

Lecture 6: Random Vectors

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The multivariate normal - diagonal covariance case

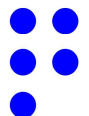
Multivariate normal - diagonal covariance case

- Take the special case of N independent random variables X_1, \dots, X_N each distributed according to a normal with known mean and variance:

$$X_i \sim N(\mu_i, \sigma_i^2)$$

- Seen as a random vector, the **joint pdf** of these variables is:

$$\begin{aligned} \mathbf{X} &= (X_1, \dots, X_N) \\ p(\mathbf{x}) &= p(x_1, \dots, x_N) = \prod_{i=1}^N p(x_i) \stackrel{\substack{\text{because} \\ \text{X}_i\text{'s are} \\ \text{independent}}}{=} \prod_{i=1}^N N(x_i | \mu_i, \sigma_i^2) \\ &= \prod_{i=1}^N (2\pi)^{-1/2} \sigma_i^{-1} \cdot \exp\left\{-\frac{1}{2} \frac{(x_i - \mu_i)^2}{\sigma_i^2}\right\} \end{aligned}$$



Multivariate normal - diagonal covariance case

- Take the special case of N independent random variables X_1, \dots, X_N each distributed according to a normal with known mean and variance:

$$X_i \sim N(\mu_i, \sigma_i^2)$$

- Seen as a random vector, the **mean** of these variables is:

$$\mathbb{E}[\mathbf{X}] = \begin{pmatrix} \mathbb{E}[X_1] \\ \vdots \\ \mathbb{E}[X_N] \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_N \end{pmatrix} = \underline{\mu}$$

$\mathbf{X} = (X_1, \dots, X_N)$

Multivariate normal - diagonal covariance case

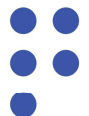
- Take the special case of N independent random variables X_1, \dots, X_N each distributed according to a normal with known mean and variance:

$$X_i \perp X_j, \text{ if } i \neq j \quad X_i \sim N(\mu_i, \sigma_i^2)$$

$\mathbb{C}[X_i, X_j] = 0, \mathbb{C}[X_i, X_i] = V[X_i] = \sigma_i^2$

- Seen as a random vector, the **covariance matrix** of these variables is:

$$\mathbb{C}[\mathbf{X}, \mathbf{X}] = \begin{bmatrix} \sigma_1^2 & & & \\ & \sigma_2^2 & & \\ & & \ddots & \\ & & & \sigma_N^2 \end{bmatrix} = \text{diag}(\sigma_1^2, \dots, \sigma_N^2)$$



Multivariate normal - diagonal covariance case

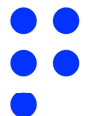
- We say that the distribution of such a random vector is a **multivariate normal** with mean vector $\boldsymbol{\mu}$ and covariance matrix $\text{diag}(\sigma_1^2, \dots, \sigma_N^2)$.

- We write:

* bottom bar denotes vector

$$\mathbf{X} \sim \mathcal{N}\left(\underline{\boldsymbol{\mu}}, \text{diag}(\sigma_1^2, \dots, \sigma_N^2)\right)$$

$$p(\underline{\mathbf{x}}) = (2\pi)^{-N/2} \underbrace{|\text{diag}(\sigma_1^2, \dots, \sigma_N^2)|^{-1/2}}_{\det(\text{diag}(\sigma_1^2, \dots, \sigma_N^2))} \cdot \exp\left\{-\frac{1}{2}(\underline{\mathbf{x}} - \underline{\boldsymbol{\mu}})^T \text{diag}(\sigma_1^2, \dots, \sigma_N^2)^{-1/2} (\underline{\mathbf{x}} - \underline{\boldsymbol{\mu}})\right\}$$



Isotropic covariance

$$p(\mathbf{x}) = (2\pi)^{-N/2} \sigma^{-N} \cdot \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^N (x_i - \mu_i)^2 \right\}$$

- For the special case where all the variances are the same and equal to σ^2 , we write:

$$\mathbf{X} \sim \mathcal{N}(\underline{\mu}, \sigma^2 \underline{\mathbf{I}})$$

- In this special case the joint PDF is:

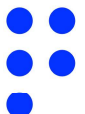
unit matrix in \mathbb{R}^N

$$\underline{\mathbf{I}} = \begin{pmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & \dots & 1 \end{pmatrix}$$

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \underline{\mu}, \sigma^2 \underline{\mathbf{I}})$$

$$= (2\pi)^{-N/2} \cdot \underbrace{|\sigma^2 \underline{\mathbf{I}}|^{-1/2}}_{\sigma^{-N}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \underline{\mu})^T \underbrace{(\sigma^2 \underline{\mathbf{I}})^{-1}}_{\sigma^{-2} \underline{\mathbf{I}}} (\mathbf{x} - \underline{\mu}) \right\}$$

$$= \dots$$

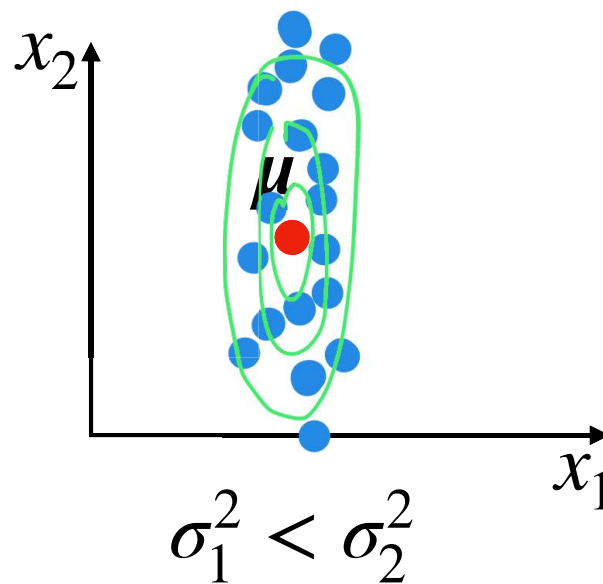
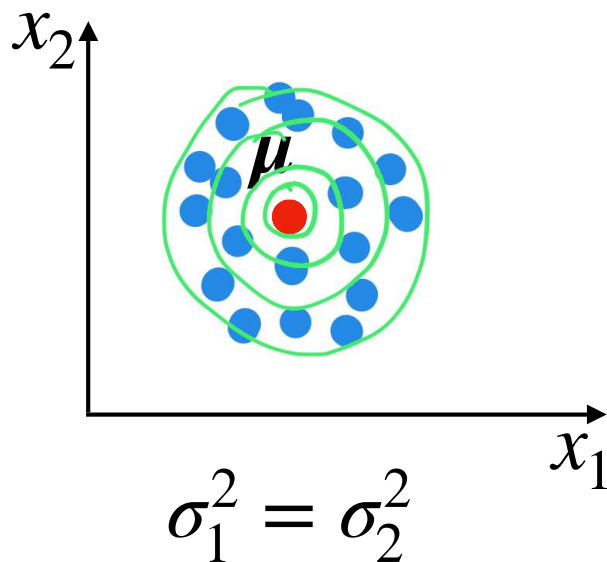


Visualizing the joint PDF of the multivariate normal with diagonal

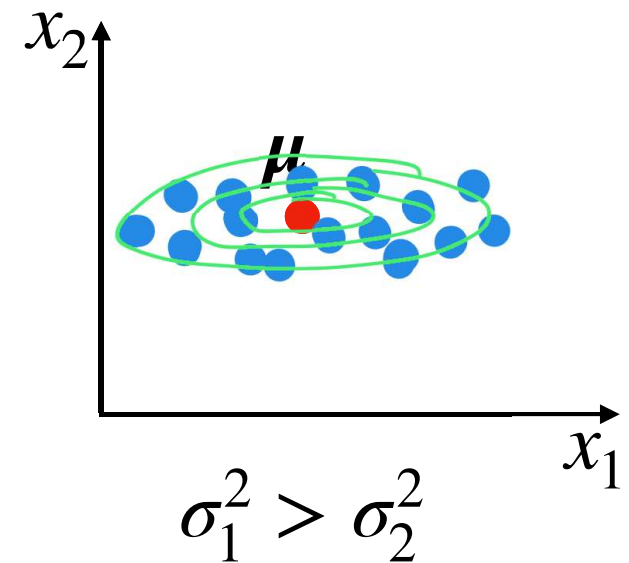
covariance

$$\mathbf{X} \sim \mathcal{N}\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{pmatrix}\right) \cdot p(\mathbf{x}) \propto \exp\left\{-\frac{1}{2\sigma_1^2}(x_1 - \mu_1)^2 - \frac{1}{2\sigma_2^2}(x_2 - \mu_2)^2\right\}$$

contour : a line $p(\mathbf{x}) = \text{const}$ proportional to



x_2 varies more



x_1 varies more

Connection to the standard normal

- Let $\mathbf{Z} \sim N(\mathbf{0}, \mathbf{I})$ be a collection of independent standard normal random variables.
- Define the random vector:

$$\mathbf{X} = \boldsymbol{\mu} + \text{diag}(\sigma_1, \dots, \sigma_N)\mathbf{Z}$$

- Then:

$$\mathbf{X} \sim N(\underline{\boldsymbol{\mu}}, \text{diag}(\sigma_1^2, \dots, \sigma_N^2))$$