

When not otherwise stated, $\{N(t) : t \geq 0\}$ is a Poisson process with rate $\lambda > 0$ and event times $\{T_n : n \geq 1\}$.

3. Let $0 < s < t$.

(a) Find the conditional distribution of $N(s)$ given $N(t)$.

(b) Compute $E[T_k | N(t)]$ for all $k \geq 1$.

$$\begin{aligned}
 a) P(N(s)=k | N(t)=n) &= \frac{P(N(s)=k, N(t)=n)}{P(N(t)=n)} \\
 &= \frac{P(N(s)=k, N(t)-N(s)=n-k)}{P(N(t)=n)} \\
 &= \frac{P(N(s)=k) P(N(t)-N(s)=n-k)}{P(N(t)=n)} \\
 &= \frac{e^{-\lambda s} (\lambda s)^k}{k!} \frac{e^{-\lambda(t-s)} (\lambda(t-s))^{n-k}}{(n-k)!} \frac{n!}{e^{-\lambda t} (\lambda t)^n} \\
 &= \binom{n}{k} \left(\frac{s}{t}\right)^k \left(1 - \frac{s}{t}\right)^{n-k} \\
 &= \binom{n}{k} \left(\frac{s}{t}\right)^k \left(1 - \frac{s}{t}\right)^{n-k}
 \end{aligned}$$

$$\text{So } \sum_{k=0}^n P(N(s)=k | N(t)=n) = \sum_{k=0}^n \binom{n}{k} \left(\frac{s}{t}\right)^k \left(1 - \frac{s}{t}\right)^{n-k}$$

$$\therefore N(s) | N(t)=n \sim \text{Bin}(n, \frac{s}{t})$$

b) Note $\{T_n \leq t\} \Rightarrow \{N(t) \geq n\}$

$$\begin{aligned}
 P(T_n > s | N(t)=n) &= P(N(s) < k | N(t)=n) \\
 &= \sum_{j=0}^{k-1} \binom{n}{j} \left(\frac{s}{t}\right)^j \left(1 - \frac{s}{t}\right)^{n-j}
 \end{aligned}$$

If $n < k$, $P(T_n > s | N(t)=n) = P(N(s) < k | N(t)=n) = 0$.

This is because $N(u) = \sum_{n=1}^{\infty} \mathbb{1}_{\{T_n \leq u\}}$ so $N(t)$ cannot equal n if $N(s) < k$ and $s < t$.

$$E[T_k | N(t)=n] = \int_0^\infty P(T_k > s | N(t)=n) ds$$

$$\begin{aligned}
 &= \int_0^\infty P(N(s) < k | N(t)=n) ds \\
 &= \int_0^t \sum_{j=0}^{k-1} \binom{n}{j} \left(\frac{s}{t}\right)^j \left(1 - \frac{s}{t}\right)^{n-j} ds \quad \text{because } s < t \\
 &= \sum_{j=0}^{k-1} \binom{n}{j} \int_0^t \left(\frac{s}{t}\right)^j \left(1 - \frac{s}{t}\right)^{n-j} ds \\
 &= \sum_{j=0}^{k-1} \binom{n}{j} \int_0^1 u^j (1-u)^{n-j} t du \quad u = \frac{s}{t} \quad du = \frac{1}{t} ds \\
 &= + \sum_{j=0}^{k-1} \binom{n}{j} \frac{\Gamma(n-j+1) \Gamma(j+1)}{\Gamma(n-j+j+2)} \quad \text{Beta normalization constant} \\
 &= + \sum_{j=0}^{k-1} \frac{n!}{j!(n-j)!} \frac{(n-j)! j!}{(n+1)!} \\
 &= + \sum_{j=0}^{k-1} \frac{1}{n+1} \\
 &= + \frac{k}{n+1}
 \end{aligned}$$

4. Let $\{M(t) : t \geq 0\}$ be a Poisson process with rate μ and event times $\{S_n : n \geq 1\}$ independent of $\{N(t)\}$ and $\{T_n\}$. Determine if the process $\{L(t) : t \geq 0\}$ is a Poisson process, and if so compute its rate.

- (a) $L(t) = N(t) + M(t)$.
 (b) $L(t)$ is the counting process on $[0, \infty)$ with event times $R_n = \min\{S_n, T_n\}$, $n \geq 1$.
 (c) $L(t)$ counts the number of intervals $(k-1, k]$, $k = 1, 2, \dots, N(t)$ for which $M(k-1, k] = 0$, i.e., on which the process $\{M(t)\}$ sees no events.

$$a) L(0) = N(0) + M(0) = 0$$

$$\begin{aligned} L(+)-L(s) &= N(t)+M(t)-(N(s)+M(s)) \\ &= N(t)-N(s)+M(t)-M(s) \end{aligned}$$

So L has stationary increments

$$G_{NH}(s) = \sum_{k=0}^{\infty} s^k \frac{e^{-\lambda}}{k!} \lambda^k = e^{-\lambda} e^{s\lambda} = e^{\lambda(s-1)}$$

$$G_{m(\tau)}(s) = e^{M(s-\tau)}$$

$$G_{N(s)}(s)G_{M(s)}(s) = e^{\lambda(s-1)} e^{M(s-1)} = e^{(\lambda+M)(s-1)} = G_{(N+M)s}(s) = G_{L(s)}(s)$$

Thus the rate is $\lambda + \mu$

$$b) \bullet L(t) = \sum_{n=1}^{\infty} I\{\min(T_n, S_n) \leq t\}$$

$$L(0) = \sum_{n=1}^{\infty} I\{\min(T_n, S_n) \leq 0\} = 0 \text{ since } T_n, S_n \geq 0$$

• WLOG assume $T_n = \min(T_n, S_n)$

$$L(t) - L(s) = \sum_{n=1}^{\infty} I\{T_n \leq t\} - I\{T_n \leq s\}$$

$$= N(t) - N(s)$$

Thus L has stationary increments

$$R_n = \min\{T_n, S_n\}$$

$$\begin{aligned} P(R_n \leq x) &= 1 - P(R_n > x) \\ &= 1 - P(T_{n+1} > x) P(S_n > x) \\ &= 1 - e^{-(x+\lambda)} t^x, \end{aligned}$$

Thus the rate is $\lambda + \mu$

$$C) \text{ From class } L(t) = \sum_{k=1}^{N(t)} I\{M(k-1, k) = 0\}$$

$$I\{M(k-1, k] = 0\} \sim \text{Ber}(e^{-\mu})$$

It follows $L(t) | N(t)=n \sim \text{Bin}(n, e^{-\mu})$

$$P(L(t)=m) = \sum_{n=0}^{\infty} P(L(t)=m | N(t)=n) P(N(t)=n)$$

$$= \sum_{n=m}^{\infty} \binom{n}{m} e^{-\lambda m} (1 - e^{-\lambda})^{n-m} \frac{e^{-\lambda t}}{n!} (\lambda + t)^m$$

$$= \frac{e^{-\lambda t - Mm}}{m!} \sum_{n=m}^{\infty} \frac{(\lambda t)^n}{(n-m)!} (1-e^{-\mu})^{n-m} \frac{(\lambda t)^m}{(\lambda + \mu)^m}$$

$$= \frac{e^{-\lambda t} (e^{-\mu t} \lambda t)^m}{m!} \sum_{k=0}^{\infty} \frac{(\lambda t + (1 - e^{-\mu t}))^k}{k!}$$

$$= \overbrace{e^{-\lambda t} e^{\lambda t + (1 - e^{-M})}}^{< k = 0} \left(\frac{e^{-\mu} \lambda t}{M!} \right)^M$$

$$(e^{-\mu t})^m = e^{-\mu m t}$$

Because $\text{Bin}(n, e^{-\mu}) = 0$ for $n < m$

$$= e^{-\mu} e^{\lambda t} \cdot \frac{(\lambda t)^m}{m!}$$

$$= \underline{(e^{-\mu} \lambda t)^m} e^{-\mu} \lambda t$$

Thus $L(t) \sim \text{Poisson}(e^{-\mu} \lambda t)$ \Rightarrow rate of L is $e^{-\mu} \lambda$

- $L(0) = 0$

- Since $L(t) \sim \text{Poisson}(e^{-\mu} \lambda t)$

$$L(t) - L(s) = L(t-s)$$