

PROBLEM #1:

For each of the following statements, determine whether it is true or false. Label "T" if it is true, otherwise label "F".

- a. If  $X_1, \dots, X_n$  are independent and are from a distribution with mean  $\mu$  and variance  $\sigma^2$ , then the sample variance  $S^2$  is an unbiased estimator of the variance  $\sigma^2$  where

$$S^2 = \frac{1}{n-1} \left( \sum_{i=1}^n (X_i - \bar{X}) \right). \quad F$$

- b. If  $X_1, \dots, X_n$  are from a distribution with mean  $\mu$  and variance  $\sigma^2$ , then the sample mean  $\bar{X}$  is an unbiased estimator of  $\mu$ . T

- c. If  $\hat{\theta}$  is a point estimator of  $\theta$  with  $E(\hat{\theta}) = \theta$ , then  $\hat{\theta}$  is an unbiased estimator of  $\theta$ . T

- d. If  $X_1, \dots, X_n$  is a random sample from a normal distribution with mean  $\mu$  and variance  $\sigma^2$ , then  $\bar{X}$  is the MVUE for  $\mu$ . T

- e. The likelihood function tells us how likely the observed sample is as a function of the possible parameter values. T

- f.  $S^2$  defined above as a point estimator of the variance  $\sigma^2$  can also be deduced from maximum likelihood estimation. F

- g. Let  $\hat{\theta}$  be the maximum likelihood estimator of the parameter  $\theta$ . Then the maximum likelihood estimator of  $\sqrt{\theta}$  is  $\sqrt{\hat{\theta}}$ . F

- h. The point estimator of the population proportion  $p$  is the sample proportion  $\hat{p} = X/n$  where  $n$  is the sample size and  $X$  is the number of "successes" in the sample. F

- i. The maximum likelihood estimator  $\hat{\theta}$  of a population parameter  $\theta$  is its MVUE. T

- j. If  $X_1, \dots, X_n$  is a random sample from a normal distribution with mean  $\mu$  and variance  $\sigma^2$ , then the maximum likelihood estimator of  $\mu$  is  $\bar{X}$  which is the same as the one obtained from the method of moments. T

SOLUTION:

- a) F  
b) T  
c) T  
d) T  
e) T  
f) F  
g) T  
h) F  
i) T  
j) T

- i) T  
j) T

**PROBLEM #2:**

Consider a random sample  $X_1, \dots, X_n$  from the shifted exponential pdf

$$f(x; \lambda, \theta) = \begin{cases} \lambda e^{-\lambda(x-\theta)}, & x \geq \theta, \\ 0, & \text{otherwise.} \end{cases}$$

Obtain the maximum likelihood estimators of  $\theta$  and  $\lambda$ . Show your work.

**SOLUTION:**

Example 7.17 textbook

$$f(x_1, \dots, x_n; \lambda) = (\lambda e^{-\lambda x_1}) \dots (\lambda e^{-\lambda x_n}) = \lambda^n e^{-\lambda \sum x_i}$$

$$\ln [f(x_1, \dots, x_n; \lambda)] = n \ln(\lambda) - \lambda \sum x_i$$

$$\frac{n}{\lambda} - \sum x_i = 0 \text{ or } \lambda = \frac{n}{\sum x_i} = \frac{1}{\bar{x}} \Rightarrow \boxed{\hat{\lambda} = \frac{1}{\bar{x}}}$$

$$\ln L = n \ln \lambda - \lambda \sum_{i=1}^n (x_i - \theta)$$

$$\frac{\partial}{\partial \lambda} \ln L = \frac{n}{\lambda} - \sum x_i$$

$$\frac{\partial}{\partial \lambda} \ln L = 0$$

$$= \frac{n}{\lambda} - \sum x_i \rightarrow \lambda = \frac{n}{\sum x_i} = \boxed{\frac{1}{\bar{x}} = \hat{\lambda}}$$

$$\boxed{\theta}$$

$$\frac{\partial}{\partial \theta} \ln L = \lambda \Rightarrow 0$$

$$x \geq \theta$$

Thus

$$\boxed{\hat{\theta} = x(1)}$$

### PROBLEM #3:

Let  $X$  denote the proportion of allotted time that a randomly selected student spends working on a certain aptitude test. Suppose the pdf of  $X$  is

$$f(x; \theta) = \begin{cases} (\theta + 1)x^\theta, & 0 \leq x \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

where  $-1 < \theta$ . A random sample of ten students yields data  $x_1 = 0.92, x_2 = 0.79, x_3 = 0.90, x_4 = 0.65, x_5 = 0.86, x_6 = 0.47, x_7 = 0.73, x_8 = 0.97, x_9 = 0.94, x_{10} = 0.77$ .

- Use the method of moments to obtain an estimator of  $\theta$ , and then compute the estimate for this data. Show your work.
- Obtain the maximum likelihood estimator of  $\theta$ , and then compute the estimate for the given data. Show your work.

$$\theta = 0.6 / .20$$

$$\hat{\theta} = 3.4$$

SOLUTION:

$$E(x) = \int x f(x) dx$$

$$\begin{aligned} E(x) &= \int_0^1 x (\theta + 1) x^\theta dx \\ &= \int_0^1 (\theta + 1) x^{\theta+1} dx \\ &= (\theta + 1) \left[ \frac{x^{\theta+2}}{\theta+2} \right]_0^1 \\ &= (\theta + 1) \left[ \frac{1^{\theta+2}}{\theta+2} - \frac{0^{\theta+2}}{\theta+1} \right] \\ &= \frac{\theta+1}{\theta+2} = \bar{x} \Rightarrow \theta+1 = \bar{x}(\theta+2) \\ &\quad \theta+1 = \bar{x}\theta + \bar{x}2 \\ &\quad \theta - \bar{x}\theta = 2\bar{x} - 1 \\ &\quad \theta(1 - \bar{x}) = 2\bar{x} - 1 \end{aligned}$$

Part A  $\Rightarrow$

$$\theta = \frac{2\bar{x} - 1}{1 - \bar{x}}$$

Part A  
DATA

$$\bar{x} = \frac{\text{Sum}}{\text{total}}$$

$$\begin{aligned} \bar{x} &= .92 + .79 + .90 + .65 + \\ &\quad .86 + .47 + .73 + .97 \\ &\quad + .94 + .77 \\ &\quad \underline{\hspace{1cm}} \\ &\quad 10 \end{aligned}$$

$$\bar{x} = .80$$

$$\theta = \frac{2(.80) - 1}{1 - .80} = \frac{.6}{.20}$$

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Example 7.17 Textbook  
B)  $n \ln(\theta + 1) + \theta \sum \ln x_i$

$$\frac{\partial}{\partial \theta} = \frac{n}{\theta+1} + \theta \sum \ln x_i = 0$$

$$\frac{n}{\theta+1} + \theta \sum \ln x_i \Rightarrow \frac{n}{\theta+1} = -\theta \sum \ln x_i$$

$$\theta+1 = \frac{-n}{\sum \ln x_i}$$

$$\hat{\theta} = \frac{-n}{\sum \ln x_i} - 1$$

PART B DATA

$$\hat{\theta} = \frac{-10}{-2.4} - 1$$

$$\hat{\theta} = 3.16$$

$$\hat{\theta} = 3.12$$

A

PROBLEM #4:

Let  $X_1, \dots, X_n$  be a random sample from the probability density function

$$f(x|\theta) = \theta x^{-2}, \quad 0 < \theta \leq x < \infty.$$

Find the MLE of  $\theta$  and the method of moments estimator of  $\theta$ , respectively. Show your work.

SOLUTION:

$$A) f(x|\theta) = \frac{\theta}{x^2}$$

$$E(x) = \int_0^{\infty} x f(x) dx$$

$$E(x) = \int_0^{\infty} x f\left(\frac{\theta}{x^2}\right) dx$$

$$= \theta \int_0^{\infty} \frac{x}{x^2} dx$$

$$= \theta \int_0^{\infty} \frac{1}{x} dx$$

$$= \theta [\log x]_0^{\infty}$$

$$E(x) = DNE$$

method of moments doesn't exist

MLE

$$L(\theta) = \prod_{i=1}^n \frac{\theta}{x_i^2}$$

$$= \frac{\theta^n}{\prod_{i=1}^n x_i^2}$$

$$\log L(\theta) = n \log \theta - 2 \sum_{i=1}^n \log x_i$$

$$\frac{d}{d\theta} \log L(\theta) = \frac{n}{\theta} > 0$$

When not 0

$$\hat{\theta} = \min \{x_i\}$$



**PROBLEM #5:**

Let  $X_1, \dots, X_n$  be a random sample with their density function being one of the following two. If  $\theta = 0$ , then

$$f(x|\theta) = \begin{cases} 1, & \text{if } 0 < x < 1, \\ 0, & \text{otherwise,} \end{cases}$$

while if  $\theta = 1$ , then

$$f(x|\theta) = \begin{cases} 1/2\sqrt{x}, & \text{if } 0 < x < 1, \\ 0, & \text{otherwise,} \end{cases}$$

Find the MLE of  $\theta$ . Show your work.

**SOLUTION:**

when  $\theta = 0$   $L(\theta|x) = \prod_{i=1}^n f(x_i|\theta)$

when  $\theta = 1$   $L(\theta|x) = \prod_{i=1}^n \frac{1}{2\sqrt{x_i}} = \frac{1}{2^n \prod_{i=1}^n x_i^{1/2}}$

$$= \frac{1}{2^n \prod_{i=1}^n x_i^{1/2}}$$

$$\Rightarrow \left( \frac{1}{4^n \prod_{i=1}^n x_i} \right)^{1/2} \Rightarrow \left( \frac{1}{4^n G} \right)^{1/2}$$

$D = 1/4$

$$= \frac{1}{(4^n G)^{1/2}} \leq \frac{1}{(4^n G)^{1/2}} \leq \frac{1}{4G} \leq \boxed{\frac{1}{4}}$$

thus  $\hat{\theta}_{MLE} = \begin{cases} 1 & G > 1/4 \\ 0 & G < 1/4 \\ 0 & G = 1/4 \end{cases}$