

HW 3

variance = 1/precision

$$a) y_i | \tau \stackrel{iid}{\sim} N(m, \frac{1}{\tau}) = \frac{1}{\sqrt{2\pi \frac{1}{\tau}}} \cdot e^{-\frac{(y_i - m)^2}{2 \cdot \frac{1}{\tau}}}$$

$$\Rightarrow p(y_1, \dots, y_n | \tau) = \prod_{i=1}^n \frac{1}{\sqrt{\frac{1}{\tau}} \sqrt{2\pi}} \cdot e^{-\frac{1}{2} \cdot (y_i - m)^2 \cdot \tau}$$

$$= \left(\frac{\sqrt{\tau}}{\sqrt{2\pi}} \right)^n \cdot e^{-\frac{\tau n}{2} \sum_{i=1}^n (y_i - m)^2}$$

$$b) \tau(\tau | y_1, \dots, y_n) \propto p(y_1, \dots, y_n | \tau) \cdot p(\tau)$$

$$= \left(\frac{\sqrt{\tau}}{\sqrt{2\pi}} \right)^n \cdot \frac{b^n}{\Gamma(a)} \tau^{a-1} \cdot e^{-b\tau} \cdot e^{-\frac{\tau n}{2} \sum_{i=1}^n (y_i - m)^2}$$

$$\begin{aligned} \text{Consider } \sum_{i=1}^n (y_i - m)^2 &= \sum_{i=1}^n (y_i - \bar{y} + \bar{y} - m)^2 \\ &= \sum_{i=1}^n [(y_i - \bar{y})^2 + (\bar{y} - m)^2 + 2(y_i - \bar{y})(\bar{y} - m)] \\ &\propto \sum_{i=1}^n [(y_i - \bar{y})^2 + 2(y_i - \bar{y})(\bar{y} - m)] \\ &= n(\bar{y} - m)^2 + 2(\bar{y} - m) \sum_{i=1}^n (y_i - \bar{y}) \\ &= n(\bar{y} - m)^2 + 2(\bar{y} - m)(n\bar{y} - n\bar{y}) \\ &= n(\bar{y} - m)^2 \end{aligned}$$

$$\begin{aligned} \Rightarrow e^{-b\tau} \cdot e^{-\frac{\tau n}{2} \sum_{i=1}^n (y_i - m)^2} &= e^{-\tau(b + \frac{n}{2} (\bar{y} - m)^2)} \\ &= e^{-b\tau + \frac{\tau n^2}{2} (\bar{y} - m)^2} = e^{-\tau(b + \frac{n}{2} (\bar{y} - m)^2)} \\ \Rightarrow p(y_1, \dots, y_n | \tau) &\propto \left(\frac{1}{\sqrt{2\pi}} \right)^n \cdot \frac{b^n}{\Gamma(a)} \cdot \tau^{a + \frac{n}{2} - 1} \cdot e^{-\tau(b + \frac{n}{2} (\bar{y} - m)^2)} \\ &\propto \tau^{a+n-1} \cdot e^{-\tau(b + \frac{n}{2} (\bar{y} - m)^2)} \end{aligned}$$

$$= \text{Gamma}\left(a + \frac{n}{2}, b + \frac{n}{2} (\bar{y} - m)^2\right)$$

$$2) \theta \sim N(m, c^2) \Rightarrow \theta | y \sim N\left(\frac{\frac{1}{c^2} m + \frac{n}{s^2} \bar{y}}{\frac{1}{c^2} + \frac{n}{s^2}}, \frac{1}{\frac{1}{c^2} + \frac{n}{s^2}}\right)$$

$$\Rightarrow E[\tilde{y}] = E[E[\tilde{y} | \theta]] = E[\theta] (= m)$$

$$\begin{aligned} 1) \text{Var}[\tilde{y}] &= E[\text{Var}[\tilde{y} | \theta]] + \text{Var}[E[\tilde{y} | \theta]] \\ &= E[s^2] + \text{Var}[\theta] = s^2 + \text{Var}[\theta] (= s^2 + c^2) \end{aligned}$$

$$\Rightarrow \tilde{y} \sim N(m, s^2 + c^2)$$

$$\Rightarrow \tilde{y} | y \sim N\left(\frac{\frac{1}{c^2}}{\frac{1}{c^2} + \frac{n}{s^2}} m + \frac{\frac{n/s^2}{1/c^2 + n/s^2}} \bar{y}, s^2 + \frac{1}{1/c^2 + n/s^2}\right)$$

4. Essentially, the argument is that since our posterior is basically just our prior updated with data, the data helps lower the variance. Let \tilde{y} be a prediction variable.

$$\text{Var}[\theta] = E[\text{Var}[\theta | \tilde{y}] + \text{Var}[E[\theta | \tilde{y}]]]$$

$$\text{Var}[\theta | y] = E[\text{Var}[(\theta | y) | \tilde{y}]] + \text{Var}[E[(\theta | y) | \tilde{y}]]$$

\Rightarrow Consider $E[\text{Var}[\theta | \tilde{y}]]$ to be the average variability when predicting the value of \tilde{y} given only the prior.

The average variability of predicting \tilde{y} using the prior tuned by the data is less than without the data, because the information lowers the uncertainty associated with our uncertain guessing of a prior $\Rightarrow E[\text{Var}[\theta | \tilde{y}]] \geq E[\text{Var}[(\theta | y) | \tilde{y}]]$

$$\Rightarrow \text{Var}[E[\theta | \tilde{y}]] = \text{Var}[\tilde{y}]$$

$$\text{Var}[E[(\theta | y) | \tilde{y}]] = \text{Var}[\tilde{y}]$$

so they're actually the same

Then, considering the 4 terms making up $\text{Var}[\theta]$ and $\text{Var}[\theta | y]$ and how they compare to each other, $\text{Var}[\theta] \geq \text{Var}[\theta | y]$.

$$5. p(\lambda | y) = \text{Gamma}(a + \sum_i y_i, b + n)$$

$$= \frac{(b+n)^{a+\sum_i y_i}}{\Gamma(a+\sum_i y_i)} \lambda^{(a+\sum_i y_i)-1} e^{-(b+n)\lambda}$$

for prior $\lambda \sim \text{Gamma}(a, b)$, which is our prior, too

$$\Rightarrow p(\tilde{y} | y) = \int \frac{(b+n)^{a+\sum_i y_i}}{\Gamma(a+\sum_i y_i)} \cdot \lambda^{(a+\sum_i y_i)-1} e^{-(b+n)\lambda} \cdot \frac{e^{-\lambda \tilde{y}}}{\tilde{y}!} d\lambda$$

$$= \frac{(b+n)^{a+\sum_i y_i}}{\tilde{y}! \Gamma(a+\sum_i y_i)} \cdot \int (\lambda)^{\tilde{y} + a + \sum_i y_i - 1} e^{-\lambda(\tilde{y} + b + n)} d\lambda$$

\downarrow Gamma

$$= \frac{(b+n)^{a+\sum_i y_i}}{\tilde{y}! \Gamma(a+\sum_i y_i)} \cdot \frac{\Gamma(\tilde{y} + a + \sum_i y_i)}{(b+n+1)^{\tilde{y} + a + \sum_i y_i}}$$

$$= \left(\frac{b+n}{b+n+1} \right)^{a+\sum y_i} \cdot \frac{1}{(b+n+1)^{\tilde{y}}} \cdot \begin{pmatrix} \tilde{y} + a + \sum y_i - 1 \\ \tilde{y} \end{pmatrix}$$

$$= \begin{pmatrix} \tilde{y} + a + \sum y_i - 1 \\ \tilde{y} \end{pmatrix} \cdot \left(\frac{b+n}{b+n+1} \right)^{a+\sum y_i} \cdot \left(1 - \frac{b+n}{b+n+1} \right)^{\tilde{y}}$$

$$\Rightarrow p(\tilde{y}^a | y) \sim \text{NB} \left(a + \sum y_i, \frac{b+n}{b+n+1} \right)$$