

$$1, a) \quad P(I_{i=1} | \pi, \mu_1, \mu_2, \sigma_1, \sigma_2, x, I_i) \propto P(x | I, \pi, \mu_1, \mu_2, \sigma_1, \sigma_2) \cdot P(I_{i=1})$$

$$= \prod_{i=1}^n P(x_i | I_{i=1}, \pi, \mu_1, \mu_2, \sigma_1, \sigma_2) \cdot \pi$$

$$\propto P(x_i | I_{i=1}, \pi, \mu_1, \mu_2, \sigma_1, \sigma_2) \cdot \pi$$

$$= \phi(x_i | \mu_1, \sigma_1^2) \cdot \pi$$

p.s.s.

$$P(I_{i=2} | \pi, \mu_1, \mu_2, \sigma_1, \sigma_2, I_{i=1}) \propto \phi(x_i | \mu_2, \sigma_2^2) (1-\pi)$$

So

$$P(I_{i=1} | \cdot) = \frac{\pi \cdot \phi(x_i; \mu_1, \sigma_1^2)}{\pi \phi(x_i; \mu_1, \sigma_1^2) + (1-\pi) \phi(x_i; \mu_2, \sigma_2^2)}$$

$$1b) \quad P(\lambda | \beta, \sigma^2, y, x) = \frac{P(\beta, \sigma^2, \lambda | y, x)}{P(\beta, \sigma^2 | y, x)} \propto P(\beta, \sigma^2, \lambda | y, x)$$

$$P(\beta, \sigma^2, \lambda | y) \propto P(y | \beta, \sigma^2, \lambda) P(\beta, \sigma^2, \lambda)$$

$$= P(y | \beta, \sigma^2, \lambda) P(\beta | \sigma^2, \lambda) P(\sigma^2 | \lambda) P(\lambda)$$

$$= \prod_{i=1}^n P(y_i | \beta, \sigma^2, \lambda) \cdot P(\beta | \sigma^2, \lambda) P(\sigma^2 | \lambda) P(\lambda)$$

$$= \prod_{i=1}^n \underbrace{\frac{1}{\sqrt{2\pi}\sigma^2} \exp\left(-\frac{1}{2\sigma^2} (y_i - x_i^T \beta)^T (y_i - x_i^T \beta)\right)}_{\text{constant with respect to } \lambda} \cdot P(\beta | \sigma^2, \lambda) P(\sigma^2 | \lambda) P(\lambda)$$

$$So, \quad P(\lambda | \beta, \sigma^2, y, x) \propto P(\beta | \sigma^2, \lambda) \underbrace{P(\sigma^2 | \lambda)}_{\text{const.}} P(\lambda)$$

$$\text{Now } \beta | \sigma^2, \lambda \sim N(0, \lambda^{-1} I_p)$$

$$\sigma^2 | \lambda \sim \text{Inv-}\chi^2 \quad \text{Does not depend on } \lambda$$

$$\lambda \sim \text{Inv-}\chi^2(\eta_0, \lambda_0)$$

$$P(\lambda | \beta, \sigma^2, y, x) \propto \prod_{i=1}^p \frac{1}{\sqrt{2\pi} \lambda} \exp\left(-\frac{1}{2 \cdot \frac{1}{\lambda}} \beta_i^2\right) \cdot \frac{\exp\left(-\frac{\eta_0 \lambda_0}{2\lambda}\right)}{\lambda^{1+\eta_0/2}}$$

$$\propto \lambda^{p/2} \exp\left(-\frac{1}{2} \sum_{i=1}^p \beta_i^2\right) \exp\left(-\frac{\eta_0 \lambda_0}{2\lambda}\right) \cdot \lambda^{-(\eta_0/2+1)}$$

$$= \lambda^{p/2 - \eta_0/2 - 1} \exp\left(-\frac{\lambda}{2} \sum_{i=1}^p \beta_i^2 - \frac{\eta_0 \lambda_0}{2\lambda}\right)$$

Not so good... Better if prior was of the form

$$\lambda^a \exp(-\lambda b)$$

This is the Gamma distribution.

The way to go.

$$\frac{1}{\lambda} \sim \text{Inv-}\chi^2(\eta_0, \lambda_0) \Rightarrow$$

$$\frac{1}{\lambda} \sim \text{Inv Gamma}\left(\frac{\eta_0}{2}, \frac{\eta_0 \lambda_0}{2}\right) \Rightarrow$$

$$\lambda \sim \text{Gamma}\left(\alpha = \frac{\eta_0}{2}, \beta = \frac{\eta_0 \lambda_0}{2}\right)$$

$$\frac{1}{\lambda} \sim \text{Inv-}\chi^2(\eta_0, \lambda_0) \Rightarrow \frac{1}{\lambda} \sim \text{Inv-Gamma} \left(\frac{\eta_0}{2}, \frac{\eta_0 \lambda_0}{2} \right)$$

$$\Rightarrow \lambda \sim \text{Gamma} \left(\frac{\eta_0}{2}, \frac{\eta_0 \lambda_0}{2} \right) \quad p(\lambda) \propto \exp\left(-\frac{\nu \chi^2}{2\lambda}\right) \lambda^{-(\nu+1)}$$

$$p(\lambda | \beta, \sigma^2, y) \propto \lambda^{p/2} \exp\left(-\lambda \frac{\sum \beta_i^2}{2}\right) \lambda^{\eta_0/2-1} \exp\left(-\lambda \frac{\eta_0 \lambda_0}{2}\right)$$

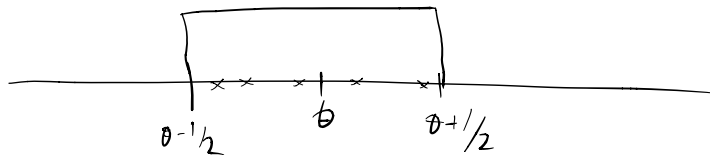
$$= \lambda^{(p+\eta_0)/2-1} \exp\left(-\lambda \left(\frac{\sum \beta_i^2 + \eta_0 \lambda_0}{2}\right)\right)$$

$$\lambda | \beta, \sigma^2, y \sim \text{Ga} \left(\frac{p+\eta_0}{2}, \frac{\sum \beta_i^2 + \eta_0 \lambda_0}{2} \right)$$

$$\frac{1}{\lambda} \sim \text{Inv-}\chi^2 \left(p+\eta_0, \frac{\sum \beta_i^2 + \eta_0 \lambda_0}{p+\eta_0} \right)$$

$$E\left(\frac{1}{\lambda}\right) \approx \frac{\sum \beta_i^2 + \eta_0 \lambda_0}{p+\eta_0} \approx \frac{\sum \beta_i^2}{p} = \text{Var}(\beta_i) \quad \text{when } p \gg \eta_0$$

2a)



$$\text{Var}(\bar{X}) = \frac{\sigma^2}{n} \quad \sigma^2 = \text{Var}(X_i) \quad X_i \sim U(\theta - 1/2, \theta + 1/2)$$

$$X \sim U(0,1) \quad \text{Var}(X) = \frac{1}{12}$$

$$\text{So } \text{Var}(\bar{X}) = \frac{1}{12n}$$

$$2b) \quad P(\theta | x_1, \dots, x_n) \propto P(x_1, \dots, x_n | \theta) P(\theta)$$

$$= \prod_{i=1}^n P(x_i | \theta) P(\theta)$$

$$= \prod_{i=1}^n \mathbb{I}(\theta - 1/2 \leq x_i \leq \theta + 1/2) \cdot 1$$



$$\begin{aligned} \theta + 1/2 \geq x_{\max} \\ \theta - 1/2 \leq x_{\min} \end{aligned} \Rightarrow \theta \in [x_{\max} - 1/2, x_{\min} + 1/2]$$



$$P(\theta | x_1, \dots, x_n) \propto 1 \quad \text{for } \theta \in [x_{\max} - 1/2, x_{\min} + 1/2]$$

= otherwise.

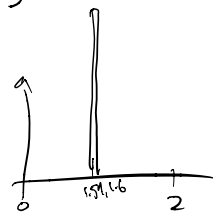
$$\theta | x_1, \dots, x_n \sim U(x_{\max} - 1/2, x_{\min} + 1/2)$$

$$2c) \quad \text{Frequentist: } \hat{\theta} = \bar{X} = 1.53$$

$$\text{Var}(\hat{\theta}) = \frac{1}{12n} = \frac{1}{12 \cdot 3} = 0.027777$$

$$\text{SD}(\hat{\theta}) = 0.1666$$

$$\text{Bayesian: } \theta | x_1, x_2, x_3 \sim U(1.59, 1.6)$$



$$3a) \quad \theta | x_1, \dots, x_n \sim \text{Beta}(\alpha+s, \beta+f)$$

$$p(\theta|x) \propto \theta^{\alpha+s-1} (1-\theta)^{\beta+f-1} \Rightarrow \ln p(\theta|x) \propto (\alpha+s-1) \ln \theta + (\beta+f-1) \ln(1-\theta)$$

$$\frac{\partial \ln p(\theta|x)}{\partial \theta} = \frac{\alpha+s-1}{\theta} + \frac{\beta+f-1}{1-\theta} (-1)$$

$$\frac{\partial \ln p(\theta|x)}{\partial \theta} = 0 \Rightarrow \frac{\alpha+s-1}{\theta} = \frac{\beta+f-1}{1-\theta}$$

$$\Rightarrow \hat{\theta} = \frac{\alpha+s-1}{\alpha+\beta+n-2}$$

$$3b) \quad \theta | x_1, \dots, x_n \sim^{\text{approx}} N(\hat{\theta}, -\mathcal{I}_{\theta}^{-1}(\hat{\theta})) \quad 1-\hat{\theta} = \frac{\beta+f-1}{\alpha+\beta+n-2}$$

$$\frac{\partial^2 \ln p(x|\theta)}{\partial \theta^2} = -\frac{\alpha+s-1}{\theta^2} + \frac{\beta+f-1}{(1-\theta)^2} (-1)$$

$$\left. \frac{\partial^2 \ln p(x|\theta)}{\partial \theta^2} \right|_{\theta=\hat{\theta}} = - \left(\frac{\alpha+s-1}{\left(\frac{\alpha+s-1}{\alpha+\beta+n-2} \right)^2} + \frac{\beta+f-1}{\left(\frac{\beta+f-1}{\alpha+\beta+n-2} \right)^2} \right)$$

$$= -(\alpha+\beta+n-2)^2 \left(\frac{1}{\alpha+s-1} + \frac{1}{\beta+f-1} \right)$$

$$= -(\alpha+\beta+n-2)^2 \left(\frac{\alpha+\beta+n-2}{(\alpha+s-1)(\beta+f-1)} \right)$$

$$= - \frac{(\alpha+\beta+n-2)^3}{(\alpha+s-1)(\beta+f-1)}$$

$$\theta | x_1, \dots, x_n \sim \text{approx } N \left(\hat{\theta} = \frac{\alpha + s - 1}{\alpha + \beta + n - 2}, \quad J_{\hat{\theta}, x}^{-1} = \frac{(\alpha + s - 1)(\beta + f - 1)}{(\alpha + \beta + n - 2)^3} \right)$$

$$\begin{aligned} \text{Check: } \theta &\sim \text{Beta}(\alpha + s, \beta + f) \\ \text{Var}(\theta) &= \frac{(\alpha + s)(\beta + f)}{(\alpha + s + \beta + f)^2 (\alpha + s + \beta + f + 1)} \\ &= \frac{(\alpha + s)(\beta + f)}{(\alpha + \beta + n)^2 (\alpha + \beta + n + 1)} \end{aligned}$$

3 c) See R-code

3 d) ———, ———

$$4 a) \quad n_A = 5 \quad n_B = 5 \quad n_C = 13$$

$$\begin{aligned} P(A | T_1, T_2) &\propto P(T_1, T_2 | A) P(A) \\ &= P(T_1 | A) \cdot P(T_2 | A) P(A) \end{aligned}$$

$$\begin{array}{ccc} P(A) & P(B) & P(C) \\ \theta_1 & \theta_2 & \theta_3 \end{array} \quad \text{From Dirichlet}$$

$$(\theta_1, \theta_2, \theta_3) | n_A, n_B, n_C \sim \text{Dirichlet}(5+1, 5+1, 13+1) \\ \text{Dirichlet}(6, 6, 14)$$

$$E(\theta_1 | n_A, n_B, n_C) = \frac{6}{23} \approx 0.26$$

$$E(\theta_2 | \cdot) = \frac{6}{23} \approx 0.26$$

$$E(\theta_3 | \cdot) = \frac{11}{23} \approx 0.48$$

$$P(T_i | A) \quad N(\mu_{iA}, 1)$$

$$\mu_{1A} | \bar{x}_1 = 1.2 \sim N(1.2, \frac{1}{5})$$

$$\mu_{1B} | \bar{x}_1 = 1.4 \sim N(1.4, \frac{1}{5})$$

$$\mu_{1C} | \bar{x}_1 = 0.7 \sim N(0.7, \frac{1}{10})$$

$$\mu_{2A} | \bar{x}_2 = 2.1 \sim N(2.1, \frac{1}{5})$$

$$\mu_{2B} | \bar{x}_2 = 3.5 \sim N(3.5, \frac{1}{5})$$

$$\mu_{2C} | \bar{x}_2 = 4.7 \sim N(4.7, \frac{1}{10})$$

$$\boxed{\mu \sim N(\bar{x}, \sigma^2)}$$

$$P(A | T_1 = 1.3, T_2 = 4.2) \propto P(T_1 = 1.3 | A) \cdot P(T_2 = 4.2 | A) \cdot P(A)$$

$$= \phi(1.3, \mu = 1.2, \sigma^2 = \frac{1}{5}) \cdot \phi(4.2, \mu = 2.1, \sigma^2 = \frac{1}{5}) \cdot \frac{6}{23}$$

$$P(B | T_1 = 1.3, T_2 = 4.2) \propto \phi(1.3 | 1.4, \frac{1}{5}) \phi(4.2, 3.5, \frac{1}{5}) \cdot \frac{6}{23}$$

$$P(C | T_1 = 1.3, T_2 = 4.2) \propto \phi(1.3 | 0.7, \frac{1}{10}) \cdot \phi(4.2 | 4.7, \frac{1}{10}) \cdot \frac{11}{23}$$