Problem 1

$$Var(Yij) = Var(M+bi+eij)$$

$$= Var(bi) + Var(eij) \quad (bi \perp eij)$$

$$= \sigma_b^2 + \sigma_e^2$$

$$(bv(Yij, Yik) = E(Yij - EYij)(Yik - EYik)$$

$$= E(Yij - M)(Yik - M)$$

$$= E(bi + eij)(bi + eik)$$

$$= E(bi^2 + bi eik + bi eij + eij eik)$$

$$= Ebi^2 + Ebi Eeik + Ebi Eeij + Eeij Eeik \quad (eij \perp eik, bi \perp ei)$$

$$= Ebi^2$$

$$= Var(bi) + (Ebi)^2$$

$$= \sigma_b^2$$

$$= \sigma_b^2$$

$$= \sigma_b^2 + \sigma_b^2$$

$$= \sigma_b^2 + \sigma_b^2$$

Both the variance and correlation is constant across visits, so this corresponds to compound symmetry pattern.

Problem 2 Question 2

Note: Answers to Questions 1 & 3 of Problem 2 are attached after this.

$$Y_{ij} = \beta_0 + a_i + b_0 * I_{(sex_i=0)} + b_1 * I_{(sex_i=1)} + \beta_1 * age_{ij} + e_{ij}$$

$$E(Y_{ij}) = E(\beta_0 + \alpha_i + b_0)(\text{sex}_{i=0}) + b_1 I_{(sex_i=1)} + \beta_1 * age_{ij} + e_{ij})$$

$$= \beta_0 + \beta_1 \cdot \alpha_2 e_{ij}$$

$$\text{var}(Y_{ij}) = \text{var}(\alpha_i + b_k + e_{ij}) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i + b_k + e_{ij}) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

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$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(b_k) + \text{var}(e_{ij}) = \sigma_0^2 + \sigma_b^2 + \sigma_e^2$$

$$\text{Var}(\alpha_i) = \text{var}(\alpha_i) + \text{var}(\beta_i) + \text{var}$$

Detween-subject covariance:

(i) if the subjects are of the same gender:

$$CoV(Y_{mj}, Y_{nk}) = E[(Y_{mj} - E_{mj})(Y_{nk} - E_{nk})]$$
 $= E[(A_m + b_k + e_{mj})(A_n + b_k + e_{nk})]$
 $= E[b_k^2]$
 $= O_b^2$

(ii) if the subjects are of different genders:

Based on the above calculation. We construct matrices A and B:

$$A = \begin{bmatrix} \sigma_a^2 t \sigma_b^2 t \sigma_e^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 \\ \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 t \sigma_e^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 \\ \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 \\ \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 + \sigma_e^2 \\ \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 + \sigma_e^2 \\ \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 & \sigma_a^2 t \sigma_b^2 + \sigma_e^2 \\ \sigma_b^2 & \sigma_b^2 & \sigma_b^2 & \sigma_b^2 \\ \sigma_b^2 & \sigma_b^2 & \sigma_b^2 & \sigma_b^2 \\ \sigma_b^2 & \sigma_b^2 & \sigma_b^2 & \sigma_b^2 \\ \sigma_b^2 & \sigma_b^2$$

Where A4x4 is the covariance within a subject, and B is the covariance between measurements from 2 subjects of the same gender.

Therefore, we get

Choy =
$$\begin{bmatrix} A B & \cdots & B \\ B A & \cdots & B \\ \vdots & \vdots & \ddots & \vdots \\ B B B & A \end{bmatrix}_{ibxib}, \quad Dgirl = \begin{bmatrix} A B & \cdots & B \\ B A & \cdots & B \\ \vdots & \vdots & \ddots & \vdots \\ B B B & A \end{bmatrix}_{ilxil},$$

and finally

p8131_hw6_xy2395

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Problem 2

2.1 Spaghetti Plot

```
# Import data
dental <-
  read.table('HW6-dental.txt', header = TRUE) %>%
  as.tibble() %>% janitor::clean_names() %>%
  mutate(gender = as.factor(gender))
# Spaghetti plot
dental %>%
ggplot(aes(x = age, y = distance, group = child, color = gender)) +
  geom_line() +
  theme_bw()
   32 -
   28
                                                                                       gender
distance <sup>54</sup>
                                                                                        — 0
                                                                                          • 1
   20
                                10
                                                       12
                                                                               14
                                           age
```

2.2 Marginal Form

Answer to this question is hand-written above.

2.3 Comparing models with different covariance patterns

For the following 3 models, we assume equal variance across measurements at different ages.

```
# Compound Symmetry covariance
compsym = gls(distance ~ gender + age,
              data = dental,
              correlation = corCompSymm(form = ~1 | child),
              method="REML")
# Exponential covariance
expo = gls(distance ~ gender + age,
           data = dental,
           correlation = corExp(form = ~1 | child),
           method = 'REML')
# Autoregressive covariance
auto1 = gls(distance ~ gender + age,
            data = dental,
            correlation = corAR1(form = ~1 | child),
            method = 'REML')
# Compare coefficient parameter estimates
bind rows(
  compsym$coefficients,
  expo$coefficients,
  auto1$coefficients,
) %>%
  mutate(CovType = c('CompSym', 'Exp', 'Auto')) %>%
  select(CovType, everything()) %>%
  knitr::kable()
```

CovType	(Intercept)	gender1	age
CompSym	15.38569	2.321023	0.6601852
Exp	15.45999	2.418714	0.6529597
Auto	15.45999	2.418714	0.6529597

The coefficient parameter estimates are similar across the 3 covariance patterns.

```
# Compare covariance estiamtes
# Compound Symmetry
compsym$sigma^2 * corMatrix(compsym$modelStruct$corStruct)[[1]]
##
            [,1]
                     [,2]
                              [,3]
                                        [,4]
## [1,] 5.316240 3.266784 3.266784 3.266784
## [2,] 3.266784 5.316240 3.266784 3.266784
## [3,] 3.266784 3.266784 5.316240 3.266784
## [4,] 3.266784 3.266784 3.266784 5.316240
# Exponential covariance
expo$sigma^2 * corMatrix(expo$modelStruct$corStruct)[[1]]
            [,1]
                     [,2]
                              [,3]
## [1,] 5.296881 3.315144 2.074839 1.298574
## [2,] 3.315144 5.296881 3.315144 2.074839
## [3,] 2.074839 3.315144 5.296881 3.315144
## [4,] 1.298574 2.074839 3.315144 5.296881
```

Autoregressive covariance

auto1\$sigma^2 * corMatrix(auto1\$modelStruct\$corStruct)[[1]]

```
## [,1] [,2] [,3] [,4]

## [1,] 5.296881 3.315144 2.074840 1.298574

## [2,] 3.315144 5.296881 3.315144 2.074840

## [3,] 2.074840 3.315144 5.296881 3.315144

## [4,] 1.298574 2.074840 3.315144 5.296881
```

The covariance estiamtes are similar across the 3 covariance patterns.