

# **Does resting energy expenditure differ between women with and without polycystic ovary syndrome (PCOS): a systematic review and meta-analysis**

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Background: Polycystic ovary syndrome (PCOS) is the most common endocrine condition affecting women of reproductive age. Women who are living with PCOS are also more likely to experience overweight and obesity and this is assumed to be a result of a lower resting energy expenditure (REE). The assumption that REE is lower in women living with PCOS has also influenced lifestyle-based interventions in this group. Evidence supporting this reduced REE compared to women without PCOS, however, is not consistent. Objectives: To systematically search and meta-analyse the existing literature to determine if resting energy expenditure differs between women with and without PCOS. Design: A systematic search was conducted using PubMed, Medline and Web of Science databases of published research from January 1990 to January 2025. Studies that measured REE in women living with PCOS, both with and without control arms of women without PCOS, were included. Results: Seventeen independent studies were identified. Study populations ranged from X to X participants, with mean ages ranging from X to X years. Conclusions: TO ADD

## **Introduction**

Polycystic ovary syndrome (PCOS) affects approximately 10% of women of reproductive age worldwide, making it the most common endocrine disorder affecting this population (Salari et al., 2024). Due to a number of factors including hyperandrogenism and alterations in insulin resistance, PCOS is believed to contribute to an increased risk of diabetes, metabolic syndrome and cardiovascular disease (Glintborg et al., 2018, 2022, 2024; Lim et al., 2019; Persson et al.,

2021), along with being a leading cause of anovulatory infertility in women (Joham et al., 2015). Furthermore, epidemiological data consistently demonstrates that women with PCOS are significantly more likely to suffer from overweight or obesity, compared to the general female population, with estimates ranging from 38% to 88% of PCOS patients falling into overweight or obese body mass index (BMI) categories (Barber, 2022; Barber & Franks, 2021).

The elevated incidence of overweight and obesity in PCOS is likely multifactorial with proposed mechanisms including blunted postprandial appetite hormone responses leading to reduced satiety and increased food cravings (Hirschberg et al., 2004; Moran et al., 2004; Stefanaki et al., 2024) and a reduced resting energy expenditure (REE) (Georgopoulos et al., 2009). Indeed, the study from Georgopoulos et al. (2009) examining REE using indirect calorimetry in women with PCOS reported that they had a significantly lower REE, ~400 kcal per day less, compared with women without PCOS. Notably, they also reported insulin resistance further influenced REE among women with PCOS, with those who were insulin resistant exhibiting an additional reduction of about ~500 kcal per day compared to women with PCOS who were not insulin resistant (Georgopoulos et al., 2009). Other studies similarly report lower BMR using indirect methods (such as prediction from bioelectrical impedance analysis or accelerometer physical activity data) in women with PCOS, also highlighting factors like insulin resistance or BMI category influence REE [Romualdi et al. (2019); churchillBasalMetabolicRate2015]. However, despite the widespread acceptance that women living with PCOS exhibit reduced REE based on studies such as these, other research has reported little to no difference in REE between women with and without PCOS (Graff et al., 2017; Larsson et al., 2016; Segal & Dunaif, 1990).

The consequences of widespread acceptance that REE is reduced in women living with PCOS should not be underestimated, particularly in light of the aforementioned incidence of overweight and obesity in this population. From a physiological perspective, if women with PCOS do exhibit a lower REE, this could imply a meaningful metabolic disadvantage that may influence dietary and nutritional guidance for weight management. However, if no such difference exists, the belief in a “slower metabolism” could instead serve as a deterrent to weight-loss efforts—efforts that are recognised as efficacious for improving PCOS symptoms (Alenezi et al., 2024; Holte et al., 1995) and are routinely recommended (Teede et al., 2023). It has been well documented that women with PCOS already experience higher rates of anxiety, depression, and lower quality of life (QOL) as a result of negative body image and weight-related concerns (Davitadze et al., 2023; Geller et al., 2025; Himelein & Thatcher, 2006; Hofmann et al., 2025). Consequently, they are often observed to engage in more frequent weight-loss attempts than women without PCOS (Pesonen et al., 2023). Clarifying the relationship between REE and PCOS may therefore help guide more accurate clinical recommendations and empower both practitioners and women with PCOS.

Therefore, to estimate the magnitude of difference in REE between women with and without PCOS, we completed a systematic review and meta-analysis of studies reporting REE in these populations.

## Methods

This systematic review and meta-analysis was pre-registered on PROSPERO ([CRD42024601434](https://www.crd42024601434)) initially on the 3<sup>rd</sup> of December 2024 and performed in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement guidelines (Page et al., 2021). The primary aim of this review was to examine the descriptive question “Does resting energy expenditure (REE) differ between women with and without polycystic ovary syndrome (PCOS)?”. We summarise and describe the studies in addition to quantitatively synthesising their results via meta-analysis.

### Search Strategy

PubMed, Web of Science, and MEDLINE databases were searched using the following Boolean search string: ((“Basal Metabolic Rate”[MeSH] OR “Energy Metabolism”[MeSH] OR “Resting Metabolic Rate” OR RMR OR “Resting Energy Expenditure” OR REE OR “Basal Metabolic Rate” OR BMR OR “resting energy” OR “basal energy expenditure”) AND (“Polycystic Ovary Syndrome”[MeSH] OR “Polycystic Ovary Syndrome” OR PCOS OR “Polycystic Ovarian Disease” OR “Stein-Leventhal Syndrome”)). Searches were limited to publications up until May 2025 when the search was completed, limited to English language articles, and Rayyan was used to manage the search and screening process. Two reviewers (RK and GK) independently screened all titles and abstracts against the predefined inclusion and exclusion criteria. Articles deemed potentially eligible by either reviewer were retrieved in full text. Full texts were then independently assessed by RK and LP to determine final eligibility. Any disagreements at either stage were resolved through discussion, and when consensus could not be reached, a third reviewer acted as an adjudicator.

### Eligibility Criteria

Studies were included in the systematic review if 1) participants were confirmed as women with PCOS between the ages of 18 to 65 years of age with or without insulin resistance; 2) otherwise healthy (e.g., non-diabetic, no cardiovascular disease); 3) had a measure of REE measured via multiple methods including direct/indirect calorimetry, doubly labelled water; and 4) trials were not retracted at the time of this analysis. Studies were excluded if they 1) used invalid or non-standard methods for measuring REE (e.g., predicted REE from body composition or accelerometer data); 2) non-peer-reviewed journal articles (including grey literature sources such as conference abstracts, theses and dissertations); and 3) were secondary analyses with the same primary outcome data as another included study.

The condition being studied was PCOS and we included observational cross-sectional design studies, in addition to studies of intervention where REE was reported for the population (and if present, the comparator i.e., women without PCOS) condition of interest. For clarity, studies

of any design were included if they reported the REE using the methods indicated for a sample of adult women with PCOS and who were otherwise healthy. This included both studies with and without samples of healthy control women without PCOS. As detailed in the statistical analysis section below, a Bayesian model with informative priors based on normative data for REE in health women without PCOS was included to provide control information indirectly where this was missing. The use of such priors is an efficient tool to incorporate historical information on a particular population in a conservative manner (Weber et al., 2021).

Following the PICO framework our eligibility criteria can be defined as follows:

- Population
  - Inclusion criteria:
    - \* Women
    - \* 18-65 y
    - \* With or without insulin resistance (IR)
- Intervention(s) or exposure(s)
  - Otherwise healthy women with PCOS
- Comparator(s) or control(s)
  - Otherwise healthy control women without PCOS
- Outcome
- Inclusion criteria
  - Resting energy expenditure (REE) measured via multiple methods including direct/indirect calorimetry, doubly labelled water.
- Exclusion criteria:
  - Studies using invalid or non-standard methods for measuring REE (e.g., predicted REE from body composition or accelerometer data)

## **Data extraction (selection and coding)**

Bibliometric data including authors, journal, and title of article were extracted. Descriptive statistics for age, body mass, fat mass, fat free mass, height, BMI, race, physical activity levels, country of investigation, information regarding glucose/insulin regulation and insulin resistance status (where available), diagnostic criteria for PCOS, and measurement method and device were extracted for each arm within each study in addition to sample size. Descriptive characteristics were then tabulated across studies for reporting.

For each arm, and observation time point if multiple observations reported (e.g., before and after an intervention), depending on what was reported by the authors we extracted the means,

medians, standard deviations, standard errors, lower and upper range values, and interquartile range for the unadjusted and/or body mass adjusted and/or fat free mass adjusted REE values. Where REE values adjusted for body mass and/or fat free mass were reported we used the reported body mass and/or fat free mass mean values for that arm to convert them to unadjusted REE values (i.e., multiplied them by body mass and/or fat free mass mean values). Where means and/or standard deviations were missing the latter were either calculated from standard errors and sample size, or all both were estimated from lower and upper range, interquartile range, median, and sample size depending on the information available using the methods of Wan et al. (2014). Further, where missing, height/body mass/BMI were estimated based on the means reported for these. The units of measurement for which REE was extracted and all REE values were converted to kcal/day. In one case (Pohlmeier et al., 2014) REE was reported relative to body mass and the unadjusted values were no longer available (confirmed by the authors) as such in this case we used the mean body mass to convert back to estimated REE unadjusted.

### **Studies with possible reporting errors**

During data extraction it was noted that several studies from the same lab/research group (Georgopoulos et al., 2009; Koika et al., 2009; Kritikou et al., 2006; Saltamavros et al., 2007) contained a number of discrepancies that seemed to be possible reporting errors. This included, based on taking the authors results as written, standard errors that implied impossible or at least incredibly unlikely standard deviations, and discrepancies in sample size reporting throughout for most variables without explanation or where this was explained the sample sizes were discrepant with the text. Further, data was not reported for the healthy control women without PCOS in three of the studies (Koika et al., 2009; Kritikou et al., 2006; Saltamavros et al., 2007), and REE was reported as an “adjusted” value whereby  $REE_{adjusted} = REE_{group\ mean} + (REE_{adjusted} - REE_{predicted})$  and the  $REE_{predicted}$  was obtained by substituting the individual lean body mass, fat mass, gender, and age in the linear regression equation generated by the data of all patients. In correspondence with the senior author we were unable to clarify the reporting discrepancies as the person responsible for the data/results was no longer contactable. The original data were also no longer available and so we could not calculate the unadjusted REE.

Given these issues we made a decision to extract the results from these studies as reported, but to conduct analyses both with and without their inclusion. Though not pre-registered, due to a lack of confidence in the results as reported, we decided to include the analysis omitting these studies as our main models in the results reported below. The results of the analysis including them are however reported in the sensitivity analysis section.

## Statistical Analysis

Statistical analysis of the data extracted was performed in R, (v 4.3.3; R Core Team, <https://www.r-project.org/>) and RStudio (v 2023.06.1; Posit, <https://posit.co/>). All code utilised for data preparation, transformations, analyses, plotting, and reporting are available in the corresponding GitHub repository [https://github.com/jamessteelei/pcos\\_ree\\_meta](https://github.com/jamessteelei/pcos_ree_meta). We cite all software and packages used in the analysis pipeline using the `grateful` package (Rodriguez-Sanchez et al., 2023) which can be seen here: [https://github.com/jamessteelei/pcos\\_ree\\_meta/blob/main/grateful-report.pdf](https://github.com/jamessteelei/pcos_ree_meta/blob/main/grateful-report.pdf). The statistical analysis plan was linked in our pre-registration (PROSPERO: [CRD42024601434](https://doi.org/10.1186/1745-6215-42024601434)) and available at the accompanying github repository. Any deviations from the pre-registration are noted below.

Given our research question our analysis was aimed at parameter estimation (Cumming, 2014) within a Bayesian meta-analytic framework (Kruschke & Liddell, 2018). For all analyses model parameter estimates and their precision, along with conclusions based upon them, are interpreted continuously and probabilistically, considering data quality, plausibility of effect, and previous literature, all within the context of each model. The `renv` package (Ushey et al., 2023) was used for package version reproducibility and a function based analysis pipeline using the `targets` package (Landau et al., 2023) was employed (the analysis pipeline can be viewed by downloading the R Project and running the function `targets::tar_visnetwork()`). Effect sizes and their variances were all calculated using the `metafor` packages `escalc()` function (Viechtbauer, 2023). The main package `brms` (Bürkner et al., 2023) was used in fitting all the Bayesian meta-analysis models. Prior and posterior draws were taken using `marginalEffects` (Arel-Bundock et al., 2023) and `tidybayes` (Kay & Mastny, 2023) packages. All visualisations are created using `ggplot2` (Wickham et al., 2023), `tidybayes`, and the `patchwork` (Pedersen, 2023) packages.

## Main Pre-registered Models

We adopted an arm-based multiple condition comparison (i.e., network) type model given that the studies included had arms of women with PCOS both with, and without, a non-PCOS control arm (Hong et al., 2016), and also in some cases multiple observations of REE in the different arms included in the study (for example, where an intervention was conducted and pre- and post-intervention REE was reported). In typical contrast-based meta-analyses data is limited to the effect sizes for paired contrasts between arms and thus studies that include both arms (i.e., relative effects between non-PCOS control vs PCOS arms); however, in arm-based analyses the data are the absolute effects within each arm and information is borrowed across studies to enable both within condition absolute, and between condition relative contrasts to be estimated. We made use of historical information regarding REE in healthy control women without PCOS by setting informative priors based on meta-analysis of large scale studies reporting normative data for REE in this population. This was included to provide control information indirectly where this was missing from particular studies. The use of

historical priors like this is an efficient tool to incorporate historical information on a particular population in a conservative manner in meta-analyses (Weber et al., 2021). From this model we focus on reporting the global grand mean estimate for the fixed between condition relative contrast for non-PCOS control vs PCOS arms as our primary estimand of interest (i.e.,  $\beta_1$  in both mean and standard deviation models). We examined both raw mean REE (i.e., the absolute mean REE in kcals per day for each arm) in addition to the between person standard deviation in REE (i.e., the absolute standard deviation in REE in kcals per day for each arm). Both models were multilevel in that they included nested random intercepts for both study, and arm within study, and in addition and in deviation from our pre-registration we also included lab as a random intercept as in some cases we had multiple studies from the same lab or research group included. Lastly, the inclusion of a random intercept for each effect size was also accidentally omitted from our pre-registration and so this is also included in the model.

### Mean REE Model

The main model for mean REE with cond representing the condition (either control or PCOS) was as follows:

$$\begin{aligned}\hat{\theta}_{ijkl} &\sim \mathcal{N}(\mu_{ijkl}, \sigma_{ijkl}^2) \\ \mu_{ijkl} &= \beta_0 + \beta_1 \text{cond}[ijk] + \alpha_{0,\text{lab}[i]} + \alpha_{1,\text{lab}[i]} \text{cond} + \alpha_{0,\text{study}[j]} + \alpha_{1,\text{study}[j]} \text{cond} + \alpha_{0,\text{arm}[k]} + \alpha_{0,\text{effect}[l]} \\ \begin{pmatrix} \alpha_{0,\text{lab}[i]} \\ \alpha_{1,\text{lab}[i]} \end{pmatrix} &\sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{\text{lab}}\right), \quad \Sigma_{\text{lab}} = \begin{pmatrix} \sigma_{0,\text{lab}}^2 & \rho_{\text{lab}} \sigma_{0,\text{lab}} \sigma_{1,\text{lab}} \\ \rho_{\text{lab}} \sigma_{0,\text{lab}} \sigma_{1,\text{lab}} & \sigma_{1,\text{lab}}^2 \end{pmatrix}, \text{ for lab } i = 1, \dots, I \\ \begin{pmatrix} \alpha_{0,\text{study}[j]} \\ \alpha_{1,\text{study}[j]} \end{pmatrix} &\sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{\text{study}}\right), \quad \Sigma_{\text{study}} = \begin{pmatrix} \sigma_{0,\text{study}}^2 & \rho_{\text{study}} \sigma_{0,\text{study}} \sigma_{1,\text{study}} \\ \rho_{\text{study}} \sigma_{0,\text{study}} \sigma_{1,\text{study}} & \sigma_{1,\text{study}}^2 \end{pmatrix}, \text{ for study } j = 1, \dots, J \\ \alpha_{0,\text{arm}[k]} &\sim \mathcal{N}(0, \sigma_{0,\text{arm}}^2), \text{ for arm } k = 1, \dots, K \\ \alpha_{0,\text{effect}[l]} &\sim \mathcal{N}(0, \sigma_{0,\text{effect}}^2), \text{ for effect } l = 1, \dots, L\end{aligned}$$

where  $\hat{\theta}_{ijkl}$  is the  $l$ th mean REE estimate from the  $k$ th arm, for the  $j$ th study, conducted by the  $i$ th lab and  $\sigma_{ijkl}^2$  is the corresponding sampling error for that estimate. The random intercepts for the  $i$ th lab,  $j$ th study,  $k$ th arm, and  $l$ th mean REE estimate are  $\alpha_{0,\text{lab}[i]}$ ,  $\alpha_{0,\text{study}[j]}$ ,  $\alpha_{0,\text{arm}[k]}$ , and  $\alpha_{0,\text{effect}[l]}$  respectively each with standard deviation of  $\sigma_{0,\text{lab}[i]}^2$ ,  $\sigma_{0,\text{study}[j]}^2$ ,  $\sigma_{0,\text{arm}[k]}^2$ , and  $\sigma_{0,\text{effect}[l]}^2$ . The parameter  $\beta_0$  represents the fixed effect estimate of REE for control conditions and  $\beta_1$  the fixed effect estimate for the offset from this for the PCOS conditions (i.e., the difference between conditions). The estimated offset was allowed to vary across both labs and studies each reflected by  $\alpha_{1,\text{lab}[i]}$  and  $\alpha_{1,\text{study}[j]}$  respectively, and these effects were also modelled as correlated with the corresponding random intercepts with covariance  $\Sigma_{\text{lab}}$  and  $\Sigma_{\text{study}}$ , and  $\text{corr}_{\text{lab}}$  and  $\text{corr}_{\text{study}}$  correlation matrices.

The priors for this model were as follows:

$$\beta_0 \sim \text{Student-}t(3, 1441.81, 84.56)$$

$$\beta_1 \sim \text{Student-}t(3, 0, 200)$$

$$\sigma_{0,\text{study}}^2 \sim \text{Half-student-}t(3, 149.89, 82.91)$$

$$\sigma_{0,\text{lab}}^2, \sigma_{1,\text{lab}}^2, \sigma_{1,\text{study}}^2, \sigma_{0,\text{arm}}^2, \sigma_{0,\text{effect}}^2 \sim \text{Half-student-}t(3, 0, 112.4)$$

$$\text{corr}_{\text{lab}} \sim \text{LKJ}(1)$$

$$\text{corr}_{\text{study}} \sim \text{LKJ}(1)$$

where the prior for  $\beta_0$ , which corresponded to the model intercept and mean REE in the control condition, was set based on meta-analysis of the mean REEs for women from two large studies of healthy people (Pavlidou et al., 2023; Velasquez, 2011) though set with a conservative degrees of freedom for the Student- $t$  distribution. The random intercept  $\sigma_{0,\text{study}}^2$  was set similarly to this. The prior for the fixed effect  $\beta_1$ , reflecting the difference between control and PCOS conditions was set based on a wide range of possible values considering the minimum and maximum values of the ranges reported in the two studies noted (i.e.,  $2492 - 908 = 1584$ ). We then set a prior that permits values approximately across this range of values with the majority of its mass centred around zero. The remaining random effects were set based on the default weakly regularising priors for `brms` and scaled to the expected response values using a half-student- $t$  distribution with 3 degrees of freedom and  $\mu = 0$ , and both correlation matrices  $\text{corr}_{\text{lab}}$  and  $\text{corr}_{\text{study}}$  were set with an LKJcorr(1) distribution.

### Standard Deviation of REE Model

The main model for the standard deviation of REE with `cond` representing the condition (either control or PCOS) was as follows:



$$\log(\hat{\theta}_{ijkl}) \sim \mathcal{N}(\mu_{ijkl}, \sigma_{ijkl}^2)$$

$$\mu_{ijkl} = \beta_0 + \beta_1 \text{cond}_{[ijkl]} + \beta_2 \tilde{m}_{[ijkl]} + \alpha_{0,\text{lab}[i]} + \alpha_{1,\text{lab}[i]} \text{cond} + \alpha_{0,\text{study}[j]} + \alpha_{1,\text{study}[j]} \text{cond} + \alpha_{0,\text{arm}[k]} + \alpha_{0,\text{effect}[l]}$$

$$m = \log(y_{i,\text{mean},[ijkl]})$$

$$\tilde{m} \sim \mathcal{N}(m, \sigma_{\log(y_{i,\text{mean},[ijkl]})}^2)$$

$$\begin{pmatrix} \alpha_{0,\text{lab}[i]} \\ \alpha_{1,\text{lab}[i]} \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{\text{lab}}\right), \quad \Sigma_{\text{lab}} = \begin{pmatrix} \sigma_{0,\text{lab}}^2 & \rho_{\text{lab}} \sigma_{0,\text{lab}} \sigma_{1,\text{lab}} \\ \rho_{\text{lab}} \sigma_{0,\text{lab}} \sigma_{1,\text{lab}} & \sigma_{1,\text{lab}}^2 \end{pmatrix}, \text{ for lab } i = 1, \dots, I$$

$$\begin{pmatrix} \alpha_{0,\text{study}[j]} \\ \alpha_{1,\text{study}[j]} \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{\text{study}}\right), \quad \Sigma_{\text{study}} = \begin{pmatrix} \sigma_{0,\text{study}}^2 & \rho_{\text{study}} \sigma_{0,\text{study}} \sigma_{1,\text{study}} \\ \rho_{\text{study}} \sigma_{0,\text{study}} \sigma_{1,\text{study}} & \sigma_{1,\text{study}}^2 \end{pmatrix}, \text{ for study } j = 1, \dots, J$$

$$\alpha_{0,\text{arm}[k]} \sim \mathcal{N}(0, \sigma_{0,\text{arm}}^2), \text{ for arm } k = 1, \dots, K$$

$$\alpha_{0,\text{effect}[l]} \sim \mathcal{N}(0, \sigma_{0,\text{effect}}^2), \text{ for effect } l = 1, \dots, L$$

where  $\log(\hat{\theta}_{ijkl})$  is the  $l$ th natural logarithm of the standard deviation of REE estimate from the  $k$ th arm, for the  $j$ th study, conducted by the  $i$ th lab and  $\sigma_{ijkl}^2$  is the corresponding sampling error for that estimate. The random intercepts for the  $i$ th lab,  $j$ th study,  $k$ th arm, and  $l$ th standard deviation of REE estimate are  $\alpha_{0,\text{lab}[i]}$ ,  $\alpha_{0,\text{study}[j]}$ ,  $\alpha_{0,\text{arm}[k]}$ , and  $\alpha_{0,\text{effect}[l]}$  respectively each with standard deviation of  $\sigma_{0,\text{lab}[i]}^2$ ,  $\sigma_{0,\text{study}[j]}^2$ ,  $\sigma_{0,\text{arm}[k]}^2$ , and  $\sigma_{0,\text{effect}[l]}^2$ . The parameter  $\beta_0$  represents the fixed effect estimate of standard deviation of REE for control conditions and  $\beta_1$  the fixed effect estimate for the offset from this for the PCOS conditions (i.e., the difference between conditions). The estimated offset was allowed to vary across both labs and studies each reflected by  $\alpha_{1,\text{lab}[i]}$  and  $\alpha_{1,\text{study}[j]}$  respectively, and these effects were also modelled as correlated with the corresponding random intercepts with covariance  $\Sigma_{\text{lab}}$  and  $\Sigma_{\text{study}}$ , and  $\text{corr}_{\text{lab}}$  and  $\text{corr}_{\text{study}}$  correlation matrices. Finally,  $\beta_2$  represents the fixed effect of the natural logarithm of the corresponding mean REE estimate  $\tilde{m}$  which is modelled as estimated with measurement error i.e.,  $m$  represents the point estimate for the  $l$ th natural logarithm of the mean REE estimate from the  $k$ th arm, for the  $j$ th study, conducted by the  $i$ th lab and  $\sigma_{\log(y_{i,\text{mean},[ijkl]})}^2$  is the corresponding sampling error for that estimate.

The priors for this model were as follows:

$$\begin{aligned}
\beta_0 &\sim \text{Student-}t(3, 5.54, 0.80) \\
\beta_1 &\sim \text{Student-}t(3, 0, 5.3) \\
\beta_2 &\sim \text{Student-}t(3, 0, 2.5) \\
\text{mean}(\tilde{m}) &\sim \text{Half-student-}t(3, 7.28, 0.62) \\
\text{sd}(\tilde{m}) &\sim \text{Half-student-}t(3, 0, 5) \\
\sigma_{0,\text{study}}^2 &\sim \text{Half-student-}t(3, 1.08, 1.06) \\
\sigma_{0,\text{lab}}^2, \sigma_{1,\text{lab}}^2, \sigma_{1,\text{study}}^2, \sigma_{0,\text{arm}}^2, \sigma_{0,\text{effect}}^2 &\sim \text{Half-student-}t(3, 0, 2.5) \\
\text{corr}_{\text{lab}} &\sim \text{LKJ}(1) \\
\text{corr}_{\text{study}} &\sim \text{LKJ}(1)
\end{aligned}$$

where the prior for  $\beta_0$ , which corresponded to the model intercept and standard deviation of REE in the control condition, was again set based on meta-analysis of the standard deviation of REEs for women from two large studies of healthy people (Pavlidou et al., 2023; Velasquez, 2011) though set with a conservative degrees of freedom for the Student- $t$  distribution. The random intercept  $\sigma_{0,\text{study}}^2$  was set similarly to this. The prior for the fixed effect  $\beta_1$ , reflecting the difference between control and PCOS conditions was set based on a wide range of possible values. Given that in many cases of variables in the field there is an approximate relationship of  $\sim 1$  for the natural logarithm of the standard deviation conditioned upon the natural logarithm of the mean (Steele et al., 2023) we set this prior to reflect the range of differences on the on the log scale (i.e.,  $\log(1584)$ ). We then set a prior that permits values approximately across this range of values with the majority of it's mass centred around zero. The prior for the fixed effect  $\beta_2$ , reflecting the relationship between the natural logarithm of the mean REE with the natural logarithm of the standard deviation of REE, was set it to be weakly informative centred on zero with a wide scale to indicate uncertainty in this outcome specifically despite the typical relationship close to  $\sim 1$ . Priors for the measurement error of the natural logarithm of the mean REE estimate  $\tilde{m}$  were again based upon meta-analysis of the aforementioned studies, though measurement error has to be positive, as does the corresponding standard deviation of this error, so we set these to conservative wide half-student- $t$  distributions. The remaining random effects were set based on the default weakly regularising priors for **brms** and scaled to the expected response values using a half-student- $t$  distribution with 3 degrees of freedom and  $\mu = 0$ , and both correlation matrices  $\text{corr}_{\text{lab}}$  and  $\text{corr}_{\text{study}}$  were set with an LKJcorr(1) distribution.

## Post-processing of models

For both models we examined trace plots along with  $\hat{R}$  values to examine whether chains have converged, and posterior predictive checks for each model to understand the model implied distributions. From each model we took draws from the posterior distributions for the conditional absolute estimates for each condition (i.e., controls and PCOS) by study incorporating random effects, the global grand mean absolute estimates for each condition ignoring random effects, and the global grand mean between condition relative contrast for controls vs PCOS conditions ignoring random effects. The between condition relative contrast for controls vs PCOS conditions corresponded to  $\beta_1$  in each model and was our primary estimand of interest; for the mean REE model this corresponded to the absolute difference in mean REE, and for the standard deviation of REE model this corresponded to the natural logarithm of the ratio of standard deviations of REE which was exponentiated (note, all log standard deviation of REE model estimates were exponentiated back to the original scale to aid interpretability). We present the full probability density functions for posterior visually, and also to calculate mean and 95% quantile intervals (QI: i.e., ‘credible’ or ‘compatibility’ intervals) for each estimate providing the most probable value of the parameter in addition to the range from 2.5% to 97.5% percentiles given our priors and data.

## Sensitivity analyses

### Pairwise contrast based models

By way of pre-registered sensitivity analysis we also conducted pairwise contrast based models where we limited the included effects to those extracted from studies including only a directly comparable control and PCOS arm at baseline. These models were both as follows:

$$\begin{aligned}\hat{\theta}_{ij} &\sim \mathcal{N}(\mu_{ij}, \sigma_{ij}^2) \\ \mu_{ij} &= \beta_0 + \alpha_{0,\text{lab}[i]} + \alpha_{0,\text{study}[j]} \\ \alpha_{0,\text{lab}[i]} &\sim \mathcal{N}(0, \sigma_{0,\text{lab}}^2), \text{ for lab } i = 1, \dots, I \\ \alpha_{0,\text{study}[j]} &\sim \mathcal{N}(0, \sigma_{0,\text{study}}^2), \text{ for study } j = 1, \dots, J\end{aligned}$$

where  $\hat{\theta}_{ijkl}$  is the pairwise effect size, either the mean difference in REE or the log coefficient of variation ratio (calculated as PCOS vs control), for the  $j$ th study, conducted by the  $i$ th lab and  $\sigma_{ijkl}^2$  is the corresponding sampling error for that effect size estimate. The random intercepts for the  $i$ th lab,  $j$ th study are  $\alpha_{0,\text{lab}[i]}$  and  $\alpha_{0,\text{study}[j]}$  respectively each with standard deviation of  $\sigma_{0,\text{lab}[i]}^2$ ,  $\sigma_{0,\text{study}[j]}^2$ . The parameter  $\beta_0$  represents the fixed effect estimate of the pairwise effect size i.e., the pooled estimate of the contrast between conditions. The priors for  $\beta_0$  in these models were set as default weakly regularising which is set on an intercept that results when internally centering all population-level predictors around zero to improve sampling efficiency and scaled to the expected response values using a student- $t$  distribution;

for the mean difference in REE this was student- $t(3, 5.5, 50.2)$  and for the log coefficient of variance ratio this was student- $t(3, -0.1, 2.5)$ . The random effects were set similarly scaled to the expected response values but using a half-student- $t$  distribution centred on zero. From these models we calculated the mean and 95% quantile intervals (i.e., ‘credible’ or ‘compatibility’ intervals) for the  $\beta_0$  (comparable to the  $\beta_1$  from the corresponding mean and standard deviation of REE arm-based models) for each effect size providing the most probable value of the parameter in addition to the range from 2.5% to 97.5% percentiles given our priors and data.

### **Additional sensitivity analyses**

As noted in the section above, “*Studies with possible reporting errors*”, we opted to conduct analysis with and without the inclusion of four studies with possible reporting errors we could not resolve (Georgopoulos et al., 2009; Koika et al., 2009; Kritikou et al., 2006; Saltamavros et al., 2007). Thus the main models described above were run with and without these studies, the main results are presented without them and the sensitivity results with them are reported separately below.

As a further sensitivity analysis, given the inclusion of some studies with interventions in women with PCOS reporting REE at multiple timepoints such as mid- and post-intervention (Bruner et al., 2006; Moran et al., 2006; Pohlmeier et al., 2014) but lacking of control women without PCOS, we opted to also conduct sensitivity analysis excluding these and only examining their baseline results in addition to the cross-sectional studies. As such, the main models described above were run without the follow-up (i.e., mid- or post-intervention timepoints) from these studies and only using the baseline results in addition to other cross-sectional studies.

## **Results**

### **Systematic Review**

Our systematic review identified 13 studies from 12 lab/research groups including 24 arms (Control arms = 9, PCOS arms = 15) and a total of 918 participants (Controls: minimum  $n = 9$ , median  $n = 29$ , maximum  $n = 54$ ; PCOS: minimum  $n = 5$ , median  $n = 28$ , maximum  $n = 266$ ). Descriptive characteristics of the arms and participants in these studies are reported in Table 1. Note, these numbers exclude the four studies previously noted with possible reporting issues that were unresolved (they are however summarised in Table 1).

Table 1: Descriptive characteristics of arms and participants for included studies

Authors	Article title	Condition	Diagnostic criteria <sup>16</sup>	Sample size	Age <sup>1</sup>	Body mass <sup>17</sup>	Height <sup>17</sup>	BMI <sup>17</sup>	Fat mass <sup>2</sup>	Fat free mass <sup>2</sup>	Race/Ethnicity	Physical activity	Country of study	Metabolic health	Measurement Method <sup>18</sup>
Segal and Dunaif (1996)	Rising metabolic rate and hyperandrogenic symptoms in polycystic ovarian syndrome.	Control		11	28 (3.32)	63.5 (4.97)				48.8 (3.32)			USA	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		9	29 (6)	66.9 (17.4)				50.4 (7.5)			USA	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 126 (SD = 24), Glucose area (0-2 h) = 104 (SD = 24), Insulin area (0-2 h) = 104 (SD = 24)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		PCOS	PCOS was diagnosed by presence of two or more clinical and/or biochemical features in the presence of chronic hyperandrogenism or anovulation.	10	25 (6.32)	84.1 (9.54)				48.7 (3.79)			USA	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 122 (SD = 43), Glucose area (0-2 h) = 104 (SD = 24), Insulin area (0-2 h) = 104 (SD = 24)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Robinson et al., (1992)	Postprandial hypermetabolism is reduced in PCOS and is associated with hyperandrogenism.	PCOS	PCOS was defined by the presence of two or more clinical and/or biochemical features in the presence of chronic hyperandrogenism or anovulation.	14	29 (6.44)	76.05 (18.05)	160.86*	27.65 (6.44)		51.12 (4.69)	Matched between groups but not reported		UK	Short, Insulin tolerance (mg/dL), Glucose slope (mg/dL/min), Peak insulin (mU/L, 4 min) = 200 (SD = 40)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		14	30.25 (4.98)	76.62 (18.03)	165.97*	27.65 (6.45)		50.12 (7.06)	Matched between groups but not reported		UK	Short, Insulin tolerance (mg/dL), Glucose slope (mg/dL/min), Peak insulin (mU/L, 4 min) = 200 (SD = 40)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Brunner et al., (2006)	Effects of exercise and women with polycystic ovary syndrome.	PCOS	Rotterdam criteria	7	32.3 (2.05)	100.3 (17.73)	166.62*	36.2 (5.20)					Canada	Fasting insulin (mU/L) = 116.7 (SD = 42.2)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		PCOS	Rotterdam criteria	5	28.4 (6.04)	94.8 (23.86)	159.45*	37.1 (7.6)					Canada	Fasting insulin (mU/L) = 230.8 (SD = 77.4)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Moran et al., (2006)	Short-term (week) hypermetabolism followed by dietary intervention in polycystic ovary syndrome	PCOS	Rotterdam criteria	34	32.02 (5.17)	96 (19.24)	165.80*	34.9 (6.70)	34.9 (8.75)	61.5 (12.24)	White/Caucasian		Australia	Fasting glucose (mg/dL) = 83.3 (SD = 10.1), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Claus et al., (2008)	Rising metabolic rate and hyperandrogenic symptoms in polycystic ovary syndrome.	PCOS	Rotterdam criteria	31	25.9 (5.3)			26.97 (5.12)					Turkey	Fasting glucose (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		29	27.1 (4.8)			26.03 (5.60)					Turkey	Fasting glucose (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Graff et al., (2013)	Dietary glycemic index is associated with hyperandrogenism and hypermetabolism in polycystic ovary syndrome.	PCOS	Rotterdam criteria	61	22.7 (6.2)			26.9 (5.6)			White/Caucasian = 87.6%	Physical activity (step/day median) = 5017 (IQR = 3058-7002)	Brazil	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 94.4 (SD = 16.7), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		44	25 (5.6)			27.1 (5.7)			White/Caucasian = 87.6%	Physical activity (step/day median) = 5017 (IQR = 3058-7002)	Brazil	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 94.4 (SD = 16.7), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Palmstein et al., (2014)	Effect of a low-glycemic diet on weight, insulin sensitivity, and hypermetabolism in polycystic ovary syndrome.	PCOS	Rotterdam criteria	10	29.6 (4.6)	105.4 (14.5)	165.46*	36.5 (4.2)	52.4 (14.8)	52.3 (10.7)	White/Caucasian = 6, Hispanic = 3, Native American = 1	Physical activity level (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	USA	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 94.4 (SD = 16.7), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Doh et al., (2016)	The Relationship between Insulin Resistance, Hyperandrogenism, and Hypermetabolism in Polycystic Ovary Syndrome: A Cross-Sectional Study.	PCOS	Rotterdam criteria	6	26 (5.19)			34.1 (3.56)	41.2 (12.45)	56.3 (4.07)	African	Engaged in sporting activities < 2 days/week	Cameroon	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		PCOS	Rotterdam criteria	8	27 (3.71)			26.4 (2.97)	23.3 (6.80)	47.4 (4.56)	African	Engaged in sporting activities < 2 days/week	Cameroon	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		10	23 (0.74)			22.5 (3.63)	17.1 (7.56)	45.9 (6.67)	African	Engaged in sporting activities < 2 days/week	Cameroon	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Larsson et al., (2016)	Dietary intake, resting energy expenditure and hypermetabolism in women with polycystic ovary syndrome.	PCOS	Modified Rotterdam criteria	72	30.2 (4.4)	79.6 (20.3)	167.12*	26.5 (7.2)					Sweden	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		30	27.8 (3.6)	70.9 (17.1)	160.77*	24.6 (5)					Sweden	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Graff et al., (2017)	Insulin resistance is associated with hypermetabolism in women with polycystic ovary syndrome.	PCOS	Rotterdam criteria	84	23.5 (6.3)			29.4 (6.4)			White/Caucasian = 92.9%	Physical activity (step/day median) = 5017 (IQR = 3058-7002)	Brazil	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 94.4 (SD = 16.7), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		54	26.2 (6.5)			27.2 (5.8)			White/Caucasian = 88.9%	Physical activity (step/day median) = 5017 (IQR = 3058-7002)	Brazil	OGTT (The oral glucose load), Fasting glucose (mg/dL) = 94.4 (SD = 16.7), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Rodrigues et al., (2017)	Low validity of predictive equations for resting energy expenditure in women with polycystic ovary syndrome.	PCOS	Rotterdam criteria	30	30.8 (5.4)	85.3 (13.1)	161.76*	32.6 (3.7)					Brazil	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Brook et al., (2017)	Assessing Energy Expenditure in Women With Polycystic Ovary Syndrome: A Cross-Sectional Study.	PCOS	Modified Rotterdam criteria	28	26.6 (5)	104.1 (19.3)	161.92*	39.9 (6.3)	51.6 (13.4)	52.5 (7.5)	White/Caucasian = 96%, African American = 4%	Physical activity level (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	USA	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
Tut et al., (2018)	Rising energy expenditure in women with polycystic ovary syndrome.	PCOS	Rotterdam criteria	206	24.9 (5.2)			26.3 (7.4)	27.1 (14.4)	40 (7.7)			Italy	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA
		Control		51	25.2 (5.6)			20.5 (2)	15.6 (5.4)	42.8 (4.6)			Italy	Hypermetabolism (mg/dL) = 102 (SD = 21), 2h OGTT = 145 (SD = 21), OGTT % glucose area (0-2 h) = 64 (SD = 12), 0-2h insulin area (0-2 h) = 104 (SD = 9)	Indiana Calorimetry; Metabolic Unit, Indiana University, Indianapolis, USA

PCOS = polycystic ovary syndrome; BMI = body mass index; OGTT = oral glucose tolerance test; HOMA-IR = homeostatic model assessment of insulin resistance

<sup>1</sup>Values are Mean (SD) unless otherwise specified; note, some have been calculated/estimated from corresponding standard error, range, IQR, median, and sample size (see data and code)<sup>17</sup>Indicates that this mean was estimated from the corresponding means for body mass/height/BMI

## Mean REE Model Results

The main model for mean REE resulted in a posterior distribution for the contrast between control and PCOS conditions with a mean point estimate of 30 kcal/day with a 95% quantile interval ranging from -47 kcal/day to 113 kcal/day suggesting there is a 95% probability that the true difference lies between these values given our priors and the data from included studies. The corresponding conditional estimates for the control condition and PCOS condition respectively were 1442 kcal/day [95%QI:1334 kcal/day to 1553 kcal/day] and 1472 kcal/day [95%QI:1359 kcal/day to 1587 kcal/day]. These results including the full visualisation of the posterior distribution, in addition to the conditional estimates by study, can be seen in Figure 1.

### Mean Resting Energy Expenditure

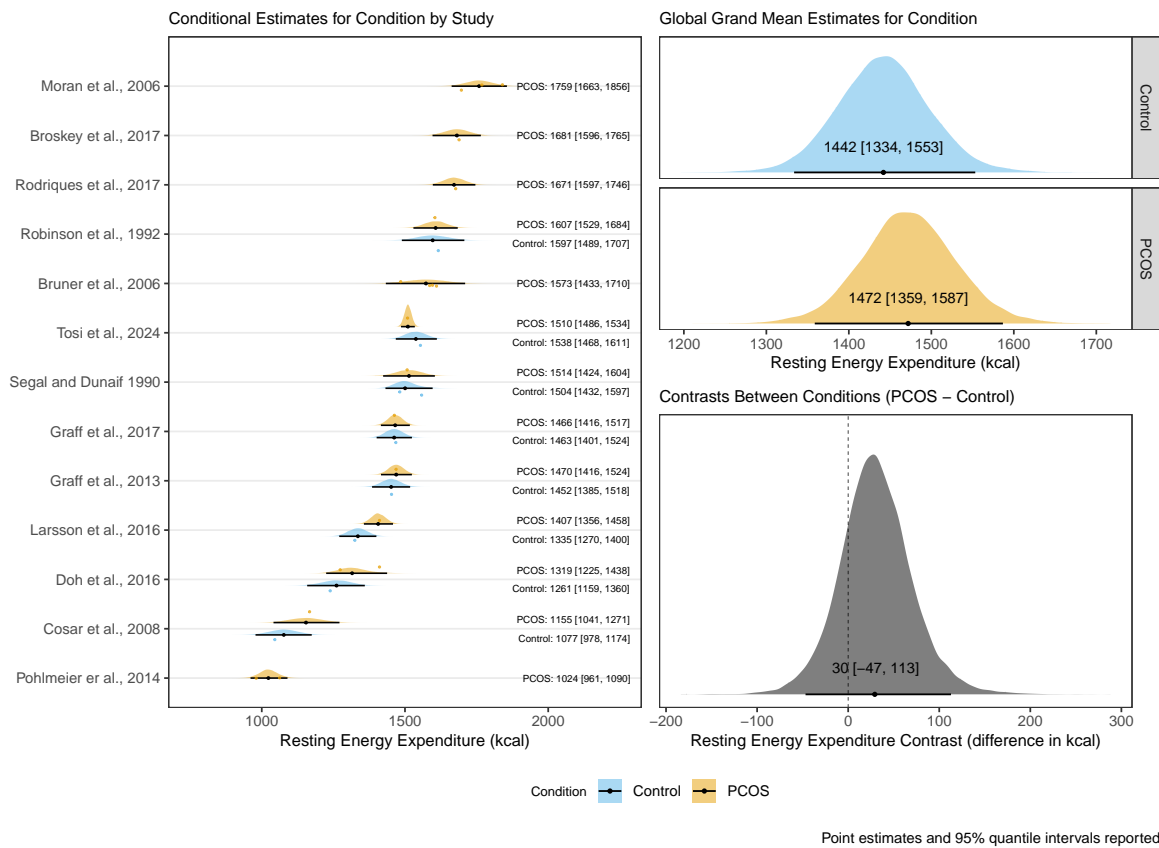


Figure 1: Posterior distribution, mean point estimates, and 95% quantile intervals for conditional estimates by study, global grand mean estimates by condition, and the contrast between conditions for mean resting energy expenditure of control women without PCOS and women with PCOS.

## Standard Deviation of REE Model Results

The main model for the between participant standard deviation of REE resulted in a posterior distribution for the contrast ratio between control and PCOS conditions with a mean point estimate of 0.98 with a 95% quantile interval ranging from 0.71 to 1.33 suggesting there is a 95% probability that the true ratio of standard deviations lies between these values given our priors and the data from included studies. The corresponding conditional estimates for the standard deviations of the control condition and PCOS condition respectively were 238 kcal/day [95%QI:178 kcal/day to 312 kcal/day] and 229 kcal/day [95%QI:169 kcal/day to 303 kcal/day]. These results including the full visualisation of the posterior distribution, in addition to the conditional estimates by study, can be seen in Figure 2.

### Standard Deviation of Resting Energy Expenditure

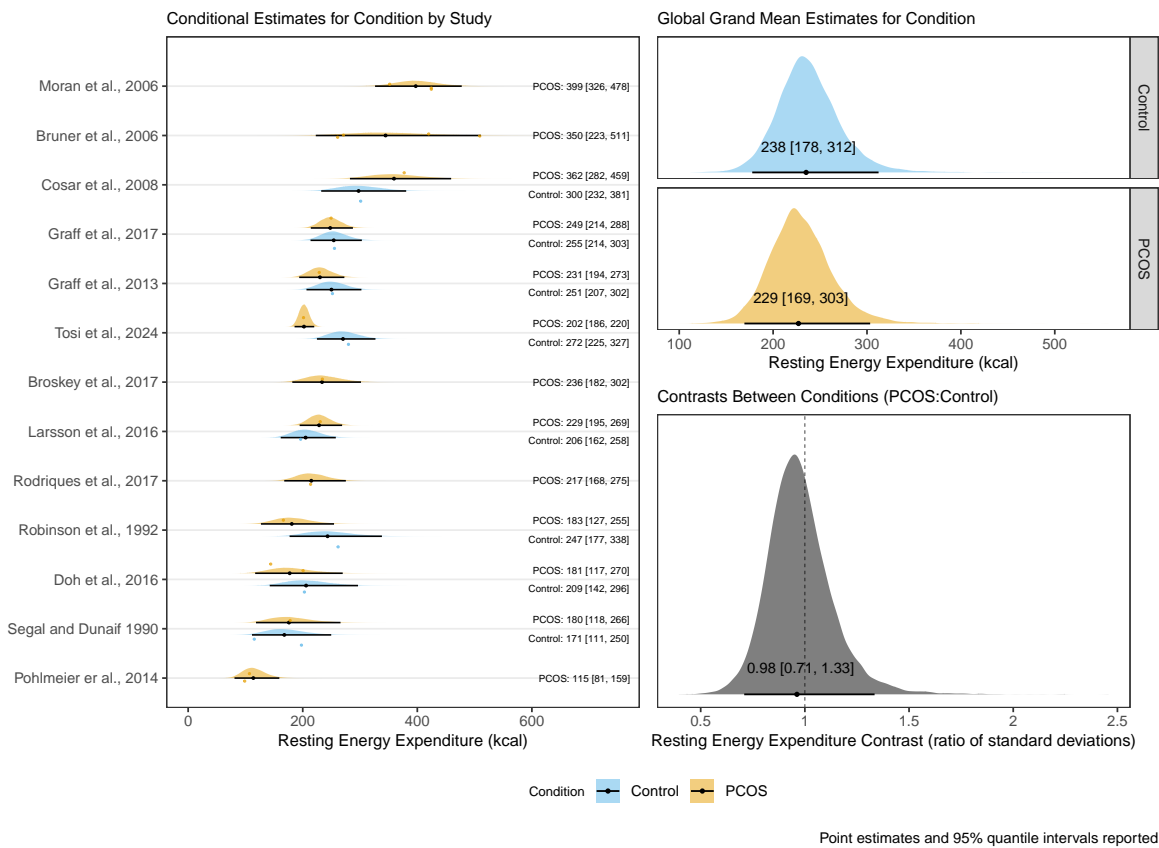


Figure 2: Posterior distribution, variance point estimates, and 95% quantile intervals for conditional estimates by study, global grand variance estimates by condition, and the contrast ratio between conditions for the standard deviation of resting energy expenditure of control women without PCOS and women with PCOS.

## **Sensitivity Analyses**

### **Pairwise Models**

For mean REE the pairwise model resulted in qualitatively similar inferences suggesting little difference between control and PCOS conditions with mean point estimate of 16 kcal/day with a 95% quantile interval ranging from -36 kcal/day to 70 kcal/day. This was similar for the standard deviation of REE with the pairwise model resulting in a contrast ratio between control and PCOS conditions with a mean point estimate of 0.9 with a 95% quantile interval ranging from 0.6 to 1.35.

### **Models including studies with possible reporting issues**

For mean REE the model including the studies noted with possible reporting issues that were not resolved (Georgopoulos et al., 2009; Koika et al., 2009; Kritikou et al., 2006; Saltamavros et al., 2007) still resulted in qualitatively similar inferences suggesting little difference between control and PCOS conditions with mean point estimate of 15 kcal/day with a 95% quantile interval ranging from -69 kcal/day to 95 kcal/day. This was similar for the standard deviation of REE with the model including these studies resulting in a contrast ratio between control and PCOS conditions with a mean point estimate of 1.21 with a 95% quantile interval ranging from 0.68 to 2.13.

### **Baseline REE Measurement Models**

For mean REE the model which included only baseline REE measurements from studies involving interventions (Bruner et al., 2006; Moran et al., 2006; Pohlmeier et al., 2014), in addition to other cross-sectional studies, also resulted in qualitatively similar inferences suggesting little difference between control and PCOS conditions with mean point estimate of 32 kcal/day with a 95% quantile interval ranging from -46 kcal/day to 118 kcal/day. This was the case also for the standard deviation of REE with this model resulting in a contrast ratio between control and PCOS conditions with a mean point estimate of 0.97 with a 95% quantile interval ranging from 0.71 to 1.36.

## **Discussion**

TO WRITE



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