

## Prop 2.1

i) Lasso solutions exist

ii)  $X\hat{\beta}_\lambda^L$  is unique

### Proof

i) Provided  $\lambda > 0$

$$\begin{aligned} \inf_{\beta: \lambda \|\beta\|_1 > \frac{1}{2n} \|Y\|_2^2} Q_\lambda(\beta) &\geq \frac{1}{2n} \|Y\|_2^2 = Q_\lambda(0) \\ &\geq \inf_{\beta: \lambda \|\beta\|_1 \leq \frac{1}{2n} \|Y\|_2^2} Q_\lambda(\beta) \quad (*) \\ &= \inf_{\beta} Q_\lambda(\beta) \end{aligned}$$

But at (\*) we are minimizing the str<sup>n</sup>  $Q_\lambda$  over a closed bounded set, so a minimiser must exist.

ii) Fix  $\lambda \geq 0$  and suppose  $\hat{\beta}^{(1)}$  and  $\hat{\beta}^{(2)}$  are two Lasso solutions with

$$Q_\lambda(\hat{\beta}^{(1)}) = Q_\lambda(\hat{\beta}^{(2)}) = c^*$$

By strict convexity  $\|\cdot\|_2^2$ ,

$$\left\| \frac{1}{2}(Y - X\hat{\beta}^{(1)}) + \frac{1}{2}(Y - X\hat{\beta}^{(2)}) \right\|_2^2 \leq \frac{1}{2} \|Y - X\hat{\beta}^{(1)}\|_2^2 + \frac{1}{2} \|Y - X\hat{\beta}^{(2)}\|_2^2$$

with equality iff  $X\hat{\beta}^{(1)} = X\hat{\beta}^{(2)}$ .

$$c^* \leq Q_\lambda\left(\frac{1}{2}\hat{\beta}^{(1)} + \frac{1}{2}\hat{\beta}^{(2)}\right)$$

$$\leq \frac{1}{2n} \left( \frac{1}{2} \|Y - X\hat{\beta}^{(1)}\|_2^2 + \frac{1}{2} \|Y - X\hat{\beta}^{(2)}\|_2^2 \right) + \lambda \left\| \frac{1}{2}\hat{\beta}^{(1)} + \frac{1}{2}\hat{\beta}^{(2)} \right\|_1 \quad (*)$$

$$\leq \frac{1}{2} Q_\lambda(\hat{\beta}^{(1)}) + \frac{1}{2} Q_\lambda(\hat{\beta}^{(2)}) = c^*$$

Therefore, we must have equality at (\*), so  $X\hat{\beta}^{(1)} = X\hat{\beta}^{(2)}$ .  $\square$

Define the spike correlation set  $\hat{E}_\lambda$  to be the set of variables  $k$  s.t.

$$\frac{1}{n} |X_k^T (Y - X\hat{\beta}_\lambda^L)| = \lambda.$$

This is well-defined as it only depends on the (unique) fitted vals. By the KKT conditions,  $\hat{E}_\lambda$  contains the set of non-zeroes of all Lasso solutions (at  $\lambda$ ).

If  $\text{rank}(X_{\hat{E}_\lambda}) = |\hat{E}_\lambda|$ , then the Lasso soln is unique:

$$X_{\hat{E}_\lambda}(\hat{\beta}_{\hat{E}_\lambda}^{(1)} - \hat{\beta}_{\hat{E}_\lambda}^{(2)}) = 0 \Rightarrow \hat{\beta}_{\hat{E}_\lambda}^{(1)} = \hat{\beta}_{\hat{E}_\lambda}^{(2)} \Rightarrow \hat{\beta}^{(1)} = \hat{\beta}^{(2)}$$

## 2.2.5 Variable selection

Noiseless linear model  $Y = X\beta^0$

$$S = \{k : \beta_k^0 \neq 0\} = \{1, \dots, s\}$$

$$N = \{1, \dots, p\} \setminus S$$

Assume  $\text{rank}(X_S) = s$ .

Thm 14

Let  $\lambda > 0$  and  $\Delta = X_N^T X_S (X_S^T X_S)^{-1} \text{sgn}(\beta_S^0)$ .

i) If  $\|\Delta\|_\infty \leq 1$  ( $\max_{k \in N} |\text{sgn}(\beta_S^0)^T (X_S^T X_S)^{-1} X_S^T X_k| \leq 1$ )

$$|\beta_k^0| > \lambda \left| \text{sgn}(\beta_S^0)^T \left\{ \left( \frac{1}{n} X_S^T X_S \right)^{-1} \right\}_k \right| \quad (k \in S) \quad (*)$$

then  $\exists$  a Lasso soln  $\hat{\beta}_\lambda^L$  with  $\text{sgn}(\hat{\beta}_\lambda^L) = \text{sgn}(\beta^0)$

ii) If  $\exists$  a Lasso soln  $\hat{\beta}_\lambda^L$  with  $\text{sgn}(\hat{\beta}_\lambda^L) = \text{sgn}(\beta^0)$

then  $\|\Delta\|_\infty \leq 1$ .

Proof: Write  $\hat{\beta} = \hat{\beta}_\lambda$ ,  $\hat{S} = \{k : \hat{\beta}_k \neq 0\}$ . KKT conditions:

$$\frac{1}{n} X^T X (\beta^0 - \hat{\beta}) = \lambda \hat{v}, \quad \|\hat{v}\|_\infty \leq 1 \text{ and } \hat{v}_S = \text{sgn}(\hat{\beta}_S)$$

Expand this

$$\frac{1}{n} \begin{pmatrix} X_S^T X_S & X_S^T X_N \\ X_N^T X_S & X_N^T X_N \end{pmatrix} \begin{pmatrix} \beta_S^0 - \hat{\beta}_S \\ -\hat{\beta}_N \end{pmatrix} = \lambda \begin{pmatrix} \hat{v}_S \\ \hat{v}_N \end{pmatrix} \quad \begin{matrix} (1) \\ (2) \end{matrix}$$

Prove (ii) first

$$\text{From (1)} \quad \beta_S^0 - \hat{\beta}_S = \lambda \left( \frac{1}{n} X_S^T X_S \right)^{-1} \text{sgn}(\beta_S^0)$$

Substituting into (2)

$$\lambda \underbrace{\frac{1}{n} X_N^T X_S \left( \frac{1}{n} X_S^T X_S \right)^{-1} \text{sgn}(\beta_S^0)}_{\Delta} = \lambda \hat{v}_N$$

Then  $\|\Delta\|_\infty \leq 1$  as  $\|\hat{v}_N\|_\infty \leq 1$

Now (i). Try  $(\hat{\beta}_S, \hat{\beta}_N) = (\beta_S^0 - \lambda(\frac{1}{n} X_S^T X_S)^{-1} \text{sgn}(\beta_S^0), 0)$

$$(\hat{v}_S, \hat{v}_N) = (\text{sgn}(\beta_S^0), \Delta)$$

Only need to check  $\text{sgn}(\hat{\beta}_S) = \text{sgn}(\beta_S^0)$ , but this follows from (\*).

## 2.2.6 Prediction estimation

Now consider  $Y = X\beta^0 + \varepsilon - 1\bar{\varepsilon}$  where  $\varepsilon_i$  are independent mean-zero sub-Gaussian with parameter  $\sigma$ .

Let  $S, s, N$  be defined as before.

### Def<sup>n</sup> 6

$$\text{Define } \phi^2 = \inf_{\substack{\beta \in \text{RP}: \|\beta_N\|_1 \leq 3\|\beta_S\|_1 \\ \|\beta_S\|_1 \neq 0}} \frac{\frac{1}{n} \|X\beta\|_2^2}{\frac{1}{s} \|\beta_S\|_1^2}$$

where compatibility factor  $\phi \geq 0$ . The compatibility condition is that  $\phi > 0$ .