SISMID 2021: Mathematical Models of Infectious Disease Day 2 Breakout Session on Heterogeneity and Herd Immunity

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1 Introduction

As we've seen in lecture, the dynamics of infectious diseases spread are influenced by population heterogeneity. This is especially true for herd immunity, which occurs when susceptible individuals in a population are indirectly protected from infection due to immunity in others. The herd immunity threshold (HIT) is the fraction of the population that is non-susceptible when an unmitigated epidemic reaches its peak. In a population with homogeneous mixing, the HIT is $1-1/R_0$, where R_0 is the basic reproduction number; for example, this translates to an HIT of 67% using an R_0 of 3.

However, population homogeneity is an unrealistic assumption, and models incorporating heterogeneity in social exposure and infection susceptibility (defined as the probability of infection given exposure) generally result in lowered HITs [1–5]. The key idea behind these models is that subpopulations important for epidemic spread (i.e., those with substantially increased susceptibility or exposure) become infected – and thus develop immunity – early on in an epidemic's course. Herd immunity for the population overall is then achieved earlier because once these individuals are no longer susceptible to infection, further epidemic spread is slowed.

We will explore this effect using seroprevalence data on SARS-CoV-2 spread in New York City (NYC) and Long Island [6], focusing in particular on disparities in infection rates across racial and ethnic groups, and using models assuming proportionate or assortative mixing (see Slides 33-38 in Dr. Lipsitch's Heterogeneity lecture). Substantial racial and ethnic disparities in infection rates, hospitalizations, and deaths have been characterized across the US [7–13], but it is unclear how these heterogeneities in risk are expected to change over time, and what implications – if any – they have on overall epidemic dynamics. Seroprevalence studies characterize past exposure by identifying SARS-CoV-2 antibodies and are more reliable and unbiased than case data, which suffer from underreporting and other biases [14]. Identifying and building structured SEIR models using observed differences in infection rates across social strata is important for understanding how variation in ex-

posure or susceptibility and social disparities are interconnected. These models are also useful for designing interventions that can both reduce disparities and disrupt overall transmission by focusing efforts on groups most affected by high transmission rates [15, 16]. Note that the types of models we will explore today, however, are meant to be illustrative and intentionally oversimplified, and the numbers generated are not intended as best estimates for HITs in these locations.

2 Code overview

Let's start by loading in (or installing) the necessary packages (chunk 1):

```
# Load libraries needed for this computing session
# For missing packages, run install.packages("package")

library(tidyverse)

library(deSolve)

library(RColorBrewer)

library(reshape2)
```

and then loading in the datasets and parameters (chunk 2). Take a second to read through the example comment to make sure you understand the format of the census and serosurvey data:

data; proportion for "other" group is given by 1 minus sum of all other groups

sero.p.loc = adjusted cumulative incidence at given location for each demographic group

sero.tested.loc = number of people tested at given location for each demographic group

```
pop.nyc <- 8398748
pop.p.nyc <- c(0.319, 0.292, 0.217, 0.141, 0.032)
sero.p.nyc <- c(0.166, 0.330, 0.252, 0.145, 0.204)
sero.tested.nyc <- c(1758, 2103, 1392, 509, 184)
# Example: There are 8398748 people in NYC, of which 31.9% are non-Hispanic white. In the serosurvey, Rosenberg et al. tested 1758 non-Hispanic whites from NYC, of which 16.6% were seropositive (have SARS-CoV-2 antibodies, indicating they were infected at some point in the past).</pre>
```

```
pop.li <- 2839436
pop.p.li <- c(0.632, 0.186, 0.093, 0.068, 0.022)
sero.p.li <- c(0.087, 0.320, 0.158, 0.084, 0.207)
sero.tested.li <- c(1599, 301, 111, 50, 50)</pre>
```

Before we run code chunk 3, let's review some of the notation we're using for the SEIR model. Compared to the single-population, homogeneous models we've already seen, it's often convenient to use matrix notation to express a more general form of the SEIR model allowing for interactions between racial and ethnic grous (or other demographic groups of interest). The groups interact through a social contact matrix that governs the cross-talk between and within groups. The system of equations looks similar to the homogeneous case:

$$\frac{d\mathbf{S}}{dt} = -(\mathbf{BI}) \circ \mathbf{S} \tag{1}$$

$$\frac{d\mathbf{E}}{dt} = (\mathbf{BI}) \circ \mathbf{S} - r\mathbf{E} \tag{2}$$

$$\frac{d\mathbf{I}}{dt} = r\mathbf{E} - \gamma\mathbf{I} \tag{3}$$

$$\frac{d\mathbf{R}}{dt} = \gamma \mathbf{I} \tag{4}$$

but now S, E, I, R are column vectors comprising the compartmental variables for each group (i.e., $S = [S_0, ..., S_p]^T$ for p demographic groups), and \circ denotes element-wise multiplication. We let

 S_0 denote non-Hispanic whites, S_1 denote Hispanics or Latinos, S_2 denote non-Hispanic African-Americans, S_3 denote non-Hispanic Asians, and S_4 denote multiracial or other demographic groups, with similar ordering for elements in vectors ${\bf E}$ through ${\bf R}$. Given a mean incubation period and mean serial interval of 5 days as suggested by empirical studies [17, 18], we set the mean latent period 1/r to be 3 days to allow for pre-symptomatic transmission and the mean infectious period $1/\gamma$ to be 4 days to coincide with the observed serial interval.

Now we can run code chunk 3, which sets up the SEIR model outlined above:

```
# SEIR model with variable exposure for each demographic group via social mixing matrix;
   also allows for varying degrees of assortativity
# epsilon = 0 gives proportionate mixing and epsilon = 1 gives fully assortative mixing
exposure.SEIR <- function(t, x, parms){  # Set up function with three arguments
   with(as.list(c(t, x, parms)),{
                                            # "with" lets us directly use the variables in
       parms
       S = c(S0, S1, S2, S3, S4)
                                    # S vector comprises susceptible counts for
          all groups...
       E = c(E0, E1, E2, E3, E4)
                                           # ... and similarly for the other compartments
          E through R
       I = c(I0, I1, I2, I3, I4)
       R = c(R0, R1, R2, R3, R4)
       beta <- (1-epsilon)*outer(activities, activities)/sum(pop*pop.p*activities) +</pre>
         epsilon*activities/(pop*pop.p)*diag(5)
                                                  # Transmission rates are dependent on...
                                                  # the contact rates of the two groups
                                                  # epsilon allows for assortative mixing
       beta <- beta/scaling.factor</pre>
                                            # Scale beta so the RO matches what we
           specified (default.RO in the parameter section)
       dS <- -beta %*% I*S
                                            # System of rate equations in matrix
          formulation
       dE <- beta %*% I*S - r*E
       dI <- r*E - g*I
       dR <- g*I
       der <- c(dS, dE, dI, dR)
                                            # Combine results into a single vector der
       list(der)
                                             # Return result as a list
   })
```

We'll present the mathematics behind the proportionate mixing versus assortative mixing models here – feel free to jump into the code first and if you have time later, come back and read through this more closely. Note that the notation is a little different compared to Dr. Lipsitch's lecture notes, but the equations are all equivalent. The simplest variable exposure model we analyze will be the *proportionate mixing* model, which assumes that the contact rate for each pair of groups is proportional to the activity level of the two groups (i.e., total number of contacts per unit time for an individual of group i) [19]. Denoting a_i as the activity level for a member of group i and a as the $1 \times p$ vector of a_i s, the ijth entry in the transmission matrix is given by:

$$\beta_{i \leftarrow j} = q \frac{a_i a_j}{\sum_k a_k N_k} \tag{5}$$

and the overall transmission matrix B can be written as:

$$\mathbf{B} = \frac{q}{\sum_{k} a_k N_k} \mathbf{a} \mathbf{a}^T \tag{6}$$

That's what's being calculated in this line of code in chunk 3:

```
activities/(pop*pop.p)*diag(5)
```

The key unknown parameters are the activity levels. More complex variable exposure models allow for assortative mixing: we can partition a fraction ϵ of contacts to be **exclusively withingroup** and distribute the rest of the contacts according to proportionate mixing (with δ_{ij} being an indicator variable that is 1 when i = j and 0 otherwise) [19]:

$$\beta_{i \leftarrow j} = (1 - \epsilon) q \frac{a_i a_j}{\sum_k a_k N_k} + \epsilon \delta_{ij} q \frac{a_i}{N_i}$$
(7)

$$\mathbf{B} = \frac{(1 - \epsilon)q}{\sum_{k} a_{k} N_{k}} \mathbf{a} \mathbf{a}^{T} + \epsilon q \operatorname{diag}(\mathbf{a} \circ 1/\mathbf{N})$$
(8)

That's what's being calculated in this line of code in chunk 3:

The main takeaway is to generally understand how these models differ – assortative mixing allows individuals from group i to preferentially make contacts with other individuals from group i, which is a more accurate depiction of many real-world interactions.

Now run code chunk 4. We won't step through all the lines here but they are commented in the code if you want to read through later (see also the Mathematical appendix for more information). There are two main things this code chunk function does: 1) it re-scales the transmission rates so the R_0 is equal to the R_0 we set back in code chunk 2, and it also plots the next-generation matrix for each model. The next-generation matrix gives the average number of infections in group i (row) caused by an infected individual in group j (column). After that, run code chunk 5. This is a helper function that solves the model, calculates the final epidemic size and HIT, and plots the epidemic trajectory.

Finally, run code chunk 6 to load in the functions we need to fit these models using the seroprevalence data. In particular, we use the data to fit the activity level parameters (total number of contacts per unit time for an individual of group i) using maximum likelihood estimation. We will talk through this and other types of model fitting methods more in depth on Day 3, so for today we'll only briefly introduce the ideas. To conduct maximum likelihood, we first assume some probability distribution underlying how the data were generated or collected. In our case, we assumed that the seroprevalence data were collected via a binomial sampling process: at a given time point t_s representing the time of the serosurvey, the number of observed seropositive cases Y_i from group i is distributed $Bin(m_i, R_i/N_i)$, where m_i is the number of people tested from group i in the serosurvey and R_i/N_i is the fraction of recovered people from the SEIR model (i.e. the "true" seroprevalence rate). With this probability distribution in mind, we can then let R optimize and find the parameters that result in the highest likelihood (intuitively, the congruence between the data and the probability model).

Great! We've run through all the functions necessary to actually begin estimating parameters, simulating models, and interpreting the results. Run code chunk 7 – we'll test everything out by fitting a model to the Long Island data assuming proportionate mixing between demographic groups. The first line will fit the model using maximum likelihood and estimate and print out the activity levels.

The output should look something like this:

```
[1] "Normalized fitted values: 1 4.31 1.956 0.916 2.475"
```

The ordering of results by racial or ethnic group is non-Hispanic whites, Hispanics or Latinos, non-Hispanic Blacks or African-Americans, non-Hispanic Asians, and multiracial or other. The interpretation is that **under this proportionate mixing model**, Hispanics or Latinos have 4.31 times the activity level relative to non-Hispanic whites, non-Hispanic Blacks or African-Americans have 1.96 times the activity level relative to non-Hispanic whites, and so on. The next line will now take these activity level parameters and simulate the full epidemic, calculating the final epidemic size, the HIT, and the cumulative incidence of each group at the point that the HIT is reached. It will also generate two plots: the next-generation matrix plot, and an epidemic trajectory plot. In the first exercise, you'll discuss these results with your group and make sure that you all see the same numbers.

3 Exercises

- A) Run all of the code chunks 1-7: this will load all of the data and functions, conduct model fitting, and simulate the full (unmitigated) epidemic for Long Island assuming proportionate mixing. What do you observe? What is the overall estimated HIT and final epidemic size for the population? At the point that the HIT is reached, what is the cumulative incidence in each demographic group? Discuss in groups your findings (make sure your answers match) and also whether or not you think the activity level estimates are reasonable. What assumptions underlie this model, and are they reasonable / unreasonable?
- B) Now run code chunk 8, which conducts model fitting and simulation for Long Island assuming a high degree of assortative mixing ($\epsilon=0.5$ i.e. half of contacts are exclusively within group, and the other half are proportionately mixed). Compare the activity levels, next-generation matrix plots, the epidemic trajectory plots, and the overall HIT values versus those in question 3A. What do you observe? Why do you think there is a difference?
- C) Now, create a new code chunk where you conduct model fitting and simulation for NYC. Do this both for a proportionate mixing model and an assortative mixing model ($\epsilon=0.5$). How do your results compare to those of Long Island? Why do you think there is a difference (compare the seroprevalence data vectors for NYC and Long Island)?
- D) Test your ideas in 3C by creating your own vector of seroprevalence values. How does the HIT change when there is greater variation in seroprevalence across groups versus when there is little or even no variation? Compared to a homogeneous population, what do you think is the effect of modeling heterogeneity on HITs and final epidemic sizes?
- E) Think through and discuss with your group how you would modify the code to incorporate the effect of non-pharmaceutical interventions (e.g. which parameters would you try to vary, for how long, by what magnitude, etc.). If you're stuck, see our Jupyter notebook (https://nbviewer.jupyter.org/github/kevincma/covid19-race-ethnicity-model/blob/main/covid19-race-ethnicity-models-notebook.ipynb) or paper (https://elifesciences.org/articles/66601) for ideas on how to implement these additions to the model.

4 Appendix

We define contacts to be interactions between individuals that allow for transmission of SARS-CoV-2 with some non-zero probability. Following the convention for age-structured transmission models [20], we defined the $p \times p$ per capita social contact matrix \mathbf{C} to consist of elements $c_{i \leftarrow j}$ at row i and column j, representing the per capita rate that individuals from group i are contacted by individuals of group j. Letting N_i be the total number of individuals in group i, the social contact matrix \mathbf{M} consists of elements $m_{i \leftarrow j} = c_{i \leftarrow j} * N_i$, which represents the average number of individuals in group i encountered by an individual in group j. In the most general case, the susceptibility to infection can vary between groups, which can be modeled by allowing the probability of infection given contact with an infected individual to vary: $\mathbf{q} = [q_0, ..., q_p]^T$. Whether to build models that allow for variable susceptibility versus exposure will depend on the demographic groups being modeled. For instance, for some diseases, age-based variation in infection rates may be due to intrinsic biological differences in susceptibility by age to infection. In contrast, because observed disparities in infection rates in US cities are strongly attributable to differences in mobility and exposure, we focused primarily on building and analyzing the variable exposure models [21–23].

The transmission matrix $\mathbf B$ is then given by $(\mathbf q \mathbf 1^T) \circ \mathbf C$, where $\mathbf 1^T$ is a 1 by p vector of 1s:

$$\mathbf{B} = \begin{bmatrix} q_0 & \dots & q_0 & \dots & q_0 \\ \vdots & & \vdots & & \vdots \\ q_4 & \dots & q_4 & \dots & q_4 \end{bmatrix} \circ \begin{bmatrix} c_{0\leftarrow 0} & \dots & c_{0\leftarrow 2} & \dots & c_{0\leftarrow 4} \\ \vdots & & \vdots & & \vdots \\ c_{4\leftarrow 0} & \dots & c_{4\leftarrow 2} & \dots & c_{4\leftarrow 4} \end{bmatrix}$$
(9)

Given mean duration of infectiousness $1/\gamma$, the next-generation matrix \mathbf{G} , representing the average number of infections in group i caused by an infected individual in group j, is given by $(q1^T)\circ\mathbf{M}/\gamma=\mathbf{N}\circ\mathbf{B}/\gamma$. R_0 for the overall population described by this structured model was calculated by computing the dominant eigenvalue of matrix \mathbf{G} . The effective reproduction number R_t at time t was calculated by computing the dominant eigenvalue of $\mathbf{G_t}=(\mathbf{q}1^T)\circ\mathbf{M_t}/\gamma=\mathbf{S_t}\circ\mathbf{B}/\gamma$, where the elements in $\mathbf{M_t}$ are given by $c_{i\leftarrow j}*S_{i,t}$ and $S_{i,t}$ is the number of susceptible individuals in group i at time t. To hold R_0 values across model types constant when calculating HITs and final epidemic sizes, we re-scaled transmission matrices to have the same dominant eigenvalue.

Calculating the HIT and final epidemic size

We defined the herd immunity threshold (HIT) for all models as the fraction of non-susceptible peo-

ple when the effective reproduction number R_t first crosses 1. In the homogenous model, where $R_t = S_t \beta/\gamma$, the analytical solution for the HIT occurs when the fraction of non-susceptible individuals equals 1-1/ R_0 . In the structured models of heterogeneous populations, the HIT was calculated via simulation: we took the dominant eigenvalue of G_t at each timestep to calculate R_t , and identified the number of non-susceptible individuals when R_t first decreased below 1. Final epidemic sizes were calculated by simulation by running the epidemic out to 1 year.

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