



13. RANDOMIZED ALGORITHMS

- ▶ *contention resolution*
- ▶ *global min cut*
- ▶ *linearity of expectation*
- ▶ *max 3-satisfiability*
- ▶ *universal hashing*
- ▶ *Chernoff bounds*
- ▶ *load balancing*

Lecture slides by Kevin Wayne

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Randomization

Algorithmic design patterns.

- Greedy.
- Divide-and-conquer.
- Dynamic programming.
- Network flow.
- **Randomization.**

in practice, access to a pseudo-random number generator



Randomization. Allow fair coin flip in unit time.

Why randomize? Can lead to simplest, fastest, or only known algorithm for a particular problem.

Ex. Symmetry breaking protocols, graph algorithms, quicksort, hashing, load balancing, Monte Carlo integration, cryptography.



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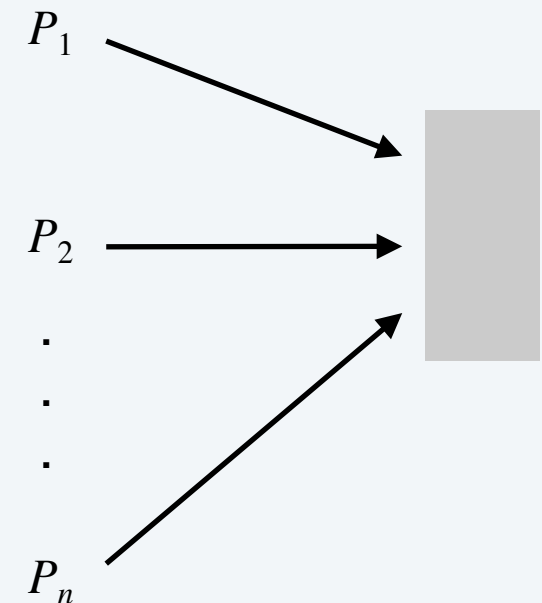
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Contention resolution in a distributed system

Contention resolution. Given n processes P_1, \dots, P_n , each competing for access to a shared database. If two or more processes access the database simultaneously, all processes are locked out. Devise protocol to ensure all processes get through on a regular basis.

Restriction. Processes can't communicate.

Challenge. Need **symmetry-breaking** paradigm.



Contention resolution: randomized protocol

Protocol. Each process requests access to the database at time t with probability $p = 1/n$.

Claim. Let $S[i, t]$ = event that process i succeeds in accessing the database at time t . Then $1 / (e \cdot n) \leq \Pr [S(i, t)] \leq 1/(2n)$.

Pf. By independence, $\Pr [S(i, t)] = p (1 - p)^{n-1}$.

process i requests access

none of remaining $n-1$ processes request access

- Setting $p = 1/n$, we have $\Pr [S(i, t)] = 1/n (1 - 1/n)^{n-1}$. ■

value that maximizes $\Pr[S(i, t)]$

between $1/e$ and $1/2$

Useful facts from calculus. As n increases from 2, the function:

- $(1 - 1/n)^n$ converges monotonically from $1/4$ up to $1/e$.
- $(1 - 1/n)^{n-1}$ converges monotonically from $1/2$ down to $1/e$.

Contention Resolution: randomized protocol

Claim. The probability that process i fails to access the database in en rounds is at most $1/e$. After $e \cdot n (c \ln n)$ rounds, the probability $\leq n^{-c}$.

Pf. Let $F[i, t]$ = event that process i fails to access database in rounds 1 through t . By independence and previous claim, we have

$$\Pr[F[i, t]] \leq (1 - 1/(en))^t.$$

- Choose $t = \lceil e \cdot n \rceil$: $\Pr[F(i, t)] \leq \left(1 - \frac{1}{en}\right)^{\lceil en \rceil} \leq \left(1 - \frac{1}{en}\right)^{en} \leq \frac{1}{e}$
- Choose $t = \lceil e \cdot n \rceil \lceil c \ln n \rceil$: $\Pr[F(i, t)] \leq \left(\frac{1}{e}\right)^{c \ln n} = n^{-c}$

Contention Resolution: randomized protocol

Claim. The probability that **all** processes succeed within $2e \cdot n \ln n$ rounds is $\geq 1 - 1/n$.

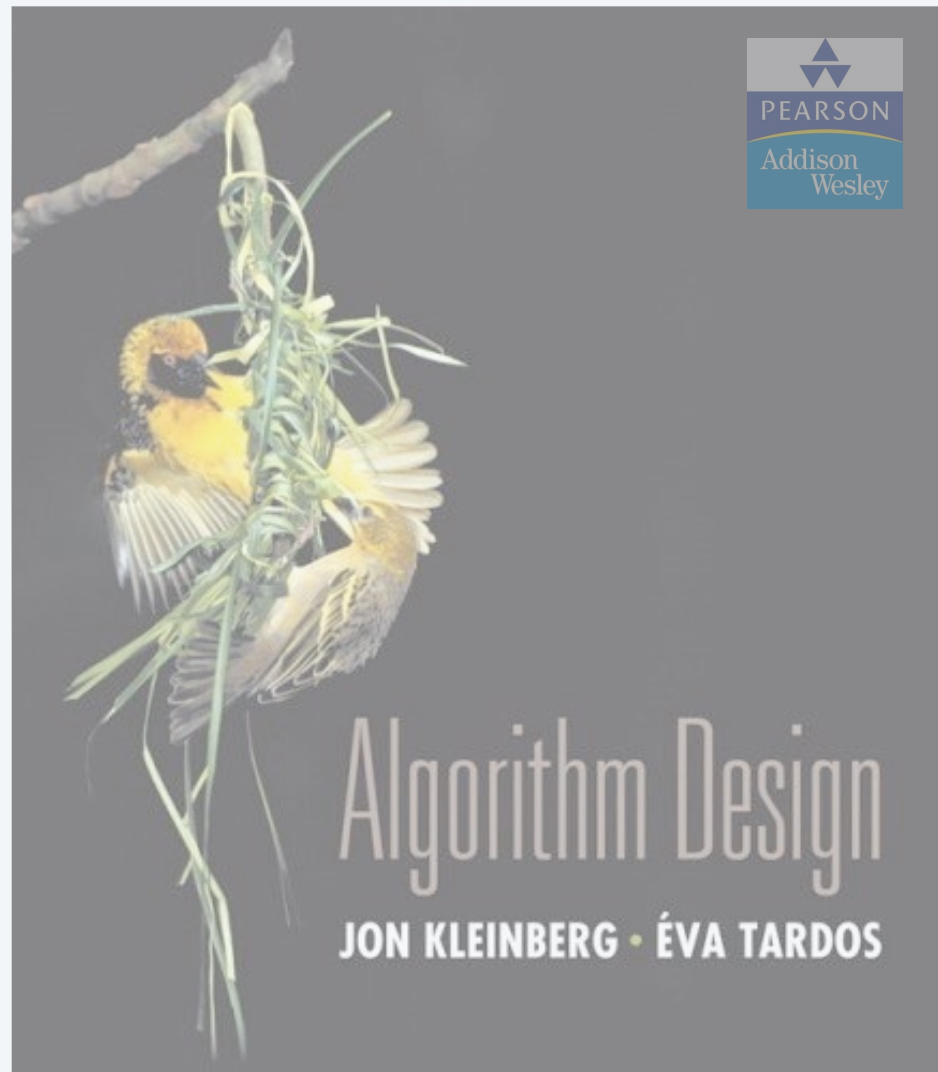
Pf. Let $F[t]$ = event that at least one of the n processes fails to access database in any of the rounds 1 through t .

$$\Pr[F[t]] = \Pr\left[\bigcup_{i=1}^n F[i,t]\right] \leq \sum_{i=1}^n \Pr[F[i,t]] \leq n\left(1 - \frac{1}{en}\right)^t$$

union bound previous slide

- Choosing $t = 2 \lceil en \rceil \lceil c \ln n \rceil$ yields $\Pr[F[t]] \leq n \cdot n^{-2} = 1/n$. ■

$$\Pr\left[\bigcup_{i=1}^n E_i\right] \leq \sum_{i=1}^n \Pr[E_i]$$



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Global minimum cut

Global min cut. Given a connected, undirected graph $G = (V, E)$, find a cut (A, B) of minimum cardinality.

Applications. Partitioning items in a database, identify clusters of related documents, network reliability, network design, circuit design, TSP solvers.

Network flow solution.

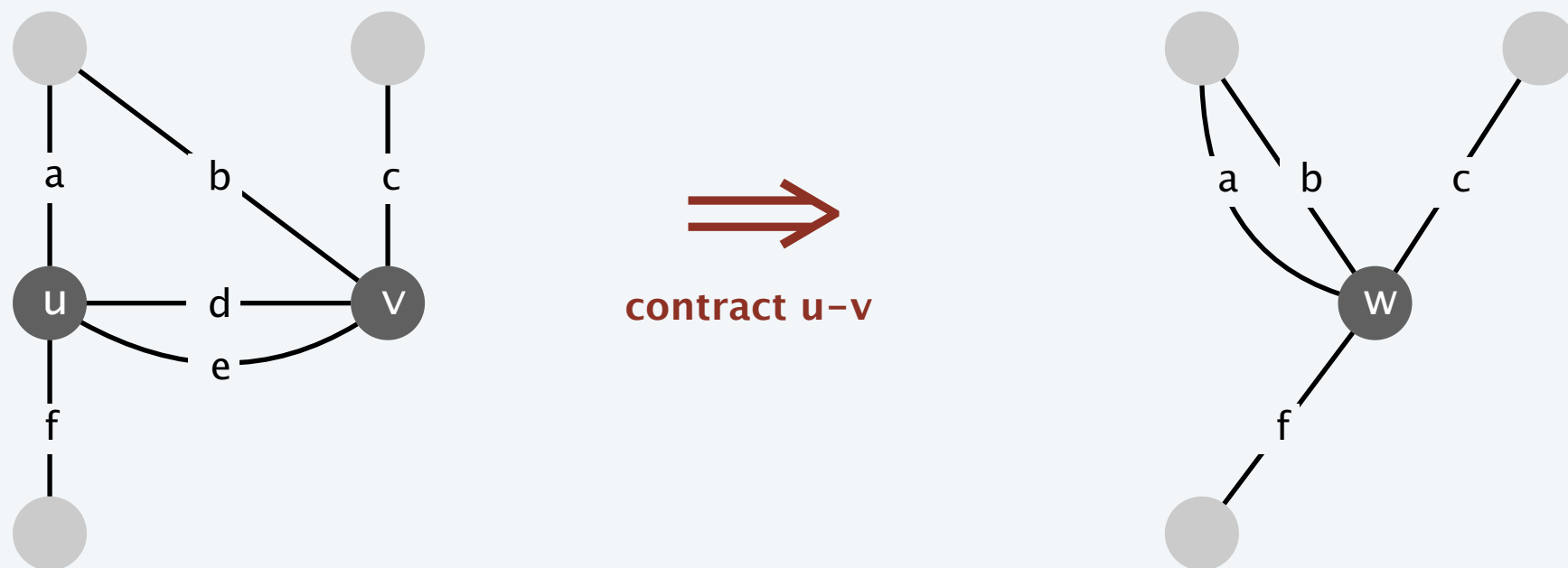
- Replace every edge (u, v) with two antiparallel edges (u, v) and (v, u) .
- Pick some vertex s and compute min s - v cut separating s from each other vertex $v \in V$.

False intuition. Global min-cut is harder than min s - t cut.

Contraction algorithm

Contraction algorithm. [Karger 1995]

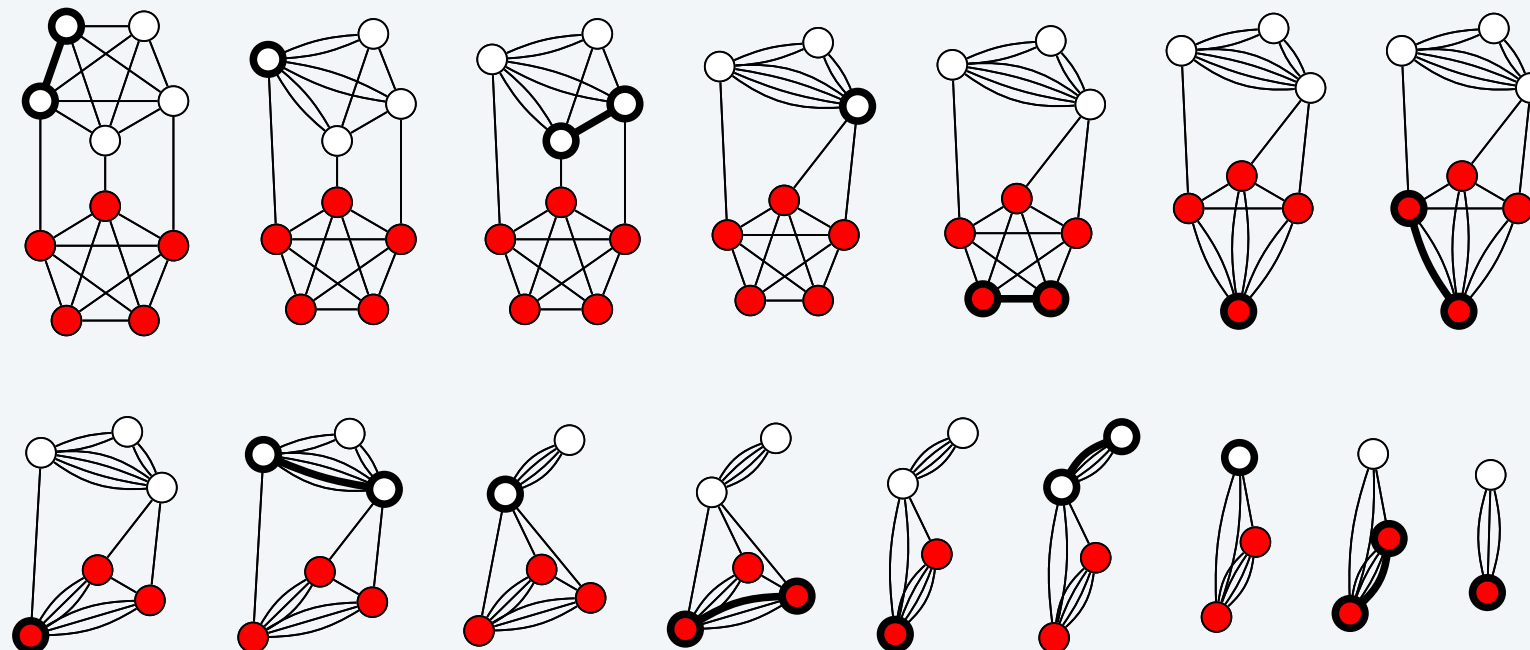
- Pick an edge $e = (u, v)$ uniformly at random.
- **Contract** edge e .
 - replace u and v by single new super-node w
 - preserve edges, updating endpoints of u and v to w
 - keep parallel edges, but delete self-loops
- Repeat until graph has just two nodes v_1 and v_2 .
- Return the cut (all nodes that were contracted to form v_1).



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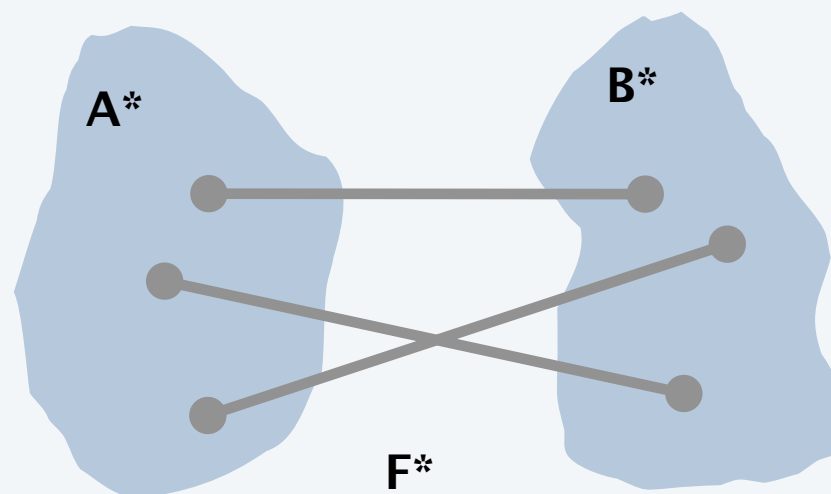
Reference: Thore Husfeldt

Contraction algorithm

Claim. The contraction algorithm returns a min cut with prob $\geq 2 / n^2$.

Pf. Consider a global min-cut (A^*, B^*) of G .

- Let F^* be edges with one endpoint in A^* and the other in B^* .
- Let $k = |F^*|$ = size of min cut.
- In **first step**, algorithm contracts an edge in F^* probability $k / |E|$.
- Every node has degree $\geq k$ since otherwise (A^*, B^*) would not be a min-cut $\Rightarrow |E| \geq \frac{1}{2} k n$.
- Thus, algorithm contracts an edge in F^* with probability $\leq 2 / n$.



Contraction algorithm

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Pf. Consider a global min-cut (A^*, B^*) of G .

- Let F^* be edges with one endpoint in A^* and the other in B^* .
- Let $k = |F^*|$ = size of min cut.
- Let G' be graph after j iterations. There are $n' = n - j$ supernodes.
- Suppose no edge in F^* has been contracted. The min-cut in G' is still k .
- Since value of min-cut is k , $|E'| \geq \frac{1}{2} k n'$.
- Thus, algorithm contracts an edge in F^* with probability $\leq 2 / n'$.
- Let E_j = event that an edge in F^* is not contracted in iteration j .

$$\begin{aligned} \Pr[E_1 \cap E_2 \cdots \cap E_{n-2}] &= \Pr[E_1] \times \Pr[E_2 \mid E_1] \times \cdots \times \Pr[E_{n-2} \mid E_1 \cap E_2 \cdots \cap E_{n-3}] \\ &\geq \left(1 - \frac{2}{n}\right) \left(1 - \frac{2}{n-1}\right) \cdots \left(1 - \frac{2}{4}\right) \left(1 - \frac{2}{3}\right) \\ &= \left(\frac{n-2}{n}\right) \left(\frac{n-3}{n-1}\right) \cdots \left(\frac{2}{4}\right) \left(\frac{1}{3}\right) \\ &= \frac{2}{n(n-1)} \\ &\geq \frac{2}{n^2} \end{aligned}$$

Contraction algorithm

Amplification. To amplify the probability of success, run the contraction algorithm many times.

Claim. If we repeat the contraction algorithm $n^2 \ln n$ times, then the probability of failing to find the global min-cut is $\leq 1 / n^2$.

with independent random choices,

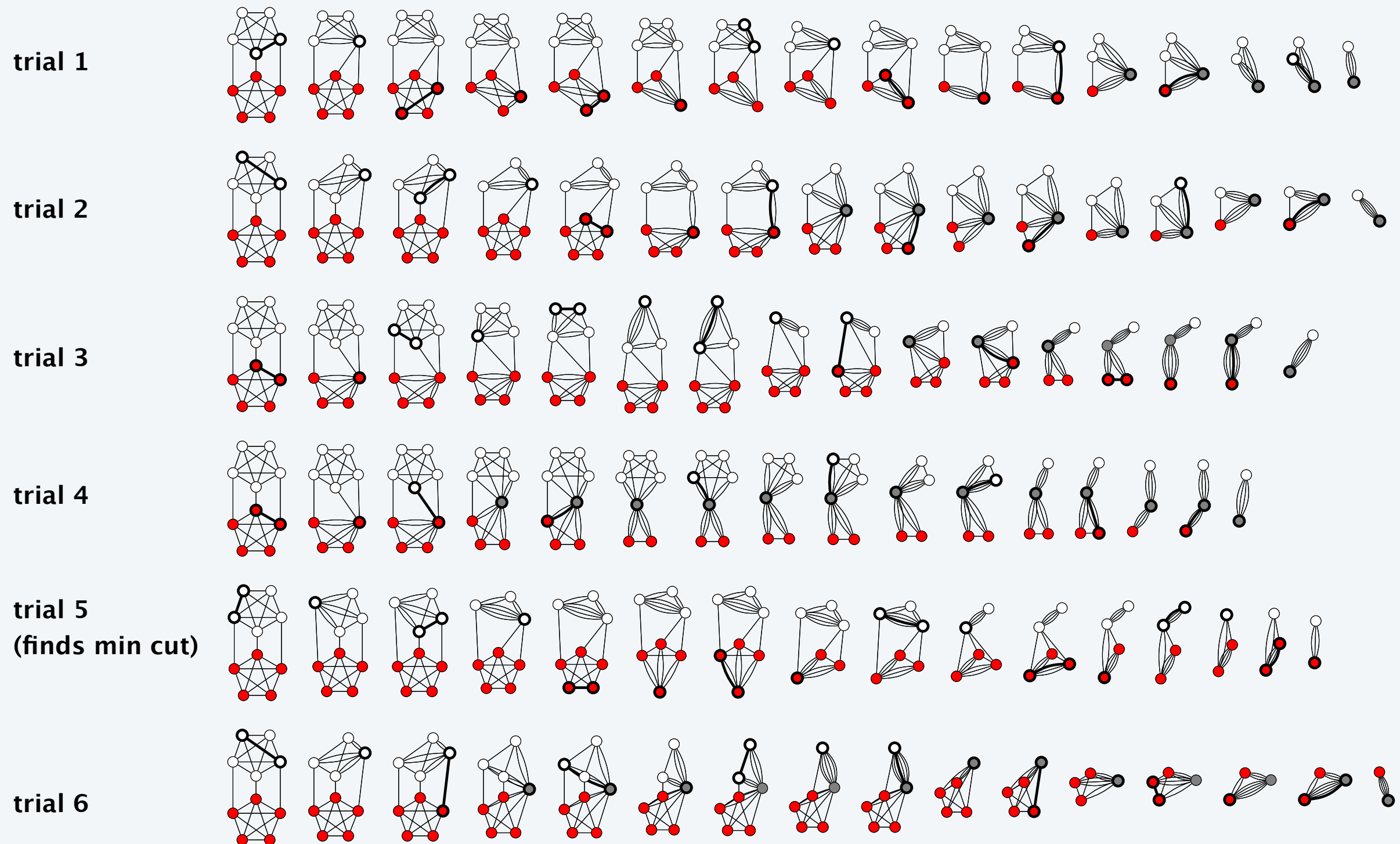


Pf. By independence, the probability of failure is at most

$$\left(1 - \frac{2}{n^2}\right)^{n^2 \ln n} = \left[\left(1 - \frac{2}{n^2}\right)^{\frac{1}{2}n^2}\right]^{2 \ln n} \leq \left(e^{-1}\right)^{2 \ln n} = \frac{1}{n^2}$$

$(1 - 1/x)^x \leq 1/e$

Contraction algorithm: example execution



...

Reference: Thore Husfeldt

Global min cut: context

Remark. Overall running time is slow since we perform $\Theta(n^2 \log n)$ iterations and each takes $\Omega(m)$ time.

Improvement. [Karger-Stein 1996] $O(n^2 \log^3 n)$.

- Early iterations are less risky than later ones: probability of contracting an edge in min cut hits 50% when $n / \sqrt{2}$ nodes remain.
- Run contraction algorithm until $n / \sqrt{2}$ nodes remain.
- Run contraction algorithm **twice** on resulting graph and return **best** of two cuts.

Extensions. Naturally generalizes to handle positive weights.

Best known. [Karger 2000] $O(m \log^3 n)$.



faster than best known max flow algorithm or
deterministic global min cut algorithm



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Expectation

Expectation. Given a discrete random variables X , its expectation $E[X]$ is defined by:

$$E[X] = \sum_{j=0}^{\infty} j \Pr[X = j]$$

Waiting for a first success. Coin is heads with probability p and tails with probability $1-p$. How many independent flips X until first heads?

$$E[X] = \sum_{j=0}^{\infty} j \cdot \Pr[X = j] = \sum_{j=0}^{\infty} j \underset{\substack{\uparrow \\ j-1 \text{ tails}}}{(1-p)^{j-1}} \underset{\substack{\uparrow \\ 1 \text{ head}}}{p} = \frac{p}{1-p} \sum_{j=0}^{\infty} j (1-p)^j = \frac{p}{1-p} \cdot \frac{1-p}{p^2} = \frac{1}{p}$$

Expectation: two properties

Useful property. If X is a 0/1 random variable, $E[X] = \Pr[X = 1]$.

Pf.
$$E[X] = \sum_{j=0}^{\infty} j \cdot \Pr[X = j] = \sum_{j=0}^1 j \cdot \Pr[X = j] = \Pr[X = 1]$$

not necessarily independent



Linearity of expectation. Given two random variables X and Y defined over the same probability space, $E[X + Y] = E[X] + E[Y]$.

Benefit. Decouples a complex calculation into simpler pieces.

Guessing cards

Game. Shuffle a deck of n cards; turn them over one at a time; try to guess each card.

Memoryless guessing. No psychic abilities; can't even remember what's been turned over already. Guess a card from full deck uniformly at random.

Claim. The expected number of correct guesses is 1.

Pf. [surprisingly effortless using linearity of expectation]

- Let $X_i = 1$ if i^{th} prediction is correct and 0 otherwise.
- Let $X =$ number of correct guesses $= X_1 + \dots + X_n$.
- $E[X_i] = \Pr[X_i = 1] = 1 / n$.
- $E[X] = E[X_1] + \dots + E[X_n] = 1 / n + \dots + 1 / n = 1. \quad \blacksquare$



linearity of expectation

Guessing cards

Game. Shuffle a deck of n cards; turn them over one at a time; try to guess each card.

Guessing with memory. Guess a card uniformly at random from cards not yet seen.

Claim. The expected number of correct guesses is $\Theta(\log n)$.

Pf.

- Let $X_i = 1$ if i^{th} prediction is correct and 0 otherwise.
- Let $X =$ number of correct guesses $= X_1 + \dots + X_n$.
- $E[X_i] = \Pr[X_i = 1] = 1 / (n - i + 1)$.
- $E[X] = E[X_1] + \dots + E[X_n] = 1/n + \dots + 1/2 + 1/1 = H(n)$. ■

↑
linearity of expectation

↑
 $\ln(n+1) < H(n) < 1 + \ln n$

Coupon collector

Coupon collector. Each box of cereal contains a coupon. There are n different types of coupons. Assuming all boxes are equally likely to contain each coupon, how many boxes before you have ≥ 1 coupon of each type?

Claim. The expected number of steps is $\Theta(n \log n)$.

Pf.

- Phase j = time between j and $j + 1$ distinct coupons.
- Let X_j = number of steps you spend in phase j .
- Let X = number of steps in total = $X_0 + X_1 + \dots + X_{n-1}$.

$$E[X] = \sum_{j=0}^{n-1} E[X_j] = \sum_{j=0}^{n-1} \frac{n}{n-j} = n \sum_{i=1}^n \frac{1}{i} = n H(n)$$



prob of success = $(n - j) / n$
 \Rightarrow expected waiting time = $n / (n - j)$



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Maximum 3-satisfiability

exactly 3 distinct literals per clause

Maximum 3-satisfiability. Given a 3-SAT formula, find a truth assignment that satisfies as many clauses as possible.

$$\begin{aligned}C_1 &= x_2 \vee \overline{x_3} \vee \overline{x_4} \\C_2 &= x_2 \vee x_3 \vee \overline{x_4} \\C_3 &= \overline{x_1} \vee x_2 \vee x_4 \\C_4 &= \overline{x_1} \vee \overline{x_2} \vee x_3 \\C_5 &= x_1 \vee \overline{x_2} \vee \overline{x_4}\end{aligned}$$

Remark. NP-hard search problem.

Simple idea. Flip a coin, and set each variable true with probability $\frac{1}{2}$, independently for each variable.

Maximum 3-satisfiability: analysis

Claim. Given a 3-SAT formula with k clauses, the expected number of clauses satisfied by a random assignment is $7k / 8$.

Pf. Consider random variable $Z_j = \begin{cases} 1 & \text{if clause } C_j \text{ is satisfied} \\ 0 & \text{otherwise.} \end{cases}$

- Let Z = number of clauses satisfied by random assignment.

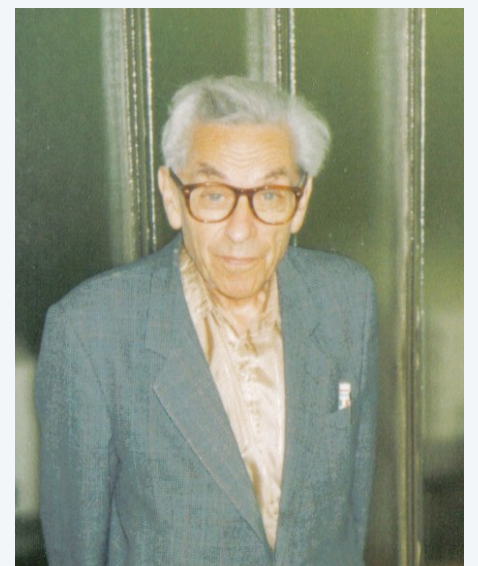
$$\begin{aligned} E[Z] &= \sum_{j=1}^k E[Z_j] \\ \text{linearity of expectation} &\nearrow \\ &= \sum_{j=1}^k \Pr[\text{clause } C_j \text{ is satisfied}] \\ &= \frac{7}{8}k \end{aligned}$$

The Probabilistic Method

Corollary. For any instance of 3-SAT, there exists a truth assignment that satisfies at least a $7/8$ fraction of all clauses.

Pf. Random variable is at least its expectation some of the time. ■

Probabilistic method. [Paul Erdős] Prove the existence of a non-obvious property by showing that a random construction produces it with positive probability!



Maximum 3-satisfiability: analysis

Q. Can we turn this idea into a $7/8$ -approximation algorithm?

A. Yes (but a random variable can almost always be below its mean).

Lemma. The probability that a random assignment satisfies $\geq 7k/8$ clauses is at least $1/(8k)$.

Pf. Let p_j be probability that exactly j clauses are satisfied;
let p be probability that $\geq 7k/8$ clauses are satisfied.

$$\begin{aligned}\frac{7}{8}k &= E[Z] = \sum_{j \geq 0} j p_j \\ &= \sum_{j < 7k/8} j p_j + \sum_{j \geq 7k/8} j p_j \\ &\leq \left(\frac{7k}{8} - \frac{1}{8}\right) \sum_{j < 7k/8} p_j + k \sum_{j \geq 7k/8} p_j \\ &\leq \left(\frac{7}{8}k - \frac{1}{8}\right) \cdot 1 + k p\end{aligned}$$

Rearranging terms yields $p \geq 1/(8k)$. ■

Maximum 3-satisfiability: analysis

Johnson's algorithm. Repeatedly generate random truth assignments until one of them satisfies $\geq 7k / 8$ clauses.

Theorem. Johnson's algorithm is a $7/8$ -approximation algorithm.

Pf. By previous lemma, each iteration succeeds with probability $\geq 1 / (8k)$. By the waiting-time bound, the expected number of trials to find the satisfying assignment is at most $8k$. ■

Maximum satisfiability

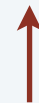
Extensions.

- Allow one, two, or more literals per clause.
- Find max **weighted** set of satisfied clauses.

Theorem. [Asano-Williamson 2000] There exists a 0.784-approximation algorithm for 3-SAT.

Theorem. [Karloff-Zwick 1997, Zwick+computer 2002] There exists a $7/8$ -approximation algorithm for version of MAX-3-SAT where each clause has **at most** 3 literals.

Theorem. [Håstad 1997] Unless $\mathbf{P} = \mathbf{NP}$, no ρ -approximation algorithm for MAX-3-SAT (and hence MAX-SAT) for any $\rho > 7/8$.



very unlikely to improve over simple randomized
algorithm for MAX-3SAT

Monte Carlo vs. Las Vegas algorithms

Monte Carlo. Guaranteed to run in poly-time, likely to find correct answer.

Ex: Contraction algorithm for global min cut.

Las Vegas. Guaranteed to find correct answer, likely to run in poly-time.

Ex: Randomized quicksort, Johnson's MAX-3-SAT algorithm.

stop algorithm
after a certain point

Remark. Can always convert a Las Vegas algorithm into Monte Carlo,
but no known method (in general) to convert the other way.

RP and ZPP

RP. [Monte Carlo] Decision problems solvable with **one-sided error** in poly-time.

can decrease probability of false negative
to 2^{-100} by 100 independent repetitions

One-sided error.

- If the correct answer is *no*, always return *no*.
- If the correct answer is *yes*, return *yes* with probability $\geq \frac{1}{2}$.

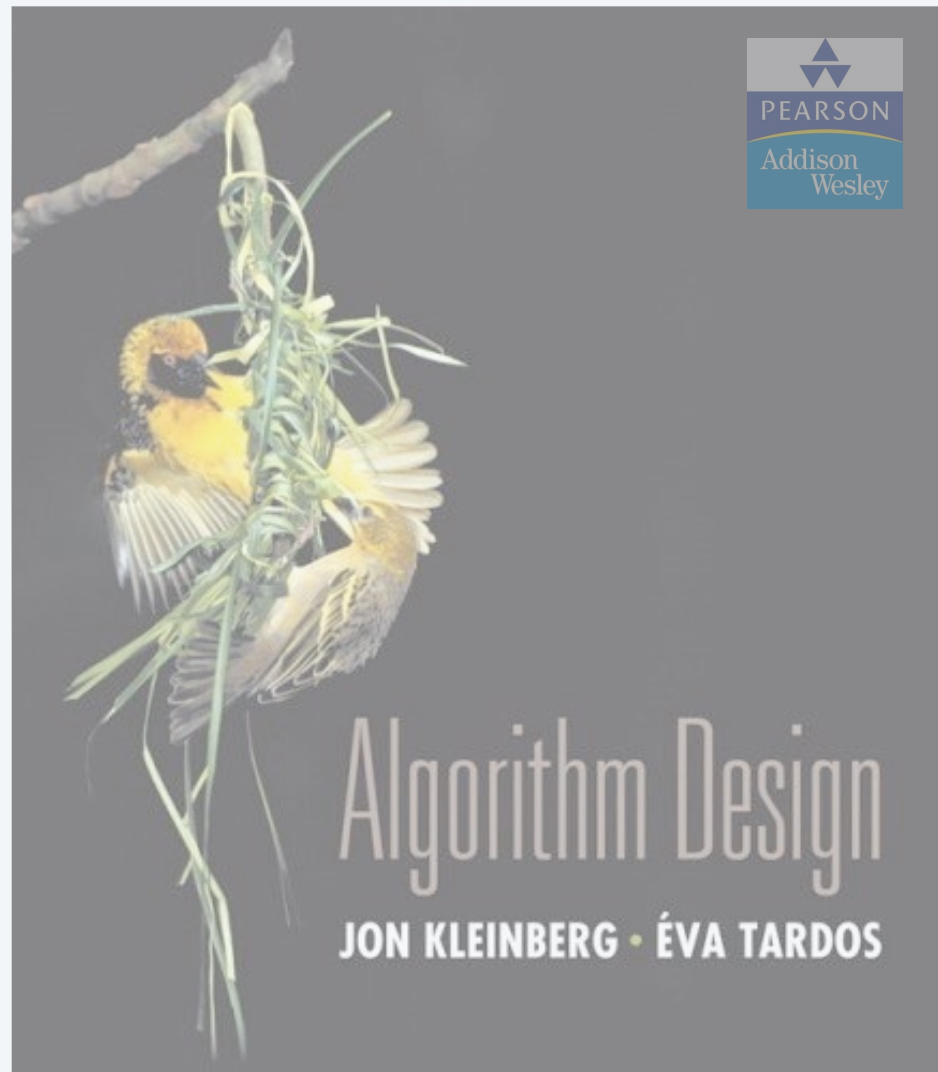
ZPP. [Las Vegas] Decision problems solvable in **expected** poly-time.

running time can be unbounded,
but fast on average

Theorem. $\mathbf{P} \subseteq \mathbf{ZPP} \subseteq \mathbf{RP} \subseteq \mathbf{NP}$.

Fundamental open questions. To what extent does randomization help?

Does $\mathbf{P} = \mathbf{ZPP}$? Does $\mathbf{ZPP} = \mathbf{RP}$? Does $\mathbf{RP} = \mathbf{NP}$?



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Dictionary data type

Dictionary. Given a universe U of possible elements, maintain a subset $S \subseteq U$ so that **inserting**, deleting, and **searching** in S is efficient.

Dictionary interface.

- `create()`: initialize a dictionary with $S = \phi$.
- `insert(u)`: add element $u \in U$ to S .
- `delete(u)`: delete u from S (if u is currently in S).
- `lookup(u)`: is u in S ?

Challenge. Universe U can be extremely large so defining an array of size $|U|$ is infeasible.

Applications. File systems, databases, Google, compilers, checksums P2P networks, associative arrays, cryptography, web caching, etc.

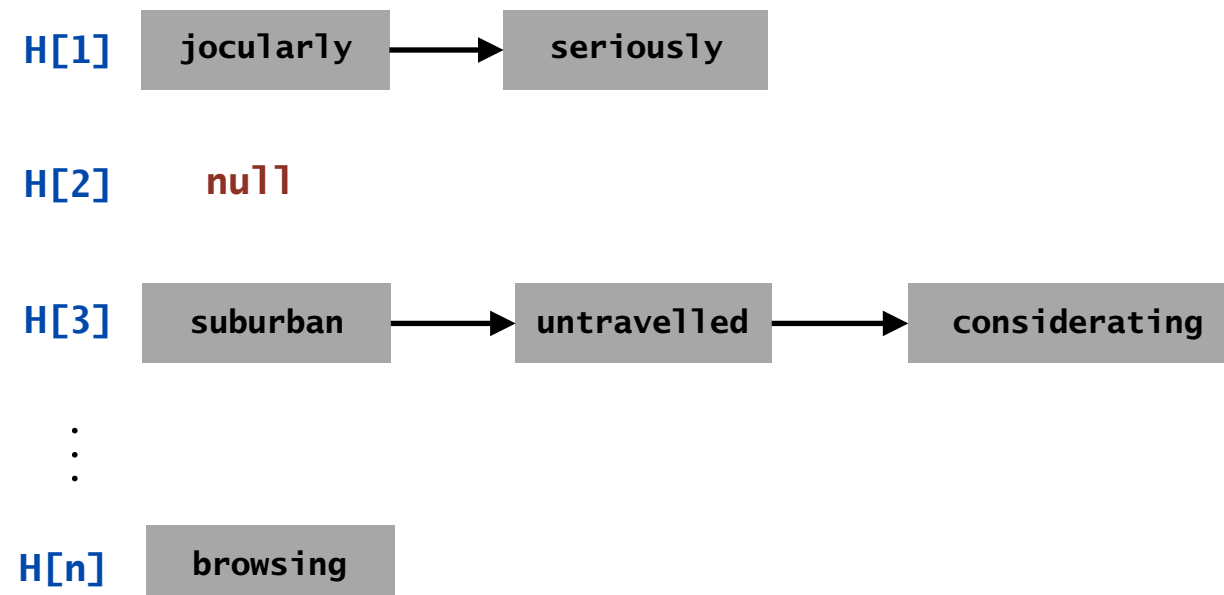
Hashing

Hash function. $h : U \rightarrow \{ 0, 1, \dots, n - 1 \}$.

Hashing. Create an array H of size n . When processing element u , access array element $H[h(u)]$.

- Collision.** When $h(u) = h(v)$ but $u \neq v$.
- A collision is expected after $\Theta(\sqrt{n})$ random insertions.
 - Separate chaining: $H[i]$ stores linked list of elements u with $h(u) = i$.

birthday paradox



Ad-hoc hash function

Ad hoc hash function.

```
int hash(String s, int n) {  
    int hash = 0;  
    for (int i = 0; i < s.length(); i++)  
        hash = (31 * hash) + s[i];  
    return hash % n;  
}
```

hash function ala Java string library


Deterministic hashing. If $|U| \geq n^2$, then for any fixed hash function h , there is a subset $S \subseteq U$ of n elements that all hash to same slot. Thus, $\Theta(n)$ time per search in worst-case.

Q. But isn't ad-hoc hash function good enough in practice?

Algorithmic complexity attacks

When can't we live with ad hoc hash function?

- Obvious situations: aircraft control, nuclear reactors.
- Surprising situations: denial-of-service attacks.



malicious adversary learns your ad hoc hash function
(e.g., by reading Java API) and causes a big pile-up in a
single slot that grinds performance to a halt

Real world exploits. [Crosby-Wallach 2003]

- Bro server: send carefully chosen packets to DOS the server, using less bandwidth than a dial-up modem
- Perl 5.8.0: insert carefully chosen strings into associative array.
- Linux 2.4.20 kernel: save files with carefully chosen names.


Hashing performance

Ideal hash function. Maps m elements **uniformly at random** to m hash slots.

- Running time depends on length of chains.
- Average length of chain = $\alpha = m / n$.
- Choose $n \approx m \Rightarrow$ on average $O(1)$ per insert, lookup, or delete.

Challenge. Achieve idealized randomized guarantees, but with a hash function where you can easily find items where you put them.

Approach. Use **randomization** in the choice of h .

 adversary knows the randomized algorithm you're using,
but doesn't know random choices that the algorithm makes

Universal hashing

Universal family of hash functions. [Carter-Wegman 1980s]

- For any pair of elements $u, v \in U$, $\Pr_{h \in H} [h(u) = h(v)] \leq 1/n$
- Can select random h efficiently.
- Can compute $h(u)$ efficiently.

 chosen uniformly at random

Ex. $U = \{a, b, c, d, e, f\}$, $n = 2$.

	a	b	c	d	e	f
$h_1(x)$	0	1	0	1	0	1
$h_2(x)$	0	0	0	1	1	1

	a	b	c	d	e	f
$h_1(x)$	0	1	0	1	0	1
$h_2(x)$	0	0	0	1	1	1
$h_3(x)$	0	0	1	0	1	1
$h_4(x)$	1	0	0	1	1	0

$$H = \{h_1, h_2\}$$

$$\Pr_{h \in H} [h(a) = h(b)] = 1/2$$

$$\Pr_{h \in H} [h(a) = h(c)] = 1$$

$$\Pr_{h \in H} [h(a) = h(d)] = 0$$

...

$$H = \{h_1, h_2, h_3, h_4\}$$

$$\Pr_{h \in H} [h(a) = h(b)] = 1/2$$

$$\Pr_{h \in H} [h(a) = h(c)] = 1/2$$

$$\Pr_{h \in H} [h(a) = h(d)] = 1/2$$

$$\Pr_{h \in H} [h(a) = h(e)] = 1/2$$

$$\Pr_{h \in H} [h(a) = h(f)] = 0$$

...

not universal


universal

Universal hashing: analysis

Proposition. Let H be a universal family of hash functions; let $h \in H$ be chosen uniformly at random from H ; and let $u \in U$. For any subset $S \subseteq U$ of size at most n , the expected number of items in S that collide with u is at most 1.

Pf. For any element $s \in S$, define indicator random variable $X_s = 1$ if $h(s) = h(u)$ and 0 otherwise. Let X be a random variable counting the total number of collisions with u .


$$E_{h \in H}[X] = E[\sum_{s \in S} X_s] = \sum_{s \in S} E[X_s] = \sum_{s \in S} \Pr[X_s = 1] \leq \sum_{s \in S} \frac{1}{n} = |S| \frac{1}{n} \leq 1$$


linearity of expectation X_s is a 0-1 random variable universal
(assumes $u \notin S$)

Q. OK, but how do we design a universal class of hash functions?

Designing a universal family of hash functions

Theorem. [Chebyshev 1850] There exists a prime between n and $2n$.

Modulus. Choose a prime number $p \approx n$.  no need for randomness here

Integer encoding. Identify each element $u \in U$ with a base- p integer of r digits: $x = (x_1, x_2, \dots, x_r)$.

Hash function. Let A = set of all r -digit, base- p integers. For each $a = (a_1, a_2, \dots, a_r)$ where $0 \leq a_i < p$, define

$$h_a(x) = \left(\sum_{i=1}^r a_i x_i \right) \bmod p$$

Hash function family. $H = \{ h_a : a \in A \}$.

Designing a universal family of hash functions


Theorem. $H = \{ h_a : a \in A \}$ is a universal family of hash functions.

Pf. Let $x = (x_1, x_2, \dots, x_r)$ and $y = (y_1, y_2, \dots, y_r)$ be two distinct elements of U .

We need to show that $\Pr[h_a(x) = h_a(y)] \leq 1 / n$.

- Since $x \neq y$, there exists an integer j such that $x_j \neq y_j$.
- We have $h_a(x) = h_a(y)$ iff

$$a_j \underbrace{(y_j - x_j)}_z = \underbrace{\sum_{i \neq j} a_i (x_i - y_i)}_m \pmod p$$

- Can assume a was chosen uniformly at random by first selecting all coordinates a_i where $i \neq j$, then selecting a_j at random. Thus, we can assume a_i is fixed for all coordinates $i \neq j$.
- Since p is prime, $a_j z = m \pmod p$ has at most one solution among p possibilities.  see lemma on next slide
- Thus $\Pr[h_a(x) = h_a(y)] = 1 / p \leq 1 / n$. ■

Number theory fact

Fact. Let p be prime, and let $z \not\equiv 0 \pmod{p}$. Then $\alpha z \equiv m \pmod{p}$ has at most one solution $0 \leq \alpha < p$.

Pf.

- Suppose α and β are two different solutions.
- Then $(\alpha - \beta)z \equiv 0 \pmod{p}$; hence $(\alpha - \beta)z$ is divisible by p .
- Since $z \not\equiv 0 \pmod{p}$, we know that z is not divisible by p ; it follows that $(\alpha - \beta)$ is divisible by p .
- This implies $\alpha = \beta$. ■

Bonus fact. Can replace "at most one" with "exactly one" in above fact.

Pf idea. Euclid's algorithm.



13. RANDOMIZED ALGORITHMS

- ▶ *contention resolution*
- ▶ *global min cut*
- ▶ *linearity of expectation*
- ▶ *max 3-satisfiability*
- ▶ *universal hashing*
- ▶ **Chernoff bounds**
- ▶ *load balancing*

Chernoff Bounds (above mean)

Theorem. Suppose X_1, \dots, X_n are independent 0-1 random variables. Let $X = X_1 + \dots + X_n$. Then for any $\mu \geq E[X]$ and for any $\delta > 0$, we have

$$\Pr[X > (1 + \delta)\mu] < \left[\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right]^\mu$$

↑
sum of independent 0-1 random variables
is tightly centered on the mean

Pf. We apply a number of simple transformations.

- For any $t > 0$,

$$\Pr[X > (1 + \delta)\mu] = \Pr[e^{tX} > e^{t(1+\delta)\mu}] \leq e^{-t(1+\delta)\mu} \cdot E[e^{tX}]$$

↑
 $f(x) = e^{tx}$ is monotone in x

↑
Markov's inequality: $\Pr[X > a] \leq E[X] / a$

- Now $E[e^{tX}] = E[e^{t \sum_i X_i}] = \prod_i E[e^{tX_i}]$

↑
definition of X

↑
independence

Chernoff Bounds (above mean)

Pf. [continued]

- Let $p_i = \Pr [X_i = 1]$. Then,

$$E[e^{tX_i}] = p_i e^t + (1 - p_i) e^0 = 1 + p_i(e^t - 1) \leq e^{p_i(e^t - 1)}$$

↑
for any $\alpha \geq 0$, $1 + \alpha \leq e^\alpha$

- Combining everything:

$$\Pr[X > (1 + \delta)\mu] \leq e^{-t(1+\delta)\mu} \prod_i E[e^{tX_i}] \leq e^{-t(1+\delta)\mu} \prod_i e^{p_i(e^t - 1)} \leq e^{-t(1+\delta)\mu} e^{\mu(e^t - 1)}$$

↑
previous slide

↑
inequality above

↑
 $\sum_i p_i = E[X] \leq \mu$

- Finally, choose $t = \ln(1 + \delta)$. ■

Chernoff Bounds (below mean)

Theorem. Suppose X_1, \dots, X_n are independent 0-1 random variables.

Let $X = X_1 + \dots + X_n$. Then for any $\mu \leq E[X]$ and for any $0 < \delta < 1$, we have

$$\Pr[X < (1 - \delta)\mu] < e^{-\delta^2 \mu / 2}$$

Pf idea. Similar.

Remark. Not quite symmetric since only makes sense to consider $\delta < 1$.



13. RANDOMIZED ALGORITHMS

- ▶ *contention resolution*
- ▶ *global min cut*
- ▶ *linearity of expectation*
- ▶ *max 3-satisfiability*
- ▶ *universal hashing*
- ▶ *Chernoff bounds*
- ▶ ***load balancing***

Load Balancing

Load balancing. System in which m jobs arrive in a stream and need to be processed immediately on m identical processors. Find an assignment that balances the workload across processors.

Centralized controller. Assign jobs in round-robin manner. Each processor receives at most $\lceil m / n \rceil$ jobs.

Decentralized controller. Assign jobs to processors uniformly at random. How likely is it that some processor is assigned "too many" jobs?


Load balancing

Analysis.

- Let X_i = number of jobs assigned to processor i .
- Let $Y_{ij} = 1$ if job j assigned to processor i , and 0 otherwise.
- We have $E[Y_{ij}] = 1/n$.
- Thus, $X_i = \sum_j Y_{ij}$, and $\mu = E[X_i] = 1$.
- Applying Chernoff bounds with $\delta = c - 1$ yields $\Pr[X_i > c] < \frac{e^{c-1}}{c^c}$
- Let $\gamma(n)$ be number x such that $x^x = n$, and choose $c = e \gamma(n)$.

$$\Pr[X_i > c] < \frac{e^{c-1}}{c^c} < \left(\frac{e}{c}\right)^c = \left(\frac{1}{\gamma(n)}\right)^{e\gamma(n)} < \left(\frac{1}{\gamma(n)}\right)^{2\gamma(n)} = \frac{1}{n^2}$$

- Union bound \Rightarrow with probability $\geq 1 - 1/n$ no processor receives more than $e \gamma(n) = \Theta(\log n / \log \log n)$ jobs.

 Bonus fact: with high probability,
some processor receives $\Theta(\log n / \log \log n)$ jobs

Load balancing: many jobs

Theorem. Suppose the number of jobs $m = 16 n \ln n$. Then on average, each of the n processors handles $\mu = 16 \ln n$ jobs. With high probability, every processor will have between half and twice the average load.

Pf.

- Let X_i, Y_{ij} be as before.
- Applying Chernoff bounds with $\delta = 1$ yields

$$\Pr[X_i > 2\mu] < \left(\frac{e}{4}\right)^{16n \ln n} < \left(\frac{1}{e}\right)^{\ln n} = \frac{1}{n^2} \qquad \Pr[X_i < \frac{1}{2}\mu] < e^{-\frac{1}{2}\left(\frac{1}{2}\right)^2(16n \ln n)} = \frac{1}{n^2}$$

- Union bound \Rightarrow every processor has load between half and twice the average with probability $\geq 1 - 2/n$. ■