

2. Find the conditional distribution $T_1, \dots, T_n | T_{n+1} = t$.

$$f_{T_1, \dots, T_n | T_{n+1}=t}(t_1, \dots, t_n, t) = \frac{f_{T_1, \dots, T_{n+1}}(t_1, \dots, t_{n+1})}{f_{T_{n+1}}(t)}$$

Note the interarrival times are distributed $\text{Exp}(\lambda)$ and the Jacobian of the following transformation is

$$J = \begin{bmatrix} 1 & & & \\ -1 & 1 & & \\ & \ddots & \ddots & \\ & & -1 & 1 \end{bmatrix} \quad \text{so } |J| = 1$$

Thus

$$\begin{aligned} f_{T_1, \dots, T_{n+1}}(t_1, \dots, t_{n+1}) &= f_{T_1, \dots, T_{n+1}}(t_1, t_2 - t_1, \dots, t_n - t_{n-1}, t_{n+1} - t_n) |J| \\ &= \prod_{k=1}^{n+1} \lambda e^{-\lambda(t_{k+1} - t_k)} \\ &= \lambda^{n+1} e^{-\lambda \sum_{k=1}^{n+1} (t_{k+1} - t_k)} \\ &= \lambda^{n+1} e^{-\lambda t_{n+1}} \end{aligned}$$

Using the transformation again and writing $T_{n+1} = \sum_{k=1}^{n+1} T_{k+1} - T_k$ it follows T_{n+1} is a sum of exponential variables and thus distributed as $T_{n+1} \sim \Gamma(n+1, \lambda)$.

Therefore

$$\begin{aligned} f_{T_1, \dots, T_n | T_{n+1}=t}(t_1, \dots, t_n, t) &= \frac{\cancel{\lambda^{n+1}} e^{-\cancel{\lambda t_{n+1}}}}{\frac{\cancel{\lambda^{n+1}} t^{n+1-1} e^{-\cancel{\lambda t_{n+1}}}}{\Gamma(n+1)}} \\ &= \frac{n!}{t^n} \end{aligned}$$

Thus $T_1, \dots, T_n | T_{n+1}=t$ is distributed as the order statistics of $\text{Uniform}[0, t]$

3. Busloads of customers arrive at an infinite server queue according to a Poisson process with rate λ . Each customer is served independently with common service time distribution having distribution function G . A bus contains j customers with probability α_j , $j = 1, 2, 3, \dots$. Let $X(t)$ denote the number of customers that have has service completed by time t .

(a) Compute $E[X(t)]$.

(b) Does $X(t)$ have a Poisson distribution?

a) Let μ be the average number of customers per bus.

Let a bus arrive at time $s < t$. Then

$$E[\text{customers from } s \text{ served by } t] = \mu G(t-s)$$

It follows

$$\begin{aligned} E[X(t)] &= \lambda \int_0^\infty \mu G(t-s) \mathbb{I}\{s \leq t\} ds \\ &= \lambda \mu \int_0^t G(t-s) ds \end{aligned}$$

b) We can view $X(t)$ as a time dependent thinning of the original Poisson process because being served can be used as a time dependent labeling. Thus $X(t) \sim \text{Pois}(\lambda \mu \int_0^t G(t-s) ds)$ by thrm in class.