

# Homework Assignment 05

Numerical Statistics Fall, 2022

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## 1.

Let  $X_1, X_2, \dots, X_n$  be independent random variables distributed according to the Bernoulli distribution with an unknown parameter  $p$  ( $0 < p < 1$ ), i.e., suppose that  $X_i$  's are independent,  $P(X_i = 1) = p$  and  $P(X_i = 0) = 1 - p$  ( $i = 1, 2, \dots, n$ ). Then the 95% confidence interval for  $p$  is

$$\frac{\bar{X}_n + \frac{(1.96)^2}{2n} \pm (1.96)\sqrt{\frac{\bar{X}_n(1-\bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}}{1 + \frac{(1.96)^2}{n}} \dots (CI_1)$$

derived from the quadratic equation

$$\{n + (1.96)^2\}p^2 - \{2n\bar{X}_n + (1.96)^2\}p + n(\bar{X}_n)^2 = 0 \dots (QE_1)$$

Denote the upper and lower limits of the 95% confidence interval ( $CI_1$ ) by  $\beta$  and  $\alpha$  ( $\alpha \leq \beta$ ), respectively. That is, let  $\beta$  and  $\alpha$  be

$$\frac{\bar{X}_n + \frac{(1.96)^2}{2n} \pm (1.96)\sqrt{\frac{\bar{X}_n(1-\bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}}{1 + \frac{(1.96)^2}{n}}, \text{ respectively.}$$

Since the parameter  $p$  is assumed to lie between 0 and 1, we might expect both  $\beta$  and  $\alpha$  to meet the condition  $0 < \alpha \leq \beta < 1$ .

(1-2) Find a necessary and sufficient condition for  $\beta < 1$ .

### Solutions.

(1-2) The necessary and sufficient condition for  $\beta < 1$  is equivalent to that of  $1 - \beta > 0$ . We show the latter.

$$\begin{aligned} 1 - \beta &= 1 - \frac{\bar{X}_n + \frac{(1.96)^2}{2n} + (1.96)\sqrt{\frac{\bar{X}_n(1-\bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}}{1 + \frac{(1.96)^2}{n}} \\ &= \frac{1 + \frac{(1.96)^2}{n} - \bar{X}_n - \frac{(1.96)^2}{2n} - (1.96)\sqrt{\frac{\bar{X}_n(1-\bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}}{1 + \frac{(1.96)^2}{n}} \\ &= \frac{1 + \frac{(1.96)^2}{2n} - \bar{X}_n - (1.96)\sqrt{\frac{\bar{X}_n(1-\bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}}{1 + \frac{(1.96)^2}{n}} \end{aligned}$$

Since  $n > 0$ ,  $1 + \frac{(1.96)^2}{n} > 0$ . So, to find the necessary and sufficient condition such that the above is

positive, it suffices to show the condition for

$$1 + \frac{(1.96)^2}{2n} - \bar{X}_n - (1.96)\sqrt{\frac{\bar{X}_n(1 - \bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}} > 0 \quad (1)$$

$$\iff 1 + \frac{(1.96)^2}{2n} - \bar{X}_n > (1.96)\sqrt{\frac{\bar{X}_n(1 - \bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}} \quad (2)$$

Because  $0 \leq \bar{X}_n \leq 1$ , Both LHS and RHS of (2) are positive. From here it follows that (2) is satisfied if and only if the LHS squared is greater than the RHS squared.

$$\iff \left(1 - \bar{X}_n + \frac{(1.96)^2}{2n}\right)^2 > \left((1.96)\sqrt{\frac{\bar{X}_n(1 - \bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}}\right)^2 \quad (3)$$

$$\iff (1 - \bar{X}_n)^2 + \frac{(1.96)^2(1 - \bar{X}_n)}{n} + \frac{(1.96)^4}{4n^2} > (1.96)^2 \left(\frac{\bar{X}_n(1 - \bar{X}_n)}{n} + \frac{(1.96)^2}{4n^2}\right) \quad (4)$$

$$\iff (1 - \bar{X}_n)^2 + \frac{(1.96)^2(1 - \bar{X}_n)}{n} + \frac{(1.96)^4}{4n^2} > \frac{(1.96)^2 \bar{X}_n(1 - \bar{X}_n)}{n} + \frac{(1.96)^4}{4n^2} \quad (5)$$

$$\iff (1 - \bar{X}_n)^2 + \frac{1}{n}(1.96)^2(1 - \bar{X}_n)(1 - \bar{X}_n) > 0 \quad (6)$$

$$\iff \underbrace{(1 - \bar{X}_n)^2}_{\geq 0} + \underbrace{\frac{1}{n}(1.96)^2(1 - \bar{X}_n)^2}_{\geq 0} > 0 \quad (7)$$

Recall that  $0 \leq \bar{X}_n \leq 1$ . For the inequality in (7) to be strict, it must be that  $\bar{X}_n \neq 1$ . Thus, the necessary and sufficient condition for the above to be strictly positive is

$$\bar{X}_n < 1$$

□

## 2.

- (2) Let  $X_1$  and  $X_2$  be independent random variables with  $E(X_1) = 3\mu$  and  $E(X_2) = -4\mu$ , where the parameter  $\mu$  is unknown. In addition, let  $Var(X_1) = 2$  and  $Var(X_2) = 5$ . Then answer the questions below.

(2-1) Find a condition on  $w_1$  and  $w_2$  such that  $\hat{\mu}_w = w_1X_1 + w_2X_2$  is an unbiased estimator of  $\mu$ . (2 points)

(2-2) Find the set of numbers  $(w_1, w_2)$  that gives the minimum value of  $Var(\hat{\mu}_w)$  provided  $\hat{\mu}_w$  is an unbiased estimator of  $\mu$ .

### Solutions.

(2-1) For  $\hat{\mu}_w = w_1X_1 + w_2X_2$  to be an unbiased estimator of  $\mu$ , we must have  $E(w_1X_1 + w_2X_2) = \mu$ .

$$\begin{aligned} E(w_1X_1 + w_2X_2) &= E(w_1X_1) + E(w_2X_2) \\ &= w_1E(X_1) + w_2E(X_2) \\ &= w_1(3\mu) + w_2(-4\mu) \\ &= (3w_1 - 4w_2)\mu = \mu \\ \iff \mu &= 0 \text{ or } 3w_1 - 4w_2 = 1 \end{aligned}$$

Thus, the condition for  $w_1, w_2$  is  $3w_1 - 4w_2 = 1 \dots \square$

(2-2) We can state the problem as

$$\underset{w_1, w_2}{\operatorname{argmin}} Var(\hat{\mu}_w) \quad \text{s.t.} \quad E(\hat{\mu}_w) = \mu$$

As we have seen from (2-1), the constraint evaluates to  $3w_1 - 4w_2 = 1$ . The objective function can be simplified thus:

$$Var(\hat{\mu}_w) = Var(w_1 X_1 + w_2 X_2)$$

Since  $X_1 \perp\!\!\!\perp X_2$ ,  $Cov(X_1, X_2) = 0$  and thus,

$$\begin{aligned} Var(w_1 X_1 + w_2 X_2) &= Var(w_1 X_1) + Var(w_2 X_2) \\ &= w_1^2 Var(X_1) + w_2^2 Var(X_2) \\ &= 2w_1^2 + 5w_2^2 \end{aligned}$$

We want to minimize this subject to the constraint  $3w_1 - 4w_2 = 1$ . The Lagrangian of the problem can be formulated as thus:

$$\min_{w_1, w_2, \lambda} \mathcal{L} = 2w_1^2 + 5w_2^2 + \lambda(1 - 3w_1 + 4w_2)$$

FOCS:

$$[w_1] : 4w_1 - 3\lambda = 0 \tag{8}$$

$$[w_2] : 10w_2 + 4\lambda = 0 \tag{9}$$

$$[\lambda] : 1 - 3w_1 + 4w_2 = 0 \tag{10}$$

From (1) and (2),  $w_1 = -\frac{15}{8}w_2$ . Plugging this into (3), we get

$$\begin{aligned} 1 + \frac{45}{8}w_2 + 4w_2 &= 0 \\ \iff \frac{77}{8}w_2 &= -1 \\ \iff w_2 &= -\frac{8}{77} \\ w_1 &= -\frac{15}{8} \cdot \left(-\frac{8}{77}\right) = \frac{15}{77} \end{aligned}$$

Since the objective function is convex, this gives us the global minimum. Hence,

$$\begin{cases} w_1 &= \frac{15}{77} \\ w_2 &= -\frac{8}{77} \end{cases} \dots \square$$