# GARCH parameters and quantiles estimation

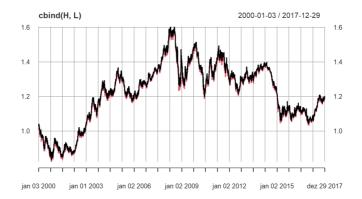
Jose Augusto Fiorucci 20/11/2020

### Input

```
x<- read.csv("~/R/EURUSD/EURUSD.csv", header=F)
colnames(x)=c("Date", "Open","High","Low", "Close","Volume","Adjusted")
x$Date=as.Date(gsub("[.]","-",x$Date))
row.names(x)=x$Date
x=xts(x[,-1],order.by = x[,1])
C_Trend = 0.95
C_Reaction = 0.50</pre>
```

#### High and Low

```
H <- Hi(x)
L <- Lo(x)
plot(cbind(H,L))</pre>
```

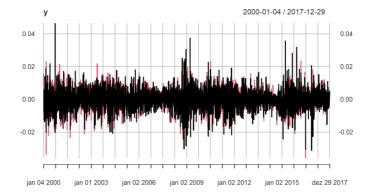


#### Returns

```
y <- cbind( diff(log(H)),  diff(log(L)) )
y <- na.omit(y)
y %>% cor() # Returns correlation
```

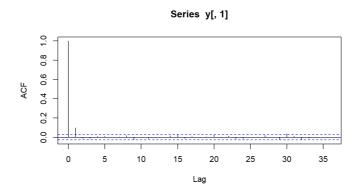
```
## High Low
## High 1.0000000 0.5579326
## Low 0.5579326 1.0000000
```

```
plot(y)
```

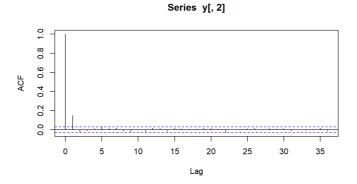


#### Autocorrelation

acf(y[,1])

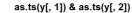


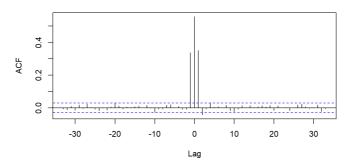
acf(y[,2])



#### **Cross correlation**

ccf(as.ts(y[,1]),as.ts(y[,2]))

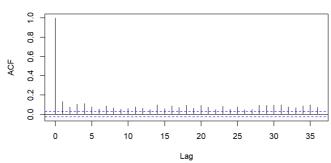




### Volatility verification

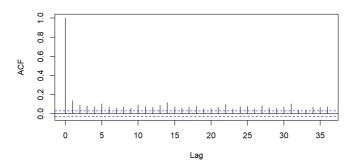
```
acf(y[,1]^2)
```





acf(y[,2]^2)

#### Series y[, 2]^2



## **Bivariate DCC-GARCH**

We will consider the DCC-GARCH to model the volatility of  $y=(r_H,r_L)'$ , where  $r_H$  and  $r_L$  denote the  $100\times$ log-returns from hight's and low's observations.

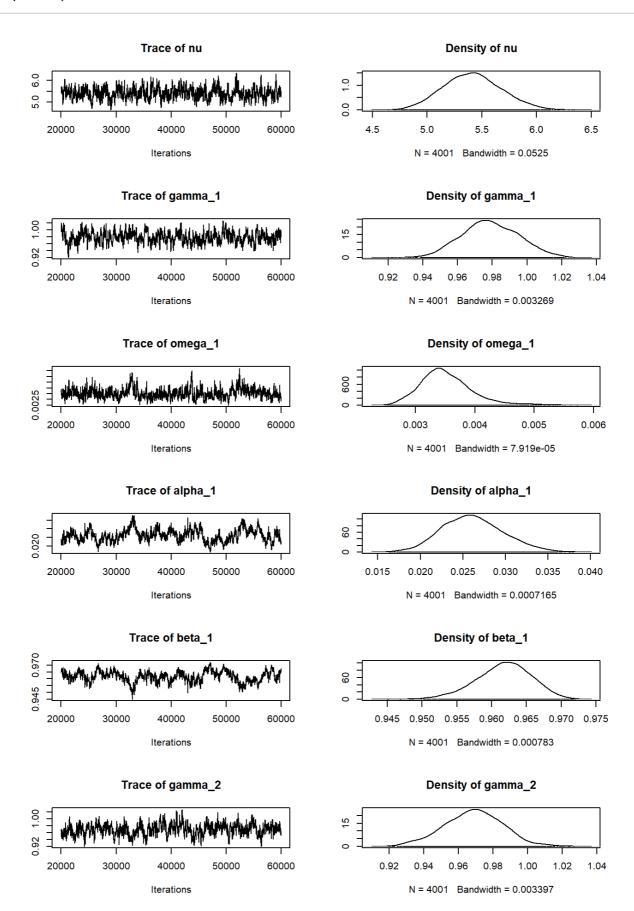
```
# returns
mY <- 100*y

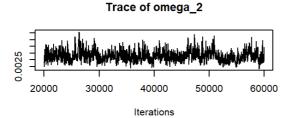
# generates the Markov Chain
start <- Sys.time()

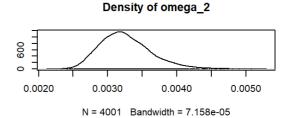
out <- bayesDccGarch(mY, control=list(print=FALSE))</pre>
```

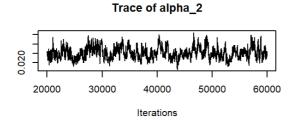
```
## Maximizing the log-posterior density function.
## Done.
## One approximation for covariance matrix of parameters cannot be directly computed through
the hessian matrix.
## Calibrating the standard deviations for simulation:
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
##
   0.33
          0.10
                 0.21
                       0.09
                               0.12
                                     0.10
                                             0.22
                                                    0.09
                                                         0.10
                                                                  0.26
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
   0.37
          0.19
                 0.22
                       0.14
                              0.16
                                     0.18
                                            0.23
                                                    0.12
                                                         0.17
                                                                  0.29
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
   0.36
           0.15
                 0.17 0.21 0.18
                                     0.15 0.21
                                                    0.19
                                                         0.14
                                                                  0.28
##
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
   0.36
           0.17
                  0.20
                         0.19
                                0.16
                                      0.19
                                             0.21
                                                    0.23 0.28
                                                                  0.26
## Computing the covariance matrix of pilot sample.
## Warning in if (class(control$cholCov) != "try-error") {: a condição tem
## comprimento > 1 e somente o primeiro elemento será usado
## Done.
## Calibrating the Lambda coefficient:
## lambda: 0.4
## Accept Rate: 0.15
## lambda: 0.32
## Accept Rate: 0.23
## Done.
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
## Done.
out2 <- increaseSim(out, nSim=50000)
## Calibrating the Lambda coefficient:
## lambda: 0.32
## Accept Rate: 0.19
## lambda: 0.256
## Accept Rate: 0.25
## Done.
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
out <- window(out2, start=20000, thin=10)
rm(out2)
end <- Sys.time()</pre>
# elapsed time
end-start
## Time difference of 4.404351 mins
```

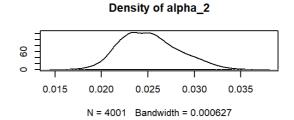
# plot Markov Chain
plot(out\$MC)

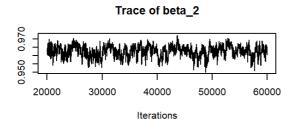


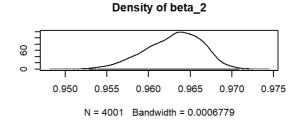


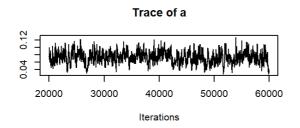


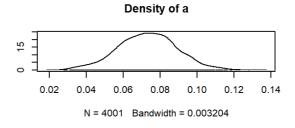


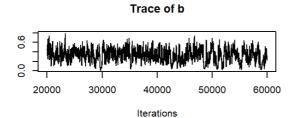


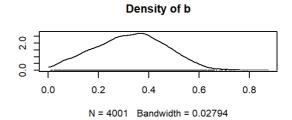












## Estimative of parameters
out\$MC %>% summary()

```
##
## Iterations = 20000:60000
## Thinning interval = 10
## Number of chains = 1
## Sample size per chain = 4001
## 1. Empirical mean and standard deviation for each variable,
##
      plus standard error of the mean:
##
##
                           SD Naive SE Time-series SE
## nu
           5.412773 0.2601698 4.113e-03
                                             1.757e-02
## gamma 1 0.978768 0.0162001 2.561e-04
                                             1.249e-03
## omega_1 0.003521 0.0004440 7.019e-06
                                             3.481e-05
## alpha_1 0.026142 0.0035511 5.614e-05
                                            5.196e-04
## beta_1 0.961705 0.0039553 6.253e-05
                                            6.108e-04
## gamma_2 0.968648 0.0168364 2.662e-04
                                             1.323e-03
## omega_2 0.003260 0.0003674 5.808e-06
                                             1.997e-05
## alpha_2 0.025180 0.0031074 4.913e-05
                                             2.697e-04
## beta_2 0.962987 0.0033597 5.311e-05
                                             2.795e-04
          0.073105 0.0158781 2.510e-04
                                             9.943e-04
## b
           0.328361 0.1384858 2.189e-03
                                             8.678e-03
##
## 2. Quantiles for each variable:
##
##
               2.5%
                         25%
                                  50%
                                           75%
                                                  97.5%
## nu
           4.920120 5.231977 5.409097 5.584851 5.938925
## gamma_1 0.948338 0.967616 0.978299 0.990394 1.010386
## omega_1 0.002792 0.003225 0.003460 0.003751 0.004617
## alpha 1 0.019722 0.023583 0.025967 0.028423 0.033595
## beta_1 0.953229 0.959219 0.961888 0.964419 0.968818
## gamma 2 0.934407 0.957154 0.968922 0.979887 1.001104
## omega_2 0.002665 0.002997 0.003222 0.003472 0.004094
## alpha_2 0.019939 0.022904 0.024913 0.027179 0.031758
## beta_2 0.955796 0.960782 0.963339 0.965448 0.968660
## a
          0.040734 0.062370 0.073216 0.083784 0.104769
## b
           0.062019 0.229617 0.332876 0.426934 0.587236
```

```
# Prepare input for the expert advisor
parEst <- summary(out)$statistics[,'Mean']</pre>
## High
#HBOP
High_{UB_{HBOP}} = qsstd(p=1-(1-C_{Trend})/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gam = 0]
#S1
High_UB_S1 = qsstd(p=1-(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['ga
mma 1'])
## I OW
Low_LB_B1 = qsstd(p=(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma
_2'])
#LBOP
Low_LB_LBOP = qsstd(p=(1-C_Trend)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_
2'])
m = matrix(NA,nrow=10,ncol=1)
rownames(m) = c("High_UB_HBOP","High_UB_S1","Low_LB_B1","Low_LB_LBOP",
               "High_omega", "High_alpha", "High_beta",
                     "Low_omega", "Low_alpha", "Low_beta")
colnames(m) = 'Value'
m["High_UB_HBOP",1] = High_UB_HBOP
m["High_UB_S1",1] = High_UB_S1
m["Low_LB_B1",1] = Low_LB_B1
m["Low_LB_LBOP",1] = Low_LB_LBOP
m["High_omega",1] = parEst["omega_1"]
m["High_alpha",1] = parEst["alpha_1"]
m["High_beta",1] = parEst["beta_1"]
m["Low_omega",1] = parEst["omega_2"]
m["Low alpha",1] = parEst["alpha 2"]
m["Low_beta",1] = parEst["beta_2"]
# Input for expert advisor
print(round(m,3))
```

```
Value
## High UB HBOP 1.971
                 0.578
## High_UB_S1
## Low LB B1
                -0.567
## Low_LB_LBOP -2.031
## High_omega
                 0.004
## High alpha
                 0.026
                 0.962
## High beta
## Low omega
                 0.003
## Low_alpha
                 0.025
## Low_beta
                 0.963
```