# Formulae for Bayesian A/B Testing: Walkthrough

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June 12, 2020

Original material

#### **Table of Contents**

- 1. A/B testing: binary outcomes
  - 1. Derivation
  - 2. Equivalent Formulas
- 2. A/B/C testing: binary outcomes
  - 1. Derivation

# A/B testing: binary outcomes

For a binary-outcome test (e.g. a test of conversion rates), the probability that B will beat A in the long run is given by:

$$\Pr(p_B>p_A) = \sum_{i=0}^{lpha_B-1} rac{B(lpha_A+i,eta_B+eta_A)}{(eta_B+i)B(1+i,eta_B)B(lpha_A,eta_A)}$$

Where:

- $\alpha_A = 1 + S_A$   $\beta_A = 1 + F_A$   $\alpha_B = 1 + S_B$
- $\beta_B$  = 1 +  $F_B$
- B is the beta function. For arbitrary  $\alpha > 0$  and  $\beta > 0$ , beta function is defines as:

$$B(lpha,eta)=rac{\Gamma(lpha)\Gamma(eta)}{\Gamma(lpha+eta)}.$$

### **Derivation**

Suppose we have two independent experimental branches (A and B) and have a Bayesian belief for each one:

$$p_A \sim ext{Beta}(lpha_A, eta_A) \ p_B \sim ext{Beta}(lpha_B, eta_B)$$

Using the pdf of the beta distribution, we can get the total probability that  $p_B$  is greater than  $p_A$  by integrating the joint distribution over all values for which  $p_B > p_A$ :

$$\Pr(p_{B} > p_{A}) = \int_{0}^{1} \Pr(p_{B} > p_{A}|p_{A}) \Pr(p_{A}) dp_{A} = \int_{0}^{1} \int_{p_{A}}^{1} \Pr(p_{B}|p_{A}) dp_{B} \Pr(p_{A}) dp_{A}$$

$$= \int_{0}^{1} \int_{p_{A}}^{1} \frac{p_{A}^{\alpha_{A}-1} (1 - p_{A})^{\beta_{A}-1}}{B(\alpha_{A}, \beta_{A})} \frac{p_{B}^{\alpha_{B}-1} (1 - p_{B})^{\beta_{B}-1}}{B(\alpha_{B}, \beta_{B})} dp_{B} dp_{A}$$
(1)

Define 
$$\mathrm{B}(x;\,a,b)=\int_0^x t^{a-1}\,(1-t)^{b-1}\,dt$$
 and  $I_x(a,b)=\frac{\mathrm{B}(x;\,a,b)}{\mathrm{B}(a,b)}.$   $I_x(a,b)$  is the regularized incomplete beta function and note that  $I_x(1,b)=\frac{\mathrm{B}(x;\,1,b)}{\mathrm{B}(1,b)}=b\cdot\int_0^x (1-t)^{b-1}dt=b\left\{\frac{1}{b}-\frac{(1-x)^b}{b}\right\}=1-(1-x)^b$ .

$$\int_{p_A}^1 \frac{p_B^{\alpha_B-1}(1-p_B)^{\beta_B-1}}{B(\alpha_B,\beta_B)} dp_B = 1 - \int_0^{p_A} \frac{p_B^{\alpha_B-1}(1-p_B)^{\beta_B-1}}{B(\alpha_B,\beta_B)} dp_B = 1 - \frac{B(p_A;\alpha_B,\beta_B)}{B(\alpha_B,\beta_B)} = 1 - I_{p_A}(\alpha_B,\beta_B). \tag{2}$$

So the equation becomes:

$$\Pr(p_B > p_A) = 1 - \int_0^1 rac{p_A^{lpha_A - 1}(1 - p_A)^{eta_A - 1}}{B(lpha_A, eta_A)} I_{p_A}(lpha_B, eta_B) dp_A.$$
 (3)

Now, there is a recursive relationship

$$I_x(a,b) = I_x(a-1,b) - \frac{x^{a-1}(1-x)^b}{(a-1)B(a-1,b)}$$
 (4)

because if we denote  $f(a,b,x)=I_x(a,b)-I_x(a-1,b)+rac{x^{a-1}(1-x)^b}{(a-1)\,B\,(a-1,b)}$  , f(a,b,0)=0 and

$$f'(a,b,x) = I'_{x}(a,b) - I'_{x}(a-1,b) + \frac{x^{a-2}(1-x)^{b-1}\{(a-1)(1-x) - bx\}}{(a-1)B(a-1,b)}$$

$$= \frac{x^{a-1}(1-x)^{b-1}}{B(a,b)} + \frac{x^{a-2}(1-x)^{b-1}}{B(a-1,b)} + \frac{x^{a-2}(1-x)^{b-1}\{(a-1)(1-x) - bx\}}{(a-1)B(a-1,b)}$$
(5)

so f'(a, b, 0) = 0 and f(a, b, x) = 0.

Using this relationship and the fact that lpha and eta are integers, we can express  $I_x$  as:

$$I_{x}(a,b) = I_{x}(a-1,b) - \frac{x^{a-1}(1-x)^{b}}{(a-1)B(a-1,b)}$$

$$= I_{x}(a-2,b) - \frac{x^{a-2}(1-x)^{b}}{(a-2)B(a-2,b)} - \frac{x^{a-1}(1-x)^{b}}{(a-1)B(a-1,b)} = \dots = 1 - (1-x)^{b} - \sum_{i=1}^{a-1} \frac{x^{a-j}(1-x)^{b}}{(a-j)B(a-j,b)}$$
(6)

Or equivalently:

$$I_x(a,b) = 1 - \sum_{i=0}^{a-1} \frac{x^i (1-x)^b}{(b+i)B(1+i,b)} \tag{7}$$

The probability integral (3) can therefore be written:

$$\begin{array}{lll} \Pr(p_B>p_A) & = & 1-\int_0^1 \frac{p_A^{\alpha_A^{-1}}(1-p_A)^{\beta_A^{-1}}}{B(\alpha_A,\beta_A)} \left(1-\sum_{i=0}^{\alpha_B-1} \frac{p_A^i(1-p_A)^{\beta_B}}{(\beta_B+i)B(1+i,\beta_B)}\right) dp_A \\ & = & 1-1+\int_0^1 \frac{p_A^{\alpha_A^{-1}}(1-p_A)^{\beta_A^{-1}}}{B(\alpha_A,\beta_A)} \sum_{i=0}^{\alpha_B-1} \frac{p_A^i(1-p_A)^{\beta_B}}{(\beta_B+i)B(1+i,\beta_B)} dp_A \\ & = & \int_0^1 \sum_{i=0}^{\alpha_B-1} \frac{p_A^{\alpha_A^{-1+i}}(1-p_A)^{\beta_A+\beta_B^{-1}}}{(\beta_B+i)B(\alpha_A,\beta_A)B(1+i,\beta_B)} dp_A \\ & = & \sum_{i=0}^{\alpha_B-1} \int_0^1 \frac{p_A^{\alpha_A^{-1+i}}(1-p_A)^{\beta_A+\beta_B^{-1}}}{(\beta_B+i)B(\alpha_A,\beta_A)B(1+i,\beta_B)} dp_A \\ & = & \sum_{i=0}^{\alpha_B-1} \frac{B(\alpha_A+i,\beta_A+\beta_B)}{(\beta_B+i)B(\alpha_A,\beta_A)B(1+i,\beta_B)} \int_0^1 \frac{p_A^{\alpha_A^{-1+i}}(1-p_A)^{\beta_A+\beta_B^{-1}}}{B(\alpha_A+i,\beta_A+\beta_B)} dp_A \end{array}$$

Finally:

$$\Pr(p_B > p_A) = \sum_{i=0}^{\alpha_B - 1} \frac{B(\alpha_A + i, \beta_A + \beta_B)}{(\beta_B + i)B(1 + i, \beta_B)B(\alpha_A, \beta_A)} \tag{8}$$

### **Equivalent Formulas**

It's possible to derive similar formulas that sum over the other three parameters:

$$\Pr(p_B > p_A) = 1 - \sum_{i=0}^{\alpha_A - 1} \frac{B(\alpha_B + i, \beta_B + \beta_A)}{(\beta_A + i)B(1 + i, \beta_A)B(\alpha_B, \beta_B)}$$
(9)

$$\Pr(p_B > p_A) = \sum_{i=0}^{\beta_A - 1} \frac{B(\beta_B + i, \alpha_A + \alpha_B)}{(\alpha_A + i)B(1 + i, \alpha_A)B(\alpha_B, \beta_B)}$$
(10)

$$\Pr(p_B>p_A)=1-\sum_{i=0}^{eta_B-1}rac{B(eta_A+i,lpha_A+lpha_B)}{(lpha_B+i)B(1+i,lpha_B)B(lpha_A,eta_A)}$$

The above formulas can be found with symmetry arguments.

## A/B/C testing: binary outcomes

It is possible to extend the binary-outcome formula to three test groups, call them A, B, and C. The probability that C will beat both A and B in the long run is:

$$\Pr(p_C > \max\{p_A, p_B\}) = \ 1 - \Pr(p_A > p_C) - \Pr(p_B > p_C) + \sum_{i=0}^{lpha_A - 1} \sum_{i=0}^{lpha_B - 1} rac{B(i+j+lpha_C, eta_A + eta_B + eta_C)}{(eta_A + i)B(1+i, eta_A)(eta_B + j)B(1+j, eta_B)B(lpha_C, eta_C)}$$

Where:

- $\alpha_X$  is one plus the number of successes for  $X \in \{A,B,C\}$
- $eta_X$  is one plus the number of failures for  $X \in \{A,B,C\}$
- $\Pr(p_X > p_C)$  is the formula for the two-group case, given by (8)

Note that this formula can be computed in  $O(\alpha_A \alpha_B)$  time (see the <u>implementation</u> section below).

#### **Derivation**

Start with a Bayesian belief for each of three experimental branches (A, B, and C):

$$egin{aligned} p_A &\sim \mathrm{Beta}(lpha_A, eta_A) \ p_B &\sim \mathrm{Beta}(lpha_B, eta_B) \ p_C &\sim \mathrm{Beta}(lpha_C, eta_C) \end{aligned}$$

Calling the pdf of the beta distribution  $f(p|\alpha,\beta)=f(p)$ , we can get the total probability that  $p_C$  is greater than both  $p_A$  and  $p_B$  by integrating the joint distribution over all values for which  $p_C>p_A$  and  $p_C>p_B$ :

$$egin{aligned} \Pr(\max\left\{p_A,p_B
ight\} < p_C) &= \int_0^1 \Pr(p_C) \Pr(p_C|\max\{p_A,p_B\} < p_C) dp_C = \int_0^1 \Pr(p_C) \Pr(p_A < p_C) \Pr(p_B < p_C) dp_C \ &= \int_0^1 \int_0^{p_C} \int_0^{p_C} f(p_A) f(p_B) f(p_C) dp_A dp_B dp_C \end{aligned}$$

Evaluating the inner two integrals, the equation becomes:

$$\Pr(p_C > \max\{p_A, p_B\}) = \int_0^1 I_{p_C}(lpha_A, eta_A) I_{p_C}(lpha_B, eta_B) f(p_C) dp_C$$
 (12)

Using the identity for  $I_X$  (7), we have:

$$\Pr(p_C > \max\{p_A, p_B\}) = \int_0^1 \left(1 - \sum_{i=0}^{lpha_A - 1} rac{p_C^i (1 - p_C)^{eta_A}}{(eta_A + i)B(1 + i, eta_A)}
ight) \left(1 - \sum_{i=0}^{lpha_B - 1} rac{p_C^i (1 - p_C)^{eta_B}}{(eta_B + i)B(1 + i, eta_B)}
ight) f(p_C) dp_C$$

Multiplying out the parenthetical terms and integrating them separately:

$$egin{aligned} \Pr(p_C > \max{\{p_A, p_B\}}) &= 1 - \int_0^1 \sum_{i=0}^{lpha_A - 1} rac{p_C^i (1 - p_C)^{eta_A}}{(eta_A + i)B(1 + i, eta_A)} f(p_C) dp_C - \int_0^1 \sum_{i=0}^{lpha_B - 1} rac{p_C^i (1 - p_C)^{eta_B}}{(eta_B + i)B(1 + i, eta_B)} f(p_C) dp_C + \int_0^1 \sum_{i=0}^{lpha_A - 1} rac{p_C^i (1 - p_C)^{eta_A}}{(eta_A + i)B(1 + i, eta_A)} \sum_{i=0}^{lpha_B - 1} rac{p_C^i (1 - p_C)^{eta_B}}{(eta_B + i)B(1 + i, eta_B)} f(p_C) dp_C \end{aligned}$$

From the previous derivation, we can rewrite the first two integrals as  $\Pr(p_A > p_C)$  and  $\Pr(p_B > p_C)$ , and consolidate the terms inside the third integral:

$$egin{array}{lll} \Pr(p_C > \max \left\{ p_A, p_B 
ight\}) &=& 1 - \Pr(p_A > p_C) - \Pr(p_B > p_C) \ &+ \int_0^1 \sum_{i=0}^{lpha_A - 1} \sum_{j=0}^{lpha_B - 1} rac{p_C^{i+j}(1 - p_C)^{eta_A + eta_B}}{(eta_A + i)(eta_B + j)B(1 + i, eta_A)B(1 + j, eta_B)} rac{p_C^{lpha_C - 1}(1 - p_C)^{eta_C - 1}}{B(lpha_C, eta_C)} dp_C \ &=& 1 - \Pr(p_A > p_C) - \Pr(p_B > p_C) \ &+ \int_0^1 \sum_{i=0}^{lpha_A - 1} \sum_{j=0}^{lpha_B - 1} rac{p_C^{i+j+lpha_C - 1}(1 - p_C)^{eta_A + eta_B + eta_C - 1}}{(eta_A + i)(eta_A + j)B(1 + i, eta_A)B(1 + j, eta_B)B(lpha_C, eta_C)} dp_C \end{array}$$

Finally, evaluating the integral we have:

$$\begin{array}{lcl} \Pr(p_C > \max\{p_A, p_B\}) & = & 1 - \Pr(p_A > p_C) - \Pr(p_B > p_C) \\ & + \sum_{i=0}^{\alpha_A-1} \sum_{j=0}^{\alpha_B-1} \frac{B(i+j+\alpha_C, \beta_A + \beta_B + \beta_C)}{(\beta_A + i)(\beta_B + j)B(1+i, \beta_A)B(1+j, \beta_B)B(\alpha_C, \beta_C)} \end{array}$$