

# Causal Inference, Time Series and Economic History

## 1. Introduction to Time Series Analysis

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# Overview

- An introduction to the course
  - Aims, format, outline and assessment
- An introduction to time series analysis
  - Primer on time series data
  - Derivation of Ordinary Least Squares (OLS) estimates
  - Assumptions under which OLS is the Best Linear Unbiased Estimator (BLUE)

## Part I: Introduction to the Course

# Aims

- To understand key quantitative methodologies in economic history
- To understand the principles of research design
  - This is a fundamental skill for research (whether it's a PhD dissertation or an academic paper)

# Format

- 6 classes
- We will study a mix of theory, simulations, paper replications and quantitative historiography
- Discussion papers are to be read *before* the seminar

# Course Outline

Week	Topic	Date	Time	Room
1	Introduction to time series analysis	Wednesday 11 May	9:30-12:30	Alfa1:1104
2	Stationarity, filtering and seasonal adjustment	Wednesday 11 May	14:30-17:30	Alfa1:1104
3	Single-equation models	Thursday 12 May	9:30-12:30	Alfa1:1104
4	Vector autoregressions	Thursday 12 May	14:30-17:30	Alfa1:1104
5	Narrative methods	Friday 13 May	9:30-12:30	Alfa1:1104
6	Instrumental variables and natural experiments	Friday 13 May	14:30-17:30	Alfa1:1104

# Assessment

- 2,500 word essay due Monday 5 September 2022 (100%)

## Part II: Introduction to Time Series Analysis



## Time Series Data ( $y_t$ )

- One unit (an individual, firm, industry, country, etc.) observed at multiple points in time, i.e.,  $i = 1, t > 1$
- Example: GDP per capita of Austria between 1870 and 1913, i.e.,  $i = 1, t = 44$
- Covered in this course

Table: GDP per Capita (1990 Int. GK\$)

Austria	
1870	1,863
1871	1,979
$\vdots$	$\vdots$
1913	3,465

## Cross Sectional Data ( $y_i$ )

- Multiple units (individuals, firms, industries, countries, etc.) observed at one point in time, i.e.,  $i > 1, t = 1$
- Example: GDP per capita of Western European countries in 1870, i.e.,  $i = 12, t = 1$
- Not covered in this course

Table: GDP per Capita (1990 Int. GK\$)

	Austria	Belgium	...	United Kingdom
1870	1,863	2,692	...	3,190

## Panel Data ( $y_{it}$ )

- Multiple units (individuals, firms, industries, countries, etc.) observed at multiple points in time, i.e.,  $i > 1, t > 1$
- Example: GDP per capita of Western European countries between 1870 and 1913, i.e.,  $i = 12, t = 44$
- Not covered in this course

Table: GDP per Capita (1990 Int. GK\$)

	Austria	Belgium	...	United Kingdom
1870	1,863	2,692	...	3,190
1871	1,979	2,682	...	3,332
$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$
1913	3,465	4,220	...	4,921

# The Time Series Regression Model

$$y_t = \alpha + \beta x_t + u_t \quad (1)$$

where  $y_t$  is the dependent variable

$x_t$  is the independent variable

$\alpha$  is the intercept

$\beta$  is the coefficient

$u_t$  is the error term

## Deriving OLS Estimates of $\alpha$ and $\beta$

The problem is to find values of  $\hat{\alpha}$  and  $\hat{\beta}$  to minimize the sum of squared residuals ( $SSR$ ):

$$SSR = \sum_{t=1}^n u_t^2 \quad (2)$$

The first step is to write equation (2) in terms of  $\hat{\alpha}$  and  $\hat{\beta}$

The time series regression model,  $y_t = \hat{\alpha} + \hat{\beta}x_t + u_t$ , shows that:

$$u_t = y_t - \hat{\alpha} - \hat{\beta}x_t \quad (3)$$

Therefore equation (2) can be re-written as:

$$SSR = \sum_{t=1}^n (y_t - \hat{\alpha} - \hat{\beta}x_t)^2 \quad (4)$$

## Deriving OLS Estimates of $\alpha$ and $\beta$

We can solve this minimisation problem using calculus by taking the partial derivative of  $SSR$ ,  $\sum_{t=1}^n (y_t - \hat{\alpha} - \hat{\beta}x_t)^2$ , with respect to  $\hat{\alpha}$  and  $\hat{\beta}$ :

$$\frac{\partial SSR}{\partial \hat{\alpha}} = -2 \sum_{t=1}^n (y_t - \hat{\alpha} - \hat{\beta}x_t) = 0 \quad (5)$$

$$\frac{\partial SSR}{\partial \hat{\beta}} = -2 \sum_{t=1}^n x_t (y_t - \hat{\alpha} - \hat{\beta}x_t) = 0 \quad (6)$$

## Deriving OLS Estimates of $\alpha$ and $\beta$

The goal now is to find an expression for  $\hat{\alpha}$ . Divide both sides of equation (5),  $-2 \sum_{t=1}^n (y_t - \hat{\alpha} - \hat{\beta}x_t) = 0$ , by -2:

$$\sum_{t=1}^n (y_t - \hat{\alpha} - \hat{\beta}x_t) = 0 \quad (7)$$

$$\sum_{t=1}^n y_t - \hat{\alpha} \sum_{t=1}^n 1 - \hat{\beta} \sum_{t=1}^n x_t = 0 \quad (8)$$

$$\sum_{t=1}^n y_t = \hat{\alpha} \sum_{t=1}^n 1 + \hat{\beta} \sum_{t=1}^n x_t \quad (9)$$

$$n\bar{y} = \hat{\alpha}n + \hat{\beta}n\bar{x} \quad (10)$$

As  $\sum_{t=1}^n y_t = n\bar{y}$ ,  $\sum_{t=1}^n x_t = n\bar{x}$

## Deriving OLS Estimates of $\alpha$ and $\beta$

Divide equation (10),  $n\bar{y}_t = \hat{\alpha}n + \hat{\beta}n\bar{x}_t$ , by  $n$ :

$$\bar{y}_t = \hat{\alpha} + \hat{\beta}\bar{x}_t \quad (11)$$

$$\hat{\alpha} = \bar{y}_t - \hat{\beta}\bar{x}_t \quad (12)$$



## Deriving OLS Estimates of $\alpha$ and $\beta$

The goal now is to find an expression for  $\hat{\beta}$ . Divide both sides of equation (6),  $-2 \sum_{t=1}^n x_t(y_t - \hat{\alpha} - \hat{\beta}x_t) = 0$ , by -2:

$$\sum_{t=1}^n x_t(y_t - \hat{\alpha} - \hat{\beta}x_t) = 0 \quad (13)$$

Replacing  $\hat{\alpha}$  in equation (13) with equation (12),  $\hat{\alpha} = \bar{y} - \hat{\beta}\bar{x}$ , gives:

$$\sum_{t=1}^n x_t[y_t - (\bar{y} - \hat{\beta}\bar{x}) - \hat{\beta}x_t] = 0 \quad (14)$$

Expanding out of the round brackets:

$$\sum_{t=1}^n x_t(y_t - \bar{y} + \hat{\beta}\bar{x} - \hat{\beta}x_t) = 0 \quad (15)$$

## Deriving OLS Estimates of $\alpha$ and $\beta$

Expanding out of the brackets again,  $\sum_{t=1}^n x_t(y_t - \bar{y} + \hat{\beta}\bar{x} - \hat{\beta}x_t) = 0$ :

$$\sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y} + \sum_{t=1}^n x_t \hat{\beta} \bar{x} - \sum_{t=1}^n x_t \hat{\beta} x_t = 0 \quad (16)$$

Bring the last two terms over to the right-hand side:

$$\sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y} = \sum_{t=1}^n x_t \hat{\beta} x_t - \sum_{t=1}^n x_t \hat{\beta} \bar{x} \quad (17)$$

$$\hat{\beta} \sum_{t=1}^n (x_t x_t - x_t \bar{x}) = \sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y} \quad (18)$$

$$\hat{\beta} = \frac{\sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y}}{\sum_{t=1}^n (x_t x_t - x_t \bar{x})} \quad (19)$$

## Deriving OLS Estimates of $\alpha$ and $\beta$

Collecting terms in equation (19),  $\hat{\beta} = \frac{\sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y}}{\sum_{t=1}^n (x_t x_t - x_t \bar{x})}$ :

$$\hat{\beta} = \frac{\sum_{t=1}^n x_t (y_t - \bar{y})}{\sum_{t=1}^n x_t (x_t - \bar{x})} \quad (20)$$

See the appendix for the details of this step:

$$\hat{\beta} = \frac{\sum_{t=1}^n (x_t - \bar{x})(y_t - \bar{y})}{\sum_{t=1}^n (x_t - \bar{x})^2} \quad (21)$$

$$\hat{\beta} = \frac{\text{Cov}(x_t, y_t)}{\text{Var}(x_t)} \quad (22)$$

## Example: Okun's Law

- Download “Okun's Law.xlsx” from Moodle
- Okun's law is an association between changes in real GDP and unemployment
- Let's try to calculate the intercept and slope in Excel and compare the results with the output in Stata

# Gauss-Markov Assumptions

- The Gauss-Markov assumptions are a set of criteria that if met mean that OLS is BLUE (Best Linear Unbiased Estimator)
- These assumptions are a crucial way to critically evaluate research design
- Violations of these assumptions has consequences for  $\hat{\beta}$  and  $se(\hat{\beta})$
- The five assumptions are listed in order of (my perceived) importance

# Gauss-Markov Assumption 1

Linear in parameters

- The time series process follows a model that is linear in parameters, such as:

$$y_t = \alpha + \beta x_t + u_t$$

$$y_t = \alpha + \beta \ln(x_t) + u_t$$

$$y_t = \alpha + \beta x_t + \gamma x_t^2 + u_t$$

- An example of a time series process that is non-linear in parameters:

$$y_t = \alpha + \beta^2 x_t + u_t$$

# Gauss-Markov Assumption 2

No perfect collinearity

- No independent variable is constant or a perfect linear combination of the others
- In other words, the independent variables must not be perfectly correlated
- For example, if we wanted to estimate:

$$y_t = \alpha + \beta x_t + \gamma z_t + u_t$$

- But:

$$z_t = \delta + \theta x_t$$

- Then we would have perfect collinearity
- Note that the absence of an error term makes the relationship exact

# Gauss-Markov Assumption 3

## Homoscedasticity

- Conditional on  $\mathbf{X}$ , the variance of  $u_t$  is the same for all  $t$ :  
$$\text{Var}(u_t \mid \mathbf{X}) = \text{Var}(u_t) = \sigma^2, t = 1, 2, \dots, n$$
- *Consequence*: Heteroscedasticity means that OLS standard errors are biased
- *Test*: Breusch-Pagan or White test for heteroscedasticity (see Wooldridge, *Introductory Econometrics*, chapters 8 and 10)
- *Potential solution*: Use an estimator that is robust to potential heteroscedasticity such as the Newey-West (1987) estimator



# Gauss-Markov Assumption 4

No serial correlation

- Conditional on  $\mathbf{X}$ , the errors in two different time periods are uncorrelated:  
 $Corr(u_t, u_s | \mathbf{X}) = 0$ , for all  $t \neq s$
- *Consequence:* Serial correlation means that OLS standard errors are biased
- *Test:* Breusch-Godfrey test (see Wooldridge, *Introductory Econometrics*, chapter 12)
- *Potential solution:* Use an estimator that is robust to potential serial correlation such as the Newey-West (1987) estimator

# Gauss-Markov Assumption 5

## Zero conditional mean

- For each  $t$ , the expected value of the error  $u_t$ , given the explanatory variables for *all* time periods, is zero:  $E(u_t | \mathbf{X}) = 0, t = 1, 2, \dots, n$
- Expressed differently,  $Cov(u_t, \mathbf{X}) = 0$
- In other words, the error term at time  $t$ ,  $u_t$ , is uncorrelated with each explanatory variable in *every* period
- *Consequence*: Non-zero conditional mean means that OLS coefficients are biased
- *Test*: Difficult to test
- *Potential solution*: Many! Some of which will be covered in this course

# Gauss-Markov Assumption 5

There are a number of reasons why the zero conditional mean assumption might fail:

1. Measurement error
2. Omitted variable bias
3. Reverse causality

# Measurement Error

- Measurement error is the difference between the observed variable and the true variable:  $e_t = x_t - x_t^*$
- For example, historical estimates of GDP are measured with error. In the United Kingdom in the late 19th century, measurement error is  $\pm 20$  per cent (Solomou and Weale, 1991)
- *Consequence:* Can lead to attenuation bias ( $|\hat{\beta}| < |\beta|$ )
- *Potential solution:* Collect more accurate data or instrumental variables

# Measurement Error

## Proof

We want to estimate the following equation:

$$y_t = \alpha + \beta x_t^* + u_t$$

But  $x_t^*$  is unobserved. As we only observe  $x_t$ , we actually estimate (as  $e_t = x_t - x_t^*$ , therefore  $x_t^* = x_t - e_t$ ):

$$y_t = \alpha + \beta(x_t - e_t) + u_t$$

$$y_t = \alpha + \beta x_t + (u_t - \beta e_t)$$

# Measurement Error

## Proof

Replacing  $u_t$  with the new residual term in our expression for the zero conditional mean assumption,  $Cov(u_t, x_t) = 0$ :

$$Cov(u_t - \beta e_t, x_t) = 0$$

Assuming  $u_t$  and  $x_t$  are uncorrelated:

$$-\beta Cov(e_t, x_t) = 0$$

Substituting  $x_t$  for  $x_t^* + e_t$  :

$$-\beta Cov(e_t, x_t^* + e_t) = 0$$

And assuming that the measurement error and the true variable are uncorrelated:

$$-\beta Cov(e_t, e_t) = 0$$

$$-\beta Var(e_t) \neq 0$$

# Measurement Error

## Example

- Download “Measurement Error.xlsx” from Moodle
- The “e\_multiplier” parameter controls the degree of time-varying measurement error
- The “e\_shifter” parameter controls the degree of time-invariant measurement error
- Let’s vary the degree of measurement error and see how  $\hat{\alpha}$  and  $\hat{\beta}$  differ from  $\alpha$  and  $\beta$

# Omitted Variable Bias

- Omitted variable bias arises when a relevant variable is omitted from the regression
- In other words, when a variable that is correlated with the dependent and independent variable is not included in the model
- *Consequence:* Omitted variable bias means that OLS coefficients are biased
- *Potential solution:* Include the omitted variable



# Omitted Variable Bias

## Proof

We want to estimate the following equation:

$$y_t = \alpha + \beta x_t + \gamma z_t + u_t$$

But instead we estimate:

$$y_t = \alpha + \beta x_t + e_t$$

where  $e_t = \gamma z_t + u_t$

Plugging  $e_t$  into our expression for the zero conditional mean assumption,

$Cov(u_t, x_t) = 0$ :

$$Cov(\gamma z_t + u_t, x_t) = 0$$

# Omitted Variable Bias

## Proof

Assuming that the population error and the included independent variable are uncorrelated:

$$\gamma \text{Cov}(z_t, x_t) = 0$$

Therefore, OLS is only unbiased if  $\gamma = 0$  (the omitted variable is not correlated with the dependent variable) or  $\text{Cov}(z_t, x_t) = 0$  (the included and omitted independent variables are not correlated with each other)

# Reverse Causality

- Reverse causality occurs when  $x_t$  not only affects but is affected by  $y_t$
- *Consequence:* Reverse causality means that OLS coefficients are biased
- *Potential solution:* Many

# Reverse Causality

## Proof

Suppose the population process is a system of equations:

$$y_t = \alpha + \beta x_t + u_t \quad (23)$$

$$x_t = \delta + \theta y_t + e_t \quad (24)$$

Consider this simple thought experiment:

1. Shock the error term in equation (23),  $u_t$
2.  $y_t$  changes in equations (23) and (24)
3.  $x_t$  changes in equations (23) and (24)

Therefore, there is a correlation between  $x_t$  and  $u_t$  that violates the zero conditional mean assumption,  $Cov(u_t, x_t) \neq 0$

# Reverse Causality

Direction of the bias in  $\hat{\beta}$

$$\hat{\beta} = \beta + \frac{\text{Cov}(u_t, x_t)}{\text{Var}(x_t)}$$

- If  $\beta$  is positive (negative) and the covariance term is positive (negative),  $\hat{\beta}$  will *overstate* the true absolute magnitude of the effect
- If  $\beta$  is positive (negative) and the covariance term is negative (positive),  $\hat{\beta}$  will *understate* the true absolute magnitude of the effect

## Next Class

- *Class discussion paper*: Edvinsson, R., 'New annual estimates of Swedish GDP, 1800-2010', *Economic History Review*, 66 (2013), pp. 1101-26.
- Stationarity, filtering and seasonal adjustment

## Further Reading

- Wooldridge, *Introductory Econometrics*, chapters 2 and 10
- Stock, J. H., and Watson, M. W., *Introduction to econometrics*, chapter 4

## Appendix: Equations (20)-(21)

Starting with the numerator in equation (21):

$$\sum_{t=1}^n (x_t - \bar{x})(y_t - \bar{y}) = \sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y} - \sum_{t=1}^n \bar{x} y_t + \sum_{t=1}^n \bar{x} \bar{y} \quad (25)$$

$$\sum_{t=1}^n x_t y_t - n\bar{x}\bar{y} - \bar{x}n\bar{y} + n\bar{x}\bar{y} \quad (26)$$

$$\sum_{t=1}^n x_t y_t - n\bar{x}\bar{y} \quad (27)$$

$$\sum_{t=1}^n x_t y_t - \sum_{t=1}^n x_t \bar{y} \quad (28)$$

$$\sum_{t=1}^n x_t (y_t - \bar{y}) \quad (29)$$

which is the numerator of equation (20)



## Appendix: Equations (20)-(21)

Moving on to the denominator in equation (21):

$$\sum_{t=1}^n (x_t - \bar{x})^2 = \sum_{t=1}^n (x_t - \bar{x})(x_t - \bar{x}) \quad (30)$$

$$\sum_{t=1}^n x_t x_t - \sum_{t=1}^n x_t \bar{x} - \sum_{t=1}^n \bar{x} x_t + \sum_{t=1}^n \bar{x} \bar{x} \quad (31)$$

$$\sum_{t=1}^n x_t x_t - n\bar{x}\bar{x} - n\bar{x}\bar{x} + n\bar{x}\bar{x} \quad (32)$$

$$\sum_{t=1}^n x_t x_t - n\bar{x}\bar{x} \quad (33)$$

$$\sum_{t=1}^n x_t x_t - \sum_{t=1}^n x_t \bar{x} \quad (34)$$

$$\sum_{t=1}^n x_t (x_t - \bar{x}) \quad (35)$$

which is the denominator of equation (20)