# GARCH parameters and quantiles estimation

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# Input

```
symbol = "BOVA11.SA"
from=as.Date('2000-01-01')
to=as.Date('2017-12-31')
C_Trend = 0.95
C_Reaction = 0.50
```

# Data download

```
x <- getSymbols.yahoo(symbol,auto.assign = FALSE, from=from, to=to)
```

# High and Low

```
H <- Hi(x)
L <- Lo(x)
C <- Cl(x)
plot(cbind(H,L))</pre>
```



#### Returns

## BOVA11.SA.High

```
y <- cbind( diff(log(H)),  diff(log(L)) )
y <- na.omit(y)
y %>% cor() # Returns correlation
## BOVA11.SA.High BOVA11.SA.Low
```

0.7085079

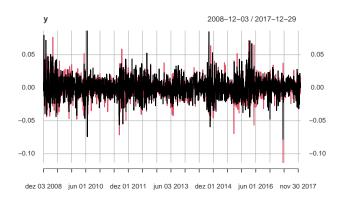
1.0000000

## BOVA11.SA.Low

0.7085079

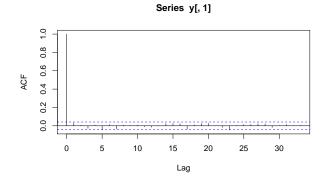
1.0000000

plot(y)

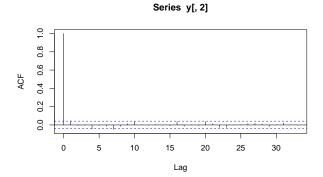


# Autocorrelation

acf(y[,1])

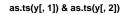


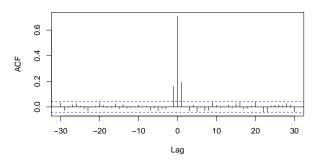
acf(y[,2])



# Cross correlation

ccf(as.ts(y[,1]),as.ts(y[,2]))

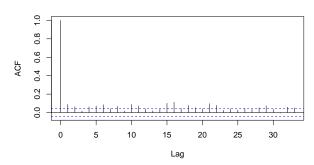




# Volatility verification

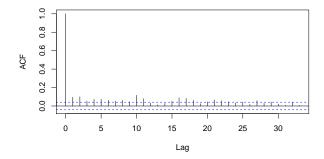
```
acf(y[,1]^2)
```

# Series y[, 1]^2



# acf(y[,2]^2)

# Series y[, 2]^2



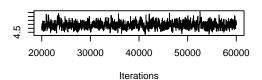
# Bivariate DCC-GARCH

We will consider the DCC-GARCH to model the volatility of  $y = (r_H, r_L)'$ , where  $r_H$  and  $r_L$  denote the  $100 \times \text{log-returns}$  from hight's and low's observations.

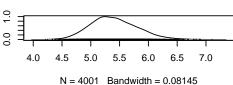
```
# returns
mY <- 100*y
# generates the Markov Chain</pre>
```

```
start <- Sys.time()</pre>
out <- bayesDccGarch(mY, control=list(print=FALSE, nPilotSim=3000))</pre>
## Maximizing the log-posterior density function.
## Done.
## Warning in if (class(control$cholCov) != "try-error") {: a condição tem
## comprimento > 1 e somente o primeiro elemento será usado
## Calibrating the Lambda coefficient:
## lambda: 0.4
## Accept Rate: 0.35
## Done.
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
## Done.
out2 <- increaseSim(out, nSim=50000)</pre>
## Calibrating the Lambda coefficient:
## lambda: 0.4
## Accept Rate: 0.37
## Done.
## Starting the simulation by one-block random walk Metropolis-Hasting algorithm.
out <- window(out2, start=20000, thin=10)</pre>
rm(out2)
end <- Sys.time()</pre>
# elapsed time
end-start
## Time difference of 1.453167 mins
## Estimative of parameters
parEst <- summary(out)$statistics[,'Mean']</pre>
# plot Markov Chain
plot(out$MC)
```

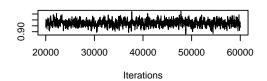




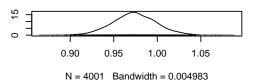
# Density of nu



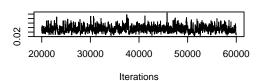
# Trace of gamma\_1



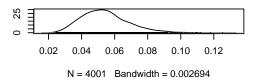
# Density of gamma\_1



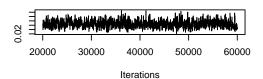
#### Trace of omega\_1



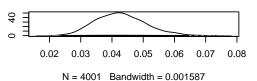
#### Density of omega\_1



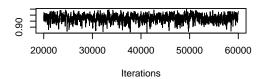
#### Trace of alpha\_1



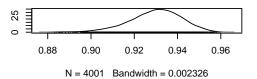
#### Density of alpha\_1



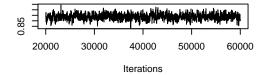
### Trace of beta\_1



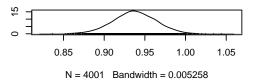
# Density of beta\_1



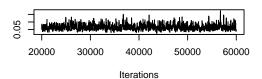
# Trace of gamma\_2



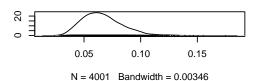
# Density of gamma\_2



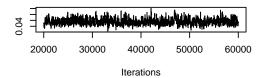
#### Trace of omega\_2



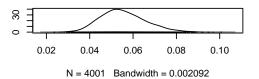
#### Density of omega\_2



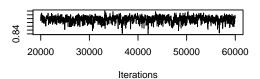
# Trace of alpha\_2



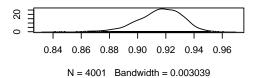
# Density of alpha\_2



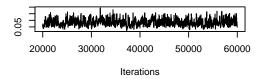
# Trace of beta\_2



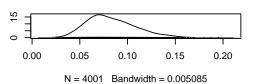
#### Density of beta\_2



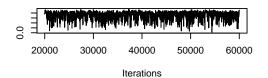
#### Trace of a



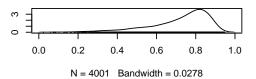
# Density of a



#### Trace of b



#### Density of b

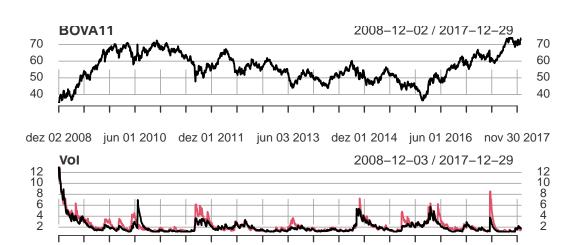


## Markov Chain convergence
out\$MC %>% geweke.diag()

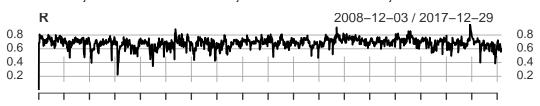
```
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
##
         nu gamma_1 omega_1 alpha_1 beta_1 gamma_2 omega_2 alpha_2
## 0.66276 0.95051 -1.06946 -0.55557 0.95487 -0.58825 -0.98527 0.01639
##
   beta 2
                  a
## 0.43885 -0.91982 1.13178
## Model adequability / residual hypothesis verification
Vol <- xts(out$H[,c("H_1,1","H_2,2")], order.by=index(y)) # estimated volatility
r <- mY / sqrt(Vol) # Standard Residuals
par(mfrow=c(3,2))
plot(r[,1], main="e_H")
plot(r[,2], main="e_L")
# Volatility test
acf(r[,1]^2, main="e_H^2")
acf(r[,2]^2, main="e_L^2")
Box.test(r[,1]^2)
##
## Box-Pierce test
##
## data: r[, 1]^2
## X-squared = 0.15303, df = 1, p-value = 0.6957
Box.test(r[,2]^2)
##
## Box-Pierce test
##
## data: r[, 2]^2
## X-squared = 0.38652, df = 1, p-value = 0.5341
# adequability
r1 <- as.numeric(r[,1])
x <- rsstd(2000, mean = 0, sd = 1, nu = parEst['nu'], xi =parEst['gamma_1'])
qqplot(x=x, y=r1, xlim=c(-5, 5), ylim=c(-5, 5), ylab="e_H",xlab="sstd")
qqline(r1)
ks.test(r1 , x)
## Warning in ks.test(r1, x): p-value will be approximate in the presence of ties
##
## Two-sample Kolmogorov-Smirnov test
##
## data: r1 and x
## D = 0.035313, p-value = 0.1424
## alternative hypothesis: two-sided
r2 <- as.numeric(r[,2])
x \leftarrow rsstd(2000, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_2'])
qqplot(x=x, y=r2, xlim=c(-5, 5), ylim=c(-5, 5), ylab="e_L", xlab="sstd")
qqline(r2)
```

```
e_H 2008-12-03 / 2017-12-29
                                                      e L 2008-12-03 / 2017-12-29
      64
20
-2
-4
-6
                                            6420-24-6
                                                                                        0
                                                 -5
                                                                                       -5
     dez 03 2008 jun 01 2012 dez 01 2015
                                                 dez 03 2008 jun 01 2012 dez 01 2015
                         e_H^2
                                                                    e_L^2
                                                ACF
                                                   0.0
            0
                 5
                     10
                          15
                              20
                                   25
                                       30
                                                            5
                                                                10
                                                                     15
                                                                         20
                                                                              25
                                                                                  30
                          Lag
                                                                      Lag
                     -2
                           0
                                 2
                                       4
                                                                -2
                                                                      0
                                                                             2
                                                                                   4
                          sstd
                                                                      sstd
ks.test(r2 , x)
## Warning in ks.test(r2, x): p-value will be approximate in the presence of ties
##
    Two-sample Kolmogorov-Smirnov test
##
##
## data: r2 and x
## D = 0.02323, p-value = 0.617
## alternative hypothesis: two-sided
## Estimative of parameters
out$MC %>% summary()
##
## Iterations = 20000:60000
## Thinning interval = 10
## Number of chains = 1
## Sample size per chain = 4001
##
  1. Empirical mean and standard deviation for each variable,
##
      plus standard error of the mean:
##
                          SD Naive SE Time-series SE
##
               Mean
                                              0.0173865
           5.40439 0.403682 0.0063820
## nu
## gamma_1 0.97378 0.024991 0.0003951
                                              0.0010712
## omega_1 0.05410 0.014011 0.0002215
                                              0.0006568
                                              0.0003122
## alpha_1 0.04275 0.007864 0.0001243
## beta_1 0.92969 0.011954 0.0001890
                                              0.0005318
## gamma_2 0.93726 0.026225 0.0004146
                                              0.0011520
```

```
## omega_2 0.06615 0.017711 0.0002800
                                           0.0007377
## alpha_2 0.05544 0.010424 0.0001648
                                            0.0004405
## beta 2 0.91581 0.015446 0.0002442
                                            0.0007026
           0.08303 0.025202 0.0003984
                                            0.0012115
## a
## b
           0.71678 0.154192 0.0024377
                                            0.0058450
##
## 2. Quantiles for each variable:
##
##
              2.5%
                       25%
                               50%
                                        75%
                                              97.5%
## nu
           4.69180 5.12251 5.37491 5.66817 6.26880
## gamma_1 0.92499 0.95726 0.97365 0.99035 1.02467
## omega_1 0.03122 0.04415 0.05274 0.06204 0.08585
## alpha_1 0.02896 0.03710 0.04231 0.04786 0.05982
## beta_1 0.90342 0.92256 0.93052 0.93801 0.95081
## gamma_2 0.88482 0.91987 0.93709 0.95479 0.98835
## omega_2 0.03797 0.05335 0.06399 0.07632 0.10532
## alpha_2 0.03743 0.04808 0.05451 0.06197 0.07806
## beta 2 0.88203 0.90650 0.91708 0.92668 0.94183
           0.04220 0.06464 0.07986 0.09918 0.13992
## a
           0.31689 0.64358 0.76210 0.82818 0.90143
## b
## Conditional Correlation
R <- xts(out$R[,2], order.by=index(y))</pre>
## Volatility
Vol <- xts(out$H[,c("H_1,1","H_2,2")], order.by=index(y))</pre>
par(mfrow=c(3,1))
plot(C, main="BOVA11")
plot(Vol)
plot(R, main="R")
```



dez 03 2008 jun 01 2010 dez 01 2011 jun 03 2013 dez 01 2014 jun 01 2016 nov 30 2017



dez 03 2008 jun 01 2010 dez 01 2011 jun 03 2013 dez 01 2014 jun 01 2016 nov 30 2017

```
# Prepare input for the expert advisor
## High
#HBOP
High_UB_HBOP = qsstd(p=1-(1-C_Trend)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_1'])
High_UB_S1 = qsstd(p=1-(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_1'])
## Low
#B1
Low_LB_B1 = qsstd(p=(1-C_Reaction)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_2'])
Low_LB_LBOP = qsstd(p=(1-C_Trend)/2, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_2'])
pH \leftarrow c(0.6, 0.65, 0.7, 0.75, 0.8, 0.85, 0.9, 0.95, 0.975, 0.99, 0.995)
qH <- round(qsstd(p=pH, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_1']),3)
names(qH) <- paste0(100*pH,"%")</pre>
pL \leftarrow 1 - pH
qL <- round(qsstd(p=pL, mean = 0, sd = 1, nu = parEst['nu'], xi = parEst['gamma_2']),3)
names(qL) <- paste0(100*pL,"%")
qC <- rbind(qH, qL)
rownames(qC) <- c("High_UB", "Low_LB")</pre>
colnames(qC) <- paste0(100*pL,"%")</pre>
m = matrix(NA,nrow=10,ncol=1)
rownames(m) = c("High_UB_HBOP","High_UB_S1","Low_LB_B1","Low_LB_LBOP",
               "High_omega", "High_alpha", "High_beta",
```

```
"Low_omega", "Low_alpha", "Low_beta" )
colnames(m) = 'Value'
m["High_UB_HBOP",1] = High_UB_HBOP
m["High_UB_S1",1] = High_UB_S1
m["Low_LB_B1",1] = Low_LB_B1
m["Low_LB_LBOP",1] = Low_LB_LBOP
m["High_omega",1] = parEst["omega_1"]
m["High_alpha",1] = parEst["alpha_1"]
m["High_beta",1] = parEst["beta_1"]
m["Low omega",1] = parEst["omega 2"]
m["Low_alpha",1] = parEst["alpha_2"]
m["Low_beta",1] = parEst["beta_2"]
# Input for expert advisor
print(qC)
          40%
                 35%
                        30%
                               25%
                                      20%
                                             15%
                                                    10%
                                                            5%
                                                                 2.5%
                                                                          1%
High_UB 0.222 0.332 0.450 0.579 0.727 0.908 1.151 1.555 1.965 2.541
Low_LB -0.185 -0.300 -0.423 -0.560 -0.718 -0.913 -1.176 -1.616 -2.066 -2.702
         0.5%
High_UB 3.018
Low_LB -3.228
print(round(m,3))
             Value
High UB HBOP 1.965
High_UB_S1
             0.579
Low_LB_B1
            -0.560
Low_LB_LBOP -2.066
High_omega
             0.054
High_alpha
             0.043
High_beta
             0.930
Low_omega
             0.066
             0.055
Low_alpha
Low_beta
             0.916
```