

5.22 Let X & Y have joint pdf
 $f(x,y) = \begin{cases} e^{-y} & 0 < x < y < \infty \\ 0 & \text{o.w.} \end{cases}$

$$\begin{aligned} M(t,s) &= E[e^{tx+sy}] = \int e^{tx+sy} f(x,y) dx dy \\ &= \int_0^{\infty} \int_0^y e^{tx+sy} e^{-y} dx dy = \int_0^{\infty} e^{(s-1)y} \cdot \frac{1}{t} [e^{tx}]_0^y dy \\ &= \int_0^{\infty} e^{(s-1)y} \cdot \frac{1}{t} (e^{ty} - 1) dy = \frac{1}{t} \int_0^{\infty} (e^{(st+1)y} - e^{(s-1)y}) dy \end{aligned}$$

If $st+1 < 0$ and $s-1 < 0$, then $M(t,s)$ exists, and

$$\begin{aligned} M(t,s) &= \frac{1}{t} \cdot \left[\frac{1}{st+1} e^{(st+1)y} - \frac{1}{s-1} e^{(s-1)y} \right]_0^{\infty} \\ &= \frac{1}{t} \left(\frac{1}{s-1} - \frac{1}{st+1} \right) \quad s < 1, t < 1-s \\ &= 1 / ((s-1) * (s+t-1)) \end{aligned}$$

5.23 Let $(X, Y) \sim \text{Bivariate Normal}(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$. Then

$$X \sim N(\mu_1, \sigma_1^2), Y \sim N(\mu_2, \sigma_2^2), \text{ and } \rho = \text{corr}(X, Y).$$

Proof: By Example 5.2, the joint MGF of (X, Y) is

$$M(t, s) = E[e^{tX + sY}] = e^{\mu_1 t + \mu_2 s + \frac{1}{2}(\sigma_1^2 t^2 + \sigma_2^2 s^2 + 2\rho\sigma_1\sigma_2 ts)}$$

Then the MGF of X is $M(t, 0) = e^{\mu_1 t + \frac{1}{2}\sigma_1^2 t^2}$, which is the MGF of a Normal (μ_1, σ_1^2) R.V. by Theorem 3.3.5.

The MGF of Y is $M(0, s) = e^{\mu_2 s + \frac{1}{2}\sigma_2^2 s^2}$, which is the MGF of a Normal (μ_2, σ_2^2) R.V. by Theorem 3.3.5.

Therefore, $X \sim N(\mu_1, \sigma_1^2)$, and $Y \sim N(\mu_2, \sigma_2^2)$.

By the discussion at the start of 5.5,

$$E[XY] = \frac{\partial^2 M}{\partial t \partial s}(0, 0) = (\mu_1 + \sigma_1^2 t + \rho\sigma_1\sigma_2 s)M(t, s)|_{(t, s) = (0, 0)}$$

$$= \frac{\partial}{\partial s}(\mu_1 + \sigma_1^2 t + \rho\sigma_1\sigma_2 s)M(t, s)|_{(t, s) = (0, 0)}$$

$$= \rho\sigma_1\sigma_2 M(0, 0) + (\mu_2 + \sigma_2^2 s + \rho\sigma_1\sigma_2 t)(\mu_1 + \sigma_1^2 t + \rho\sigma_1\sigma_2 s)M(t, s)|_{(t, s) = (0, 0)}$$

$$= \rho\sigma_1\sigma_2 + \mu_2\mu_1$$

$$\text{so } \text{cov}(X, Y) = E[XY] - E[X]E[Y] = \rho\sigma_1\sigma_2 + \mu_1\mu_2 - \mu_1\mu_2 = \rho\sigma_1\sigma_2, \text{ and}$$

$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{Var}(X)}\sqrt{\text{Var}(Y)}} = \frac{\rho\sigma_1\sigma_2}{\sigma_1\sigma_2} = \rho.$$

5.25 Let X_1 and X_2 be independent, $X_i \sim N(\mu_i, \sigma_i^2)$,
 $Y_1 = X_1$, $Y_2 = X_1 + X_2$.

(a) The MGF of X_i is $M_i(t_i) = e^{\mu_i t_i + \frac{1}{2} \sigma_i^2 t_i^2}$, and
 because X_1 and X_2 are independent their joint MGF
 factors: $M(t_1, t_2) = M_1(t_1) M_2(t_2)$
 $= e^{\mu_1 t_1 + \mu_2 t_2 + \frac{1}{2} (\sigma_1^2 t_1^2 + \sigma_2^2 t_2^2)}$

The joint MGF of Y_1, Y_2 then, is

$$\begin{aligned} N(t_1, t_2) &= E[e^{t_1 Y_1 + t_2 Y_2}] = E[e^{t_1 X_1 + t_2 X_1 + t_2 X_2}] \\ &= E[e^{(t_1 + t_2) X_1 + t_2 X_2}] = M(t_1 + t_2, t_2) \\ &= e^{\mu_1(t_1 + t_2) + \mu_2 t_2 + \frac{1}{2} (\sigma_1^2 (t_1 + t_2)^2 + \sigma_2^2 t_2^2)} \\ &= e^{\mu_1 t_1 + (\mu_1 + \mu_2) t_2 + \frac{1}{2} (\sigma_1^2 t_1^2 + (\sigma_1^2 + \sigma_2^2) t_2^2 + 2\sigma_1^2 t_1 t_2)} \end{aligned}$$

Which is the MGF of a Bivariate Normal $(\mu'_1, \mu'_2, \sigma_1'^2, \sigma_2'^2, \rho)$

(b) R.V. with $\mu'_1 = \mu_1$, $\mu'_2 = \mu_1 + \mu_2$, $\sigma_1'^2 = \sigma_1^2$, $\sigma_2'^2 = \sigma_1^2 + \sigma_2^2$,

$$\rho = \frac{\sigma_1^2}{\sigma_1' \sigma_2'} = \frac{\sigma_1'^2}{\sigma_1' \sigma_2'} = \frac{\sigma_1'}{\sigma_2'} = \frac{\sigma_1}{\sqrt{\sigma_1^2 + \sigma_2^2}} = \text{corr}(Y_1, Y_2).$$

$$E[Y_1] = \mu_1 \text{ and } E[Y_2] = \mu_1 + \mu_2$$

(c) By Theorem 5.4.8,

$$\begin{aligned} Y_2 | Y_1 = y_1 &\sim N\left(\mu_1 + \mu_2 + \frac{\sigma_1'}{\sigma_2'} \frac{\sigma_2'}{\sigma_1'} (y_1 - \mu_1), \sigma_2'^2 (1 - \rho^2)\right) \\ &\sim N(\mu_2 + y_1, \sigma_2'^2 (1 - \frac{\sigma_1'^2}{\sigma_2'^2})) \sim N(\mu_2 + y_1, \sigma_2'^2 - \sigma_1'^2) \\ &\sim N(y_1 + \mu_2, \sigma_2^2) \end{aligned}$$