

csci-5454-hw6

Jake Krol

October 2024

1 Question 1

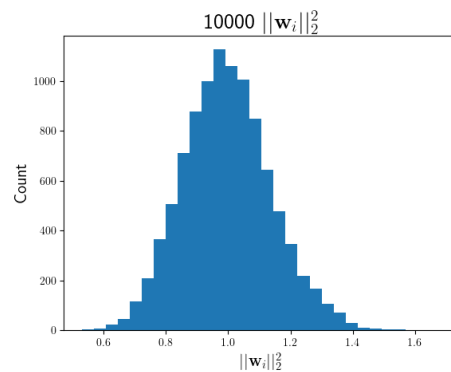
1.1 Question 1a

A fixed random vector, \mathbf{x} , was generated by taking 1000 samples of a standard normal distribution, then the vector was divided by its L2 norm to set $\|\mathbf{x}\|_2 = 1$.

$$\sqrt{\mathbf{x} \cdot \mathbf{x}} = \sqrt{x_1^2 + \dots + x_d^2} = \|\mathbf{x}\|_2$$

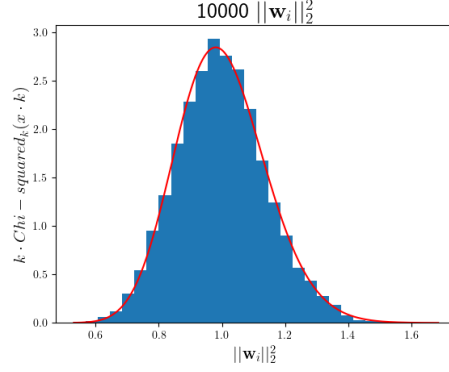
$N = 10000$ Johnson-Lindenstrauss (JL) matrices were generated and multiplied by \mathbf{x} to produce N JL-transformed vectors $\mathbf{w}_1 + \dots + \mathbf{w}_N$. The squared L2 norm $\|\mathbf{w}_i\|_2^2$, was calculated for N transformations.

$$\begin{aligned}\mathbf{w}_i &= G_i \mathbf{x} \\ \|\mathbf{w}_i\|_2^2 &= \mathbf{w}_i \cdot \mathbf{w}_i\end{aligned}$$



1.2 Question 1b

Each element of a JL-transformed vector is a sum of i.i.d. scaled Gaussian random variables $w_i = \sum_{j=1}^d x_j \cdot g_j$ (where w_i is an element of JL-transformed vector, \mathbf{w}), and the squared L2 norm involves summing squared Gaussians which results in a Chi-squared random variable.



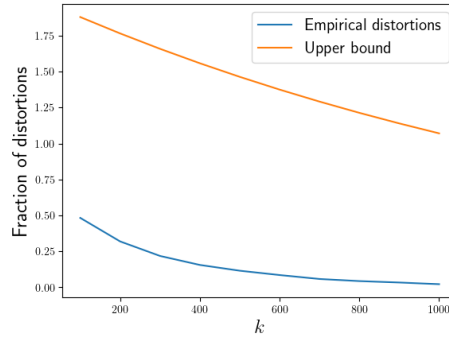
Since $E[g_j] = 0$ and $VAR[g_j] = 1$ the expected value is $E[w_i] = \sum_{j=1}^d x_j E[g_j] = 0$ and, given independence, the variance is $VAR[w_i] = E[w_i^2] - E[w_i]^2 = E[w_i^2] - 0 = \sum_{j=1}^d VAR[x_j \cdot g_j] = \sum_{j=1}^d x_j^2 \cdot VAR[g_j] = \sum_{j=1}^d x_j^2 \cdot 1 = x_1^2 + \dots + x_d^2 = \|x\|_2^2$. By linearity of expectation, the expected value of the sum of squared Gaussians is also $\|x\|_2^2$.

$$\left\| \frac{1}{\sqrt{k}} Gx \right\|_2^2 = \frac{1}{k} \sum_{i=1}^k w_i^2$$

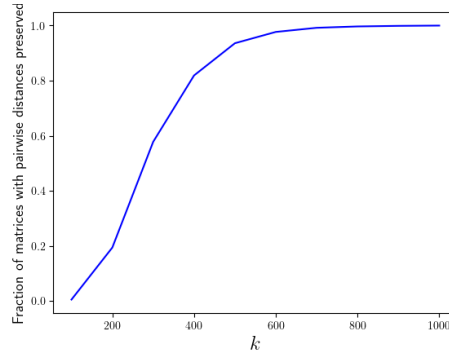
$$E\left[\frac{1}{k} \sum_{i=1}^k w_i^2\right] = E\left[\frac{1}{k} \sum_{i=1}^k \|x\|_2^2\right] = \|x\|_2^2$$

1.2.1 Question 1c

The number of distortions from the empirical simulation and the upper bound $2e^{\frac{-k\epsilon^2}{4}}$ for various k shows the empirical count of distortions is much less than the upper bound. The fraction of distortions represents the probability of $\|\mathbf{w}\|_2^2$ deviating greater than $\epsilon = 0.05$ from the original norm $\|\mathbf{x}\|_2^2$, and the results demonstrate well that this probability will not exceed the upper bound.



1.2.2 Question 1d



1.3 Question 2

1.3.1 Question 2a

The algorithm returns HEADS with $\frac{1}{2^k}$ Pr by leveraging k independent Bernoulli random variables. By using k independent Bernoulli random variables $X_1 \dots X_k$ each with $p = \frac{1}{2}$ the intersection of all outcomes being HEADS is $p^k = \frac{1}{2^k}$. In contrast, the complementary event is handled by returning TAILS immediately if any of the k tosses are not HEADS.

Algorithm 1 BIASED-COIN(k)

```

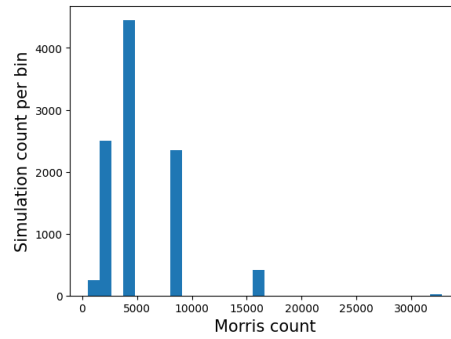
 $i = 0$ 
while  $i < k$  do
     $c = FLIP(p = \frac{1}{2})$  ▷ flip fair coin
    if  $c == \text{HEADS}$  then
         $i = i + 1$ 
    else
        return TAILS
    end if
end while
return HEADS

```

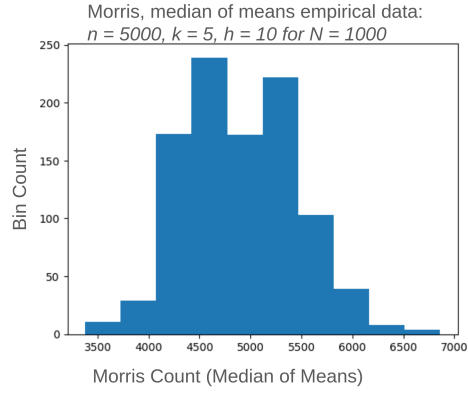
1.3.2 Question 2b

	Theory	$h = 10^4$ simulations
Expected (Avg) Estimated Count	5000	≈ 5032
Standard Deviation	$\sqrt{\frac{n(n-1)}{2}} \approx 3535$	≈ 3583

Table 1: Morris counter empirical results and theory



1.3.3 Question 2c



	Theory	$h = 10(\text{corrected}), k = 5$
Expected (Avg) Estimated Count	5000	≈ 4883
Standard Deviation	- do not bother-	≈ 552

Table 2: Morris counter median of means

The variance of the median of means Morris counter (part 2c) is less than the variance for the 10^4 individual Morris counters (part 2b). Particularly, using samples means, $k = 5$ super counters, results in lower variance for each trial. Taking the median super-counter from each trial also aids in centering the final outcome since medians are rank-based which can mitigate the undesired weight toward high magnitude outliers.