

LDA

$$X = \{x_1, \dots, x_d\}$$

$$Y = \{0, 1\}$$

$$\textcircled{1} \quad P[Y=0] = P[Y=1] = \frac{1}{2}$$

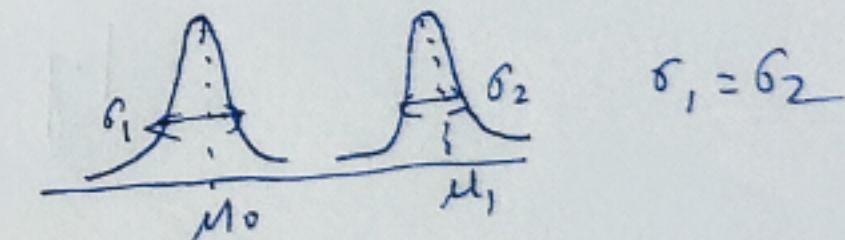
\textcircled{2} $P[X=x|Y=y]$ is a Gaussian Distribution

$$P[X=x|Y=y] = \frac{1}{(2\pi)^{d/2} |\Sigma_y|^{1/2}} e^{-\frac{1}{2} (x-\mu_y)^T \Sigma_y^{-1} (x-\mu_y)}$$

$$y=0, 1$$

\textcircled{3} The covariance matrixes of the Gaussian distributions are the same

$$\Sigma_0 = \Sigma_1$$



$$h_{\text{Bayes}}(x) = \underset{y \in \{0, 1\}}{\operatorname{argmax}} P[Y=y | X=x] \quad P[X=0 | X=x] \quad P[X=1 | X=x]$$

$$= \underset{y \in \{0, 1\}}{\operatorname{argmax}} \frac{P[Y=y] P[X=x | Y=y]}{P[X=x]} \quad 0.2 \quad 0.8$$

$$= \underset{y \in \{0, 1\}}{\operatorname{argmax}} P[X=x | Y=y]$$

$$\max_{y \in \{0, 1\}} P[x = x | y = y]$$

$y = 1$ if

$$P[x = x | y = 1] > P[x = x | y = 0]$$

$y = 0$ if $P[x = x | y = 1] < P[x = x | y = 0]$

$$\Rightarrow P[x = x | y = 1] > P[x = x | y = 0]$$

$$\frac{P[x = x | y = 1]}{P[x = x | y = 0]} > 1$$

$$\log \left[\frac{P[x = x | y = 1]}{P[x = x | y = 0]} \right] > 0$$

$$\log \left[\frac{\frac{(2\pi)^{d/2} |\Sigma|^{1/2}}{e^{-\frac{1}{2}(x - \mu_1)^T \Sigma^{-1} (x - \mu_1)}}}{\frac{(2\pi)^{d/2} |\Sigma|^{1/2}}{e^{-\frac{1}{2}(x - \mu_0)^T \Sigma^{-1} (x - \mu_0)}}} \right] > 0$$

$$\log \left(e^{-\frac{1}{2}(x - \mu_1)^T \Sigma^{-1} (x - \mu_1)} \right) - \log \left[e^{-\frac{1}{2}(x - \mu_0)^T \Sigma^{-1} (x - \mu_0)} \right] > 0$$

$$-(x - \mu_1)^T \Sigma^{-1} (x - \mu_1) + (x - \mu_0)^T \Sigma^{-1} (x - \mu_0) > 0$$

$$(x - \mu_0)^T \Sigma^{-1} (x - \mu_0) - (x - \mu_1)^T \Sigma^{-1} (x - \mu_1) > 0$$

$$(x^T - \mu_0^T) \Sigma^{-1} (x - \mu_0) - (x^T - \mu_1^T) \Sigma^{-1} (x - \mu_1) > 0$$

~~$$x^T \cancel{\Sigma^{-1}} x - x^T \Sigma^{-1} \mu_0 - \mu_0^T \Sigma^{-1} x + \mu_0^T \Sigma^{-1} \mu_0$$~~

~~$$- x^T \cancel{\Sigma^{-1}} x + x^T \Sigma^{-1} \mu_1 + \mu_1^T \Sigma^{-1} x - \mu_1^T \Sigma^{-1} \mu_1 > 0$$~~

~~$$-2 \cancel{x^T} - 2 \mu_0^T \Sigma^{-1} x + 2 \mu_1^T \Sigma^{-1} x + \mu_0^T \Sigma^{-1} \mu_0 - \mu_1^T \Sigma^{-1} \mu_1 > 0$$~~

$$2(\mu_1^T - \mu_0^T) \Sigma^{-1} x + \mu_0^T \Sigma^{-1} \mu_0 - \mu_1^T \Sigma^{-1} \mu_1 > 0$$

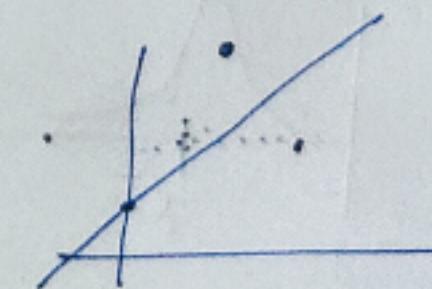
$$\omega^T x + b > 0$$

$$\omega = \cancel{\mu_0^T} \Sigma^{-1} (\mu_1 - \mu_0)$$

$$b = \mu_0^T \Sigma^{-1} \mu_0 - \mu_1^T \Sigma^{-1} \mu_1$$

$$\checkmark \omega^T x + b \geq 0 \Rightarrow y = 1$$

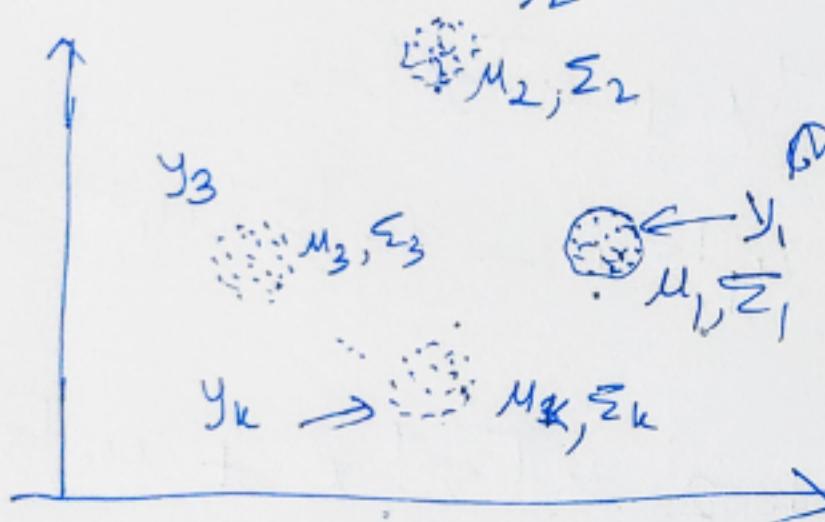
$$\omega^T x + b < 0 \Rightarrow y = 0$$



Expectation Maximization Algorithm.

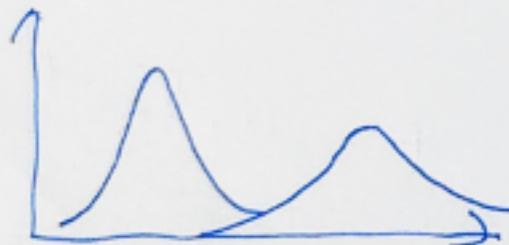
$$S = \{x_1, \dots, x_n\}$$

$$\max_{\theta} \ell(S, \theta) = \log (P_{\theta}(x=x))$$



Mixture of Gaussian Distributions

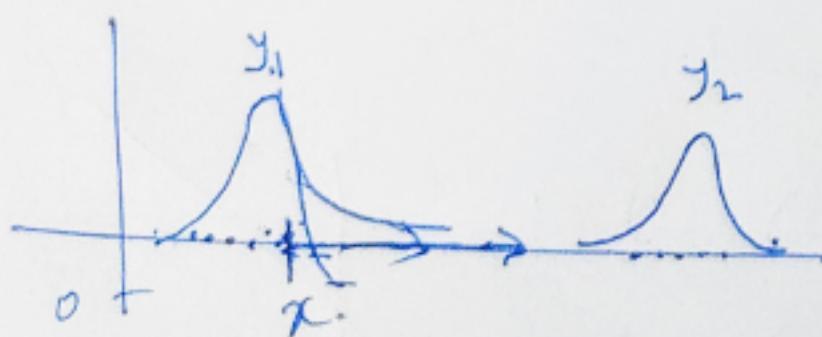
$$\Theta = (\mu_1, \mu_2, \dots, \mu_k, \Sigma_1, \Sigma_2, \dots, \Sigma_k)$$



Consider each Gaussian distribution separately. Let us consider k random variables y_1, y_2, \dots, y_k (Latent Variables)

$$P[Y=y] = c_y \quad y \in \{1, 2, \dots, k\}$$

$$\sum_{y=1}^k c_y = 1$$



We choose x on the basis of the value y according to a Gaussian distribution

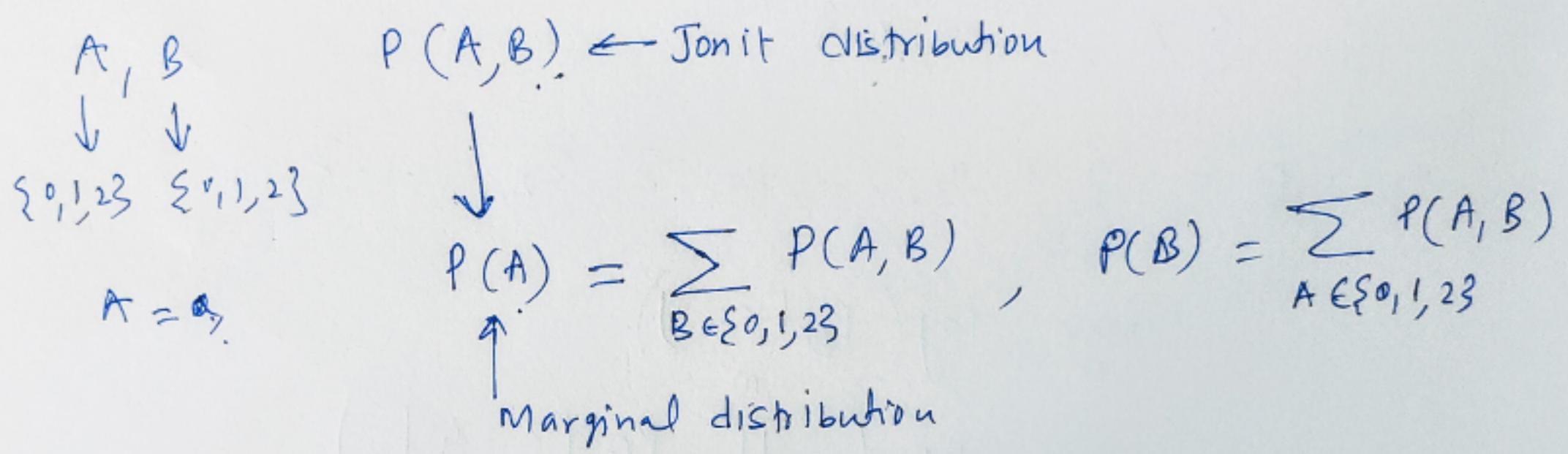
$$P[X=x | Y=y] = \frac{1}{(2\pi)^{d/2} |\Sigma_y|^{1/2}} e^{-\frac{1}{2} (x - \mu_y)^T \Sigma_y^{-1} (x - \mu_y)}$$

$$P[X=x | Y=1] >$$

$$P[X=x | Y=2]$$

$$\max L(s; \theta) = \max_{\theta} \frac{P_\theta[x=x]}{\cdot}$$

Marginal Distribution



②

$P_\theta(x, y)$

$$P_\theta(x) = \sum_{y=1}^k P_\theta(x=x, Y=y)$$

$$P_\theta[x=x] = \sum_{y=1}^k P_\theta[x=x, Y=y]$$

$$P[x=x, Y=y] = P[Y=y] P[x=x | Y=y].$$

$$P_\theta[x=x] = \sum_{y=1}^k \underbrace{P[Y=y]}_{\cdot} P[x=x | Y=y] \quad \text{--- } ①$$

$$P_{\theta}[X = x] = \sum_{y=1}^k c_y \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} (x - \mu_y)^T \Sigma_y^{-1} (x - \mu_y)} \quad \textcircled{D}$$

$$S = \{x_1, \dots, x_m\}$$

$$\Theta = \{\mu_1, \mu_2, \dots, \mu_k, \Sigma_1, \Sigma_2, \dots, \Sigma_k\}$$

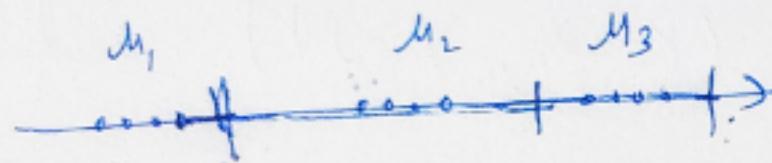
$$\max_{\theta} \log \left(\prod_m P_{\theta}[x = x^m] \right)$$

$$\max_{\theta} \log \left(\prod_{i=1}^m P_{\theta}[x = x_i] \right)$$

$$\max_{\theta} \sum_{i=1}^m \log \left(P_{\theta}[x = x_i] \right)$$

from (D) $\Rightarrow \max_{\theta} \sum_{i=1}^m \log \left(\sum_{y=1}^k P[y=y] P[x=x_i | y=y] \right)$

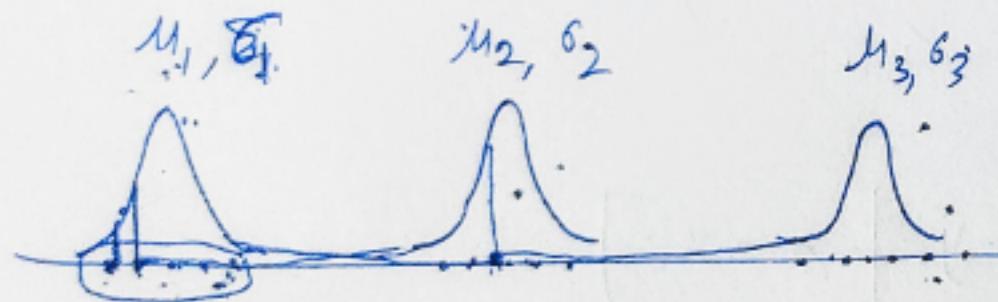
$$\max_{\theta} \sum_{i=1}^m \log \left[\sum_{y=1}^k P_{\theta}[x = x_i, y = y] \right]$$



$$\sum_{i=1}^k \sum_{x \in G_i} \|x - \mu_i\|_2^2$$

$$\sum_{i=1}^k \sum_{j=1}^m c_{j,i} \|x - \mu_i\|_2^2$$

$c_{j,i} = 1$ if j^{th} point belongs
to the i^{th} cluster
 $= 0$



$$y_i \sim G(\mu_i, \sigma_i) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu_i)^2}{2}}$$

- If we know the source distribution of all the points then we can easily determine the parameter
- If we know the parameters for all the distributions then we can easily determine the source distribution for each point
 - (1, 1)
 - (3, 1)
 - (5, 1)

$$F(Q, \theta) = \sum_{i=1}^m \sum_{y=1}^k Q_{i,y} \log (P_\theta(x=x_i, Y=y))$$

$$Q_{i,y} = P_\theta(Y=y | x=x_i)$$

$$\left(\begin{array}{c} \cancel{\theta^{(t)}} \\ \theta^{(0)} \end{array} \right), \quad \left(\begin{array}{c} \theta^{(1)}, Q^{(1)} \end{array} \right), \quad \left(\begin{array}{c} \theta^{(2)}, Q^{(2)} \end{array} \right), \quad \dots$$

$$\begin{aligned} Q^{(t+1)} &= P_{\theta^{(t)}}[Y=y | x=x_i] \\ &= \frac{P_{\theta^{(t)}}[x=x_i | y=y]}{P[x=x_i]} \\ &= \frac{c_y^{(t)} e^{-\frac{1}{2} \|x_i - y\|^2}}{z_i} \quad \sum = 1 \end{aligned}$$

$$\theta^{(t+1)} = \max_{\theta} F(\phi^{(t+1)}, \theta)$$

$$= \max_{\theta} \sum_{i=1}^m \sum_{y=1}^k P_{\theta^{(t)}}[Y=y \mid X=x_i] \left[\log(c_y) - \frac{1}{2} \|x_i - u_y\|^2 \right]$$

$$\frac{\partial}{\partial \theta_y} F(\phi^{(t+1)}, \theta) = 0$$

\downarrow
 $y = 1, 2, \dots, k$

$$u_y = \frac{\sum_{i=1}^m P_{\theta^{(t)}}[Y=y \mid X=x_i] x_i}{\sum_{i=1}^m P_{\theta^{(t)}}[Y=y \mid X=x_i]}$$

