

## **Assignment 3**

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**Tutorial Day and Time:** Friday 2:15 PM – 4:15 PM

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## Question 1

1ai)

$H_0: m_X = m_Y$

$H_1: m_X \neq m_Y$

```
binom.test(12, 17)

##
##  Exact binomial test
##
## data:  12 and 17
## number of successes = 12, number of trials = 17, p-value = 0.1435
## alternative hypothesis: true probability of success is not equal to 0.5
## 95 percent confidence interval:
##  0.4404173 0.8968645
## sample estimates:
## probability of success
##                0.7058824
```

The p-value > 0.05 so we cannot reject the null hypothesis that  $m_X = m_Y$  at the significance level of 5%.

1aii)

$H_0: m_X = m_Y$

$H_1: m_X \neq m_Y$

```
x = c(26.1, 26.6, 27.4, 27.5, 27.8, 28.1, 28.4, 29.5, 29.8, 30.4, 30.4, 31.2,
31.5, 32.9, 33.6, 34.1, 35.9)
y = c(27.4, 28.1, 22.9, 31.3, 16.3, 50.1, 20.0, 24.6, 23.3, 19.3, 24.4, 24.4,
29.5, 27.6, 21.7, 25.4, 39.4)

wilcox.test(x, y, paired = TRUE)

##
##  Wilcoxon signed rank exact test
##
## data:  x and y
## V = 124, p-value = 0.02322
## alternative hypothesis: true location shift is not equal to 0
```

The p-value < 0.05 so we reject the null hypothesis that  $m_X = m_Y$  at the significance level of 5% and can conclude that there is sufficient evidence to show that the location of X and Y differ.

### 1aiii)

$H_0: \mu_X = \mu_Y$

$H_1: \mu_X \neq \mu_Y$

```
x = c(26.1, 26.6, 27.4, 27.5, 27.8, 28.1, 28.4, 29.5, 29.8, 30.4, 30.4, 31.2,
31.5, 32.9, 33.6, 34.1, 35.9)
y = c(27.4, 28.1, 22.9, 31.3, 16.3, 50.1, 20.0, 24.6, 23.3, 19.3, 24.4, 24.4,
29.5, 27.6, 21.7, 25.4, 39.4)

t.test(x, y, paired = TRUE)

##
## Paired t-test
##
## data: x and y
## t = 1.6402, df = 16, p-value = 0.1205
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.954790 7.484202
## sample estimates:
## mean of the differences
## 3.264706
```

The p-value > 0.05 so we cannot reject the null hypothesis that  $\mu_X = \mu_Y$  at the significance level of 5%.

### 1b)

The sign test has a larger type II error rate / smaller power and so it is plausible that the null hypothesis has incorrectly not been rejected in this test. For the t-test we have made the assumption that the differences between X and Y are normally distributed and since we have a small sample size of only 17, it is plausible that these differences do not follow a normal distribution. It would be more appropriate to give more consideration to the outcome of the Wilcoxon signed-rank test in this case, which simply assumes that the differences between X and Y are continuous and follow a symmetrical distribution, which is a reasonable assumption under the null hypothesis. Therefore, there is mild evidence that X and Y differ in location, however, further testing with a larger sample would be required to make stronger conclusions.

1c)

```
B = 20000
n = 17
numRejectionsSign = 0
numRejectionsWilcoxon = 0
numRejectionsT = 0

for (i in 1:B) {
  numSuccesses = 0
  sampleDifference = rnorm(n, 3, 5)
  for (number in sampleDifference) {
    if (sign(number) == 1) {
      numSuccesses = numSuccesses + 1
    }
  }
  if (binom.test(numSuccesses, n)$p.value < 0.05) {
    numRejectionsSign = numRejectionsSign + 1
  }
  if (wilcox.test(sampleDifference)$p.value < 0.05) {
    numRejectionsWilcoxon = numRejectionsWilcoxon + 1
  }
  if (t.test(sampleDifference)$p.value < 0.05) {
    numRejectionsT = numRejectionsT + 1
  }
}

powerSign = numRejectionsSign / B
powerWilcoxon = numRejectionsWilcoxon / B
powerT = numRejectionsT / B

cat("Simulated power of sign test: ", powerSign, "\n")
Simulated power of sign test: 0.4859

cat("Simulated power of Wilcoxon test: ", powerWilcoxon, "\n")
Simulated power of Wilcoxon test: 0.60815

cat("Simulated power of t-test: ", powerT, "\n")
Simulated power of t-test: 0.64395
```

## Question 2

2a)

```
germinations = c(3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 17)
count = c(1, 2, 2, 4, 10, 16, 9, 11, 13, 4, 7, 1)

experiments = data.frame(germinations, count)

data = rep(experiments$germinations, experiments$count)

p1 = prop.test(sum(data), 80*30)

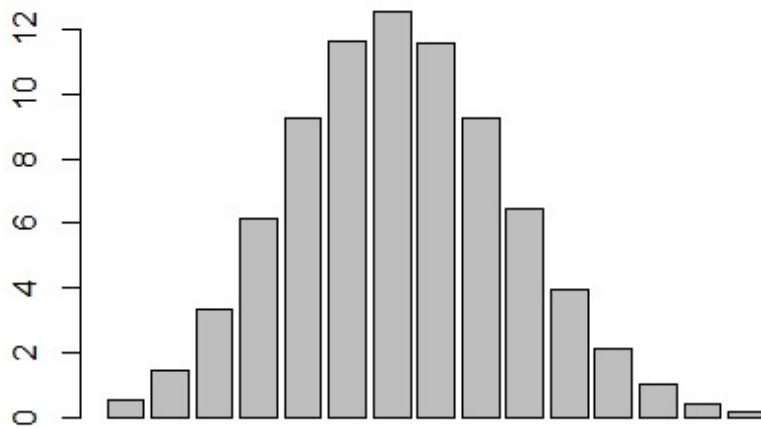
p1

##
## 1-sample proportions test with continuity correction
##
## data:  sum(data) out of 80 * 30, null probability 0.5
## X-squared = 362.7, df = 1, p-value < 2.2e-16
## alternative hypothesis: true p is not equal to 0.5
## 95 percent confidence interval:
##  0.2871084 0.3243539
## sample estimates:
##           p
## 0.3054167

prop.estimate = as.numeric(p1$estimate)
```

2b)

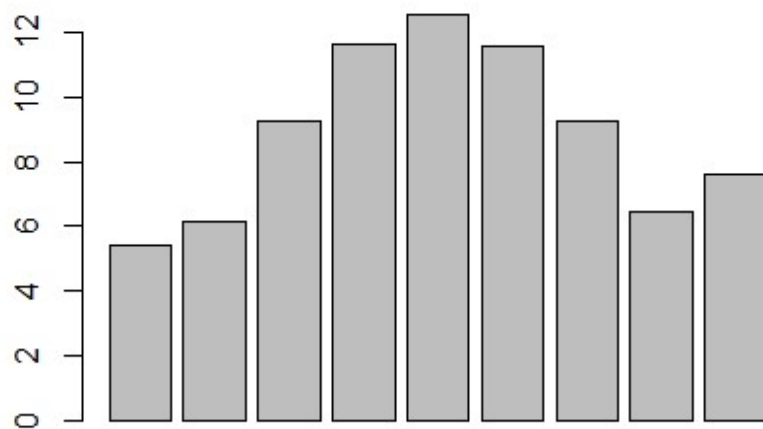
```
barplot(dbinom(3:17, 30, prop.estimate) * 80)
```



```
X1 <- cut(data, breaks = c(0, 5.5, 6.5, 7.5, 8.5, 9.5, 10.5, 11.5, 12.5, Inf))
T1 <- table(X1)
grouped.data <- as.numeric(T1)

p <- c(pbinom(5, 30, prop.estimate), dbinom(6:12, 30, prop.estimate), 1 - pbinom(12, 30, prop.estimate))

barplot(p * 80)
```



2c)

```
chi1 = chisq.test(x=grouped.data, p = p)
```

```
chi1
```

```
##
```

```
## Chi-squared test for given probabilities
```

```
##
```

```
## data: grouped.data
```

```
## X-squared = 5.924, df = 8, p-value = 0.6557
```

```
X.squared = unname(chi1$statistic)
```

```
# recalculate the p-value using the correct degrees of freedom
```

```
1 - pchisq(unname(X.squared), length(grouped.data) - 2)
```

```
## [1] 0.5486503
```

The p-value  $0.5487 > 0.05$  and so there is insufficient evidence to conclude that there is a difference between a Binomial distribution and the distribution of the number of germinations of seeds of the tested plant.

## Question 4

4a)

```
Angle = c(rep(seq(0,30,10), each=10))

Panel = c(rep(rep(1:5, each = 2), 4))

Power = c(42.3, 41.4, 42.2, 40.3, 37.6, 35.7, 36.8, 34.9, 45.8, 43.7, 42.1, 40.2, 42.1, 40.3, 38.4, 36.5, 38.0, 37.1, 45.2, 43.1, 42.6, 40.8, 42.7, 40.8, 38.6, 36.7, 40.2, 38.3, 46.9, 44.8, 43.6, 41.5, 43.8, 41.9, 41.9, 39.8, 42.9, 40.8, 45.4, 43.5)

data = data.frame(Angle, Panel, Power)

model1 = lm(Power ~ factor(Angle) + factor(Panel), data = data)

anova(model1)

## Analysis of Variance Table
##
## Response: Power
##              Df Sum Sq Mean Sq F value    Pr(>F)
## factor(Angle)  3  36.890   12.297    5.8632  0.002614 **
## factor(Panel)  4 235.522   58.880   28.0748 4.602e-10 ***
## Residuals     32  67.113    2.097
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Let  $\alpha$  = the effect of angle elevation on power output.

Let  $\beta$  = the effect of panel type on power output.

$$\mu_{ij} = \mu + \alpha_i + \beta_j$$

$$H_{0A}: \alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = 0$$

$$H_{1A}: \bar{H}_{0A}$$

$$F \text{ value} = 5.8632$$

$$p\text{-value} = 0.002614$$

Assumptions:

- There is no statistical interaction between the factors and thus factor effects are additive.
- We have random samples drawn independently of each other from the different populations, each having a normal distribution.



- All populations have the same variance,  $\sigma^2$ .

0.002614 < 0.05. Therefore, there is sufficient evidence to conclude at the 5% level of significance that the mean power output of solar panels varies between the different angles of elevation and thus that the angle of elevation influences mean power output.

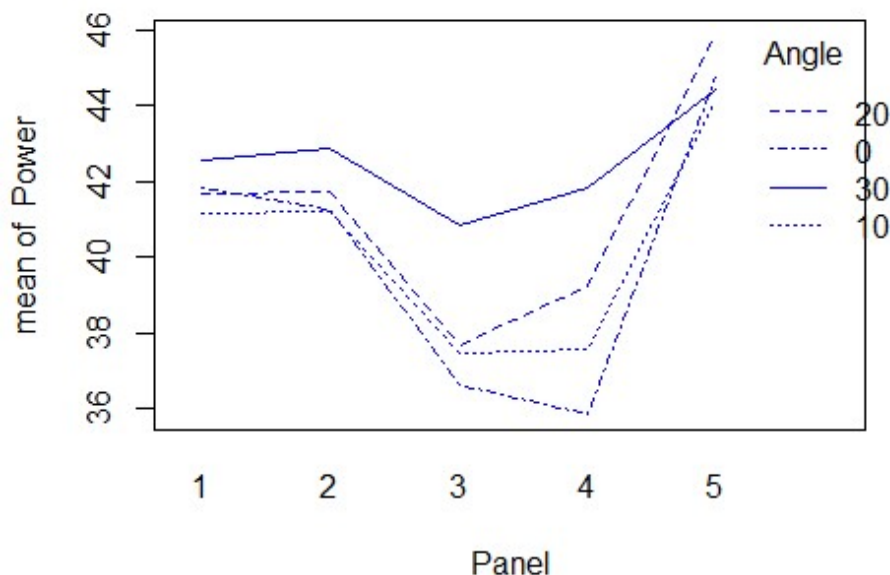
4b)

```
model2 = lm(Power ~ factor(Angle) * factor(Panel), data = data)

anova(model2)

## Analysis of Variance Table
##
## Response: Power
##
##           Df Sum Sq Mean Sq F value    Pr(>F)
## factor(Angle)      3  36.890   12.297   6.9610  0.002163 **
## factor(Panel)       4 235.522   58.880  33.3317 1.383e-08 ***
## factor(Angle):factor(Panel) 12  31.782    2.649   1.4993  0.204458
## Residuals          20  35.330    1.767
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

with(data, interaction.plot(Panel, Angle, Power, col = "blue"))
```



0.2045 > 0.05. Therefore, there is insufficient evidence to conclude at the the 5% level of significance that there is interaction between panel type and the angle of elevation.