# GARCH parameters and quantiles estimation

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## Input

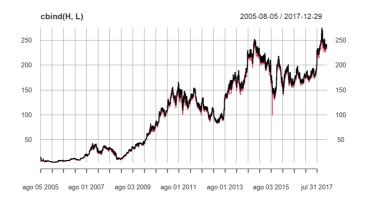
```
symbol = "BIDU"
from=as.Date('2000-01-01')
to=as.Date('2017-12-31')
C_Trend = 0.95
C_Reaction = 0.50
```

#### Data download

```
x <- getSymbols.yahoo(symbol,auto.assign = FALSE, from=from, to=to)
```

#### High and Low

```
H <- Hi(x)
L <- Lo(x)
C <- Cl(x)
plot(cbind(H,L))</pre>
```

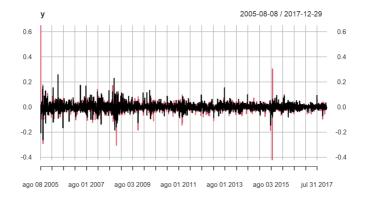


#### Returns

```
y <- cbind( diff(log(H)),  diff(log(L)) )
y <- na.omit(y)
y %>% cor() # Returns correlation
```

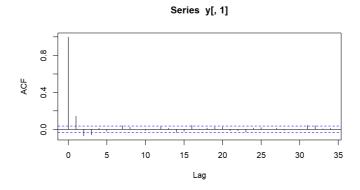
```
## BIDU.High BIDU.Low
## BIDU.High 1.0000000 0.6907737
## BIDU.Low 0.6907737 1.0000000
```

plot(y)

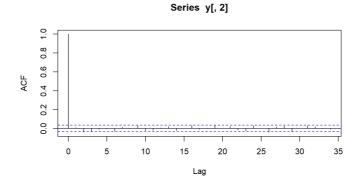


### Autocorrelation

acf(y[,1])

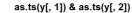


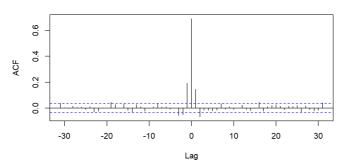
acf(y[,2])



#### **Cross correlation**

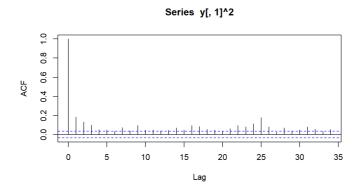
ccf(as.ts(y[,1]),as.ts(y[,2]))



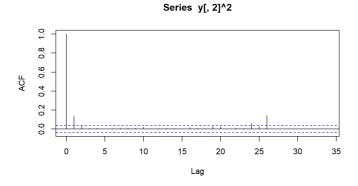


### Volatility verification

acf(y[,1]^2)



acf(y[,2]^2)



# **Bivariate DCC-GARCH**

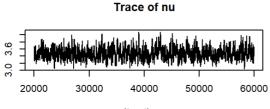
We will consider the DCC-GARCH to model the volatility of  $y=(r_H,r_L)'$ , where  $r_H$  and  $r_L$  denote the  $100\times$ log-returns from hight's and low's observations.

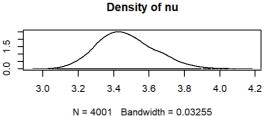
```
# returns
mY <- 100*y

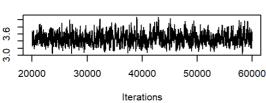
# generates the Markov Chain
start <- Sys.time()

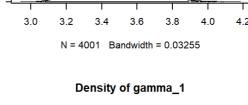
out <- bayesDccGarch(mY, control=list(print=FALSE, nPilotSim=3000))</pre>
```

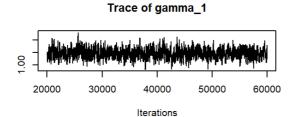
```
## Maximizing the log-posterior density function.
## Done.
## One approximation for covariance matrix of parameters cannot be directly computed through
the hessian matrix.
## Calibrating the standard deviations for simulation:
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
##
   0.31
          0.11
                 0.19 0.12 0.14
                                     0.11 0.18
                                                   0.14
                                                         0.16
                                                                  0.45
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
   0.32
          0.20
                0.20
                       0.20 0.23
                                     0.20 0.18
                                                   0.22
                                                          0.15
                                                                  0.46
## Accept Rate:
## phi_1 phi_2 phi_3 phi_4 phi_5 phi_6 phi_7 phi_8 phi_9 phi_10 phi_11
                 0.19
                       0.19
                               0.22
                                     0.20 0.18
                                                   0.23 0.17
          0.19
## Computing the covariance matrix of pilot sample.
## Warning in if (class(control$cholCov) != "try-error") {: a condição tem
## comprimento > 1 e somente o primeiro elemento será usado
## Done.
## Calibrating the Lambda coefficient:
## lambda: 0.4
## Accept Rate: 0.44
## Done.
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
## Done.
out2 <- increaseSim(out, nSim=50000)
## Calibrating the Lambda coefficient:
## lambda: 0.4
## Accept Rate: 0.47
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
## Done.
out <- window(out2, start=20000, thin=10)
rm(out2)
end <- Sys.time()</pre>
# elapsed time
end-start
## Time difference of 7.961156 mins
## Estimative of parameters
parEst <- summary(out)$statistics[,'Mean']</pre>
# plot Markov Chain
plot(out$MC)
```

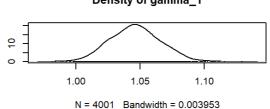


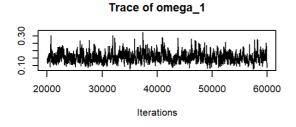


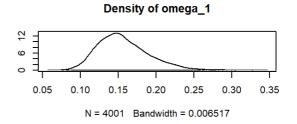


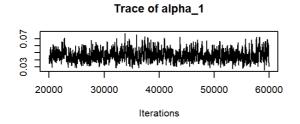


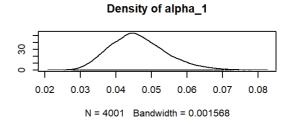


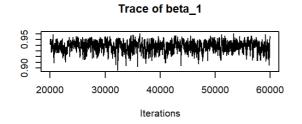


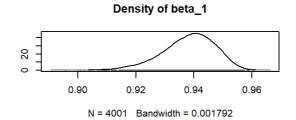


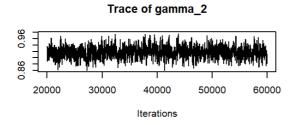


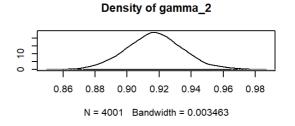


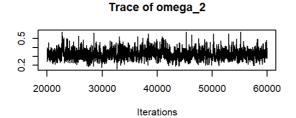


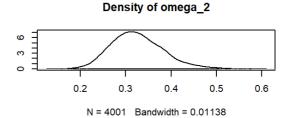


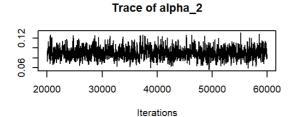


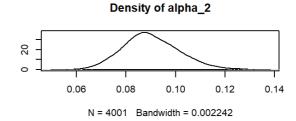


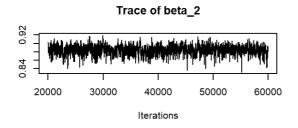


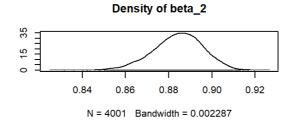


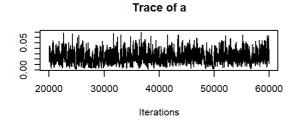


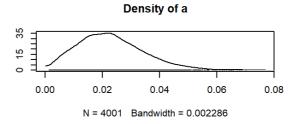


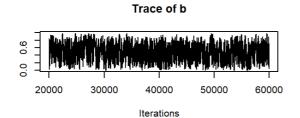


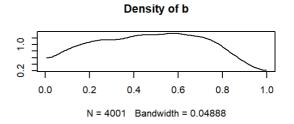












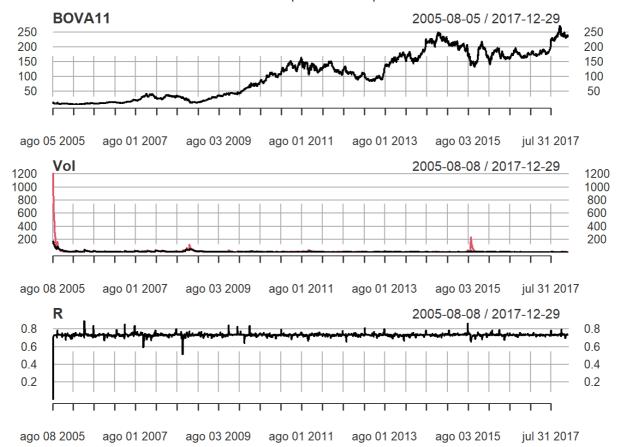
## Estimative of parameters
out\$MC %>% summary()

```
##
## Iterations = 20000:60000
## Thinning interval = 10
## Number of chains = 1
## Sample size per chain = 4001
##
## 1. Empirical mean and standard deviation for each variable,
##
      plus standard error of the mean:
##
##
                         SD Naive SE Time-series SE
## nu
           3.46953 0.161295 0.0025500
                                           0.0070830
## gamma 1 1.04634 0.019685 0.0003112
                                           0.0007400
## omega_1 0.15656 0.033671 0.0005323
                                         0.0016143
## alpha_1 0.04628 0.007827 0.0001237
                                         0.0003444
## beta 1 0.93806 0.009085 0.0001436
                                         0.0004384
## gamma_2 0.91714 0.017624 0.0002786
                                          0.0006315
## omega 2 0.32553 0.057318 0.0009062
                                          0.0019317
## alpha_2 0.09006 0.011183 0.0001768
                                         0.0004378
## beta_2 0.88423 0.011487 0.0001816
                                         0.0004219
## a
          0.02300 0.011366 0.0001797
                                          0.0004206
## b
           0.47363 0.242272 0.0038302
                                           0.0087305
##
## 2. Quantiles for each variable:
##
##
                        25%
                                50%
                                              97.5%
               2.5%
                                        75%
## nu
           3.182443 3.35403 3.45468 3.57429 3.81363
## gamma_1 1.009616 1.03277 1.04611 1.05902 1.08500
## omega_1 0.102939 0.13249 0.15214 0.17577 0.23124
## alpha 1 0.032849 0.04073 0.04558 0.05114 0.06375
## beta 1 0.917443 0.93263 0.93905 0.94452 0.95311
## gamma 2 0.882842 0.90543 0.91696 0.92843 0.95364
## omega_2 0.228322 0.28534 0.32091 0.36091 0.45425
## alpha_2 0.069952 0.08238 0.08922 0.09727 0.11381
## beta_2 0.860157 0.87698 0.88495 0.89217 0.90495
## a
          0.004653 0.01479 0.02202 0.02997 0.04853
## b
           0.049212 0.27342 0.47727 0.67022 0.90255
## Conditional Correlation
R <- xts(out$R[,2], order.by=index(y))</pre>
```

```
## Conditional Correlation
R <- xts(out$R[,2], order.by=index(y))

## Volatility
Vol <- xts(out$H[,c("H_1,1","H_2,2")], order.by=index(y))

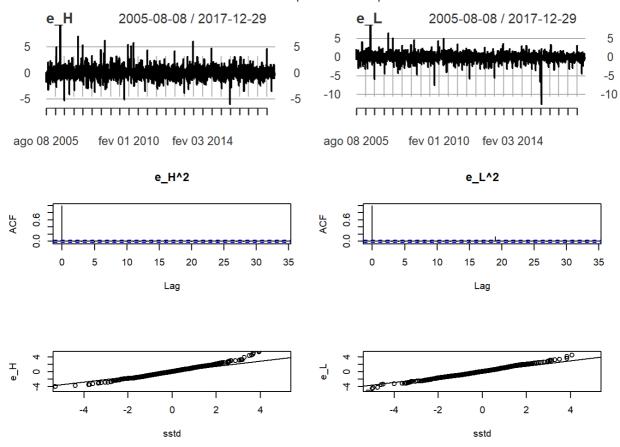
par(mfrow=c(3,1))
plot(C, main="BOVA11")
plot(Vol)
plot(R, main="R")</pre>
```



```
## Standard Residuals
r <- mY / sqrt(Vol)

par(mfrow=c(3,2))

plot(r[,1], main="e_H")
plot(r[,2], main="e_L")
acf(r[,1]^2, main="e_H^2")
acf(r[,2]^2, main="e_L^2")
r1 <- as.numeric(r[,1])
x <- rsstd(2000, mean = 0, sd = 1, nu = parEst['nu'], xi =parEst['gamma_1'])
qqplot(x=x, y=r1, xlim=c(-5, 5), ylim=c(-5, 5), ylab="e_H",xlab="sstd")
qqline(r1)
r2 <- as.numeric(r[,2])
x <- rsstd(2000, mean = 0, sd = 1, nu = parEst['nu'], xi =parEst['gamma_2'])
qqplot(x=x, y=r2, xlim=c(-5, 5), ylim=c(-5, 5), ylab="e_L",xlab="sstd")
qqline(r2)</pre>
```



```
# Prepare input for the expert advisor
## High
#HBOP
High UB HBOP = qsstd(p=1-(1-C Trend)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gam
ma_1'])
#S1
High_UB_S1 = qsstd(p=1-(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['ga
mma_1'])
## Low
#B1
Low_LB_B1 = qsstd(p=(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma
_2'])
#LBOP
Low_LB_LBOP = qsstd(p=(1-C_Trend)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_
2'])
pH <- c(0.6, 0.65, 0.7, 0.75, 0.8, 0.85, 0.9, 0.95, 0.975, 0.99, 0.995)
qH \leftarrow round(qsstd(p=pH, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma 1']),3)
names(qH) <- paste0(100*pH,"%")</pre>
pL <- 1 - pH
qL <- round(qsstd(p=pL, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_2']),3)
names(qL) <- paste0(100*pL,"%")
qC <- rbind(qH, qL)
rownames(qC) <- c("High_UB", "Low_LB")</pre>
colnames(qC) <- paste0(100*pL,"%")</pre>
m = matrix(NA,nrow=10,ncol=1)
rownames(m) = c("High_UB_HBOP", "High_UB_S1", "Low_LB_B1", "Low_LB_LBOP",
               "High_omega", "High_alpha", "High_beta",
                      "Low_omega", "Low_alpha", "Low_beta")
colnames(m) = 'Value'
m["High UB HBOP",1] = High UB HBOP
m["High UB S1",1] = High UB S1
m["Low_LB_B1",1] = Low_LB_B1
m["Low_LB_LBOP",1] = Low_LB_LBOP
m["High omega",1] = parEst["omega 1"]
m["High_alpha",1] = parEst["alpha_1"]
m["High_beta",1] = parEst["beta_1"]
m["Low omega",1] = parEst["omega 2"]
m["Low alpha",1] = parEst["alpha 2"]
m["Low beta",1] = parEst["beta 2"]
# Input for expert advisor
print(qC)
```

#### print(round(m,3))

```
Value
High_UB_HBOP 1.978
             0.478
High_UB_S1
Low_LB_B1
            -0.467
Low_LB_LBOP -2.027
High_omega
             0.157
High_alpha
             0.046
High_beta
             0.938
Low_omega
             0.326
Low_alpha
             0.090
Low_beta
              0.884
```