarding how agents overreact to new information at individual levels, but the information is no longer the real-time realization of the variable but a noisy signal of it. Since the mechanism is explicitly discussed in previous sections, I omit the detailed derivations here.

3.4.1 Summary of predictions of DENI

- Population forecast errors are zero on average across time. The sign of the auto-correlation depends on the relative size of the positive serial correlation induced by rigidity and the negative one induced by overreaction.
- Average uncertainty is now the same as NI due to the noisiness of the signal, and it increases with the noisiness of signals and volatility of inflation per se.
- Disagreement is positive on average due to the presence of private information. It also has positive auto-correlation and higher variance across time than FIRE.

3.5 An inflation process with stochastic volatility (SV)

This section considers an alternative data-generating process in which the volatility of shocks to inflation are stochastic, following the unobservable components/stochastic volatility (UCSV, or SV for short) model by [Stock and Watson \(2007\)](#). One of the widely acknowledged findings from the paper is that the univariate inflation process is better described by a two-component model with time-varying volatility.

For the purpose of this paper, the SV assumption has two additional purposes. First, a simple inflation process with constant volatility is not consistent with the time-varying pattern of observed forecast uncertainty as well as its comovement with other moments such as disagreement and the size of forecast errors, as seen in [Figure 2.4](#). Second, for the purpose of understanding expectation formation, extending the inflation process to allow for stochastic volatility also serves as a sensitivity test of various theories of expectation formation with respect to the underlying generating process of the inflation.

In particular, UCSV assumes that inflation consists of a permanent ζ and transitory component η . Time variations in the relative size of the volatility of two components σ_ζ^2 and σ_η^2 drive time variations of the persistence of inflation shocks. The logged volatility of two components themselves evolve as random walks subject to shocks μ_ζ and μ_η .

$$\begin{aligned}
y_t &= \zeta_t + \eta_t, \quad \text{where } \eta_t = \sigma_{\eta,t} \nu_{\eta,t} \\
\zeta_t &= \zeta_{t-1} + z_t, \quad \text{where } z_t = \sigma_{z,t} \nu_{\epsilon,t} \\
\log \sigma_{\eta,t}^2 &= \log \sigma_{\eta,t-1}^2 + \mu_{\eta,t} \\
\log \sigma_{z,t}^2 &= \log \sigma_{z,t-1}^2 + \mu_{\epsilon,t}
\end{aligned} \tag{31}$$

The shocks to the level of the two components η_t , and z_t , and those to their volatility, $\mu_{\eta,t}$ and $\mu_{z,t}$, are drawn from the following distributions, respectively. The only parameter of the model is γ , which determines the smoothness of the time-varying volatility.

$$\begin{aligned}
\nu_t &= [\nu_{\eta,t}, \nu_{z,t}] \sim N(0, I) \\
\mu_t &= [\mu_{\eta,t}, \mu_{z,t}]' \sim N(0, \gamma I)
\end{aligned} \tag{32}$$

The information set necessary for forecasting is different in SV from that in an AR(1) process. Consider first the benchmark case of FIRE. At time t , the FIRE agent sees the most recent and past realization of all stochastic variables as of t , including y_t , ζ_t , η_t , $\sigma_{\eta,t}$, $\sigma_{z,t}$. Using the superscript $*sv$ to denote the FIRE benchmark prediction under the stochastic volatility, and suppressing the individual subscript i (because there is no disagreement in FIRE), the h -period-ahead forecast of inflation is equal to the contemporaneous realization of the permanent component, $\epsilon_t \equiv \zeta_t$.

$$\bar{y}_{t+h|t}^{*sv} = \zeta_t \tag{33}$$

Under FIRE, forecast error is simply the cumulative sum of unrealized permanent and transitory shocks from t to $t+h$, which is equal to the following. And, disagreement is zero across agents in FIRE.

$$\overline{FE}_{t+h|t}^{*sv} = - \sum_{s=1}^h (\eta_{t+s} + z_{t+s}) \tag{34}$$

The h -step-ahead conditional variance or the forecast uncertainty is time-varying, as the volatility is stochastic now. It depends on the realizations of the component-specific volatility as of t .

$$\begin{aligned}
\overline{Var}_{t+h|t}^{*sv} &= \sum_{s=1}^h E_t(\sigma_{\eta,t+s}^2) + E_t(\sigma_{z,t+s}^2) \\
&= \sigma_{\eta,t}^2 \sum_{s=1}^h \exp^{-0.5s\gamma} + \sigma_{z,t}^2 \exp^{-0.5h\gamma}
\end{aligned} \tag{35}$$

SESV Under the sticky expectation (SE), an agent whose most recent up-to-date update happened in $t - \tau$ only have seen the realizations of y , ζ , η , σ_{η} , σ_z till $t - \tau$. The

average forecast is hence the weighted average of all past realizations of the permanent component up to t .

$$y_{t+h|t-\tau}^{sesv} = \sum_{\tau=0}^{\infty} \lambda(1-\lambda)^{\tau} \zeta_{t-\tau} \quad (36)$$

The distribution of lagged updating is also reflected in the average forecast uncertainty. The population average uncertainty is a weighted average of FIRE uncertainty at $t, t-1 \dots t-\tau \dots t-\infty$. (Equation 37) The key difference in SV from AR(1) is that the average uncertainty exhibits positive serial correlation under SV. Expectations being sticky further increases the positive serial correlation compared to that in FIRE due to the lag in updating of the shocks to the volatility. The predictions regarding both forecast errors and disagreements under SV are the same.

$$Var_{t+h|t}^{sesv} = \sum_{\tau=0}^{\infty} \lambda(1-\lambda)^{\tau} Var_{t+h|t-\tau}^{*sv} \quad (37)$$

NISV Under noisy information (NI), in order to forecast future y , the agent at time t needs to form her best nowcast of the permanent component ζ_t , denoted as $\bar{\zeta}_{t|t}$, using noisy signals and Kalman filtering. We assume again that the noisy signals of ζ_t consists of a public signal s_t^{pb} and a private signal $s_{i,t}^{pr}$ containing noises around the true realization of ζ_t . Following a long tradition of modeling the signaling-extraction problem in this two-component context, we further assume the public signal $s_t^{pb} = y_t$, meaning the inflation realization itself is the public signal of the permanent component. Accordingly, the transitory shock η_t is equivalent to the realized noise of the public signal ϵ_t .

$$y_{t+h|t}^{nisv} = \bar{\zeta}_{t|t} = (1 - P_t^{sv} H) y_{t+h-1|t-1}^{nisv} + P_{z,t}^{sv} z_t \quad (38)$$

In the above equation, Kalman gain $P_t^{sv} = [P_{z,t}^{sv}, P_{\xi,t}^{sv}]$ is a function of forecasting uncertainty $Var_{t|t-1}^{svni}$, the constant noisiness of private signal σ_{ξ} and that of public signal, $\sigma_{z,t}$, the time-varying volatility of the transitory component of the inflation.

The prediction regarding forecast errors and disagreements remain the same as in AR(1). What is different under time-varying volatility is that there is no steady-state Kalman gain and uncertainty that are independent of time because the underlying volatility of the variable is time-varying. This also implies that the rigidity induced by the noisiness of information is state-dependent. At each period, the agents in the economy will update their forecast based on the realized volatility. In periods with high (low) fundamental volatility, the Kalman gain from noisy signals is larger (smaller) thus the agents will be more (less) responsive to the new information. There is no such state-dependence of rigidity in SE.

The mechanisms of DE and DENI exactly mimic that under AR(1) except that the average volatility is time-varying now.

3.6 Comparing theories

We summarize the distinctions across different theories in their moment conditions. The three theories all predict average population FEs across time to be zero and have higher variances than in FIRE. The distinction lies in that two rigidity models SE and NI imply a positive serial correlation while DE implies a negative serial correlation. DENI allows the two forces to counteract each other. Therefore, the sign of the serial correlation is ambiguous.

In terms of disagreement across agents, both SE and NI predict non-zero disagreement, positive serial correlation, and higher volatility. Non-zero disagreement can only arise in DE only if heterogeneity in overreaction is introduced. But this still leads to a zero serial correlation. The prediction of DENI is the same as that of NI.

Forecast uncertainty also contains useful distinctions in moment conditions across theories. In particular, although SE and NI both predict higher uncertainty than FIRE, the magnitudes of that depend on the size of structural parameters in the respective model. In contrast, baseline DE predicts equal uncertainty to FIRE. DENI has similar predictions as NI.

The general takeaway is that not only the first moment such as forecast error but also higher moments, disagreement and uncertainty contains restrictions to identify the model parameters within each theory. We will utilize these moment conditions to estimate each theory in Section 5.

4 Reduced-form Tests of FIRE

As a credible starting benchmark, I first reproduce a number of reduced-form statistical tests of FIRE only using information from forecast errors primarily following [Mankiw et al. \(2003\)](#), and the results are reported in Table 11 in Appendix 6. Consistent with the existing findings, the results reject the null hypothesis of unbiasedness in forecasts, non-serial correlation of non-overlapping forecast errors, and efficient use of information in forecasting.

This section presents a number of new tests relying on uncertainty in Table 4 and 5, in the spirit of forecasting efficiency by [Nordhaus \(1987\)](#). It is an extension of revision tests on mean forecasts by [Fuhrer \(2018\)](#) to the forecast uncertainty.

Table 4 focuses on estimating forecasting efficiency using revisions of mean forecasts and uncertainty, hereafter referred to as revision-based tests. In plain words, the revision from 1-year-ahead forecast to nowcast of current-year inflation is efficient if the following two criteria are satisfied: (1) forecast revision does not depend on past information, including the past revisions; (2) the drop in uncertainty is sufficiently rapid to reflect the uncertainty of all realized shocks.

The mean revision test by [Fuhrer \(2018\)](#) takes the following form (using 1 period as an example):

$$y_{i,t+1|t+1} - y_{i,t+1|t} = \alpha + \beta(y_{i,t+1|t} - y_{i,t+1|t-1}) + \epsilon_{i,t+1} \quad (39)$$

In the above equation $\beta = 0$ according to FIRE, because rational forecast revision

only responds to newly realized shocks, thus it is not predictable by past revisions.²⁹ Since we have four vintages of the forecasts from SPF, the above specification can include lagged revisions up to 4 quarters.

The test with uncertainty simply replaces the revision of forecast with revision in uncertainty, as shown below.

$$\text{var}_{i,t+1|t+1} - \text{var}_{i,t+1|t} = \alpha^{\text{var}} + \beta^{\text{var}}(\text{var}_{i,t+1|t} - \text{var}_{i,t+1|t-1}) + \zeta_{i,t+1} \quad (40)$$

This regression follows from Equation 13 for SE and Equation 24 for NI. The autocorrelation coefficient β^{var} speaks to the speed of the drop in uncertainty, which takes zero value of zero under FIRE and one with perfect rigidity. Depending on the model, one can interpret it as the particular structural parameter of rigidity.

The top panel in Table 4 presents the results for the mean forecast. Following Fuhrer (2018), I include the median forecast available at time t and $t-1$ as an indicator of past information for the revision regression. In the first column of each panel, I report the regression on a constant.

The mean revision in forecast is mildly negative and significant for CPI forecast. The second to fourth columns of each panel in Table 4 checks autocorrelation of revisions, including different lags. Revisions of forecasts are serially correlated over 4 quarters, and the coefficients are all positive and significant. Also, the median forecasts as the past information always predict a negative revision with significant coefficients. This is evidence against the null hypothesis of FIRE and my estimates are comparable with those by Fuhrer (2018).

The bottom panel reports autoregression results for revision in uncertainty. Again, the first column first test the mean revision against the null being zero. For professional forecasters, the mean revisions in uncertainty are negative (0.25-0.3 percentage points equivalence in standard deviation of uncertainty) and statistically significant, confirming our observation from Figure 4 that forecasters are more certain about current inflation compared to her previous year forecast. However, for households at the population level, moving from 3-year-ahead inflation to 1-year-ahead inflation 2-years later, the drop in uncertainty is not significantly different from zero, suggesting inefficient forecasting and inconsistent with the FIRE.

The second to fourth column shows a positive serial correlation of revision in uncertainty for both CPI and PCE forecasts. The revision to CPI seems more efficient as serial correlation is with only one-quarter lag. For PCE, the revisions in uncertainty are serially correlated with all past three quarters.

Table 5 presents the results with the revision replaced with change in mean forecasts and uncertainty, i.e. from $y_{t|t-1}$ to $y_{t+1|t}$. As we have discussed in Section 3, the auto-correlation of change in mean forecast and uncertainty do not bear testable predictions from FIRE. But if the forecasts and uncertainty are persistent in its first difference, it may imply that the agent does not react to the news and newly realized shocks sufficiently. In addition, the auto-correlation regressions of this kind is a useful characterization of the time series dynamics of forecasts. With the variable being the

²⁹Adding $y_{t+1|t}$ to both sides of Equation 39 gives an equivalent null hypothesis used by Fuhrer (2018): coefficient of regression of $y_{t+1|t+1}$ on $y_{t+1|t}$ is $1 - \beta = 1$.

Table 4: Tests of Revision Efficiency Using Mean Revision and Uncertainty

Test 1. Revision efficiency of mean forecast																		
SPF CPI									SPF PCE				SCE					
	Mean revision	t-1	t-1-t-2	t-1-t-3	Mean revision	t-1	t-1-t-2	t-1-t-3		Mean revision	t-1	t-1-t-2	t-1-t-3		Mean revision	t-1	t-1-t-2	t-1-t-3
L.InfExp_Mean_rv		0.482*** (0.037)	0.406*** (0.040)	0.400*** (0.044)		0.569*** (0.050)	0.473*** (0.069)	0.456*** (0.080)	L.InfExp_Mean_rv		0.986*** (0.048)	0.844*** (0.111)	0.861*** (0.114)					
L2.InfExp_Mean_rv			0.153*** (0.036)	0.121** (0.039)			0.164** (0.051)	0.133* (0.058)	L2.InfExp_Mean_rv			0.130 (0.115)	0.119 (0.125)					
L3.InfExp_Mean_rv				0.054 (0.045)				0.080 (0.040)	L3.InfExp_Mean_rv			0.047 (0.116)	0.250 (0.157)					
SPFCPI.ct50		0.393** (0.120)	0.370* (0.142)	0.354* (0.149)					L4.InfExp_Mean_rv				-0.195 (0.156)					
SPFPCE.ct50						0.380** (0.120)	0.341* (0.134)	0.340* (0.143)	L5.InfExp_Mean_rv				0.091 (0.141)					
									L6.InfExp_Mean_rv				-0.155 (0.114)					
Const	-0.070* (0.036)	-0.858** (0.249)	-0.809** (0.290)	-0.773* (0.307)	-0.056 (0.033)	-0.732** (0.222)	-0.654** (0.244)	-0.644* (0.263)	Const	-0.032 (0.045)	0.015 (0.018)	0.025 (0.017)	0.014 (0.018)					
N	1765	1501	1295	1136	1513	1275	1086	945	N	86	85	83	80					
R2	0.000	0.281	0.302	0.296	0.000	0.373	0.381	0.375	R2	0.000	0.858	0.876	0.884					
Test 2. Revision efficiency of uncertainty																		
	Mean revision	t-1	t-1-t-2	t-1-t-3	Mean revision	t-1	t-1-t-2	t-1-t-3		Mean revision	t-1	t-1-t-2	t-1-t-3		Mean revision	t-1	t-1-t-2	t-1-t-3
L.InfExp_Var_rv		0.339*** (0.074)	0.166 (0.093)	0.185* (0.089)		0.395*** (0.057)	0.352*** (0.066)	0.285*** (0.070)	L.InfExp_Var_rv		0.749*** (0.074)	0.770*** (0.151)	0.844*** (0.167)					
L2.InfExp_Var_rv			0.162* (0.078)	0.217** (0.073)			0.206*** (0.058)	0.114 (0.058)	L2.InfExp_Var_rv			0.014 (0.174)	-0.097 (0.203)					
L3.InfExp_Var_rv				0.231** (0.079)				0.271*** (0.069)	L3.InfExp_Var_rv			-0.005 (0.123)	-0.218 (0.175)					
Constant	-0.087*** (0.008)	-0.053*** (0.007)	-0.053*** (0.007)	-0.031** (0.009)	-0.076*** (0.006)	-0.048*** (0.006)	-0.035*** (0.006)	-0.026*** (0.005)	L4.InfExp_Var_rv				0.378* (0.153)					
									L5.InfExp_Var_rv				-0.134 (0.231)					
									L6.InfExp_Var_rv				0.066 (0.144)					
									Const	0.020 (0.157)	-0.006 (0.103)	0.037 (0.103)	0.037 (0.106)					
N	1696	1435	1237	1082	1622	1370	1171	1013	N	86	85	83	80					
R2	0.000	0.132	0.092	0.210	0.000	0.146	0.211	0.263	R2	0.000	0.567	0.602	0.639					

Standard errors are clustered by date. *** p<0.001, ** p<0.01 and * p<0.05.

first difference, the panel structure of SCE and SPF allows for calculating changes in individual levels for greater sample size, especially for households. Besides autoregression, I also report the constant estimate of the changes in the first column of each sub-panel.

The most noticeable pattern for both professionals and households, and for both mean forecast and uncertainty, is that past change predict future changes with universally negative coefficients across different horizons. Most of the negative coefficients are statistically significant in the 1% level. For instance, one percentage point of increase in SPF's CPI forecast in the previous quarter predicts around 0.29 percentage points decrease in the next quarter. This negative correlation is smaller in size but remain significant into further past, i.e. -0.24 for two quarters lag and -0.1 for three quarters lag. Data on SCE is monthly, so lags are included up to six months. The negative correlation between past and current changes are all negative and significant for households. The sizes of the correlation coefficients are comparable with professional forecasts when the monthly coefficients are converted to their quarterly average, i.e. -0.3 to -0.4.

Such an auto-correlation of change in mean forecast is very much reflected in the same regression for uncertainty, as reported in the bottom panel of Table 5. For SPF of CPI and PCE, respectively, one unit increase in uncertainty about 1-year-ahead inflation in the previous quarter predicts around a 0.39 and 0.44 unit of the drop in the next quarter. The effect holds up to two quarters for professionals and 5 months for households.

These evidence suggest that both the mean forecast and uncertainty of individuals are mean-reverting. An essentially equivalent explanation is that both series are realizations of noisy signals around their respective long-run mean. This will lead to the exact negative correlation of the first differences we have seen. The mean-reverting patterns might also explain what is observed from Figure 3, according to which, there are no significant changes in the distribution across different years.

The second noticeable result lies in the constant regressions reported in the first column of each sub-panel in Table 5. It implies that households constantly lower their mean forecasts as well as uncertainty from month to month, while professional forecasts do not behave in such a pattern. In particular, the constant regression of the change in the mean forecast for SCE gives an estimated coefficient of -0.05 which is significant in the 5% level. Individual households' 1-year-ahead inflation expectations keep being downward adjusted each month compared to their previous answer. What is more interesting is that their uncertainty about 1-year-ahead inflation also decreases each month. The size of the downward adjustment is -1.39 unit and statistically significant in the level of 0.1%. This negative significant and constant coefficient remains throughout all auto-regressions, implying it is not driven by time-varying changes.

The most natural explanation for this, that repeatedly surveyed households have become more informative about inflation over time. Given the unconditional forecast errors of inflation by households are positive, a downward adjustment of inflation stands for a less-biased forecast. ³⁰

³⁰The possibility that the surveys' information set is influenced by the survey itself, or "learning through surveys" (Kim and Binder, 2020), is a double-edged sword. On one hand, this poses a

Table 5: Weak Tests of Revision Efficiency Using Change in Forecasts and Uncertainty

Test 3. Weak. Efficiency of change in forecast	SPF CPI			SPF PCE			SCE		
	Mean change	t-1	t-1-t-2	t-1-t-3	Mean revision	t-1	t-1-t-2	t-1-t-3	t-1-t-2
L.InfExp_Mean_ch	-0.249*** (0.062)	-0.308*** (0.054)	-0.311*** (0.060)	-0.324*** (0.089)	-0.289*** (0.060)	-0.408*** (0.005)	-0.569*** (0.012)	-0.612*** (0.008)	-0.569*** (0.012)
L2.InfExp_Mean_ch		-0.194*** (0.064)	-0.177** (0.060)	-0.100 (0.081)	-0.100 (0.081)		-0.305*** (0.007)	-0.385*** (0.012)	-0.305*** (0.007)
L3.InfExp_Mean_ch			-0.075* (0.036)	0.014 (0.064)			-0.126*** (0.007)	-0.241*** (0.012)	-0.126*** (0.007)
							-0.142*** (0.012)	-0.142*** (0.012)	-0.142*** (0.012)
							-0.078*** (0.011)	-0.078*** (0.011)	-0.078*** (0.011)
							-0.038*** (0.008)	-0.038*** (0.008)	-0.038*** (0.008)
Constant	-0.019 (0.019)	-0.015 (0.019)	-0.021 (0.020)	0.017 (0.021)	0.026 (0.024)	-0.034** (0.013)	0.002 (0.017)	0.020 (0.019)	0.002 (0.017)
N	4286	3519	2581	1791	1538	85166	67555	20894	43489
R2	0.000	0.072	0.103	0.089	0.077	0.088	0.187	0.259	0.187

Test 4. Weak. Efficiency of change in uncertainty	Mean change			Mean change			Mean change		
	t-1	t-1-t-2	t-1-t-3	t-1	t-1-t-2	t-1-t-3	t-1	t-1-t-2	t-1-t-3
L.InfExp_Var_ch	-0.381*** (0.071)	-0.542*** (0.075)	-0.601*** (0.069)	-0.352*** (0.043)	-0.428*** (0.053)	-0.513*** (0.054)	-0.396*** (0.008)	-0.572*** (0.010)	-0.655*** (0.016)
L2.InfExp_Var_ch		-0.280*** (0.059)	-0.428*** (0.061)		-0.237*** (0.044)	-0.430*** (0.044)		-0.307*** (0.011)	-0.430*** (0.018)
L3.InfExp_Var_ch			-0.293*** (0.078)			-0.382*** (0.059)		-0.124*** (0.006)	-0.269*** (0.016)
								-0.155*** (0.015)	-0.155*** (0.015)
								-0.079*** (0.012)	-0.079*** (0.012)
								-0.029** (0.009)	-0.029** (0.009)
Constant	-0.002 (0.005)	-0.001 (0.006)	0.006 (0.005)	0.003 (0.004)	0.004 (0.005)	0.008 (0.004)	-0.717*** (0.064)	-0.593*** (0.072)	-0.448*** (0.076)
N	1685	1439	1251	1629	1406	1079	88052	69979	21476
R2	0.000	0.130	0.240	0.000	0.125	0.286	0.000	0.187	0.271

Standard errors are clustered by date. *** p<0.001, ** p<0.01 and * p<0.05.

In summary, the major additional insights that arise from the empirical tests of this section is that rigidity of incorporating new information in forming expectations imply noticeable inefficiency of revisions in forecasts and drop in uncertainty.

5 Model Estimation and Sensitivity Analysis

5.1 SMM Estimation

The reduced-form tests in Section 4 are sufficient in rejecting the null hypothesis of FIRE. But there are two limitations with these tests in terms of identifying differences among non-FIRE theories. First, the coefficient estimates from the reduced-form regression cannot always be mapped into a structural parameter of the particular model. Second, even if it does so, the tests fall short of simultaneously utilizing all the restrictions across moments implied by a particular non-FIRE theory, as discussed in great detail in Section 3. In this section, I undertake a structural estimation that jointly accounts for cross-moment restrictions.

Since many of the moment conditions cannot be easily derived as a closed-form function of parameters, I adopt the simulated method of moment (SMM). In a nutshell, the estimation chooses the best set of model parameters by minimizing the weighted distances between the data moments and the model-simulated moments. For a given process of inflation, and a particular theory of expectation formation, the vector of the parameters estimates is defined as the minimizer of the following objective function.

$$\hat{\Omega}^o = \underset{\{\Omega^o \in \Gamma^o\}}{\operatorname{argmin}} (M_{\text{data}} - F^o(\Omega^o, H))W(M_{\text{data}} - F^o(\Omega^o, H))'$$

where Ω^o stands for the parameters of the particular pair of theory of expectation and inflation process, i.e. $o \in \{se, ni, de, deni\} \times \{ar, sv\}$. Γ^o represents the corresponding parameter space respecting the model-specific restrictions. M_{data} is a vector of the unconditional moments that is computed from data on expectations and inflation. F^o is the simulated model moments under the theory pair o . W is the weighting matrix used for the SMM estimation. I report estimation results using the 2-step feasible SMM approach, in which the inverse of the variance-covariance matrix from the 1st-step estimation using identity matrix is used as the W in the second step, which has been shown to give asymptotically efficient estimates of the model parameters.

Crucially, notice that the model-implied moments F^o are not just a function of model parameters Ω^o , but also a function of the corresponding information set available to the forecasters. I use H to represent the historical realizations of the variables in the agents' information set that are used as the inputs for forecasts. Although the real-time history is the same across models, the mapping between the history to data moments depends on model specifics. For instance, although real-time inflation is

methodological challenge to the survey designers of expectations as to if surveys can objectively elicit the “true” expectations held by the respondents. On the other hand, researchers can use the survey as a meaningful intervention tool to identify the effect of factors such as information provision and attention. Recent examples of this type of research includes: [Coibion et al. \(2018\)](#) for firms and [Coibion et al. \(2019\)](#) for households.

the only variable in the information set for AR(1) process, the information set in SV contains both the permanent component of the inflation and the realized levels of volatility. Since the different components are not directly observed from historical data, I estimate it using the Markov Chain Monte Carlo (MCMC) procedure developed by [Stock and Watson \(2007\)](#) in this context. The estimated time-varying permanent and transitory volatility of both Core CPI and headline CPI is shown in Appendix 5.

It is also important to mimic the information set that was truly available to the agents at each point of the time in history³¹. Therefore, I use the real-time data on historical inflation that was publicly available at each point of the time instead of the historical data released later, since it is well known among macroeconomists that the latter typically incorporates many rounds of revisions over time. I obtain the data from the Real-Time Data Research Center hosted by Philadelphia Fed³².

The estimation is also specific to the choices of moments included in computing the distances. I focus on the unconditional moments of (independent of the time) at the population level, defined in Table 1. In particular, they include the mean (FE), variance($FEVar$) and auto-covariance($FEATV$) of population forecast error, the mean($Disg$), variance($DisgVar$) and auto-covariance($DisgATV$) of disagreement, and the mean (Var), variance($VarVar$) and auto-covariance($VarATV$) of uncertainty. When the joint estimation is done, two unconditional moments of the inflation are used, the variance($InfVar$) and auto-covariance($InfVar$). Table 2 reports the size of these moments computed for both SPF and SCE.

The model-implied moment conditions also implicitly depend on the parameters of the inflation process for a given model. This point is illustrated well in [Bordalo et al. \(2020\)](#). For instance, the observed overreaction in DE is lower for an AR(1) process with higher persistence. In recovering the model parameters associated with expectation formation, it is important to take into account the information contained in expectation data regarding the process of inflation per se. To handle this, I undertake both 2-step and joint estimation. The former refers to the first externally estimating the inflation process and then estimating expectation formation separately treating the inflation parameter as given. The latter refers to jointly estimating parameters of inflation and expectation.

These alternative specifications of the estimation also serve as a model sensitivity analysis with respect to the following criteria: (1) different choices of moments; (2) AR(1) and SV for the process of inflation. (3) two-step and jointly. (4) for both professionals and households. A reasonable theory of expectation formation ought to be relatively robust to these criteria.

5.2 Moments matching and parameter estimates

We first focus on the professionals' expectations as a benchmark, since professionals may exhibit less severe deviation from FIRE than households do, according to much previous empirical evidence. A first look at the data moments of SPF, as reported in Table 2, already provide clues as to the degree of the deviations from FIRE benchmark

³¹For the importance of using real-time data to study survey forecasts, see [Faust and Wright \(2008\)](#), [Faust and Wright \(2009\)](#) and so on.

³²<https://www.philadelphiafed.org/research-and-data/real-time-center/real-time-data/>.

predictions laid out in Section 3. When separately estimated, the core CPI inflation during 1995-2020 followed an AR(1) process with the persistence parameter $\hat{\rho} = 0.98$ and volatility $\hat{\sigma}_\omega = 0.22$. The FIRE benchmark predicts zero forecast error, zero disagreement, an uncertainty level identical to the volatility of inflation shocks $\sigma_\omega^2 = 0.04$, and zero time-variation and zero auto-correlation of these population forecasting moments. But in the data, quarterly professional forecasts of the core CPI inflation have a mildly positive average forecast error of 0.136 percentage points, a degree of Disg of 0.183, and an Var of 0.242, both of which are larger than the inflation volatility of $\sigma_\omega^2 = 0.04 \approx 0.22^2$ estimated from the same period. The average uncertainty is also greater than the square of forecast error or its variance 0.133, unlike the FIRE prediction.

Table 6 presents the SMM estimates of different model parameters for professionals. For each theory, I estimate the theory both in two steps or jointly using expectations and inflation data moments. Different rows within each panel report the estimates depending on various choices of moments used for estimation: forecast errors only (FE), forecast error and disagreement (FE+Disg), and the two plus uncertainty (FE+Disg+Var).

Table 6: SMM Estimates of Different Models: Professionals

SE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\lambda}$	ρ	σ	$\hat{\lambda}$	ρ	σ				
FE	0.35	0.99	0.23	0.18	1	0.01				
FE+Disg	0.3	0.99	0.23	0.18	1	0				
FE+Disg+Var	0.32	0.99	0.23	0.21	1	0.02				
NI										
Moments Used	2-Step Estimate				Joint Estimate					
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ		
FE	2.73	3	0.99	0.23	3	3	1	0.02		
FE+Disg	3	2.95	0.99	0.23	3	3	1	0.02		
FE+Disg+Var	3	3	0.99	0.23	1.97	1.17	1	0		
DE										
Moments Used	2-Step Estimate				Joint Estimate					
	$\hat{\theta}$	σ_{θ}	ρ	σ	$\hat{\theta}$	σ_{θ}	ρ	σ		
FE	0.25	1.57	0.99	0.23	1.39	2.83	0.9	0.17		
FE+Disg	1.12	1.62	0.99	0.23	0.05	1.73	1	0.01		
FE+Disg+Var	0.29	1.81	0.99	0.23	1.19	1.68	0.91	0.16		
DENI										
Moments Used	2-Step Estimate					Joint Estimate				
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ
FE	0.3	0.02	2.56	0.99	0.23	-1.14	2.75	0.12	1	0
FE+Disg	-1.21	3	0.09	0.99	0.23	-0.82	2.45	2.83	1	0.01
FE+Disg+Var	-0.8	3	0.09	0.99	0.23	1.67	3	3	1	0.03

5.2.1 Cross-moment consistency within a theory of expectation

Among four models under consideration, SE and NI outperform DE and DENI in terms of its within-model robustness against the targeted moments, as shown in the estimation of professional forecasts in Table 6.

For SE, the estimated quarterly updating rate is between 0.3-0.35 across different combinations of moments. Such a medium degree of information rigidity seems to reflect the counteracting evidence both for and against information rigidity. On one hand, average forecast errors of a much higher variance ($\text{FEVar} = 0.133$) than the inflation volatility $0.22^2 \approx 0.04$, and a positive disagreement of 0.183 and higher uncertainty of 0.242 than the conditional volatility are all in line with the patterns of the information rigidity in the form of SE. On the other hand, only a mild serial correlation in average forecast errors (with an auto-covariance $\text{FEATV} = 0.097$) and a small sized disagreement ($\text{Disg} = 0.183$) suggests the degree of the rigidity is not extremely large.

The NI estimates of noisiness of both public and private signals, despite their cross-moment consistency, turn out to be so large in magnitudes that they often hit the externally set upper bounds of 3 in estimation. These are highly noisy signals compared to the conditional standard deviation of inflation shocks $\sigma_\omega = 0.22$, suggesting that dispersed and noisy information does play a role in generating rigidity seen both in forecast errors and uncertainty, as well as the observed disagreement. Bringing in information from uncertainty does put some “discipline” on the parameters, generating smaller, albeit still large, noisiness of both signals.

Compared to the two rigidity models, DE estimates are more sensitive to moments restrictions yet the estimations all confirm the existence of a positive mass of overreactive agents. With only information from forecast error, the estimated overreaction parameter of DE θ is around 0.25 on average for all forecasters. Using disagreement helps identify the population dispersion in the degree of overreaction σ_θ^2 , which is estimated to be 1.62, but doing so also leads to a significantly different estimate of average overreaction $\hat{\theta}$ of 1.12. This is so because disagreements not only depends on the dispersion of overreaction, but also the average degree to which agents overreact. While, incorporating information from uncertainty, the estimated $\hat{\theta}$ and σ_θ reverse to resemble those based on only forecast errors.

DENI sees to bear even less cross-moment consistency. The implied overreaction parameter spans positive to negative values for professionals, depending on the specific moments that are used. This partly reflects the fact DENI has the most parameters among all theories, and hence a larger degree of freedom of the model parameterization. The problem also manifests itself in non-convergence in the estimation of households although additional information from uncertainty, in theory, could help identify parameters.

5.2.2 Interactions between expectation formation and inflation process

In all four models, estimated parameters of expectation formation vary when one jointly estimates expectation and inflation process parameters. For instance, with the benchmark AR(1) process, both the persistence of the shock to inflation ρ and the overall volatility of the inflation shock σ_ω determine what the FIRE forecasts moments

should be. Furthermore, model-specific forecasting moments are not only a function of the model parameters, but also the inflation process parameters itself. The differences between 2-step estimation and joint estimation reveal such inter-dependence.

For SE, letting professional forecasts reveal information about the inflation process leads to a more persistent (close to unit root) and lower conditional volatility. The joint estimate of the persistence parameter ρ becomes 1,0 and the inflation volatility becomes 0.0 – 0.02, compared to 0.99 and 0.23 when estimated separately. Consequently, the joint estimation produces a lower updating rate λ (higher rigidity) around 0.18 – 0.21. This is consistent with the underlying mechanisms of SE. For households, the mutual dependence is directly observed in terms of uncertainty. Joint estimation produces an identical degree of persistence, but a lower conditional volatility. This further requires a higher degree of rigidity to allow the model to match the observed level of uncertainty.

NI also shows some sensitivity toward the estimation procedure, especially when uncertainty is included for targeted moments. The joint estimation reveals a more persistent inflation process, $\rho = 1.0$ and a smaller size of conditional volatility close to zero. The implied noisiness of both signals correspondingly become smaller in the joint estimation. The mutual dependence between noisiness of signals and volatility of inflation shocks is not surprising because, in NI, both shocks to inflation itself or to the signals contribute to forecast errors and forecast uncertainty. Although the two are clearly distinguishable from the point of view of the modeler, the distinction between a shock to inflation itself and simply to signals may be indistinguishable for the agents who try to form the forecasts.

DE estimates in joint estimation demonstrate substitution between the persistence ρ and overreaction $\hat{\theta}$. A higher persistence in joint estimation comes with a lower degree of overreaction and vice versa. This echoes the finding in Afrouzi et al. (2020) about the inter-dependence between the implied persistence in forecasting and underlying degree of overreaction.

DENI performs the worst in terms of its sensitivity against the estimation procedure, implying it is very dependent on the persistence and volatility of the inflation process. Taking the most likely credible estimate that utilizes information from all moments, the joint estimation implies a significantly higher degree of overreaction $\hat{\theta} = 1.67$, much noisy private signals σ_{pr} from 0.09 to greater than 3, together with a more persistent inflation process with a smaller conditional volatility.

5.2.3 Professionals versus households

In contrast with professionals, raw household forecasts see a more substantial deviation from FIRE in every dimension. During the sample period of 2013-2020, the monthly headline CPI inflation is estimated to have a persistence parameter of 0.98 and volatility of 0.41.³³ The household forecasts had an average forecast error of 1.77, a disagreement of 2.585, and an average uncertainty of 1.75. Note that these are computed based on residuals by first controlling for the individual fixed effects. This follows

³³The higher volatility compared to that of core CPI is due to a higher frequency and inclusion of more volatile items in the CPI basket.

from the emerging evidence³⁴ that individual-specific effects such as demographics and experience play important roles in driving systematic differences in expectations. After controlling for this ex-ante heterogeneity in beliefs, households' expectations exhibit similar yet still more distorted expectations than that of professionals.

Table 7: SMM Estimates of Different Models: Households

SE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\lambda}$	ρ	σ	$\hat{\lambda}$	ρ	σ				
FE	0.35	0.98	0.41	0.2	0.98	0.16				
FE+Disg	0.35	0.98	0.41	0.2	0.98	0.1				
FE+Disg+Var	0.35	0.98	0.41	0.2	0.98	0.1				
NI										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ		
FE	3	0.45	0.98	0.41	2.98	0.28	0.95	0.24		
FE+Disg	1.55	0.36	0.98	0.41	3	0.28	0.95	0.24		
FE+Disg+Var	2.58	0.97	0.98	0.41	2.9	1.02	0.96	0.24		
DE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\theta}$	σ_{θ}	ρ	σ	$\hat{\theta}$	σ_{θ}	ρ	σ		
FE	-0.47	0.81	0.98	0.41	-0.61	4.15	0.95	0.25		
FE+Disg	-0.3	2.08	0.98	0.41	-0.08	2.13	0.95	0.25		
FE+Disg+Var	-0.35	2.08	0.98	0.41	0.36	2.07	0.9	0.36		
DENI										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ
FE	N/A	N/A	N/A	0.98	0.41	-0.39	1.86	0.03	0.95	0.24
FE+Disg	N/A	N/A	N/A	0.98	0.41	N/A	N/A	N/A	N/A	N/A
FE+Disg+Var	N/A	N/A	N/A	0.98	0.41	0.35	0	1.32	0.96	0.22

Table 7 presents the estimates for households after excluding the fixed effects on forecasts, forecast errors, and uncertainty. Overall, the estimates are surprisingly similar to that for professionals and within each model, in some sense, exhibits even more model consistency across moment conditions and estimation procedures. Most noticeably, the updating rate in SE falls into a very small range of values of 0.31 – 0.35 of each month. This is extremely close to our estimates for professionals. It is well documented in the literature that households are more irrational than professionals. But the results of our estimates show that the major differences are not least mostly due to the differences in updating rate in information. The second intuitive pattern revealed from the households estimates is that compared to those of professionals, the NI estimates of households show equally noisy public signals and more precise private signals across all moments, and according to both 2-step and join estimations.

³⁴Malmendier and Nagel (2015), Das et al. (2017), D'Acunto et al. (2019).

For DE, unlike professionals, the canonical DE model-based estimates suggest a negative, instead of positive, value of the overreaction parameter on the average, in conjunction with a large dispersion that gives a positive mass of overreactive agents.

The same as that of professionals, the household estimates of DENI are equally sensitive across targeted moments.

5.2.4 Alternative inflation process with stochastic volatility

Table 8 and 9 show the estimation of the model-specific parameters allowing the alternative inflation process featuring stochastic volatility in two separate unobserved components of different persistence. Compared to the benchmark AR(1) process, there are two crucial implications for expectation formation. The first is that now the SV model admits time-varying volatility, which has more potential to be consistent with the time-varying pattern of the forecast uncertainty, primarily in the expectations of the household. The second is that given the permanent component is a random walk, the shock to the permanent component is permanent instead of persistent. We examine if the alternative inflation process accommodates similar estimates of the model-specific parameters.

An overarching finding is that SV process of inflation significantly improves the within-model consistency of almost each model and for both types of agents.

Among all theories, SE gives the closest parameter estimates to that of the benchmark. Surprisingly enough, admitting stochastic volatility reveals an almost identical information rigidity in SE, i.e. an updating rate of 0.34 – 0.35 for both households and professionals. This is not a mechanical coincidence, since the SVSE model assumes agents do not only infrequently update realized shocks but also the shocks to the volatility, as explicitly discussed in the section 3.5. Therefore, the dynamics of uncertainty seen from data do provide useful additional information now to identify information rigidity than the benchmark model. This suggests that SE has a very good consistency against the assumed inflation process. Such information rigidity is also indirectly confirmed by the revealed underreaction according to DESV estimates. On average, both households and professionals underreact instead of overreacting to inflation news, even under the stochastic volatility model of inflation. More specifically, professionals show slightly more underreaction with SV than the benchmark estimates, and households show less.

The estimates for NI augmented with *SV* are a lot different from that of benchmark estimates, but appear to be more sensible, benefiting from the modified inflation process. This ultimately comes from a more sensible assumption about the form in NISV that what agents do not perfectly observe are the unobserved permanent component instead of the inflation itself. The estimated noisiness of public signals are close to the size of the average conditional volatility of inflation, by assumption of the NISV model. The private signals of households remain extremely large (2.37-2.9), and significantly larger than that of professionals (0.04).

5.3 The scoring card of different theories

Table 10 reports my evaluation of the four theories under consideration based on four sensitivity criteria laid out in the previous section. According to this evaluation, the

Table 8: SMM Estimates of Different Models under Stochastic Volatility: Professionals

SE			
Moments Used	2-Step Estimate		
	$\hat{\lambda}$		
FE	0.35		
FE+Disg	0.34		
FE+Disg+Var	0.34		
NI			
Moments Used	2-Step Estimate		
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	
FE	0.27	0.04	
FE+Disg	0.27	0.04	
FE+Disg+Var	0.27	0.05	
DE			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	σ_{θ}	
FE	-0.39	0.69	
FE+Disg	-0.39	0.7	
FE+Disg+Var	-0.39	0.7	
DENI			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$
FE	1.38	3	1.72
FE+Disg	3	3	2.03
FE+Disg+Var	3	3	2.28

SE seems to capture the average behavior of expectations better than the other two theories. DE shows a similar degree of performance according to these criteria, but the average estimation does not show overreaction in line with the premise of the theory, instead, underreaction, consistent with the existence of the information rigidity.

In comparison, NI and DENI perform less well in terms of these four criteria, except for the sensitivity against the moment conditions used for estimation. It is worth giving some diagnoses of the crucial challenges of the two theories. The central challenge that makes NI unable to explain the data well is that the level of information rigidity consistent with NI formulation requires too big sizes of the nosiness of the signals. And it is also very sensitive to the assumed level of inflation volatility and the degree of its persistence. Such a problem also exists in the DENI model, which in theory could reconcile the observed coexistence of information rigidity and overreaction better, as argued by [Bordalo et al. \(2020\)](#).

Table 9: SMM Estimates of Different Models under Stochastic Volatility: Households

SE			
Moments Used	2-Step Estimate		
	$\hat{\lambda}$		
FE	0.35		
FE+Disg	0.35		
FE+Disg+Var	0.35		
NI			
Moments Used	2-Step Estimate		
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	
FE	0.44	2.37	
FE+Disg	0.44	2.61	
FE+Disg+Var	0.44	2.9	
DE			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	σ_{θ}	
FE	-0.05	0.58	
FE+Disg	-0.05	0.56	
FE+Disg+Var	-0.05	0.56	
DENI			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$
FE	N/A	N/A	N/A
FE+Disg	0.71	0.01	0.95
FE+Disg+Var	N/A	N/A	N/A

Table 10: Scoring card of different theories

Criteria	SE	NI	DE	DENI
Sensitive to moments used for estimation?	No	No	Yes	No
Sensitive to the assumed inflation process?	No	Yes	Yes	Yes
Sensitive to two-step or joint estimate?	Yes	Yes	Yes	Yes
Sensitive to the type of agents?	Yes	Yes	Yes	Yes

6 Conclusion

Most studies on expectation formation that document how it deviates from the FIRE benchmark have focused on the first moment, namely the mean forecasts and the cross-sectional dispersion of the forecasts. However, this paper has shown that the surveyed forecasting uncertainty by professionals and households provides useful information for understanding the exact mechanisms of expectation formation. It not only provides additional reduced-form testing results of rejecting FIRE, such as persistent disagreements in forecasting uncertainties and its inefficient revisions, but also provides additional moment restrictions to any particular model of expectation forma-

tion, which helps identify differences across theories.

At least three lines of questions remain unresolved in this paper and require future research. First, this paper focuses on a selective list of models of expectation formation, and inevitably omits various others likewise proven to match certain aspects of surveyed inflation expectations, such as adaptive learning ([Marcet and Sargent, 1989](#); [Evans and Honkapohja, 2012](#)), experience-based learning ([Malmendier and Nagel, 2015](#)), heterogeneous models ([Patton and Timmermann, 2010](#); [Farmer et al., 2021](#)), and asymmetric attention ([Kohlhas and Walther, 2021](#)). It would be a fruitful attempt to explore the corresponding predictions of these models about uncertainty. Second, throughout the analysis, we maintained the normality/symmetric assumptions of the shocks and ignored beliefs in tail events or even higher moments. It would be natural to explore how different theories of expectation formation may contain different predictions on tail beliefs. Finally, although this paper focuses only on macroeconomic expectations regarding inflation, it is worth asking if the belief formation regarding individual variables such as income bears similar mechanisms and matches the observed empirical patterns of surveyed expectations and risks.³⁵

³⁵A few recent studies on income/wage/unemployment/job-search expectations: [Mueller et al. \(2021\)](#); [Wang \(2022\)](#); [Koşar and Van der Klaauw \(2022\)](#); [Jäger et al. \(2022\)](#); [Caplin et al. \(2023\)](#).

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Appendix

Detailed derivation

SE

$$\begin{aligned}
\bar{Var}_t(y_{t+h}) &= \sum_{\tau=0}^{+\infty} \underbrace{\lambda(1-\lambda)^\tau}_{\text{fraction who does not update until } t-\tau} \underbrace{Var_{t|t-\tau}(y_{t+h})}_{\text{Variance of most recent update at } t-\tau} \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \sum_{s=1}^{h+\tau} \rho^{2(s-1)} \sigma_\omega^2 \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau (1 + \rho^2 + \dots + \rho^{2(h+\tau-1)}) \sigma_\omega^2 \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{\rho^{2(h+\tau)} - 1}{\rho^2 - 1} \sigma_\omega^2 \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{\rho^{2(h+\tau)}}{\rho^2 - 1} \sigma_\omega^2 - \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{1}{\rho^2 - 1} \sigma_\omega^2 \\
&= \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \rho^{2\tau} \frac{\rho^{2h}}{\rho^2 - 1} \sigma_\omega^2 - \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{1}{\rho^2 - 1} \sigma_\omega^2 \\
&= \sum_{\tau=0}^{+\infty} \lambda((1-\lambda)\rho^2)^\tau \frac{\rho^{2h}}{\rho^2 - 1} \sigma_\omega^2 - \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{1}{\rho^2 - 1} \sigma_\omega^2 \\
&= \frac{\lambda}{(1-\rho^2 + \lambda\rho^2)} \sum_{\tau=0}^{+\infty} (1-\rho^2 + \lambda\rho^2)(1-(1-\rho^2 + \lambda\rho^2))^\tau \frac{\rho^{2h}}{\rho^2 - 1} \sigma_\omega^2 \\
&\quad - \sum_{\tau=0}^{+\infty} \lambda(1-\lambda)^\tau \frac{1}{\rho^2 - 1} \sigma_\omega^2 \\
&= \left(\frac{\lambda\rho^{2h}}{(1-\rho^2 + \lambda\rho^2)(\rho^2 - 1)} - \frac{1}{\rho^2 - 1} \right) \sigma_\omega^2 \\
&= \left(\frac{\lambda\rho^{2h}}{1-\rho^2 + \lambda\rho^2} - 1 \right) \frac{\sigma_\omega^2}{\rho^2 - 1} \\
&= \left(\frac{\lambda\rho^{2h} - 1 + \rho^2 - \lambda\rho^2}{1-\rho^2 + \lambda\rho^2} \right) \frac{\sigma_\omega^2}{\rho^2 - 1}
\end{aligned} \tag{41}$$

NI

The steady-state nowcasting uncertainty Var_{ss}^{ni} is solved using the updating equation (Equation 23).

$$\begin{aligned}
\text{Var}_{t|t}^{ni} &= \text{Var}_{t|t-1}^{ni} - \text{Var}_{t|t-1}^{ni} H' (H \text{Var}_{t|t-1}^{ni} H' + \Sigma^v)^{-1} H \text{Var}_{t|t-1}^{ni} \\
&\rightarrow \text{Var}_{t|t}^{ni} = \rho^2 (\text{Var}_{t-1|t-1}^{ni} + \sigma^2) \\
&- \rho^2 (\text{Var}_{ss}^{ni} + \sigma^2) H' (H \rho^2 (\text{Var}_{ss}^{ni} + \sigma^2) H' + \Sigma^v)^{-1} H \text{Var}_{ss}^{ni} \\
&\rightarrow \text{Var}_{ss}^{ni} = \rho^2 (\text{Var}_{ss}^{ni} + \sigma_\omega^2) \\
&- \rho^2 (\text{Var}_{ss}^{ni} + \sigma_\omega^2) H' (H \rho^2 (\text{Var}_{ss}^{ni} + \sigma_\omega^2) H' + \Sigma^v)^{-1} H \text{Var}_{ss}^{ni}
\end{aligned} \tag{42}$$

DE

$$\begin{aligned}
FE_{i,t+h|t}^{de} &= y_{i,t+h|t}^{de} - y_{t+h} \\
&= \rho^h y_t - y_{t+h} + \theta_i (\rho^h y_t - y_{i,t+h|t-1}^{de}) \\
&= \rho^h y_t - y_{t+h} + \theta_i (\rho^h y_t - y_{t+h} - FE_{i,t+h|t-1}^{de}) \\
&= FE_{t+h|t}^* + \theta_i (\rho^h y_t - y_{t+h} - FE_{i,t+h|t-1}^{de}) \\
&= (1 + \theta_i) FE_{t+h|t}^* - \theta_i FE_{i,t+h|t-1}^{de} \\
&= (1 + \theta_i) FE_{t+h|t}^* - \theta_i (\rho FE_{i,t+h-1|t-1}^{de} - \omega_{t+h}) \\
&= (1 + \theta_i) FE_{t+h|t}^* - \theta_i \rho FE_{i,t+h-1|t-1}^{de} + \theta_i \omega_{t+h} \\
&= (1 + \theta_i) FE_{t+h|t}^* + \theta_i (\omega_{t+h} - \rho FE_{i,t+h-1|t-1}^{de}) \\
&= (1 + \theta_i) FE_{t+h-1|t}^* + (1 + \theta_i) (-\omega_{t+h}) + \theta_i (\omega_{t+h} - \rho FE_{i,t+h-1|t-1}^{de}) \\
&= (1 + \theta_i) FE_{t+h-1|t}^* - \omega_{t+h} - \theta_i \rho FE_{i,t+h-1|t-1}^{de} \\
&= FE_{t+h|t}^* + \theta_i FE_{t+h-1|t}^* - \theta_i \rho FE_{i,t+h-1|t-1}^{de} \\
&= FE_{t+h|t}^* + \theta_i (FE_{t+h-1|t}^* - \rho FE_{i,t+h-1|t-1}^{de})
\end{aligned} \tag{43}$$

Reduced-form tests with forecast errors

The FE-based null-hypothesis of FIRE utilize the moment restrictions on forecast errors. In plain words, the null hypotheses of the three tests are the following. First, since the forecasts are on average unbiased according to FIRE, forecast errors across agents should converge to zero in a large sample. Second, forecast errors of non-overlapping forecasting horizon are not serially correlated. Third, forecast errors cannot be predicted by any information available at the time of the forecast, including the mean forecast itself and other variables that are in the agent's information set. This follows from Equation 2. In addition, I include what is called a weak version of the FE-based test which explores the serial correlation of forecast errors in overlapping periods, i.e. 1-year-ahead forecasts within one year. The forecast errors are correlated to the extent of the realized shocks in the overlapping periods. So the positive serial correlation does not directly violate FIRE. But the correlation of overlapping forecast errors still contains useful information about the size of the realized shocks.

Individual-level data are used whenever possible, utilizing the panel structure of both surveys. Since test 2 and 3 requires individual forecasts in vintages that are more

than one year apart while SCE only surveys each household for 12 months, the two tests are done with the population average expectations for SCE. Also, the regressions are adjusted accordingly depending on the quarterly and monthly frequency of SPF and SCE. Since these regressions are based on 1-year inflation in overlapping periods, Newy-West standard error is computed for hypothesis testing.

First, all three forecast series easily reject the null hypothesis of unbiasedness at the significance level of 0.1%. There are upward biases in both professional forecasts of core PCE inflation and households' forecast of headline inflation ³⁶, while at the same time professionals underpredicted core CPI inflation over the entire sample period. This was primarily driven by the under-prediction of the inflation over the recent two years since the Pandemic.

Second, the average point forecast one year ago predicts the forecast errors of both groups at the significance level of 0.1%. For headline CPI inflation, for instance, one percentage point inflation forecast corresponds to 0.35 percentage points of the forecast errors one year later. Thus, test 2 in Table 11 easily rejects the second hypothesis test of FIRE that past information does not predict future forecast errors. This suggests that both types of agents inefficiently utilize all information when making the forecasts.

Third, forecast errors are positively correlated with the forecast errors one year ago, with a significant coefficient ranging from 0.35 to 0.572. A higher positive autocorrelation coefficient of forecast errors by households is consistent with the common finding that households are subject to more information rigidity than attentive professionals.

Lastly, test 4 in Table 11 presents a higher serial correlation of forecast errors produced within a year. For SPF forecasts, the serial correlation does not exist beyond 2 quarters, implying the relative efficiency of professional forecasts. For the households, the forecast errors are more persistent over the entire year, in that current forecast errors are correlated with all past forecast errors over the past three quarters. Although the persistence of 1-year forecast errors within one year does not directly violate FIRE, the fact that households' forecast errors are more persistent than professionals' indicates that the former group is subject to a higher degree of rigidity than the latter one.

Estimates of stochastic volatility model of inflation

Figure 5 plots the estimated stochastic volatility of permanent and transitory component of inflation, respectively, as specified in Equation 31, using the same estimation method of Stock and Watson (2007).

Structural estimation covering Pandemic sample (2020-2023)

Table 13, 12 report the estimates for households and professionals, respectively, using only sample till March 2022 for AR(1) inflation process. Table 15, 14 report the estimates based on the alternative UCSV model of inflation process.

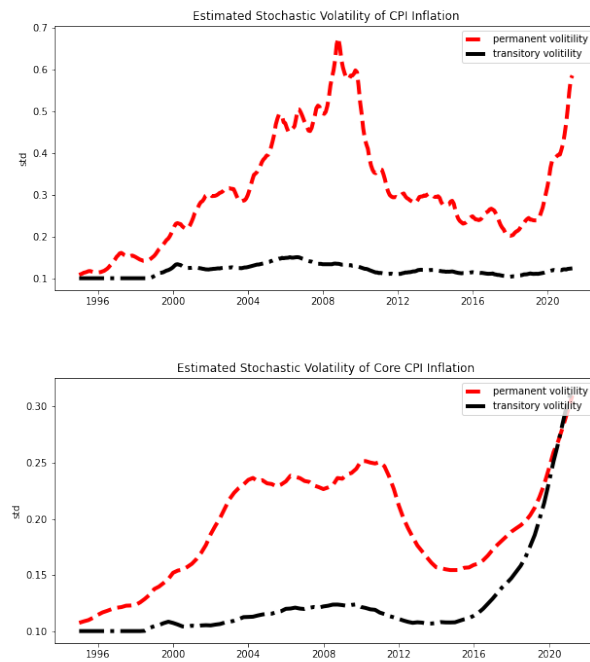
³⁶Coibion et al. (2018) finds the same upward bias for firms' managers.

Table 11: Tests of Rationality and Efficiency Using Forecast Errors

	SPF CPI	SPF PCE	SCE
Test 1: Bias			
Constant	-3.021*** (0.242)	0.460*** (0.047)	1.673*** (0.008)
N	5510	1610	112668
Test2: FE Depends on past information			
Forecast 1-yr before	0.350*** (0.035)	0.460*** (0.047)	4.190*** (0.659)
Constant	-3.452*** (0.386)	-2.333*** (0.192)	-12.92*** (2.213)
N	3945	1610	84
R ²	0.828	0.826	0.311
Test3: FE of non-overllaping forecast horizons are serially correlated			
Forecast Error 1-year before	0.350*** (0.035)	0.460*** (0.047)	0.572** (0.195)
Constant	0.314 (0.231)	-1.351*** (0.156)	-0.149 (0.445)
N	3945	1610	84
R ²	0.828	0.826	0.0957
Time FE	Yes	Yes	No
Test4: Overlapping FE are serially correlated			
Forecast Error 1-q before	0.502*** (0.060)	0.551*** (0.075)	0.327*** (0.010)
Forecast Error 2-q before	0.0901 (0.064)	0.231*** (0.060)	0.341*** (0.024)
Forecast Error 3-q before	0.146* (0.065)	0.0693 (0.052)	0.333*** (0.023)
Constant	1.147*** (0.224)	-0.356*** (0.058)	0.509*** (0.035)
N	2971	1338	4432
R ²	0.890	0.903	0.243

Note: white standard errors reported in the parentheses of estimations. *** p<0.001, ** p<0.01 and * p<0.05.

Figure 5: Stochastic Volatility of Inflation



Note: this figure plots the estimated stochastic volatility of permanent and transitory components of monthly headline CPI (top) and quarterly core CPI inflation (bottom) using the same approach as in [Stock and Watson \(2007\)](#).

Table 12: SMM Estimates of Different Models: Professionals

SE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\lambda}$	ρ	σ	$\hat{\lambda}$	ρ	σ				
FE	0.35	0.99	0.23	0.2	1	0.06				
FE+Disg	0.34	0.99	0.23							
FE+Disg+Var	0.35	0.99	0.23							
NI										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ		
FE	2.49	3	0.99	0.23	1.46	3	1	0.09		
FE+Disg	1.89	3	0.99	0.23	0	1.45	1	0.03		
FE+Disg+Var	0.39	3	0.99	0.23	0	0.28	0.9	0.45		
DE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\theta}$	σ_{θ}	ρ	σ	$\hat{\theta}$	σ_{θ}	ρ	σ		
FE	-0.07	0.61	0.99	0.23	-0.43	1.25	1	0.06		
FE+Disg	-0.29	1.87	0.99	0.23	0.97	1.67	0.91	0.44		
FE+Disg+Var	0.67	1.67	0.99	0.23	0.73	1.63	0.9	0.46		
DENI										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ
FE	0.05	0.61	0.31	0.99	0.23	0.77	2.74	3	0.99	0.11
FE+Disg	1.09	3	2.74	0.99	0.23	1.09	3	3	0.99	0.12
FE+Disg+Var	1.09	3	3	0.99	0.23	1.25	3	3	0.99	0.13

Table 13: SMM Estimates of Different Models: Households

SE										
Moments Used	2-Step Estimate			Joint Estimate						
	$\hat{\lambda}$	ρ	σ	$\hat{\lambda}$	ρ	σ				
FE	0.35	0.98	0.41	0.21	0.95	0.64				
FE+Disg	0.34	0.98	0.41							
FE+Disg+Var	0.35	0.98	0.41							
NI										
Moments Used	2-Step Estimate				Joint Estimate					
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ		
FE	0.01	2.89	0.98	0.41	0.12	0	0.95	0.65		
FE+Disg	0.02	2.91	0.98	0.41	0.02	2.05	0.95	0.63		
FE+Disg+Var	3	3	0.98	0.41	3	3	0.92	0.81		
DE										
Moments Used	2-Step Estimate				Joint Estimate					
	$\hat{\theta}$	σ_{θ}	ρ	σ	$\hat{\theta}$	σ_{θ}	ρ	σ		
FE	0.95	0.72	0.98	0.41	1.05	0.86	0.95	0.62		
FE+Disg	1.04	1.16	0.98	0.41	1.05	1.15	0.95	0.62		
FE+Disg+Var	1.04	1.16	0.98	0.41	1.12	1.14	0.93	0.75		
DENI										
Moments Used	2-Step Estimate					Joint Estimate				
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	ρ	σ
FE	0.79	0.02	2.98	0.98	0.41	0.78	0.02	0.74	0.95	0.62
FE+Disg				0.98	0.41	0.74	0.01	1.89	0.9	0.86
FE+Disg+Var	1.62	3	0.65	0.98	0.41	2.42	3	1.68	0.94	0.67

Table 14: SMM Estimates of Different Models under Stochastic Volatility: Professionals

SE			
Moments Used	2-Step Estimate		
	$\hat{\lambda}$		
FE	0.35		
FE+Disg	0.35		
FE+Disg+Var	0.35		
NI			
Moments Used	2-Step Estimate		
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	
FE	N/A	N/A	
FE+Disg	0.02	0	
FE+Disg+Var	N/A	N/A	
DE			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	σ_{θ}	
FE	-0.29	0.67	
FE+Disg	-0.29	0.75	
FE+Disg+Var	-0.29	0.75	
DENI			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$
FE	3	0.27	0.27
FE+Disg	3	2.84	0.24
FE+Disg+Var	3	3	0.25

Table 15: SMM Estimates of Different Models under Stochastic Volatility: Households

SE			
Moments Used	2-Step Estimate		
	$\hat{\lambda}$		
FE	0.35		
FE+Disg	0.35		
FE+Disg+Var	0.35		
NI			
Moments Used	2-Step Estimate		
	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$	
FE	0.05	2.96	
FE+Disg	0.05	2.99	
FE+Disg+Var	0.05	2.82	
DE			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	σ_{θ}	
FE	1.28	0.84	
FE+Disg	1.28	1.2	
FE+Disg+Var	1.28	1.2	
DENI			
Moments Used	2-Step Estimate		
	$\hat{\theta}$	$\hat{\sigma}_{pb}$	$\hat{\sigma}_{pr}$
FE	N/A	N/A	N/A
FE+Disg	0.71	0.01	0.95
FE+Disg+Var	N/A	N/A	N/A