

Chebyshev Inequality. Let X be a r.v. with variance σ^2 and $E\{X\} = \mu$. Then

$$P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2},$$

Proof. $|X - \mu| \geq k\sigma$ if $(X - \mu)^2 \geq k^2\sigma^2$, so $P(|X - \mu| \geq k\sigma) = P((X - \mu)^2 \geq k^2\sigma^2)$.

Let $u(x) = (x - \mu)^2$ and $C = k^2\sigma^2$, then by Theorem 2.4.6

$$P(u(X) \geq C) \leq \frac{E[u(X)]}{C}$$

or

$$P((X - \mu)^2 \geq k^2\sigma^2) \leq \frac{E\{(X - \mu)^2\}}{k^2\sigma^2} = \frac{1}{k^2}, \text{ since}$$

$$\sigma^2 = \text{var}\{X\} = E\{(X - \mu)^2\}.$$

Therefore

$$P(|X - \mu| \geq k\sigma) = P((X - \mu)^2 \geq k^2\sigma^2) \leq \frac{1}{k^2}.$$

2.23. Let X be a random variable with pmf $p(x) = \frac{x}{8}$ if $x = 1, 2, 5$ and $p(x) = 0$ otherwise

$$\begin{aligned} (a) E[X] &= \sum_x x p(x) = 1 \cdot p(1) + 2 p(2) + 5 p(5) \\ &= \frac{1}{8} + \frac{4}{8} + \frac{25}{8} \\ &= \frac{30}{8} = \frac{15}{4} = 3.75 \end{aligned}$$

$$(b) \text{Var}[X] = E[X^2] - E[X]^2,$$

$$\begin{aligned} E[X^2] &= \sum_x x^2 p(x) = 1 p(1) + 4 p(2) + 25 p(5) \\ &= \frac{1}{8} + \frac{8}{8} + \frac{125}{8} \\ &= \frac{134}{8} = \frac{67}{4} \end{aligned}$$

$$\therefore \text{Var}[X] = \frac{67}{4} - \left(\frac{15}{4}\right)^2 = \frac{268}{16} - \frac{225}{16} = \frac{43}{16} = 2.6875$$

$$(c) E[2X+3] = 2E[X] + 3 = 2 \cdot \frac{15}{4} + 3 = \frac{15}{2} + \frac{6}{2} = \frac{21}{2} = 10.5$$

2.24 Let X be continuous with pdf $f(x) = \begin{cases} 3x^2 & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$

$$(a) E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_0^1 x 3x^2 dx = \frac{3}{4} x^4 \Big|_0^1 = \frac{3}{4}.$$

$$(b) \text{Var}(X) = E[X^2] - E[X]^2, \text{ and}$$

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_0^1 x^2 3x^2 dx = \frac{3}{5} x^5 \Big|_0^1 = \frac{3}{5}, \text{ so}$$

$$\text{Var}(X) = \frac{3}{5} - \left(\frac{3}{4}\right)^2 = \frac{3}{5} - \frac{9}{16} = \frac{3}{80} = 0.0375$$

$$(c) E[X^r] = \int_{-\infty}^{\infty} x^r f(x) dx = 3 \int_0^1 x^r x^2 dx = \frac{3}{3+r} x^{3+r} \Big|_0^1 = \frac{3}{3+r}$$

$$(d) E[3X - 5X^2 + 1] = 3E[X] - 5E[X^2] + 1 \\ = 3 \cdot \frac{3}{4} - 5 \cdot \frac{3}{5} + 1 = \frac{1}{4}$$

2.25. Let X be continuous with pdf $f(x) = \begin{cases} x^{-2} & 1 \leq x < \infty \\ 0 & \text{otherwise} \end{cases}$.

(a) If $E[X]$ exists it would need to be

$$\int_{-\infty}^{\infty} f(x) \cdot x dx = \int_1^{\infty} x \cdot x^{-2} dx = \int_1^{\infty} x^{-1} dx = \lim_{x \rightarrow \infty} \ln x,$$

and the limit does not exist, so $E[X]$ does not exist.

(b) If $E[X^{-1}]$ exists, then it equals

$$\int_{-\infty}^{\infty} f(x) \cdot x^{-1} dx = \int_1^{\infty} x^{-2} \cdot x^{-1} dx = \int_1^{\infty} x^{-3} dx = -\frac{1}{2} x^{-2} \Big|_1^{\infty} = \frac{1}{2}, \text{ so}$$

$E[X^{-1}]$ exists.

(c) If $E[X^K]$ exists, then it must equal

$$\int_{-\infty}^{\infty} x^K f(x) dx = \int_1^{\infty} x^K x^{-2} dx = \int_1^{\infty} x^{K-2} dx, \text{ but in}$$

general $\int_1^{\infty} x^p dx$ converges iff $p < -1$, so $E[X^K]$

exists iff $K-2 < -1$, or iff $K < 1$.

2.34 Let $M_X(t) = \frac{1}{8}e^t + \frac{1}{4}e^{2t} + \frac{5}{8}e^{5t}$ be the MGF of X .

(a) Let Y be a random variable with pmf $f(y) = \begin{cases} \frac{y}{8} & y=1,2,5 \\ 0 & \text{otherwise} \end{cases}$.

Then the MGF of Y $M_Y(t) = E[e^{Yt}]$

$$= \sum_y e^{yt} f(y)$$

$$= \frac{1}{8}e^t + \frac{1}{4}e^{2t} + \frac{5}{8}e^{5t}$$

$$= M_X(t).$$

Therefore, the distribution of X is the same as the distribution of Y because $M_X = M_Y$.

$$(b) P(X=2) = P(Y=2) = f(2) = \frac{1}{4}.$$

2.36 Let X be a continuous random variable with pdf

$$f(x) = \begin{cases} e^{-x-2} & -2 < x \\ 0 & \text{otherwise} \end{cases}$$

(a) If $M_X(t)$ is the MGF of X , then

$$\begin{aligned} M_X(t) &= E[e^{xt}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_{-2}^{\infty} e^{tx} e^{-x-2} dx \\ &= \int_{-2}^{\infty} e^{x(t-1)-2} dx \\ &= \frac{1}{t-1} e^{x(t-1)-2} \Big|_{-2}^{\infty} \\ &= \frac{1}{1-t} e^{-2t} \quad \text{if } t-1 < 0 \\ &\quad \text{or } t < 1 \end{aligned}$$

$$\begin{aligned} (b) \quad E[X] &= M'_X(0) = \frac{-2e^{-2t}(1-t) - e^{-2t}(-1)}{(1-t)^2} \Big|_{t=0} \\ &= -1 \end{aligned}$$

$$\begin{aligned} E[X^2] &= M''_X(0) = \frac{d}{dt} \left[\frac{-2e^{-2t}}{1-t} + \frac{e^{-2t}}{(1-t)^2} \right]_{t=0} \\ &= \frac{d}{dt} \left(e^{-2t} \left((1-t)^{-2} - 2(1-t)^{-3} \right) \right)_{t=0} \\ &= \left[-2e^{-2t} \left((1-t)^{-2} - 2(1-t)^{-3} \right) + e^{-2t} \left(2(1-t)^{-3} - 2(1-t)^{-2} \right) \right]_{t=0} \\ &= -2(-1) = 2 \end{aligned}$$