**Assignment 1** Graded Student James La Fontaine **Total Points** 21 / 21 pts Question 1 (no title) **13** / 13 pts **3** / 3 pts 1.1 (no title) **→ + 3 pts** Point adjustment (no title) **7** / 7 pts 1.2 + 7 pts Point adjustment 1.3 (no title) 3 / 3 pts **→ + 3 pts** Point adjustment Question 2 (no title) **3** / 3 pts 2.1 (no title) **3** / 3 pts + 3 pts Point adjustment Question 3 (no title) **5** / 5 pts 3.1 (no title) **5** / 5 pts **▶ + 5 pts** Point adjustment

No questions assigned to the following page.				

## MAST30027 Modern Applied Statistics Assignment 1

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Tutorial: Wed 1-2PM, Yidi Deng

Question assigned to the following page: 1.1

## Question 1

```
(a)  \begin{array}{l} > {\bf library}({\bf faraway}) \\ > {\bf data}({\bf orings}) \\ > \\ > {\bf logL} = {\bf function}({\bf betas}\,,\,\,{\bf orings}) \; \{ \\ + & {\bf eta} = {\bf cbind}(1,\,\,{\bf orings\$temp}) \; \% \# {\bf betas} \\ + & {\bf return}\,\,({\bf sum}({\bf orings\$damage}\,*\,\,{\bf log}({\bf pnorm}({\bf eta})) + (6\,\,-\,\,{\bf orings\$damage}) \\ & & *\,\,\,{\bf log}(1\,\,-\,\,{\bf pnorm}({\bf eta})) \;)) \\ + \; \} \\ > \\ > & ({\bf betahat} = {\bf optim}({\bf c}(10\,,\,\,-.1),\,\,{\bf logL}\,,\,\,{\bf orings=orings}\,, \\ & {\bf control=list}\,({\bf fnscale}\,=\!-1)) {\bf \$par}) \\ [1] \quad 5.5917242 \quad -0.1058008 \\ \\ \hat{\beta}_0 = 5.5917 \\ \\ \hat{\beta}_1 = -0.1058 \\ \end{array}
```

Question assigned to the following page: 1.2

$$\frac{\eta_{i} = \beta_{0} + \beta_{1}, \xi_{2}}{\xi\left(\left(\beta_{0}, \beta_{1}\right) = \zeta + \frac{1}{2}\left[\left(\gamma_{i} \mid 0\right) + \left(\beta_{0} + \beta_{1}\right) + \left(\beta_{0} - \gamma_{1}\right) \mid 0\right] \quad \text{where} \quad \frac{3}{2} \frac{\eta_{i}}{\beta_{0}} = 1}{f(\eta_{i}) + \frac{1}{2}\left[\left(\gamma_{i} \mid 0\right) + \left(\beta_{0} - \gamma_{1}\right) \mid 0\right]} \quad \text{where} \quad \frac{3}{2} \frac{\eta_{i}}{\beta_{0}} = 1$$

$$\frac{3}{2} \frac{\xi\left(\beta_{0}, \beta_{1}\right)}{\beta_{0}} = \sum_{i} \left[\gamma_{i} \frac{f(\eta_{i})}{f(\eta_{i})} + \left(\gamma_{i} - \delta\right) \frac{f(\eta_{i})}{1 - f(\eta_{i})}\right] \quad \frac{3}{2} \frac{\eta_{i}}{\beta_{i}} = t;$$

$$= \sum_{i} \left[\gamma_{i} \frac{f(\eta_{i})}{f(\eta_{i})} + \left(\gamma_{i} - \delta\right) \frac{f(\eta_{i})}{1 - f(\eta_{i})}\right] \quad + \left(\gamma_{i} - \delta\right) \frac{\left(1 - f(\eta_{i})\right) \frac{f'(\eta_{i}) - f(\eta_{i})}{2} - f(\eta_{i})}{\left(1 - f(\eta_{i})\right)^{2}}\right]$$

$$= \sum_{i} \left[\gamma_{i} \times \frac{f(\eta_{i}) f'(\eta_{i}) - f(\eta_{i})^{2}}{f'(\eta_{i})} + \left(\gamma_{i} - \delta\right) \frac{f'(\eta_{i}) - f(\eta_{i}) f'(\eta_{i}) + f(\eta_{i})^{2}}{\left(1 - f(\eta_{i})\right)^{2}}\right]$$

$$= \sum_{i} \left[\gamma_{i} \times \frac{f'(\eta_{i}) f'(\eta_{i}) - f(\eta_{i})^{2}}{f'(\eta_{i})} + \left(\gamma_{i} - \delta\right) \frac{f'(\eta_{i}) - f(\eta_{i}) f'(\eta_{i}) - f(\eta_{i})^{2}}{\left(1 - f(\eta_{i})\right)^{2}}\right]$$

$$= \sum_{i} \left[\gamma_{i} \times \frac{f'(\eta_{i}) f'(\eta_{i}) - f(\eta_{i})^{2}}{f'(\eta_{i})} + \left(\gamma_{i} - \delta\right) \frac{f'(\eta_{i}) - f(\eta_{i}) f'(\eta_{i}) - f(\eta_{i})^{2}}{\left(1 - f(\eta_{i})\right)^{2}}\right]$$

$$E\left(\frac{3^{\frac{1}{2}}L[\rho_{1}h_{2}]}{3^{\frac{1}{2}}\rho_{3}}\right) = \sum_{i} \left[\frac{6}{p_{i}} \times \frac{p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) - \frac{1}{p_{i}}(\eta_{i})^{2}}{p_{i}} - 6\left(1 - p_{i}\right) \frac{f^{i}(\eta_{i}) - p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) + f^{i}(\eta_{i})^{2}}{(1 - p_{i})^{2}}\right]$$

$$= \sum_{i} \left[\frac{6}{p_{i}} \times \frac{p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) - \frac{1}{p_{i}}(\eta_{i})^{2}}{p_{i}} - 6 \cdot \frac{\frac{1}{p_{i}}(\eta_{i}) - p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) + \frac{1}{p_{i}}(\eta_{i})^{2}}{(1 - p_{i})^{2}}\right]$$

$$= C\sum_{i} \left[\frac{p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) - p_{i}}{p_{i}} \frac{f^{i}(\eta_{i}) - \frac{1}{p_{i}} f^{i}(\eta_{i}) + \frac{1}{p_{i}} f^{i}(\eta_{i})^{2}}{p_{i}} - \frac{\frac{1}{p_{i}} \frac{p_{i}}{p_{i}} f^{i}(\eta_{i}) + \frac{1}{p_{i}} f^{i}(\eta_{i})^{2}}{p_{i}} \frac{p_{i}}{p_{i}} - \frac{1}{p_{i}} \frac{p_{i}}{p_{i}} \frac{p_{i}}{p_{i}} - \frac{1}{p_{i}} \frac{p_{i}}{p_{i}} \frac{p_{i}}{p_{i}} \frac{p_{i}}{p_{i}} - \frac{1}{p_{i}} \frac{p_{i}}{p_{i}} \frac{p_{i}}{p_{i}}$$

Question assigned to the following page: 1.2

```
> iprobit = function(x) pnorm(x)
> phat = iprobit(betahat[1] + orings$temp*betahat[2])
> I11 = 3/pi * sum(dexp(qnorm(phat)^2) / (phat*(1-phat)))
> I12 = 3/pi * sum(orings$temp * dexp(qnorm(phat)^2) / (phat*(1-phat)))
> I22 = 3/pi * sum(orings$temp^2 * dexp(qnorm(phat)^2) / (phat*(1-phat)))
>
> Iinv = solve(matrix(c(I11, I12, I12, I22), 2, 2))
> betahat[1] + c(-1,1)*qnorm(0.975)*sqrt(Iinv[1,1])
[1] 2.239762 8.943686
> betahat[2] + c(-1,1)*qnorm(0.975)*sqrt(Iinv[2,2])
[1] -0.15784670 -0.05375481
```

95% Confidence Interval for  $\hat{\beta_0}$ : (2.2398, 8.9437)

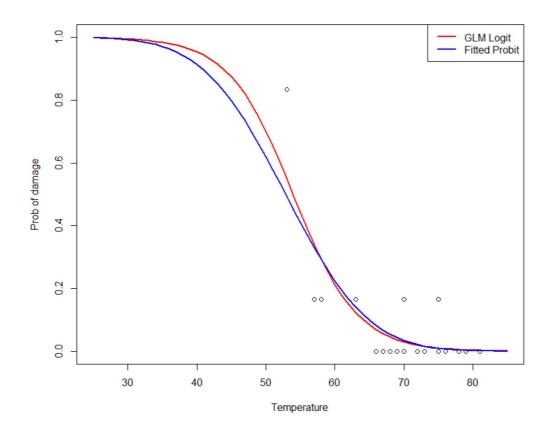
95% Confidence Interval for  $\hat{\beta}_1$ : (-0.1578, -0.0538)

Question assigned to the following page: <u>1.3</u>

```
(c)
> logL.F = function(betas, orings) {
       eta = cbind(1, orings$temp) %*% betas
       \mathbf{return} \ (\mathbf{sum}(\ \mathtt{orings\$damage}\ *\ \mathbf{log}(\mathbf{pnorm}(\ \mathtt{eta}\ \mathtt{)})\ +\ (6\ -\ \mathtt{orings\$damage})
+
                                                             * log(1 - pnorm(eta)))
+ }
>
> logL.R = function(beta0, orings) {
       eta = beta0
       return (sum(orings$damage * log(pnorm(eta)) + (6 - orings$damage)
+
                                                             * log(1 - pnorm(eta)) ))
+ }
>
> (betahat.F = optim(c(10, -.1), logL.F, orings=orings,
                                        control=list(fnscale=-1))par)
[1] 5.5917242 -0.1058008
> (betahat.R = optim(c(5), logL.R, orings=orings, control=list(fnscale=-1))$par)
[1] -1.40625
> (LR = -2*(logL.R(betahat.R, orings) - logL.F(betahat.F, orings)))
[1] 20.76711
> pchisq(LR, df=1, lower=FALSE)
[1] \ 5.186617e{-06}
p-value < 0.05, so we reject H_0: \beta_1 = 0
(d)
> si2 = matrix(c(1, 31), 1, 2) \% \% Iinv \% matrix(c(1, 31), 2, 1)
> etahat = betahat [1] + betahat [2] *31
> \text{eta_l} = \text{etahat} - \mathbf{qnorm}(0.975) * \mathbf{sqrt}(\text{si2})
> \text{eta_r} = \text{etahat} + \text{qnorm}(0.975)*\text{sqrt}(\text{si}2)
> c(eta_l, eta_r)
[1] 0.5557902 4.0680114
> iprobit (etahat)
[1] 0.9896084
> c(iprobit(eta_l), iprobit(eta_r))
[1] 0.7108229 0.9999763
\hat{p} = 0.9896
95% Confidence Interval for \hat{p}: (0.7108, 0.9999)
```

No questions assigned to the following page.				

```
(e)
> ilogit = function(x) exp(x)/(1+exp(x))
> iprobit = function(x) pnorm(x)
> \ \log\_{mod} = \ glm(cbind(damage\,,\ 6-damage)\ \tilde{\ }\ temp\,,\ family=binomial(link="logit")\,,
                      orings)
  \operatorname{plot}(\operatorname{damage}/6 \ \widetilde{\ } \operatorname{temp}, \operatorname{orings}, \operatorname{xlim}=\mathbf{c}(25,85), \operatorname{ylim}=\mathbf{c}(0,1),
         xlab="Temperature", ylab="Prob_of_damage")
  x = seq(25,85,1)
> \ lines(x, \ ilogit(log\_mod\$coefficients[1] \ + \ log\_mod\$coefficients[2]*x), \ col="red",
> lines(x, iprobit(betahat[1] + betahat[2]*x), col="blue", lwd=2)
  legend(x = "topright",
+
            legend = c("GLM_Logit", "Fitted_Probit"),
+
            lty = \mathbf{c}(1),
            col = c("red", "blue"),
+
            lwd = 2)
```



Question assigned to the following page: <u>2.1</u>

## Question 2

```
(a)
> library (faraway)
> missing = with (pima, missing <- glucose==0 | diastolic==0 | triceps==0
                                                      | bmi == 0)
> pima_subset = pima[!missing, c(6,9)]
> str(pima_subset)
'data.frame':
                  532 obs. of 2 variables:
 \$ \  \, \text{bmi} \  \, : \  \, \text{num} \quad \, 33.6 \quad 26.6 \quad 28.1 \quad 43.1 \quad 31 \quad 30.5 \quad 30.1 \quad 25.8 \quad 45.8 \quad 43.3 \quad \dots
 $ test: int 1 0 0 1 1 1 1 1 1 0 ...
> pima_mdl = glm(test ~ bmi, family=binomial(link="logit"),
                      pima_subset)
> phat = ilogit (pima_mdl$coefficients[2]*7)
> log_odds = logit(phat)
> as.numeric(log_odds)
[1] 0.6980179
logit(\hat{p}) = 0.6980
(b)
> phat_l = ilogit ((pima_mdl\$coefficients[2]*7 - qnorm(0.975)
                     * summary(pima_mdl)$coefficients[2, 2] * 7))
> phat_r = ilogit ((pima_mdl\$coefficients[2]*7 + qnorm(0.975)
                     * summary(pima_mdl)$coefficients[2, 2] * 7))
> CI_logodds = c(logit(phat_l), logit(phat_r))
> as.numeric(CI_logodds)
[1] 0.4883237 0.9077121
95% Confidence Interval for logit(\hat{p}): (0.4883, 0.9077)
```

No questions assigned to the following page.				

## Question 3

$$\nu > 0$$
  $\lambda > 0$ 

$$\int (x; \nu, \lambda) = \frac{\lambda^{\nu}}{\Gamma(\nu)} x^{\nu-1} e^{-\lambda x} \qquad x > 0$$

$$\begin{cases}
\exists (x) = \frac{\lambda^{\nu}}{T(\nu)} x^{\nu-1} e^{-\lambda x} \\
= e^{-\lambda} \left[ -\lambda x + (\nu - 1) \log x + \nu \log \lambda - \log T(\nu) \right] \\
= e^{-\lambda} \left[ -\lambda x + \nu \log \lambda + (\nu - 1) \log x - \log T(\nu) \right] \\
= e^{-\lambda} \left[ \frac{x(-\frac{\lambda}{\nu}) - \log (\frac{\nu}{\lambda}) - \log (\nu)}{\frac{1}{\nu}} + (\nu - 1) \log x - \log T(\nu) \right] \\
= e^{-\lambda} \left[ \frac{x(-\frac{\lambda}{\nu}) - \log (\frac{\nu}{\lambda}) - \log (\nu)}{\frac{1}{\nu}} + \frac{-\log (\nu) + (1 - \frac{1}{\nu}) \log x - \frac{1}{\nu} \log T(\nu)}{\frac{1}{\nu}} \right] \\
= e^{-\lambda} \left[ \frac{x(-\frac{\lambda}{\nu}) - \log (\frac{\nu}{\lambda})}{\frac{1}{\nu}} + \frac{\log (\nu) + (1 - \frac{1}{\nu}) \log x - \frac{1}{\nu} \log T(\nu)}{\frac{1}{\nu}} \right] \\
= e^{-\lambda} \left[ \frac{x(-\frac{\lambda}{\nu}) - \log (\frac{\nu}{\lambda}) + \log (\frac{\nu}{\lambda}) - \log (\frac{\nu}{\lambda}) + \log (\frac{\nu}{\lambda}) - \log (\frac{\nu}{\lambda})}{\frac{1}{\nu}} \right] \\
= e^{-\lambda} \left[ \frac{x(-\frac{\lambda}{\nu}) - \log (\frac{\nu}{\lambda}) + \log (\frac{\nu}{\lambda}) - \log (\frac{\nu}$$

The gamma distribution is an exponential family

Question assigned to the following page: <u>3.1</u>

$$\Theta = -\frac{\lambda}{\nu} \qquad b(\theta) = \log(-\frac{1}{\theta}) \qquad \alpha(\phi) = \frac{1}{\nu}$$
Variance function  $\nu(M) = b''((b')^{-1}(M))$ 

$$b''(\theta) = \frac{1}{-\frac{1}{\theta}} \times \frac{1}{\theta^2} = -\frac{1}{\theta} \quad , \quad (b')^{-1}(M) = -\frac{1}{M}$$

$$b'''(\theta) = \frac{1}{\theta^2}$$

$$b^{*}(\theta) = \frac{1}{\theta^{2}}$$

$$\Rightarrow v(\mu) = \frac{1}{(-\frac{1}{\mu})^{2}}$$

$$= \frac{1}{\frac{1}{\mu^{2}}} = \mu^{2}$$

$$\Rightarrow$$
 canonical link  $g(\mu) = -\frac{1}{\mu}$