Multi-level multi-state modelling applied to hospital admission in mexican patients with COVID-19

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Abstract

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Author summary

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Introduction

The SARS-CoV-2 pandemic was declared a Public Health Emergency of International Concern on January 30, 2020 by the World Health Organization. The Mexican Health Authorithies declared the first lockdown on March 26 with 585 cases and 8 deaths reported for COVID-19 [1]; at the end of the lockdown (june 5th 2020) the total cases were 110,026 and 13 170 deaths. Until November 1, Mexico is the fourth country in death rates of SARS-CoV-19 (106,765 deaths), with 1,122,362 incident cases [2].

Over time it has become clear that the presence of comorbidities such as hypertension, diabetes, obesity and smoking are factors that increase the serious illness that leads to hospitalization and in 25% of the cases they required admission and intubation to the intensive care unit [3]. Mexico ranks second in obesity among OECD countries, with almost 72.5% obesity among the adult population, which is associated with the high prevalence of type 2 diabetes, estimated at 13% of the adult population in 2017, which is the highest rate among OECD countries; hypertension is also one of the higher chronic diseases among adult population 30% [4]. The high prevalence of this

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comorbidities besides the lack of a functional health care system is believed to be the main reason why the severe cases and deaths rates in the country are so high.

There have been different efforts to understand how patients with comorbilities may face the COVID-19 disease; the work by [5] aims to identify the risk factors in order to propose a clinical score to predict COVID-19 lethality, including the different factors like diabetes and obesity among mexican population. This lead to believe that obesity mediates 49.5% of the effect of diabetes on COVID-19 lethality. Early-onset diabetes conferred an increased risk of hospitalization and obesity conferred an increased risk for intensive care unit admission and intubation.

After onset of infection there is a period of time between symptom detection and hospitalization. The time elapsed before patients approach hospitals could be excessively long. Once patients are admitted to hospital, there is also a period of time between the admission and death. Estimation of lengths of these times through a multilevel model could enable a better information system to estimate incidence and transmission rates, particularly at regional level since differences have already been established.

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This work considers a multi-state model under a Bayesian framework to estimate times between symptom detection and hospitalization and between hospitalization and death. Data used in the modeling comes from the official database by the Mexican Ministry of Health; the analysis provides of general overview of hospitalizations in each state of the country and the different health institutions within. Variables affecting the patient's final outcome such as the aforementioned comorbidites are included in the model. Additionally regional heterogeneity is accounted trough nested models that consider the regional contribution and also the health service provider. Other efforts in recent literature [6] have considered more states (hospitalization-ICU, ICU-death, ICU-discharged). They were able to asses whether improvements in patient outcomes have been sustained, finding evidence that median hospital stays have lengthened. Unfortunately the data available for Mexico does not have the required information to include such states. Nevertheless, we believe this model could provide a better information system to estimate incidence and transmission rates, particularly important while new variants and increased transmission rates are present.

The fact that different final outcomes could be related to patient's late hospitalization, hence suggesting that the average patient waits until the symptoms are severe to seek professional healthcare, needs to be further investigated.

Methods and materials

Data Source and Study Population

We conducted a prevalence study the official database from the Mexican Ministry of Health, this data provides a overview of hospital admissions, deaths and the period of time between hospitalizations and first symptoms between March and December 2020. The data analyzed included mexican adult population diagnosed with COVID-19 in the whole country; the exclusion criteria were the observations with incomplete data about hospital admission, symptoms or comorbidities. Additionally patients whose time of initial symptoms was captured as the day the were admitted to hospital were removed, since this time was likely to be unknown. After applying exclusion criteria a total sample of 1200 registers of adult patients belonging to any healthcare institution, either private or public in the 32 states of Mexico was selected, preserving the population characteristics. JUAN PABLO ESTE NUMERO DE MUESTRA ES CORRECTO?

Comorbidities that could worsen the patient outcome, affecting the times we aim to model; such as diabetes, hypertension, obesity, chronic obstructive pulmonary disease (COPD), asthma, immunosuppression and chronic kidney disease were included as linear

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predictors for each state in the model and so two different relevant groups were statistically significant:

 $x_{death} = (\sim Diabetes + COPD + Obesity + Hypertension + Diabetes * Obesity * Hypertension + Kanada + Value + Value$

About 87% of the population in Mexico belongs to some healthcare institution, but during this pandemic mexican government has established a list of hospitals designated to treat COVID-19 patients without any affiliation distinction. In this study we identified 6 different healthcare providers which were classified as sectors IMSS, ISSSTE, SEDENA/SEMAR/PEMEX, SSA, ESTATALES (healthcare provider within each state) this 5 are the public care provider and the sixth sector is private hospitals. It is worth mentioning that following the national hospital transformation plan [7], IMSS alone has transformed about 260 medical units to treat Covid-19 patients [8].

Modelling

We developed four different Bayesian models for trajectories of interest namely, Symptoms-Hospitalization and Hospitalization-Death, in which non-informative initial distributions were used, located near 0, to improve convergence. Additionally a QR reparameterization for the covariable matrix was used, that is, if X is an $n \times m$ covariable matrix, corresponding to the aforementioned comorbidities, X = QR, where Q is an ortogonal matrix and R is an upper triangular matrix. In practice, considering X = Q'R' where $Q' = Q\sqrt{n-1}$ and $R' = \frac{1}{\sqrt{n-1}}R$ is convinient. Hence if ζ is the N linear predictor vector such that $\zeta = X\beta$, with β a K coefficient vector, then $\zeta = X\beta = QR\beta = Q'R'\beta$. We used $\zeta = Q'R'\beta$ for numerical stability.

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Each model has different levels on it, as more levels were included it was possible to see how different the results were according to the new variable added. The four levels of information were:

- 1. Data of death and hospitalizations which are assumed independent.
- 2. This model is based in the first model, but here the information is slipt up by the different states of Mexico.
- 3. Third model takes the information of the second one and adds the type/institution in which the patients were treated.
- 4. The last models adds the information to the third one by each state of the Mexican republic and so all the variables are accounted in the last one

To choose the model that we will describe in this analysis we considered the leave-one-out cross-validation (LOO) proposed by [6], which estimates pointwise out-of-sample prediction accuracy, using the log-likelihood evaluated at the posterior simulations of the parameter values, Table 1.

Model	elpd leave one out	p leave one out
Model 1	-75473.3 (187.8)	26.9(2.2)
Model II	-75352.2 (186.3)	84.1 (4.7)
Model III	-75284.9 (186.6)	92.5(4.9)

Table 1. Expected log pointwise predictive density for a new data set and effective number of parameters (standard deviation).

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Patient i corresponds to the i-th row of vectors M and H for deaths and hospitalizations, respectively. We considered a different set of covariables in each case and we assumed deaths and hospitalizations are independent and the model is given as

$$M \sim Weibull(\alpha, \eta)$$

$$H \sim Weibull(\alpha, \upsilon)$$

$$\eta = \exp\left(-\frac{\mu_m + \mathbf{Q}^* \vartheta}{\alpha}\right)$$

$$\upsilon = \exp\left(-\frac{\mu_h + \mathbf{Q}^{**} \theta}{\alpha}\right)$$

$$\alpha = \exp(\alpha_r * \tau_\alpha)$$

$$\alpha_r \sim N(0, 1)$$

$$\mu_m, \mu_h \sim N(0, \tau_\mu)$$

$$\vartheta, \theta \sim U(-\infty, \infty)$$

where Q^* and Q^{**} are matrices of standarized covariables for deaths and hospitalizations respectively and τ_{α} and τ_{μ} are given positive values. This model is described in grey in Figure REF.

Model II: two levels

The second model is based on the first one, an additional level is added to account for each state of Mexico to model deaths. The hospitalization H remains unchanged and for each state $l=1,\ldots,32$ and patient i, deaths are summarized in matrix M_l the model is defined as

$$\begin{split} M_l &\sim Weibull(\alpha, \eta) \\ H &\sim Weibull(\alpha, \upsilon) \\ \eta &= \exp\left(-\frac{\mu_m + \mu_l^r + \mathbf{Q}^* \vartheta}{\alpha}\right), l = 1, ..., 32 \\ \upsilon &= \exp(-\frac{\mu_h + \mathbf{Q}^{**} \theta}{\alpha}) \\ \mu_l &= \sigma * \mu_l^r \\ \alpha &= \exp(\alpha^r * \tau_\alpha) \\ \alpha^r &\sim N(0, 1) \\ \mu_l^r &\sim N(0, 1) \\ \sigma &\sim t_3^+(0, 1) \\ \mu_m, \mu_h &\sim N(0, \tau_\mu) \\ \vartheta, \theta &\sim U(-\infty, \infty) \end{split}$$

where Q^* and Q^{**} are matrices of standarized covariables for deaths and hospitalizations respectively and τ_{α} and τ_{μ} are given positive values. This model is described in green in Figure REF.

Model III: Three levels Based on Model II, we consider a third level to include the type of health service where patients are hospitalized, k, for patient i in state l we have $M_{l,k}$ for the corresponding death matrix and the model is given as

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$$\begin{split} M_{l,k} &\sim Weibull(\alpha, \eta) \\ H &\sim Weibull(\alpha, \upsilon) \\ \eta &= \exp\left(-\frac{\mu_m + \mu_l^r + \mu_k^r + \mathbf{Q}^* \vartheta}{\alpha}\right), l = 1, ..., 32, k = 1, ..., 5 \\ \upsilon &= \exp\left(-\frac{\mu_h + \mathbf{Q}^{**} \theta}{\alpha}\right) \\ \mu_l &= \sigma_l * \mu_l^r, l = 1, 2 \\ \mu_k &= \sigma_l * \mu_k^r \\ \alpha &= \exp(\alpha^r * \tau_\alpha) \\ \alpha^r &\sim N(0, 1) \\ \mu_l^r, \mu_k^r &\sim N(0, 1) \\ \sigma_l &\sim t_3^+(0, 1) \\ \mu_m, \mu_h &\sim N(0, \tau_\mu) \\ \vartheta, \theta &\sim U(-\infty, \infty) \end{split}$$

where Q^* and Q^{**} are matrices of standarized covariables for deaths and hospitalizations respectively and τ_{α} and τ_{μ} are given positive values. This model is described in yellow in Figure REF.

Model IV:

Based as well on Modelo II, however, we consider index j = (l, k) where l accounts for the l-th state and k for the type of health service, $j \in \{1, ..., 153\}$ where the distribution for deaths is given by

$$M_j \sim Weibull\left(\alpha, exp\left(-\frac{Q^{**}\theta^m + \mu_k + \mu_l}{\alpha}\right)\right)$$

Results

The parameters were estimated using Stan DAR QUIZA MÁS DETALLES. We show results for model threes which performed better in terms of the likelihood and showed good convergence of all parameters. The posterior 0.95 credibility intervals for parameters of interest at different levels of the model are shown in Figures xx and xx. It is worth pointing out that we are displaying the log hazard ratio, hence positive values for parameters will point to increasing risks for the corresponding transition and level.

Increased risk for hospitalization was observed at the global population level for chronic renal disease, whereas for death such was the case for COPD and the interaction of diabetes:hypertension:obesity.

This study has shown that there are differences in mortality between the states without accounting for institution, if $\beta>0$ the risk increases whereas $\beta<0$ decreases the risk of the event, and it is related to the prompt time of death or viceversa. (Image Mu_1Jer1) shows the states in which the overall rate of mortality is higher such as Campeche, Colima, Guanajuato, Hidalgo, Jalisco, Morelos, Nayarit, Oaxaca, Puebla, Tabasco and Veracruz. The difference might be linked to the late hospitalization of patients

fig(JER1_mort) displays evidence that 5 days after hospitalization there is a peak on mortality rate, which could be related due the late hospitalization of patients with mild symptoms who developed "happy hypoxemia", that is extremely low blood oxygenation,

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but without sensation of dyspnea [9]; in Wuhan a cohort of patients infected with severe acute respiratory syndrome coronavirus 2 (SARS-COV-2) among of 62% with severe disease and 46% who ended up intubated, ventilated or dead did not present dyspnea [10].

Regarding the 6 helth provider sectors included in the analysis differences were also found. While State managed hospitals and private sector showed lower risks, in contrast the IMSS seems to be the one with the highest risk (fig Mu_12Jer2). Although it is worth mentioning that following the national hospital transformation plan [7], IMSS alone has transformed about 260 medical units to treat Covid-19 patients [8].

Disscusion

Through time it has been proved that the presence of comorbidities such as diabetes, hypertension, obesity, chronic obstructive pulmonary disease (COPD), asthma, immuno-suppression and chronic kidney disease are associated to a worse outcome for the patients diagnosed with COVID-19. Particularly to those who are hospitalized in the ICU due intubation. To our knowledge this is the first study in Mexico which analyzed the time elapsed between the patient 's first symptoms, hospitalization and death; all this analysis can break down to the different states of the republic and the healthcare institutions within them. Because of this it is possible to distinguish those for which the risk of hospitalization and death increases. Mexican population is the third place in obesity among the OCDE countries such thing as this one is one of the main reason why the severe cases due COVID-19 are so high.

One of the main problems is the precarious situation of the public healthcare system which universal coverage is estimated about 87% of the mexican population. It's clear that mexican healthcare has overrun during this pandemic and has appealed to private health providers to cope with the treatment of COVID-19 patients. Among all health care providers (sectors) the IMSS is the one with the highest risk, however it is the largest healthcare provider across Mexico with hospitals from level 2-level 4 of which "Siglo XXI" is a country-leader in investigation and innovative treatments and procedures. In March, 2020 a list of hospitals that were "converted" [7] to the treatment of COVID-19 were listed 38 hospital among which 18 were IMSS hospitals only; after one year 960 hospitals across the country were converted to treat patients with COVID-19 of these 289 (30.10%) belongs to IMSS [8].

This analysis can be helpful to a regional level to improve healthcare assistance, additionally it could be useful to inform statistical estimation of parameters for an epidemiological model.

This study has shown that there are differences in mortality between the different states of the republic; there are states in which the overall rate of mortality is higher due the late hospitalization of patients such as Veracruz, Nuevo Leon, San Luis Potosi, Guanajuato, Chiapas and Mexico City. Breaking down this analysis to state level we found a higher risk of hospitalization, specifically in Veracruz HR which historically has been unsteady regarding the public healthcare system in sectors like IMSS, ISSSTE, SEDENA/SEMAR/PEMEX.

One of the main limitations of the study is the reduced number of states we were able to include in the models due to the lack of information regarding dates of dismissal of recovered patient's from hospitalization.

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